

## 1 Equivalence of $(C[0, 1], \|\cdot\|_\infty)$ -valued random variables and stochastic processes indexed by $[0, 1]$ with state space $\mathbb{R}$ and continuous sample paths

**Proposition 1.1** (The “one-dimensional subsets” of  $C[0, 1]$  generate its Borel  $\sigma$ -algebra)

Let  $(C[0, 1], \|\cdot\|_\infty)$  be the metric space of continuous  $\mathbb{R}$ -valued functions defined on the closed unit interval equipped with the supremum norm. For each  $t \in [0, 1]$ , let  $\text{ev}_t : C[0, 1] \rightarrow \mathbb{R} : x \mapsto x(t)$ . Define:

$$\mathcal{S} := \left\{ \text{ev}_t^{-1}(H) \subset C[0, 1] \mid \begin{array}{l} t \in [0, 1] \\ H \in \mathcal{O} \end{array} \right\} \subset \mathcal{P}(C[0, 1]).$$

Then,  $\mathcal{S}$  generates the Borel  $\sigma$ -algebra  $\mathcal{B} := \mathcal{B}(C[0, 1], \|\cdot\|_\infty)$  of the metric space  $(C[0, 1], \|\cdot\|_\infty)$ ; in other words,

$$\sigma(\mathcal{S}) = \mathcal{B}.$$

**PROOF** First, note that  $\sigma(\mathcal{S}) \subset \mathcal{B}$ . Indeed, recall that, for each  $t \in [0, 1]$ ,  $\text{ev}_t : C[0, 1] \rightarrow \mathbb{R}$  is continuous, hence  $(\mathcal{B}, \mathcal{O})$ -measurable, by Corollary B.4. In particular,  $\text{ev}_t^{-1}(H) \in \mathcal{B}$ , for each  $t \in [0, 1]$  and  $H \in \mathcal{O}$ . Thus,  $\mathcal{S} \subset \mathcal{B}$ ; hence,  $\sigma(\mathcal{S}) \subset \mathcal{B}$ .

It remains to establish the reverse inclusion. To this end, first observe that, for each  $x \in C[0, 1]$  and each  $\varepsilon > 0$ , we have

$$\overline{B(x, \varepsilon)} = \bigcap_{r \in \mathbb{Q} \cap [0, 1]} \left\{ y \in C[0, 1] \mid |y(r) - x(r)| \leq \varepsilon \right\} = \bigcap_{r \in \mathbb{Q} \cap [0, 1]} \text{ev}_r^{-1}([x(r) - \varepsilon, x(r) + \varepsilon]),$$

which shows that  $\sigma(\mathcal{S})$  contains all the closed balls in  $C[0, 1]$ . On the other hand, recall that, in any metric space, every open ball can be expressed as a countable union of closed balls; indeed, for any  $y$  in the given metric space, and any  $\delta > 0$ , we have:

$$B(y, \delta) = \bigcup_{n \in \mathbb{N}} \overline{B\left(y, \delta - \frac{1}{n}\right)}.$$

We thus see that  $\sigma(\mathcal{S})$  contains all the open balls in  $C[0, 1]$ . By the separability of  $C[0, 1]$  and Theorem C.1, we see that every open subset of  $C[0, 1]$  can be expressed as a countable union of open balls. Hence,  $\sigma(\mathcal{S})$  in fact contains all the open subsets of  $C[0, 1]$ , which immediately yields  $\mathcal{B} \subset \sigma(\mathcal{S})$ . This proves  $\sigma(\mathcal{S}) = \mathcal{B}$ .  $\square$

### Theorem 1.2

Suppose:

- $(\Omega, \mathcal{A})$  is a measurable space.
- Let  $(C[0, 1], \|\cdot\|_\infty)$  denote the metric space of continuous  $\mathbb{R}$ -valued functions defined on the compact unit interval equipped with the supremum norm.
- Let  $\mathcal{B} := \mathcal{B}(C[0, 1], \|\cdot\|_\infty)$  denote the Borel  $\sigma$ -algebra of the metric space  $(C[0, 1], \|\cdot\|_\infty)$ .
- Let  $\mathcal{O}$  denote the Borel  $\sigma$ -algebra of  $\mathbb{R}$  (equipped with usual Euclidean metric).
- $X : \Omega \rightarrow C[0, 1]$  is a function with domain  $\Omega$  and codomain  $C[0, 1]$ , but otherwise arbitrary.
- For each  $t \in [0, 1]$ , let  $\text{ev}_t : C[0, 1] \rightarrow \mathbb{R} : x \mapsto x(t)$ .

# Donsker's Theorems (Functional Central Limit Theorems)

- For each  $t \in [0, 1]$ , let  $X_t := \text{ev}_t \circ X$ . In other words,  $X_t : \Omega \rightarrow \mathbb{R} : \omega \mapsto \text{ev}_t(X(\omega)) = X(\omega)(t)$ .

Then,  $X$  is  $(\mathcal{A}, \mathcal{B})$ -measurable if and only if, for each  $t \in [0, 1]$ ,  $X_t$  is  $(\mathcal{A}, \mathcal{O})$ -measurable.

PROOF

( $\Rightarrow$ )

It is trivial to see that, for each  $t \in [0, 1]$ ,  $\text{ev}_t : (C[0, 1], \|\cdot\|_\infty) \rightarrow (\mathbb{R}, |\cdot|) : x \mapsto x(t)$  is continuous. Recall that continuous maps are necessarily Borel measurable; see Corollary B.4. Hence,  $\text{ev}_t : (C[0, 1], \|\cdot\|_\infty) \rightarrow (\mathbb{R}, |\cdot|)$  is  $(\mathcal{B}, \mathcal{O})$ -measurable, for each  $t \in [0, 1]$ . Now, suppose  $X : \Omega \rightarrow C[0, 1]$  is  $(\mathcal{A}, \mathcal{B})$ -measurable. Then, for each  $t \in [0, 1]$ , the composition  $X_t := \text{ev}_t \circ X$  is  $(\mathcal{A}, \mathcal{O})$ -measurable, as required.

( $\Leftarrow$ )

Suppose that, for each  $t \in [0, 1]$ ,  $X_t := \text{ev}_t \circ X$  is  $(\mathcal{A}, \mathcal{O})$ -measurable. We seek to establish that  $X : (\Omega, \mathcal{A}) \rightarrow (C[0, 1], \mathcal{B})$  is  $(\mathcal{A}, \mathcal{B})$ -measurable. To this end, let

$$\mathcal{S} := \left\{ \text{ev}_t^{-1}(H) \subset C[0, 1] \mid \begin{array}{l} t \in [0, 1] \\ H \in \mathcal{O} \end{array} \right\} \subset \mathcal{P}(C[0, 1]).$$

Then, note that the  $(\mathcal{A}, \mathcal{B})$ -measurability of  $X$  follows immediately from Theorem B.3, Proposition 1.1, and the following

**Claim:**  $X^{-1}(\mathcal{S}) \subset \mathcal{A}$ .

Proof of Claim: Every set in  $\mathcal{S}$  has the form  $\text{ev}_t^{-1}(H)$ , for some  $t \in [0, 1]$  and some  $H \in \mathcal{O}$ . Note that

$$X^{-1}(\text{ev}_t^{-1}(H)) = (\text{ev}_t \circ X)^{-1}(H) = X_t^{-1}(H) \in \mathcal{A},$$

where the last containment follows immediately from the  $(\mathcal{A}, \mathcal{O})$ -measurability hypothesis on  $X_t$ , for each  $t \in [0, 1]$ . This shows that  $X^{-1}(\mathcal{S}) \subset \mathcal{A}$  and completes the proof of the Claim.

The proof of the Theorem is now complete. □

## Theorem 1.3

Suppose:

- $(\Omega, \mathcal{A}, \mu)$  is a probability space.
- Let  $(C[0, 1], \|\cdot\|_\infty)$  denote the metric space of continuous  $\mathbb{R}$ -valued functions defined on the closed unit interval equipped with the supremum norm.
- $X : \Omega \rightarrow C[0, 1]$  is a function with domain  $\Omega$  and codomain  $C[0, 1]$ , but otherwise arbitrary.
- For each  $t \in [0, 1]$ , let  $\text{ev}_t : C[0, 1] \rightarrow \mathbb{R} : x \mapsto x(t)$ .
- For each  $t \in [0, 1]$ , let  $X_t := \text{ev}_t \circ X$ . In other words,  $X_t : \Omega \rightarrow \mathbb{R} : \omega \mapsto \text{ev}_t(X(\omega)) = X(\omega)(t)$ .

Then, the following are equivalent:

- (i)  $X$  is a  $(C[0, 1], \|\cdot\|_\infty)$ -valued random variable.
- (ii) For each  $t \in [0, 1]$ ,  $X_t$  is an  $\mathbb{R}$ -valued random variable.
- (iii)  $\{X_t : \Omega \rightarrow \mathbb{R}\}_{t \in [0, 1]}$  is a stochastic process indexed by the closed unit interval defined on the probability space  $(\Omega, \mathcal{A}, \mu)$  with state space  $\mathbb{R}$  and continuous sample paths.

## 2 Donsker's Theorem for $(C[0, 1], \|\cdot\|_\infty)$

### Proposition 2.1

- Let  $\xi_1, \xi_2, \dots : \Omega \rightarrow \mathbb{R}$  be a sequence of independent and identically distributed  $\mathbb{R}$ -valued random variables defined on the probability space  $(\Omega, \mathcal{A}, \mu)$ , with expectation value zero and common finite variance  $\sigma^2 > 0$ .
- Define the random variables:

$$\begin{cases} S_0 & : \Omega \rightarrow \mathbb{R} : \omega \mapsto 0, & \text{and} \\ S_n & : \Omega \rightarrow \mathbb{R} : \omega \mapsto \sum_{i=1}^n \xi_i(\omega), & \text{for each } n \in \mathbb{N}. \end{cases}$$

- For each  $n \in \mathbb{N}$ , define  $X^{(n)} : \Omega \rightarrow C[0, 1]$  as follows:

$$X^{(n)}(\omega)(t) := \frac{1}{\sigma \cdot \sqrt{n}} \left\{ S_{i-1}(\omega) + n \left( t - \frac{i-1}{n} \right) \xi_i(\omega) \right\}, \text{ for each } \omega \in \Omega, t \in \left[ \frac{i-1}{n}, \frac{i}{n} \right], i = 1, 2, 3, \dots, n.$$

- For each  $n \in \mathbb{N}$  and each  $t \in [0, 1]$ , define  $X_t^{(n)} : \Omega \rightarrow \mathbb{R}$  as follows:

$$X_t^{(n)}(\omega) := X^{(n)}(\omega)(t), \text{ for each } \omega \in \Omega.$$

Then, the following statements are true:

- (i) For each  $\omega \in \Omega$  and each  $n \in \mathbb{N}$ ,

$$X^{(n)}(\omega) \left( \frac{i}{n} \right) = \frac{1}{\sigma \cdot \sqrt{n}} \cdot S_i(\omega), \text{ for } i = 0, 1, 2, \dots, n.$$

- (ii) For each  $\omega \in \Omega$  and each  $n \in \mathbb{N}$ ,

$$X^{(n)}(\omega)(t) \text{ is the linear interpolation from } \frac{1}{\sigma \cdot \sqrt{n}} S_{i-1}(\omega) \text{ to } \frac{1}{\sigma \cdot \sqrt{n}} S_i(\omega) \text{ over } t \in \left[ \frac{i-1}{n}, \frac{i}{n} \right],$$

where  $i = 1, 2, \dots, n$ .

- (iii) For any  $0 \leq t_0 < t_1 < t_2 < \dots < t_k \leq 1$ ,

$$\left( X_{t_1}^{(n)} - X_{t_0}^{(n)}, \dots, X_{t_k}^{(n)} - X_{t_{k-1}}^{(n)} \right) \xrightarrow{d} N \left( \mu = \mathbf{0}, \Sigma = \text{diag}(t_1 - t_0, \dots, t_k - t_{k-1}) \right), \text{ as } n \rightarrow \infty.$$

- (iv) For any  $0 \leq t_1, t_2, \dots, t_k \leq 1$ ,

$$\left( X_{t_1}^{(n)}, X_{t_2}^{(n)}, \dots, X_{t_k}^{(n)} \right) \xrightarrow{d} N \left( \mu = \mathbf{0}, \Sigma = \left[ \min\{t_i, t_j\} \right]_{1 \leq i, j \leq k} \right), \text{ as } n \rightarrow \infty.$$

PROOF

- (i) Obvious.

- (ii) Obvious.

# Donsker's Theorems (Functional Central Limit Theorems)

(iii) First, note that, for each  $\omega \in \Omega$ ,  $n \in \mathbb{N}$ , and  $t \in [0, 1]$ , we have

$$X_t^{(n)}(\omega) = \frac{1}{\sigma \cdot \sqrt{n}} \left\{ S_{[nt]}(\omega) + (nt - [nt]) \cdot \xi_{[nt]+1}(\omega) \right\},$$

where  $\lfloor \cdot \rfloor : \mathbb{R} \rightarrow \mathbb{Z}$ , defined by

$$\lfloor x \rfloor := \max \left\{ k \in \mathbb{Z} \mid k \leq x \right\}, \quad \text{for each } x \in \mathbb{R},$$

is the round-down function. We next state three Claims, whose proofs will be given below. We note that the desired conclusion follows readily from Claim 3 and the Cramér-Wold Theorem (Theorem 1.9(iii), p.56, [3]); hence the present proof is complete once we establish the three Claims below.

**Claim 1:** If  $\{a_n\}_{n \in \mathbb{N}}$  is a sequence of non-negative integers and  $\{b_n\}_{n \in \mathbb{N}} \subset \mathbb{N}$  a sequence of positive integers satisfying:

$$a_n < b_n, \text{ for sufficiently large } n \in \mathbb{N}, \quad \text{and} \quad \lim_{n \rightarrow \infty} \frac{b_n - a_n}{n} = c > 0,$$

then

$$\frac{1}{\sigma \cdot \sqrt{n}} \cdot \sum_{i=1+a_n}^{b_n} \xi_i \xrightarrow{d} \sqrt{c} \cdot Z, \quad \text{where } Z \sim N(0, 1).$$

**Claim 2:** For each fixed  $t \in [0, 1]$ ,

$$W(t)_n := \frac{1}{\sigma \cdot \sqrt{n}} \cdot (nt - [nt]) \cdot \xi_{[nt]+1} \xrightarrow{d} 0.$$

**Claim 3:** For  $0 \leq t_0 < t_1 < t_2 < \dots < t_k \leq 1$ , and arbitrary  $c_1, c_2, \dots, c_k \in \mathbb{R}$ ,

$$\sum_{i=1}^k c_i \left( X_{t_i}^{(n)} - X_{t_{i-1}}^{(n)} \right) \xrightarrow{d} N \left( 0, \sum_{i=1}^k c_i^2 (t_i - t_{i-1}) \right), \quad \text{as } n \rightarrow \infty.$$

Proof of Claim 1: Note that, for sufficiently large  $n \in \mathbb{N}$ , we may write

$$\frac{1}{\sigma \cdot \sqrt{n}} \cdot \sum_{i=1+a_n}^{b_n} \xi_i = \frac{\sqrt{b_n - a_n}}{\sqrt{n}} \cdot \left( \frac{1}{\sigma \cdot \sqrt{b_n - a_n}} \cdot \sum_{i=1+a_n}^{b_n} \xi_i \right).$$

Since, by hypothesis, that

$$\lim_{n \rightarrow \infty} \frac{b_n - a_n}{n} = c > 0,$$

Claim 1 follows by Slutsky's Theorem (Example 6, p.40, [2]), once we establish the following:

$$\frac{1}{\sigma \cdot \sqrt{b_n - a_n}} \cdot \sum_{i=1+a_n}^{b_n} \xi_i \xrightarrow{d} N(0, 1), \quad \text{as } n \rightarrow \infty.$$

# Donsker's Theorems (Functional Central Limit Theorems)

We establish the above convergence by invoking the Lindeberg Central Limit Theorem (Theorem 1.15, §1.5.5, p.67, [3]). In the present context, the Lindeberg Condition is the following:

$$\lim_{n \rightarrow \infty} \frac{1}{B_n^2} \cdot E \left[ \sum_{i=1+a_n}^{b_n} \xi_i^2 \cdot I_{\{|\xi_i| \geq \varepsilon S_n\}} \right] = 0, \quad \text{for each } \varepsilon > 0,$$

where

$$B_n^2 := \text{Var} \left[ \sum_{i=1+a_n}^{b_n} \xi_i \right] = (b_n - a_n) \sigma^2 > 0.$$

The last equality used the hypothesis that  $\xi_1, \xi_2, \dots$  are independent and identically distributed with common finite variance  $0 < \sigma^2 < \infty$ . Hence, for each  $\varepsilon > 0$ ,

$$\begin{aligned} \frac{1}{B_n^2} \cdot E \left[ \sum_{i=1+a_n}^{b_n} \xi_i^2 \cdot I_{\{|\xi_i| \geq \varepsilon B_n\}} \right] &= \frac{1}{(b_n - a_n) \sigma^2} \cdot (b_n - a_n) \cdot E \left[ \xi_1^2 \cdot I_{\{|\xi_1| \geq \varepsilon \sigma \sqrt{b_n - a_n}\}} \right] \\ &= \frac{1}{\sigma^2} \cdot E \left[ \xi_1^2 \cdot I_{\{|\xi_1| / \sigma \geq \sqrt{b_n - a_n}\}} \right] \rightarrow 0, \quad \text{as } n \rightarrow \infty, \end{aligned}$$

since  $\lim_{n \rightarrow \infty} \sqrt{b_n - a_n} = \infty$  and  $\sigma^2 = E[\xi_1^2]$  is finite. This verifies that the Lindeberg Condition indeed holds in the present context, and completes the proof of Claim 1.

Proof of Claim 2: First, note that  $E[W(t)_n] = 0$ . We now argue that  $W(t)_n \xrightarrow{p} 0$ . To this end, let  $\varepsilon > 0$  be given. Then,

$$\begin{aligned} \varepsilon^2 \cdot P(|W(t)_n| \geq \varepsilon) &\leq E[W(t)_n^2 \cdot I_{\{|W(t)_n| \geq \varepsilon\}}] \\ &\leq E[W(t)_n^2] = \text{Var}(W(t)_n) = \text{Var} \left[ \frac{1}{\sigma \cdot \sqrt{n}} \cdot (nt - \lfloor nt \rfloor) \cdot \xi_{\lfloor nt \rfloor + 1} \right] \\ &= \frac{1}{n \cdot \sigma^2} \cdot (nt - \lfloor nt \rfloor)^2 \cdot \text{Var}(\xi_{\lfloor nt \rfloor + 1}) = \frac{1}{n \cdot \sigma^2} \cdot (nt - \lfloor nt \rfloor)^2 \cdot \sigma^2 \\ &\leq \frac{1}{n}, \end{aligned}$$

which implies

$$\lim_{n \rightarrow \infty} P(|W(t)_n| \geq \varepsilon) = 0, \quad \text{for each } \varepsilon > 0,$$

i.e.  $W(t)_n \xrightarrow{p} 0$ , as  $n \rightarrow \infty$  (Definition 2, Chapter 1, [2]), which is equivalent to  $W(t)_n \xrightarrow{d} 0$ , as  $n \rightarrow \infty$  (by Theorem 1, Chapter 1 and Theorem 2, Chapter 2, [2]). This proves Claim 2.

Proof of Claim 3: Let  $0 \leq t_0 < t_1 < t_2 < \dots < t_k \leq 1$ , and  $c_1, c_2, \dots, c_k \in \mathbb{R}$  be arbitrary. Observe that:

$$\begin{aligned} &\sum_{i=1}^k c_i \left( X_{t_i}^{(n)} - X_{t_{i-1}}^{(n)} \right) \\ &= \sum_{i=1}^k \frac{c_i}{\sigma \cdot \sqrt{n}} \left\{ S_{\lfloor nt_i \rfloor} - S_{\lfloor nt_{i-1} \rfloor} \right\} + \sum_{i=1}^k \frac{c_i}{\sigma \cdot \sqrt{n}} \left\{ (nt_i - \lfloor nt_i \rfloor) \cdot \xi_{\lfloor nt_i \rfloor + 1} - (nt_{i-1} - \lfloor nt_{i-1} \rfloor) \cdot \xi_{\lfloor nt_{i-1} \rfloor + 1} \right\} \\ &= \sum_{i=1}^k \frac{c_i}{\sigma \cdot \sqrt{n}} \left\{ \sum_{j=1+\lfloor nt_{i-1} \rfloor}^{\lfloor nt_i \rfloor} \xi_j \right\} + \sum_{i=1}^k c_i \left\{ W(t_i)_n - W(t_{i-1})_n \right\} \\ &= \sum_{i=1}^k c_i Y_i^{(n)} + \sum_{i=1}^k c_i \left\{ W(t_i)_n - W(t_{i-1})_n \right\} \end{aligned}$$

# Donsker's Theorems (Functional Central Limit Theorems)

By Claim 2 and Slutsky's Theorem (Corollary, p.40, [2]),

$$\sum_{i=1}^k c_i \left\{ W(t_i)_n - W(t_{i-1})_n \right\} \xrightarrow{d} 0, \text{ as } n \rightarrow \infty. \quad (2.1)$$

Next, note that since  $\xi_1, \xi_2, \xi_3, \dots$  are independent, we see that, for each fixed  $n \in \mathbb{N}$ ,

$$Y_i^{(n)} := \frac{1}{\sigma \cdot \sqrt{n}} \cdot \sum_{j=1+\lfloor nt_{i-1} \rfloor}^{\lfloor nt_i \rfloor} \xi_j, \quad i = 1, 2, 3, \dots, k,$$

are independent. Now, since  $0 \leq t_{i-1} < t_i \leq 1$ , it follows that  $\lfloor nt_{i-1} \rfloor < \lfloor nt_i \rfloor$  for sufficiently large  $n \in \mathbb{N}$ . In addition,

$$\begin{aligned} \frac{\lfloor nt_i \rfloor - \lfloor nt_{i-1} \rfloor}{n} &= \frac{\lfloor nt_i \rfloor}{n} - \frac{\lfloor nt_{i-1} \rfloor}{n} = \left( \frac{nt_i}{n} + \frac{\lfloor nt_i \rfloor - nt_i}{n} \right) - \left( \frac{nt_{i-1}}{n} + \frac{\lfloor nt_{i-1} \rfloor - nt_{i-1}}{n} \right) \\ &= t_i - t_{i-1} + \frac{\lfloor nt_i \rfloor - nt_i}{n} - \frac{\lfloor nt_{i-1} \rfloor - nt_{i-1}}{n}, \end{aligned}$$

which implies

$$\left| \frac{\lfloor nt_i \rfloor - \lfloor nt_{i-1} \rfloor}{n} - (t_i - t_{i-1}) \right| = \left| \frac{\lfloor nt_i \rfloor - nt_i}{n} - \frac{\lfloor nt_{i-1} \rfloor - nt_{i-1}}{n} \right| \leq \frac{2}{n} \rightarrow 0, \text{ as } n \rightarrow \infty.$$

Thus,

$$\lim_{n \rightarrow \infty} \frac{\lfloor nt_i \rfloor - \lfloor nt_{i-1} \rfloor}{n} = t_i - t_{i-1} > 0.$$

Thus, by Claim 1, we see that, for each  $i = 1, 2, \dots, k$ ,

$$Y_i^{(n)} := \frac{1}{\sigma \cdot \sqrt{n}} \cdot \sum_{j=1+\lfloor nt_{i-1} \rfloor}^{\lfloor nt_i \rfloor} \xi_j \xrightarrow{d} \sqrt{t_i - t_{i-1}} \cdot N(0, 1) = N(0, t_i - t_{i-1}), \text{ as } n \rightarrow \infty. \quad (2.2)$$

By (2.1), (2.2), Proposition A.1, and Slutsky's Theorem (Corollary, p.40, [2]), we now see that

$$\sum_{i=1}^k c_i \left( X_{t_i}^{(n)} - X_{t_{i-1}}^{(n)} \right) = \sum_{i=1}^k c_i Y_i^{(n)} + \sum_{i=1}^k c_i \left\{ W(t_i)_n - W(t_{i-1})_n \right\} \xrightarrow{d} N \left( 0, \sum_{i=1}^k c_i^2 (t_i - t_{i-1}) \right).$$

This completes the proof of Claim 3.

(iv) Let  $t_0 := 0$ , hence,  $X_{t_0}^{(n)} \equiv 0$  for each  $n \in \mathbb{N}$ . We thus have, for each  $n \in \mathbb{N}$ ,

$$\begin{pmatrix} X_{t_1}^{(n)} \\ X_{t_2}^{(n)} \\ \vdots \\ X_{t_k}^{(n)} \end{pmatrix} = \underbrace{\begin{bmatrix} 1 & 0 & 0 & \cdots & \cdots & 0 & 0 \\ 1 & 1 & 0 & \cdots & \cdots & 0 & 0 \\ 1 & 1 & 1 & \cdots & \cdots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \ddots & 0 & 0 \\ \vdots & \vdots & \vdots & \ddots & \ddots & 0 & 0 \\ 1 & 1 & 1 & \cdots & \cdots & 1 & 0 \\ 1 & 1 & 1 & \cdots & \cdots & 1 & 1 \end{bmatrix}}_T \cdot \begin{pmatrix} X_{t_1}^{(n)} \\ X_{t_2}^{(n)} - X_{t_1}^{(n)} \\ \vdots \\ X_{t_k}^{(n)} - X_{t_{k-1}}^{(n)} \end{pmatrix}.$$

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By (iii), we know that

$$\begin{pmatrix} X_{t_1}^{(n)} \\ X_{t_2}^{(n)} - X_{t_1}^{(n)} \\ \vdots \\ X_{t_k}^{(n)} - X_{t_{k-1}}^{(n)} \end{pmatrix} \xrightarrow{d} \begin{pmatrix} Z_{t_1} \\ Z_{t_2-t_1} \\ \vdots \\ Z_{t_k-t_{k-1}} \end{pmatrix} \sim N\left(\mu = \mathbf{0}, \Sigma = \text{diag}(t_1, t_2 - t_1, \dots, t_k - t_{k-1})\right), \text{ as } n \rightarrow \infty.$$

Since the map  $\mathbb{R}^k \rightarrow \mathbb{R}^k : x \mapsto T \cdot x$  is continuous, we see immediately by Slutsky's Theorem (Theorem 6(a), p.39, [2]) that

$$\begin{pmatrix} X_{t_1}^{(n)} \\ X_{t_2}^{(n)} \\ \vdots \\ X_{t_k}^{(n)} \end{pmatrix} \xrightarrow{d} T \cdot \begin{pmatrix} Z_{t_1} \\ Z_{t_2-t_1} \\ \vdots \\ Z_{t_k-t_{k-1}} \end{pmatrix}, \text{ as } n \rightarrow \infty.$$

Since the map  $\mathbb{R}^k \rightarrow \mathbb{R}^k : x \mapsto T \cdot x$  is an invertible linear automorphism on  $\mathbb{R}^k$ , we see that

$$L = \begin{pmatrix} L_{t_1} \\ L_{t_2} \\ \vdots \\ L_{t_k} \end{pmatrix} := T \cdot \begin{pmatrix} Z_{t_1} \\ Z_{t_2-t_1} \\ \vdots \\ Z_{t_k-t_{k-1}} \end{pmatrix}$$

is still an  $\mathbb{R}^k$ -valued Gaussian random variable, and it clearly has expectation value  $\mathbf{0} \in \mathbb{R}^k$ , since each of  $Z_{t_1}, Z_{t_2-t_1}, \dots, Z_{t_k-t_{k-1}}$  has expectation value  $0 \in \mathbb{R}$ . It remains only to compute the covariance matrix of the  $\mathbb{R}^k$ -valued Gaussian random variable  $L$ . To this end, consider  $1 \leq i \leq j \leq k$ , i.e.  $t_i \leq t_j$ . Then, using the alternative notation  $Z_{t_1-t_0} := Z_{t_1}$ , we have

$$\begin{aligned} \text{Cov}(L_{t_i}, L_{t_j}) &= \text{Cov}(Z_{t_1} + Z_{t_2-t_1} + \dots + Z_{t_i-t_{i-1}}, Z_{t_1} + Z_{t_2-t_1} + \dots + Z_{t_j-t_{j-1}}) \\ &= \text{Cov}\left(\sum_{a=1}^i Z_{t_a-t_{a-1}}, \sum_{b=1}^j Z_{t_b-t_{b-1}}\right) = \sum_{a=1}^i \sum_{b=1}^j \text{Cov}(Z_{t_a-t_{a-1}}, Z_{t_b-t_{b-1}}) \\ &= \sum_{a=1}^i \text{Cov}(Z_{t_a-t_{a-1}}, Z_{t_a-t_{a-1}}) = \sum_{a=1}^i \text{Var}(Z_{t_a-t_{a-1}}) = \sum_{a=1}^i (t_a - t_{a-1}) \\ &= (t_1 - t_0) + (t_2 - t_1) + \dots + (t_{i-1} - t_{i-2}) + (t_i - t_{i-1}) \\ &= t_i = \min\{t_i, t_j\}, \end{aligned}$$

as required. □

## A Technical Lemmas

Note that  $X_n \xrightarrow{d} X$  and  $Y_n \xrightarrow{d} Y$  does NOT in general imply  $X_n + Y_n \xrightarrow{d} X + Y$ . But the implication does hold if  $X_n$  and  $Y_n$  are independent for each  $n \in \mathbb{N}$ , and both  $X$  and  $Y$  are Gaussian random variables, as the following Proposition shows.

**Proposition A.1** *Let  $k \in \mathbb{N}$  be fixed. Suppose:*

- For each  $n \in \mathbb{N}$ ,

$$Y_1^{(n)}, Y_2^{(n)}, \dots, Y_k^{(n)} : \Omega^{(n)} \longrightarrow \mathbb{R}$$

*are independent  $\mathbb{R}$ -valued random variables defined on the probability space  $\Omega^{(n)}$ .*

- For each  $i = 1, 2, \dots, k$ ,

$$Y_i^{(n)} \xrightarrow{d} N(\mu_i, \sigma_i^2), \quad \text{as } n \longrightarrow \infty.$$

Then, for any  $c_1, c_2, \dots, c_k \in \mathbb{R}$ ,

$$\sum_{i=1}^k c_i Y_i^{(n)} \xrightarrow{d} N\left(\sum_{i=1}^k c_i \mu_i, \sum_{i=1}^k c_i^2 \sigma_i^2\right), \quad \text{as } n \longrightarrow \infty.$$

**PROOF** Let  $Y^{(n)} := \sum_{i=1}^k c_i Y_i^{(n)}$ . Let  $\varphi_X$  denote the characteristic function of a  $\mathbb{R}$ -valued random variable  $X$ . Then,

$$\begin{aligned} \varphi_{Y^{(n)}}(t) &= \varphi_{\sum_{i=1}^k c_i Y_i^{(n)}}(t) \\ &= \prod_{i=1}^k \varphi_{c_i Y_i^{(n)}}(t), \quad \text{since } Y_1^{(n)}, \dots, Y_k^{(n)} \text{ are independent} \\ &= \prod_{i=1}^k \varphi_{Y_i^{(n)}}(c_i t) \\ &\longrightarrow \prod_{i=1}^k \exp\left\{\sqrt{-1} \mu_i (c_i t) - \frac{1}{2} \sigma_i^2 (c_i t)^2\right\} \\ &= \exp\left\{\sqrt{-1} \left(\sum_{i=1}^k c_i \mu_i\right) t - \frac{1}{2} \left(\sum_{i=1}^k c_i^2 \sigma_i^2\right) t^2\right\}, \quad \text{as } n \longrightarrow \infty, \end{aligned}$$

where the second and third equalities follow from the properties of characteristic functions of random variables (see p.21, [2]), while the expression of the limit follows from the fact that the characteristic function  $\varphi_Z$  of a random variable  $Z$  with distribution  $N(\mu, \sigma^2)$  is

$$\varphi_Z = \exp\left\{\sqrt{-1} \mu t - \frac{1}{2} \sigma^2 t^2\right\}.$$

The Proposition now follows immediately from the Lévy-Cramér Continuity Theorem (Theorem 1.9(ii), p.56, [3]).  $\square$

## B Continuous maps are Borel measurable

**Lemma B.1 (The pre-image of a  $\sigma$ -algebra is itself a  $\sigma$ -algebra.)**

*Suppose  $\Omega$  is a non-empty set,  $(X, \mathcal{X})$  is a measurable space, and  $f : \Omega \longrightarrow X$  is a map from  $\Omega$  into  $X$ . Then,*

$$f^{-1}(\mathcal{X}) := \{f^{-1}(V) \subset \Omega \mid V \in \mathcal{X}\}$$

*is a  $\sigma$ -algebra of subsets of  $\Omega$ .*



# Donsker's Theorems (Functional Central Limit Theorems)

Study Notes

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PROOF

$$\Omega \in f^{-1}(\mathcal{X}) \quad f(\Omega) \subset X \implies \Omega = f^{-1}(X) \in f^{-1}(\mathcal{X}).$$

$f^{-1}(\mathcal{X})$  is closed under complementations Let  $V \in \mathcal{X}$ . Then,  $X \setminus V \in \mathcal{X}$ , and

$$\Omega \setminus f^{-1}(V) = \{ \omega \in \Omega \mid f(\omega) \notin V \} = \{ \omega \in \Omega \mid f(\omega) \in X \setminus V \} = f^{-1}(X \setminus V) \in f^{-1}(\mathcal{X}),$$

which shows that  $f^{-1}(\mathcal{X})$  is indeed closed under complementations.

$f^{-1}(\mathcal{X})$  is closed countable unions Let  $V_1, V_2, \dots \in \mathcal{X}$ . Then,  $\bigcup_{i=1}^{\infty} V_i \in \mathcal{X}$ , and

$$\bigcup_{i=1}^{\infty} f^{-1}(V_i) = \left\{ \omega \in \Omega \mid \begin{array}{c} f(\omega) \in V_i \\ \text{for some } i \in \mathbb{N} \end{array} \right\} = f^{-1}\left(\bigcup_{i=1}^{\infty} V_i\right) \in f^{-1}(\mathcal{X}),$$

which proves that  $f^{-1}(\mathcal{X})$  is indeed closed under countable unions.

This concludes the proof that that  $f^{-1}(\mathcal{X})$  is a  $\sigma$ -algebra of subsets of  $\Omega$ . □

**Lemma B.2 (The push-forward of a  $\sigma$ -algebra is itself a  $\sigma$ -algebra.)**

Suppose  $(\Omega, \mathcal{A})$  is a measurable space,  $X$  is a non-empty set, and  $f : \Omega \rightarrow X$  is a map from  $\Omega$  into  $X$ . Then,

$$\mathcal{F} := \{ V \subset X \mid f^{-1}(V) \in \mathcal{A} \}$$

is a  $\sigma$ -algebra of subsets of  $X$ .

PROOF

$$\underline{X \in \mathcal{F}} \quad f^{-1}(X) = \Omega \in \mathcal{A} \implies X \in \mathcal{F}.$$

$\mathcal{F}$  is closed under complementations  $V \in \mathcal{F} \implies f^{-1}(V) \in \mathcal{A} \implies f^{-1}(X \setminus V) = \Omega \setminus f^{-1}(V) \in \mathcal{A} \implies X \setminus V \in \mathcal{F}$ , which proves that  $\mathcal{F}$  is indeed closed under complementations.

$\mathcal{F}$  is closed under countable unions

$$\begin{aligned} V_1, V_2, \dots \in \mathcal{F} &\implies f^{-1}(V_1), f^{-1}(V_2), \dots \in \mathcal{A} \\ &\implies f^{-1}\left(\bigcup_{i=1}^{\infty} V_i\right) = \bigcup_{i=1}^{\infty} f^{-1}(V_i) \in \mathcal{A} \\ &\implies \bigcup_{i=1}^{\infty} V_i \in \mathcal{F}, \end{aligned}$$

which proves that  $\mathcal{F}$  is indeed closed under countable unions. □

**Theorem B.3**

Suppose  $(\Omega, \mathcal{A})$  and  $(X, \mathcal{X})$  are measurable spaces, and  $f : \Omega \rightarrow X$  is a map from  $\Omega$  into  $X$ . Then,  $f$  is  $(\mathcal{A}, \mathcal{X})$ -measurable if there exists  $\mathcal{S} \subset \mathcal{X}$  satisfying the following conditions:

- $\mathcal{S}$  generates  $\mathcal{X}$ , i.e.  $\sigma(\mathcal{S}) = \mathcal{X}$ , and
- $f^{-1}(S) \in \mathcal{A}$ .

PROOF By Lemma B.2,

$$\mathcal{F} := \{ V \subset X \mid f^{-1}(V) \in \mathcal{A} \}$$

is a  $\sigma$ -algebra of subsets of  $X$ . By hypothesis,  $\mathcal{S} \subset \mathcal{F}$ ; hence,  $\mathcal{X} = \sigma(\mathcal{S}) \subset \mathcal{F}$ . Thus,  $f^{-1}(\mathcal{X}) \subset \mathcal{A}$ ; equivalently,  $f$  is  $(\mathcal{A}, \mathcal{X})$ -measurable.  $\square$

**Corollary B.4 (Continuous maps are Borel measurable.)**

*Suppose  $X_1, X_2$  are topological spaces, and  $\mathcal{B}_1, \mathcal{B}_2$  are their respective Borel  $\sigma$ -algebras. Then, every continuous map  $f : X_1 \rightarrow X_2$  is  $(\mathcal{B}_1, \mathcal{B}_2)$ -measurable.*

## C Topology

**Theorem C.1 (Appendix M3, [1])**

*Suppose  $S$  is a metric space. Then, the following conditions are equivalent:*

- (i)  *$S$  is separable.*
- (ii) *The topology of  $S$  has a countable basis.*
- (iii) *Every open cover of **each subset** of  $S$  has a countable subcover.*

## References

- [1] BILLINGSLEY, P. *Convergence of Probability Measures*, second ed. John Wiley & Sons, 1999.
- [2] FERGUSON, T. S. *A Course in Large Sample Theory*, first ed. Texts in Statistical Science. CRC Press, 1996.
- [3] SHAO, J. *Mathematical Statistics*, second ed. Springer, 2003.