## 1 Scheffé's S-Method: A Special Case

Let  $X_1, \ldots, X_r$  be independent normal  $\mathbb{R}$ -valued random variables with common known variance  $\sigma^2 = 1$ , and  $E(X_i) = \xi_i$ . We seek to find simultaneous confidence intervals for the collection of linear functionals

$$f_{\mathbf{u}}(\xi) = \mathbf{u} \bullet \xi = \sum_{i=1}^{r} u_i \, \xi_i,$$

indexed by  $\mathbf{u} = (u_1, u_2, \dots, u_r)$  in the unit sphere  $S^{r-1}$  in  $\mathbb{R}^r$ , i.e.

$$\|\mathbf{u}\|^2 = \sum_{i=1}^r u_i^2 = 1.$$

In other words, we seek two functions  $L, M: S^{r-1} \times \mathbb{R}^r \longrightarrow \mathbb{R}$  such that, for each  $\mathbf{x} \in \mathbb{R}^r$ , the set

$$S(\mathbf{x}; L, M) := \left\{ \zeta \in \mathbb{R}^r \mid L(\mathbf{u}, \mathbf{x}) \le \mathbf{u} \bullet \zeta \le M(\mathbf{u}, \mathbf{x}), \text{ for each } \mathbf{u} \in S^{r-1} \right\}$$

satisfies:

$$P_{\zeta}(\{\mathbf{x} \in \mathbb{R}^r \mid \zeta \in S(\mathbf{x}; L, M)\}) = \gamma, \text{ for each } \zeta \in \mathbb{R}^r.$$

**Proposition 1.1** Given any such  $S(\mathbf{x}; L, M)$ , there exists some constant c > 0 such that

$$S(\mathbf{x}; L, M) = S(\mathbf{x}; c) := \left\{ \zeta \in \mathbb{R}^r \mid \|\mathbf{u} \bullet (\mathbf{x} - \zeta)\| \le c, \text{ for each } \mathbf{u} \in S^{r-1} \right\}$$

PROOF We first make the following

**CLAIM 1:** Without loss of generality, we may assume that the functions L and M satisfy:

$$L(\mathbf{u}; \mathbf{x}) = -M(-\mathbf{u}; \mathbf{x}).$$

Indeed, since  $\mathbf{u} \in S^{r-1} \iff -\mathbf{u} \in S^{r-1}$ , we have that, for any  $\mathbf{u} \in S^{r-1}$ ,

$$L(\mathbf{u}, \mathbf{x}) \le \mathbf{u} \bullet \zeta \le M(\mathbf{u}, \mathbf{x}) \implies L(-\mathbf{u}, \mathbf{x}) \le -\mathbf{u} \bullet \zeta \le M(-\mathbf{u}, \mathbf{x}) \implies -M(-\mathbf{u}, \mathbf{x}) \le \mathbf{u} \bullet \zeta \le -L(-\mathbf{u}, \mathbf{x})$$

Hence, the inequalities  $L(\mathbf{u}, \mathbf{x}) \leq \mathbf{u} \cdot \zeta \leq M(\mathbf{u}, \mathbf{x})$  in fact imply:

$$\widetilde{L}(\mathbf{u}; \mathbf{x}) := \max\{L(\mathbf{u}; \mathbf{x}), -M(-\mathbf{u}; \mathbf{x})\} \le \mathbf{u} \bullet \zeta \le \min\{M(\mathbf{u}; \mathbf{x}), -L(-\mathbf{u}; \mathbf{x})\} =: \widetilde{M}(\mathbf{u}; \mathbf{x}).$$

Thus, we may replace the original bounding functions L and M with the tighter  $\widetilde{L}$  and  $\widetilde{M}$ , respectively. Furthermore, note that

$$\widetilde{L}(\mathbf{u}; \mathbf{x}) := \max\{L(\mathbf{u}; \mathbf{x}), -M(-\mathbf{u}; \mathbf{x})\} = -\min\{-L(\mathbf{u}; \mathbf{x}), M(-\mathbf{u}; \mathbf{x})\} = -\widetilde{M}(-\mathbf{u}; \mathbf{x})$$

This completes the proof of CLAIM 1.

**CLAIM 2:** The functions L and M are "invariant" with respect to the orthogonal group

$$O(r) := \left\{ Q \in \operatorname{GL}(\mathbb{R}^r) \mid Q^{\operatorname{T}} \cdot Q = Q \cdot Q^{\operatorname{T}} = I_r \right\}$$

in the following sense:

$$L(Q\mathbf{u}; Q\mathbf{x}) = L(\mathbf{u}; \mathbf{x}), \text{ and } M(Q\mathbf{u}; Q\mathbf{x}) = M(\mathbf{u}; \mathbf{x}), \text{ for each } \mathbf{u} \in S^{r-1}, \mathbf{x} \in \mathbb{R}^r, Q \in O(r).$$

## Scheffé's S-Method for Simultaneous Confidence Intervals

Kenneth Chu Study Notes August 12, 2012

## A Equivariance

**Definition A.1** Suppose:

- $X:(\Omega,\mathcal{A},\mu)\longrightarrow (\mathcal{X},\mathcal{B})$  is a random variable.
- $G_{\mathcal{X}} \subset \operatorname{Aut}(\mathcal{X}, \mathcal{B})$  is a subgroup of the automorphism group  $\operatorname{Aut}(\mathcal{X}, \mathcal{B})$  of the codomain  $(\mathcal{X}, \mathcal{B})$  of X.
- $G_{\mathcal{T}} \in \operatorname{Aut}(\mathcal{T})$  is a subgroup of the automorphism group  $\operatorname{Aut}(\mathcal{T})$  of the decision space  $\mathcal{T}$ .
- $h: G_{\mathcal{X}} \longrightarrow G_{\mathcal{T}}$  is a homomorphism of groups. Thus,  $G_{\mathcal{X}}$  induces an action of  $\mathcal{T}$  via h.
- $T:(\mathcal{X},\mathcal{B})\longrightarrow (\mathcal{T},\mathcal{C})$  be a measurable map. T will be called a **statistic**.

The statistic T is said to be  $(G_{\mathcal{X}}, h)$ -equivariant if the following condition holds:

$$T \circ g = h(g) \circ T$$
; equivalently  $T(g(x)) = h(g)[T(x)]$  for each  $x \in \mathcal{X}, g \in G_{\mathcal{X}}$ 

## References