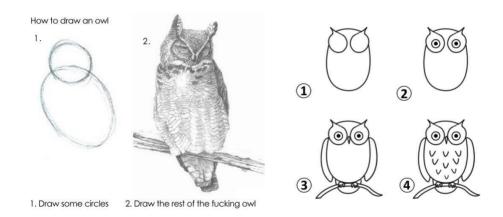
## An introduction to bayesian data analysis with PyMC3

Sean Meling Murray and Solveig Masvie

inmeta

17. februar 2020

## Googling bayesian analysis vs this presentation



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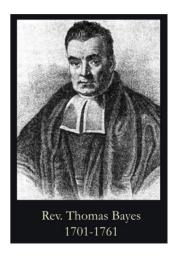
## Road map

- 1. Theory
  - (a) The basics of Bayesianism
- (b) Markov chain Monte Carlo methods (MCMC)
- 2. Practice
  - (a) Probabilistic programming with PyMC3

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## What is Bayesian data analysis?

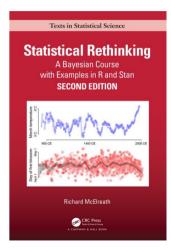
"A Bayesian is one who, vaguely expecting a horse, and catching a glimpse of a donkey, strongly believes he has seen a mule."



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## It's just counting!

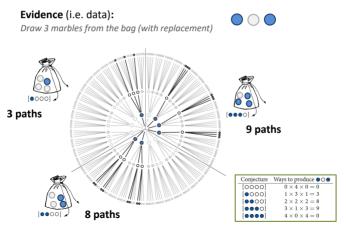
- Richard McElreath: "Bayesian inference is just counting."
- Count all the ways observed data could have arisen according to assumptions
- Assumptions that can arise in more ways are more consistent with the data, and therefore more plausible



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### The Garden of Forking Data

#### Counting possibilities



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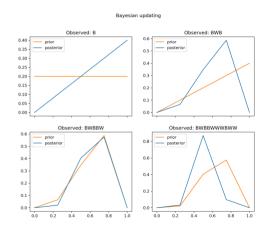
## Priors, posteriors and likelihoods

Prior  $p(\theta)$ : Encodes assumptions about  $\theta$ Likelihood  $p(y|\theta)$ : How were the observed data generated? Posterior  $p(\theta|y)$ : Assumptions

Posterior  $p(\theta | y)$ : Assumptions about  $\theta$  consistent with data

$$p(\theta|y) = \frac{p(\theta)p(y|\theta)}{\int p(y|\theta)p(\theta)d\theta}$$
$$\approx p(\theta)p(y|\theta)$$

 $posterior \propto prior \cdot likelihood$ 



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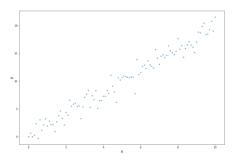
#### Model

- In the Bayesian analysis we are trying to describe the process that generated the data.
- ullet We assume that the process can be expressed as a parametic model M(input, ullet)

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## Example

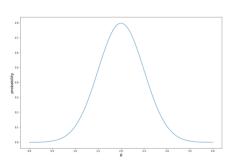
The oh so boring linear regression example....

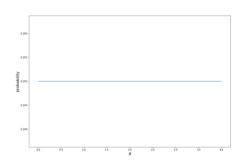


Model: 
$$y_i = \theta x_i + \varepsilon_i, \quad \varepsilon \sim N(0, \sigma^2)$$

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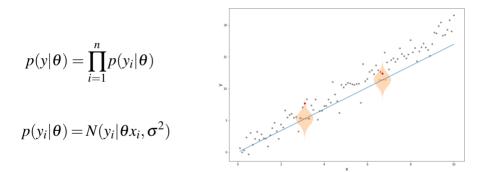
## **Example: Priors**





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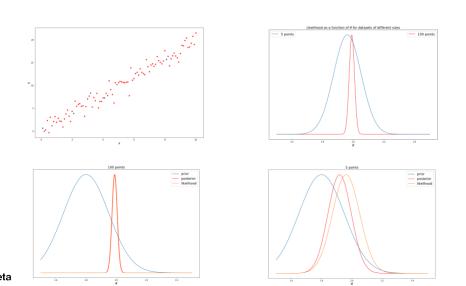
## Example: Likelihood



What is the probability of my data given the model and parameters I've chosen?

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## Example: Size of dataset



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## Towards a Bayesian workflow

- 1. Design the model (tell the data story)
  - (a) Describe data generating process:  $p(y|\theta) = \theta^{\sum_i \mathbf{1}(y_i=1)} (1-\theta)^{\sum_i \mathbf{1}(y_i\neq 1)}$
- (b) Encode assumptions about parameters:  $p(\theta) = \frac{1}{5}$
- 2. Condition on the data (update step)
  - (a)  $p(\theta|y) \propto p(\theta)p(y|\theta)$
- 3. Evaluate the model (critique), and either
  - (a) be happy, it worked as planned, or
  - (b) return to step 1

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## The Garden of Forking Data, revisited

```
import numpy as np
## Observed data
data = np.array([1, 0, 1])
## Conjectures (possible values for theta)
# arrav([0., 0.25, 0.5, 0.75, 1.1)
theta = np.linspace(0, 1, num=5)
# Prior distribution over theta
# array([0.2, 0.2, 0.2, 0.2, 0.2])
uniform prior = np.repeat(0.2, repeats=5)
# Likelihood function
def likelihood(data, theta):
    num blue = data.sum()
    num white = len(data) - num blue
    return theta**num blue * (1 - theta) **num white
# Unnormalised posterior over theta
# array([0., 0.009375, 0.025, 0.028125, 0.])
unnormalised posterior = uniform prior * likelihood(data, theta)
# Posterior distribution over theta
# arrav([0., 0.15, 0.4, 0.45, 0.1)
posterior = unnormalised posterior / unnormalised posterior.sum()
```

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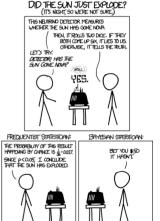
## The Garden of Forking Data, revisited

# import numpy as np # Number of paths consistent with conjectures forking\_data = np.array([0, 3, 8, 9, 0]) # array([0, 0.15, 0.4, 0.45, 0.] forking\_data / forking\_data.sum()

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## The Frequentist vs. Bayesian debacle

- Frequentist statistics
  - ▶ Probability = limiting frequency
  - ▶ Uncertainty arises from sampling variation
- Bayesian statistics
  - ► Frequency ≠ probability
  - Uncertainty arises from ignorance
- Intuitive way to represent and take into account uncertainty
- Easily incorporates prior information



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## But priors are subjective! Buu! Hiss!

- Non-informative vs. informative priors
  - ► Forking data example: Uniform(0,1) vs.  $Uniform(\frac{1}{4},\frac{3}{4})$
- Test your assumptions, e.g. using prior predictive checks
- More about priors:
  - ► Gelman et al. (2017): https://arxiv.org/abs/1708.07487
  - ► Bayesian Methods for Hackers: https://tinyurl.com/v2yzyqv
- For more Bayesian challenges, see e.g. Gelman and Yao (2020): https://tinyurl.com/sbr2tev

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## Telling the data story

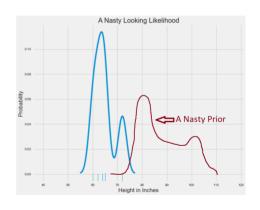
"People think in terms of stories - thus the unreasonable power of the anecdote to drive decision-making, well-founded or not. Existing analytics largely fails to provide this kind of story; instead, numbers seemingly appear out of thin air."

Beau Cronin, Why Probabilistic Programming Matters (2013)

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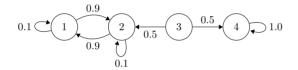
## How to sample from intractable posteriors?

- In textbooks, nice, we can sample directly from well-behaved, analytical posteriors
- In the real world, often impossible to sample directly from p(θ|y)
- A general class of algorithms called Markov chain Monte
   Carlo let us approximately sample from posterior



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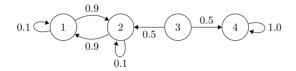
#### Markov chains



- A Markov chain defines a joint probability distribution over a sequence of events
- Used to model events in discrete and continuous time (manufacturing processes, queuing systems, etc.)
- Typically interested in the long-term distribution of being in a certain state given an initial starting state

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## Properties of a Markov chain



- Memoryless (Markov property):  $P(X_t | X_{t-1}, X_{t-2}, \dots, X_1) = P(X_t | X_{t-1})$
- Under certain conditions, guaranteed to have a unique, limiting distribution (aka. stationary distribution)
  - ▶ Long run proportion of time spent in each state
  - ► For more details, check
    https://github.com/smu095/presentations/

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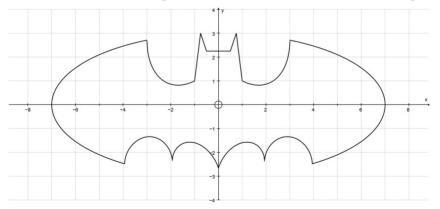
#### Monte Carlo simulations

- Monte Carlo (MC) simulations are just a way to approximate numerical results using repeated random sampling
- Main idea:
  - Generate *N* samples  $x_1, \ldots, x_N$  from p(x), approximate *f* using the empirical distribution of  $\{f(x_n)\}_{n=1}^N$
  - $\triangleright \mathbb{E}[f] = \int f(x)p(x) dx \approx \frac{1}{N} \sum_{n=1}^{N} f(x_n) = \hat{f}$
- MC estimates converge thanks to Law of Large Numbers
  - $\blacktriangleright (\hat{f} \mathbb{E}[f]) \to \mathcal{N}\left(0, \frac{\mathit{Var}[f]}{\mathit{N}_{\mathit{eff}}}\right) \text{ as } N \to \infty$

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## Application: Evaluate integrals

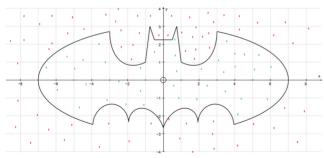
• Calculate difficult integrals, such as the area of the Batman sign



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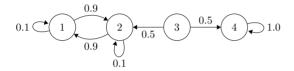
## Application: Evaluate integrals

- Repeatedly sample  $(u_1, u_2) \sim \text{Uniform}(-\frac{1}{2}, \frac{1}{2})$
- Calculate area A of Batman sign as  $A = \text{area of rectangle} \times \frac{\text{green dots}}{\text{all dots}}$



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## Properties of a Markov chain



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#### Monte Carlo simulations

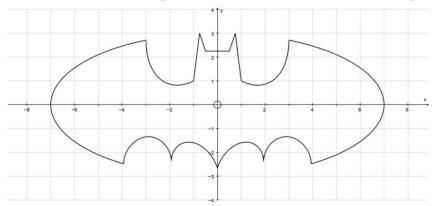
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$$\triangleright (\hat{f} - \mathbb{E}[f]) \to \mathcal{N}\left(0, \frac{Var[f]}{N_{eff}}\right) \text{ as } N \to \infty$$

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## Application: Evaluate integrals

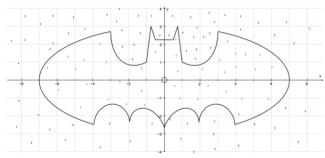
• Calculate difficult integrals, such as the area of the Batman sign



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## Application: Evaluate integrals

- Repeatedly sample  $(u_1, u_2) \sim \text{Uniform}(-\frac{1}{2}, \frac{1}{2})$
- Calculate area A of Batman sign as  $A = \text{area of rectangle} \times \frac{\text{green dots}}{\text{all dots}}$



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#### Markov chain Monte Carlo methods

- Main idea: Construct a Markov chain over the parameter space where the stationary distribution is the posterior  $p(\theta | y)$ 
  - ► For more details, check
    https://github.com/smu095/presentations/
- Randomly move around the parameter space such that the fraction of time spent at each randomly sampled parameter value is proportional to the true target density  $p(\theta | y)$ .
- Use the sequence of parameters generated by Markov chain to calculate MC approximations of any quantity of interest

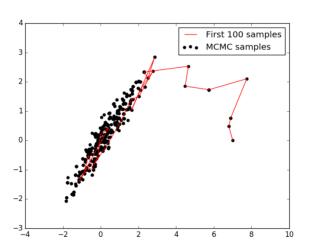
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## Algorithmic view of MCMC

- 1. Start at current position.
- 2. Propose new position.
- 3. Accept/reject new position based on position's adherence to data and prior distributions.
  - (a) If reject: Remain in your current position, return to step 1.
  - (b) If accept: Move to the new position, return to step 1.
- 4. After large number of iterations, return sequence of accepted positions.

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#### MCMC illustration



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## Metropolis algorithm

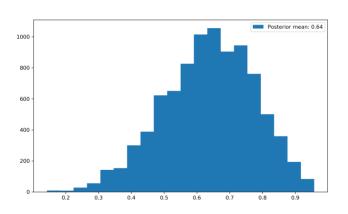
- 1. Start at some initial parameter value  $\theta_0$
- 2. For t = 1 to T:
  - (a) Sample  $\theta^*$  from a proposal distribution  $q(\theta^* | \theta_{t-1})$   $\theta^* = \theta_{t-i} + \mathcal{N}(0, \text{jump scale})$
- (b) Evaluate  $\alpha_{\theta^*} = \frac{p(\theta^*|y)}{p(\theta_{t-1}|y)}$ .
- (c) Accept/reject  $\theta^*$ If  $\alpha_{\theta^*} \ge 1$  set  $\theta_t = \theta^*$ Else draw  $u \sim \textit{Uniform}(0,1)$  set  $\theta_t = \begin{cases} \theta^*, & \text{if } u < \alpha_{\theta^*} \\ \theta_{t-1}, & \text{o.w.} \end{cases}$

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## Metropolis in Python

```
"""Metropolis sampling from posterior of mean"""
<<<<<< HEAD
import numpy as np
# Marbles data
data = np.array([1, 0, 1])
def log_p(theta, x):
    N = len(x)
    k = np.sum(x)
    logprior = np.repeat(0.2, len(x))
    loglik = k*np.log(theta + 10e-8) + (N - k)*np.log(1 - theta + 10e-8)
    return np.sum(logprior + loglik)
# Initial value for theta
theta = 0.5
import numpy as np
data = np.arrav([1, 0, 1]) # Observations
def log_p(theta, x): # Proportional log posterior
    N, k = len(x), np.sum(x)
    logprior = np.repeat(0.2, N)
    loglik = k*np.log(theta + 10e-8) + (N - k)*np.log(1 - theta + 10e-8)
    return np.sum(logprior + loglik)
```

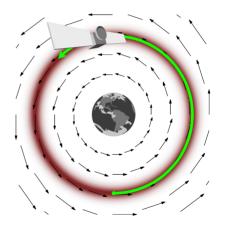
## Metropolis in Python



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## Sampling algorithms

- Many different sampling algorithms exist
- Currently, the state-of-the-art is Hamiltonian Monte Carlo (HMC)
  - ► See e.g. Betancourt (2017): https://arxiv.org/abs/1701.0



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## Probabilistic programming in Python

- PyStan
- PyMC3
- Pyro
- TensorFlow Probability









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## PyMC3

- Python package for for fitting Bayesian models using modern methods (MCMC, variational inference, etc.)
- Well-documented, large suite of statistical models
- Highly flexible, but easy to use
- Uses Theano-backend (not good)
- PyMC4 in development, uses TensorFlow Probability-backend



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