## Ecology Appendix S1 Posterior distribution and Gibbs sampler Modeling abundance, distribution, movement, and space use with camera and telemetry data

Richard B. Chandler<sup>1</sup>\*, Daniel A. Crawford<sup>2</sup>, Elina P. Garrison<sup>3</sup>, Karl V. Miller<sup>1</sup>, Michael J. Cherry<sup>2</sup>

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<sup>&</sup>lt;sup>1</sup> Warnell School of Forestry and Natural Resources, University of Georgia

<sup>&</sup>lt;sup>2</sup> Caesar Kleberg Wildlife Research Institute at Texas A&M University-Kingsville

<sup>&</sup>lt;sup>3</sup> Florida Fish and Wildlife Conservation Commission

 $<sup>{\</sup>rm *Corresponding\ author:\ rchandler@warnell.uga.edu}$ 

## Posterior distribution

The posterior distribution of the joint spatial capture-recapture movement model (with constant  $\sigma$  and data augmentation) is

$$\left\{ \prod_{i=1}^{M} p(y_i^{\text{cap}}|z_i, p) \left\{ \prod_{t=1}^{T} \left\{ \prod_{j=1}^{J} p(y_{ijt}|z_i, \lambda_0, \sigma_{\text{det}}, \mathbf{u}_{it}) \right\} p(\mathbf{u}_{it}|\mathbf{u}_{i,t-1}, \mathbf{s}_i, \rho, \sigma) \right\} p(\mathbf{s}_i) p(z_i|\psi) \right\} \times p(p) p(\lambda_0) p(\sigma_{\text{det}}) p(\rho) p(\sigma) p(\psi) \tag{1}$$

where

$$p(y_i^{\text{cap}}|z_i, p) = \text{Bern}(y_i|z_i \times p)$$

$$p(y_{ijt}|z_i, \lambda_0, \sigma_{\text{det}}, \mathbf{u}_{it}) = \text{Pois}(y_{ijt}|z_i \times \lambda_{ijt}^{\text{det}})$$

$$p(\mathbf{u}_{it}|\mathbf{u}_{i,t-1}, \mathbf{s}_i, \rho, \sigma) = \begin{cases} \text{Norm}(\mathbf{u}_{it}|\mathbf{s}_i + (\mathbf{u}_{i,t-1} - \mathbf{s}_i)\rho, \text{diag}(\sigma^2 - \sigma^2\rho^2)) & \text{for } t > 1\\ \text{Norm}(\mathbf{u}_{it}|\mathbf{s}_i, \text{diag}(\sigma^2)) & \text{for } t = 1 \end{cases}$$

$$p(\mathbf{s}_i) = \text{Unif}(\mathcal{M})$$

$$p(z_i|\psi) = \text{Bern}(z_i|\psi)$$

and the other probability distributions are priors for the parameters. Note that some or all of the  $\mathbf{u}_{it}$  locations could be observed.

## Gibbs sampler

Sampling from the joint posterior is computationally challenging because of the latent movement paths for the M-n augmented individuals. The burden can be reduced by marginalizing these latent paths while retaining the activity centers  $\{s_i\}$ . This is accomplished using Eq. 6 in the manuscript and the probability density:

$$p(0|\lambda_0, \sigma_{\text{det}}, \rho, \sigma, \mathbf{s}_i, z_i) = \text{Bern}(0|\tilde{p}_i)$$

The Gibbs sampler begins by initializing the unknown parameters and then sampling from the following full conditional distributions.

Use Metropolis-Hastings (MH) to sample from:

$$p(\rho, \sigma|\cdot) \propto \left\{ \prod_{i=1}^{n} \prod_{t=1}^{T} p(\mathbf{u}_{it}|\mathbf{u}_{i,t-1}, \mathbf{s}_{i}, \rho, \sigma) \right\} \left\{ \prod_{i=n+1}^{M} p(0|\lambda_{0}, \sigma_{\text{det}}, \rho, \sigma, \mathbf{s}_{i}, z_{i}) \right\} p(\rho)p(\sigma)$$

Use MH to sample from:

$$p(\lambda_0, \sigma_{\text{det}}|\cdot) \propto \left\{ \prod_{i=1}^n \prod_{j=1}^J \prod_{t=1}^T p(y_{ijt}|z_i, \lambda_0, \sigma_{\text{det}}, \mathbf{u}_{it}) \right\} \left\{ \prod_{i=n+1}^M p(0|\lambda_0, \sigma_{\text{det}}, \rho, \sigma, \mathbf{s}_i, z_i) \right\} p(\lambda_0) p(\sigma_{\text{det}})$$

Sample directly from

$$p(\psi|\boldsymbol{z}) = \text{Beta}\left(1 + \sum_{i=1}^{M} z_i, 1 + M - \sum_{i=1}^{M} z_i\right)$$

For i = n + 1, ..., M, use MH to sample from

$$p(z_i|\cdot) \propto p(y_i^{\text{cap}}|z_i, p) \left\{ \prod_{i=1}^J \prod_{t=1}^T p(y_{ijt}|z_i, \lambda_0, \sigma_{\text{det}}, \mathbf{u}_{it}) \right\} p(0|\lambda_0, \sigma_{\text{det}}, \rho, \sigma, \mathbf{s}_i, z_i) p(z_i|\psi)$$

Use MH (or direct draw from beta full conditional) to sample from

$$p(p|\cdot) \propto \prod_{i=1}^{M} p(y_i^{\text{cap}}|z_i \times p)p(p)$$

For i = 1, ..., n, use MH to sample from

$$p(\mathbf{s}_i|\cdot) \propto p(\mathbf{u}_{it}|\rho,\sigma,\mathbf{s}_i)p(\mathbf{s}_i)$$

For  $i = n + 1, \dots, M$ , use MH to sample from

$$p(\mathbf{s}_i|\cdot) \propto p(0|\lambda_0, \sigma_{\text{det}}, \rho, \sigma, \mathbf{s}_i, z_i)p(\mathbf{s}_i)$$

For i = 1, ..., n and for cases where  $\mathbf{u}_{it}$  is not observed, use MH to sample from

$$p(\mathbf{u}_{it}|\cdot) \propto p(\mathbf{u}_{i,t+1}|\mathbf{u}_{i,t},\mathbf{s}_i,\rho,\sigma)p(\mathbf{u}_{i,t}|\mathbf{u}_{i,t-1},\mathbf{s}_i,\rho,\sigma)\prod_{i=1}^{J}p(y_{ijt}|\lambda_0,\sigma_{\mathrm{det}},\mathbf{u}_{it})$$