Inference for Regression with Variables Generated by AI or Machine Learning

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Outline

1. Introduction

- 2. Setup and Use Cases
- 3. Two-Step Inference is Biased
- 4. How to Do Valid Inference
- 5. Application: Remote Work and Wage Inequality
- 6. Application: CEO Time Use and Firm Performance
- 7. Application: Central Bank Communication
- 8. Conclusio

Motivation

Economists now routinely generate variables by AI/ML methods

- quantify unstructured data (text, images, ...)
- measure subtle concepts (uncertainty, sentiment, ...)
- generate variables previously too costly, labor-intensive, or infeasible to collect

The generated variables are inputs to downstream econometric models

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Almost all papers use a two-step strategy:

- 1. Generate estimate $\hat{\theta}_i$ of true variable θ_i using Al/ML algorithm
- 2. Plug estimates $(\hat{\theta}_i)$ into an econometric model, treating $\hat{\theta}_i$ as regular numeric data

This Paper

1. Two-step strategy leads to invalid inference: Cls have right width but wrong centering (bias)

Bias depends on relative importance of

- (a) measurement error in $\hat{\theta}_i$
- (b) sampling error in downstream model

Valid two-step inference requires (a) \ll (b)

This is not the case in most leading applications

2. Two solutions: bias correction + one-step strategy

NB: Measurement error is AI/ML-generated variables in non-classical.

3. Shows empirical relevance in several empirical applications

Related Work

Recent work (mainly stats/poli sci) has pointed out potential for generated variables to cause problems

- General ML-generated variables: Fong and Tyler (2021), Allon et al. (2023), Angelopoulos et al. (2023a, 2023b), Zhang et al. (2023), Zrnic and Candès (2024), and Miao and Lu (2024)
- Variables generated by LLMs: Egami et al. (2023, 2024), Ludwig et al. (2025), Carlson and Dell (2025).

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These works propose bias corrections assuming a validation subsample in which $(Y_i, \theta_i, \hat{\theta}_i)$ are observed

- But one can estimate the model using (Y_i, θ_i) in validation sample! AI/ML gen. vars only helpful for efficiency
- Not feasible in most economic use cases where θ_i is truly latent (uncertainty, sentiment, ...)

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Setup

Want: perform inference on γ and/or α in the model

$$Y_i = \gamma^T \theta_i + \alpha^T \mathbf{q}_i + \varepsilon_i, \qquad \mathbb{E}\left[\varepsilon_i | \theta_i, \mathbf{q}_i\right] = 0,$$

- θ_i is a latent variable of economic interest
- **q**_i are observed numeric covariates
- Unstructured/high-dim dataset \mathbf{x}_i available for estimating $\mathbf{ heta}_i$

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Two-Step Strategy:

- 1. Estimate $\hat{\theta}_i$ of θ_i obtained from unstructured data \mathbf{x}_i
- 2. Regress Y_i on $\hat{\theta}_i$ and \mathbf{q}_i . Perform inference treating $\hat{\theta}_i$ as regular numeric data.

Example 1: AI/ML-Generated Labels

- ML algorithms often deployed to impute missing observations from unstructured data.
 Goldsmith-Pinkham and Shue (2023), Adams-Prassl et. al. (2023), Argyle et al. (2025), and Wu and Yang (2024)
- Leading use case: missing θ_i is binary (e.g., race indicator)
- Generate estimate $\hat{\theta}_i$ of θ_i using unstructured data \mathbf{x}_i (e.g., voter registration data)
- Regress Y_i on $\hat{\theta}_i$ and controls \mathbf{q}_i

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- Regress Y_i on $\hat{\theta}_i$ and controls \mathbf{q}_i
- Measurement error due to misclassification error:

$$\mathsf{Pr}(heta_i = 1 | \mathbf{x}_i, \mathbf{q}_i)
eq \mathsf{Pr}(\hat{ heta}_i = 1 | \mathbf{x}_i, \mathbf{q}_i)$$

Example 2: Dimensionality Reduction

- · Obtain low-dimensional representation of unstructured data which is plugged into regression:
 - <u>Text data:</u> Hansen McMahon Prat (2018); Mueller and Rauh (2018); Larsen and Thorsrud (2019); Thorsrud (2020); Bybee Kelly Manela Xiu (2024); Ash Morelli Vannoni (2025)
 - Survey data: Bandiera Prat Hansen Sadun (2020); Draca and Schwarz (2020)
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- Obs i is a V-dim vector of feature counts \mathbf{x}_i
- Factor structure on multinomial probabilities (as in probabilistic latent semantic analysis/LDA):

$$\mathbf{x}_i | (C_i, \boldsymbol{\vartheta}_i) \sim \mathsf{Multinomial}(C_i, \mathbf{B}^T \boldsymbol{\vartheta}_i)$$

- $\mathbf{B}^T = [\beta_1, \dots, \beta_K]$, each $\beta_k \in \Delta^{V-1}$ is a topic
- observation-specific topic weights $artheta_i \in \Delta^{K-1}$
- subset of interest: $\theta_i = S\vartheta_i$

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- Measurement error due to upstream sampling error in $\hat{m{ heta}}_i$

Example 3: Indices

- Several influential works generate indices by classifying documents + aggregating
 Baker Bloom Davis (2016), Caldara and Iacoviello (2022), Gorodnichenko Pham Talavera (2023)
- Each month observe C_i documents (e.g., set of newspapers)
- Of these, X_i are classified as pertaining to concept (e.g., policy uncertainty)
- Latent true uncertainty $\theta_i \in [0,1]$
- Naive estimator: $\hat{\theta}_i = X_i/C_i$ (cf. BBD's EPU measure)

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- **Problem**: $\hat{\theta}_i$ is a signal of θ_i
- e.g., could change set of newspapers and get a different (but related) measure

Example 3: Indices

• Topic model representation:

$$\mathbf{x}_i | (C_i, \boldsymbol{\vartheta}_i) \sim \mathsf{Multinomial}(C_i, \mathbf{B}^T \boldsymbol{\vartheta}_i),$$
 for $\mathbf{x}_i = (X_i, C_i - X_i)^T$,
$$\mathbf{B}^T = \underbrace{\begin{bmatrix} \beta_1 & \beta_0 \\ (1 - \beta_1) & (1 - \beta_0) \end{bmatrix}}_{\mathsf{misclass. rates}}, \quad \boldsymbol{\vartheta}_i = \begin{bmatrix} \theta_i \\ 1 - \theta_i \end{bmatrix}$$

Measurement error due to misclassification error and upstream sampling error

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Asymptotics: General Case

• Consider a sequence of DGPs for $(Y_i, \theta_i, \hat{\theta}_i, \mathbf{q}_i, \mathbf{x}_i)_{i=1}^n$ indexed by sample size n, in which

$$\frac{1}{\sqrt{n}}\sum_{i=1}^n \hat{\theta}_i(\hat{\theta}_i - \theta_i)^T \to_{\rho} \kappa \Omega,$$

(expressions are DGP-specific)

- Scalar $\kappa \geq 0$ measures the importance of measurement error relative to sampling error
- Positive κ allows both sampling error and measurement error to play a role
- ullet Reflects prevailing trend: increasingly large data sets + increasingly accurate algorithms

Asymptotics: κ and Ω

• ML-generated binary labels:

$$\sqrt{n} imes \underbrace{\mathbb{E}\left[\hat{ heta}_i(1- heta_i)
ight]}_{ ext{false-positive rate}}
ightarrow \kappa, \qquad \mathbf{\Omega} = 1$$

• Topic models:

$$\sqrt{n} \times \mathbb{E}\left[\frac{1}{C_i}\right] \to \kappa, \qquad \mathbf{\Omega} = \mathbf{S}(\mathbf{B}\mathbf{B}^T)^{-1}\mathbf{B}\operatorname{diag}(\mathbf{B}^T\mathbb{E}[\vartheta_i])\mathbf{B}^T(\mathbf{B}\mathbf{B}^T)^{-1}\mathbf{S}^T - \mathbb{E}\left[\boldsymbol{\theta}_i\,\boldsymbol{\theta}_i^T\right]$$

Theorem on Two-Step Inference

Theorem: Two-Step Inference is Invalid Unless $\kappa = 0$

1. OLS estimator $\hat{\psi} = (\hat{\gamma}, \hat{\alpha})$ of $\psi = (\gamma, \alpha)$ from regressing Y_i on $\hat{\xi}_i = (\hat{\theta}_i, \mathbf{q}_i)$ has asy dist

$$\sqrt{n} \begin{pmatrix} \hat{\boldsymbol{\psi}} - \boldsymbol{\psi} \end{pmatrix} \rightarrow_{d} N \begin{pmatrix} -\kappa \mathbb{E}[\boldsymbol{\xi}_{i} \, \boldsymbol{\xi}_{i}^{T}]^{-1} \begin{pmatrix} \boldsymbol{\Omega} & \boldsymbol{0} \\ \boldsymbol{0} & \boldsymbol{0} \end{pmatrix} \boldsymbol{\psi} \,, \underbrace{\mathbb{E}[\boldsymbol{\xi}_{i} \, \boldsymbol{\xi}_{i}^{T}]^{-1} \mathbb{E}[\varepsilon_{i}^{2} \boldsymbol{\xi}_{i} \, \boldsymbol{\xi}_{i}^{T}] \mathbb{E}[\boldsymbol{\xi}_{i} \, \boldsymbol{\xi}_{i}^{T}]^{-1}}_{=: \mathbf{V}} \end{pmatrix}$$

where $\boldsymbol{\xi}_i = (\boldsymbol{\theta}_i, \mathbf{q}_i)$ are the "true" covariates

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where $\boldsymbol{\xi}_i = (\boldsymbol{\theta}_i, \mathbf{q}_i)$ are the "true" covariates

2. Eicker–Huber–White standard errors are consistent for all $\kappa \geq 0$:

$$\hat{\mathbf{V}} := \left(\frac{1}{n} \sum_{i=1}^{n} \hat{\boldsymbol{\xi}}_{i} \hat{\boldsymbol{\xi}}_{i}^{T}\right)^{-1} \left(\frac{1}{n} \sum_{i=1}^{n} \hat{\varepsilon}_{i}^{2} \hat{\boldsymbol{\xi}}_{i} \hat{\boldsymbol{\xi}}_{i}^{T}\right) \left(\frac{1}{n} \sum_{i=1}^{n} \hat{\boldsymbol{\xi}}_{i} \hat{\boldsymbol{\xi}}_{i}^{T}\right)^{-1} \rightarrow_{p} \mathbf{V}$$

Implications

- $\kappa \in (0, \infty)$: two-step inference is **biased**
 - degree of bias is increasing in κ (relative importance of measurement vs sampling error)
 - no variance distortion, unlike generated regressors
- $\kappa=0$: two-step inference is **valid**: treat $\hat{\theta}_i$ as if they are the true latent θ_i
- Take-away: if κ is large, consider using resources to improve precision of $\hat{\pmb{\theta}}_i$

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- $\kappa = 0$: two-step inference is **valid**: treat $\hat{\theta}_i$ as if they are the true latent θ_i
- Take-away: if κ is large, consider using resources to improve precision of $\hat{\theta}_i$
- To the extent empirical papers flag concerns about 2-step inference, usually about std errors
- Common intuition is wrong: problem is measurement error not standard errors

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How to Do Valid Inference

1. Explicit Bias Correction: use analytical expressions in Theorem to adjust two-step estimates/Cls

Advantage: Simple and scalable

Disadvantage: Not feasible in complex models; poor approximation with large κ

2. One-Step Strategy: MLE using joint likelihood for upstream IR model + regression model

Advantage: General purpose and flexible

Disadvantage: More computationally demanding

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Hansen Lambert Bloom Davis Sadun Taska (WP, 2023)

- Consider n = 16,315 SD food+accom sector (NAICS code 72) job postings from January 2022
- Regress log wages Y_i on ML-generated remote work indicator $\hat{ heta}_i$
- Fixed effects for SOC code (job type) and full/part time

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- Regress log wages Y_i on ML-generated remote work indicator $\hat{ heta}_i$
- Fixed effects for SOC code (job type) and full/part time
- Estimate FPR from a subsample of size m=1,000 postings, $\widehat{FPR} \approx 0.009$
- For 1-step: use 3 component Gaussian mixture for errors $arepsilon_i| heta_i$

Two-Step Estimates Smaller

| | | No Fixed Effects | | | With Fixed Effects | | |
|--------|-------|------------------|----------------|-------|--------------------|----------------|--|
| | Est. | Std Err | 95% CI | Est. | Std Err | 95% CI | |
| OLS | 0.648 | 0.024 | [0.599, 0.697] | 0.363 | 0.021 | [0.321, 0.406] | |
| ВС | 1.052 | 0.140 | [0.777, 1.326] | 0.641 | 0.099 | [0.446, 0.836] | |
| 1-Step | 0.563 | 0.016 | [0.532, 0.595] | 0.448 | 0.017 | [0.415, 0.480] | |

Corrected CIs to the right of Two-Step CIs

| | | No Fixed Effects | | | With Fixed Effects | | |
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Bandiera Hansen Prat Sadun (JPE, 2020)

- Time-use survey data for 916 CEOs
- 654 combinations of activities (e.g., meeting with suppliers) in 15min intervals
- LDA with K=2: 2 types of CEO behaviors β_1 (leaders) and β_2 (managers).
- Two-step strategy: regress log sales Y_i on leader weight $\hat{\theta}_{i,1}$ and firm characteristics \mathbf{q}_i .

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Original Paper: $\hat{\kappa} = 0.44$ (average $C_i = 88.4$).

Modified Sample: draw 10% of activities for each CEO (without replacement) $\longrightarrow \hat{\kappa} = 4.26$.

Similar Estimates in Full Sample

Table 1: Estimates of Impact of CEO Behavior on Firm Performance

| Sample | Estimation Strategy | | | |
|---------------|---------------------|-----------------|----------------|--|
| | Two-Step | Bias Correction | Joint | |
| Full | 0.405 | 0.474 | 0.402 | |
| | [0.224, 0.585] | [0.294, 0.655] | [0.240, 0.603] | |
| 10% Subsample | 0.227 | 1.054 | 0.439 | |
| | [-0.038, 0.492] | [0.789, 1.319] | [0.153, 0.711] | |

Difference in Subsample

Table 2: Estimates of Impact of CEO Behavior on Firm Performance

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Central Bank Communication

• Does written central bank communication drive long rates? Estimate

$$Y_i = \gamma \theta_i + \alpha' \mathbf{q}_i + u_i$$

- Y_i is the path factor from Gürkaynak, Sack, and Swanson (2005) (mkt perceptions of future rates)
- θ_i is a hawkish/dovish index (cf. Gorodnichenko, Pham, Talavera (2023))
- **q**_i are controls (including shadow short rate)

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- qi are controls (including shadow short rate)
- Hawkish/dovish index:
 - ullet classify FOMC sentences as hawkish/dovish/neutral using fine-tuned BERT + aggregate
 - sentiment estimate

$$\hat{\theta}_i = \frac{N_i^H - N_i^D}{N_i^H + N_i^D}$$

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Compare 1 and 2 step methods over 02/1995-06/2023

Central Bank Communication: One-Step Effect Size 3x Larger

| | Estimation Strategy | | |
|------------------------|---------------------|-----------------|--|
| | Two-Step | One-Step | |
| Sentiment (θ_i) | 0.039 | 0.114 | |
| | [0.012, 0.066] | [0.027, 0.198] | |
| Policy Rate (q_i) | -0.004 | -0.003 | |
| | [-0.011, 0.003] | [-0.011, 0.004] | |
| $eta_{f 0}$ | | 0.009 | |
| | | [0.001, 0.026] | |
| eta_1 | | 0.676 | |
| | | [0.585, 0.768] | |
| Observations | 200 | 200 | |
| R^2 | 0.0425 | 0.1429 | |

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Conclusion

- Empirical work routinely uses AI/ML algorithms to generate new variables
- Common empirical practice leads to invalid inference
- We propose two solutions: bias correction + one-step strategy
- Illustrate important differences in simulations + applications
- Works in progress: specific methods tailored to important use cases
 - $\bullet\,$ VARs and impulse response analysis w/ Hansen and Shin