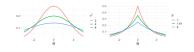
Introduction to Machine Learning

Regularization Bayesian Priors





Learning goals

- RRM is same as MAP in Bayes
- Gaussian/Laplace prior corresponds to L2/L1 penalty

RRM VS. BAYES I

We already created a link between max. likelihood estimation and ERM.

Now we will generalize this for RRM.

Assume we have a parameterized distribution $p(y|\theta, \mathbf{x})$ for our data and a prior $q(\theta)$ over our param space, all in Bayesian framework.



From Bayes theorem:

$$p(\theta|\mathbf{x},y) = \frac{p(y|\theta,\mathbf{x})q(\theta)}{p(y|\mathbf{x})} \propto p(y|\theta,\mathbf{x})q(\theta)$$

RRM VS. BAYES II

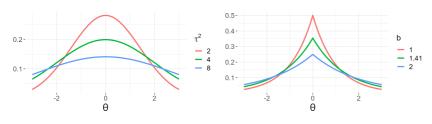
The maximum a posteriori (MAP) estimator of θ is now the minimizer of

$$-\log p(y\mid \boldsymbol{\theta},\mathbf{x})-\log q(\boldsymbol{\theta}).$$

- Again, we identify the loss $L(y, f(\mathbf{x} \mid \theta))$ with $-\log(p(y|\theta, \mathbf{x}))$.
- If $q(\theta)$ is constant (i.e., we used a uniform, non-informative prior), the second term is irrelevant and we arrive at ERM.
- If not, we can identify $J(\theta) \propto -\log(q(\theta))$, i.e., the log-prior corresponds to the regularizer, and the additional λ , which controls the strength of our penalty, usually influences the peakedness / inverse variance / strength of our prior.



RRM VS. BAYES III





- L2 regularization corresponds to a zero-mean Gaussian prior with constant variance on our parameters: $\theta_i \sim \mathcal{N}(0, \tau^2)$
- L1 corresponds to a zero-mean Laplace prior: $\theta_j \sim Laplace(0, b)$. Laplace(μ , b) has density $\frac{1}{2b} \exp(-\frac{|\mu-x|}{b})$, with scale parameter b, mean μ and variance $2b^2$.
- In both cases, regularization strength increases as variance of prior decreases: more prior mass concentrated around 0 encourages shrinkage.
- Elastic-net regularization corresponds to a compromise between Gaussian and Laplacian priors
 Zou and Hastie 2005
 Hans 2011

EXAMPLE: BAYESIAN L2 REGULARIZATION I

We can easily see the equivalence of L2 regularization and a Gaussian prior:

• Gaussian prior $\mathcal{N}_d(\mathbf{0}, diag(\tau^2))$ with uncorrelated components for θ :

$$q(\theta) = \prod_{j=1}^{d} \phi_{0,\tau^2}(\theta_j) = (2\pi\tau^2)^{-\frac{d}{2}} \exp\left(-\frac{1}{2\tau^2} \sum_{j=1}^{d} \theta_j^2\right)$$

MAP:

$$\begin{split} \hat{\theta}^{\text{MAP}} &= & \arg\min_{\boldsymbol{\theta}} \left(-\log p \left(y \mid \boldsymbol{\theta}, \mathbf{x} \right) - \log q(\boldsymbol{\theta}) \right) \\ &= & \arg\min_{\boldsymbol{\theta}} \left(-\log p \left(y \mid \boldsymbol{\theta}, \mathbf{x} \right) + \frac{d}{2} \log(2\pi\tau^2) + \frac{1}{2\tau^2} \sum_{j=1}^{d} \theta_j^2 \right) \\ &= & \arg\min_{\boldsymbol{\theta}} \left(-\log p \left(y \mid \boldsymbol{\theta}, \mathbf{x} \right) + \frac{1}{2\tau^2} \|\boldsymbol{\theta}\|_2^2 \right) \end{split}$$

• We see how the inverse variance (precision) $1/\tau^2$ controls shrinkage



EXAMPLE: BAYESIAN L2 REGULARIZATION II

- DGP $y = \theta + \varepsilon$ where $\varepsilon \sim \mathcal{N}(0, 1)$ and $\theta = 1$; with Gaussian prior on θ , so $\mathcal{N}(0, \tau^2)$ for $\tau \in \{0.25, 0.5, 2\}$
- For n = 20, posterior of θ and MAP can be calculated analytically
- Plotting the *L*2 regularized empirical risk $\mathcal{R}_{reg}(\theta) = \sum_{i=1}^{n} (y_i \theta)^2 + \lambda \theta^2$ with $\lambda = 1/\tau^2$ shows that ridge solution is identical with MAP
- In our simulation, the empirical mean is $\bar{y} = 0.94$, with shrinkage toward 0 induced in the MAP

