

Laboratório Nacional de Computação Científica
Programa de Pós-Graduação em Modelagem Computacional

Generalized Lambda Distribution for Uncertainty Quantification of Large-scale Spatio-temporal Models

Noel Moreno Lemus

Petrópolis, RJ - Brasil

Abril de 2018

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Thesis submitted to the examining committee
in partial fulfillment of the requirements for
the degree of Doctor of Sciences in Computa-
tional Modeling.

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Approved by:

Prof. Fábio André Machado Porto,
D.Sc.
(Presidente)

Prof. Fernando Alves Rochinha, D.Sc.

Prof. Hugo de La Cruz, Ph.D.

Prof. Artur Ziviani, Ph.D.

Petrópolis, RJ - Brasil
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Dedication

To my little and special family.

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O autor manifesta reconhecimentos às pessoas e instituições que colaboraram para a execução de seu trabalho.

“Essentially, all models are wrong, but some are useful.”
(George Edward Pelham)

Abstract

Segundo a ??, 3.1-3.2), o resumo deve ressaltar o objetivo, o método, os resultados e as conclusões do documento. A ordem e a extensão destes itens dependem do tipo de resumo (informativo ou indicativo) e do tratamento que cada item recebe no documento original. O resumo deve ser precedido da referência do documento, com exceção do resumo inserido no próprio documento. (...) As palavras-chave devem figurar logo abaixo do resumo, antecedidas da expressão Palavras-chave:, separadas entre si por ponto e finalizadas também por ponto.

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Abstract

Large-scale spatio-temporal simulations with quantified uncertainty enable scientists/decision-makers to precisely assess the degree of confidence of their simulation-based predictions. This uncertainty could be quantified or characterized in different ways, from the use of low order statistical moments (the most commonly used), to the evaluation of a complete PDF (a most complete approach). The latter provides a more comprehensive description of the uncertainty leading to aware decisions. However, fitting PDFs to the data is computational intensive. Moreover, due to heterogeneity the uncertainty computed in regions of the dataset is hampered by the existence of different PDF types.

In this thesis, we propose a new method to quantify the uncertainty in large-scale spatio-temporal models based on the Generalized Lambda Distribution (GLD). GLD is a family of PDFs that nicely models the heterogeneity of uncertainty as discussed above. It is specified by 4 parameters that simplifies PDFs comparisons easing analytical processing, such as clustering. We show how the dataset modeled through GLDs can be used to answer queries, such as: *(i)* how to group the output of the UQ process based on the similarity of the uncertainty?, *(ii)* what is the uncertainty in some spatio-temporal locations not previously analysed?, *(iii)* what is the uncertainty of an specific spatio-temporal region?, *(iv)* how to compare two regions as a function of its uncertainty?, and *(v)* what is the less uncertain model from a set of models? The proposed method has been tested in realistic use cases from various scientific areas. Additionally, an R package has been implemented with all the functionalities discussed in the thesis.

Keywords: Uncertainty Quantification, Large-scale spatio-temporal models, Big Data, Generalized Lambda Distribution

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List of abbreviations and acronyms

UQ	Uncertainty Quantification
FP	Forward Problem
<i>QoI</i>	Quantity of Interest
<i>GLD</i>	Generalized Lambda Distribution
p.f.	percentile function
r.v.s	random variables
GLDEX	r package to compute the GLD

List of symbols

Γ	Letra grega Gama
Λ	Lambda
ζ	Letra grega minúscula zeta
\in	Pertence

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1 Introduction

The rapid growth of high-performance computing and the advances in numerical techniques in the last two decades have provided an unprecedented opportunity to explore complex physical phenomena using large-scale spatio-temporal modeling and simulation. At the same time, scientific community is leaving behind the traditional deterministic approach, which offers point predictions with no associated uncertainty (JOHNSTONE et al., 2016); to include Uncertainty Quantification (UQ) as a common practice in their researches.

Large-scale spatio-temporal simulations with quantified uncertainty enable scientists to make precise statements about the degree of confidence they have in their simulation-based predictions. These approaches find practical applicability in models for predicting the behavior of weather, hurricane forecasts (TOBERGTE; CURTIS, 2013), subsurface hydrology (BARONI; TARANTOLA, 2014), geology (GUERRA et al., 2016), nuclear reactor design, financial portfolios (CHEN; FLOOD; SOWERS, 2008), and biological phenomena, just to name a few. They also allow to study physical phenomena that are impossible to assess experimentally, for example: simulate nuclear accidents, or the conditions that some spatial vehicle will find at landing in Mars, and so on. The success of these techniques has made them increasingly important tools for high impact predictions and decision making.

UQ includes different aspects that warranty the predictive fidelity of a numerical simulation, such as the uncertainty in the experimental data, which is used for defining the parameter values of a model; the propagation of uncertain parameters through the model; and the choice of the model itself. UQ is a complex process that covers the following main tasks: (i) uncertainty characterization (CRESPO; KENNY; GIESY, 2014), also called model calibration (FARRELL, 2015) or statistical inverse problem (ESTACIO-HIROMS; PRUDENCIO, 2012); (ii) sensitivity analysis; (iii) forward problem or uncertainty propagation; and (iv) model selection.

This paper is focused on *forward propagation*, whose objective is to quantify the uncertainties in model output(s) propagated from uncertain inputs. The targets of *forward propagation* analysis can be: (i) evaluate low-order moments (i.e. mean and variance) of the outputs, (ii) evaluate the reliability of the outputs, and/or (iii) assess the complete probability distribution (PDF) of the outputs.

When dealing with large-scale spatio-temporal models, a huge amount of data is generated as a result of the simulation process. Indeed, on each spatio-temporal location $(s_i, t_j) \in \mathcal{S} \times \mathcal{T} \subseteq \mathbb{R}^3 \times \mathbb{R}$, usually more than 10^4 simulations are performed. Then, the size

of the output dataset is in the order of $N_s \times N_t \times N_{sim}$, where: N_s is the number of spatial locations, N_t is the number of time steps, and N_{sim} is the number of simulations. An example of the volume of data generated by these simulations is given in the experimental section ?? of this paper, where the output dataset is about 2.4 TB. This turn *forward propagation* in a data intensive problem.

Another important aspect, which is often not taken into account, is that the uncertainty need to be quantified in some way that can be used after, to answer questions that arise in the *UQ* context. In that sense, assess the complete *PDF* could be the best way to quantify uncertainty, because if you can find the *PDF* that best fit the dataset with reasonably accurately, you can get all the statistical properties under one roof. At the same time, we can substitute the original data by the *PDFs*, which represents a huge reduction in the volume of data to manipulate.

Contradictorily, statistical moments (e.g. mean and standard deviation) are possibly the most used ways to quantify the uncertainty, despite the fact that they doesn't have information about the manner in which the data are distributed (LAMPASI; Di Nicola; PODESTA, 2006). This is because of the difficulty to find the *PDF* that best fit a dataset (KARIAN; DUDEWICZ, 2011), even more, when dealing with large-scale spatio-temporal models where the *PDF* needs to be derived on each spatio-temporal location, and therefore the *forward propagation* problem becomes time consuming and computationally intensive too.

However, the use of low order moments alone prevents us from making accurate analysis with respect to the uncertainty. They are not enough neither for the characterization nor for the quantification of the uncertainty, and questions such as:

- What is the uncertainty in the spatio-temporal region $\mathcal{S}_i \times \mathcal{T}_j$ associated to the *QoI* q_k and a computational model \mathcal{M}_m ?
- How to compare different spatio-temporal regions $\mathcal{S}_i \times \mathcal{T}_j$ with respect to the uncertainty?
- What is the less uncertain model from the set of models $\mathcal{M} = \mathcal{M}_1, \mathcal{M}_2, \dots, \mathcal{M}_m$, to predict the value of a *QoI* q_k , over a spatio-temporal region $\mathcal{S}_i \times \mathcal{T}_j$?

can be poorly answered. So, we emphasize that only the characterization of the uncertainty by using the *PDF* allows aware decisions.

A first effort to try to estimate the *PDFs* on large-scale spatio-temporal simulations was done by (LIU et al.,) Ji et. al. in ***Parallel Computation of PDFs on Big Spatial Data Using Spark***. They propose a new solution to efficiently compute the *PDFs* in parallel using Spark, through three methods: data grouping, machine learning

prediction and sampling. The main drawback of the proposed approach is that you should try many different distributions, to find the PDF that best fits the dataset on each specific spatio-temporal location. Another drawback is that, as we mentioned above, the uncertainty needs to be quantified in the way that facilitates its further use; and the heterogeneity of the functions used in the approach doesn't facilitate it.

To face these challenges, in this paper we propose a general framework to quantify the uncertainty in large-scale spatio-temporal models. It uses a data-driven approach and combines the generalized lambda distribution (*GLD*), clusters algorithms and information entropy, for helping researchers to answer the above questions and many others that arise in *UQ* context. Our proposal provides a generally applicable and easy-to-use tool that supports the representation and analysis of uncertainty, as was suggested in the "*Workshop on Quantification, Communication, and Interpretation of Uncertainty in Simulation and Data Science*" (TOBERGTE; CURTIS, 2013).

In order to illustrate the use of the proposed framework, a case study is discussed. The main results obtained are: (i) the *GLD* good fits for more than the 80 % of the dataset, (ii) the use of the *GLD* allows to include clustering algorithms to group the spatio-temporal locations with similar uncertainty, (iii) the centroids of the clusters can be used as a faithful representation of the rest of the spatio-temporal locations, which significantly reduces the data corresponding to the simulation outputs, (iv) with the use of these centroids we can characterize the uncertainty in any spatio-temporal region as a mixture of *GLDs*.

The rest of the paper is organized as follows: Section ?? gives the theoretical foundations of *UQ* and highlights some interesting aspects included in our proposal. Section ?? describes the principal characteristics of the *GLD* that make it suitable for this proposal. Section ?? presents the proposed approach, the workflow we implement and some considerations of the implementation. Section ?? presents a use case and discusses the results. This use case allows us to explain our approach in the context of a real problem, which facilitates its understanding. Section ?? covers the related works and finally, section 9 concludes the paper and proposes some future works.

1.1 Research Objectives

The main objective of this thesis is a new method to quantify the uncertainty in large-scale spatio-temporal models based on the Generalized Lambda Distribution (*GLD*).

To achieve that goal the following research questions need to be answered:

RQ1. how to group the output of the UQ process based on the similarity of the uncertainty?

RQ2. what is the uncertainty in some spatio-temporal locations not previously analysed?

RQ3. what is the uncertainty of an specific spatio-temporal region?

RQ4. how to compare two regions as a function of its uncertainty?

RQ5. what is the less uncertain model from a set of models?

RQ1 and **RQ2** are answered in chapter 5, while **RQ3**, **RQ4** and **RQ5** are answered in chapter 7. In chapter 8 all the questions are answered again for all the use cases.

1.2 Highlights of the Dissertation

1.3 Organization of the Dissertation

The structure of the remainder of this thesis is outlined for reference.

Chapter 2 background of UQ.

Chapter 3 Ji paper.

Chapter 4 GLD.

Chapter 5 GLD clustering and kriging.

Chapter 6 Workflow.

Chapter 7 Use cases.

Chapter 8 Conclusions and future works.

2 Uncertainty Quantification Background

“UQ cannot tell you that your model is ‘right’ or ‘true’, but only that, if you accept the validity of the model (to some quantified degree), then you must logically accept the validity of certain conclusions (to some quantified degree)”
(Sullivan, 2015)

In this chapter, we summarize some definitions in UQ context that are important to understand the rest of the document. Also, different ways of representation of the uncertainty are discussed, with a brief justification of those we are going to use in the thesis. A general workflow of the UQ process is presented where we contextualize our contributions. A detailed discussion about forward propagation is presented. And finally, we summarize the chapter.

Hablar de todos los tipos de incertaza, de incerteza en datos, en modelos, parametrica, hasta llegar a lo que a nosotros nos interesa.

2.1 Definitions

2.1.1 Errors vs Uncertainties

2.1.2 Aleatoric vs Epistemic Uncertainty

It is sometimes assumed that uncertainty can be classified into two categories, aleatoric and epistemic, (KIUREGHIAN; DITLEVSEN, 2009) although the validity of this categorization is open to debate.

Aleatoric uncertainty arises from an inherent randomness in the properties or behavior of the system under study. For example, the weather conditions at the time of a reactor accident are inherently random with respect to our ability to predict the future. Other examples include the variability in the properties of a population of weapon components and the variability in the possible future environmental conditions that a weapon component could be exposed to. Alternative designations for aleatory uncertainty include variability, stochastic, irreducible and type A. (HELTON, 2009)

Epistemic uncertainty derives from a lack of knowledge about the appropriate value to use for a quantity that is assumed to have a fixed value in the context of a particular analysis. For example, the pressure at which a given reactor containment would fail for a specified set of pressurization conditions is fixed but not amenable to being unambiguously

defined. Other examples include minimum voltage required for the operation of a system and the maximum temperature that a system can withstand before failing. Alternative designations for epistemic uncertainty include state of knowledge, subjective, reducible and type B. (HELTON, 2009)

2.2 Uncertainty Representation

An immediate challenge in the development of an appropriate treatment of uncertainty is the selection of a mathematical structure to be used in its representation (HELTON et al., 2010). Traditionally, probability theory has provided this structure [48-55]. However, in the last several decades, additional mathematical structures for the representation of uncertainty such as evidence theory [56-63], possibility theory [64- 70], fuzzy set theory [71-75], and interval analysis [76-81] have been introduced. This introduction has been accompanied by a lively discussion of the strengths and weaknesses of the various mathematical structures for the representation of uncertainty [82-90]. For perspective, several comparative discussions of these different approaches to the representation of uncertainty are available [72; 91-98]

This section briefly summarizes some of this approaches, and discuss in more details probability theory as this is the main one used in the rest of the thesis.

2.2.1 Interval Analysis

2.2.2 Variance

2.2.3 Information Entropy

The concept of information entropy was first defined by Shannon (1948) in a study performed to identify the amount of information required to transmit English text. The underlying idea was that, given the probabilities of letters occurring in the English alphabet, it is possible to derive a measure describing the missing information to determine the full text of a partially transmitted message, where information is understood as the information required to identify the message, not the information of the message itself. Based on several theoretical considerations, Shannon derived the following equation to classify a measure of the missing information, often referred to as information entropy:

$$H = - \sum_i^N p_i \log p_i \quad (2.1)$$

The information entropy H is defined as the sum of the product of the probability p for each possible outcome i of N , total possible outcomes, with its logarithm. The minimum value is 0, because $\log 1 = 0$.

2.2.3.1 Information entropy in a spatio-temporal context

For each spatio-temporal region, the information entropy can be described as:

$$H(s, t) = - \sum_{m=1}^M p_m(s, t) \log p_m(s, t) \quad (2.2)$$

where s denotes the location of the subregion, M represents the number of possible (exclusive) members the subregion may contain, and t is the physical time.

2.2.3.2 Information entropy as a measure of uncertainty

Based on 2.2.3 and 2.2.3.1, if the possible outcomes of the model and the probability of each outcome on each (s, t) , are known, then the information entropy could be used as a qualitative measure of the uncertainty of the model output. For example, in a spatio-temporal region (s, t) where the outcome is always the same, the information entropy is 0, because the outcome is known. On the other hand, in the worse case where all the outcomes have the same probability in (s, t) , the entropy is maximum and the uncertainty too.

2.2.4 Probability Theory

2.3 Methods for Uncertainty Propagation

2.3.1 Sampling Methods

2.3.1.1 Monte Carlo

Monte Carlo simulations (MCS) provide the most robust and straightforward way to solve PDEs with random coefficients. In the case of (22.2), for instance, they consist of (i) generating multiple realizations of the input parameters a and b , (ii) solving deterministic PDEs for each realization, and (iii) evaluating ensemble statistics or PDFs of these solutions. MCS do not impose limitations on statistical properties of input parameters, entail no modifications of existing deterministic solvers, and are ideal for parallel computing (HIGDON, 2017).

2.4 Software and Tools for UQ

Currently, advances in uncertainty propagation and assessment have been paralleled by a growing number of software tools for uncertainty analysis, but none has gained recognition for a universal applicability, including case studies with spatial models and spatial model inputs. (K. Sawicka; SOIL, 2016)

These include both free software, like OpenTURNS (Andrianov et al., 2007), DACOTA (Adams et al., 2009) and DUE (Brown and Heuvelink, 2007), commercial, like COSSAN (Schuëller and Pradlwarter, 2006), or free, but written for a licenced software, e.g. SAFE (Pianosi et al., 2015) or UQLab (Marelli and Sudret, 2014) toolboxes for MATLAB. A broad review of existing software packages is available in Bastin et al. (2013). To the best of our knowledge, however, none of the existent software is specifically designed to be extended by the environmental science community. The use of powerful but complex languages like C++ (e.g. Dakota), Python (e.g. OpenTURNS) or Java (e.g. DUE) often discourages relevant portions of the non-highly-IT trained scientific community from the adoption of otherwise powerful tools. spup-R package ([K. Sawicka; SOIL, 2016](#)). De aqui saque lo de arriba tambien, aunque lo de arriba lo puedo buscar en sus respectivos papers y hablar un poco de cada uno de ellos.

2.5 Summary

2.6 Concepts

high-dimensional parameter spaces computationally demanding forward models nonlinearity and/or complexity in the forward model

2.7 Ideas a usar

HPC and computational modeling play a dominant role in shaping the methodological developments and research in uncertainty qualification. Depending on the complexity of the uncertainty qualification investigation, anywhere from 10^2 to 10^8 runs of the computational model may be required. Thus, uncertainty qualification investigations may require extreme-computing environments (e.g., exascale) to obtain results in a useful time frame, even if a single run of the computational model does not require such resources.

Advances in computing over the past few decades—both in availability and power—have led to an explosion in computational models available for simulating a wide variety of complex physical (and social) systems. These complex models—which may involve millions of lines of code, and require extreme-computing resources—have led to numerous scientific discoveries and advances. This is because these models allow simulation of physical processes in environments and conditions that are difficult or even impossible to access experimentally. However, scientists’ abilities to quantify uncertainties in these model-based predictions lag well behind their abilities to produce these computational models. This is largely because such simulation-based scientific investigations present a set of challenges that is not present in traditional investigations.

Until recently, the original approach of describing model parameters using single values has been retained, and consequently the majority of mathematical models in use today provide point predictions, with no associated uncertainty. (JOHNSTONE et al., 2016)

a 'typical' UQ problem involves one or more mathematical models for a process of interest, subject to some uncertainty about the correct form of, or parameter values for, those models.

Often, though not always, these uncertainties are treated probabilistically.

but how will you actually go about evaluating that expected value when it is an integral over a million-dimensional parameter space? Practical problems from engineering and the sciences can easily have models with millions or billions of inputs (degrees of freedom).

the language of probability theory is a powerful tool in describing uncertainty

UQ cannot tell you that your model is 'right' or 'true', but only that, if you accept the validity of the model (to some quantified degree), then you must logically accept the validity of certain conclusions (to some quantified degree). (SULLIVAN, 2015)

"UQ studies all sources of error and uncertainty, including the following: systematic and stochastic measurement error; ignorance; limitations of theoretical models; limitations of numerical representations of those models; limitations of the accuracy and reliability of computations, approximations, and algorithms; and human error. A more precise definition is UQ is the end-to-end study of the reliability of scientific inferences."

UQ is not a mature field like linear algebra or single-variable complex analysis, with stately textbooks containing well-polished presentations of classical theorems bearing August names like Cauchy, Gauss and Hamilton. Both because of its youth as a field and its very close engagement with applications, UQ is much more about problems, methods and 'good enough for the job'. There are some very elegant approaches within UQ, but as yet no single, general, over-arching theory of UQ.

In

Probability theorists usually denote the sample space of a probability space by Ω ; PDE theorists often use the same letter to denote a domain in \mathbb{R}^n on which a partial differential equation is to be solved. In UQ, where the worlds of probability and PDE theory often collide, the possibility of confusion is clear. Therefore, this book will tend to use Θ for a probability space and \mathbf{X} for a more general measurable space, which may happen to be the spatial domain for some PDE.

3 Parallel Computation of PDFs on Large-scale Spatio-temporal Models

3.1 Introduction

3.2 Architecture for Computing PDFs in Spark

3.3 Experimental Evaluation

4 The Generalized Lambda Distribution

“There are good reasons for using the GLD distribution methods... GLD fits have been used successfully in many fields ... Try the GLD first and stop there if the results are acceptable.”
(Karian and Dudewicz, 2011)

Fitting statistical distribution to data (real or simulated), is an important task in uncertainty quantification forward problem. When fitting data, one typically first selects a general class, or family, of distributions and then finds values for the distributional parameters that best match the observed data (LAKHANY; MAUSSER, 2000). One of this families is the Generalized Lambda Distribution (*GLD*), originally proposed by Ramberg and Schmeiser in 1974, as a generalization of the Tukey’s distribution (1947). The *GLD* has tested

In this chapter a review of the principal characteristics of the *GLD*

it has been used in Monte Carlo simulations (Hogben, 1963), the modeling of empirical distributions (Okur, 1988; Ramberg et al., 1979), in response time modeling (Au-Yeung et al., 2004; Kumaran and Achary, 1996), supply chain (Ganesalingam and Kumar, 2001) and in the sensitivity analysis of robust statistical methods (Shapiro et al., 1968).

4.1 The Generalized Lambda Distribution

The generalized lambda distribution is a continuous distribution defined in terms of its quantile function. Various parameterizations exist (see Section 4.1.3), but the most popular are the proposed by Ramberg and Schmeiser in 1974, Section 4.1.1; and the proposed by Freimer et al. in 1988, Section 4.1.2.

4.1.1 The Ramberg and Schmeiser Parameterization

The Generalized Lambda Distribution (*GLD*) was proposed by Ramberg and Schmeiser in 1974 as an extension of the Tukey’s distribution. It is represented by the quantile function:

$$Q_{RS}(y|\lambda) = Q_{RS}(y|\lambda_1, \lambda_2, \lambda_3, \lambda_4) = \lambda_1 + \frac{y^{\lambda_3} - (1 - y)^{\lambda_4}}{\lambda_2} \quad (4.1)$$

Table 1 – Support regions of the GLD and conditions on the parameters given by the RS parameterization to define a valid distribution function (KARIAN; DUDEWICZ, 2011). The support regions are displayed in Fig. 1. Note that there are no conditions on λ_1 to obtain a valid distribution.

Region	λ_2	λ_3	λ_4	$Q(0)$	$Q(1)$
1	< 0	≤ -1 $-1 < \lambda_3 < 0$ $\frac{(1-\lambda_3)^{1-\lambda_3}(\lambda_4-1)^{\lambda_4-1}}{(\lambda_4-\lambda_3)^{\lambda_4-\lambda_3}} = \frac{-\lambda_3}{\lambda_4}$	≥ 1 > 1	$-\infty$	$\lambda_1 + \frac{1}{\lambda_2}$
2	< 0	≥ 1 > 1 $\frac{(1-\lambda_4)^{1-\lambda_4}(\lambda_3-1)^{\lambda_3-1}}{(\lambda_3-\lambda_4)^{\lambda_3-\lambda_4}} = \frac{-\lambda_4}{\lambda_3}$	≤ -1 $-1 < \lambda_4 < 0$	$\lambda_1 - \frac{1}{\lambda_2}$	∞
3	> 0	> 0 $= 0$ > 0	> 0 > 0 $= 0$	$\lambda_1 - \frac{1}{\lambda_2}$ λ_1 $\lambda_1 - \frac{1}{\lambda_2}$	$\lambda_1 + \frac{1}{\lambda_2}$ $\lambda_1 + \frac{1}{\lambda_2}$ λ_1
4	< 0	< 0 $= 0$ < 0	< 0 < 0 $= 0$	$-\infty$ λ_1 $-\infty$	∞ ∞ λ_1

where $Q_{RS} = F^{-1}$ is the quantile function for probabilities y , with $y \in [0, 1]$; λ_1 and λ_2 are the location and scale parameters, and λ_3 and λ_4 determine the skewness and kurtosis of the $GLD(\lambda_1, \lambda_2, \lambda_3, \lambda_4)$.

The probability density function of the GLD at the point $x = Q_{RS}(y)$ is given by:

$$f(x) = f(Q_{RS}(y)) = \frac{\lambda_2}{\lambda_3 y^{\lambda_3-1} + \lambda_4 (1-y)^{\lambda_4-1}} \quad (4.2)$$

In order to have a valid distribution function, the probability density function $f(x)$ need to be positive for all x and integrates to one over the allowed domain:

$$f(x) \geq 0 \quad (4.3)$$

$$\int f(x) dx = 1 \quad (4.4)$$

This impose complex constraints on the parameters and support regions of the RS parameterization, as summarized in table 1 and figure 1.

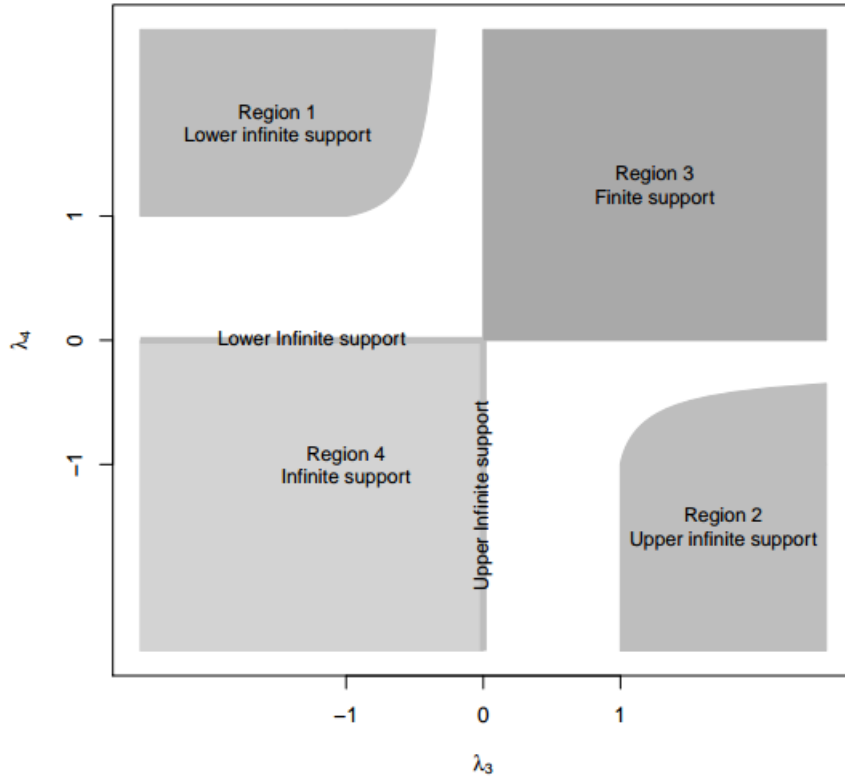


Figure 1 – Support regions of the GLD in the RS parameterization that produce valid statistical distributions.

4.1.2 The FMKL Parameterization

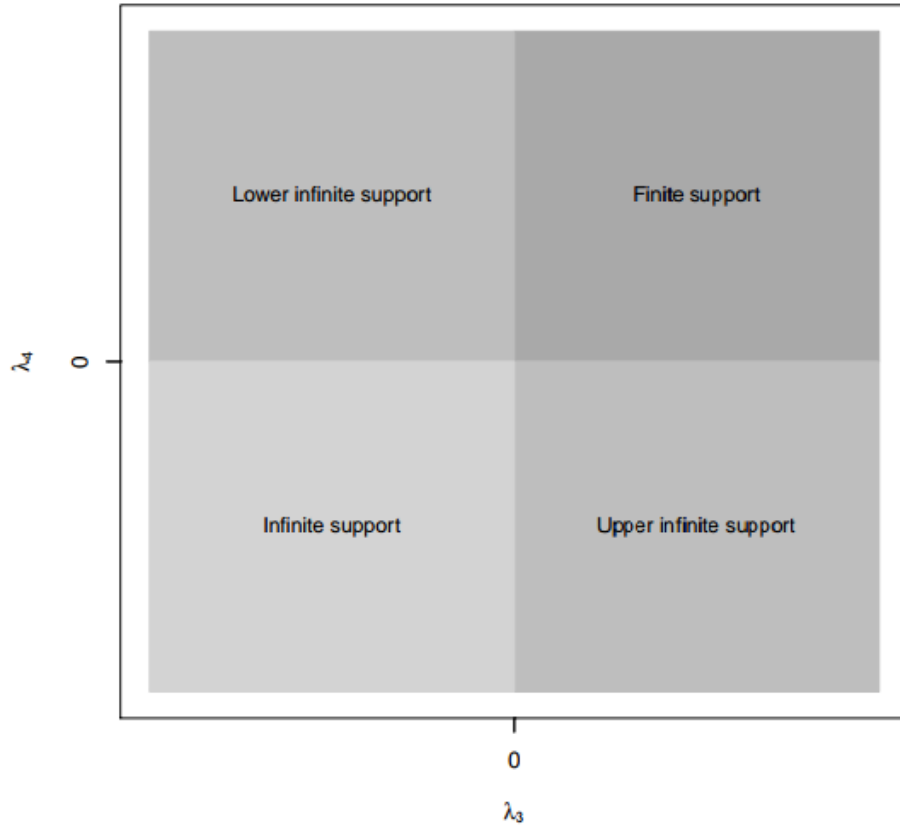
To circumvent the constraints on the *RS* parameter values, Freimer et al. ([FREIMER; LIN; MUDHOLKAR, 1988](#)) introduced a new parameterization called *FKML*, equation 4.5.

$$Q_{FMKL}(y|\lambda_1, \lambda_2, \lambda_3, \lambda_4) = \lambda_1 + \frac{1}{\lambda_2} \left[\frac{y^{\lambda_3} - 1}{\lambda_3} - \frac{(1 - y)^{\lambda_4} - 1}{\lambda_4} \right] \quad (4.5)$$

As in the previous parameterization, λ_1 and λ_2 are the location and scale parameters, but in this one λ_3 and λ_4 are the tail index parameters. The advantage over the previous parameterization is that the only constraint on the parameters is that λ_2 must be positive. Figure 2 displays the support regions of the *GLD* in the *FKML* parameterization, table 2.

Table 2 – Support regions of the *GLD* given by the *FMKL* parameterization (MARCONDES; PEIXOTO; MAIA, 2017).

λ_3	λ_4	$Q(0)$	$Q(1)$
> 0	> 0	$\lambda_1 - \frac{1}{\lambda_2 \lambda_3}$	$\lambda_1 + \frac{1}{\lambda_2 \lambda_4}$
> 0	≤ 0	$\lambda_1 - \frac{1}{\lambda_2 \lambda_3}$	∞
≤ 0	> 0	$-\infty$	$\lambda_1 + \frac{1}{\lambda_2 \lambda_4}$
≤ 0	≤ 0	$-\infty$	∞

Figure 2 – Support regions of the *GLD* in the *FMKL* parameterization that produce valid statistical distributions.

The probability density function of the *FMKL-GLD* at the point $x = Q_{FMKL}(y)$ is given by (SU, 2015):

$$f(x) = f(Q_{FMKL}(y)) = \frac{\lambda_2}{y^{\lambda_3-1} + (1-y)^{\lambda_4-1}} \quad (4.6)$$

Although both the *RS* and *FMKL GLD* are generalizations of Tuckey's Lambda Distribution, they are not equivalent, so that the distribution fitted by one parametrization to a dataset differs in general from the one fitted by the other (MARCONDES; PEIXOTO; MAIA, 2017). Both of these representations can present a wide variety of shapes and therefore are utilized in practice; however, it is generally the *FMKL GLD* preferred due

to the ease in its use (CORLU; METERELLIYOZ, 2016). In this thesis, we also use the *FMKL GLD* representation.

4.1.3 Other Parameterizations

One of the criticisms of the *GLD* is that its skewness is expressed in terms of both tail indices λ_3 and λ_4 . In one approach addressing this concern, a five-parameter *GLD* was introduced by Joiner et al. (JOINER; ROSENBLATT, 1971), which, expressed in the *FKML* parameterization, can be written as,

$$Q_{JR}(y|\lambda_1, \lambda_2, \lambda_3, \lambda_4) = \lambda_1 + \frac{1}{2\lambda_2} \left[(1 - \lambda_5) \frac{y^{\lambda_3} - 1}{\lambda_3} - (1 + \lambda_5) \frac{(1 - y)^{\lambda_4} - 1}{\lambda_4} \right] \quad (4.7)$$

It has λ_1 and λ_2 as the location and scale parameters, and an asymmetry parameter, λ_5 , which weights each side of the distribution and the two tail indices, λ_3 and λ_4 . The conditions on the parameters are $\lambda_2 > 0$ and $-1 < \lambda_5 < 1$. The drawback of this parameterization is that the additional parameter can make the estimation of the parameter values even more difficult.

In (CHALABI; DIETHELM; SCOTT, 2012) the authors introduce a new parameterization of the *GLD* that transform the *FMKL* parameterization, equation 4.5 in terms of an asymmetry and steepness parameter without adding a new variable. Its major advantage is that provides an intuitive interpretation of its parameters. A new **R** package called **gldist** was implemented with the new *GLD* parameterization, along with the parameter estimation methods they present in his work. The problem with this parameterization is that the **R** package was removed from the official repository because of the code is out of date.

4.2 FMKL GLD Shapes

Both *RS* and *FMKL GLD* can describe a variety of shapes, examples: U-shaped, bell shaped, triangular, and exponentially (SU, 2007). At the same time they also provide good fits to many well know distributions. In the case of *RS GLD* and extensive study can be found in (KARIAN; DUDEWICZ, 2011), for the *FMKL GLD* see (FREIMER; LIN; MUDHOLKAR, 1988).

Those properties of the *GLD* are important to our purpose for two reasons: first we don't need previous knowledge to use the *GLD* to fit any dataset, that is exactly the case in large-scale spatio-temporal models, and second the *GLD* can be compared and grouped based on its shapes, that allow us to answer the **RQ1** as we show in Chapter 5.

As the *FMKL GLD* parameterization is the one selected to be used in this thesis, in the next sub-sections we present a brief review of its shapes and how this parameterization

fit some well-known distributions.

The shape of the *GLD* are dependent of its λ values. In the case of the *FMKL GLD* parameterization, Freimer et al. (FREIMER; LIN; MUDHOLKAR, 1988) classify the shapes into five categories depending on the variety of distributions that can be represented by the several combinations of the shape parameters λ_3 and λ_4 . In particular, Class-I ($\lambda_3 < 1, \lambda_4 < 1$) represents unimodal densities with continuous tails. This class is subdivided in I_a ($\lambda_3, \lambda_4 \leq 1/2$), I_b ($1/2 < \lambda_3 < 1, \lambda_4 \leq 1/2$) and I_c ($1/2 < \lambda_3 < 1, 1/2 < \lambda_4 < 1$). In I_a we find distributions such as *Gaussian(Normal)*, *Beta(2,3)* and $\Gamma(\alpha = 5)$; in I_b $\Gamma(\alpha = 3)$ and *Lognormal*($\sigma = 0.5$); and in I_c distributions as the example of Class-I in Figure 3.

Class-II ($\lambda_3 > 1, \lambda_4 < 1$) represents monotone pdfs similar to the *Exponential* distribution, *Beta(1,2)* or *Lognormal*($\sigma = 1.0$). Class-III ($1 < \lambda_3 < 2, 1 < \lambda_4 < 2$) represents U-shaped densities with truncated tails, Class-IV ($\lambda_3 > 2, 1 < \lambda_4 < 2$) represents S-shaped densities, and Class-V ($\lambda_3 > 2, \lambda_4 > 2$) represents unimodal densities with truncated tails. Figure 3 provides the shapes that are represented by the parameters indicated in Table 3.

Table 3 – Examples of the five categories of distributions the *FMKL GLD* can represent.

	λ_1	λ_2	λ_3	λ_4
Class-I	0	1	0.5	0.6
Class-II	0	1	2	0.5
Class-III	0	1	1.5	1.5
Class-IV	0	1	2.5	1.5
Class-V	0	1	3	3

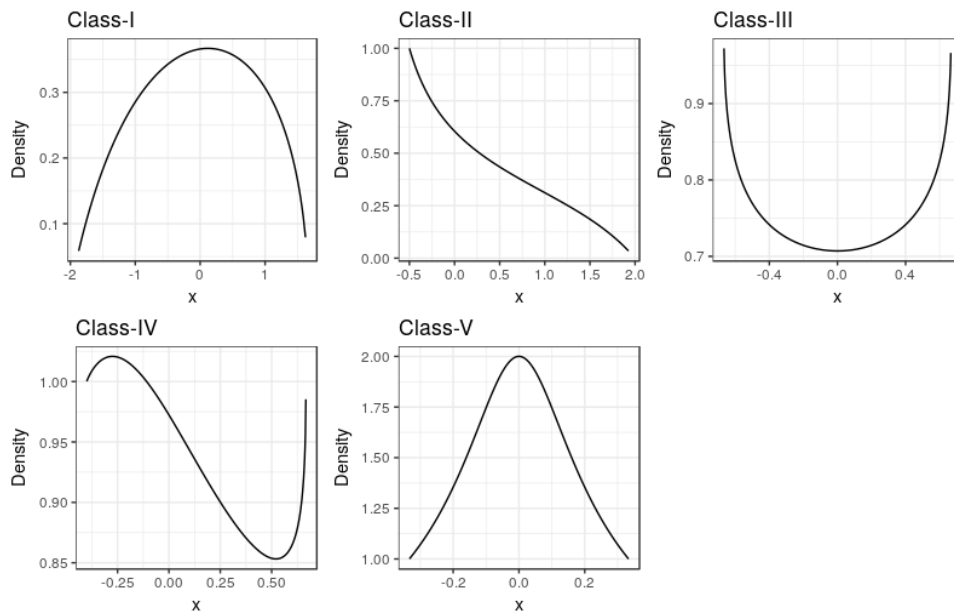


Figure 3 – Examples of the five categories of shapes the *FMKL GLD* can represent.

Figure 4 show the five categories of the *FMKL GLD* shapes in (λ_3, λ_4) space. There are two regions in this figure that were left out of the analysis, the regions with $(\lambda_3 < 1, \lambda_4 > 1)$ and $(1 < \lambda_3 < 2, \lambda_4 > 2)$. Those regions are symmetric to region II and IV respectively, see Figure 5.

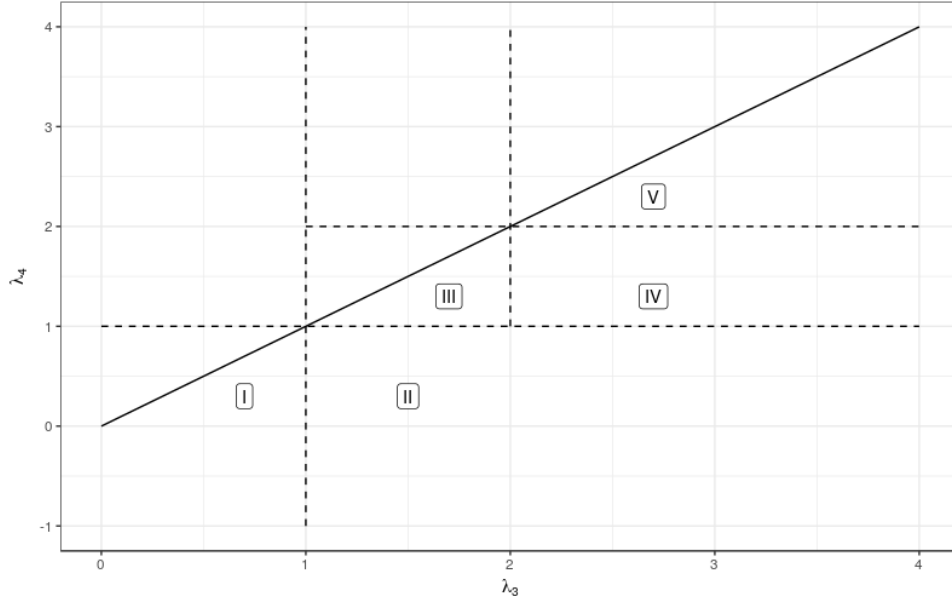


Figure 4 – The five categories of shapes of the *FMKL GLD* in the (λ_3, λ_4) space.

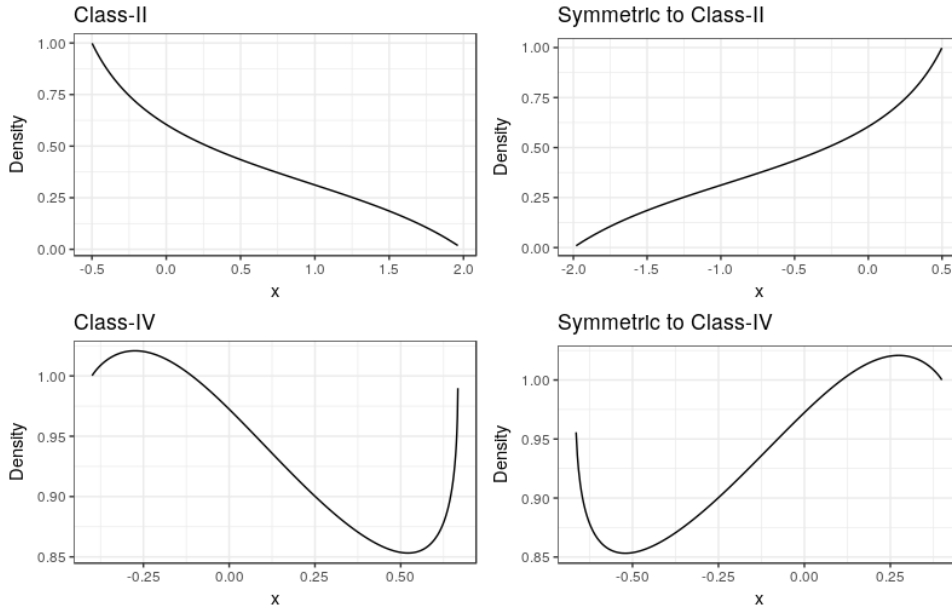


Figure 5 – Symmetry of the regions $(\lambda_3 < 1, \lambda_4 > 1)$ and $(1 < \lambda_3 < 2, \lambda_4 > 2)$ with respect to region II and IV.

The results presented in this section are extremely important to our approach because they are the bases of the clustering algorithms we propose in Chapter 5.

4.3 Numerical Methods to Fit the GLD to Data

Given a random sample $x_1, x_2, x_3, \dots, x_n$, the basic problem in fitting a statistical distribution to this data is that of approximating the distribution from which the sample was obtained. If it is known, because of theoretical considerations, that the distribution is of a certain type (e.g., a gamma distribution with unknown parameters), then through moment matching, or some other means, one can determine a specific distribution that fits the data. This, however, is generally not the case and, in the absence of any knowledge regarding the distribution, it makes sense to appeal to a flexible family of distributions and choose a specific member of that family (KARIAN; DUDEWICZ, 2011).

There are two different parameter estimation philosophies, **direct estimation methods**, such as least-squares estimation with order statistics [Ozturk and Dale, 1985] and with percentiles [Karian and Dudewicz, 1999, 2000; Fournier et al., 2007; King and MacGillivray, 2007; Karian and Dudewicz, 2003]; the methods of moments [Ozturk and Dale, 1982; Gilchrist, 2000], L-moments [Gilchrist, 2000; Karvanen and Nuutinen, 2008], and trimmed L-moments [Asquith, 2007]; and the goodness-of-fit method with histograms (??) and with maximum likelihood estimation (SU, 2007). On the other side, **stochastic methods** have been introduced with various estimators such as goodness-of-fit [Lakhany and Mausser, 2000] or the starship method [King and MacGillivray, 1999].

Moreover, (LAMPASI; Di Nicola; PODESTA, 2006)

Numerical maximum log likelihood estimation for generalized lambda distributions (LODZIENSIS, 2013) (LAKHANY; MAUSSER, 2000) (FOURNIER et al., 2007) (MARCONDES; PEIXOTO; MAIA, 2017) Review and Genetic Algorithms approach (CORLU; METERELLIYOZ, 2016)

4.4 GLD Approximations of Some Well-Known Distributions

For the $GLD(\lambda_1, \lambda_2, \lambda_3, \lambda_4)$ to be useful for fitting distributions to data, it should be able to provide good fits to many of the distributions the data may come from. In (KARIAN; DUDEWICZ, 2011), the authors explore how the GLD fit sixteen well-known distributions using the RS GLD parameterization. Here we explore how the GLD fit eight distributions, but using the $FMKL$ parameterization.

In Table 4 we show the original distribution, the four λ values of the fit and the result of apply a Kolmogorov-Smirnov test to validate the good of the fit.

Table 4 – GLD Approximations of 8 Well-Known Distributions

Distribution	Parameters	λ_1	λ_2	λ_3	λ_4	KS-test
Normal	$N(0, 1)$	-0.04263	1.49039	0.13787	0.12571	951
Uniform	$U(0, 1)$	0.46250	2.16223	1.00008	0.8614	912
Exponential	$\theta = 1$	0.49150	1.40546	1.44813	-0.10419	923
Chi-Square	$nu = 5$	4.141853	0.486702	0.508298	-0.045440	911
Gamma	$(\alpha = 5, \theta = 3)$	1.535078	1.846312	0.410183	0.027492	885
Weibull	$(\alpha = 1, \beta = 5)$	2.684657	0.263865	1.413826	-0.067658	940
Lognormal	$(\mu = 0, \sigma = 1/3)$	0.984696	4.516254	0.324879	-0.074348	903
Beta	$(\beta_3 = \beta_4 = 1)$	0.50092	2.00702	0.99505	1.00060	906

As was expected the results are slightly different to those presented by Karian et al., as we use other parameterization. The KS-test value was over 900 in seven cases and near 900 in the case of the Gamma distribution. This result suggests us that the fit was good. All the fit provides (λ_3, λ_4) values that match the shape region each distribution belongs to.

4.5 Fitting Mixture Distributions Using a Mixture of Generalized Lambda Distributions

Esto esta en (TOBERGTE; CURTIS, 2013)

Maximum Log Likelihood Estimation using EM Algorithm and Partition Maximum Log Likelihood Estimation for Mixtures of Generalized Lambda Distributions (SU, 2011). En este articulo presenta ejemplos bien extraños.

Fitting the GLD and compare with the normal mixture (NING; GAO; DUDEWICZ, 2008)

4.6 GLD Random Variate Generation

An important thing to take into account when we substitute the raw data produced as an output of a simulation process, by its *PDF* is that those *PDFs* need to allow us to reproduce the original data as close as possible. The outcome produced by a particular *PDF* is known as random variate, its definition is:

Definition 4.1. A **random variate** is a particular outcome of a random variable. The random variates which are other outcomes of the same random variable might have different values.

Random variates are used when simulating processes driven by random influences. One of the important applications of the *GLD* has been the generation of random variables for Monte Carlo studies (Mustafa Inchasi, 2016).

This fact is justified by the following theorem, enunciated by Karian and Dudewicz (2010).

Theorem 4.2. *If $Q_X(y)$ is the percentile function of a random variable X , and U is a uniform random variable on $(0, 1)$ then $Q_X(U)$ has the same PDF as does X .*

For a proof, also see p. 156 of Karian and Dudewicz (1999). The percentile function is not available in a closed (or easy-to-work-with) form for many of the most important distributions, such as the normal distribution. However, the GLD is (see sections ?? and ??) defined by its p.f., which is a simple-to-calculate expression.

Thus, **r.v.s for a simulation study can easily be generated from any distribution that can be modeled by a GLD.**

Example 4.3. Suppose we have modeled an important **r.v.** by an approximate standard normal distribution X . We show in Section 4.4 that a close fit to the standard normal is available via the *RS-GLD* with

$$(\lambda_1, \lambda_2, \lambda_3, \lambda_4) = (0, 0.1975, 0.1349, 0.1349) \quad (4.8)$$

and this *GLD* has **p.f.**

$$Q(y) = \frac{y^{0.1349} - (1 - y)^{0.1349}}{0.1975} \quad (4.9)$$

Thus, if U_1, U_2, \dots are independent uniform **r.v.s** on $(0, 1)$, then

$$Q(U_1), Q(U_2), \dots \quad (4.10)$$

are independent and (approximately) $N(0, 1)$ **r.v.s** for the simulation study at hand.

This theorem means that, independently of the nature of the dataset (Normal, Exponential, etc.), when we fit a GLD to it, we can proceed similarly to the example above. That is, we just need to generate a stream of independent uniform **r.v.s** on $(0, 1)$, and then evaluate the equation 4.10. There are a number of good sources of independent uniform r.v.s on $(0, 1)$ (KARIAN; DUDEWICZ, 2011).

This is an important property of the *GLD* that allow us to substitute the raw data by the four lambdas of the *GLD* that best fit it (if the fit is a good one), with the warranty that if we need to go back, the *GLD* could generate a good representation of the original data.

4.7 GLD and Uncertainty Quantification

4.7.1 Related Works

A solution to determining the reliability of products Using Generalized Lambda Distribution ([MOVAHEDI; LOTFI; NAYYERI, 2013](#))

Fundamental Reference ([LAMPASI; Di Nicola; PODESTA, 2006](#))

4.7.2 Relevance of GLD in Uncertainty Quantification

The use of the *GLD* to quantify the uncertainty is justified because:

- the *GLD* fits the *PDF* of a wide variety of datasets, including those that follow distributions such as normal, uniform, Student's t, U-shaped, exponential, etc;
- no prior knowledge is needed to fit the *GLD* to a dataset, which is practical and suitable for automatic and software procedures;
- the *PDF* is completely characterized by the four parameters of the *GLD*, which represents a reduction in the amount of data that must be stored for post-processing;
- the shape of the *GLD* is governed by its parameters, so the *GLDs* can be grouped based on their shapes, which is especially useful for further queries; and
- in cases where mixture of distributions are needed, *GLD* mixtures could be a very good option;
- a *GLD* is good for random variate generation.

([NING; GAO; DUDEWICZ, 2008](#))

4.8 The GLDEX R package

In the implementation of our approach, we use the **GLDEX R** package ([SU, 2007](#)). The GLDEX R package provides fitting algorithms with two objectives: (i) to provide a smoothing device to fit distributions to data using the weight and unweighted discretised approach based on the bin width of the histogram; (ii) to provide a definitive fit to the data set using the maximum likelihood estimation.

The GLDEX package also provides diagnostic tests to examine the quality of fit through the resample Kolmogorov-Smirnoff test, quantile plots and comparison of the mean, variance, skewness and kurtosis between the empirical data and the fitted distribution.

The GLDEX package is used in this thesis to: (i) fit the *GLD* distribution to a dataset on each spatio-temporal location; (ii) examine the quality of the fit; (iii) sampling any spatio-temporal location based on its *GLD*.

4.9 Conclusions

5 Clustering Uncertain Data Based on GLD Similarity

In chapter 4 we exposed the two most important parametrizations of the *GLD* and select the *FMKL* as the one to use for the rest of the thesis. In this parameterization λ_1 represent the location of the *GLD* and is directly related to the mean of the distribution. λ_2 is the scale, directly related to the standard deviation; and λ_3 and λ_4 represent the left and right tails of the distribution. Combinations of λ_3 and λ_4 can be used to estimate the skewness and kurtosis of the distribution.

The uncertainty can be characterized in many different ways as we mention in chapter 2, but from the *GLD* point of view, λ_2 , λ_3 and λ_4 are the responsables of this. So, in this chapter we try to answer the research question 1, we formulate in the introduction:

RQ1. how to group the output of the UQ process based on the similarity of the uncertainty?

First of all, in section 5.1 a brief review of some related works is performed. In this section, some advantage and drawbacks are highlighted, and some considerations about the possibilities of the use of the *GLD* to solve some of the drawbacks are commented. Next, in section 5.2 our hypothesis about the use of the *GLD* to clusterized uncertain data, is presented. Sections 5.3 and 5.4 present two synthetics datasets and the results of the clustering technique. Those results help us to validate our hypothesis. Finally, section 5.5 summarize and discuss the main results of the chapter.

5.1 Related Works

(JIANG et al., 2011)

5.2 Clustering Based on GLD

Our hypothesis is that, as the *GLD* shape is characterized by λ_2 , λ_3 and λ_4 , and this shape change slowly with the change in the λ_i values, we can group the uncertainty using clustering algorithms above λ_2 , λ_3 and λ_4 .

To test our hypothesis, we generate two synthetic datasets using 4 different probability density functions: Gaussian, Exponential, Uniform and Gamma. The structure of

the datasets is represented in 5.1.

$$S(x_i, \langle v_j \rangle) i = 1.....n, j = 1.....m \quad (5.1)$$

where:

- n represent the number of objects of the dataset and,
- m represent the size of each object.

For example, the first object could be a Gaussian distribution with size = 1000, mean = 0 and standard deviation = 2, figure ?? . The datasets are described in details in sections 5.3 and 5.4.

5.2.1 Fit the GLD to a dataset

When we generate a synthetic dataset, the next step is to find the *GLD* that best fit $\langle v_j \rangle$ on each x_i . As the fitting process is computationally intensive we implement a parallel algorithm using **R**. The pseudo-code is shown in algorithm 1.

Algorithm 1 Fitting the GLD to a synthetic dataset

```

1: function GLDFIT( $S(x_i, \langle v_1, v_2, \dots, v_n \rangle)$ )
2:    $\langle \lambda_1, \lambda_2, \lambda_3, \lambda_4 \rangle \leftarrow \text{FIT.GLD.LM}(\langle v_1, v_2, \dots, v_n \rangle)$ 
3:    $isValid_{(x_i)} \leftarrow \text{VALIDITYCHECK}(\langle \lambda_3, \lambda_4 \rangle)$ 
4:   if  $isValid_{(x_i)}$  then
5:      $[pvalue, D]_{(x_i)} \leftarrow \text{KS}(\langle \lambda_1, \lambda_2, \lambda_3, \lambda_4 \rangle_{(x_i)})$ 
6:   if  $pvalue_{(x_i)} > 0.05$  then
7:      $\text{STORELAMBDAS}(\langle \lambda_1, \lambda_2, \lambda_3, \lambda_4 \rangle, x_i)$ 

```

The algorithm receive a dataset represented by 5.1 and, for each position x_i , call a function **fit.gld.lm** from the **R** package **GLDEX** presented in section 4.8, line 2 of the algorithm 1. In line 3 we check the validity of the *GLD* returned by the function (remember from chapter 4 that the *GLD* is not always valid). In line 5 a good-of-fit test is perform to be sure that each *GLD* is a good representation for the dataset in x_i . Finally all the *GLD* with $pvalue > 0.05$ are stored to be used in the next section.

The final result of this process is a new dataset with the form:

$$S(x_i, \langle \lambda_1, \lambda_2, \lambda_3, \lambda_4 \rangle) i = 1.....n \quad (5.2)$$

5.2.2 Clustering the GLD

The clustering algorithm is trivial because the idea we try to test is that we can clusterize the uncertain data, using a simple k-means with a Euclidean distance over the λ_i space.

The dataset 5.2 is modified to remove λ_1 . Then, a k-means algorithm is used first over $\langle \lambda_2, \lambda_3, \lambda_4 \rangle$ and second over $\langle \lambda_3, \lambda_4 \rangle$. The results are discussed in sections 5.3 and 5.4.

5.3 Synthetic Data I

To generate the first synthetic data set we use 11 probability density functions, where 5 are Gaussian, 5 Exponential, and one Uniform, figures 6, 7 and 8. The standard deviation of the 5 Gaussian distributions is $0.05 * i$, with $i = 1, 2, 3, 4, 5$, and we generate 90 samples of each distribution. This is, the first 90 objects where generated from a Gaussian distribution with standard deviation 0.05, and so on. Similarly, the rate of the 5 Exponential distributions is i , with $i = 1, 2, 3, 4, 5$, and again we generate 90 samples of each one. Finally, 100 samples of a Uniform distribution between $[0, 1]$ were generated. In resume, we have 1000 objects, where the first 450 were sampled from a Gaussian distributions, the next 450 from an Exponential and the last 100 from a Uniform distribution. As we generate a synthetic dataset in this way, we have the ground truth of the clustering in the dataset. This ground truth is used to evaluate the clustering quality of our algorithms.

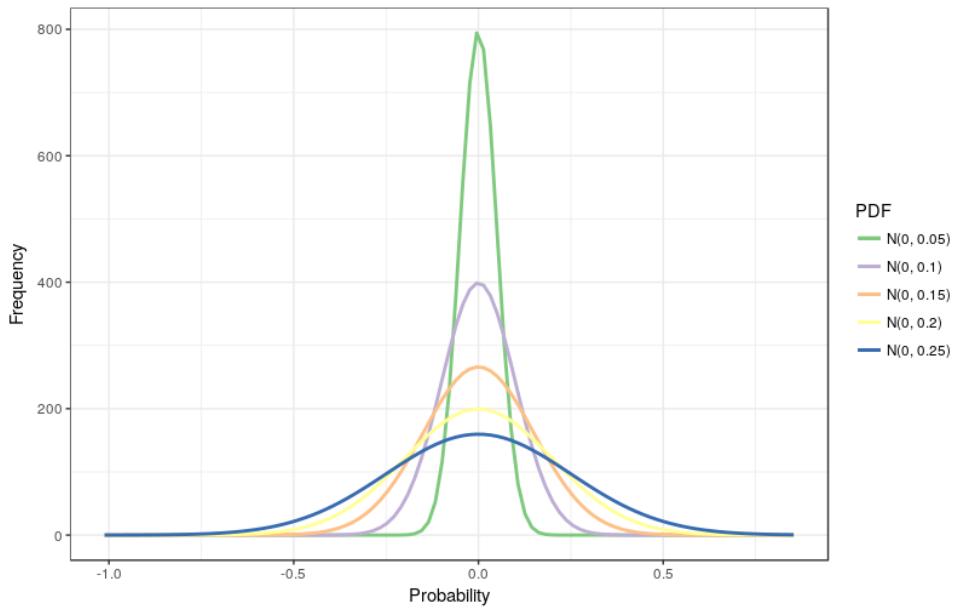


Figure 6 – Gaussian (Normal) distributions used to generate the synthetic dataset.

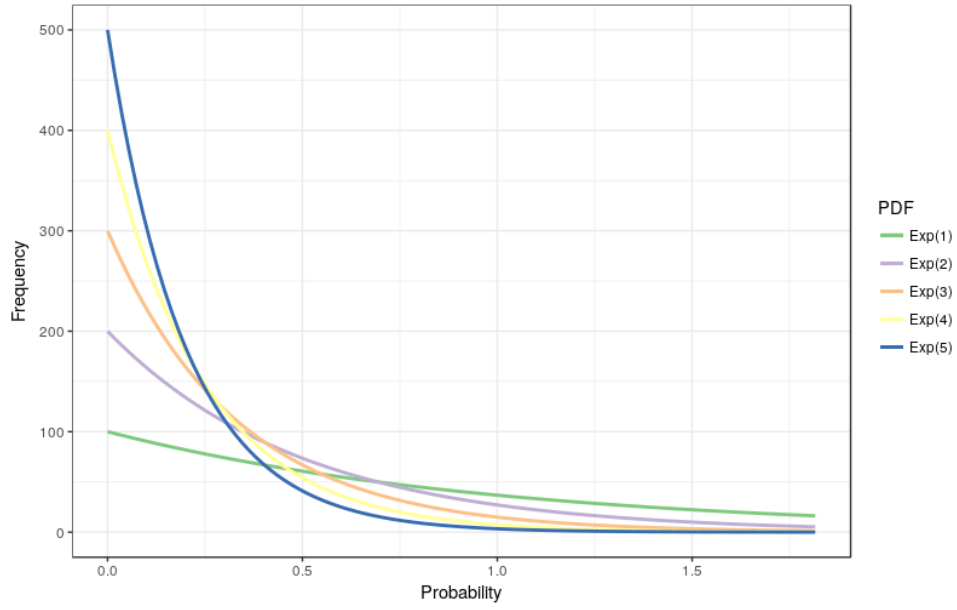


Figure 7 – Exponential distributions used to generate the synthetic dataset.

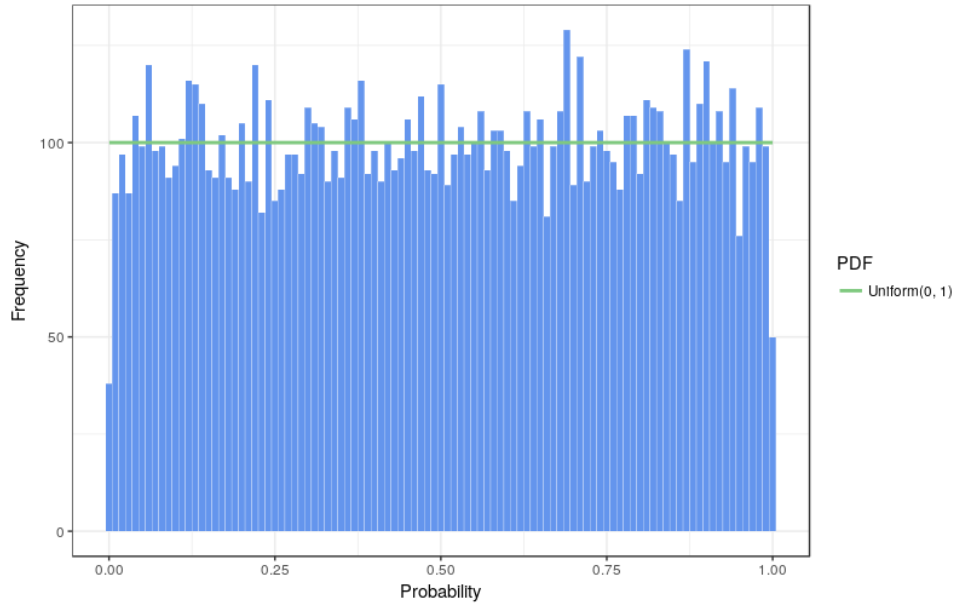


Figure 8 – Uniform distribution used to generate the synthetic dataset.

This dataset could be represented as a multidimensional array where for each position x_i , we have 1000 values v_j , equation 5.3. In this case i and j vary from 1 to 1000 casually.

$$S(x_i, < v_j >) i, j = 1, 2, \dots, 1000 \quad (5.3)$$

The fitting algorithm proposed in subsection 5.2.1 is applied over 5.3. The good-of-fit test return that all the *GLDs* are good fit for its corresponding distribution. As a result

the dataset 5.4 is generated.

$$S(x_i, < \lambda_1, \lambda_2, \lambda_3, \lambda_4 >) i = 1.....1000 \quad (5.4)$$

5.3.1 Clustering using λ_2 , λ_3 and λ_4

As we mention above, our idea is to test what happen if we use a simple k-means with euclidean distance over the λ_2 , λ_3 and λ_4 values of the *GLDs*. Similar to the paper (JIANG et al., 2011), as we use 11 *PDFs* to generate the synthetic dataset I, we expect that the k-mean algorithm will return 11 clusters as well (one for each distribution). Then 11 is the number we use with the k-means algorithm.

In figure 9 and table 5 the distribution of the clusters returned by the k-means is shown.

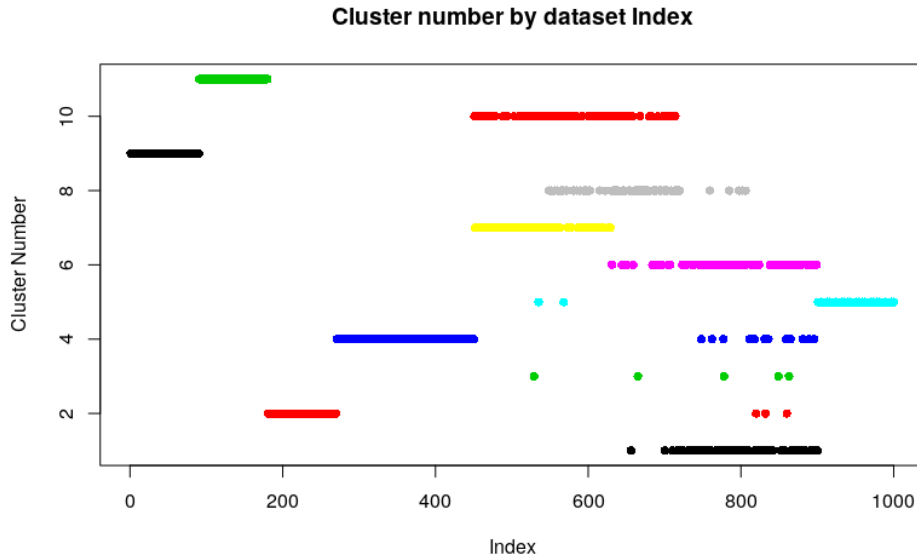


Figure 9 – Distribution of the clusters using k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*.

Remembered, the first 450 elements are normal distributions, the next 450 are exponential and the last 100 are uniform. Looking to those regions in general, the first observation is that we have 23 false positives, three in cluster 2, 18 in cluster 4 and two in cluster 5. The second observation is that, the normal distributions were grouped in 4 clusters (2, 4, 9 and 11), cluster 2 group perfectly it 90 elements with 2 false positives, clusters 9 and 11 group exactly its 90 elements each. The cluster 4 group the last 180 elements of the Normal distribution, with 18 false positives

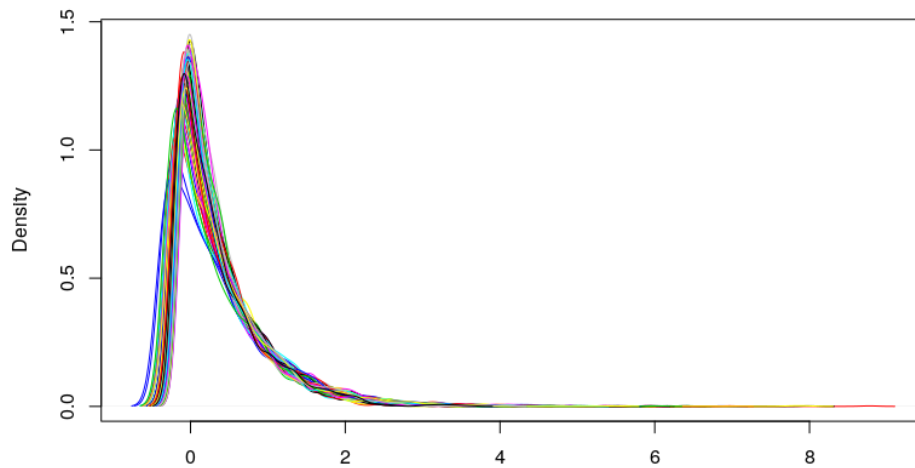
The Uniform distribution was grouped totally in cluster 5, with two false positives as was mention above. The last observation is that the algorithm can't separate the 5

Table 5 – Distribution of the clusters using k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*.

Cluster	Type of Distribution	No. of Elements
1	Exponential	82
2	Normal	93
3	Exponential	5
4	Normal	198
5	Uniform	102
6	Exponential	83
7	Exponential	91
8	Exponential	82
9	Normal	90
10	Exponential	84
11	Normal	90

Exponential distributions, but this is not a bad result as we will show soon.

In figures between 10 and 20 we show the *PDFs* of all the distributions that belongs to the same cluster. If we take a look at figures 10, 11, 12, 18, 19 and 20 we see that the exponential distribution was well grouped. Really the problem is that the rate value of $0.05 * i$ used to generate the exponential distribution does not have such a big difference between one and another.

Figure 10 – Cluster 1 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

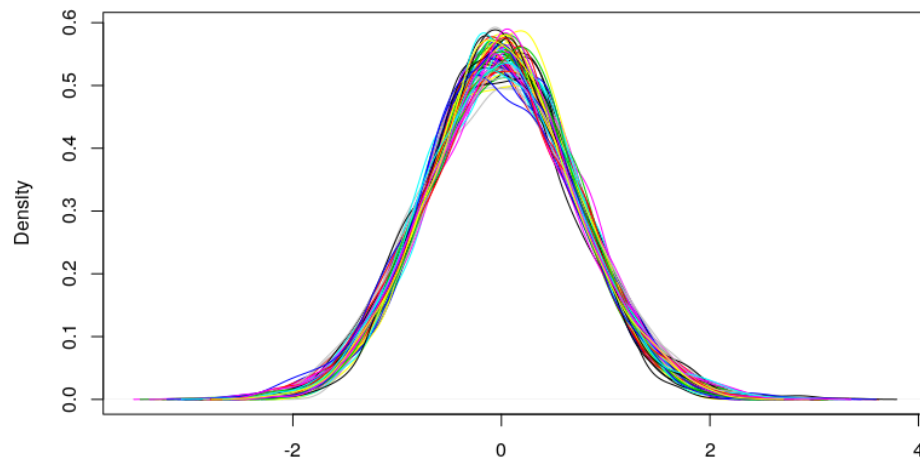


Figure 11 – Cluster 2 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

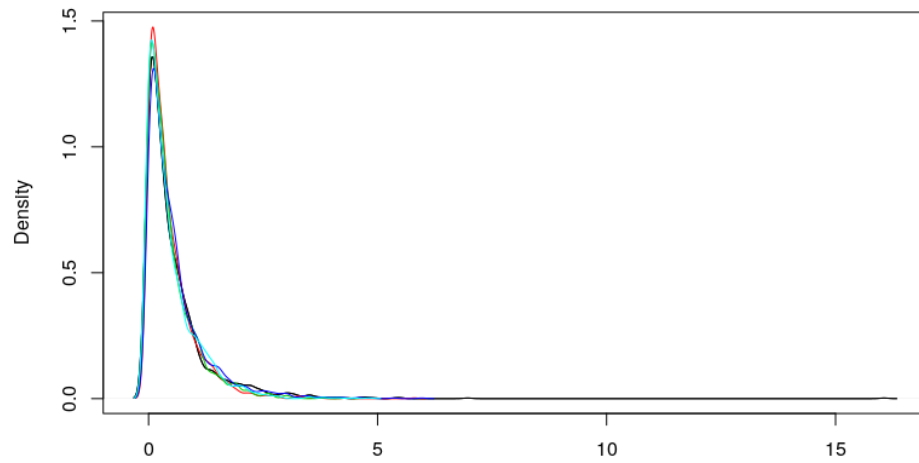


Figure 12 – Cluster 3 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

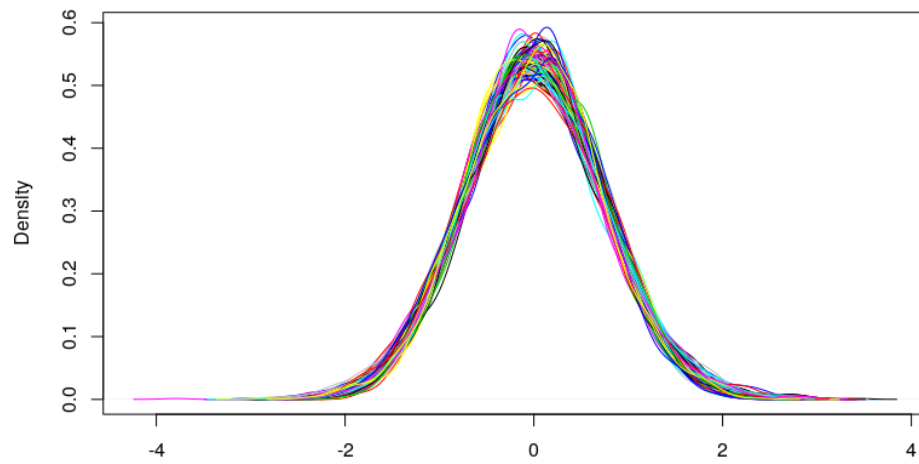


Figure 13 – Cluster 4 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

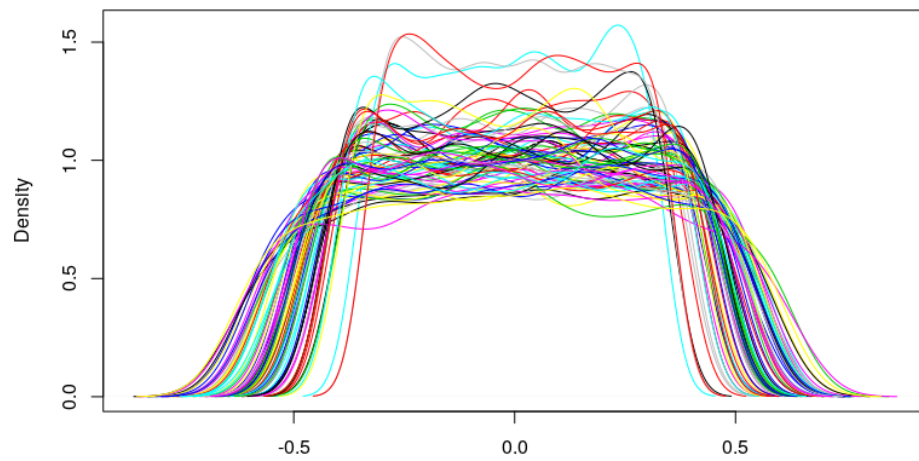


Figure 14 – Cluster 5 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

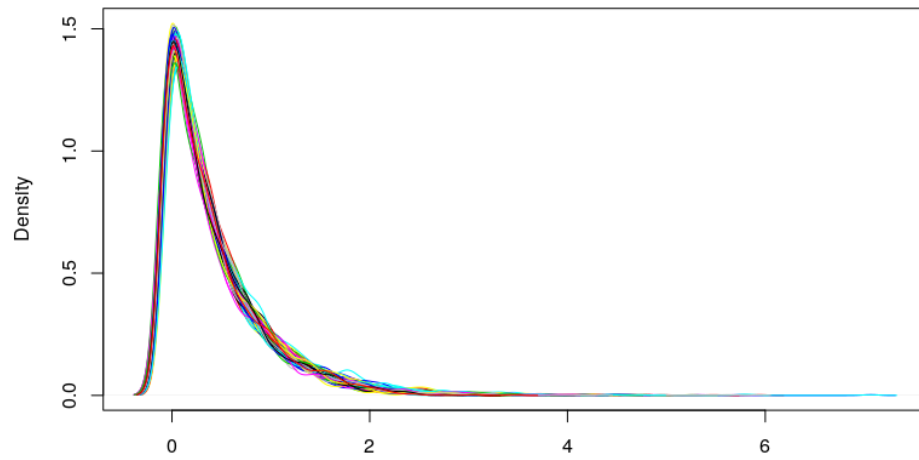


Figure 15 – Cluster 6 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

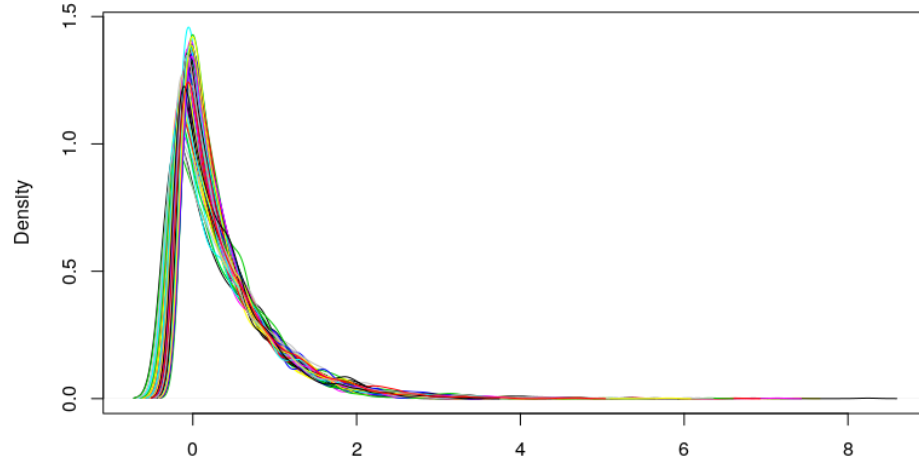


Figure 16 – Cluster 7 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

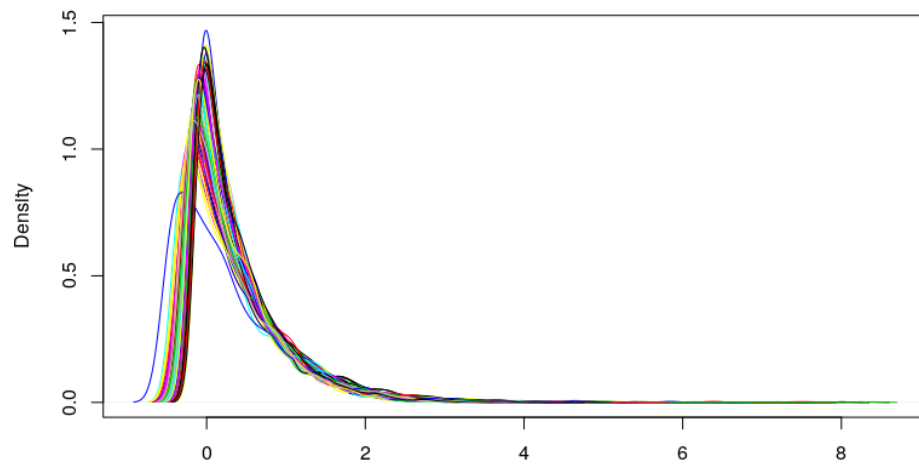


Figure 17 – Cluster 8 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

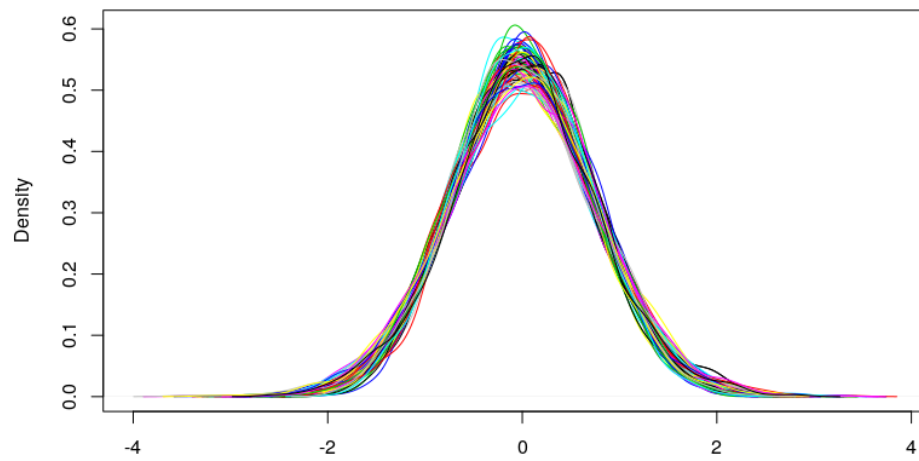


Figure 18 – Cluster 9 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

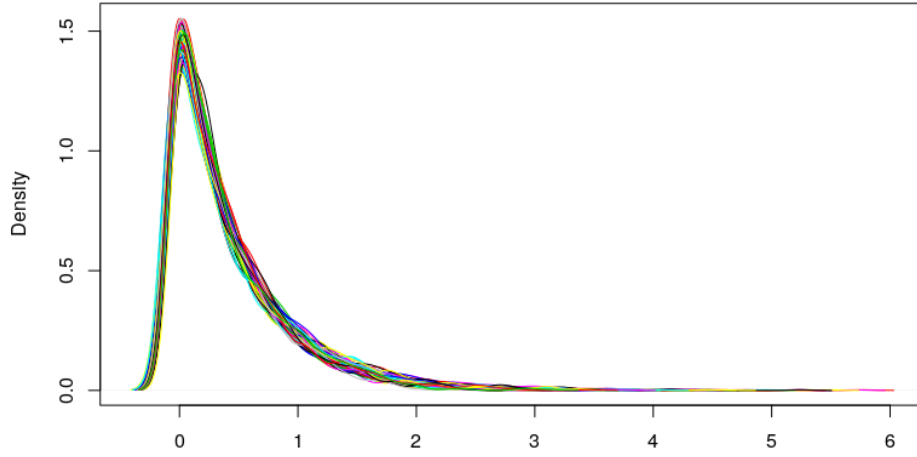


Figure 19 – Cluster 10 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

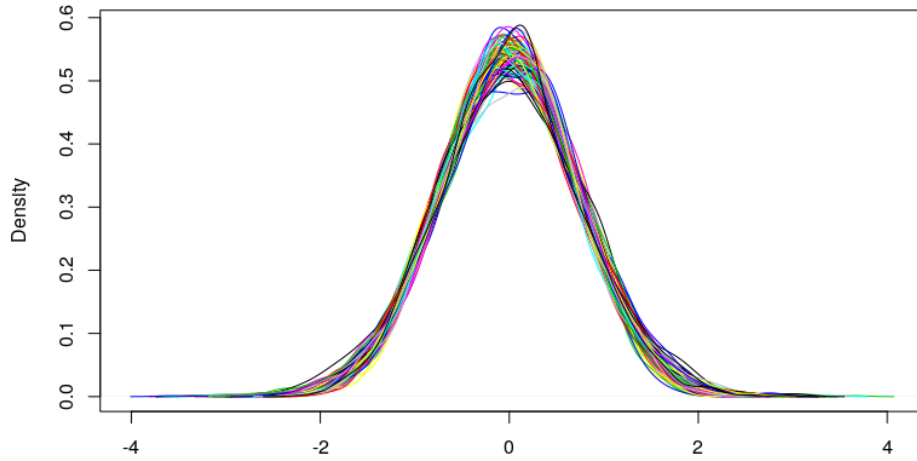


Figure 20 – Cluster 11 returned by the k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*, synthetic dataset I.

Another interesting result is shown in figures 21 and 22. As we can see, clusters 2, 4, 9 and 11 that represent the Normal distribution are all at the same region over the λ_3 and λ_4 space, near $\lambda_3 = 0$ and $\lambda_4 \in [0, 0.3]$. Similarly cluster 5, that represents the Uniform distribution is on the top left of the λ_3 and λ_4 space, $\lambda_3 \in [0, 0.3]$ and $\lambda_4 \in [0.7, 1.5]$. And finally the rest of the clusters that represent the Exponential distribution are distributed

in the bottom of the λ_3 and λ_4 space, $\lambda_3 \in [0.2, 7]$ and $\lambda_4 \in [-0.1, 0.1]$. As we see in the rest of the thesis, this result is repeated in all the use cases.

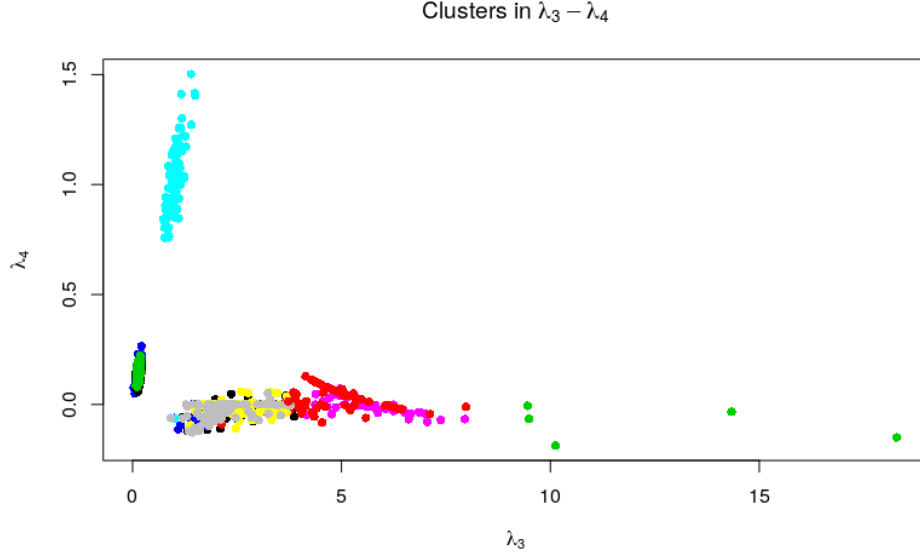


Figure 21 – Distribution of the clusters over the λ_3 and λ_4 space.

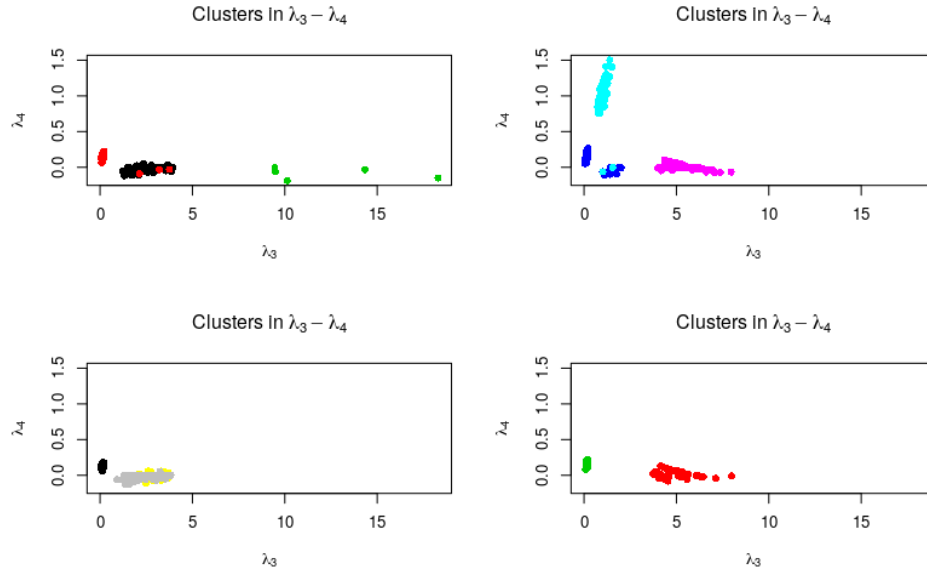


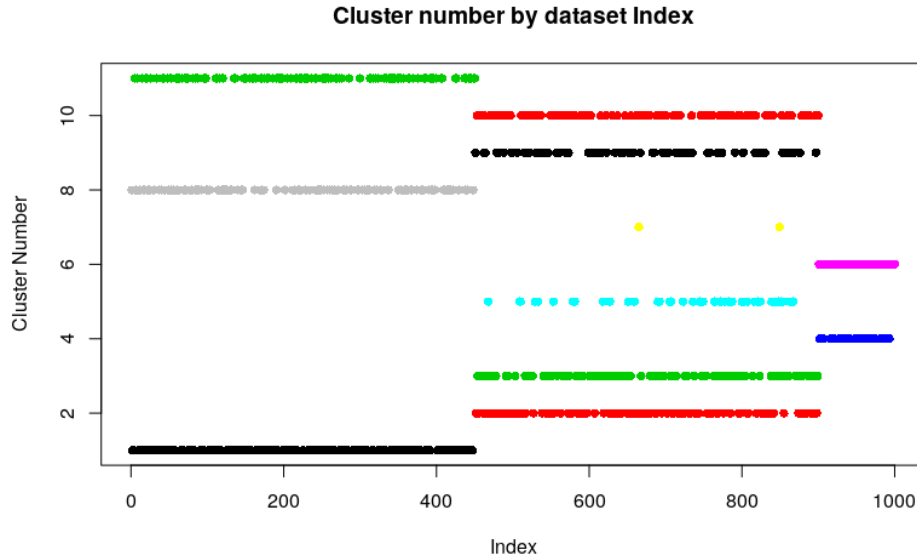
Figure 22 – Distribution of the clusters over the λ_3 and λ_4 space. In the top left corner: clusters 1, 2 and 3. Top right corner: clusters 4, 5 and 6. Bottom left: clusters 7, 8 and 9. Bottom right: clusters 10 and 11.

5.3.2 Clustering using λ_3 and λ_4

In this section we proceed similar to section 5.3.1, but the k-means algorithm run over λ_3 and λ_4 . The distribution of the clusters is shown in figure 23 and table 6.

Table 6 – Distribution of the clusters using k-means over the λ_3 and λ_4 values of the *GLDs*.

Cluster	Type of Distribution	No. of Elements
1	Normal	197
2	Exponential	118
3	Exponential	110
4	Uniform	35
5	Exponential	41
6	Uniform	65
7	Exponential	2
8	Normal	131
9	Exponential	74
10	Exponential	105
11	Normal	122

Figure 23 – Distribution of the clusters using k-means over the λ_3 and λ_4 values of the *GLDs*.

As we don't use λ_2 here, is clear that the algorithm can't distinguish the distributions by its standard deviation. But, as the shape of the *GLD* is defined by λ_3 and λ_4 , what we expect is that the algorithm can separate the objects by type of distribution. As we see in figure 23 this is exactly what we get, there is no any false positive in this case, the three regions (Normal, Exponential and Uniform) are identified by the k-means.

Clusters 1, 8 and 11 group all the Normal distributions, clusters 4 and 6 group the Uniform and the rest group the Exponential.

In the λ_3 and λ_4 space the behavior is very similar at the one we get in subsection 5.3.1, figures 24 and 25. Again the distributions are concentrated near the same (λ_3, λ_4)

values.

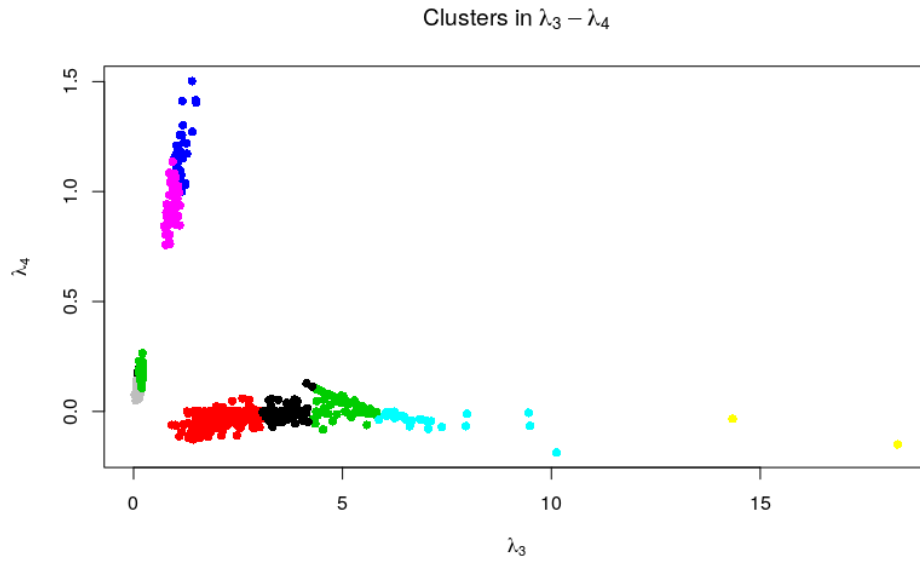


Figure 24 – Distribution of the clusters over the λ_3 and λ_4 space.

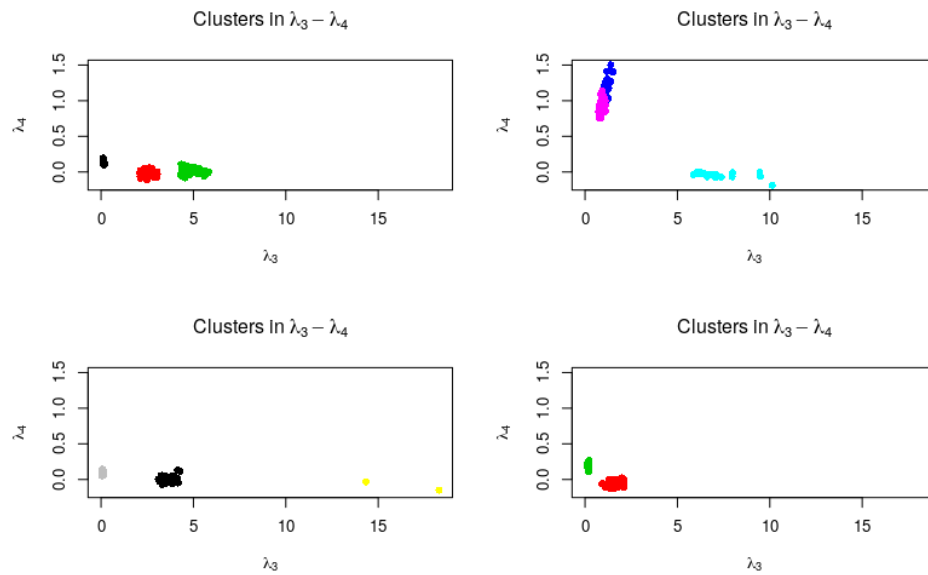


Figure 25 – Distribution of the clusters over the λ_3 and λ_4 space. In the top left corner: clusters 1, 2 and 3. Top right corner: clusters 4, 5 and 6. Bottom left: clusters 7, 8 and 9. Bottom right: clusters 10 and 11.

5.4 Synthetic Data II

The second synthetic dataset is similar to the first one, here we include 5 Gamma distributions, between the Exponential and the Uniform, figure 26. The shape of the

Gamma distribution is i , with $i = 1, 2, 3, 4, 5$. This dataset have 1450 objects, where the first 450 were sampled from a Gaussian distributions, the next 450 from an Exponential, the next 450 are Gamma, and the last 100 from a Uniform distribution. As we use 16 different distributions, this is the number of clusters to be used with the k-means algorithm.

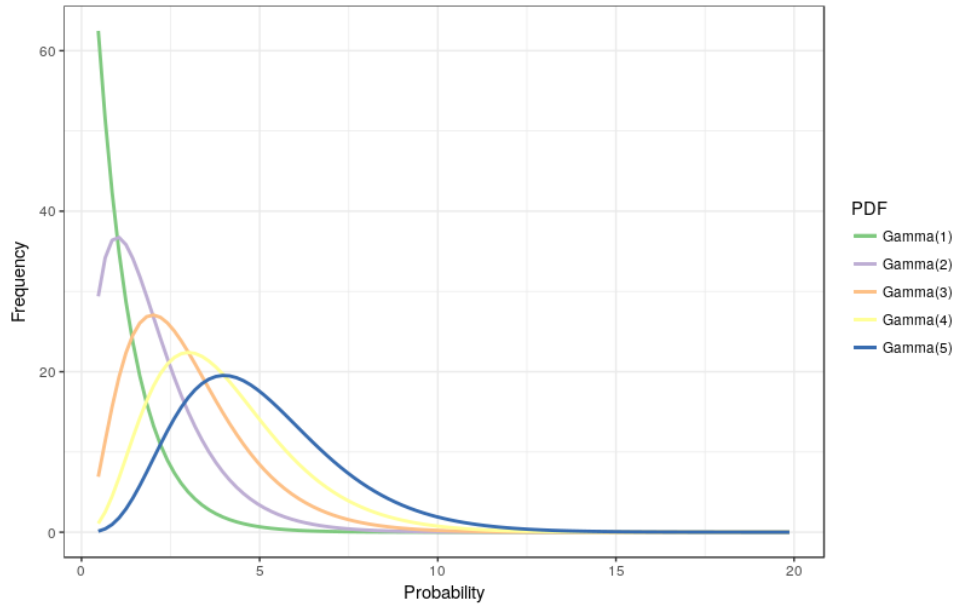


Figure 26 – Gamma distributions used to generate the synthetic dataset.

Similar to the dataset I, the fitting algorithm proposed in subsection 5.2.1 is applied over dataset II. The good-of-fit test return that all the *GLDs* are good fit for its corresponding distribution.

5.4.1 Clustering using λ_2 , λ_3 and λ_4

The distribution of the clusters returned by the k-means algorithm is shown in figure 27 and table 7.

Table 7 – Distribution of the clusters using k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*.

Cluster	Type of Distribution	No. of Elements
1	Normal	90
2	Exponential	44
3	Uniform	45
4	Normal	179
5	Gamma	60
6	Uniform	55
7	Exponential	87
8	Exponential	58
9	Exponential	67
10	Exponential	74
11	Exponential	25
12	Exponential	90
13	Normal	96
14	Normal	90
15	Gamma	30
16	Gamma	360

Figure 27 – Distribution of the clusters using k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*.

In general the results are very similar to the results of the section 5.3, but we get less false positives, 5 in total. 3 false positives in cluster 4 and 2 false positives in cluster 13. The normal distribution was grouping again in for clusters: 1, 4, 13 and 14. The uniform distribution was grouping in clusters 3 and 6 without false positives. The gamma distribution introduced here was grouped in clusters 5, 15 and 16, without false positives. And finally the rest of the clusters are for the exponential distribution.

The projection of the clusters over the λ_3 and λ_4 space is shown in figure 28. The two clusters of the uniform distribution are located again in the top-left region of the figure. The normal distribution is located in the same place, near $\lambda_3 = 0$ and $\lambda_4 \in [0, 0.3]$. The exponential distribution is distributed in the bottom of the λ_3 and λ_4 space. The gamma distribution is overlapped together with the normal distribution.

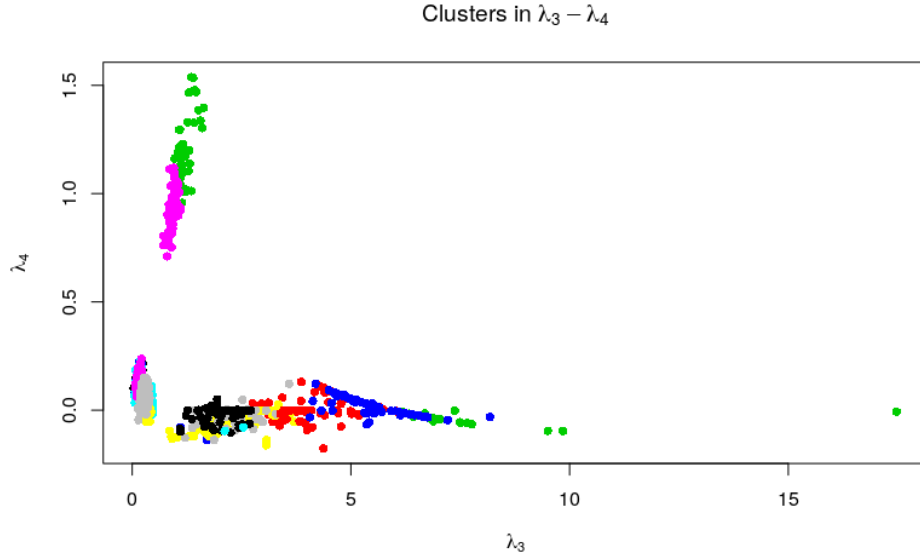


Figure 28 – Distribution of the clusters over the λ_3 and λ_4 space.

5.4.2 Clustering using λ_3 and λ_4

The distribution of the clusters returned by the k-means when using the values of λ_3 and λ_4 to group the second synthetic dataset are shown in figure 29 and table 8.

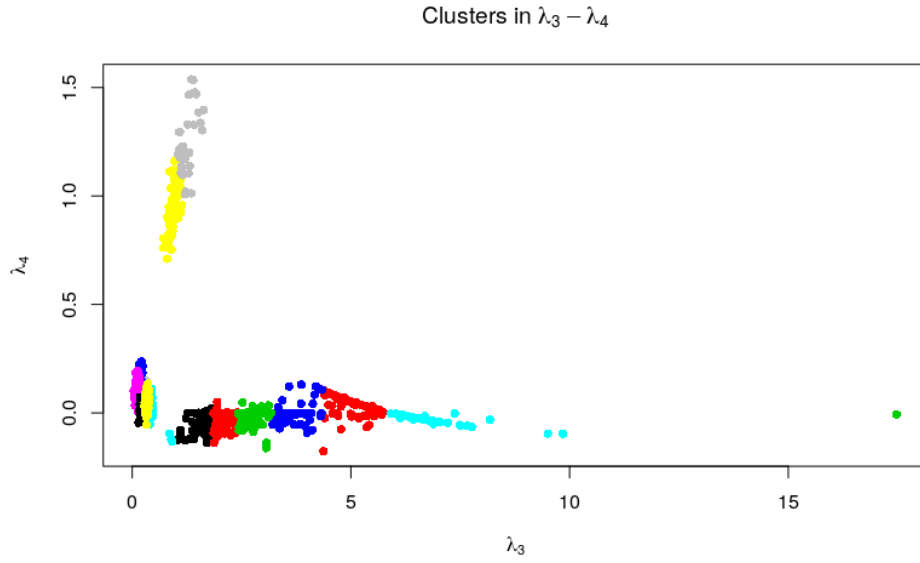
A few false positives are observed in clusters 5, 6 and 12, but nothing to worry about. Again the regions of the four distribution families are perfectly separated by the algorithm.

Table 8 – Distribution of the clusters using k-means over the λ_3 and λ_4 values of the *GLDs*.

Cluster	Type of Distribution	No. of Elements
1	Exponential	64
2	Exponential	126
3	Exponential	1
4	Exponential	57
5	Gamma	83
6	Normal	139
7	Uniform	67
8	Gamma	148
9	Gamma	108
10	Exponential	75
11	Exponential	80
12	Normal	112
13	Exponential	44
14	Normal	201
15	Gamma	112
16	Uniform	33

Figure 29 – Distribution of the clusters using k-means over the λ_2 , λ_3 and λ_4 values of the *GLDs*.

The projection of the clusters over the λ_3 and λ_4 space is show in figure 30.

Figure 30 – Distribution of the clusters over the λ_3 and λ_4 space.

5.5 Conclusions

In this chapter, we explore clustering uncertain data based on the similarity between their distributions. The idea is to answer the research question 1 *"how to group the output of the UQ process based on the similarity of the uncertainty?"*. Different to the approaches reported in the literature, we propose the use of a k-means algorithm with euclidean distance over the λ_2 , λ_3 and λ_4 space of the *GLD*.

The approach was tested over two synthetic datasets and the results of the test were exactly what we expect.

6 Kriging of the GLD parameters

6.1 Spatio-temporal Interpolation

6.2 Kriging over GLD

6.3 Use Case

6.4 Conclusions

7 Our Approach

In the more general case the computational model $\mathbf{q} = \mathcal{M}(\boldsymbol{\theta})$ represents the spatio-temporal evolution of a complex systems, and the *QoI* \mathbf{q} could be represented as:

$$\mathbf{Q} = (\mathbf{q}(s_1, t_1), \mathbf{q}(s_2, t_2), \dots, \mathbf{q}(s_n, t_n)) \quad (7.1)$$

where:

- $(s_1, t_1), (s_2, t_2), \dots, (s_n, t_n) \in \mathcal{S} \times \mathcal{T} \subseteq \mathbb{R}^3 \times \mathbb{R}$ represent a set of distinct spatio-temporal locations, and
- $\mathbf{q}(s_i, t_j)$ represents a value of the *QoI* at the spatio-temporal location (s_i, t_j)

We could have many *QoI*, but for simplicity here we are going to consider just one.

In the presence of a stochastic problem, on each spatio-temporal location (s_i, t_j) we have many realizations of $q(s_i, t_j)$. A structure of a database to store this information can be modeled as:

$$S(s_i, t_j, \text{simId}, q(s_i, t_j)) \quad (7.2)$$

where *simId* represents the *id* of one simulation (realization).

The first step of our approach consists of trying to find the *GLD* that best fits our simulations on each spatio-temporal location. The algorithms are described in the next section.

7.1 Fitting a GLD to a spatio-temporal dataset

Given a random sample $q_1, q_2, q_3, \dots, q_n$, the basic problem in fitting a statistical distribution to these data is that of approximating the distribution from which the sample was obtained. In our approach we divide this step in three task:

- Fit the *GLD* to the data.
- Evaluate the validity of the resulting *GLD* on each spatio-temporal location.
- Perform a ks-test to evaluate the quality of the fit on each spatio-temporal location.

The fitting process has been implemented following the algorithm 2. Before starting the fitting process, we need to group all the simulations that correspond to the same spatio-temporal location (s_i, t_j) . As a result we get a new dataset $S^*(s_i, t_j, \langle q_1, q_2, \dots, q_n \rangle)$, where $q_i, 1 \leq i \leq n$, represents a vector of all the values of q at point (s_i, t_j) .

7.1.1 Fitting process

Now, for each spatio-temporal location $(s_i, t_j) \in \mathcal{S} \times \mathcal{T}$ we use a function of the GLDEX R package described in section ??, to fit the *GLD* to a vector $\langle q_1, q_2, \dots, q_n \rangle$, line 2 of algorithm 2. As a result of this task we get the lambda values of the *GLD* that best fit the dataset at each spatio-temporal location, equation 7.3.

$$S'(s_i, t_j, GLD(\lambda_1, \lambda_2, \lambda_3, \lambda_4)) \quad (7.3)$$

7.1.2 GLD validity check

As we mention in section ?? the *GLD* is not always valid, it depends of the λ_3 and λ_4 values. The evaluation of the validity of the *GLD* is straightforward, if λ_3 and λ_4 are in the gray regions of Figure ?? the *GLD* is not valid, on the other case is valid.

The validity check is performed in line 3 of the algorithm 2, and as a result we get:

$$S_{\text{validity}}(s_i, t_j, \text{valid}(s_i, t_j)), \quad (7.4)$$

where:

$$\text{valid}(s_i, t_j) = \begin{cases} 1 & \text{if GLD is valid in } (s_i, t_j) \\ 0 & \text{otherwise} \end{cases} \quad (7.5)$$

7.1.3 Quality of the fit

Now at the remaining points, where the *GLD* is valid, we need to evaluate how good is the fit. That is, we evaluate whether the *GLD* (PDF) correctly describes the dataset. We use here the Kolmogorov-Smirnov test (KS-test). The KS-test determines if two datasets differ significantly. In this case, these datasets are: the original dataset and a second one generated using the fitted *GLD*. As a result, this test returns two values: a Kolmogorov-Smirnoff Distance (D); and a p-value, line 5 of algorithm 2. The distance D is the maximum distance between both cumulative density functions (CDF), as shown in Figure 31. A small distance means that both, the dataset and the fitted PDF, are similar.

The second value, the p-value, is a more robust test, as it helps us to determine the significance of our results. Suppose we have two hypotheses, the null hypothesis is that our PDF is a good fit to our dataset, and the alternative hypothesis is that it is not. Then, a small p-value (typically ≤ 0.05) indicates strong evidence against the null hypothesis, so you reject the null hypothesis. A large p-value (> 0.05) indicates weak evidence against the null hypothesis, so you fail to reject the null hypothesis. p-values very close to the cutoff (0.05) are considered to be marginal (could go either way).

At the end of this task we have two new multidimensional arrays with the values of \mathcal{D} and p-value on each spatio-temporal locations.

$$S_{\mathcal{D}}(s_i, t_j, \mathcal{D}(s_i, t_j)) \quad (7.6)$$

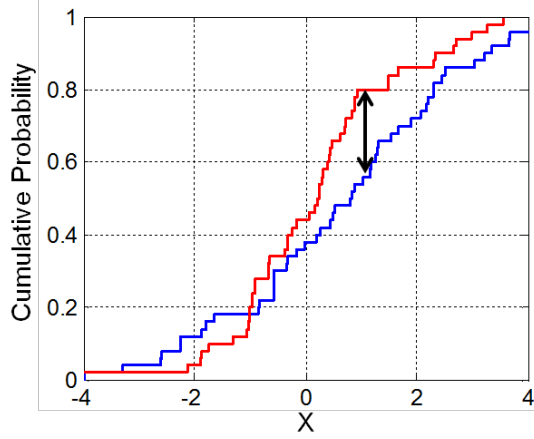


Figure 31 – Illustration of the two-sample Kolmogorov–Smirnov statistic. Red and blue lines each correspond to an empirical distribution function, and the black arrow is the two-sample KS statistic.

$$S_{pvalue}(s_i, t_j, pvalue(s_i, t_j)) \quad (7.7)$$

Finally, in line 7 of the algorithm 2 we store the lambda values of those *GLDs* that are valid and return p-values greater than 0.05.

Algorithm 2 Fitting the GLD to a spatio-temporal dataset

```

1: function GLDFIT( $S(s_i, t_j, < q_1, q_2, \dots, q_n >)$ )
2:    $< \lambda_1, \lambda_2, \lambda_3, \lambda_4 > \leftarrow \text{FIT.GLD.LM}(< q_1, q_2, \dots, q_n >)$ 
3:    $isValid_{(s_i, t_j)} \leftarrow \text{VALIDITYCHECK}(< \lambda_3, \lambda_4 >)$ 
4:   if  $isValid_{(s_i, t_j)}$  then
5:      $[pvalue, D]_{(s_i, t_j)} \leftarrow \text{KS}(< \lambda_1, \lambda_2, \lambda_3, \lambda_4 >_{(s_i, t_j)})$ 
6:   if  $pvalue_{(s_i, t_j)} > 0.05$  then
7:      $\text{STORELAMBDAS}(< \lambda_1, \lambda_2, \lambda_3, \lambda_4 >, s_i, t_j)$ 

```

7.2 Clustering the GLD based on its lambda values

In section ?? we discussed the different shapes of the GLD and define the regions of the (λ_3, λ_4) space where the shapes are similar. In Figure ??, we show how similar values of λ_3 and λ_4 lead to similar shapes. This fact suggests that one can clusterize the *GLD* based on its lambda values. The result of this clusterization are groups of *GLDs* with similar shapes (behaviors).

In addition to λ_3 and λ_4 , which represent the right and left tails of the distribution, we have also to consider λ_2 , as the latter represents the dispersion of the distribution.

Then, in this step of our workflow, we clusterize the *GLDs* using λ_2 , λ_3 and λ_4 values. The final result of this step is:

$$S_{\mathcal{C}}(s_i, t_j, GLD_k, clusterID) \quad (7.8)$$

where: *clusterID* represents the ID of the cluster to which the *GLD* at the spatio-temporal location (s_i, t_j) belongs.

With the *GLD* clusterized, we can use this result to characterize the uncertainty in a particular spatio-temporal region, or to measure numerically the corresponding uncertainty. In subsections 7.3 and 7.4, we describe how those approaches are implemented (see Figure 32).

7.3 Use of GLD mixture to characterize the uncertainty in an spatio-temporal region

One of the main advantages of assessing the complete probability distribution of the outputs, in place of low order moments (mean and standard deviation), is that we can use the *PDFs* to answer queries. For example, suppose we want to know the mean and standard deviation in a particular spatio-temporal region $(\mathcal{S}_i \times \mathcal{T}_j)$, or we want to observe graphically the distribution of the raw data generated in the simulation process in a spatio-temporal region.

Let us consider the second query. Up to this point, we have discussed the fit of *GLDs* that characterize the uncertainty at each spatio-temporal locations (s_i, t_j) , and the cluster to which the *GLD* at that particular spatio-temporal location would belong to. If we consider the clusterization of *GLD* to be of good quality, we can pick the *GLD* at the centroid of each cluster as a representative of all its members. In this context, in a particular spatio-temporal region, each cluster may be qualified with a weight given by:

$$w_k = \frac{1}{N} \sum_{i=1}^S \sum_{j=1}^T w(s_i, t_j) \quad (7.9)$$

where:

$$w(s_i, t_j) = \begin{cases} 1 & \text{if } clusterID(s_i, t_j) = k \\ 0 & \text{otherwise} \end{cases} \quad (7.10)$$

and N is the number of points in the region $(\mathcal{S}_i \times \mathcal{T}_j)$.

The weight w_k is the frequentist probability of occurrence of the cluster k in the region, and complies with the conditions outlined in section ?? that $w_k \geq 0$ and $\sum w_k = 1$.

Remember that the mixture of the *GLDs* can be written as:

$$f(x) = \sum_{k=1}^K w_k GLD(\lambda_1, \lambda_2, \lambda_3, \lambda_4) \quad (7.11)$$

So, if we have the weights and a representative GLD for each cluster, we have the mixture of GLD that characterizes the uncertainty in the spatio-temporal region $(\mathcal{S}_i \times \mathcal{T}_j)$.

The GLD mixture process is summarized in algorithm 3.

Algorithm 3 GLD mixture in a region $(\mathcal{S}_i \times \mathcal{T}_j)$

```

1: function GLDMIXTURE( $\mathcal{S}_i \times \mathcal{T}_j, C_{\mathcal{S}_i \times \mathcal{T}_j}$ )
2:   for each  $p_i$  in  $(\mathcal{S}_i \times \mathcal{T}_j)$  do
3:      $c \leftarrow \text{cluster}(p_i)$ 
4:      $w_c = w_c + 1$ 
5:      $N = N + 1$ 
6:   end for
7:   return  $\frac{1}{N} \sum_c^{C_{\mathcal{S}_i \times \mathcal{T}_j}} w_c * c.\text{getGLD}()$ 

```

7.4 Information entropy as a measure of the uncertainty in an spatio-temporal region

Now, what happen if we want to measure the uncertainty quantitatively? As we mention in subsection 2.2.3.2 the information entropy is useful in this context. The limitation we mention in that section is solved here, because we can use the different clusters we got in section 7.2 as the different outcomes of the system. The equation 2.2 can be rewritten as follow:

$$H(s, t) = - \sum_{c=1}^C p_c(s, t) \log p_c(s, t) \quad (7.12)$$

where c represent a particular cluster of the total number of clusters C , and $p_c(s, t)$ represent the probability of occurrence of the cluster c in the spatio-temporal region (s, t) .

Algorithm 4 Information Entropy in a region $(\mathcal{S}_i \times \mathcal{T}_j)$

```

1: function GLDMIXTURE( $\mathcal{S}_i \times \mathcal{T}_j, C_{\mathcal{S}_i \times \mathcal{T}_j}$ )
2:   for each  $p_i$  in  $(\mathcal{S}_i \times \mathcal{T}_j)$  do
3:      $c \leftarrow \text{cluster}(p_i)$ 
4:      $w_c = w_c + 1$ 
5:      $N = N + 1$ 
6:   end for
7:    $p_c(s, t) = \frac{w_c}{N}$ 
8:    $H(s, t) \leftarrow - \sum_{c=1}^C p_c(s, t) \log p_c(s, t)$ 
9:   return  $H(s, t)$ 

```

Algorithm 4 computes the Information Entropy in a region $C_{(\mathcal{S}_i \times \mathcal{T}_j)}$. In lines 2 to 7, we compute the probability of each cluster in the region. Using this probability we compute the information entropy $H(s, t)$, line 8, and finally we return the result in line 9.

7.5 Information entropy and model selection

7.6 UQ Proposed Dataflow

Summarizing, we proposed a workflow to quantify the uncertainty in large-scale spatio-temporal models, figure 32. The workflow is divided into three main steps, the fitting process, the clustering and the queries. We illustrate the use of the workflow with two queries, sections 7.3 and 7.4.

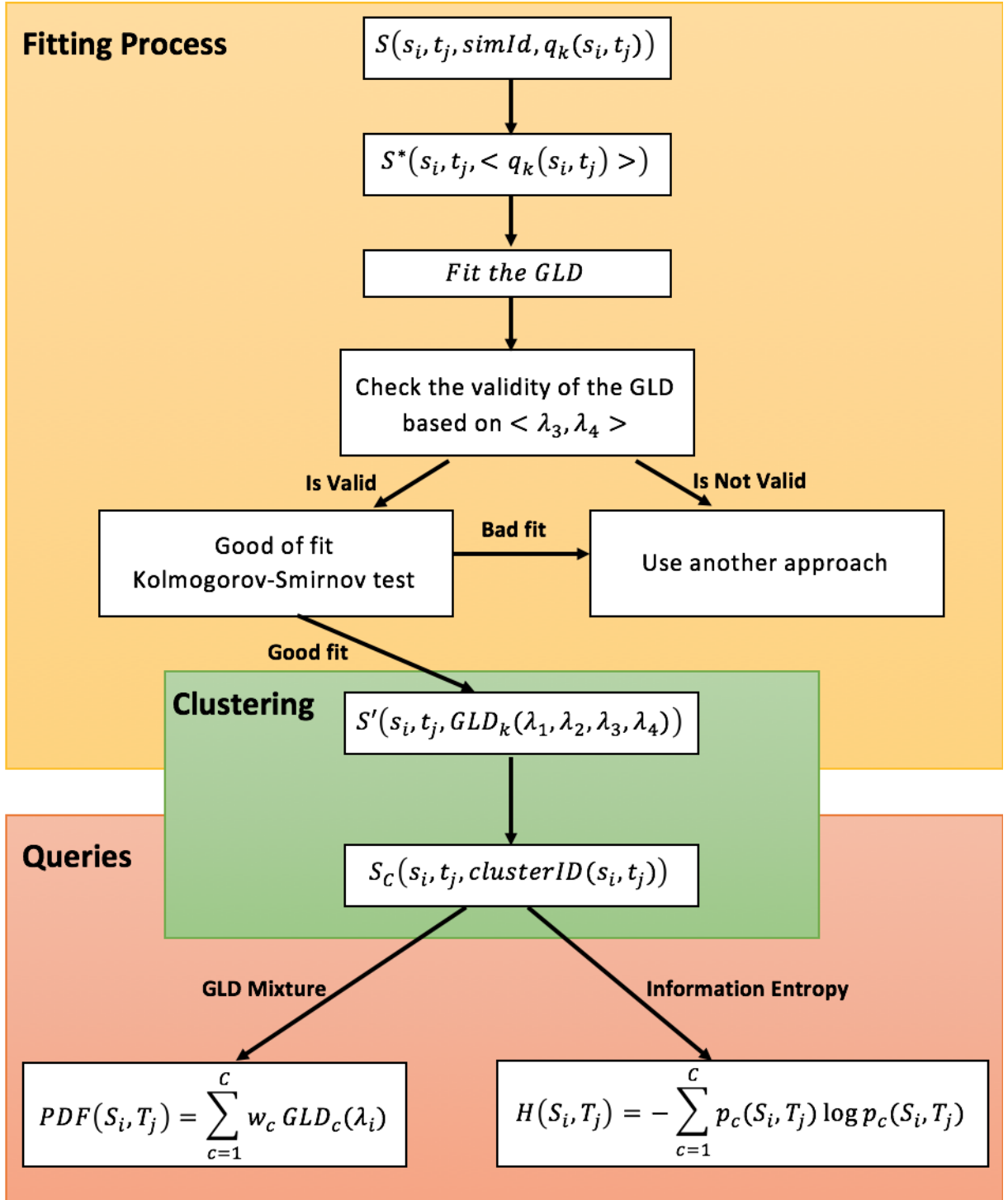


Figure 32 – Proposed workflow. The workflow was divided in three steps, (a) the fitting process, (b) the clustering of the GLDs and, (c) the queries over the results of the clustering process.

7.7 LaSST-UQ R package

7.8 Conclusions

8 Use Cases

In the present chapter we are going to test the UQMS in three different scenarios, spatial only domain, section 8.1, spatio-temporal domain, section 8.2, and finally a multidisciplinary system, section 8.3.

8.1 Case Study: Wave Propagation Problem

8.1.1 The Dataset

In the HPC4e benchmark, the models have been designed as a set of 16 layers with constant physical properties. The top layer delineates the topography and the other 15 delineate different layer interface surfaces or horizons. To generate a single cube with dimensions $250 \times 501 \times 501$ we can use the values provided in the benchmark. For example, to generate a cube in the $v_p(m/s)$ variable we can use the fixed values of Table 9.

The first slice of this cube is shown in Figure 33.

Layer	$v_p(m/s)$
1	1618.92
2	1684.08
3	1994.35
4	2209.71
5	2305.55
6	2360.95
7	2381.95
8	2223.41
9	2712.06
10	2532.22
11	2841.03
12	3169.31
13	3252.35
14	3642.28
15	3659.22
16	4000.00

Table 9 – Values of v_p used in the generation of a single velocity field cube.

Layer	PDF Family	Parameters
1	Gaussian	[1619, 711.2]
2	Gaussian	[3368, 711.2]
3	Gaussian	[8839, 711.2]
4	Gaussian	[7698, 301.5]
5	Lognormal	[7723, 294.7]
6	Lognormal	[7733, 292.2]
7	Lognormal	[7658, 312.1]
8	Lognormal	[3687, 368.7]
9	Exponential	[3949, 394.9]
10	Exponential	[5983, 711.2]
11	Exponential	[3520, 352.0]
12	Exponential	[3155, 315.5]
13	Uniform	[2541, 396.4]
14	Uniform	[2931, 435.3]
15	Uniform	[2948, 437.0]
16	Uniform	[3289, 471.1]

Table 10 – PDFs and its parameters used to sampling the v_p , to generate n velocity models.

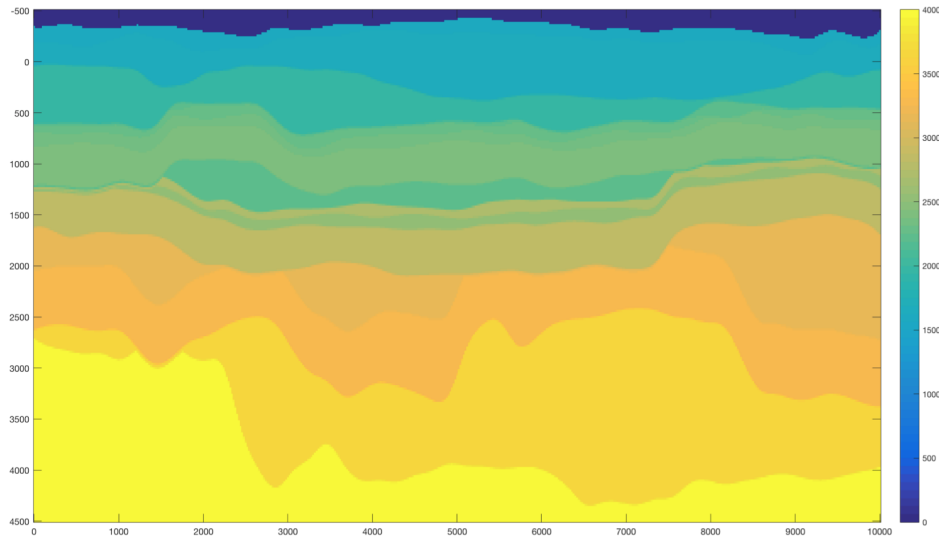


Figure 33 – One slice of the $250 \times 501 \times 501$ cube. In the slice we can distinguish between the different layers.

Now as our purpose is to study the uncertainty in the output as a result of the propagation of the input uncertainty throughout the model, we cannot use this benchmark as it is. We need the input, $v_p(m/s)$ in this case, to be uncertain. In order to achieve so, we compute $v_p(m/s)$ as a random variable with the *PDFs* shown in Table 10.

Then, using a Monte Carlo method we generate a sampling of 1000 realizations of the $v_p(m/s)$ variable, Figure 34; and using a Matlab script provided by the HPC4e

benchmark we simulate 1000 times, one for each realization, and generate 1000 cubes (230 GB) as an output. The resulting cubes are $250 \times 501 \times 501$ multi-dimensional arrays.

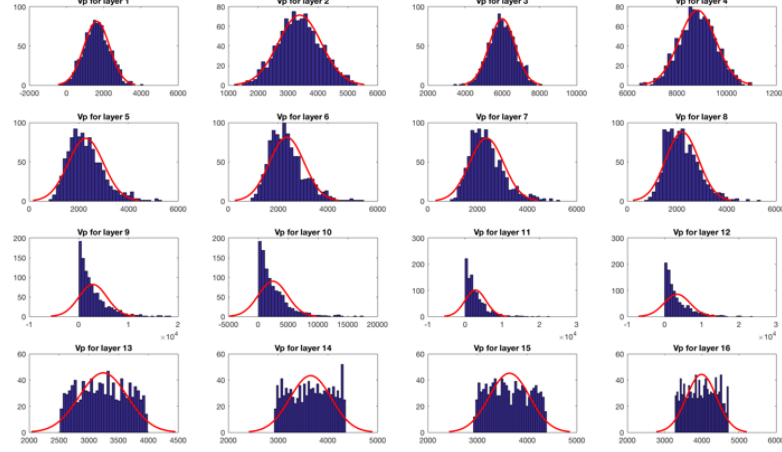


Figure 34 – Histograms of the 1000 samplings generated using Monte Carlo method and the PDFs reported in Table 10.

To simplify the computational process and visualize the results, we select the slice 200 to be used here, then we have 1000 realizations of a slice with size of 250×501 . The equation 7.2 can be simplified because we have two dimensions in space and don't have time domain, then our dataset can be represented as $S(x_i, y_j, simId, v_p(x_i, y_j))$. In this new representation (x_i, y_j) are the 2D coordinates and $v_p(x_i, y_j)$ is the velocity value at point (x_i, y_j) . $simId$ still represents the Id of the simulation and its range here is between 1 and 1000.

Now that we have an experimental dataset we can start to apply our workflow, step by step.

8.1.2 Fitting the GLD

The first step is to find the *GLD* that best fits the dataset at each spatial location. Running the algorithm proposed in Section 7.1.1 we get as a result a new 2D array:

$$S'(x_i, y_j, GLD(\lambda_1, \lambda_2, \lambda_3, \lambda_4)) \quad (8.1)$$

The raw data is reduced and our dataset is characterized by four lambda values at each spatial location. Now we need to assess the validity of the *GLDs* and how well they fit the dataset. Those analyses are described in sections 8.1.3 and 8.1.4.

8.1.3 GLD validity check

Once the algorithm to check the validity of the *GLD* is run on the experimental dataset, we obtain as a result that the *GLD* is valid in all the (x_i, y_j) space.

8.1.4 Quality of the fit

The next step is to check how good is the fit. To do this we use an algorithm that returns the D and p -value for the KS-test at each spatial location. As we show in figure 35, and remember that with a p -value > 0.05 we cannot reject the null hypothesis, we conclude that the fit of the GLD is acceptable in most cases. To be more exact, the p -value was greater than 0.05 in 82 % of the spatial locations, figure 36.

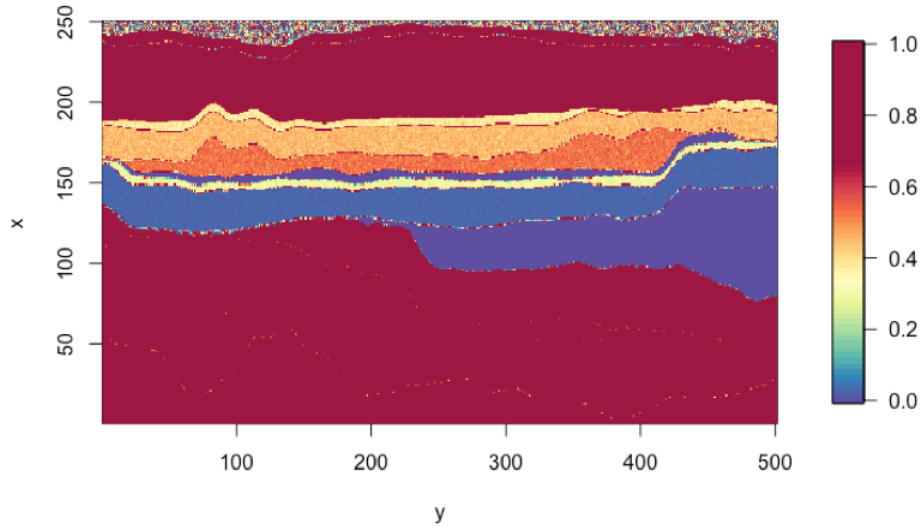


Figure 35 – Goodness of the fit based on the p -value returning by the KS-test. p -value > 0.05 represent a good fit of the GLD to the dataset at (x_i, y_j) .

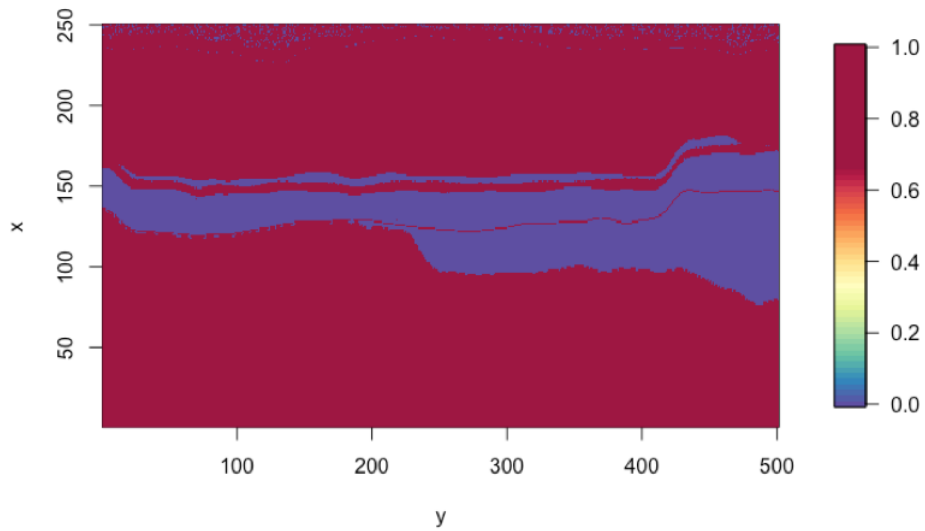


Figure 36 – The red color shows where the p -value was greater than 0.05.

If we consider the distance D , returned by the KS-test, the result is similar, figure 37. We can see a blue region that is common in figures 36 and 37. This region is where

the quality of the GLD fit is below a threshold. On those cases, some *GLD* extensions proposed in (KARIAN; DUDEWICZ, 2011) could be used.

As the main purpose of this paper is to demonstrate the utility of the use of the *GLD* in *UQ*, then we are not going to deep in other algorithms to solve this particular problem.

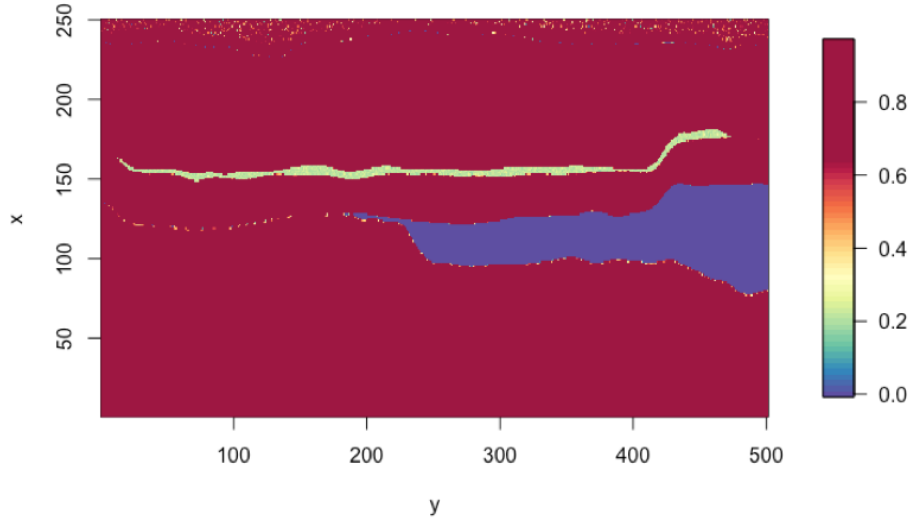


Figure 37 – Kolmogorov-Smirnoff Distance (D). The red regions represent where the GLD fits well.

8.1.5 Clustering

At this point we have our dataset characterized by the schema depicted by Equation 8.1, then using a clustering algorithm, such as k-means, we are going to group the GLDs based on its $(\lambda_2, \lambda_3, \lambda_4)$ values, as those are the values that describe the shape of the distribution at each point of the dataset.

In this paper we use k-means algorithm with $n = 10$, where n is the number of clusters to be made. This is an arbitrary value, we are investigating other algorithms as DBSCAN and what are the ϵ of this algorithm that warranty a good clusterization, but discussing alternative GLDs clustering algorithms is beyond the scope of this paper.

Once the clustering algorithm has been applied, a new dataset is produced, where for each spatial location we have a label that indicates the cluster the GLD at each position belongs to (see the schema at Equation 8.2), Figure 38. Note that, in Figure 38, the blue region corresponding to cluster 11 is not a cluster itself. It is rather the region where the *GLD* is not valid, see section 8.1.4.

$$S_C(x_i, y_j, clusterID, GLD_{x_i, y_j}) \quad (8.2)$$

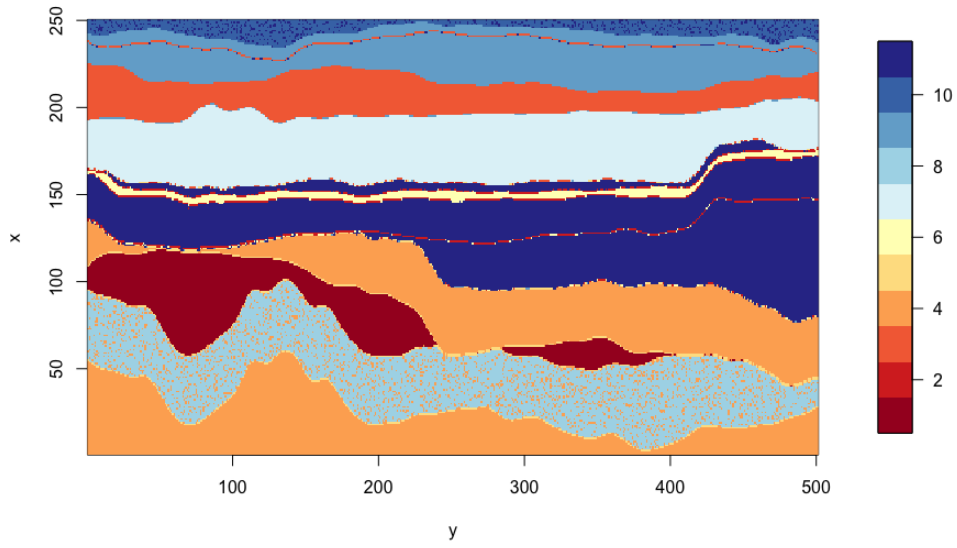


Figure 38 – Result of the clusterization using k-means with $n = 10$.

If we visually compare Figure 33 with Figure 38, we observe a close similarity between the two. It is clear that they can not be equal because we are talking about a slice of a deterministic model, and the result of making clusters on 1000 realizations of a stochastic model, but as the model used here is very linear, this is the result we expect.

Another interesting result is shown in Figure 39, where we plot the clusters in (λ_3, λ_4) space. As we mention in section ??, the shape of the *GLD* depends on the values of λ_3 and λ_4 . In this scenario, the expected result is that the members of the same cluster share similar values of λ_3 and λ_4 . This is exactly the result we can observe in Figure 39.

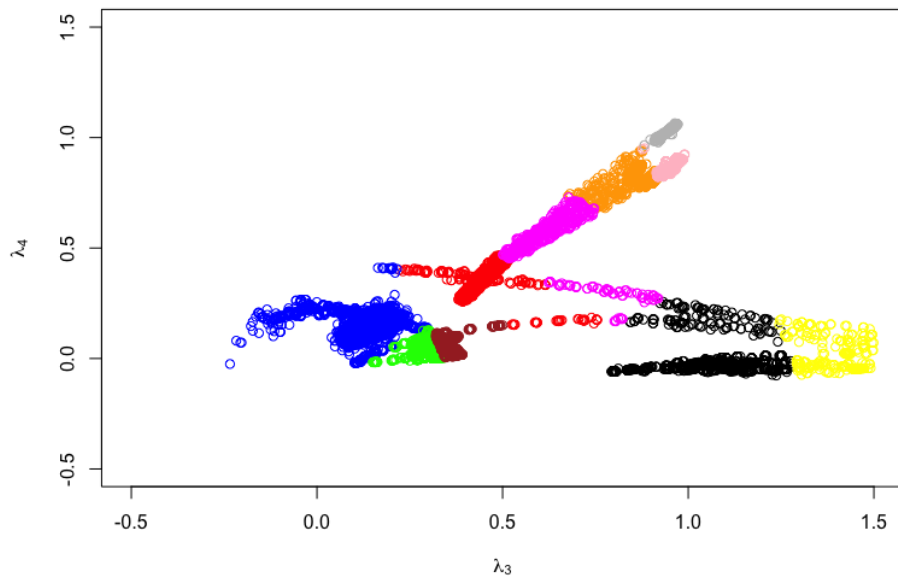


Figure 39 – Distribution of the clusters in the (λ_3, λ_4) space. The points that belongs to a same cluster are one near the others, as was expected.

To further corroborate this fact, in Figures 40 to 49 we show the *PDFs* of 60 members of the 10 clusters. Visually assessing the figures we have an idea of how similar are the shapes of the members of a same cluster and how dissimilar are the shapes of the members of different clusters. This suggests that our approach is valid. A product of these observations is that we can pick one member of each cluster (the centroid) as a representative of all the members of this cluster, Table 11. The selected member is going to be used to answer the queries in the next sections.

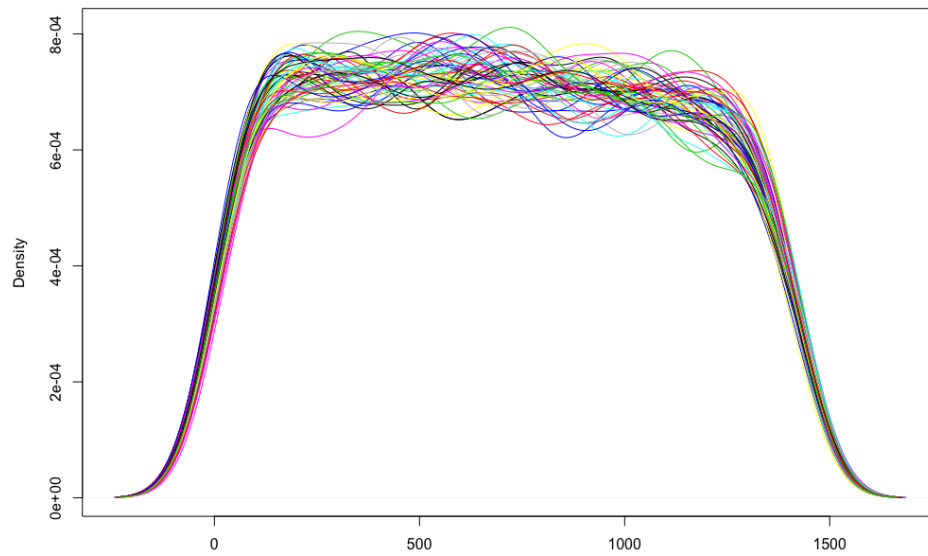


Figure 40 – *PDFs* of 60 members of the cluster 1.

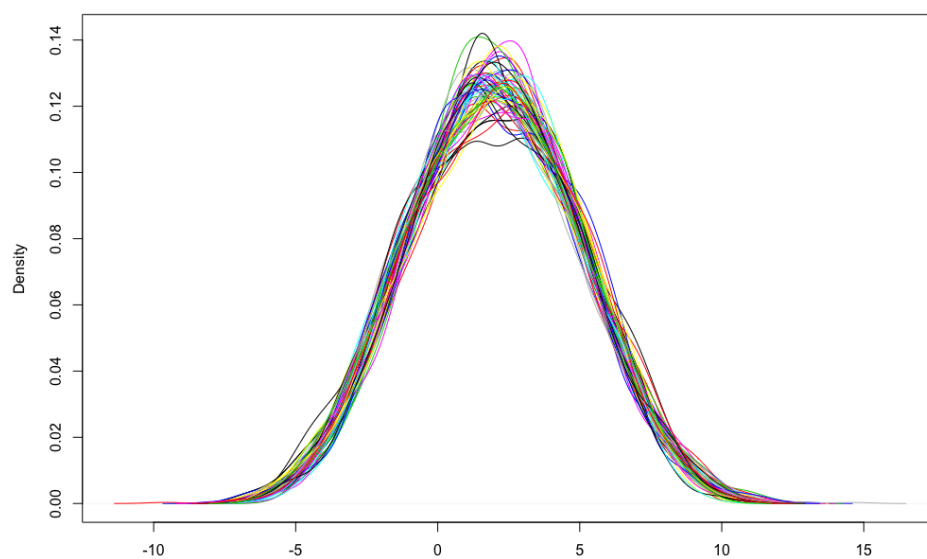


Figure 41 – *PDFs* of 60 members of the cluster 2.

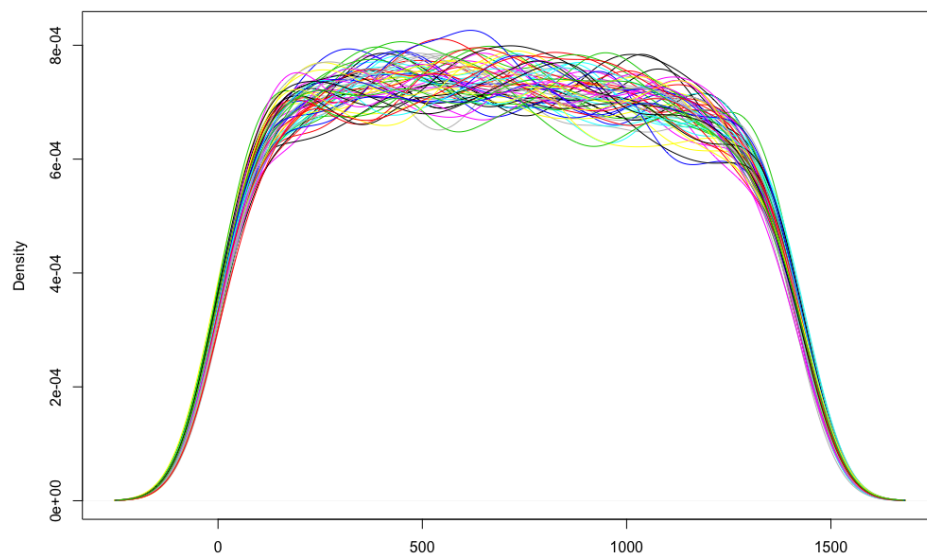


Figure 42 – *PDFs* of 60 members of the cluster 3.

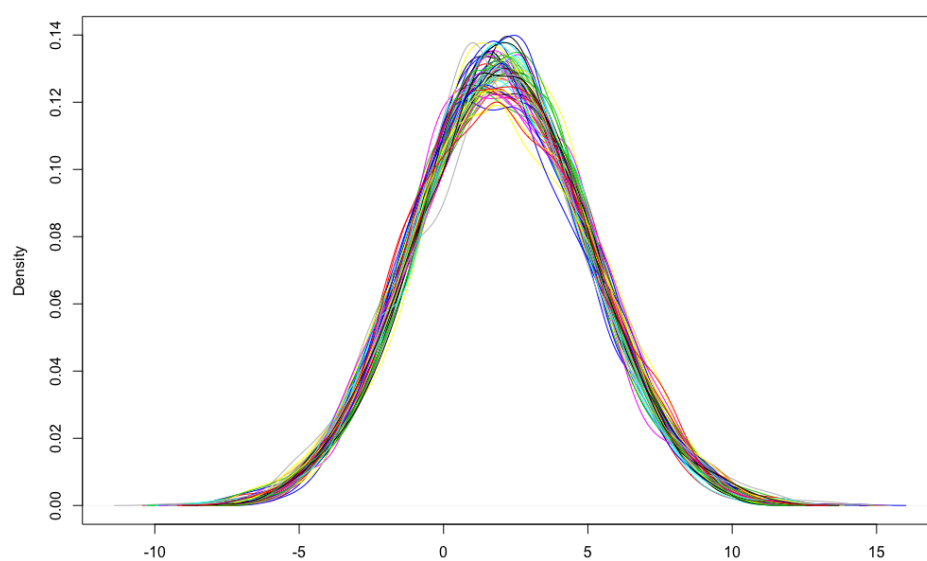


Figure 43 – *PDFs* of 60 members of the cluster 4.

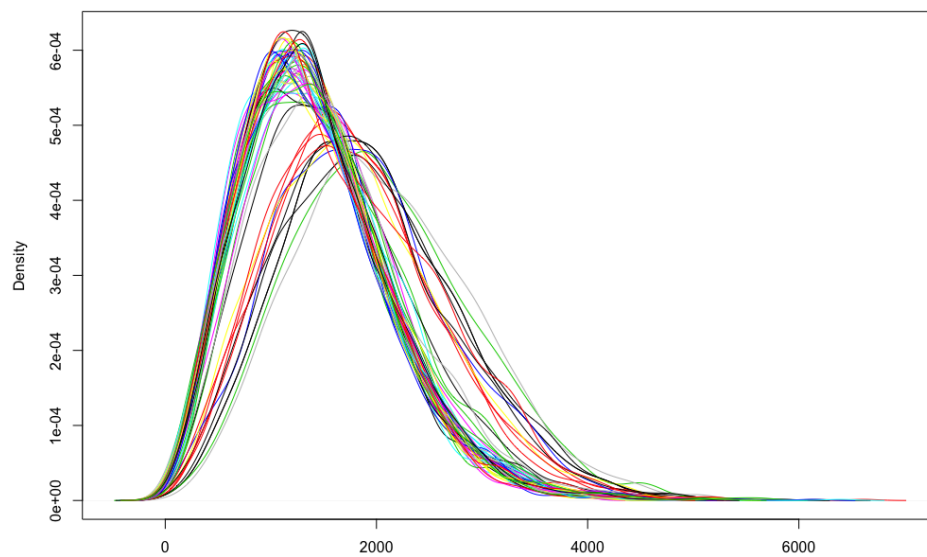


Figure 44 – *PDFs* of 60 members of the cluster 5.

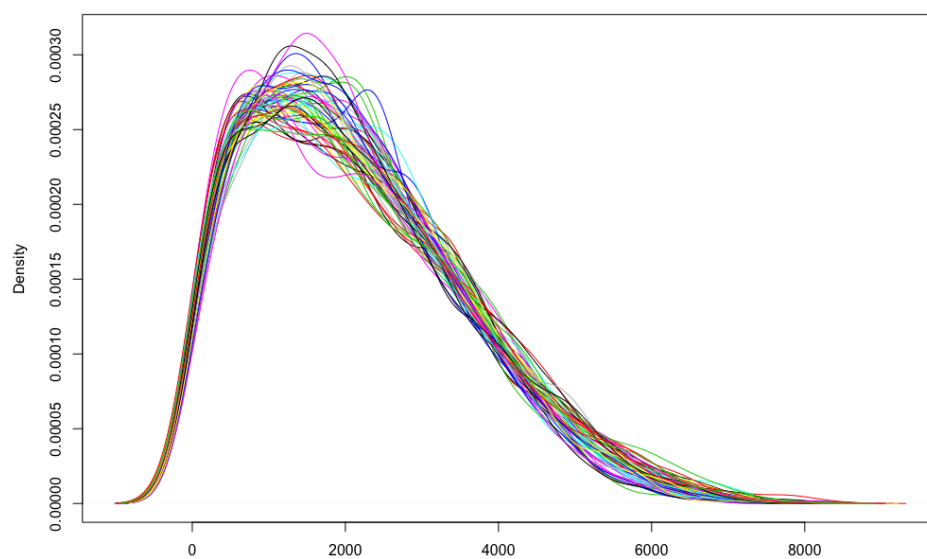


Figure 45 – *PDFs* of 60 members of the cluster 6.

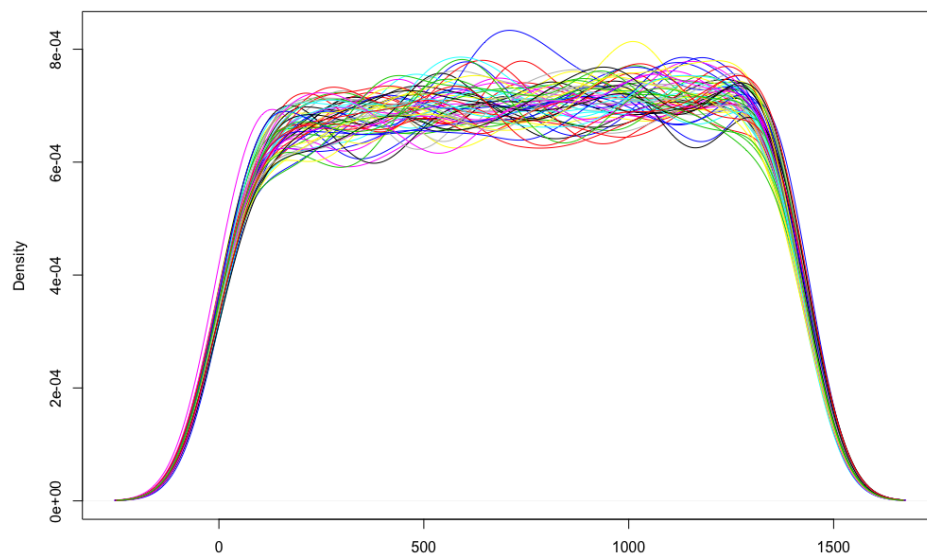


Figure 46 – *PDFs* of 60 members of the cluster 7.

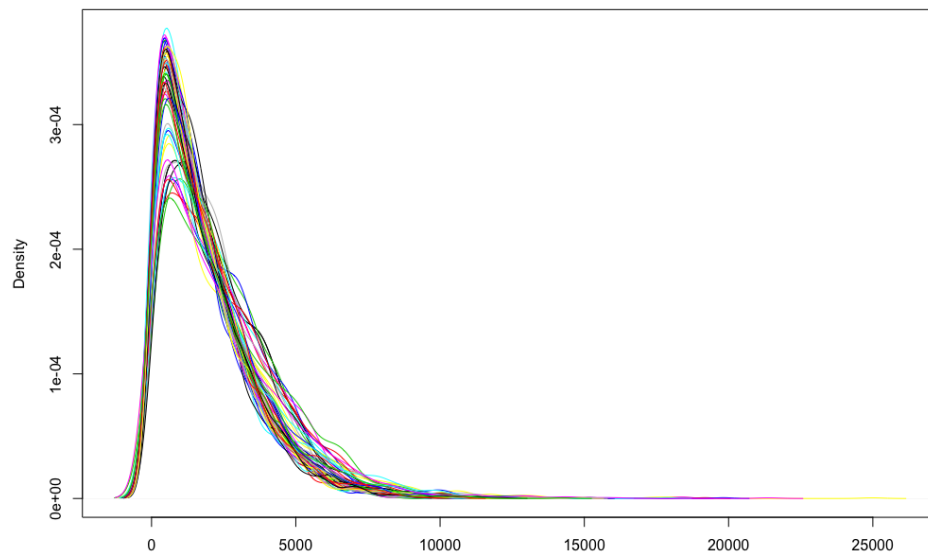


Figure 47 – *PDFs* of 60 members of the cluster 8.

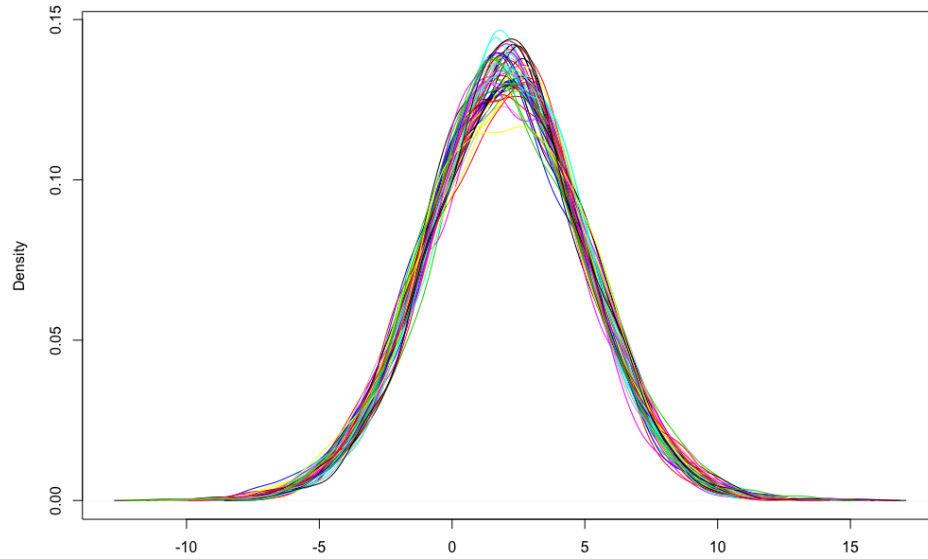
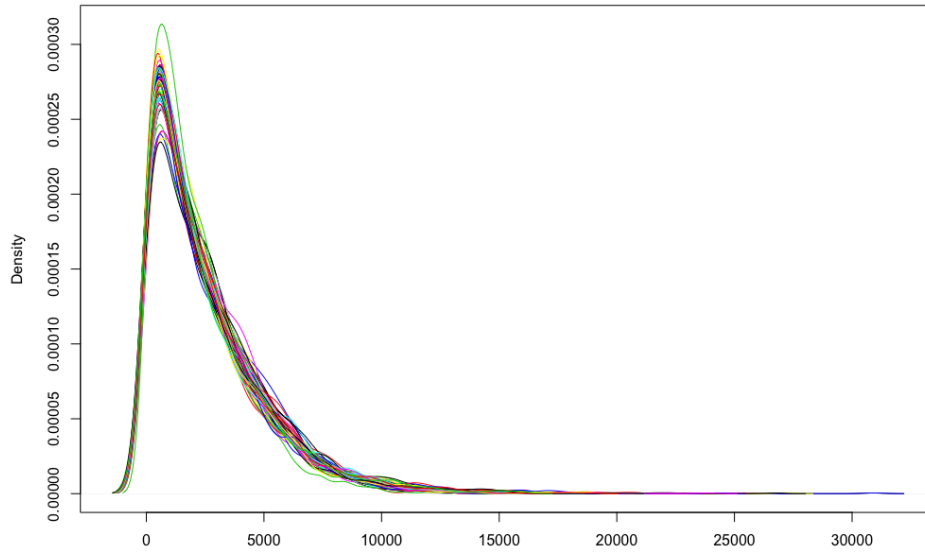


Figure 48 – *PDFs* of 60 members of the cluster 9.

Cluster	λ_2	λ_3	λ_4
1	0.0013937313	0.9585829	1.04696461
2	0.0005291388	1.1633978	-0.07162550
3	0.0020630696	0.1349486	0.17305941
4	0.0016238358	0.8653824	0.83857646
5	0.0027346929	0.5084664	0.39199164
6	0.0003894541	1.4076354	-0.01925743
7	0.0021972784	0.3253562	0.01493809
8	0.0015421749	0.9491101	0.86699555
9	0.0018672401	0.2176002	0.17862024
10	0.4856397733	0.1404140	0.14011298

Table 11 – Centers of the clusters.

Figure 49 – *PDFs* of 60 members of the cluster 10.

The 125250 points of the slice are distributed through the clusters following the histogram of the figure 50 and Table 12.

Cluster	No. of members
1	27217
2	15223
3	6749
4	3421
5	1353
6	25853
7	1374
8	18103
9	12051
10	13156

Table 12 – Distribution of the clusters.

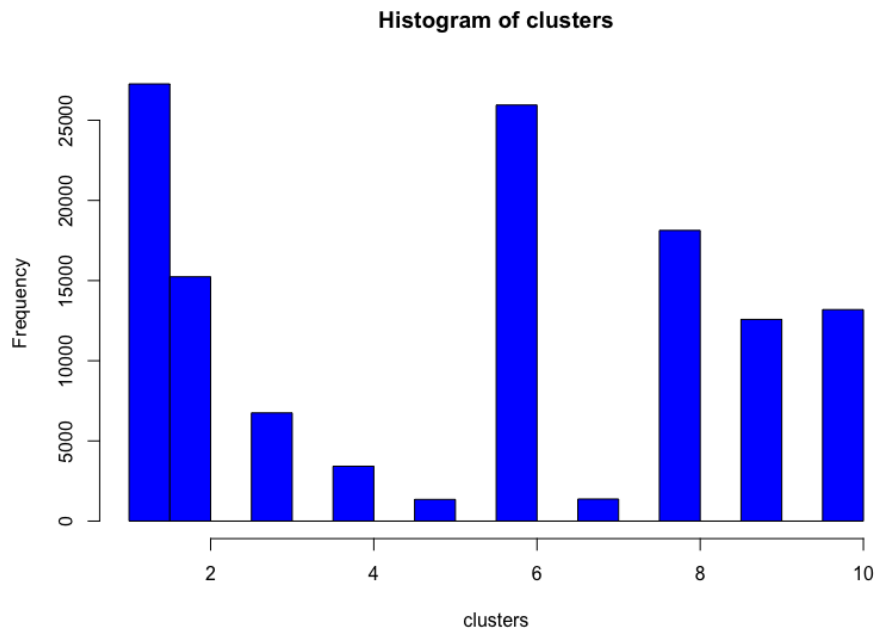


Figure 50 – Distribution of the clusters.

8.1.6 Spatio-temporal queries

At this point, the initial dataset is summarized as depicted by the schema in equation 8.2. It can be used to answer queries and to validate our approach, comparing the results with the raw data.

First of all we select four spatio-temporal regions of the dataset where the clusters suggest us different behaviors. The regions are shown in Figure 51 and the values of $[x_1, x_2], [y_1, y_2]$ that define the regions are shown in Table 13.

Region	x_1	x_2	y_1	y_2
Region 1	210	250	0	40
Region 2	150	250	50	150
Region 3	0	75	100	200
Region 4	0	250	300	400

Table 13 – Analysis Regions.

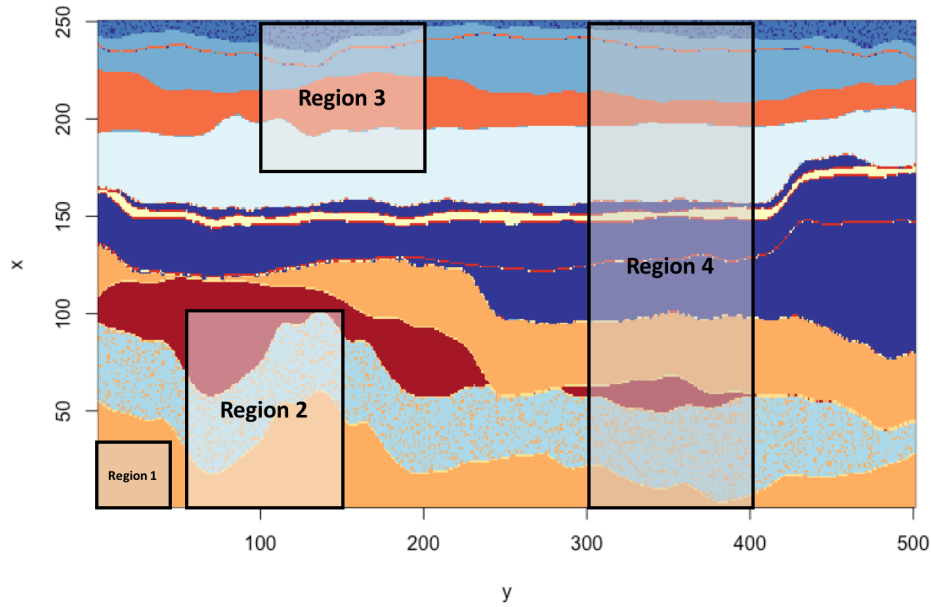


Figure 51 – Analysis Regions.

With these four regions we assess the adoption of the *GLD* mixture to obtain the *PDF* that characterizes the uncertainty in an specific region, section 8.1.6.1; and in section 8.1.6.2. We use the Information Entropy to assign a value that measures the uncertainty at each region. In section 8.1.6.1, we expect the *GLD* mixture to characterize well the raw data; and in 8.1.6.2 we hope that the information entropy is zero in region 1 and increases between regions 2, 3 and 4.

8.1.6.1 GLD mixture

The experiment here is to use the representative *GLDs* at each cluster and the weight associated to it in the region. Using these parameters we can build a *GLD mixture* that characterizes the uncertainty on that region. Here we use the algorithm described in section ??.

First of all we query the region to find the clusters represented inside it, and how are they distributed. Below we show the R codes to query the four regions. The retrieved results are shown in Table 14.

Cluster	Region 1	Region 2	Region 3	Region 4
1	0	2250	0	979
2	0	0	0	268
3	0	0	2596	1468
4	1640	4467	0	5173
5	0	149	0	269
6	0	0	0	416
7	0	0	1967	3920
8	0	3335	0	3432
9	0	0	1918	3280
10	0	0	901	583

Table 14 – Distribution of the clusters by regions.

Metrics	Region 1	Region 2	Region 3	Region 4
p-value	0.73	0.56	0.34	0.08

Table 15 – p-values by regions.

```

> clRegion1 = clByRegion(210, 250, 0, 40)
> clRegion2 = clByRegion(150, 250, 50, 150)
> clRegion3 = clByRegion(0, 75, 100, 200)
> clRegion4 = clByRegion(0, 250, 300, 400)

```

If we divide the columns of Table 14 by the sum of the elements of each column we get the weight needed to formulate the *mixed GLDs*. It is clear that the *GLD* in region 1 is represented by the *GLD* of cluster 4. On the other 3 cases we get:

$$\begin{aligned}
GLD_{region1} &= GLD_{c4} \\
GLD_{region2} &= 0.22GLD_{c1} + 0.44GLD_{c4} + 0.014GLD_{c5} \\
&\quad + 0.33GLD_{c8} \\
GLD_{region3} &= 0.34GLD_{c3} + 0.26GLD_{c7} + 0.25GLD_{c9} \\
&\quad + 0.12GLD_{c10} \\
GLD_{region4} &= 0.22GLD_{c1} + 0.44GLD_{c4} + 0.014GLD_{c5} \\
&\quad + 0.33GLD_{c8}
\end{aligned}$$

Now we need to evaluate if the *mixture of GLDs* describes well the uncertainty in the regions. To do this we perform the same *ks-test* used to evaluate the goodness of the fit and described in Section 7.1.3.

Based on the *p-value*, Table 15, we can conclude that in all 4 regions the *mixture of GLDs* is a good fit to the raw data.

entropy	Region 1	Region 2	Region 3	Region 4
value	0	1.122243	1.41166	2.024246

Table 16 – Information Entropy by regions.

8.1.6.2 Information Entropy

Now we are going to evaluate what happens with the information entropy. Based on the distribution of clusters inside the regions, table 14; we can compute the entropy. In this case we use an R function called *entropy*, implemented in the r-package of the same name (HAUSSER; STRIMMER, 2008).

As we expect, Table 16, the entropy in region 1 is zero, because the region contains only members of the cluster 4. On the other regions the entropy increases from region 2 to region 4, as we expected.

It is clear that the information entropy is a very good and simple measure of the uncertainty, and here it is demonstrated its utility combined with the *GLD*.

The first one is a geophysical tests for wave propagation problems

As a first case study we use the “HPC4E Seismic Test Suite”, a collection of four 3D models and sixteen associated tests that can be downloaded freely at the project’s website (<https://hpc4e.eu/downloads/datasets-and-software>). The models include simple cases that can be used in the development stage of any geophysical imaging practitioner (developer, tester ...) as well as extremely large cases that can only be solved in a reasonable time using ExaFLOPS supercomputers. The models are generated to the required size by means of a Matlab/Octave script and hence can be used by users of any OS or computing platform. The tests can be used to benchmark and compare the capabilities of different and innovative seismic modelling approaches, hence simplifying the task of assessing the algorithmic and computational advantages that they pose.

In our case, we are going to use the “HPC4E Seismic Test Suite” as a case study of the proposed UQMS. As we mention in the introduction of this chapter this model is a spatial only domain problem, because we are going to consider a multidimensional array as an Input and a multidimensional array as an output, but of them time independent.

8.1.7 Mathematical Formulation

8.1.8 Model and Dataset Description

The models have been designed as a set of 16 layers with constant physical properties. The top layer delineates the topography and the other 15 different layer interface surfaces or horizons. In the following, an interface horizon is associated with properties that apply

Table 17 – Layer constant properties and their depth range. “Star” layers are only used in the flat case, in substitution of their non-star equivalents

Layer Id	Vp (m/s)	Vs (m/s)	Density (Kg/m3)	Max. depth (m)	Min. depth (m)
1	1618.92	500.00	1966.38	-135.55	-476.35
2	1684.08	765.49	1985.88	41.50	-394.90
3					
4					
5					
6					
7					
8					
9					
10					
11					
12					
13					
14					
15					
16					
2*					
3*					

to the layer that exists between itself and the immediately next layer horizon. The model covers an area of 10 x 10 x 5 km, with maximum topography at about 500 m and maximum depth at about 4500 m. The layer horizons have been sampled very finely with 1.6667 m spacing so that a highly accurate representation can be honored at high frequencies. For simulation schemes based on unstructured grids, the layer horizons can be used easily to constrain model blocks. For simulation schemes based upon Cartesian grids, a simple script is provided that can generate 3D grids for any desired spatial sampling. Table 17 shows the properties of each of the layers included in the models.

8.1.9 Adding uncertainty into the model

The “HPC4E Seismic Test Suite” does not provide uncertainty sources, because all the input parameters of the model have fixed values. Then, to the purpose of our work we need to add some uncertainties into the inputs. Let’s suppose the variable V_p is uncertain. As this variable have 16 different values, one for each layer, we can consider it as a random vector, equation 8.3. We associate to each of the V_{p_i} a Normal distribution with μ_i equal to the value reported in Table 17 and $\sigma = 2$.

$$V_p = \langle V_{p_i}, \mathcal{N}(\mu_i, \sigma_i) \rangle \quad (8.3)$$

8.2 Case Study: Austin, queso library

8.3 Case Study: Multidisciplinary System (NASA)

8.4 Case Study: Spatio-temporal Nicholson-Bailey model

Este esta en el software uqlab, en la carpeta Doc Manuals

9 Conclusions and Future Works

9.1 Revisiting the Research Questions

9.2 Significance and Limitations

9.3 Open Problems and Future Work

9.4 Final Considerations

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Appendix

APPENDIX A – uqms R package

A.1 Título da seção

Aqui temos uma seção dentro do Apêndice.

APPENDIX B – Ideas

B.0.1 Variance, Information and Entropy

Variance.

Information and Entropy.

B.0.2 Information Gain, Distances and Divergences

B.1 Sensitivity Analysis

Sensitivity analysis is the systematic study of how model inputs—parameters, initial and boundary conditions—affect key model outputs. Depending on the application, one might use local derivatives or global descriptors such as Sobol’s functional decomposition or variance decomposition. Also, the needs of the application may range from simple ranking of the importance of inputs to a response surface model that predicts the output given the input settings. Such sensitivity studies are complicated by a number of factors, including the dimensionality of the input space, the complexity of the computational model, limited forward model runs due to the computational demands of the model, the availability of adjoint solvers or derivative information, stochastic simulation output, and high-dimensional output. Challenges in sensitivity analysis include dealing with these factors while addressing the needs of the application. (??)

$$E = mc^2 \tag{B.1}$$

APPENDIX C – Título do apêndice C

Annex

ANNEX A – Título do anexo A

A.1 Título da seção

Aqui temos uma seção dentro do Anexo.

ANNEX B – Título do anexo B

ANNEX C – Título do anexo C