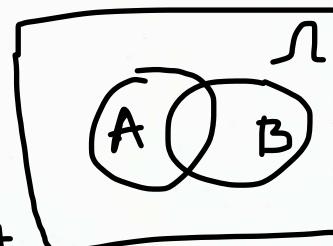


Lecture 2.

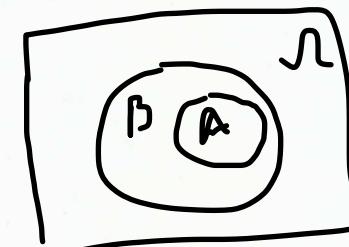
2.1

Rules of Probability

1. $P(\bar{A}) = 1 - P(A), A \in \mathcal{F}$
 $P(\bar{A} \cup A) = P(\Omega) = 1, A \cap \bar{A} = \emptyset$
 $P(\bar{A}) + P(A) = 1 \Rightarrow P(\bar{A}) = 1 - P(A)$ ■
2. $P(\emptyset) = 0$
 proved earlier
3. $P(A \cup B) = P(A) + P(B) - P(A \cap B)$ (Addition law)
 $\Omega = (\bar{A} \cup B) \cup (A \cap B^c) \cup (A \cap B) \cup (B \cap A^c)$
 $\bar{A} \cup B, A \cap B^c, A \cap B, B \cap A^c$ are all disjoint


$$P(A \cup B) = P(\underbrace{\bar{A} \cup B}_{P(A)} + \underbrace{A \cap B^c}_{P(B^c)} + \underbrace{A \cap B}_{P(A \cap B)} + \underbrace{B \cap A^c}_{P(B \cap A^c)})$$

$$P(A \cup B) = P(A) + P(B) - P(A \cap B)$$

■
4. If $A \subseteq B$, then $P(A) \leq P(B)$
 $B = A \cup (A^c \cap B)$
 $A^c \cap B$ disjoint

 $P(B) = P(A) + P(A^c \cap B) \geq_0 \Rightarrow P(A) \leq P(B)$ ■

$$5. P(\bigcup_{i=1}^n A_i) = \sum_{i=1}^n P(A_i) - \sum_{\substack{i > j \\ i > j > k}} P(A_i \cap A_j \cap A_k) + \sum_{\substack{i > j > k \\ i > j > k > l}} P(A_i \cap A_j \cap A_k \cap A_l) - \dots$$

... $+ (-1)^{n-k} P(A_1 \cap A_2 \cap \dots \cap A_n)$

$$P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(A \cap B) - P(B \cap C) - P(A \cap C) + P(A \cap B \cap C)$$

$$6. P(\bigcup_{i=1}^n A_i) \leq \sum_{i=1}^n P(A_i)$$

Proof: By math induction

$$\textcircled{1} \quad n=2, P(A_1 \cup A_2) = P(A_1) + P(A_2) - P(A_1 \cap A_2) \leq P(A_1) + P(A_2)$$

\textcircled{2} Assume for $n=k$, i.e.

$$P(\bigcup_{i=1}^k A_i) \leq \sum_{i=1}^k P(A_i)$$

$$\textcircled{3} \quad \begin{aligned} \text{Prove for } n=k+1 \\ P(\bigcup_{i=1}^{k+1} A_i) &= P\left[\left(\bigcup_{i=1}^k A_i\right) \cup A_{k+1}\right] \stackrel{\textcircled{2}}{\leq} P\left(\bigcup_{i=1}^k A_i\right) + P(A_{k+1}) \\ &\leq \sum_{i=1}^k P(A_i) + P(A_{k+1}) = \sum_{i=1}^{k+1} P(A_i) \end{aligned}$$

Ex. In a lottery there are 10 tickets numbered 1, 2, ..., 10. Two numbers are drawn for prizes. You hold tickets 1 and 2. $P(\text{you win at least one prize}) = ?$

Sol'n: $A: \{\#1 \text{ is winning}\}$
 $B: \{\#2 \text{ is winning}\}$

$$\frac{\binom{2}{2}}{\binom{10}{2}} = \frac{1}{45}$$

$$P(A \cup B) = P(A) + P(B) - P(A \cap B) = \frac{1}{5} + \frac{1}{5} - \frac{1}{5} \cdot \frac{1}{9}$$

$$= \frac{2}{5} - \frac{1}{45} = \frac{17}{45}$$

Conditional Probability

2.3

Suppose we performed a chance experiment, and we don't know the outcome (ω), but we have some partial information (event A) about ω .

Question: Given this partial information what's the probability that the outcome is in some event B?

Ex. Toss a coin 3 times.

$$B = \{ \text{2 or more heads} \} = \{ \text{HHH, HHT, HTH, THH} \}$$

$$A = \{ \text{1st outcome is H} \} = \{ \text{HHH, HHT, HTH, HTT} \}$$

$$P(B|A) = \frac{3}{4} = \frac{3/8}{4/8} = \frac{P(A \cap B)}{P(A)}$$

$$A \cap B = \{ \text{HHH, HHT, HTH} \}$$

Def. Given a probability space (Ω, \mathcal{F}, P) and events $A, B \in \mathcal{F}$, $P(A) > 0$. The **conditional probability** of B given the information that A has occurred is

$$P(B|A) = \frac{P(A \cap B)}{P(A)}$$

Ex. we toss a die. $P(\text{observing 6} | \text{outcome is even})$ [2-4]

$$A = \{\text{even}\}, B = \{6\}$$

$$P(B|A) = \frac{P(A \cap B)}{P(A)} = \frac{\frac{1}{6}}{\frac{1}{2}} = \frac{1}{3}$$

Theorem. If $A \in \mathcal{F}$, $P(A) > 0$, then (Ω, \mathcal{F}, Q) is a probability space, where $Q : \mathcal{F} \rightarrow [0, 1]$ is defined by $Q(B) = P(B|A)$.

$$P(\bar{B}|A) = 1 - P(B|A), A, B \in \mathcal{F}, P(A) > 0$$

$$\begin{aligned} P(B_1 \cup B_2 | A) &= P(B_1 | A) + P(B_2 | A) - \\ &\quad - P(B_1 \cap B_2 | A) \end{aligned}$$

Multiplication rule

- For any two events A and B , $P(A \cap B) = P(B|A)P(A)$
- For any 3 events A, B , and C , $P(A \cap B \cap C) = P(A)P(B|A)P(C|AB)$
- In general,

$$P\left(\bigcap_{i=1}^n A_i\right) = P(A_1)P(A_2|A_1)P(A_3|A_1 \cap A_2) \cdots P\left(A_n \mid \bigcap_{i=1}^{n-1} A_i\right)$$

Ex. An urn contains 10 balls, 3 blue and 7 white. we draw a ball and note its color. Then we replace it and add one more of the same color. we repeat this process 3 times. What is the probability that the first 2 balls are blue and the third one is white?

2.5

Sol'n:

$$A_1 = \{1^{\text{st}} \text{ ball is blue}\}$$

$$A_2 = \{2^{\text{nd}} \text{ ball is blue}\}$$

$$A_3 = \{3^{\text{rd}} \text{ is white}\}$$

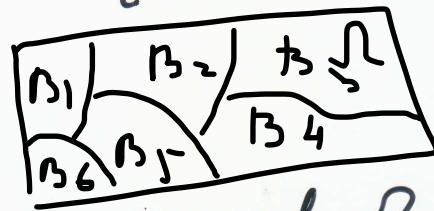
$$\begin{aligned} P(A_1 \cap A_2 \cap A_3) &= P(A_1)P(A_2|A_1)P(A_3|A_1 \cap A_2) \\ &= \frac{3}{10} \cdot \frac{4}{11} \cdot \frac{7}{12} = \frac{7}{110} \end{aligned}$$

Law of Total Probability.

Def. For a probability space (Ω, \mathcal{F}, P) , a **partition** of Ω is a countable collection $\{B_i\}$ of events such that

$$B_i \in \mathcal{F}, B_i \cap B_j = \emptyset \text{ for } i \neq j \quad \text{and} \quad \bigcup_i B_i = \Omega$$

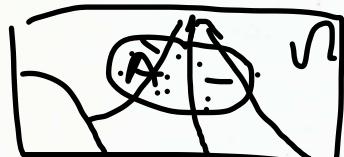
$$\Omega = \bigcup_{i=1}^6 B_i$$



Theorem. If $\{B_1, B_2, \dots, B_k\}$ is a partition of Ω and $A \in \mathcal{F}$, then for all $i = 1, \dots, k$, such that $P(B_i) > 0$,

$$P(A) = \sum_{i=1}^k P(A|B_i) P(B_i)$$

$$\forall A \in \mathcal{F}$$



Proof: $A \in \mathcal{F}, A = A \cap \Omega$

$$\Omega = \bigcup_{i=1}^k B_i, A = A \cap (\bigcup B_i)$$

$$= \underbrace{(A \cap B_1)}_{\text{disjoint}} \cup \underbrace{(A \cap B_2)}_{\text{disjoint}} \cup \dots \cup \underbrace{(A \cap B_k)}$$

$$P(A) = P(A \cap B_1) + \dots + P(A \cap B_k)$$

$$= P(A|B_1)P(B_1) + \dots + P(A|B_k)P(B_k) = \sum_{i=1}^k P(A|B_i)P(B_i)$$

Bayes' Rule

Let $\{B_1, B_2, \dots, B_k\}$ be a partition of Ω such that $P(B_i) > 0, i=1, k$. Then

$$P(B_j|A) = \frac{P(A|B_j)P(B_j)}{\sum_{i=1}^k P(A|B_i)P(B_i)}, A \in \mathcal{F}$$

Proof:

$$P(B_j|A) = \frac{P(A \cap B_j)}{P(A)} = \frac{P(A|B_j)P(B_j)}{\sum_{i=1}^k P(A|B_i)P(B_i)}$$

■

Ex. There are 2 lists of applicants for jobs.

List 1 contains the names of 5 women and 2 men.

List 2 contains the names of 2 women and 6 men.

A name is randomly selected from list 1 and added to list 2. A name is then randomly selected from list 2. Given that the name selected is that of a man, what is the probability that a woman's name was originally selected from list 1?

Sol'n: $A = \text{man's name selected from list 2}$
 $B = \text{woman's name ...}$ list 1

$$P(B|A) = ? \quad \Omega = B \cup \bar{B}, B_1 = B, B_2 = \bar{B}$$

$$P(B|A) = \frac{P(A|B)P(B)}{P(A|B)P(B) + P(A|\bar{B})P(\bar{B})}$$

$$= \frac{2}{3} \cdot \frac{5}{7} = \frac{\left(\frac{2}{3} \cdot \frac{5}{7} + \frac{7}{9} \cdot \frac{2}{7} \right)}{22} = \frac{15}{22}$$

Independence

Ex. Roll a die twice.

A : 3 or less on first roll = $\{(1, \cdot), (2, \cdot), (1, \cdot)\}$
 B : sum is odd = $\{(1, 2), (1, 4), (1, 6), \dots\}$

$$P(A) = \frac{18}{36} = \frac{1}{2} = P(B), A \cap B = \{(1, 2), (3, 4), (3, 6), \dots\}$$

$$P(A \cap B) = \frac{1}{4}, P(A) = P(B)$$

$$P(B|A) = \frac{P(B \cap A)}{P(A)} = \frac{1}{2} = \frac{P(A \cap B)}{P(A)} = P(A|B)$$

- If occurrence of one event does not affect the probability that the other occurs, then these events are independent.

Def. Events A and B are **independent** if

$$P(A \cap B) = P(A)P(B) \Leftrightarrow P(A|B) = P(A)$$

$$\qquad \qquad \qquad P(B|A) = P(B)$$

Note: Independence \neq disjoint !!!

Ex. Toss a die

A : observe an odd number

$$P(A) = P(B) = \frac{1}{2}$$

B : observe an even number

$$A \cap B = \emptyset$$

C : observe a 1 or 2

$A, B - A$ is disjoint

$$P(A \cap B) = 0 \neq P(A)P(B) = \frac{1}{4}$$

\Rightarrow not independent

$$P(C) = \frac{1}{3}, A \cap C = \{(1)\}, P(A \cap C) = \frac{1}{6}$$

$$= P(A) \cdot P(C) = \frac{1}{2} \cdot \frac{1}{3} = \frac{1}{6}$$

In general,

A collection of events $\{A_1, A_2, \dots, A_n\}$ is (mutually) independent if for any subcollection $\{A_{i_1}, A_{i_2}, \dots, A_{i_m}\}$

$$P(A_{i_1} \cap A_{i_2} \cap \dots \cap A_{i_m}) = P(A_{i_1}) P(A_{i_2}) \cdots P(A_{i_m})$$

Note: Pairwise independence does not guarantee mutual independence.

Ex. Roll a die twice

A : 1st die odd

$$P(A) = P(B) = P(c) = \frac{1}{2}$$

B : 2nd die odd

C : sum is odd

$$P(A \cap B) = \frac{1}{4} = P(A) P(B)$$

$$P(A \cap C) = \frac{1}{4} = P(A) P(C)$$

$$P(B \cap C) = \frac{1}{4} = P(B) P(C)$$

$\Rightarrow A, B, C$ are pairwise independent

$$P(A \cap B \cap C) = P(\emptyset) = 0 \neq P(A) P(B) P(C)$$

Claim: If A, B are independent so are A, \bar{B} and \bar{A}, B , and \bar{A}, \bar{B}

L 2.9

Proof:

Indicator function: $I_A(x) = \begin{cases} 1, & x \in A \\ 0, & x \notin A \end{cases}$

Fact: $A \subset B \Leftrightarrow I_A(x) \leq I_B(x)$
iff

Proof: $\Rightarrow A \subset B$

let $x \in A \Rightarrow x \in B$

$$I_A(x) = 1 \leq I_B(x) = 1$$

$$x \notin A \Rightarrow I_A(x) = 0$$

$$I_B(x) \geq 0 = I_A(x)$$

$$\Rightarrow I_A(x) \leq I_B(x) \quad \forall x$$

$$\Leftarrow I_A(x) \leq I_B(x)$$

let $x \in A$, need: $x \in B$

$$\text{Suppose } x \notin B \Rightarrow I_B(x) = 0$$

$$0 = I_B(x) \geq I_A(x) = 1 \rightarrow \leftarrow \Rightarrow A \subset B$$