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#43

a)

Majority decoding: based on the assumption that the largest number of occurrences of a symbol was the transmitted symbol.

$X$ : # of error in a message

$$X \sim b(5, 0.2)$$

$P(\text{the message will be wrong when coded})$

$$= P(X \geq 3) = \sum_{x=3}^5 \binom{5}{x} (0.2)^x (0.8)^{5-x}$$

b)

All of the transmission in a message are independent,  
digit

#50

$X$ : # of head result while flipping an biased coin for 10 times

$$X \sim b(10, p)$$

c)

$$P(\text{first 3 outcomes: h.t.t} \mid X=6) = \frac{P(\text{first 3 outcomes: h.t.t}, X=6)}{P(X=6)}$$

$$= \frac{P(h)p \cdot P(t) \cdot P(t) \cdot P(h)^5}{\binom{10}{6} p^6 (1-p)^4}$$

$$= \frac{\binom{7}{5} p^5 (1-p)^2}{\binom{10}{6} p^6 (1-p)^4} = \frac{\binom{7}{5}}{\binom{10}{6}} = \frac{1}{10}$$

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cb)

$$P(\text{first three outcomes: t, h, t} \mid X=b) = \frac{P(\text{first three outcomes: t, h, t}, X=b)}{P(X=b)}$$

$$= \frac{(1-p)^2 p \binom{7}{5} p^5 (1-p)^2}{\binom{10}{6} p^6 (1-p)^4} = \frac{1}{10}$$

# 53.

ca)

$$P(\text{both partners were born on } 9/30) = \left(\frac{1}{365}\right)^2$$

$X$ : # of couple that both born on 9/30 within 80000 couples

$$X \sim b(80000, \left(\frac{1}{365}\right)^2) \xrightarrow[n \rightarrow \infty]{d} \text{poisson}(\lambda = 0.6)$$

$$\lambda = 80000 \times \left(\frac{1}{365}\right)^2 = np$$

$$\Rightarrow P(X \geq 1) = 1 - e^{-0.6}$$

cb)

$P(\text{both partners celebrated their birthday on same day of the yr})$

$$= \left(\frac{1}{365}\right)^2 \times 365 = \frac{1}{365}$$

$Y$ : # of couple celebrated their BD on same day of the yr  
within 80000 couples,

$$Y \sim b(80000, \frac{1}{365}) \xrightarrow[n \rightarrow \infty]{d} \text{poisson}(\lambda = \frac{80000}{365})$$

$$P(X \geq 1) = 1 - e^{-\frac{80000}{365}}$$

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64. 令 r.v.  $X \sim \text{Poi}(1)$ . 則  $Y = 4X \sim \text{Poi}(4)$ .

$$(a) P(Y=8) = 1 - P(Y < 8) = 1 - \sum_{i=0}^7 \frac{e^{-4} 4^i}{i!}$$

$$(b). \text{ 令 (a) 中 } 1 - \sum_{i=0}^7 \frac{e^{-4} 4^i}{i!} = p$$

令 r.v.  $W \sim \text{Bin}(12, p)$ , 總共有 12 月. 每月發生機率都

$$P(W \geq 2) = 1 - P(W=0) - P(W=1) = 1 - \binom{12}{0} p^0 (1-p)^{12} - \binom{12}{1} p^1 (1-p)^{11} \\ = 1 - (1-p)^{12} - 12p(1-p)^{11}$$

16. 題意為令今天所在的月份為第一個月, 而第一次  
出現有  $\geq 8$  自殺的月份為第  $i$  月份. 請問機率為?  
令 r.v.  $Z \sim \text{Geo}(p)$ .

$$P(Z=i) = (1-p)^{i-1} p, i \geq 1$$

假設: 每月事件發生皆獨立. (課本 p158. 8.1 第一行).

74. (a) 令  $X \sim \text{Bin}(5, \frac{2}{3})$ , 總共 5 人. 每人願意來的機率為  $\frac{2}{3}$ .

$\therefore$  interviewer 想面 5 人. 故全部人都要來.

$$P(X=5) = \binom{5}{5} \left(\frac{2}{3}\right)^5 \left(\frac{1}{3}\right)^0 = \frac{32}{243}$$

(b) 令  $Y \sim \text{Bin}(8, \frac{2}{3})$ .  $\therefore$  interviewer 想面 5 人. 故 5, 6, 7, 8 人都來

$$P(Y \geq 5) = \sum_{y=5}^8 \binom{8}{y} \left(\frac{2}{3}\right)^y \left(\frac{1}{3}\right)^{8-y} \quad (\text{面試數})$$

需達 5 人.

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74. [Cont.]

(c). interviewer 剛好 5 人面試. 意謂 前 5 人中  
只面試了 4 人. 而第 6 人必面.

於是 (a) 以設  $X \sim \text{Bin}(5, \frac{2}{3})$ .

$$\text{則 } p(X=4) \times \frac{2}{3} = \binom{5}{4} \left(\frac{2}{3}\right)^4 \left(\frac{1}{3}\right) \times \frac{2}{3} = \binom{5}{4} \left(\frac{2}{3}\right)^5 \left(\frac{1}{3}\right) = \frac{160}{729} \\ (0.2195) \text{ 耳}$$

(d). 同 (c) 想 =  $Z$ .

令  $Z \sim \text{Bin}(6, \frac{2}{3})$ .

$$\text{則 } p(Z=4) \times \frac{2}{3} = \binom{6}{4} \left(\frac{2}{3}\right)^4 \left(\frac{1}{3}\right)^2 \times \frac{2}{3} = \binom{6}{4} \left(\frac{2}{3}\right)^5 \left(\frac{1}{3}\right)^2 = \frac{160}{729} \\ (0.2195) \text{ 耳}$$

28. 令  $X \sim \text{Geo}(p)$ . 要做到  $n$  次 selection 亦即 前  
(n-1) 次皆沒抽到  $W \geq B$ .

$$\text{其中 } p = \frac{\binom{4}{2} \binom{4}{2}}{\binom{8}{4}} = \frac{36}{70} = \frac{18}{35} \approx 0.5143$$

$$\text{則 } p(X=n) = \left(\frac{17}{35}\right)^{n-1} \frac{18}{35} = 18(17)^{n-1} / 35^n \text{ 耳}$$

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# Theoretical Exercises

#10

$$X \sim b(n, p)$$

$$E\left(\frac{1}{X+1}\right) = \sum_{x=0}^n \left(\frac{1}{x+1}\right) \frac{n!}{x!(n-x)!} p^x (1-p)^{n-x}$$

$$= \frac{1}{p(n+1)} \sum_{x=0}^n \frac{(n+1)!}{(x+1)!(n-x)!} p^{x+1} (1-p)^{n-x}$$

$$\stackrel{z=x+1}{=} \frac{1}{p(n+1)} \sum_{z=1}^{n+1} \frac{(n+1)!}{z!(n+1-z)!} p^z (1-p)^{n+1-z}$$

$$= \frac{1}{p(n+1)} \left( (1-p+p)^{n+1} - \frac{(n+1)!}{0!(n+1-0)!} p^0 (1-p)^{n+1-0} \right)$$

$$= \frac{1}{p(n+1)} (1 - (1-p)^{n+1})$$

#16

$$X \sim \text{Poisson}(\lambda)$$

$$\frac{P(X=\tilde{n})}{P(X=\tilde{n}-1)} = \frac{\frac{\lambda^{\tilde{n}} e^{-\lambda}}{\tilde{n}!}}{\frac{\lambda^{\tilde{n}-1} e^{-\lambda}}{(\tilde{n}-1)!}} = \frac{\lambda}{\tilde{n}}$$

$\frac{\lambda}{\tilde{n}} \geq 1$  if  $\lambda \geq \tilde{n}$ , indicates that  $P(X=\tilde{n})$  increase monotonically as  $\tilde{n}$  increases,  $\rightarrow \textcircled{1}$

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$$\frac{P(\tilde{n}+1)}{P(\tilde{n})} = \frac{e^{-\lambda} \frac{\lambda^{\tilde{n}+1}}{(\tilde{n}+1)!}}{e^{-\lambda} \frac{\lambda^{\tilde{n}}}{\tilde{n}!}} = \frac{\tilde{n}}{\lambda}$$

$\Rightarrow \frac{\tilde{n}}{\lambda} \leq 1$  if  $\tilde{n} \leq \lambda$  indicates that  $P(X=\tilde{n})$

decreases monotonically,

as  $\tilde{n}$  decreases  $\rightarrow$  (2)

by (1), (2),

$P(X=\tilde{n})$  reaches its maximum when  $\tilde{n}$  is the largest integer not exceeding  $\lambda$ .

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(theoretical exercise)

19.  $X \sim \text{poi}(\lambda)$ ,  $X=0, 1, 2, \dots$ 

$$E(X^n) = \sum_{x=0}^{\infty} x^n \frac{e^{-\lambda} \lambda^x}{x!} = \sum_{x=0}^{\infty} x^{n-1} \frac{e^{-\lambda} \lambda^x}{(x-1)!} = \lambda \sum_{x=0}^{\infty} x^{n-1} \frac{e^{-\lambda} \lambda^{x-1}}{(x-1)!}$$

$$= \lambda \sum_{x=1}^{\infty} x^{n-1} \frac{e^{-\lambda} \lambda^{x-1}}{(x-1)!} = \lambda \sum_{y=0}^{\infty} (y+1)^{n-1} \frac{e^{-\lambda} \lambda^y}{y!}$$

(當  $x=0$  時, 求和為 0) (令  $y=x-1$ )

$$= \lambda \sum_{x=0}^{\infty} (x+1)^{n-1} \frac{e^{-\lambda} \lambda^x}{x!} = \lambda E[(X+1)^{n-1}]$$

(再將  $y$  換回  $x$ )

$$\text{故 } E[X^3] = \lambda E[(X+1)^2] = \lambda E(X^2) + 2\lambda E(X) + \lambda$$

$$= \lambda (\text{Var}(X) + [E(X)]^2) + 2\lambda^2 + \lambda$$

$$= \lambda^3 + 3\lambda^2 + \lambda \quad \text{已知 } \text{Var}(X) = E(X) = \lambda$$

20. 題意為每輪丟  $n$  個銅板, 若  $n$  個銅板中有至少一個正面則停止. 若無正面則再丟一輪, 直到出現至少一個正面.

$$\text{故 } P(X=1) = \binom{n}{1} p(1-p)^{n-1} + (1-p)^n \binom{n}{1} p(1-p)^{n-1} + (1-p)^{2n} \binom{n}{1} p(1-p)^{n-1} + \dots$$

$$= \binom{n}{1} p(1-p)^{n-1} [1 + (1-p)^n + (1-p)^{2n} + \dots]$$

$$= \binom{n}{1} p(1-p)^{n-1} \frac{1}{1-(1-p)^n} = np(1-p)^{n-1} \frac{1}{1-(1-p)^n}$$

$$(p = \frac{\lambda}{n}) = \lambda (1 - \frac{\lambda}{n})^{n-1} \frac{1}{1 - (1 - \frac{\lambda}{n})^n} \xrightarrow{n \rightarrow \infty} \frac{\lambda e^{-\lambda}}{1 - e^{-\lambda}} \quad \# (b).$$