## 1 Variance estimation for multi-stage sampling

Let U be a finite population and  $\mathcal{P}(U)$  the power set of U. Let  $p: \mathcal{S} \longrightarrow (0,1]$  be a r-stage sampling design  $(r \geq 2)$ , where  $\mathcal{S} \subset \mathcal{P}(U)$  is the set of all admissible samples under the design p. We express the hierarchical structure of the population U, with respect to the r-stage design p, as follows:

$$U = \bigsqcup_{i \in U^{(1)}} U_i^{(2+)} = \bigsqcup_{i \in U^{(1)}} \bigsqcup_{a \in U_i^{(2)}} U_{ia}^{(3)} = \cdots = \bigsqcup_{i \in U^{(1)}} \cdots \bigsqcup_{b \in U_{i\cdots}^{(r-1)}} U_{i\cdots b}^{(r)}$$

$$(1.1)$$

where  $U^{(1)}$  is the set of all primary sampling units (PSU), and for each PSU  $i \in U^{(1)}$ ,  $U_i^{(2)}$  denotes the set of all secondary sampling units (SSU) contained in  $i \in U^{(1)}$ , and for each SSU  $a \in U_i^{(2)}$ ,  $U_{ia}^{(3)}$  denotes the set of all tertiary sampling units (TSU) contained in  $a \in U_i^{(2)}$ , and so on. Similarly, we express the hierarchical structure of every admissible sample  $s \in \mathcal{S}$  as follows:

$$s = \bigsqcup_{i \in s^{(1)}} s_i^{(2+)} = \bigsqcup_{i \in s^{(1)}} \bigsqcup_{a \in s_i^{(2)}} s_{ia}^{(3)} = \cdots$$
 (1.2)

Let  $y:U\longrightarrow \mathbb{R}$  be a population characteristic. Let T be its population total, i.e.

$$T = \sum_{u \in U} y_u = \sum_{i \in U^{(1)}} T_i = \sum_{i \in U^{(1)}} \sum_{a \in U_{ia}^{(2)}} T_{ia} = \cdots = \sum_{i \in U^{(1)}} \cdots \sum_{u \in U_{i...}^{(r)}} y_u$$
 (1.3)

#### Theorem 1.1

If  $\widehat{T}_i: \mathcal{S}_i^{(2+)} \longrightarrow \mathbb{R}$  is an unbiased estimator for  $T_i$ , for each PSU  $i \in U^{(1)}$ , then the random variable  $\widehat{T}: \mathcal{S} \longrightarrow \mathbb{R}$  defined as follows:

$$\widehat{T}(s) := \sum_{i \in s^{(1)}} \frac{\widehat{T}_i(s_i^{(2+)})}{\pi_i^{(1)}} \tag{1.4}$$

is a design-unbiased estimator for

$$T = \sum_{i \in U^{(1)}} T_i \tag{1.5}$$

If the r-stage sampling design has invariant and independent subsampling, then the design-variance of  $\widehat{T}$  is given by:

$$\operatorname{Var}\left[\widehat{T}\right] = \underbrace{\operatorname{Var}^{(1)}\left(E^{(2+)}(\widehat{T} \mid s^{(1)})\right)}_{\operatorname{V}_{\operatorname{PSU}}} + \underbrace{E^{(1)}\left(\operatorname{Var}^{(2+)}(\widehat{T} \mid s^{(1)})\right)}_{\operatorname{V}_{\operatorname{subsampling}}}, \tag{1.6}$$

where

$$\operatorname{Var}^{(1)}\left(E^{(2+)}(\widehat{T} \mid s^{(1)})\right) = \sum_{i \in U^{(1)}} \sum_{j \in U^{(1)}} \Delta_{ij}^{(1)} \frac{T_i}{\pi_i^{(1)}} \frac{T_j}{\pi_j^{(1)}}, \text{ and}$$

$$\tag{1.7}$$

$$E^{(1)}\left(\operatorname{Var}^{(2+)}(\widehat{T} \mid s^{(1)})\right) = \sum_{i \in U^{(1)}} \frac{V_i}{\pi_i^{(1)}}, \tag{1.8}$$

with

$$V_{i} := \operatorname{Var}^{(2+)} \left[ \widehat{T}_{i} \right] \quad \text{and} \quad \Delta_{ij}^{(1)} := \begin{cases} \pi_{i}^{(1)} \left( 1 - \pi_{i}^{(1)} \right), & \text{if } i = j \\ \pi_{ij}^{(1)} - \pi_{i}^{(1)} \pi_{j}^{(1)}, & \text{if } i \neq j \end{cases}$$

$$(1.9)$$

February 4, 2015

Furthermore, if  $\widehat{V}_i: \mathcal{S}_i^{(2+)} \longrightarrow \mathbb{R}$  is an unbiased estimator for  $V_i:=\operatorname{Var}^{(2+)}\left[\widehat{T}_i\right]$ , and  $\pi_i^{(1)}>0$ ,  $\pi_{ij}^{(1)}>0$  for any PSUs  $i,j\in U^{(1)}$ , then

$$\widehat{\text{Var}}\Big[\widehat{T}\Big](s) := \sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{\widehat{T}_i(s_i^{(2+)})}{\pi_i^{(1)}} \cdot \frac{\widehat{T}_j(s_j^{(2+)})}{\pi_j^{(1)}} + \sum_{i \in s^{(1)}} \frac{\widehat{V}_i(s_i^{(2+)})}{\pi_i^{(1)}}, \tag{1.10}$$

$$\widehat{\text{Var}}^{(1)} \Big[ \widehat{T} \Big] (s) := \sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{\widehat{T}_i(s_i^{(2+)})}{\pi_i^{(1)}} \cdot \frac{\widehat{T}_j(s_j^{(2+)})}{\pi_j^{(1)}} + \sum_{i \in s^{(1)}} \left( 1 - \frac{1}{\pi_i^{(1)}} \right) \frac{\widehat{V}_i(s_i^{(2+)})}{\pi_i^{(1)}}$$

$$(1.11)$$

$$= \sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{\widehat{T}_i(s_i^{(2+)})}{\pi_i^{(1)}} \cdot \frac{\widehat{T}_j(s_j^{(2+)})}{\pi_j^{(1)}} + \sum_{i \in s^{(1)}} \frac{\widehat{V}_i(s_i^{(2+)})}{\pi_i^{(1)}} - \sum_{i \in s^{(1)}} \frac{\widehat{V}_i(s_i^{(2+)})}{\left(\pi_i^{(1)}\right)^2}$$
(1.12)

$$\widehat{\text{Var}}^{(2+)} \Big[ \widehat{T} \Big] (s) := \sum_{i \in s^{(1)}} \frac{\widehat{V}_i(s_i^{(2+)})}{\left(\pi_i^{(1)}\right)^2}$$
(1.13)

 $are \ unbiased \ estimators \ for \ Var\Big[\ \widehat{T}\ \Big], \ \ V_{PSU} := Var^{(1)}\Big(E^{(2+)}(\ \widehat{T}\ \Big|\ s^{(1)})\Big), \ and \ \ V_{subsampling} := E^{(1)}\Big(Var^{(2+)}(\ \widehat{T}\ \Big|\ s^{(1)})\Big), \ respectively.$ 

#### Corollary 1.2

$$\widehat{\operatorname{Var}}^{(1)} \left[ \widehat{T} \right] (s) = \widehat{\operatorname{Var}} \left[ \widehat{T} \right] (s) - \widehat{\operatorname{Var}}^{(2+)} \left[ \widehat{T} \right] (s)$$
(1.14)

Proof of Theorem 1.1

$$\operatorname{Var}^{(1)} \left[ E^{(2+)} \left[ \ \widehat{T} \ \middle| \ s^{(1)} \ \right] \right] = \operatorname{Var}^{(1)} \left[ E^{(2+)} \left[ \sum_{i \in s^{(1)}} \frac{\widehat{T}_i(s_i^{(2+)})}{\pi_i^{(1)}} \ \middle| \ s^{(1)} \ \right] \right]$$

$$= \operatorname{Var}^{(1)} \left[ \sum_{i \in s^{(1)}} \frac{E^{(2+)} \left[ \ \widehat{T}_i(s_i^{(2+)}) \ \right]}{\pi_i^{(1)}} \right]$$

$$= \operatorname{Var}^{(1)} \left[ \sum_{i \in s^{(1)}} \frac{T_i}{\pi_i^{(1)}} \right]$$

$$= \sum_{i \in U^{(1)}} \sum_{j \in U^{(1)}} \Delta_{ij}^{(1)} \frac{T_i}{\pi_i^{(1)}} \frac{T_j}{\pi_j^{(1)}}$$

$$E^{(1)} \left[ \operatorname{Var}^{(2+)} \left[ \ \widehat{T} \ \middle| \ s^{(1)} \ \right] \right] = E^{(1)} \left[ \operatorname{Var}^{(2+)} \left[ \sum_{i \in s^{(1)}} \frac{\widehat{T}_i(s_i^{(2+)})}{\pi_i^{(1)}} \ \middle| \ s^{(1)} \ \right] \right]$$

$$= E^{(1)} \left[ \sum_{i \in s^{(1)}} \operatorname{Var}^{(2+)} \left[ \frac{\widehat{T}_i(s_i^{(2+)})}{\pi_i^{(1)}} \right] \right]$$

$$= E^{(1)} \left[ \sum_{i \in s^{(1)}} \frac{\operatorname{Var}^{(2+)} \left[ \widehat{T}_i(s_i^{(2+)}) \right]}{\left(\pi_i^{(1)}\right)^2} \right]$$

$$= E^{(1)} \left[ \sum_{i \in s^{(1)}} \frac{V_i/\pi_i^{(1)}}{\pi_i^{(1)}} \right]$$

$$= \sum_{i \in U^{(1)}} \frac{V_i}{\pi_i^{(1)}}$$

$$E\left(\widehat{\operatorname{Var}}^{(2+)}(\widehat{T})\right) = E\left(\sum_{i \in s^{(1)}} \frac{\widehat{V}_{i}(s_{i}^{(2+)})}{\left(\pi_{i}^{(1)}\right)^{2}}\right) = E^{(1)}\left(E^{(2+)}\left(\sum_{i \in s^{(1)}} \frac{\widehat{V}_{i}(s_{i}^{(2+)})/\pi_{i}^{(1)}}{\pi_{i}^{(1)}} \middle| s^{(1)}\right)\right)$$

$$= E^{(1)}\left(\sum_{i \in s^{(1)}} \frac{E^{(2+)}\left[\widehat{V}_{i}(s_{i}^{(2+)})\middle| s^{(1)}\right]/\pi_{i}^{(1)}}{\pi_{i}^{(1)}}\right) = E^{(1)}\left(\sum_{i \in s^{(1)}} \frac{V_{i}/\pi_{i}^{(1)}}{\pi_{i}^{(1)}}\right) = \sum_{i \in U^{(1)}} \frac{V_{i}}{\pi_{i}^{(1)}}$$

$$= E^{(1)}\left(\operatorname{Var}^{(2+)}(\widehat{T}\middle| s^{(1)})\right) = \operatorname{V}_{\mathrm{PSU}}$$

Similarly,

$$E\left(\sum_{i \in s^{(1)}} \frac{\widehat{V}_i(s_i^{(2+)})}{\pi_i^{(1)}}\right) = E^{(1)}\left(E^{(2+)}\left(\sum_{i \in s^{(1)}} \frac{\widehat{V}_i(s_i^{(2+)})}{\pi_i^{(1)}} \middle| s^{(1)}\right)\right) = E^{(1)}\left(\sum_{i \in s^{(1)}} \frac{E^{(2+)}\left[\widehat{V}_i(s_i^{(2+)}) \middle| s^{(1)}\right]}{\pi_i^{(1)}}\right)$$

$$= E^{(1)}\left(\sum_{i \in s^{(1)}} \frac{V_i}{\pi_i^{(1)}}\right) = \sum_{i \in U^{(1)}} V_i$$

Next, observe that

$$E\left[\sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{\widehat{T}_{i}(s_{i}^{(2+)})}{\pi_{i}^{(1)}} \cdot \frac{\widehat{T}_{j}(s_{j}^{(2+)})}{\pi_{j}^{(1)}}\right] = E^{(1)} \left(E^{(2+)} \left(\sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{\widehat{T}_{i}(s_{i}^{(2+)})}{\pi_{i}^{(1)}} \cdot \frac{\widehat{T}_{j}(s_{j}^{(2+)})}{\pi_{j}^{(1)}} \right| s^{(1)}\right)\right)$$

$$= E^{(1)} \left(\sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{E^{(2+)} \left[\widehat{T}_{i}(s_{i}^{(2+)}) \cdot \widehat{T}_{j}(s_{j}^{(2+)}) \mid s^{(1)}\right]}{\pi_{i}^{(1)} \pi_{j}^{(1)}}\right)\right)$$

Now, observe (the key technical observation) that

$$E^{(2+)} \left[ |\widehat{T}_i(s_i^{(2+)}) \cdot \widehat{T}_j(s_j^{(2+)})| | s^{(1)}| \right] = E^{(2+)} \left[ |\widehat{T}_i(s_i^{(2+)}) \cdot \widehat{T}_j(s_j^{(2+)})| \right] = \begin{cases} |\operatorname{Var}^{(2+)}(\widehat{T}_i) + E^{(2+)}(\widehat{T}_i)| + E^{(2+)}(\widehat{T}_i)^2, & \text{if } i = j, \\ E^{(2+)}(\widehat{T}_i) \cdot E^{(2+)}(\widehat{T}_i) + E^{(2+)}(\widehat{T}_i)^2, & \text{if } i \neq j \end{cases}$$

Hence.

$$E\left[\sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{\widehat{T}_{i}(s_{i}^{(2+)})}{\pi_{i}^{(1)}} \cdot \frac{\widehat{T}_{j}(s_{j}^{(2+)})}{\pi_{j}^{(1)}}\right] = E^{(1)}\left(\sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{E^{(2+)}\left[\widehat{T}_{i}(s_{i}^{(2+)}) \cdot \widehat{T}_{j}(s_{j}^{(2+)})\right] s^{(1)}}{\pi_{i}^{(1)}\pi_{j}^{(1)}}\right)$$

$$= E^{(1)}\left(\sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{T_{i}}{\pi_{i}^{(1)}} \cdot \frac{T_{j}}{\pi_{j}^{(1)}} + \sum_{i \in s^{(1)}} \frac{(1 - \pi_{i}^{(1)})V_{i}/\pi_{i}^{(1)}}{\pi_{i}^{(1)}}\right)$$

$$= \sum_{i \in U^{(1)}} \sum_{j \in U^{(1)}} \Delta_{ij}^{(1)} \cdot \frac{T_{i}}{\pi_{i}^{(1)}} \cdot \frac{T_{j}}{\pi_{j}^{(1)}} + \sum_{i \in U^{(1)}} \frac{(1 - \pi_{i}^{(1)})V_{i}/\pi_{i}^{(1)}}{\pi_{i}^{(1)}} \cdot V_{i}$$

$$= \sum_{i \in U^{(1)}} \sum_{j \in U^{(1)}} \Delta_{ij}^{(1)} \cdot \frac{T_{i}}{\pi_{i}^{(1)}} \cdot \frac{T_{j}}{\pi_{j}^{(1)}} + \sum_{i \in U^{(1)}} \frac{V_{i}}{\pi_{i}^{(1)}} - \sum_{i \in U^{(1)}} V_{i}$$

February 4, 2015

We may now establish that

$$E\left[\widehat{\operatorname{Var}}\left(\widehat{T}\right)\right] = E\left[\sum_{i \in s^{(1)}} \sum_{j \in s^{(1)}} \frac{\Delta_{ij}^{(1)}}{\pi_{ij}^{(1)}} \cdot \frac{\widehat{T}_{i}(s_{i}^{(2+)})}{\pi_{i}^{(1)}} \cdot \frac{\widehat{T}_{j}(s_{j}^{(2+)})}{\pi_{j}^{(1)}} + \sum_{i \in s^{(1)}} \frac{\widehat{V}_{i}(s_{i}^{(2+)})}{\pi_{i}^{(1)}}\right]$$

$$= \left\{\sum_{i \in U^{(1)}} \sum_{j \in U^{(1)}} \Delta_{ij}^{(1)} \cdot \frac{T_{i}}{\pi_{i}^{(1)}} \cdot \frac{T_{j}}{\pi_{j}^{(1)}} + \sum_{i \in U^{(1)}} \frac{V_{i}}{\pi_{i}^{(1)}} - \sum_{i \in U^{(1)}} V_{i}\right\} + \left\{\sum_{i \in U^{(1)}} V_{i}\right\}$$

$$= \sum_{i \in U^{(1)}} \sum_{j \in U^{(1)}} \Delta_{ij}^{(1)} \cdot \frac{T_{i}}{\pi_{i}^{(1)}} \cdot \frac{T_{j}}{\pi_{j}^{(1)}} + \sum_{i \in U^{(1)}} \frac{V_{i}}{\pi_{i}^{(1)}}$$

$$= \operatorname{Var}\left(\widehat{T}\right)$$

Lastly, note that

$$\widehat{\operatorname{Var}}^{(1)} \left\lceil \widehat{T} \right\rceil (s) \ = \ \widehat{\operatorname{Var}} \left\lceil \widehat{T} \right\rceil (s) \ - \ \widehat{\operatorname{Var}}^{(2+)} \left\lceil \widehat{T} \right\rceil (s)$$

Hence,

$$\begin{split} E \bigg[ \widehat{\text{Var}}^{(1)} \Big( \widehat{T} \Big) \bigg] &= E \bigg[ \widehat{\text{Var}} \Big( \widehat{T} \Big) \bigg] - E \bigg[ \widehat{\text{Var}}^{(2+)} \Big( \widehat{T} \Big) \bigg] \\ &= \text{Var} \bigg[ \widehat{T} \bigg] - E^{(1)} \Big( \text{Var}^{(2+)} (\widehat{T} \, \Big| \, s^{(1)} ) \Big) \\ &= \text{Var}^{(1)} \Big( E^{(2+)} (\widehat{T} \, \Big| \, s^{(1)} ) \Big) + E^{(1)} \Big( \text{Var}^{(2+)} (\widehat{T} \, \Big| \, s^{(1)} ) \Big) - E^{(1)} \Big( \text{Var}^{(2+)} (\widehat{T} \, \Big| \, s^{(1)} ) \Big) \\ &= \text{Var}^{(1)} \Big( E^{(2+)} (\widehat{T} \, \Big| \, s^{(1)} ) \Big) = \text{V}_{\text{subsampling}} \end{split}$$

# 2 Variance estimation for four-stage sampling, with SRSWOR at each stage

First, recall that for a simple random sampling without replacement (SRSWOR), with fixed sample size n from a population of size N, the first- and second-order selection probabilities are given by:

$$\pi_i = \frac{n}{N} \text{ and } \pi_{ij} = \frac{n(n-1)}{N(N-1)},$$
(2.1)

for any distinct units i, j in the population. The Horvitz-Thompson estimator of the population total of a population characteristic y is, by definition:

$$\widehat{T}_y^{\text{HT}}(s) := \frac{N}{n} \sum_{k \in s} y_k = w \cdot \sum_{k \in s} y_k, \quad \text{where } w := \frac{N}{n}.$$

$$(2.2)$$

The design variance of  $\widehat{T}_y^{\mathrm{HT}}$  is given by:

$$\operatorname{Var}\left[\widehat{T}_{y}^{\operatorname{HT}}\right] = \sum_{i \in U} \sum_{j \in U} \Delta_{ij} \frac{y_{i}}{\pi_{i}} \frac{y_{j}}{\pi_{j}} = \cdots = N^{2} \left(1 - \frac{n}{N}\right) \cdot \frac{1}{n} \left\{\frac{1}{N - 1} \sum_{k \in \underline{U}} \left(y_{k} - \overline{y}_{\underline{U}}\right)^{2}\right\}$$
(2.3)

$$= N^{2} \left( 1 - \frac{n}{N} \right) \cdot \frac{1}{n} \cdot \operatorname{SVar} \left( \left\{ y_{k} \right\}_{k \in \mathbf{U}} \right) = (nw)^{2} \left( 1 - \frac{1}{w} \right) \cdot \frac{1}{n} \cdot \operatorname{SVar} \left( \left\{ y_{k} \right\}_{k \in U} \right)$$
(2.4)

$$= nw(w-1) \cdot \text{SVar}(\{y_k\}_{k \in \mathbf{U}}) = N \cdot (w-1) \cdot \text{SVar}(\{y_k\}_{k \in \mathbf{U}}), \qquad (2.5)$$

where  $\overline{y}_U := \frac{1}{N} \sum_{k \in U} y_k$ . Recall also that a design-unbiased estimator of  $\widehat{T}_y^{\text{HT}}$  is given by:

$$\widehat{\operatorname{Var}}\left[\widehat{T}_{y}^{\operatorname{HT}}\right](s) = \sum_{i \in s} \sum_{j \in s} \frac{\Delta_{ij}}{\pi_{ij}} \frac{y_{i}}{\pi_{i}} \frac{y_{j}}{\pi_{j}} = N^{2} \left(1 - \frac{n}{N}\right) \cdot \frac{1}{n} \left\{\frac{1}{n-1} \sum_{k \in s} (y_{k} - \overline{y}_{s})^{2}\right\}$$

$$(2.6)$$

$$= N^{2} \left(1 - \frac{n}{N}\right) \cdot \frac{1}{n} \cdot \operatorname{SVar}(\{y_{k}\}_{k \in s})$$

$$(2.7)$$

$$= nw(w-1) \cdot \text{SVar}(\{y_k\}_{k \in \mathbf{s}}) = N \cdot (w-1) \cdot \text{SVar}(\{y_k\}_{k \in \mathbf{s}})$$
 (2.8)

With the above observations, applying Theorem 1.1 recursively to  $\widehat{\text{Var}}\Big[\widehat{T}\Big]$ ,  $\widehat{V}_i^{(2+)}$ ,  $\widehat{V}_{ia}^{(3+)}$ , and  $\widehat{V}_{iac}^{(4)}$  immediately yields the following:

### Corollary 2.1

For a four-stage sampling design with invariant and independent subsampling, where sampling random sampling without replacement (SRSWOR) is used at each stage, we have

$$\widehat{\text{Var}}\Big[\widehat{T}\Big](s) = N^{(1)} \left(w^{(1)} - 1\right) \text{SVar}\left(\left\{\widehat{T}_{i}\right\}_{i \in s^{(1)}}\right) + w^{(1)} \sum_{i \in s^{(1)}} \widehat{V}_{i}^{(2+)} 
\widehat{V}_{i}^{(2+)} = N_{i}^{(2)} \left(w_{i}^{(2)} - 1\right) \text{SVar}\left(\left\{\widehat{T}_{ia}\right\}_{a \in s_{i}^{(2)}}\right) + w_{i}^{(2)} \sum_{a \in s_{i}^{(2)}} \widehat{V}_{ia}^{(3+)} 
\widehat{V}_{ia}^{(3+)} = N_{ia}^{(3)} \left(w_{ia}^{(3)} - 1\right) \text{SVar}\left(\left\{\widehat{T}_{iac}\right\}_{c \in s_{ia}^{(3)}}\right) + w_{ia}^{(3)} \sum_{a \in s_{ia}^{(3)}} \widehat{V}_{iac}^{(4)} 
\widehat{V}_{iac}^{(4)} = N_{iac}^{(4)} \left(w_{iac}^{(4)} - 1\right) \text{SVar}\left(\left\{y_{u}\right\}_{u \in s^{(4)}}\right)$$
(2.9)

(2.11)

$$\widehat{\text{Var}}^{(1)} \left[ \widehat{T} \right] (s) = n^{(1)} w^{(1)} \left( w^{(1)} - 1 \right) \text{SVar} \left( \left\{ \widehat{T}_i \right\}_{i \in s^{(1)}} \right) + w^{(1)} \left( 1 - w^{(1)} \right) \sum_{i \in s^{(1)}} \widehat{V}_i^{(2+)}$$
(2.10)

$$\widehat{\text{Var}}^{(2+)} \left[ \widehat{T} \right] (s) = \left( w^{(1)} \right)^2 \sum_{i \in s^{(1)}} \widehat{V}_i^{(2+)}$$
(2.12)