

Tutorial 3 Solutions

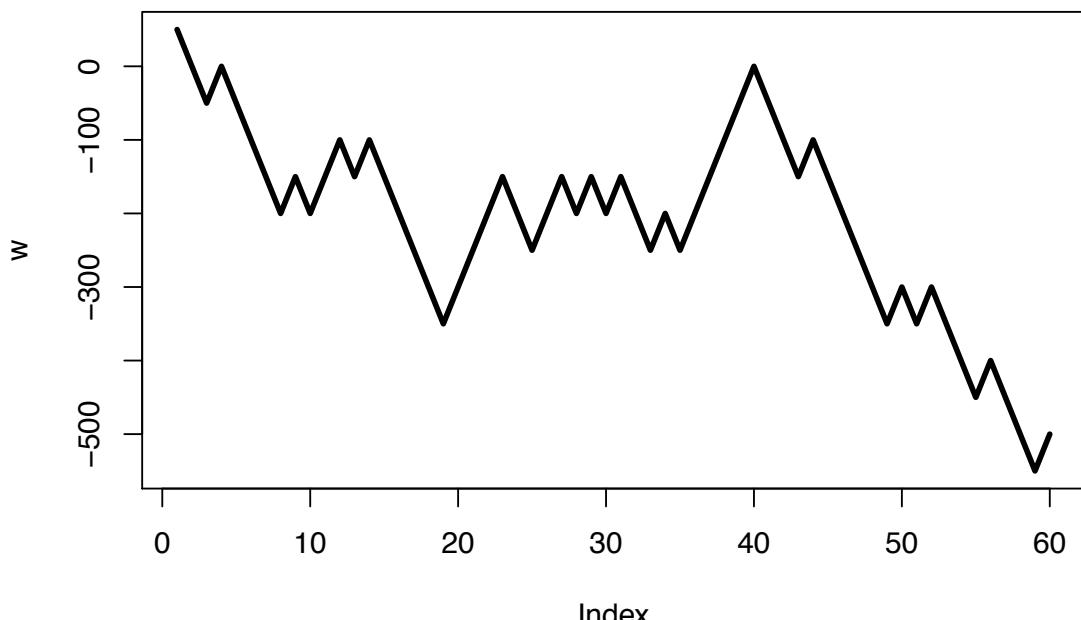
STAT 3013/4027/8027

-
1. SI: 2.6, 2.8, 2.10 **Ans.** See the handwritten pages.

- **Question 2: R code:**

```
##  
set.seed(10)  
x <- rbinom(60, 1, 0.5)  
x[x==0] <- -1  
x <- x*50  
  
w <- cumsum(x)  
  
##  
plot(w, type="l", lwd=3, main="A Realization of the Drunkard's Walk")
```

A Realization of the Drunkard's Walk



```
## Let's examine the sampling distribution of W  
set.seed(10)  
S <- 10000  
W <- rep(0, S)  
  
for(s in 1:S){
```

```

x <- rbinom(60, 1, 0.5)
x[x==0] <- -1
x <- x*50
W[s] <- sum(x)
}

##  

mean(W)

## [1] 1.13  

var(W)

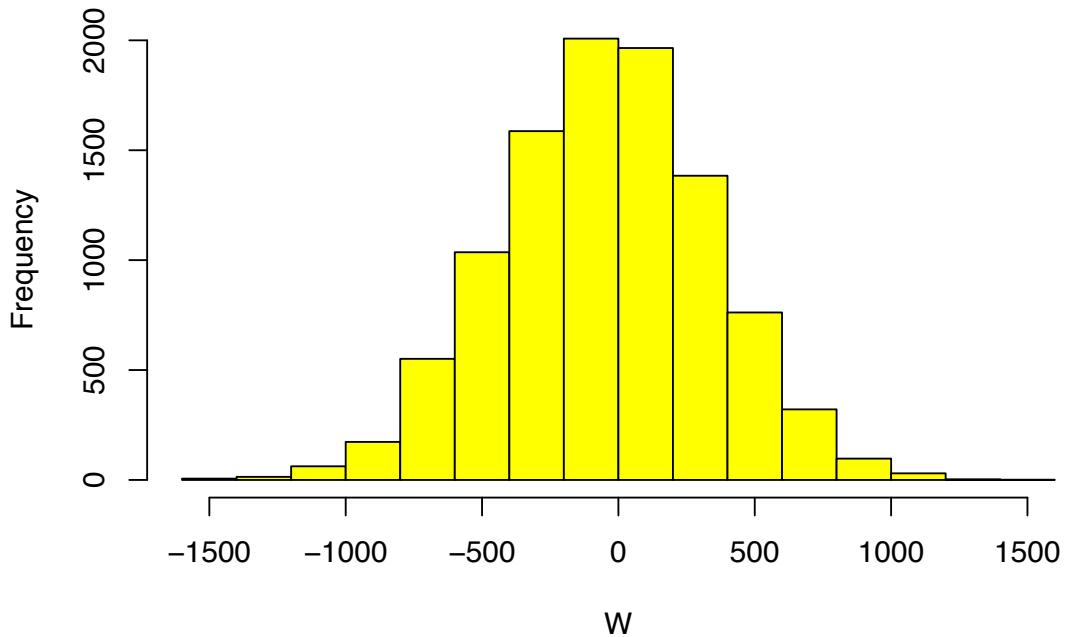
## [1] 150870.8  

##  

hist(W, col="yellow")

```

Histogram of W



- **Question 3**

Let's first work out the exact answer.

$$\begin{aligned}
\int_0^1 \cos(2\pi x) dx &= \frac{1}{2\pi} \sin(2\pi x) \Big|_0^1 \\
&= 0
\end{aligned}$$

- Note. Suppose we consider $U \sim \text{Uniform}(0,1)$. Let's look at the $E[U]$.

$$E[U] = \int_0^1 u f(u) du = \int_0^1 u du$$

- Now let's look at the expected value of the function: $\cos(2\pi u)$:

$$E[g(U)] = \int_0^1 \cos(2\pi u) du$$

So our integral of interest is the expected value of the function. We can approximate that via simulation.

$$\hat{I}(g(u)) = \frac{1}{S} \sum_1^S \cos(2\pi u)$$

```
set.seed(1001)
S1 <- 100

u <- runif(S1)
g.u <- cos(2*pi*u)
mean(g.u)

## [1] 0.05382692
```

```
set.seed(1001)
S2 <- 1000

u <- runif(S2)
g.u <- cos(2*pi*u)
mean(g.u)

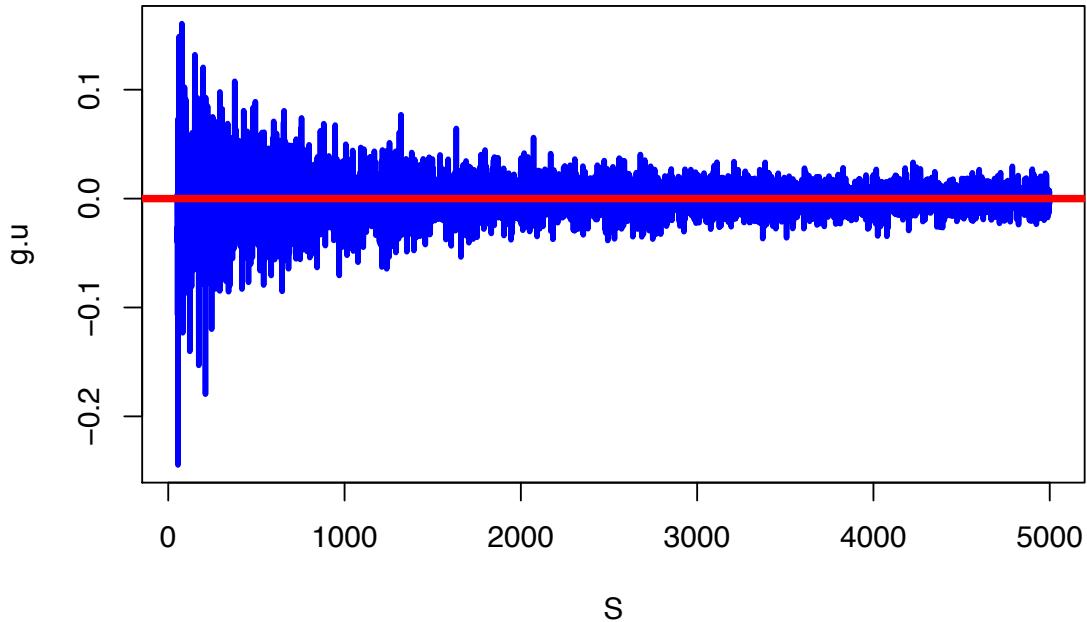
## [1] 0.006431414
```

- Now let's plot for increasing values of S :

```
S <- 50:5000
g.u <- rep(0, length(S))

c <- 1
for(s in S){
  g.u[c] <- mean(cos(2*pi* runif(s)))
  c <- c+1
}

plot(S, g.u, type="l", lwd=3, col="blue")
abline(h=0, col="red", lwd=4)
```



- As we can work out the $E[g(u)]$ and the $V[g(u)]$ we could also use the CLT theorem for calculations based on the $\frac{1}{S} \sum_1^S \cos(2\pi u)$.
- **Question 4**
- For this question we can't work out an analytical solution (some type of approximation must be performed).

$$\begin{aligned} E[I(g(u))] &= \int_0^1 \cos(2\pi u^2) du \\ &\approx \frac{1}{S} \sum_1^S \cos(2\pi u^2) \end{aligned}$$

```
set.seed(1001)
S1 <- 100

u <- runif(S1)
g.u <- cos(2*pi*u^2)
mean(g.u)

## [1] 0.3031969

set.seed(1001)
S2 <- 1000

u <- runif(S2)
g.u <- cos(2*pi*u^2)
mean(g.u)
```

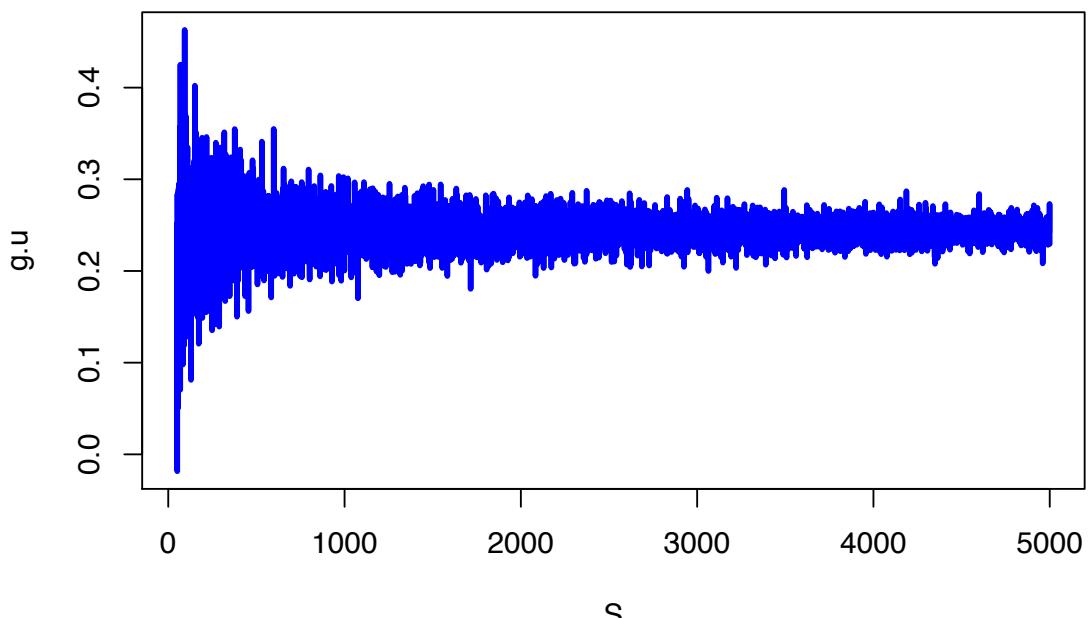
```

## [1] 0.2592541
S <- 50:5000
g.u <- rep(0, length(S))

c <- 1
for(s in S){
  g.u[c] <- mean(cos(2*pi* runif(s)^2))
  c <- c+1
}

plot(S, g.u, type="l", lwd=3, col="blue")

```



GJJ 2.6 $x_1, \dots, x_n \sim \text{Bernoulli}(\theta)$

$$f(x; \theta) = \theta^x (1-\theta)^{1-x}$$

CR lower bound for estimators of θ :

$$V(\hat{\theta}) \geq I_{\theta}^{-1} ; \quad I_{\theta} = -E\left(\frac{\partial^2 \ell}{\partial \theta^2}\right)$$

$$= E\left(\left(\frac{\partial \ell}{\partial \theta}\right)^2\right)$$

$$\Rightarrow \ell(\theta) = \prod_{i=1}^n \theta^{x_i} (1-\theta)^{1-x_i}$$
$$= \theta^{\sum x_i} (1-\theta)^{n-\sum x_i}$$

$$\ell(\theta) = \sum x_i \log(\theta) + (n-\sum x_i) \log(1-\theta)$$

$$\frac{\partial \ell}{\partial \theta} = \frac{\sum x_i}{\theta} - (n-\sum x_i) \frac{1}{1-\theta}$$

$$\frac{\partial^2 \ell}{\partial \theta^2} = -\frac{\sum x_i}{\theta^2} - \frac{(n-\sum x_i)}{(1-\theta)^2}$$

$$\Rightarrow E\left(\frac{\partial^2 \ell}{\partial \theta^2}\right) = -\frac{E(\sum x_i)}{\theta^2} - \frac{E((n-\sum x_i))}{(1-\theta)^2}$$

$$E(\sum x_i) = \sum E(x_i) = \sum \theta = n\theta$$

$$= -\frac{n\theta}{\theta^2} - \frac{n-n\theta}{(1-\theta)^2} = -\frac{n\cancel{\theta}}{\theta^2} - \frac{n(1-\cancel{\theta})}{(1-\theta)^2}$$

$$= -\frac{n}{\theta} - \frac{n}{(1-\theta)} = \frac{-n(1-\theta) - n\theta}{\theta(1-\theta)}$$

$$= \frac{-n}{\theta(1-\theta)} \quad \therefore I_{\theta} = -E\left(\frac{\partial^2 \ell}{\partial \theta^2}\right)$$

$$= -\left(\frac{n}{\theta(1-\theta)}\right)$$

$$= \frac{n}{\theta(1-\theta)}$$

$$C-R ZB = \frac{\theta(1-\theta)}{n}$$

Now the C-R ZB for an estimator
of $\theta^2 = g(\theta)$:

$$\begin{aligned} V(\hat{\theta}^2) &\leq \left(\left(\frac{\partial g}{\partial \theta}\right)^2 I_{\theta}^{-1}\right) \\ &= \frac{(2\theta)^2 \theta(1-\theta)}{n} \end{aligned}$$

$$= \frac{4\theta^3(1-\theta)}{n}$$

CTJ Q 2.8.) $x_1, \dots, x_n \stackrel{iid}{\sim} N(\mu, \theta = \sigma^2)$; assume μ is known.

$$S^2 = \frac{\sum_{i=1}^n (x_i - \bar{x})^2}{n-1}, \text{ note: } W = \frac{(n-1)S^2}{\sigma^2} \sim \chi_{n-1}^2$$

For the CRLB does this make a difference in this case? Why or why not?

$$E(W) = n-1$$

$$V(W) = 2(n-1)$$

$$E\left(\frac{(n-1)}{\sigma^2} S^2\right) = \frac{(n-1)}{\sigma^2} E(S^2) = (n-1)$$

$$E(S^2) = \sigma^2 \therefore S^2 \text{ is an unbiased estimator of } \sigma^2$$

$$V\left(\frac{(n-1)}{\sigma^2} S^2\right) = \frac{(n-1)^2}{(\sigma^2)^2} V(S^2) = 2(n-1)$$

$$V(S^2) = \frac{2(\sigma^2)^2}{(n-1)}$$

Now let's determine the CRLB:

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\theta}} \exp\left(-\frac{1}{2\theta}(x_i - \mu)^2\right) = \prod_{i=1}^n \frac{1}{\sqrt{2\pi\theta}} \exp\left(-\frac{1}{2\theta}(x_i - \mu)^2\right) \\ &= (2\pi)^{-n/2} \theta^{-n} \exp\left(-\frac{1}{2\theta} \sum (x_i - \mu)^2\right) \end{aligned}$$

$$\ell(\theta) = -\frac{n}{2} \log(2\pi) - \frac{n}{2} \log(\theta) - \frac{1}{2\theta} \sum (x_i - \mu)^2$$

$$\Rightarrow \ell'(\theta) = -\frac{n}{2} \frac{1}{\theta} + \frac{1}{2\theta^2} \sum (x_i - \mu)^2$$

$$\ell''(\theta) = \frac{n}{2} \frac{1}{\theta^2} - \frac{1}{\theta^3} \sum (x_i - \mu)^2$$

$$-\mathbb{E}(\ell''(\theta)) = -\frac{n}{2} \frac{1}{\theta^2} + \frac{1}{\theta^3} \mathbb{E}(\sum (x_i - \mu)^2)$$

$$= -\frac{n}{2} \frac{1}{\theta^2} + \frac{1}{\theta^3} \underbrace{\sum \mathbb{E}((x_i - \mu)^2)}_{=\theta} = -\frac{n}{2} \frac{1}{\theta^2} + \frac{n}{\theta^3} = n\theta$$

$$= -\frac{n}{2} \frac{1}{\theta^2} + \frac{n}{2\theta^2} = \frac{n}{2\theta^2} \Rightarrow \text{CRLB} = \frac{2\theta^2}{n}$$

- For finite n : $CRLB = \frac{2\theta^2}{n} = \frac{2\sigma^4}{n} < \frac{2\sigma^4}{n-1} = V(S^2)$

However, as $n \rightarrow \infty$ they both go to 0.

- Now we want to know the value 'c' which minimizes the MSE of:

$$c \sum_{i=1}^n (x_i - \bar{x})^2$$

- We know $MSE(\hat{\theta}) = V(\hat{\theta}) + [Bias(\hat{\theta})]^2$

$$\Rightarrow V(c \sum (x_i - \bar{x})^2) = V(c \frac{(n-1)}{(n-1)} \sum (x_i - \bar{x})^2)$$

$$= V(c(n-1) S^2) = c^2(n-1)^2 \frac{2\sigma^4}{(n-1)}$$

$$= c^2(n-1) 2\sigma^4$$

$$\Rightarrow E(c \sum (x_i - \bar{x})^2) = E(c(n-1) S^2)$$

$$= c(n-1) E(S^2) = c(n-1) \sigma^2$$

$$\therefore MSE(c \sum (x_i - \bar{x})^2) = c^2(n-1) 2\sigma^4 + [c(n-1)\sigma^2 - \sigma^2]^2$$

minimize MSE wrt c

$$\frac{d}{dc} = 2c(n-1) 2\sigma^4 + 2c(n-1)^2 \sigma^4 - 2c(n-1)\sigma^4 + \sigma^4$$

$$= (n-1)\sigma^4 [4c + 2c(n-1) - 2] = 0$$

$$\Rightarrow 4c + 2cn - 2c - 2 = 0$$

$$\Rightarrow c + cn - 1 = 0$$

$$\Rightarrow c(1+n) = 1 \Rightarrow c = \frac{1}{n+1}$$

Check to
see that
we have
a minimum

$$\frac{d^2}{dc^2} = (n-1) \sigma^n [n + 2(n-1)] > 0$$

for $n \geq 2$

∴ we have a minimum.

GJJ Q 2.10) $x_1, \dots, x_n \stackrel{iid}{\sim} N(\theta, \sigma^2)$, where σ^2 is known.

Lemma 2.1] Under our standard regularity conditions we have:

$$I = -E\left(\frac{\partial^2 \ell}{\partial \theta^2}\right)$$

$$\begin{aligned} L(\theta) &= \prod_{i=1}^n (2\pi\sigma^2)^{-\frac{1}{2}} \exp\left(-\frac{1}{2\sigma^2}(x_i - \theta)^2\right) \\ &= (2\pi\sigma^2)^{-\frac{n}{2}} \exp\left(-\frac{1}{2\sigma^2} \sum (x_i - \theta)^2\right) \end{aligned}$$

$$\ell(\theta) = -\frac{n}{2} \log(2\pi\sigma^2) - \frac{1}{2\sigma^2} \sum (x_i - \theta)^2$$

$$\ell'(\theta) = +\frac{2}{2\sigma^2} \sum (x_i - \theta) = \frac{1}{\sigma^2} \left[\sum x_i - n\theta \right]$$

$$\ell''(\theta) = -\frac{n}{\sigma^2} \Rightarrow I = -E\left(-\frac{n}{\sigma^2}\right) = \frac{n}{\sigma^2}$$

$$\therefore CRLB = \frac{1}{I} = \frac{\sigma^2}{n}$$