

APPLIED STATISTICS

Variable Selection

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Overview

- Motivation
- Sequential Variable Selection
- Variable Selection Among All Subsets
- Cross Validation for Variable Selection Results
- Multicollinearity

References

1. **F.L. Ramsey and D.W. Schafer** (2012)
Chapter 12 of *The Statistical Sleuth*
2. The slides are made by **R Markdown**.
<http://rmarkdown.rstudio.com>

Motivation

There are two prime reasons for variable selection:

1. Simple models with less variables are preferable to complex models with more variables.
2. Including unnecessary variables in a model results in a loss of precision
⇒ overfitting.

Variable selection involves choosing a subset of explanatory variables to construct the multiple linear regression model.

Because if the explanatory variables selected in MLR are determined, then the MLR model with those explanatory variables is given.

Hence, sometimes we also call model selection.

Different subsets of explanatory variables determine different models. We call those models candidate models.

Motivation Example: Significance Depends on Other Explanatory Variables in the Model (Con'd)

Suppose we are interested in predicting ANU students' 2nd year GPA (Y) given their 1st year GPA (X_1) and UAC score (X_2). The following regression line is fit:

$$\mu\{Y|X_1, X_2\} = \beta_0 + \underline{\beta_1 X_1} + \underline{\beta_2 X_2}. \quad (1)$$

Based on the data, the p -values for the t -tests of whether $\beta_j = 0$ versus $\beta_j \neq 0$ for $j = 1, 2$ are 0.15 and 0.20, respectively.

Does this mean that we **do not need to select both X_1 and X_2 in the model?** NO!

The test for β_2 tells us whether X_2 is needed in the model that already contains X_1 , i.e., does X_2 offer any information about mean GPA over and above that of X_1 ?

The meaning of the coefficient of an explanatory variable depends on what other explanatory variables have been included in the regression.

Motivation Example: Significance Depends on Other Explanatory Variables in the Model (Con'd)

If we fit the following two models:

$$\mu\{Y|X_1\} = \alpha_0 + \alpha_1 X_1 \text{ and } \mu\{Y|X_2\} = \gamma_0 + \gamma_2 X_2.$$

For both models, the p-values for the t-tests of $\alpha_1 = 0$ versus $\alpha_1 \neq 0$ and $\gamma_2 = 0$ versus $\gamma_2 \neq 0$ can be computed. Based on the data, the results of the p-values are 0.01 and 0.02, respectively.

Hence at least one of X_1 and X_2 is needed in the model.

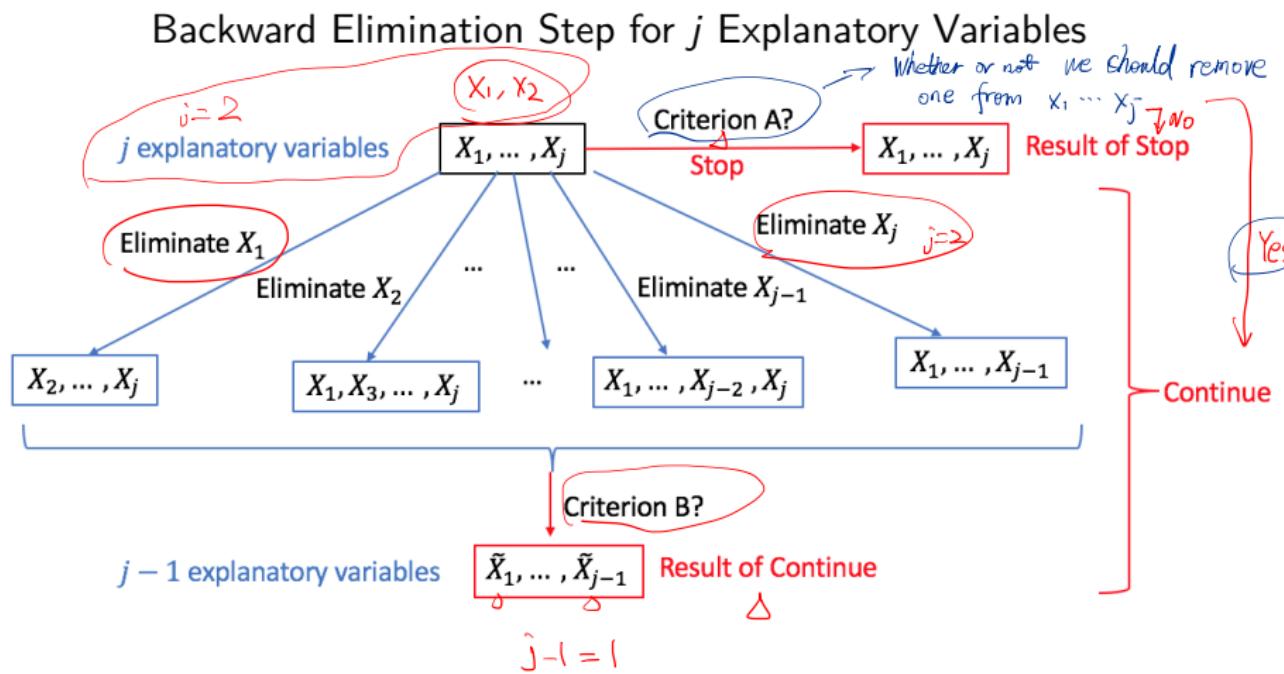
In this example X_1 and X_2 are probably highly correlated so we might expect this to be the case. The following F -test of model (1) avoids this problem.

H_0 : none of X_1 and X_2 is needed in the model \leftrightarrow X

H_a : at least one of X_1 and X_2 is needed in the model.

However, the F -test does not answer: which of X_1 and X_2 should be selected in the model.

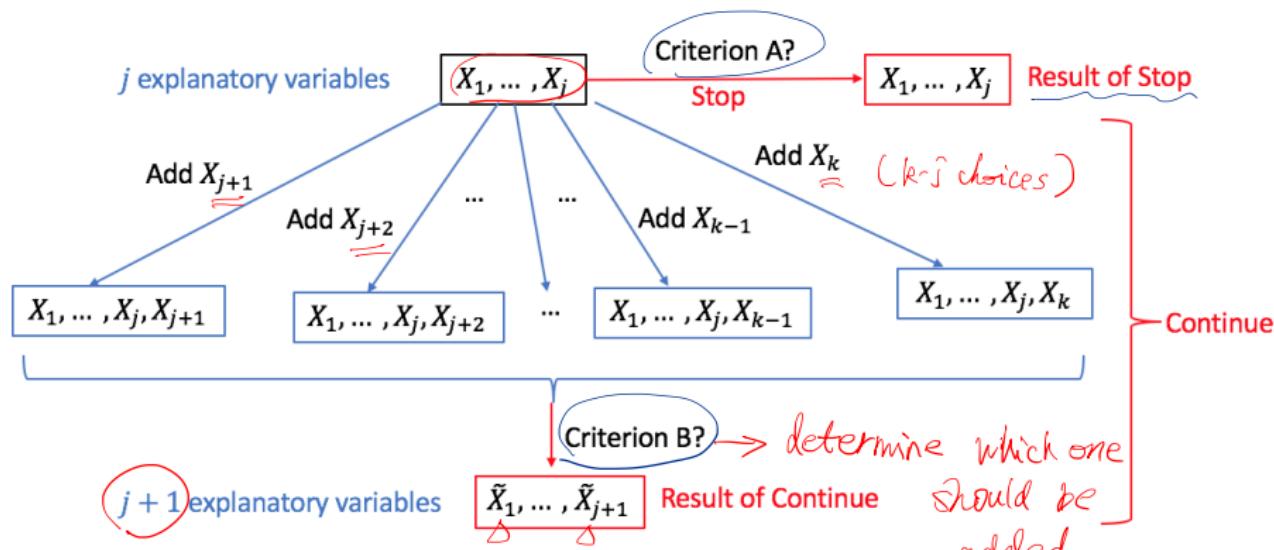
Sequential Variable Selection - Backward Elimination Steps



Sequential Variable Selection - Forward Selection Steps

0 → stop
↓ continue
1 → stop
↓ continue

Forward Selection Step for j Explanatory Variables



Suppose we have k explanatory variables X_1, \dots, X_k in total.

Sequential Variable Selection

The idea behind sequential techniques is a sequential search through all possible combinations of variables by either adding or removing a single explanatory variable from the current candidate model at each step.

In order to accomplish the sequential variable selection, we need to determine Criterion A and Criterion B in forward selection steps and backward elimination steps.

Usually the criterion depends on statistical measures.

Sequential Variable Selection by Using F -Statistic

We start from using F -statistic as the measure. The F -statistic is

$$F\text{-Stat} = \frac{(SSE_{\text{reduced}} - SSE_{\text{full}})/d}{\hat{\sigma}_{\text{full}}^2},$$

which is used to test

H_0 : the **reduced model** is appropriate \Leftrightarrow

H_a : the **full model** is appropriate.

1 variable
less than
the full model
in sequential
variable selection
Hence, $d=1$.

The p -value of F -test $< \alpha$ (usually 0.05) \Leftrightarrow F -Stat is too large.

\Rightarrow

Reject H_0 .

\Rightarrow

The **full model** is preferred.

Otherwise, the **reduced model** is preferred

Sequential Variable Selection by Using F -Statistic (Con'd)

One can verify that the p -value of the F -test $< 0.05 \Leftrightarrow F\text{-Stat is larger than approximately 4}$, especially when sample size n is very large.

only for $d=1$.

Hence a predefined "cut-off" for F -Stat is usually 4.

$d=1$

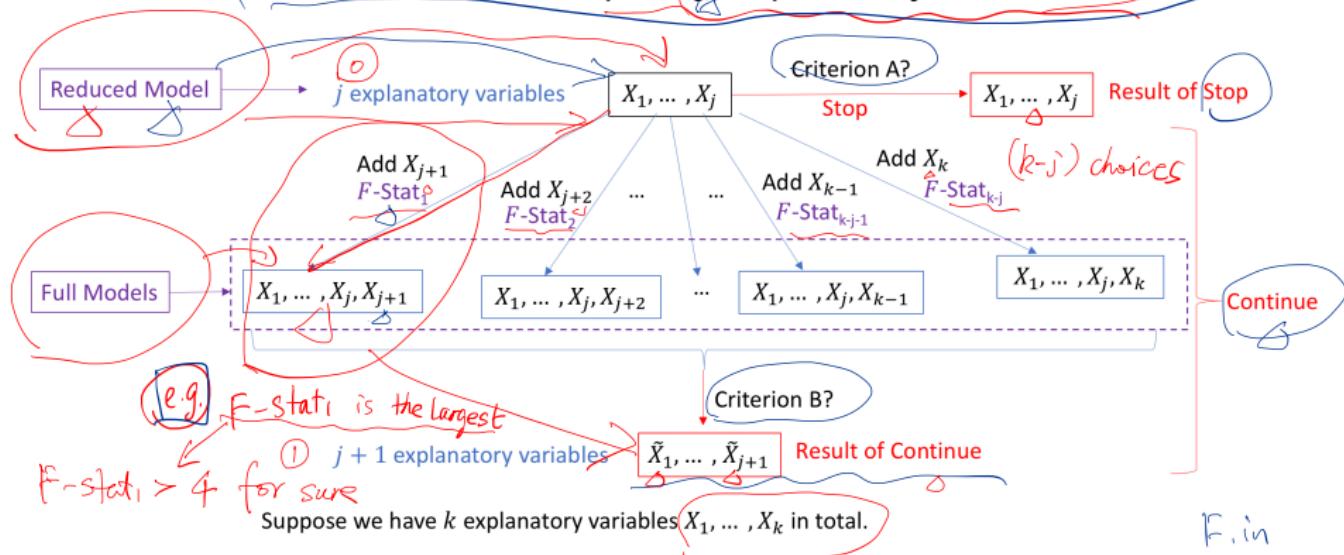
$F\text{-stat} > 4 \Rightarrow p\text{-value} < 0.05 \Rightarrow \text{Reject } H_0.$
\Rightarrow
The full model is preferred.

$F\text{-stat} < 4 \Rightarrow p\text{-value} > 0.05 \Rightarrow \text{Not reject } H_0.$
\Rightarrow
The reduced model is preferred.

We will explain the reduced model and the full model in forward selection steps and backward elimination steps.

Forward Selection Step by Using F -Statistic

Forward Selection Step for j Explanatory Variables

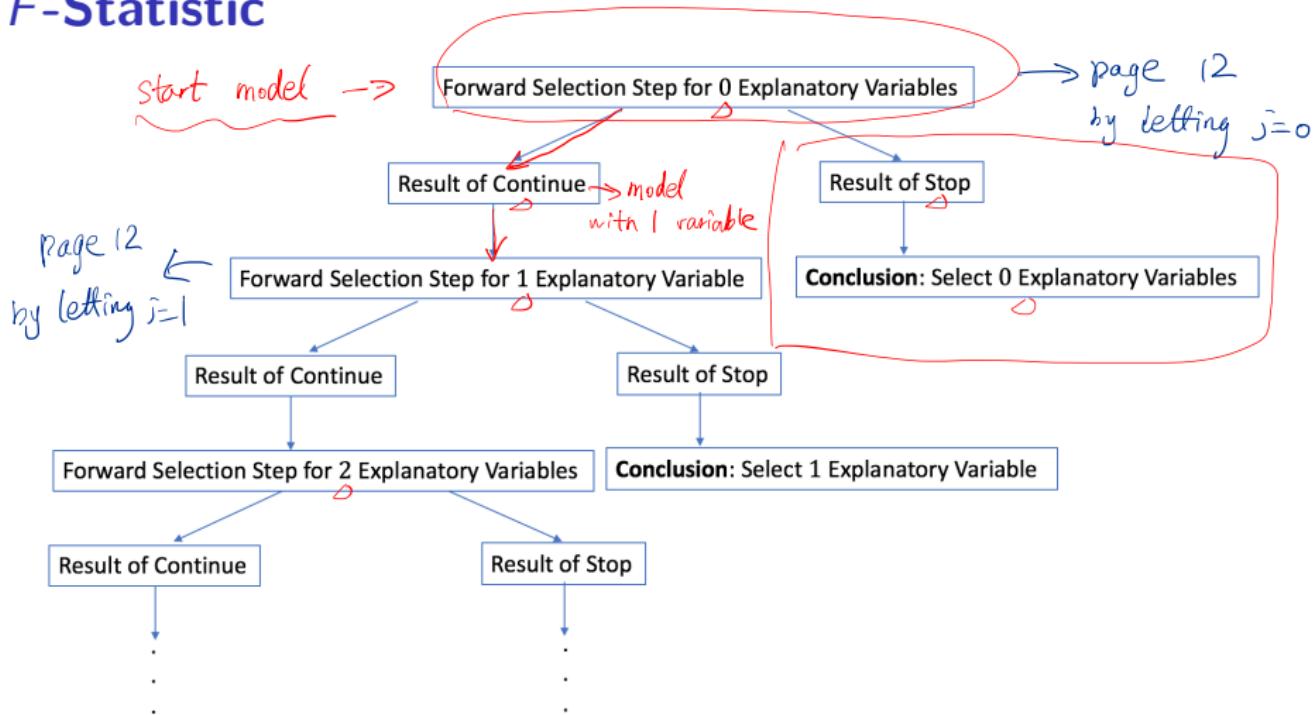


Criterion A: if $\max\{F\text{-Stat}_1, F\text{-Stat}_2, \dots, F\text{-Stat}_{k-j-1}, F\text{-Stat}_{k-j}\} < 4$ or $j = k$, then **Stop**; otherwise **Continue**.

is not satisfied

Criterion B: $\tilde{X}_1, \dots, \tilde{X}_{j+1}$ are those variables such that the corresponding F -Stat is the largest.

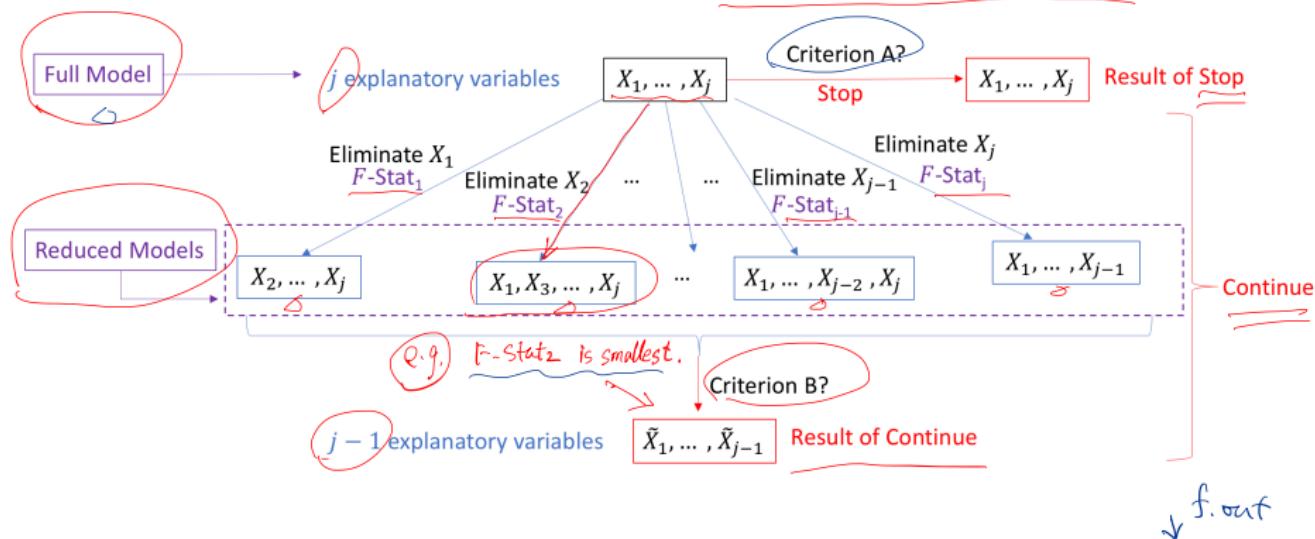
Complete Forward Selection Procedure by Using F-Statistic



Keep doing on the above procedures, until the first time we obtain the **Result of Stop**.

Backward Elimination Step by Using F -Statistic

Backward Elimination Step for j Explanatory Variables

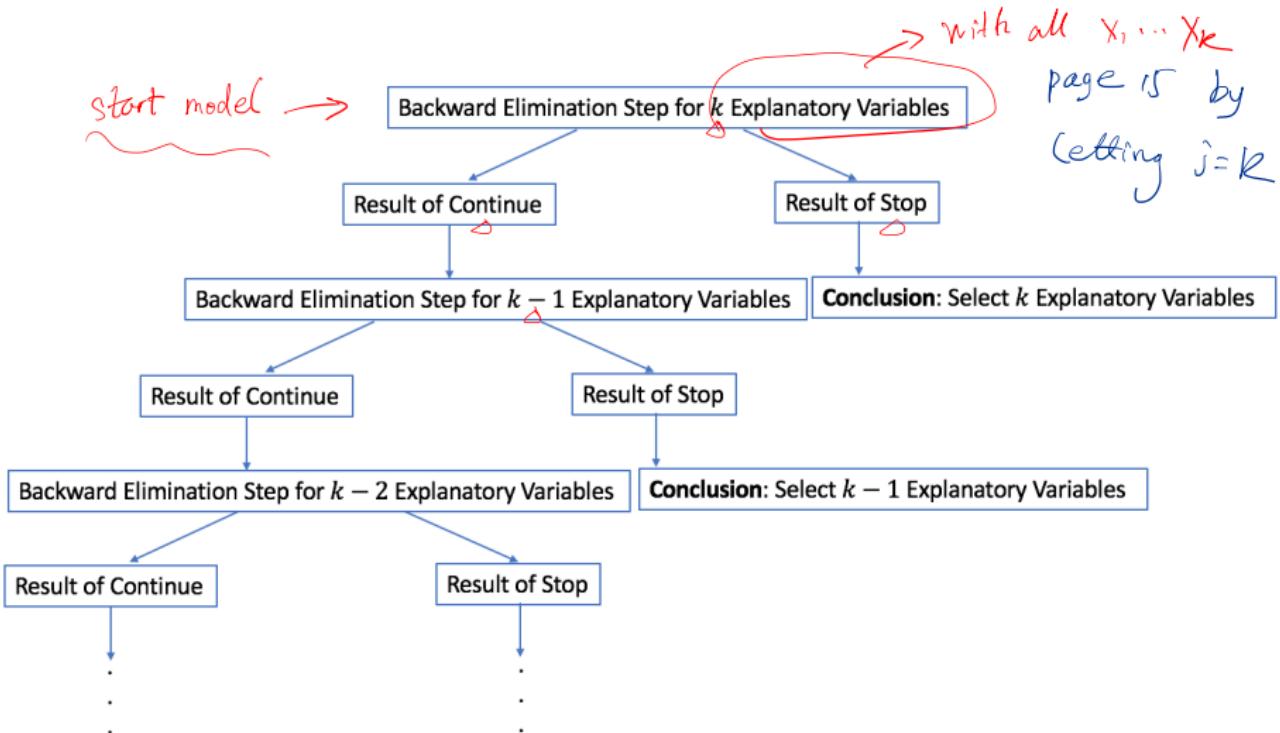


Criterion A: if $\min\{F\text{-Stat}_1, F\text{-Stat}_2, \dots, F\text{-Stat}_{j-1}, F\text{-Stat}_j\} > 4$ or
 $j = 0$, then Stop; otherwise Continue.

$$\min \{ F\text{-Stat}_1, \dots, F\text{-Stat}_j \} < 4$$

Criterion B: $\tilde{X}_1, \dots, \tilde{X}_{j-1}$ are those variables such that the corresponding F -Stat is the smallest.

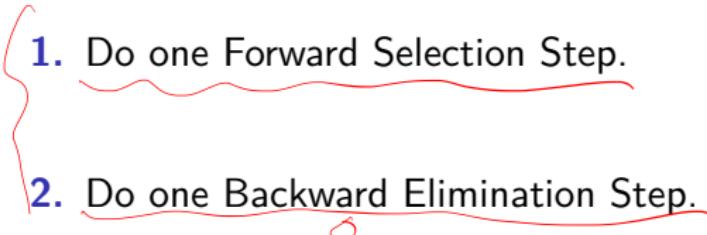
Complete Backward Elimination Procedure by Using F-Statistic



Keep doing on the above procedures, until the first time we obtain the Result of Stop.

Stepwise Selection Procedure by Using F-Statistic

This constitutes a combination of backward elimination and forward selection. Each step consists of the following steps. (Any starting model can be chosen, e.g., the model with no explanatory variables, or the model with all the explanatory variables.)

- 
1. Do one Forward Selection Step.
 2. Do one Backward Elimination Step.

Repeat Steps 1 and 2 until no explanatory variables can be added or removed.

Remark: The above three methods will sometimes lead to different variable selection results.

Example: SAT Scores

```
rm(list=ls())
library('Sleuth3')
SATdata=case1201
head(SATdata)
```

##	State	SAT	Takers	Income	Years	Public	Expend	Rank
## 1	Iowa	1088	3	326	16.79	87.8	25.60	89.7
## 2	SouthDakota	1075	2	264	16.07	86.2	19.95	90.6
## 3	NorthDakota	1068	3	317	16.57	88.3	20.62	89.8
## 4	Kansas	1045	5	338	16.30	83.9	27.14	86.3
## 5	Nebraska	1045	5	293	17.25	83.6	21.05	88.5
## 6	Montana	1033	8	263	15.91	93.7	29.48	86.4

SATdata=SATdata[-29,] #removing Alaska

Y<-SATdata[,2]

X<-SATdata[,-c(1,2)] → We drop 1st 2nd columns of SATdata.

X<-as.matrix(X) → dataframe

→ matrix

Example: SAT Scores (Con'd)

```
#install.packages('wle')
library(wle) #need to load this library!

## Loading required package: circular

## Warning: package 'circular' was built under R version 3.3.2

## 
## Attaching package: 'circular'

## The following objects are masked from 'package:stats':
## 
##     sd, var

mle.stepwise(Y ~ X, f.in=4, f.out=4, type="Forward")

## Call: mle.stepwise(formula = Y ~ X, type = "Forward", f.in = 4, f.out = 4)
## 
## Forward selection procedure
## 
## F.in: 4

## Last 4 iterations:
##   (Intercept) XTakers XIncome XYears XPublic XExpend XRank
## [1,] 1 0 0 0 0 0 162.400
## [2,] 1 0 0 0 0 1 1 22.820
## [3,] 1 0 0 1 0 1 1 16.320
## [4,] 1 0 1 1 0 1 1 5.595
```

see page 12

Forward Selection Step for 0 Variables
Result of Continue

Forward selection step for 1 variables
Result of Continue

Result of stop

conclusion

Example: SAT Scores (Con'd)

```
full=lm(Y-SATdata$Rank) #full model  
reduced = lm(Y - 1) #reduced model  
anova(reduced,full,test='F')  
    △ △ △
```

← 1st row of the table

```
## Analysis of Variance Table  
##  
## Model 1: Y ~ 1  
## Model 2: Y ~ SATdata$Rank  
##   Res.Df   RSS Df Sum of Sq      F    Pr(>F)  
## 1     48 245376  
## 2     47 55079  1    190297 162.38 < 2.2e-16 ***  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
full=lm(Y-SATdata$Rank+SATdata$Expend) #full model  
reduced = lm(Y-SATdata$Rank) #reduced model  
anova(reduced,full,test='F')
```

← 2nd row of the table

```
## Analysis of Variance Table  
##  
## Model 1: Y ~ SATdata$Rank  
## Model 2: Y ~ SATdata$Rank + SATdata$Expend  
##   Res.Df   RSS Df Sum of Sq      F    Pr(>F)  
## 1     47 55079  
## 2     46 36815  1    18265 22.822 1.849e-05 ***  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

Example: SAT Scores (Con'd)

```
mle.stepwise(Y~X,f.in=4,f.out=4,type="Backward")
```

```
##  
## Call:  
## mle.stepwise(formula = Y ~ X, type = "Backward", f.in = 4, f.out = 4)
```

```
##  
##  
## Backward selection procedure  
##
```

```
## F.out: 4
```

```
##  
## Last 2 iterations:
```

```
## (Intercept) XTakers XIncome XYears XPublic XExpend XRank  
## [1,] 1 0 1 1 1 1 1 0.000818  
## [2,] 1 0 1 1 0 1 1 0.770000
```

/L18

Example: SAT Scores (Con'd)

```
mle.stepwise(Y~X,f.in=4,f.out=4,type="Stepwise")
```

```
##  
## Call:  
## mle.stepwise(formula = Y ~ X, type = "Stepwise", f.in = 4, f.out = 4)  
##  
##
```

```
## Stepwise selection procedure  
##  
## F.in: 4  
## F.out: 4
```

```
## Last 5 iterations:
```

	(Intercept)	XTakers	XIncome	XYears	XPublic	XExpend	XRank	
## [1,]	1	0	0	0	0	0	1	1.624e+02
## [2,]	1	0	0	0	0	1	1	2.282e+01
## [3,]	0	0	0	0	0	1	1	1.551e-03
## [4,]	0	0	0	0	1	1	1	1.102e+01
## [5,]	0	0	0	1	1	1	1	6.412e+00

$$1.624 \times 10^2 = 162.4$$

Backward step based on 1 variable
is hidden

Sequential Variable Selection by Using Other Statistical Measures

Recall that in order to accomplish the sequential variable selection, we need to determine Criterion A and Criterion B in forward selection steps and backward elimination steps.

Usually the criterion depends on statistical measures.

We start from using F-statistic as the measure and we adopt the F-test idea to determine Criterion A and Criterion B.

However, other statistical measures can also be used to determine Criterion A and Criterion B. But the idea is different from the F-test.

Idea of Variable Selection by Using Other Statistical Measures

Recall that we pursue good fitting for a MLR model, and SSE (deviance) measures the goodness of fit for MLR.

Based on the definition of SSE (deviance), the smaller the SSE (deviance) is, the better fitting of a model.

Hence, one goal for MLR is to find an appropriate model with smaller SSE.

However, the goal for variable selection is to find a small number of explanatory variables if possible to construct MLR.

The above two goals are contradicted since more explanatory variables results in the decrease in SSE.

Hence, an appropriate variable selection criterion should provide a compromise between how well the model fits the data and the number of explanatory variables \Rightarrow a standard to determine statistical measures.

Idea of Variable Selection by Using Adjusted R-squared

all variables $x_1 \dots x_k$

Adjusted R-squared can be considered as one possible statistical measure for variable selection. For the following MLR model

$$\mu\{Y|X_1, \dots, X_j\} = \beta_0 + \beta_1 X_1 + \dots + \beta_j X_j, \text{ we have}$$

$$\text{Adjusted } R^2 = 1 - \frac{\text{SSE}/\{n - (j+1)\}}{\text{SST}/(n-1)}.$$

$(x_1 \dots x_j)$ $(x_2 \dots x_{j+1})$

For two candidate models, if their number of the explanatory variables j is the same, then the model with smaller SSE, or equivalently larger adjusted R-squared, is preferred.

For two candidate models, if their SSE is the same, then the model with smaller number of explanatory variables j , or equivalently larger adjusted R-squared, is preferred.

Hence, the variable selection criterion based on adjusted R-squared is: the model with larger adjusted R-squared is preferred.

Idea of Variable Selection by Using AIC and BIC

AIC (Akaike Information Criterion) and BIC (Bayesian) can also be considered. For the following MLR model

$$\mu\{Y|X_1, \dots, X_j\} = \beta_0 + \beta_1 X_1 + \dots + \beta_j X_j, \text{ we define}$$

$$\left\{ \begin{array}{l} \text{AIC} = n \left\{ \log \left(\frac{\text{SSE}}{n} \right) + 1 + \log(2\pi) \right\} + 2 \times (j+1) \text{ and} \\ \text{BIC} = n \left\{ \log \left(\frac{\text{SSE}}{n} \right) + 1 + \log(2\pi) \right\} + \overset{n>2}{\text{log}(n)} \times (j+1). \end{array} \right.$$

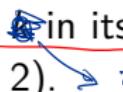
For two candidate models, if their number of the explanatory variables j is the same, then the model with smaller SSE, or equivalently **smaller AIC (or BIC)**, is preferred.

For two candidate models, if their SSE is the same, then the model with smaller number of explanatory variables j , or equivalently **smaller AIC (or BIC)**, is preferred.

Hence, the variable selection criterion based on AIC (or BIC) is: **the model with smaller AIC (or BIC) is preferred.**

Idea of Variable Selection by Using Other Statistical Measures

Adjusted R-squared, AIC and BIC all compromise how well the model fits the SSE and the number of explanatory variables (j).

Compared to AIC, BIC assigns a larger weight to the number of explanatory variables in its expression (usually the sample size n is large such that $\log(n) > 2$). 

Hence BIC usually prefers the model with less explanatory variables compared to AIC.

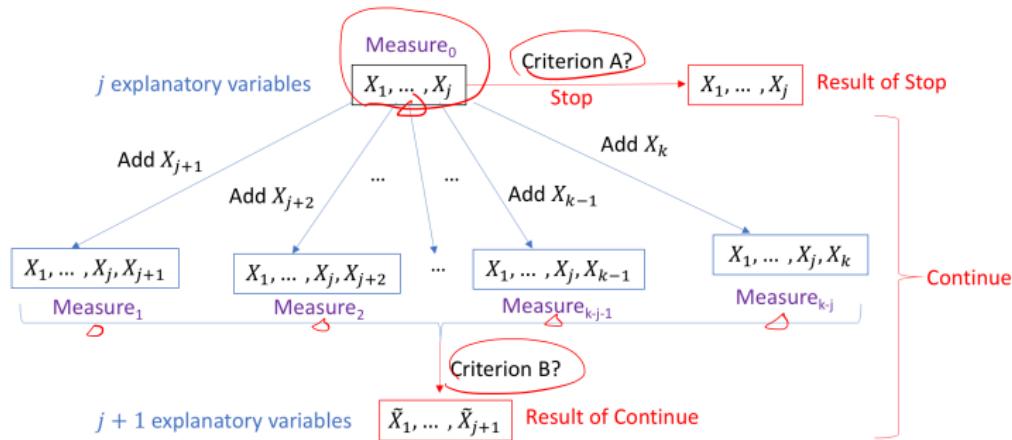
In the following, we will introduce the sequential variable selection procedure based on one of the above three measures.

Let Measure = $-1 \times$ Adjusted R^2 , or Measure = AIC, or Measure = BIC in the following.

Then, the variable selection criterion based on "Measure" is: the model with smaller "Measure" is preferred.

Forward Selection Step by Using Other Statistical Measures

Forward Selection Step for j Explanatory Variables



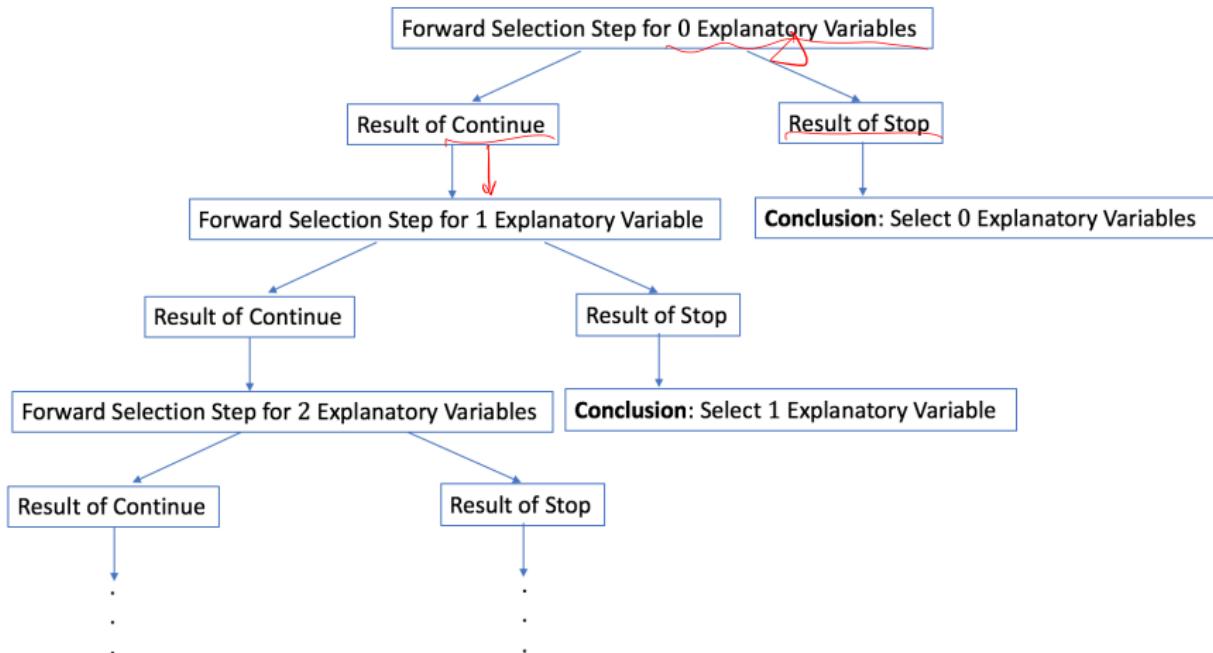
Suppose we have k explanatory variables X_1, \dots, X_k in total.

Criterion A: if

$\min\{Measure_1, Measure_2, \dots, Measure_{k-j-1}, Measure_{k-j}\} > Measure_0$ or
 $j = k$, then Stop; otherwise Continue.

Criterion B: $\tilde{X}_1, \dots, \tilde{X}_{j+1}$ are those variables such that the corresponding "Measure" is the smallest.

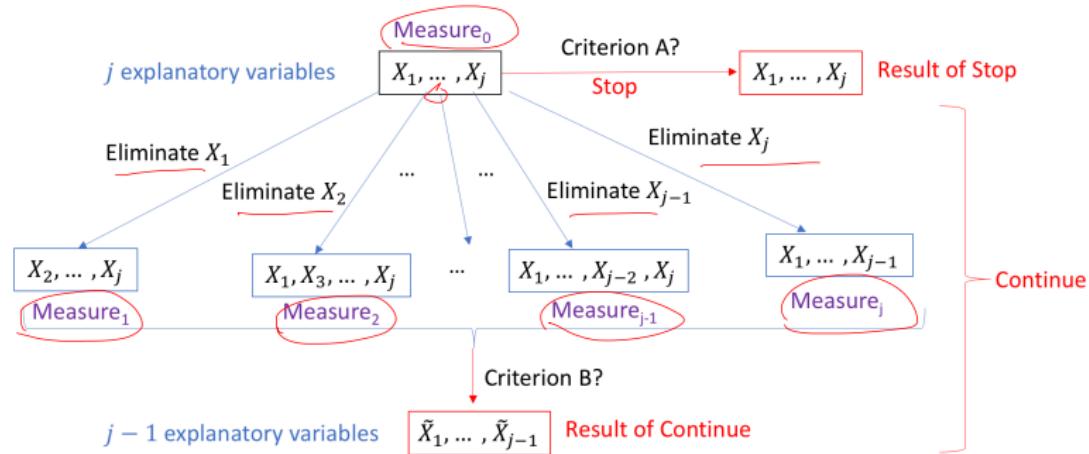
Complete Forward Selection Procedure by Using Other Statistical Measures



Keep doing on the above procedures, until the first time we obtain the **Result of Stop**.

Backward Elimination Step by Using Other Statistical Measures

Backward Elimination Step for j Explanatory Variables

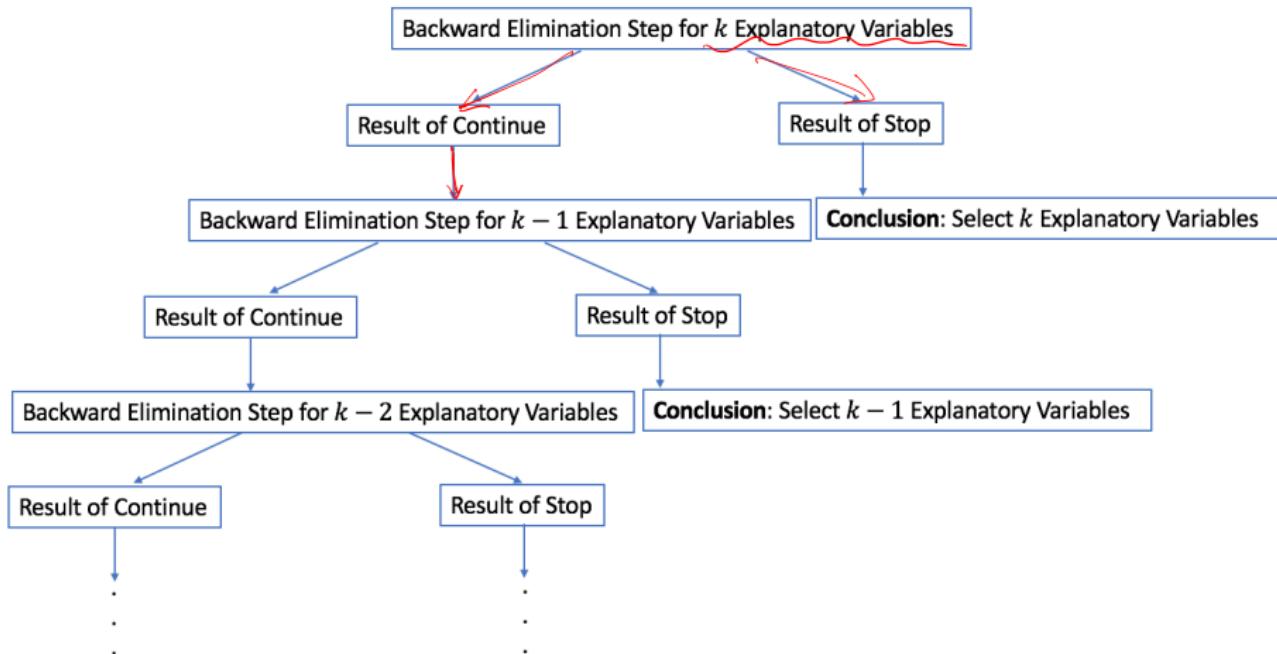


Criterion A: if

$\min\{\text{Measure}_1, \text{Measure}_2, \dots, \text{Measure}_{j-1}, \text{Measure}_j\} > \text{Measure}_0$ or $j = 0$,
then Stop; otherwise Continue.

Criterion B: $\tilde{X}_1, \dots, \tilde{X}_{j+1}$ are those variables such that the corresponding
"Measure" is the smallest.

Complete Backward Elimination Procedure by Using Other Statistical Measures



Keep doing on the above procedures, until the first time we obtain the **Result of Stop**.

Stepwise Selection Procedure by Other Statistical Measures

This constitutes a combination of backward elimination and forward selection. Each step consists of the following steps. (Any starting model can be chosen, e.g., the model with no explanatory variables, or the model with all the explanatory variables.)

1. Do one Forward Selection Step.
2. Do one Backward Elimination Step.

Repeat Steps 1 and 2 until no explanatory variables can be added or removed.

Remark: The above three methods will sometimes lead to different variable selection results.

Example: SAT Scores (Con'd)

```
X<-data.frame(X)
fit<-lm(Y~.,data=X)
#install.packages('MASS')
library(MASS)
```

Warning: package 'MASS' was built under R version 3.3.2

```
#Backward AIC
a=stepAIC(fit,direction="backward",data=X)
```

```
summary(a)

##
## Call:
## lm(formula = Y ~ Income + Years + Expend + Rank, data = X)
##
## Residuals:
##    Min      1Q  Median      3Q     Max
## -47.005 -15.548   1.759  13.534  51.808
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)
## (Intercept) -255.2894    88.6074 -2.881 0.006103 **
## Income       0.2412     0.1020  2.365 0.022479 *
## Years        18.9447    5.1563  3.674 0.000644 ***
## Expend       3.3851     0.7775  4.354 7.87e-05 ***
## Rank         9.3764     0.6589 14.229 < 2e-16 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 23.34 on 44 degrees of freedom
## Multiple R-squared:  0.9023, Adjusted R-squared:  0.8934
## F-statistic: 101.6 on 4 and 44 DF,  p-value: < 2.2e-16
```

model with all variables

dataset for X variables

return a lot of R output.

final result for backward selection

each backward step

Example: SAT Scores (Con'd)

```
#Backward BIC  
n=length(Y) → sample size  
a=stepAIC(fit,direction="backward",data=X, k=log(n))
```

```
summary(a)
```

```
##  
## Call:  
## lm(formula = Y ~ Income + Years + Expend + Rank, data = X)  
##  
## Residuals:  
##      Min       1Q   Median       3Q      Max  
## -47.005 -15.548   1.759  13.534  51.808  
##  
## Coefficients:  
##             Estimate Std. Error t value Pr(>|t|)  
## (Intercept) -255.2894    88.6074 -2.881 0.006103 **  
## Income       0.2412     0.1020  2.365 0.022479 *  
## Years        18.9447    5.1563  3.674 0.000644 ***  
## Expend       3.3851     0.7775  4.354 7.87e-05 ***  
## Rank         9.3764     0.6589 14.229 < 2e-16 ***  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##  
## Residual standard error: 23.34 on 44 degrees of freedom  
## Multiple R-squared:  0.9023, Adjusted R-squared:  0.8934  
## F-statistic: 101.6 on 4 and 44 DF,  p-value: < 2.2e-16
```

assigns weight $\log(n)$ to $(j+1)$

↑

BIC formula

Example: SAT Scores (Con'd)

```
#Stepwise AIC  
a=stepAIC(fit,direction="both",data=X)
```

```
summary(a)
```

```
##  
## Call:  
## lm(formula = Y ~ Income + Years + Expend + Rank, data = X)  
##  
## Residuals:  
##      Min      1Q  Median      3Q     Max  
## -47.005 -15.548    1.759   13.534   51.808  
##  
## Coefficients:  
##             Estimate Std. Error t value Pr(>|t|)  
## (Intercept) -255.2894    88.6074 -2.881 0.006103 **  
## Income      0.2412     0.1020   2.365 0.022479 *  
## Years       18.9447     5.1563   3.674 0.000644 ***  
## Expend      3.3851     0.7775   4.354 7.87e-05 ***  
## Rank        9.3764     0.6589  14.229 < 2e-16 ***  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##  
## Residual standard error: 23.34 on 44 degrees of freedom  
## Multiple R-squared:  0.9023, Adjusted R-squared:  0.8934  
## F-statistic: 101.6 on 4 and 44 DF,  p-value: < 2.2e-16
```

Example: SAT Scores (Con'd)

#Forward AIC

fit<-lm(Y~1,data=X) → fitted model for 0 variables
a=stepAIC(fit,direction="forward",scope=list(lower=-1,upper=~Takers + Income + Years + Public + Expend
+ Rank))

summary(a)

All Variables

```
##  
## Call:  
## lm(formula = Y ~ Rank + Expend + Years + Income, data = X)  
##  
## Residuals:  
##      Min      1Q  Median      3Q     Max  
## -47.005 -15.548   1.759  13.534  51.808  
##  
## Coefficients:  
##             Estimate Std. Error t value Pr(>|t|)  
## (Intercept) -255.2894    88.6074 -2.881 0.006103 **  
## Rank         9.3764     0.6589 14.229 < 2e-16 ***  
## Expend       3.3851     0.7775  4.354 7.87e-05 ***  
## Years        18.9447     5.1563  3.674 0.000644 ***  
## Income        0.2412     0.1020  2.365 0.022479 *  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##  
## Residual standard error: 23.34 on 44 degrees of freedom  
## Multiple R-squared:  0.9023, Adjusted R-squared:  0.8934  
## F-statistic: 101.6 on 4 and 44 DF,  p-value: < 2.2e-16
```

Variable Selection Among All Subsets \rightarrow different method

Variable selection involves choosing a subset of explanatory variables to construct the multiple linear regression model.

For instance, consider we have two explanatory variables X_1 and X_2 in total.
The subsets are

$$\emptyset, \underline{\{X_1\}}, \underline{\{X_2\}}, \text{ and } \underline{\underline{\{X_1, X_2\}}}. \quad \begin{matrix} \leftarrow & \text{one from these} \\ & \text{four subsets} \end{matrix}$$

Different subsets of explanatory variables determine different models. We call those models candidate models. \leftrightarrow a subset

The methods of variable selection among all subsets fit all possible candidate models (e.g., all the four subsets in the above example).

A value of the statistical measure (e.g., adjusted R-squared, AIC and BIC) is computed for each candidate model.

Let Measure = $-1 \times \text{Adjusted } R^2$, or Measure = AIC, or Measure = BIC.

Then, the model with the smallest "Measure" is preferred.

Variable Selection Among All Subsets (Con'd)

list all subsets

$x_1 \dots x_{100}$

2^{100}

subsets

The variable selection among all subsets is a search through all possible subsets of variables, in order to obtain the resulting model with the smallest "Measure", which is an alternative method for variable selection and is different from the sequential variable selection techniques.

→ is computational less complicated.

The sequential techniques is a sequential search by either adding or removing a single explanatory variable from the current candidate model at each step. The "Measure" is used to determine Criterion A and Criterion B in forward selection steps and backward elimination steps.

subsets x_1, x_2 all variable

\emptyset

→ measure 1

$\{x_1\}$

→ fitted model → measure 2

$\{x_2\}$

→ measure 3

$\{x_1, x_2\}$

→ measure 4

forward

\emptyset

forward steps

path example:

$\emptyset \rightarrow \{x_1\} \rightarrow \{x_1, x_2\}$

continue
e.g. $\{x_1\}$

stop → conclusion

backward example:
 $\{x_1, x_2\} \rightarrow \{x_2\} \rightarrow \emptyset$

/L20 forward steps

Idea of Variable Selection by Using C_p -Statistic

For the following MLR model with **all** the explanatory variables

$$\mu\{Y|X\} = \beta_0 + \beta_1 X_1 + \cdots + \beta_k X_k, \text{ where } X = (X_1, \dots, X_k),$$

we can compute $\hat{\sigma}_{\text{all}}^2$.

We consider a new statistical measure C_p -statistic. For the following candidate model

→ subset $\{x_1 \dots x_j\}$

size of a candidate model is $j+1$

$$\mu\{Y|X_1, \dots, X_j\} = \beta_0 + \beta_1 X_1 + \cdots + \beta_j X_j,$$

we can compute its SSE. The C_p -statistic is defined as

$$C_p = (j+1) + (n-j-1) \frac{\text{SSE}/(n-j-1) - \hat{\sigma}_{\text{all}}^2}{\hat{\sigma}_{\text{all}}^2} = \boxed{\frac{\text{SSE}}{\hat{\sigma}_{\text{all}}^2} + 2(j+1) - n.}$$

stays the same for different $(x_1 \dots x_j)$

Idea of Variable Selection by Using C_p -Statistic (Con'd)

For two candidate models, if their number of the explanatory variables j is the same, then the model with smaller SSE, or equivalently **smaller** C_p , is preferred.

For two candidate models, if their SSE is the same, then the model with smaller number of explanatory variables j , or equivalently **smaller** C_p , is preferred.

Hence, the variable selection criterion based on C_p -statistic is: the model with **smaller** C_p is preferred.

C_p -statistic also compromises how well the model fits the data (SSE) and the number of explanatory variables (j).

The variable selection among all subsets by using C_p -statistic is a search through all possible subsets of variables, in order to obtain the resulting model with the smallest C_p . ATC, BCC / largest Adj R^2

Cross Validation for Variable Selection Results

The above variable selection techniques will sometimes lead to different variable selection results.

Cross validation is a process of “checking” the selected model.

The original dataset is split into two parts:

obs 1 ... n

1. a training dataset: for model fitting and variable selection;



variable selection result

2. a test dataset (hold-out sample): for checking selection results and finding the best model.

↓ perform prediction

This procedure allows the performance of a model to be gauged on data that were not used to fit the model.

Mean Squared Prediction Error

We usually use the predictive performance of the test dataset to check the selection results and to find the best model.

A measure of predictive ability is given by the mean squared prediction error (MSPE):

$$\text{MSPE} = \frac{1}{n_{\text{test}}} \sum_{\ell=1}^{n_{\text{test}}} (Y_\ell - \hat{Y}_\ell)^2, \text{ where}$$

\downarrow
real
 \downarrow
prediction

obs: $\ell=1, 2, \dots, n_{\text{test}}$

- Y_ℓ is the observed ℓ -th response from the test dataset.

\approx
 $x_{1,\text{new}}, x_{2,\text{new}}, \dots, x_{j,\text{new}}$

- \hat{Y}_ℓ is the predicted value of Y_ℓ based on the regression model constructed by the training dataset.

$$\text{prediction } \hat{Y}_{\text{new}} = \hat{\beta}_0 + \hat{\beta}_1 x_{1,\text{new}} + \dots + \hat{\beta}_j x_{j,\text{new}}$$

$\hat{Y}(y|x_1, \dots, x_j) = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \dots + \hat{\beta}_j x_j$
 $\downarrow \quad \downarrow \quad \downarrow$
computed based on

- n_{test} is the number of the observations of the test dataset.

The best model is the model with the smallest MSPE.

Example: SAT Scores (Con'd)

```
#install.packages('leaps')  
library(leaps)
```

```
## Warning: package 'leaps' was built under R version 3.3.2
```

```
#help(leaps) #get some info on the leaps command
```

```
X=cbind(log(SATdata[,3]),SATdata[,4:8])
```

```
colnames(X)=c('x1','x2','x3','x4','x5','x6')
```

```
Y=SATdata[,2]
```

```
length(X[,1])
```

→ 1st column

```
## [1] 49 → sample size
```

```
#Training data
```

```
Xtraining=X[1:40,]
```

```
Ytraining=Y[1:40]
```

```
#Test data
```

```
Xtest=X[41:49,]
```

```
Ytest=Y[41:49]
```

subset with 1 variable
 $\{x_1\}$
 $\{x_2\}$
 \vdots
 $\{x_6\}$ # : 6

Example: SAT Scores (Con'd)

#Training data

```
Cpsel=leaps(Xtraining, Ytraining, method='Cp', nbest=2)
```

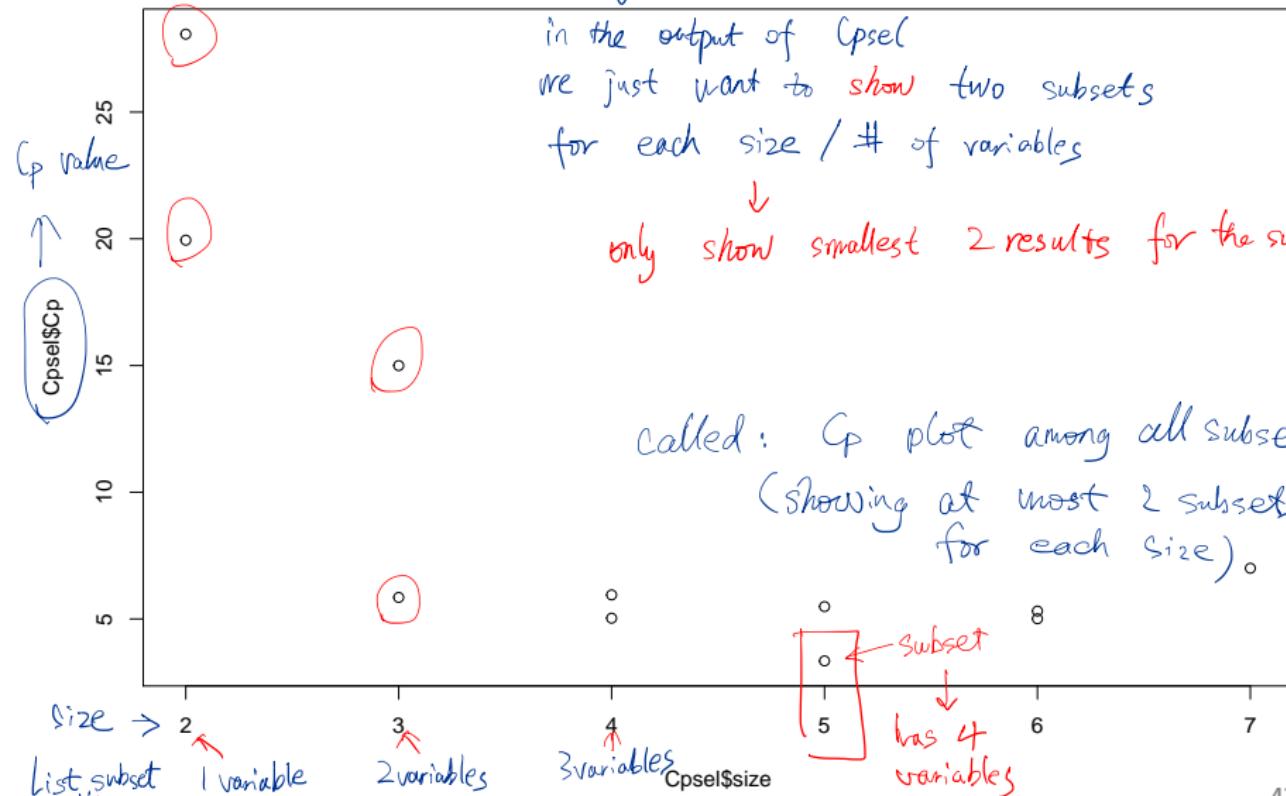
size : # of variables + 1
in a candidate model

(Cp, nbest=2)

in the output of Cpsel

we just want to show two subsets
for each size / # of variables

only show smallest 2 results for the subsets



Example: SAT Scores (Con'd)

```
cbind(Cpsel$which, Cpsel$size, Cpsel$Cp)
```

```
##   x1 x2 x3 x4 x5 x6
## 1 1 0 0 0 0 0 2 19.946845
## 1 0 0 0 0 0 1 2 28.076894
## 2 1 0 0 0 1 0 3 5.848913
## 2 1 0 0 1 0 0 3 14.996836
## 3 1 0 0 0 1 1 4 5.029187
## 3 1 0 0 1 1 0 4 5.948021
## 4 1 0 1 0 1 1 5 3.343196
## 4 1 0 1 1 1 0 5 5.487590
## 5 1 0 1 1 1 1 6 5.006153
## 5 1 1 1 0 1 1 6 5.310103
## 6 1 1 1 1 1 1 7 7.000000
```

of variables

size

Cp

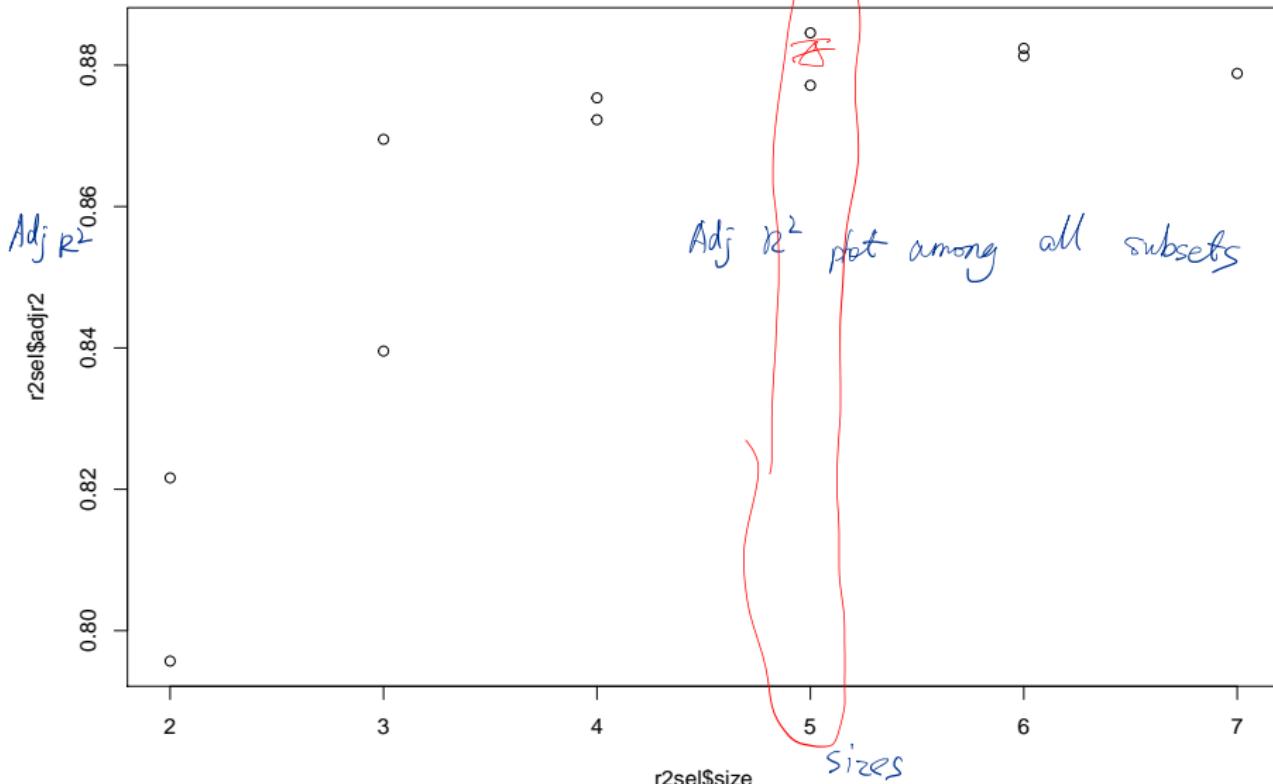
#Variable Selection Result: size=5, 1 0 1 0 1 1, x1, x3, x5, x6

size = 5

selected

Example: SAT Scores (Con'd)

```
r2sel=leaps(Xtraining,Ytraining,method='adjr2",nbest=2)  
plot(r2sel$size,r2sel$adjr2)
```



Example: SAT Scores (Con'd)

```
cbind(r2sel$which, r2sel$size, r2sel$adjr2)
```

```
##   1 2 3 4 5 6  
## 1 1 0 0 0 0 2 0.8216154  
## 1 0 0 0 0 0 1 2 0.7956931  
## 2 1 0 0 0 1 0 3 0.8695092  
## 2 1 0 0 1 0 0 3 0.8395530  
## 3 1 0 0 0 1 1 4 0.8753745  
## 3 1 0 0 1 1 0 4 0.8722821  
## 4 1 0 1 0 1 1 5 0.8845738  
## 4 1 0 1 1 1 0 5 0.8771504  
## 5 1 0 1 1 1 1 6 0.8823800  
## 5 1 1 1 0 1 1 6 0.8812968  
## 6 1 1 1 1 1 1 7 0.8788383
```

↓
size

#Variable Selection Result: size=5, 1 0 1 0 1 1, x1,x3,x5,x6

selection result



Example: SAT Scores (Con'd)

#Cross Validation

```
fitall=lm(Ytraining~.,data=Xtraining)
summary(fitall)
```

```
##  
## Call:  
## lm(formula = Ytraining ~ ., data = Xtraining)  
##  
## Residuals:  
##      Min       1Q   Median       3Q      Max  
## -45.963 -12.873   2.109  10.679  44.054  
##  
## Coefficients:  
##             Estimate Std. Error t value Pr(>|t|)  
## (Intercept) 484.882137 264.914499  1.830  0.0762 .  
## x1          -35.153528 13.983161  -2.514  0.0170 *  
## x2           0.009766  0.124502   0.078  0.9380  
## x3          11.750673  5.979973   1.965  0.0579 .  
## x4           0.278738  0.500545   0.557  0.5814  
## x5           2.669543  0.882795   3.024  0.0048 **  
## x6           3.637652  2.306533   1.577  0.1243  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##  
## Residual standard error: 20.28 on 33 degrees of freedom  
## Multiple R-squared:  0.8975, Adjusted R-squared:  0.8788  
## F-statistic: 48.15 on 6 and 33 DF, p-value: 6.3e-15
```

real value prediction for test dataset
#MPSEfull
mean((Ytest-predict(fitall,Xtest))^2) → MSPE

```
## [1] 1558.079
```

Example: SAT Scores (Con'd)

```
fitselect=lm(Ytraining~.,data=Xtraining[,c(1,3,5,6)])
summary(fitselect)

##
## Call:
## lm(formula = Ytraining ~ ., data = Xtraining[, c(1, 3, 5, 6)])
##
## Residuals:
##    Min      1Q  Median      3Q     Max 
## -48.60 -12.90   2.60  10.56  43.06 
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) 438.7462   232.7437   1.885  0.06774 .
## x1          -31.3364    11.2500  -2.785  0.00857 ** 
## x3           11.4040    5.7976   1.967  0.05715 .  
## x5            2.8932    0.7569   3.822  0.00052 *** 
## x6            4.4030    1.8371   2.397  0.02202 *  
## ---        
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 19.79 on 35 degrees of freedom
## Multiple R-squared:  0.8964, Adjusted R-squared:  0.8846 
## F-statistic: 75.72 on 4 and 35 DF,  p-value: < 2.2e-16
```

#MPSEselect
mean((Ytest-predict(fitselect,Xtest[,c(1,3,5,6)]))^2)

[1] 1417.981

model based on selection

Multicollinearity – Motivating Example

In a study of primary school reading ability, an educator wished to relate a measure of a student's reading ability (Y) to age (X_1) and gender (female/male). 

Gender is a categorical variable with two categories.

To allow for it in MLR, we can consider the following two possible indicator variables.

$$\begin{cases} X_2 = 1 \text{ if } \underline{\text{female}}; 0 \text{ otherwise;} \\ X_3 = 1 \text{ if } \underline{\text{male}}; 0 \text{ otherwise.} \end{cases}$$

It is worth noting that $\underline{X_2 + X_3 = 1}$.

Actually we only have the information of X_1 and X_2 . The information of X_3 is redundant since $X_3 = 1 - X_2$, which is included in X_2 . This phenomenon is called **multicollinearity**.

As a consequence, drop one of the indicator variables (X_2 or X_3) and fit the model.

Multicollinearity

In general, suppose we have explanatory variables X_1, \dots, X_k in total, if one of the explanatory variables X_j have the situation

$$\underline{X_j} \approx c_0 + c_1 \underline{X_1} + \dots + c_{j-1} \underline{X_{j-1}} + c_{j+1} \underline{X_{j+1}} + \dots + c_k \underline{X_k},$$

for some constants $c_0, c_1, \dots, c_{j-1}, c_{j+1}, \dots, c_k$ (which are not all equal to 0), then those explanatory variables are said to be multicollinear.

The MLR based on explanatory variables X_1, \dots, X_k is said to have a multicollinearity problem.

Multicollinearity can result in

$$\begin{pmatrix} 1 & X_{1,1} & \cdots & X_{1,n} \\ 1 & X_{2,1} & \cdots & X_{2,n} \\ \vdots & \vdots & & \vdots \\ 1 & X_{k,1} & \cdots & X_{k,n} \end{pmatrix}$$



- For the $n \times (k + 1)$ design matrix $\underline{\mathbb{X}}$, the matrix inverse $(\underline{\mathbb{X}}^\top \underline{\mathbb{X}})^{-1}$ may not exist (cannot be computed), and thus the LS estimates may not be obtained.

$$(\underline{\mathbb{X}}^\top \underline{\mathbb{X}})^{-1} \underline{\mathbb{X}}^\top \underline{y}$$

- Even if sometimes $(\underline{\mathbb{X}}^\top \underline{\mathbb{X}})^{-1}$ still exists, the LS estimates are highly unstable and imprecise \Rightarrow SEs of the estimators are large and a lot of hypothesis testing results are not significant.

Variance Inflation Factors (VIF)

A quantitative measure of the multicollinearity is given by the variance inflation factors (VIF).

The VIF associated with the j -th explanatory variable is:

$$\text{VIF}_j = \frac{1}{1 - R_j^2}$$

is not from MLR
 $\mu Y | X_j$

where R_j^2 is the R-squared by regressing X_j on $X_1, \dots, X_{j-1}, X_{j+1}, \dots, X_k$.

The “rule of thumb” cut-off for VIF is 10.

If one of VIF_j 's is larger than 10, then the MLR based on explanatory variables X_1, \dots, X_k has a multicollinearity problem.

In this case, all the parameter estimates associated with explanatory variables with large VIF's will have low precision (very large standard errors \Rightarrow insignificant).

How to deal with the multicollinearity problem?

Backward Elimination to Deal with Multicollinearity

①

An explanatory variable X_j with $\text{VIF}_j > 10$ should be eliminated.

However, usually two or more VIFs are larger than 10.

Can we eliminate all of them in the model? NO!

We can only drop one at each time. This is very similar to the backward elimination for variable selection.

In this case, which one of the explanatory variables should be eliminated?

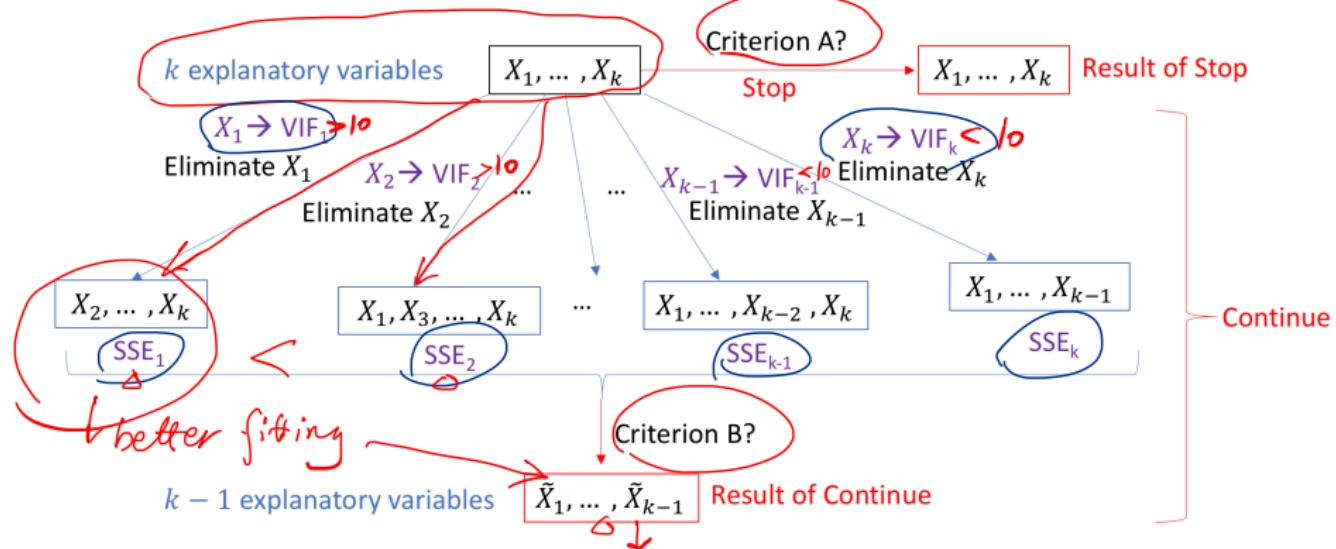
③

Intuitively, the criterion is: the resulting model after dropping the explanatory variable **has the best fitting (smallest SSE or deviance)**.

Later in this course, we will introduce an alternative approach, principal components analysis (PCA), to deal with multicollinearity.

Backward Elimination Step

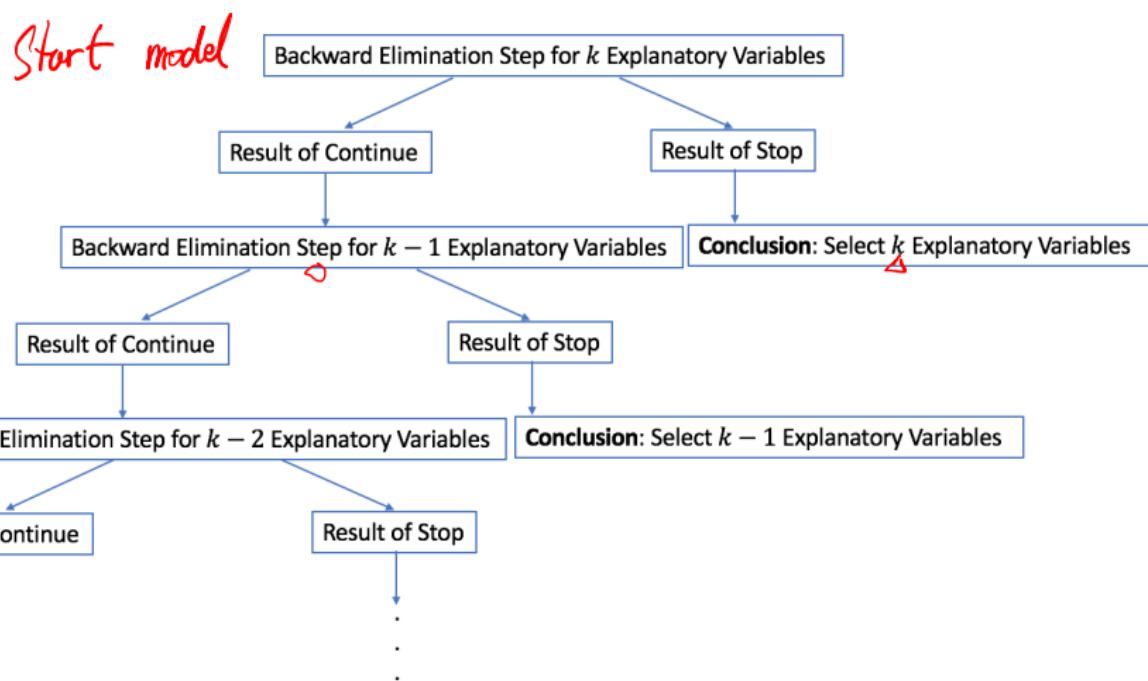
Backward Elimination Step for k Explanatory Variables



Criterion A: if $\max\{\text{VIF}_1, \text{VIF}_2, \dots, \text{VIF}_{k-1}, \text{VIF}_k\} < 10$ or $k = 0$, then Stop; otherwise Continue.

Criterion B: $\tilde{X}_1, \dots, \tilde{X}_{k-1}$ are those variables such that under the condition $\text{VIF} > 10$, the corresponding SSE is the smallest.

Complete Backward Elimination Procedure to Deal with Multicollinearity



Keep doing on the above procedures, until the first time we obtain the **Result of Stop**.

Example: SAT Scores (Con'd)

```
Y<-SATdata[,2]
X<-SATdata[,-c(1,2)]
fit=lm(Y~.,data=X)
summary(fit)

##
## Call:
## lm(formula = Y ~ ., data = X)
##
## Residuals:
##    Min      1Q  Median      3Q     Max 
## -51.388 -13.268   1.758  13.496  51.024 
##
## Coefficients:
##             Estimate Std. Error t value Pr(>|t|)    
## (Intercept) -203.92618 192.87664 -1.057  0.29642  
## Takers       0.01833  0.64099  0.029  0.97732  
## Income        0.18058  0.14807  1.220  0.22944  
## Years         16.53592  5.95782  2.775  0.00819 ** 
## Public       -0.44299  0.52053 -0.851  0.39957  
## Expend        3.72998  0.88504  4.214  0.00013 *** 
## Rank          9.78937  1.93456  5.060 8.75e-06 *** 
## ---      
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 
##
## Residual standard error: 23.68 on 42 degrees of freedom
## Multiple R-squared:  0.904, Adjusted R-squared:  0.8903 
## F-statistic: 65.95 on 6 and 42 DF,  p-value: < 2.2e-16
```



Example: SAT Scores (Con'd)

```
#install.packages('car')  
library(car)
```

Warning: package 'car' was built under R version 3.3.2

\leftarrow lm() result

```
vif(fit)
```

	Takers	Income	Years	Public	Expend	Rank
##	17.399186	3.201245	1.469017	2.173352	1.535438	13.917392

#Continue

#Try Dropping Takers

```
X1=X[,-1]
```

```
fit1=lm(Y~.,data=X1)
```

```
deviance(fit1) → SSE1
```

```
## [1] 23546.74
```

#Try dropping Rank

```
X2=X[,-6]
```

```
fit2=lm(Y~.,data=X2)
```

```
deviance(fit2) → SSE2
```

```
## [1] 37901.82
```

#Based on SSE, we choose to eliminate TAKERS. The resulting VIFs after elimination are:

vif(fit1) → model with variables except for Takers

	Income	Years	Public	Expend	Rank
##	2.328161	1.446115	2.061058	1.531747	2.311335

#Stop

backward step for 6 variables

(A) ↓ Continue

(B) Drop Takers

↓

backward step for 5 variables

(A) ↓ Stop

VIFs > 10