

Convolution.

Suppose X, Y jointly distributed r.v's. We want to find the probability/density function of $Z = X + Y$.

Discrete case:

Let X, Y have joint probability function $P_{X,Y}(x,y)$.

$Z = z$ whenever $X = x$ and $Y = z - x$

$$P_Z(z) = P(Z=z) = \sum_x P_{X,Y}(x, z-x)$$

If X, Y are independent,

$$P_Z(z) = \sum_x P_X(x) P_Y(z-x) \leftarrow \begin{matrix} P_Z = P_X \circ P_Y \\ \text{convolution} \end{matrix}$$

Ex. $X \sim \text{Poisson}(\lambda_1)$, $Y \sim \text{Poisson}(\lambda_2)$, X and Y are independent. Find the distribution of $X+Y$.

Sol'n: $Z = X+Y$

$$(a+b)^n = \sum_{k=1}^n \binom{n}{k} a^k b^{n-k}$$

$$P_Z(z) = \sum_x P_X(x) P_Y(z-x)$$

$$= \sum_x \frac{\lambda_1^x e^{-\lambda_1}}{x!} \frac{\lambda_2^{z-x} e^{-\lambda_2}}{(z-x)!}$$

$$= e^{-\lambda_1-\lambda_2} \sum_x \frac{\lambda_1^x \lambda_2^{z-x}}{x! (z-x)!} \frac{z!}{z!}$$

$$= \frac{e^{-\lambda_1-\lambda_2}}{z!} \sum_x \binom{z}{x} \lambda_1^x \lambda_2^{z-x}$$

$$= \frac{e^{-\lambda_1-\lambda_2}}{z!} (\lambda_1 + \lambda_2)^z \sim \text{Poisson}(\lambda_1 + \lambda_2)$$

Continuous case:

Suppose X, Y are r.v.'s with joint density function $f_{X,Y}(x, y)$. We want to find the density functions of $Z = X + Y$.

$$\begin{aligned}
 F_Z(z) &= P(Z \leq z) = P(X + Y \leq z) \\
 &= \int_{-\infty}^{\infty} \int_{-\infty}^{z-x} f_{X,Y}(x, y) dy dx \\
 &= \left(\begin{array}{l} y = u - x \\ dy = du \end{array} \right) = \int_{-\infty}^{\infty} \int_{-\infty}^z f_{X,Y}(x, u-x) du dx \\
 &= \int_{-\infty}^z \left[\int_{-\infty}^x f_{X,Y}(x, u-x) dx \right] du
 \end{aligned}$$

$$f_Z(z) = \int_{-\infty}^z f_{X,Y}(x, z-x) dx$$

If $\int_{-\infty}^z f_{X,Y}(x, z-x) dx$ is continuous at z

If X, Y are independent,

$$\begin{aligned}
 f_Z(z) &= \int_{-\infty}^{\infty} f_X(x) f_Y(z-x) dx \\
 &= f_X \circ f_Y \rightarrow \text{convolution}
 \end{aligned}$$

Ex. $X, Y \sim \text{iid } \exp(\lambda)$. Find the density of

$$Z = X+Y.$$

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}, \quad f_Y(y) = \begin{cases} \lambda e^{-\lambda y}, & y \geq 0 \\ 0, & \text{otherwise} \end{cases}$$

$$\begin{aligned} f_Z(z) &= \int_0^z f_X(x) f_Y(z-x) dx \\ &= \int_0^z \lambda e^{-\lambda x} \cdot \lambda e^{-\lambda(z-x)} dx = \lambda^2 e^{-\lambda z} \int_0^z dx \\ &= \begin{cases} \lambda^2 e^{-\lambda z}, & z \geq 0 \\ 0, & z < 0 \end{cases} \Rightarrow Z \sim \text{Gamma}(2, \lambda) \end{aligned}$$

More on Normal Distribution.

If $Z \sim N(0, 1)$, then

$$\varphi_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}, \quad -\infty < z < \infty$$

$$X = \sigma Z + \mu \Rightarrow X \sim N(\mu, \sigma^2)$$

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, \quad -\infty < x < \infty$$

$$Z = \frac{X-\mu}{\sigma} \sim N(0, 1)$$

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Let $X, Y \sim \text{iid } N(0, 1)$, $Z = X + Y$

$$f_Z(z) = f_X \circ f_Y(z) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-x^2/2} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{(z-x)^2}{2}} dx \quad (\textcircled{E})$$

$$-\frac{x^2}{2} - \frac{(z-x)^2}{2} = -\frac{x^2}{2} - \frac{z^2}{2} + zx - \frac{x^2}{2}$$

$$= -\frac{z^2}{4} - \frac{x^2}{4} + zx - x^2 = -\frac{z^2}{4} - \left(x - \frac{z}{2}\right)^2$$

$$\textcircled{E} \quad \frac{1}{2\pi} e^{-\frac{z^2}{4}} \int_{-\infty}^{\infty} e^{-(x - \frac{z}{2})^2} dx \cdot \frac{\sqrt{\pi}}{\sqrt{\pi}}$$

$$N\left(\frac{z}{2}, \frac{1}{2}\right)$$

$$= \frac{1}{2\sqrt{\pi}} e^{-\frac{z^2}{4}} \underbrace{\int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi \cdot \frac{1}{2}}} e^{-\frac{(x - \frac{z}{2})^2}{2 \cdot \frac{1}{2}}} dx}_{\approx 1}$$

$$= \frac{1}{\sqrt{4\pi}} e^{-\frac{z^2}{4}}$$

$$Z \sim N(0, 2)$$

In general,

$$X_1, \dots, X_n \sim \text{iid } N(0, 1)$$

$$X_1 + \dots + X_n \sim N(0, n)$$

Ex. $X_1, X_2 \sim \chi^2_{(1)}$, $Y = X_1 + X_2$. Find density of Y . [9.5]

$$f_{X_1}(x) = \frac{x^{-1/2} e^{-x/2}}{\Gamma(1/2)}$$

$$f_Y(y) = \int_0^y \frac{x^{-1/2} e^{-x/2}}{\sqrt{2} \sqrt{\pi}} \frac{(y-x)^{-1/2} e^{-(y-x)/2}}{\sqrt{2} \sqrt{\pi}} dx$$

$$= \frac{1}{2\pi} e^{-y/2} \int_0^y x^{-1/2} (y-x)^{-1/2} dx$$

$$u = \frac{x}{y} \Rightarrow y du = dx$$

$$= \frac{1}{2\pi} e^{-y/2} \int_0^1 (uy)^{-1/2} (y-uy)^{-1/2} y du$$

$$= \frac{1}{2\pi} e^{-y/2} \int_0^1 u^{-1/2} (1-u)^{-1/2} du \quad (=)$$

$$\text{Beta}(\alpha, \beta) = \int_0^1 u^{\beta-1} (1-u)^{\alpha-1} du = \frac{\Gamma(\alpha) \Gamma(\beta)}{\Gamma(\alpha+\beta)}$$

$$(=) \frac{1}{2\pi} e^{-y/2} \frac{\Gamma(\frac{1}{2}) \Gamma(\frac{1}{2})}{\Gamma(\frac{1}{2} + \frac{1}{2})} = \frac{1}{2\pi} e^{-y/2} \frac{\sqrt{\pi} \sqrt{\pi}}{1}$$

$$= \frac{1}{2} e^{-y/2} \Rightarrow Y \sim \text{Gamma}(1, \frac{1}{2})$$

Cauchy Distribution: The standard Cauchy distribution can be expressed as the ratio of two standard Normal r.v's.

Let $X, Y \sim \text{iid } N(0,1)$. Then $Z = \frac{X}{Y} \sim \text{Cauchy}$.

What's the density function of Z ?

Multivariable Transformations.

Recall: If X is a r.v. with $f_X(\cdot)$ and $U = h(X)$, where h is either \nearrow or \searrow

$$\text{Then } f_U(u) = f_X(h^{-1}(u)) \left| \frac{d h^{-1}(u)}{du} \right|$$

Suppose X and Y are jointly continuous r.v.'s, and $U_1 = X+Y$, $U_2 = X-Y$
 $f_{U_1, U_2} = ?$

Theorem. Suppose X, Y are continuous r.v.'s with joint df $f_{X,Y}(x,y)$ and that for all (x,y) s.t. $f_{X,Y}(x,y) > 0$

$$u_1 = h_1(x,y) \text{ and } u_2 = h_2(x,y)$$

is a 1-1 transformation from (x,y) to (u_1, u_2) with inverse

$$x = h_1^{-1}(u_1, u_2) \text{ and } y = h_2^{-1}(u_1, u_2).$$

$$\text{Then } f_{U_1, U_2}(u_1, u_2) = f_{X,Y}\left(h_1^{-1}(u_1, u_2), h_2^{-1}(u_1, u_2)\right) |J|$$

where $J = \det \begin{bmatrix} \frac{\partial h_1}{\partial u_1} & \frac{\partial h_1}{\partial u_2} \\ \frac{\partial h_2}{\partial u_1} & \frac{\partial h_2}{\partial u_2} \end{bmatrix} \rightarrow \text{Jacobian}$

Ex. $X, Y \sim \text{iid } N(0, 1)$

$$U_1 = X+Y, U_2 = X-Y$$

$$f_{U_1, U_2} = ?$$

$$f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}, f_Y(y) = \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}, f_{X,Y} = \frac{1}{2\pi} e^{-\frac{x^2+y^2}{2}}$$

$$U_1 = X+Y = h_1(X, Y), U_2 = X-Y = h_2(X, Y)$$

$$x = \frac{U_1 + U_2}{2} = h_1^{-1}(U_1, U_2), y = \frac{U_1 - U_2}{2} = h_2^{-1}(U_1, U_2)$$

$$J = \det \begin{bmatrix} \frac{\partial h_1^{-1}}{\partial U_1} & \frac{\partial h_1^{-1}}{\partial U_2} \\ \frac{\partial h_2^{-1}}{\partial U_1} & \frac{\partial h_2^{-1}}{\partial U_2} \end{bmatrix} = \det \begin{bmatrix} 1/2 & 1/2 \\ 1/2 & -1/2 \end{bmatrix} = \frac{1}{2} \left(-\frac{1}{2}\right) - \frac{1}{2} \cdot \frac{1}{2} = -\frac{1}{2}$$

$$\begin{aligned} f_{U_1, U_2}(U_1, U_2) &= f_{X, Y}(h_1^{-1}, h_2^{-1}) \left| \begin{bmatrix} U_1 \\ U_2 \end{bmatrix} \right| \\ &= \frac{1}{2\pi} e^{-\frac{1}{2} \left(\frac{U_1+U_2}{2}\right)^2} e^{-\frac{1}{2} \left(\frac{U_1-U_2}{2}\right)^2} \left| -\frac{1}{2} \right| \\ &= \frac{1}{4\pi} e^{-\frac{U_1^2}{4} - \frac{U_2^2}{4}} = \frac{e^{-\frac{U_1^2}{4}}}{\underbrace{\sqrt{4\pi}}_{f_{U_1}(U_1)}} \cdot \frac{e^{-\frac{U_2^2}{4}}}{\underbrace{\sqrt{4\pi}}_{f_{U_2}(U_2)}} \\ U_1, U_2 &\sim N(0, 2) \end{aligned}$$

Density of Quotient.

Suppose X, Y are independent continuous r.v's and we are interested in the density of

$$Z = \frac{X}{Y}.$$

$$U_1 = \frac{X}{Y}, \quad U_2 = Y$$

$$h_1(x, y) = \frac{x}{y}, \quad h_2(x, y) = y$$

$$x = u_1, u_2 = h_1^{-1}(u_1, u_2), \quad y = u_2 = h_2^{-1}(u_1, u_2)$$

$$J = \det \begin{bmatrix} u_2 & u_1 \\ 0 & 1 \end{bmatrix} = u_2 \cdot 1 - 0 \cdot u_1 = u_2$$

$$f_{U_1, U_2}(u_1, u_2) = f_{X, Y} \left(h_1^{-1}(u_1, u_2), h_2^{-1}(u_1, u_2) \right) |J|$$

$$= f_{X, Y}(u_1, u_2, u_2) |u_2|$$

$$= f_X(u_1, u_2) f_Y(u_2) |u_2|$$

$$\boxed{f_{U_1}(u_1) = \int_{-\infty}^{\infty} f_X(u_1, u_2) f_Y(u_2) |u_2| du_2}$$

$$X, Y \sim \text{iid } \mathcal{N}(0, 1)$$

$$Z = \frac{X}{Y}$$

$$U_1 = \frac{X}{Y}, \quad U_2 = Y$$

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$$f_{U_1, U_2}(u_1, u_2) = f_X(u_1, u_2) f_Y(u_2) |u_2|$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{(u_1 u_2)^2}{2}} \cdot \frac{1}{\sqrt{2\pi}} e^{-\frac{u_2^2}{2}} |u_2|$$

$$f_{U_1}(u_1) = \int_{-\infty}^{\infty} \frac{1}{2\pi} e^{-\frac{u_2^2}{2}} (u_1^2 + 1) |u_2| du_2$$

even

$$= 2 \int_0^{\infty} \frac{1}{2\pi} e^{-\frac{u_2^2}{2}} (u_1^2 + 1) u_2 du_2$$

$$w = \frac{u_2^2}{2} (u_1^2 + 1)$$

$$\frac{1}{u_1^2 + 1} du_2 = u_2 du_2$$

$$= \frac{1}{\pi} \int_0^{\infty} \frac{1}{u_1^2 + 1} e^{-w} dw = \frac{1}{\pi} \frac{1}{u_1^2 + 1} \left[\int_0^{\infty} e^{-w} dw \right]$$

$$Z = \frac{X}{Y} \Rightarrow f_Z(z) = \frac{1}{\pi} \frac{1}{z^2 + 1} \xleftarrow{\text{Cauchy d.f}}$$

t-distribution.

Beta Distribution

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