

Williams College ECON 379:

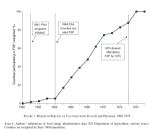
Program Evaluation for International Development

Module 7: Two-Way Fixed Effects

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Diff-in-Diff with Staggered Timing

Example: counties introduced food stamps at different times





Notes: Authors' tabulations of food starm patients trained to the process of the shading corresponds to the county FSP start data, where darkers basing indicases later county implementation.

source: Almond, Hoynes, and Schanzenbach (AER, 2016)

Example: states adopted Medicaid at different times

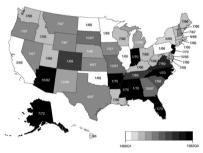


Figure 2.

Medicaid Adoption by Quarter Notes: Adoption dates come from the Department of Health Education and Welfare (1970)

& Social Security Administration (2013). The map is shaded relative to the quarter of adoption and states are labeled with the month and year of adoption.

source: Boudreaux, Golberstein, and McAlpine (Journal of Health Economics, 2016)

Example: counties opening community health centers



FIGURE 3, ESTABLISHMENT OF COMMUNITY HEALTH CENTERS BY COUNTY OF SERVICE DELIVERY, 1965–1980

Note: Dates are the first year that a CHC was established in the county.

Source: Information on CHCs drawn from NACAP and PHS reports.

source: Bailey and Goodman-Bacon (AER, 2015)

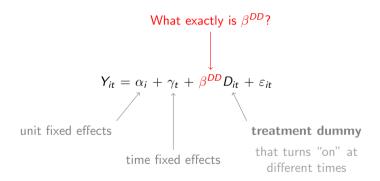
Example: African countries democratized at different times



FIGURE A.2. Geographical distribution of democratized countries since 1990. Black-colored countries are democratized since 1990; grey-colored countries the other countries in the sample for infant mortality analysis. Democratized countries include the Comoros, tiny islands to the northwest of Madaeascar, which may not be visible as black-colored.

source: Kudamatsu (JEEA, 2012)

Two-Way Fixed Effects Estimates of β^{DD}



What exactly is β^{DD} in TWFE?

	t = 1	t = 2	t = 3	t = 4	t = 5
ID 1	3	3	5	5	5
ID 2	3	3	5	5	5
ID 3	1	1	4	4	4
ID 4	1	1	4	4	4

treatment

comparison

$$\Rightarrow \beta^{DD} = 5 - 3 - (4 - 1)$$

What exactly is β^{DD} in TWFE?

	t = 1	t=2	t = 3	t = 4	t = 5	
ID 1	3	3	5	5	5	
ID 2	3	3	7	7	7	١
ID 3	1	1	4	4	4	,
ID 4	1	1	4	4	4	

treatment

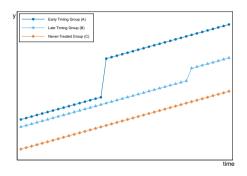
comparison

$$\Rightarrow \beta^{DD} = 6 - 3 - (4 - 1) \leftarrow \text{average treatment effect}$$

Multiple Treatment and Comparison Groups

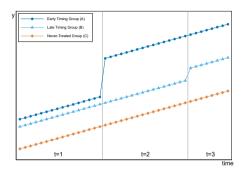
	t = 1	t=2	t = 3	t = 4	t = 5	t = 6	t = 7	t = 8
ID 1	0	0	1	1	1	1	1	1
ID 2	0	0	1	1	1	1	1	1
ID 3	0	0	0	0	0	1	1	1
ID 4	0	0	0	0	0	1	1	1
ID 5	0	0	0	0	0	0	0	0
ID 6	0	0	0	0	0	0	0	0

Decomposition into Timing Groups



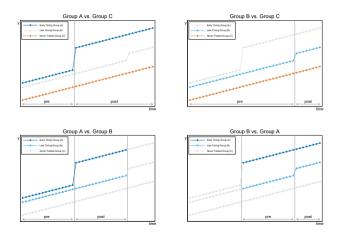
Panel with variation in treatment timing can be decomposed into distinct **timing groups** reflecting observed onset of treatment

Decomposition into Timing Groups



Example: with three timing groups (one of which is never treated), can construct three timing windows (pre, middle, post or t = 1, 2, 3)

Decomposition into Standard 2 × 2 DDs



Two-Way Fixed Effects β^{DD} as a Weighted Sum

The two-way fixed effects estimator β^{DD} is a weighted sum of all possible 2 \times 2 diff-in-diff estimators across timing groups

- Some use an already-treated group as comparison
 - Creates problems if treatment effect grows/changes over time
 - If treatment effect gets increases over time within treated units, using already-treated units biases estimates of treatment effect
 - Bias intuitively similar to continuous treatment case
 - When treated units in comparison group, we are relying on assumptions about functional form (of treatment effect)

Easiest way to see this \rightarrow unpack TWFE diff-in-diff estimator (as a weighted sum of Y values – but what are the weights?)

Two-Way Fixed Effects as Univariate Regression

Two-way fixed effects is equivalent to univariate regression:

$$\tilde{Y}_{it} = \alpha + \tilde{D}_{it} + \epsilon_{it}$$

(just the mean across i and t)

Two-Way Fixed Effects as Univariate Regression

Two-way fixed effects is equivalent to univariate regression:

$$\tilde{Y}_{it} = \alpha + \tilde{D}_{it} + \epsilon_{it}$$

where
$$ilde{Y}_{it} = Y_{it} - ar{Y}_t - \left(ar{Y}_i - ar{ar{Y}}\right)$$

and
$$ilde{D}_{it} = D_{it} - ar{D}_t - (ar{D}_i - ar{ar{D}})$$

⇒ Treatment "dummy" now continuous

$$\Rightarrow eta^{ extit{OLS}} = \sum_i ilde{Y}_{it} \left(ilde{D}_{it} - ar{ ilde{D}}_{it}
ight) \left(rac{1}{\sum_i \left(ilde{D}_{it} - ar{ ilde{D}}_{it}
ight)^2}
ight)$$

Two-Way Fixed Effects as Univariate Regression

Two-way fixed effects is equivalent to univariate regression:

where
$$ilde{Y}_{it}=Y_{it}-ar{Y}_t-ig(ar{Y}_i-ar{ar{Y}}ig)$$
 and $ilde{D}_{it}=D_{it}-ar{D}_t-ar{D}_t-ar{D}_i$

⇒ Treatment "dummy" now continuous

$$\Rightarrow eta^{OLS} = \sum_{i} \tilde{Y}_{it} \left(\frac{\tilde{D}_{it}}{\tilde{D}_{it}} - \frac{\tilde{\bar{D}}_{it}}{\tilde{D}_{it}} \right) \left(\frac{1}{\sum_{i} \left(\tilde{D}_{it} - \tilde{\bar{D}}_{it} \right)^{2}} \right)$$

 \Rightarrow Observations where $\tilde{D}_{it} - \tilde{D}_{it} < 0$ in comparison group

	t = 1	t = 2	t = 3	t = 4
ID 1	0	1	1	1
ID 2	0	0	0	1

Two timing groups, D_{it} is either 0 or 1

 \Rightarrow If we regressed Y_{it} on D_{it} , β_{OLS} is weighted sum of Y_{it} values

	t = 1	t = 2	t = 3	t = 4
ID 1	0	1	1	1
ID 2	0	0	0	1
$ar{ar{D}}_t$	0	0.5	0.5	1

	t=1	t = 2	t = 3	t = 4
ID 1	0	1	1	1
ID 2	0	0	0	1
$ar{D}_t$	0	0.5	0.5	1

D_{it} (treatment dummy)

	t = 1	t = 2	t = 3	t = 4
ID 1	0	1	1	1
ID 2	0	0	0	1
$ar{\mathcal{D}}_t$	0	0.5	0.5	1

$$\begin{array}{c|cccc} \bar{D}_i & \bar{\bar{D}} & \bar{D}_i - \bar{\bar{D}} \\ \hline 0.75 & 0.5 & 0.25 \\ \hline 0.25 & 0.5 & -0.25 \\ \hline \end{array}$$

	$D_{it}-ar{\mathcal{D}}_t$				
	t = 1	t = 2	t = 3	t = 4	
ID 1	0	0.5	0.5	0	
ID 2	0	-0.5	-0.5	0	
$ar{ar{D}_t}$	0	0.5	0.5	1_	

$$\begin{array}{c|cccc} \bar{D}_i & \bar{\bar{D}} & \bar{D}_i - \bar{\bar{D}} \\ \hline 0.75 & 0.5 & 0.25 \\ \hline 0.25 & 0.5 & -0.25 \\ \hline \end{array}$$

$$ilde{D}_{it} = D_{it} - ar{D}_t - \left(ar{D}_i - ar{ar{D}}\right) \longrightarrow \text{need to divide by}$$
 $t = 1 \quad t = 2 \quad t = 3 \quad t = 4$
 $\text{ID 1 } -0.25 \quad 0.25 \quad 0.25 \quad -0.25$
 $\text{ID 2 } 0.25 \quad -0.25 \quad -0.25 \quad 0.25$
 $\text{to get OLS weights}$

$$\tilde{D}_{it} = D_{it} - \bar{D}_t - \left(\bar{D}_i - \bar{\bar{D}}\right) \\
t = 1 \quad t = 2 \quad t = 3 \quad t = 4 \\
ID 1 \quad -0.25 \quad 0.25 \quad 0.25 \quad -0.25 \\
ID 2 \quad 0.25 \quad -0.25 \quad -0.25 \quad 0.25$$

Treated cells are not all getting positive weight

$$ilde{D}_{it} = D_{it} - ar{D}_t - \left(ar{D}_i - ar{ar{D}}\right)$$
 $t = 1$ $t = 2$ $t = 3$ $t = 4$
 $ID \ 1$ -0.25 0.25 0.25 -0.25
 $ID \ 2$ 0.25 -0.25 -0.25 0.25

Treated cells are not all getting positive weight

⇒ Later treated periods in early adopter units negatively weighted

	Y_{it}				
	t = 1	t = 2	t = 3	t = 4	
ID 1	0	10	10	10	
ID 2	0	0	0	10	

Let
$$Y_{it} = \gamma_i + \lambda_t + \delta_{it}$$

Treated cells

	Y_{it}				
	t = 1	t = 2	t = 3	t = 4	
ID 1	0	10	10	10	
ID 2	0	0	0	10	

Let
$$Y_{it} = \delta_{it}$$

 $\Rightarrow Y_{it}$ is treatment effect

Treated cells

 $\longrightarrow eta_{OLS} = 10$

homogeneous impacts:

$$E[\beta_{OLS}] = ATE$$

Treated cells

 $\longrightarrow \beta_{OLS} = 6$ heterogeneous impacts:

 $E[\beta_{OLS}] = \text{weighted ATE}$

Treated cells

heterogeneous impacts: $E[\beta_{OLS}] = \text{weighted ATE}$

 $\rightarrow \beta \alpha s = 6$

(with mysterious weights)

Treated cells

Treated cells

Practical Implications

Using diff-in-diff to evaluate staggered programs is complicated

- Avoid variation in treatment timing?
- A (substantial) never-treated group, many pre-periods helps
- Calculate your weights

Look at your data to assess common trends after treatment

- Graph the 2×2 diff-in-diffs
- Remove time FEs and plot the jumps at treatment

New tools for assessing the validity of diff-in-diff in these settings:

• Stata commands: bacondecomp and fuzzydid

