HW 2

MFE 407: Empirical Methods in Finance

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Group 6

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Problem 1 AR(p) Process:

```
1.(a)
phi0 <- 0
phi1 <- 1.1
phi2 <- -0.25

size <- 10000
w <- rep(0, size)

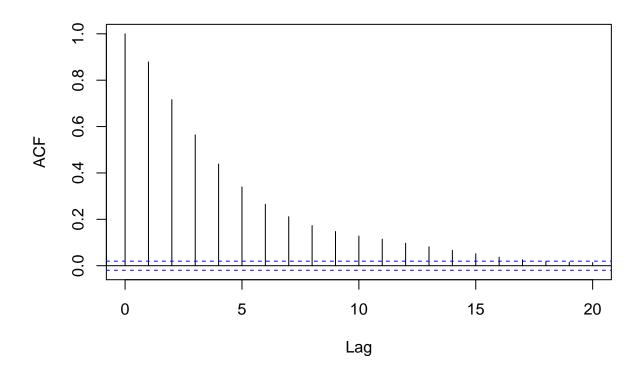
w[1] <- 0.5
w[2] <- 0.1

epsilon <- rnorm(size, 0,1)

for(i in 3:size) {
    w[i] = phi0 + phi1 * w[i-1] + phi2 * w[i-2] + epsilon[i]
}

acf(w, lag.max = 20)</pre>
```

Series w



```
1.(b)

x1 <- (phi1 + sqrt(phi1^2 + 4 * phi2))/(-2 * phi2)

x2 <- (phi1 - sqrt(phi1^2 + 4 * phi2))/(-2 * phi2)

w1 <- 1/x1

w2 <- 1/x2

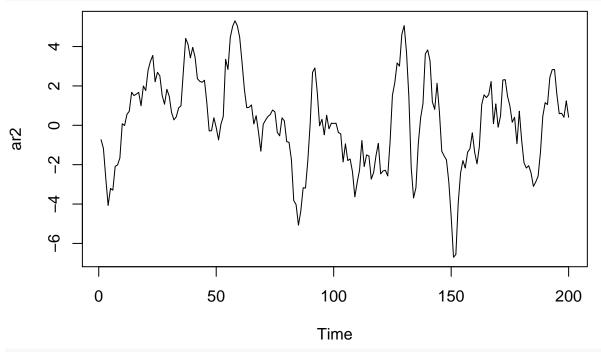
w1

## [1] 0.3208712
```

[1] 0.7791288

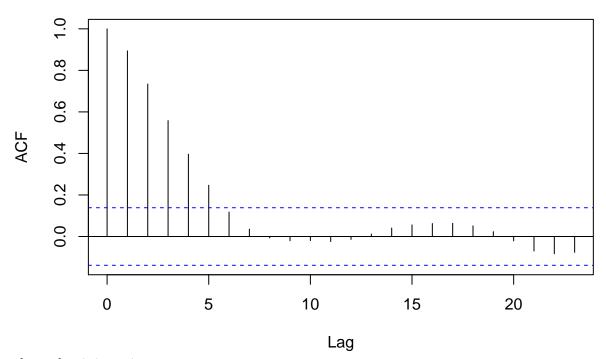
The two roots are real, so the process is stationary. Also based on the plot, the series converges Also plot the graph and acf:

```
ar2 <- arima.sim(model = list(order = c(2,0,0), ar=c(1.1,-0.25)), n = 200)
plot(ar2)</pre>
```



acf(ar2)

Series ar2



shows that it is stationary.

```
(c) After expanding the series: \frac{\partial X_t}{\partial \varepsilon} = \phi_1^6 + 5\phi_1^4\phi_2 + 6\phi_1^2\phi_2^2 + \phi_2^3 multiplier1 <- phi1^6 + 5 * phi1^4 *phi2 + 6 * phi1^2 * phi2^2 + phi2^3 cat("the multiplier is", multiplier1)
```

This

the multiplier is 0.379561

```
(d)
phi1 <- 0.9
phi2 <- 0.8
x1 <- (phi1 + sqrt(phi1^2 + 4 * phi2))/(-2 * phi2)
x2 <- (phi1 - sqrt(phi1^2 + 4 * phi2))/(-2 * phi2)
multiplier2 <- phi1^6 + 5 * phi1^4 * phi2 + 6 * phi1^2 * phi2^2 + phi2^3
cat("The multiplier is", multiplier2)
```

```
## The multiplier is 6.778241
cat("The two roots are", 1/x1,1/x2)
```

The two roots are -0.5512492 1.451249

One of the roots are bigger than 1, hence the process is not stationary.

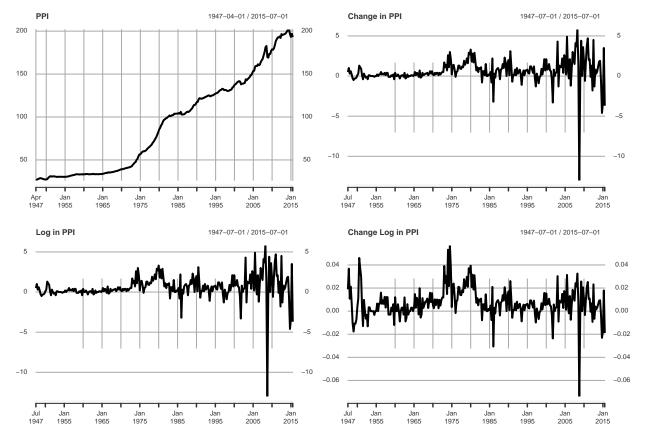
Problem 2

1.

```
library(lubridate)
```

##

```
## Attaching package: 'lubridate'
## The following object is masked from 'package:base':
##
##
       date
library(xts)
## Loading required package: zoo
##
## Attaching package: 'zoo'
## The following objects are masked from 'package:base':
##
##
       as.Date, as.Date.numeric
library(forecast)
## Warning: package 'forecast' was built under R version 3.5.2
par(mfrow=c(2,2))
ppi_raw <- read.csv("PPIFGS.csv")</pre>
ppi_xts <- xts(x=as.double(ppi_raw$VALUE), ymd(ppi_raw$DATE))</pre>
#2.a
plot(ppi_xts, main ="PPI", cex=0.3)
#2.b
ppi_diff <- diff(ppi_xts)[-1]</pre>
plot(ppi_diff, main = "Change in PPI", cex=0.3)
#2.c
ppi_log <- log(ppi_xts)</pre>
plot(ppi_diff, main = "Log in PPI", cex=0.3)
#2.d
ppi_diff_log <- diff(ppi_log)[-1]</pre>
plot(ppi_diff_log, main = "Change Log in PPI", cex=0.3)
```

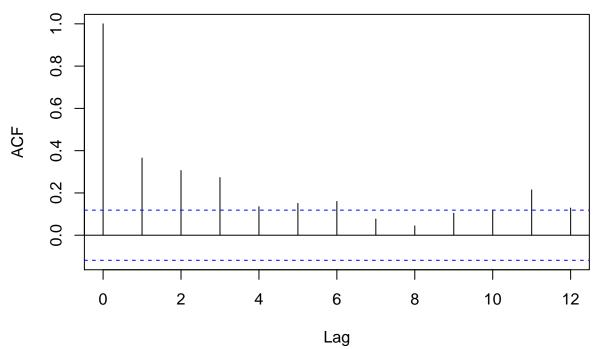


2. Change Log in PPI 2.(d) is covariance stationary. 2(a) has no constant mean 2(b), 2(c) don't have constant variance.

3.

acf(ppi_diff_log, lag.max = 12)

Series ppi_diff_log



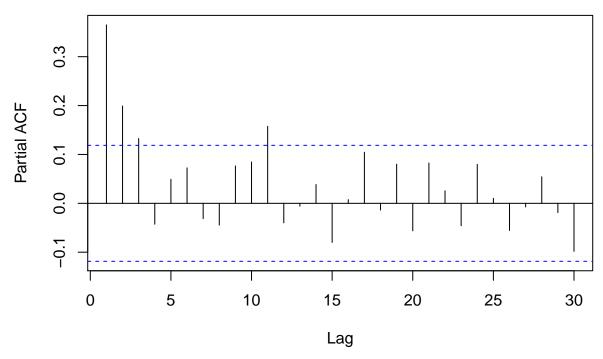
ACF of changes of log(PPI) converges as lags increase and starts to diminish below one standard deviation (s.d) when lag is greater than 4. There might exist some seasonality in prior periods as we observe ACF greater than 1 s.d for lag = 11.

The

4.

pacf(ppi_diff_log, lag.max = 30)

Series ppi_diff_log



We tried to plot PACF with 30 lags, and identify that there is no seasonality in the data. On the PACF plot, it is still significant after 3 lags and also significant on 11th lag. We would like to try different models (AR(3), AR(1:3,11)) to see which one fits better.

5.

```
Choose AR(3) and AR with lag 1,2,3 and 11
```

Coefficent for Ar3 is:

ar3\$coef

```
## ar1 ar2 ar3 intercept
## 0.268731607 0.160715955 0.140186620 0.007322805
```



```
cat("Coefficent for AR with lag on (1,2,3,11) is:\n")
```

Coefficent for AR with lag on (1,2,3,11) is:

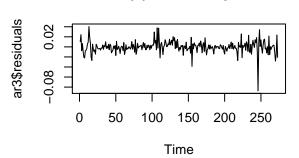
```
ar12311$coef
```

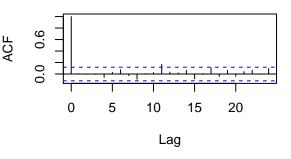
```
## ar1 ar2 ar3 ar4 ar5 ar6
## 0.252303148 0.146834382 0.139532331 0.000000000 0.000000000 0.000000000
## ar7 ar8 ar9 ar10 ar11 intercept
```

```
cat("Standard error for Ar3 is:\n")
## Standard error for Ar3 is:
sqrt(ar3$sigma2)
## [1] 0.0117843
cat("################n")
cat("Standard error for AR with lag on (1,2,3,11) is:\n")
## Standard error for AR with lag on (1,2,3,11) is:
sqrt(ar12311$sigma2)
## [1] 0.01157444
ar3roots <-polyroot(c(1,-ar3$coef[1:3]))</pre>
ar12311roots <-polyroot(c(1,-ar12311$coef[1:11]))</pre>
cat("######################\n")
cat("Mod for Ar3 is:\n")
## Mod for Ar3 is:
Mod(1/ar3roots)
## [1] 0.7409665 0.4349647 0.4349647
cat("Mod for AR with lag on (1,2,3,11) is:\n")
## Mod for AR with lag on (1,2,3,11) is:
Mod(1/ar12311roots)
## [1] 0.8336369 0.8502784 0.8502784 0.8688015 0.8428316 0.8428316 0.9304305
  [8] 0.8688015 0.8346560 0.8336369 0.8346560
As the mod of the characteristic roots are smaller than 1, so we conclude that these models are stationary
5.b Plot both AR plots and ACF to show that residual is white noise and has no autocorrelation
par(mfrow=c(2,2))
plot(ar3$residuals, main="AR(3) Residual plot")
acf(ar3$residuals, main="ACF AR(3) Residual plot")
plot(ar12311$residuals, main="AR(1:3,11) Residual plot")
acf(ar12311$residuals, main="ACF AR(1:3,11) Residual plot")
```

AR(3) Residual plot

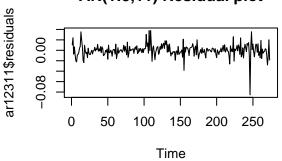
ACF AR(3) Residual plot

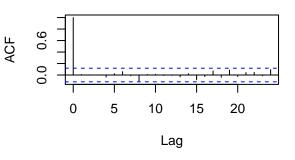




AR(1:3,11) Residual plot

ACF AR(1:3,11) Residual plot





5.c
Box.test(ar3\$residuals, lag=8, type='Ljung')

```
##
## Box-Ljung test
##
## data: ar3$residuals
## X-squared = 5.3458, df = 8, p-value = 0.7201
Box.test(ar3$residuals, lag=12, type='Ljung')
```

```
##
## Box-Ljung test
##
## data: ar3$residuals
## X-squared = 13.829, df = 12, p-value = 0.3118
AIC(ar3)
```

```
## [1] -1639.746
BIC(ar3)
```

[1] -1621.699
Box.test(ar12311\$residuals, lag=8, type='Ljung')

##
Box-Ljung test
##
data: ar12311\$residuals

```
## X-squared = 4.6285, df = 8, p-value = 0.7964
Box.test(ar12311$residuals, lag=12, type='Ljung')
##
## Box-Ljung test
##
## data: ar12311$residuals
## X-squared = 4.7273, df = 12, p-value = 0.9665
AIC(ar12311)
## [1] -1647.148
BIC(ar12311)
## [1] -1625.491
Choose AR with lag on 1,2,3,11 because of lower AIC & BIC and p-value greater than 5% significance level.
  6. Fit AR(3) and AR model with lag 1,2,3 and 11
ppi_xts2005 <- ppi_xts["1947-04-01/2005-12-31"]
ppi_diff_log_2005 <- diff(log(ppi_xts2005))[-1]</pre>
num <- length(ppi_diff_log) - length(ppi_diff_log_2005)</pre>
traindata <- ppi_diff_log["2005-12-31/2015"]</pre>
# fit for the models
ar3_2005<- arima(ppi_diff_log_2005, order=c(3,0,0))
ar12311_2005<- arima(ppi_diff_log_2005, order=c(11,0,0),
               fixed=c(NA,NA,NA,0,0,0,0,0,0,0,NA,NA),transform.pars = FALSE)
fitar3 <- forecast(ar3 2005, h=39)
fitar12311 <-forecast(ar12311_2005, h=num)
e_ar3 <- sum((as.double(traindata) - fitar3$mean)^2) /num</pre>
e_ar12311 <- sum((as.double(traindata) - fitar12311$mean)^2) /num
cat("MSPE AR(3) is:",e_ar3,"\n")
## MSPE AR(3) is: 0.000338503
cat("MSPE AR(1,2,3,11) is:",e_ar12311,"\n")
## MSPE AR(1,2,3,11) is: 0.0003393177
Simulate Random Walk (39 steps)
e_rw \leftarrow rep(0,10000)
for(j in 1: 10000){
  result <- rep(0,num+1)
 result[1] <- last(ppi_xts2005)</pre>
 for(i in 2:40){
   result[i] = result[i-1] + rnorm(1,0,1)
  }
```

```
rw <- diff(log(result))
e_rw[j] <- sum((as.double(traindata) - rw)^2) /num
}
cat("Average mean of MSPE of 10000 random walk is:", mean(e_rw),"\n")</pre>
```

 $\mbox{\tt \#\#}$ Average mean of MSPE of 10000 random walk is: 0.0003867735

Hence our model is definitely better than random walk.