Sequential Monte-Carlo Filtering/Smoothing for State and Parameter Estimation in General Nonlinear State Space Systems

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Agenda

- Problem statement: Nonlinear State Space Systems (NSSS)
- State and parameter estimation in NSSS
- Expectation Maximization (EM) for joint state and parameter estimation in NSSS
- Handling non-linearity: Sequential Monte Carlo (SMC)
- Particle Filter/Smoother for joint state and parameter estimation

Nonlinear State Space Systems

We consider a family of discrete-time and affine NSSSs:

$$\mathbf{x}_k = f_{\theta^{\times}}(\mathbf{x}_{k-1}, \mathbf{u}_k) + \mathbf{v}_k$$

 $\mathbf{y}_k = g_{\theta^{\times}}(\mathbf{x}_k) + \mathbf{w}_k$

- Where:
 - \mathbf{x}_k : hidden state at the k-th step-time
 - \mathbf{y}_k : output observation at the k-th step-time
 - f,g: parametric state transition and output models with parameters θ^x and θ^y , respectively
 - \mathbf{u}_k : external input
 - $\mathbf{v}_k, \mathbf{w}_k$: additive noise

NSSS in Time-series Modeling with Missing Data

- Considering \mathbf{y}_k as a vector of output observations at the k-th step:
 - Time-series may suffer from missing-data where some of outputs are not measurable
 - Weather-data: daily max/min temperature is available, but the associated time-of-year may be missing
- NSSS: model time-series as a state-space system with missing-data being represented by hidden-states
- Training the model, infer missing-data from the received observations

NSSS in Model Reference Control

The goal in Model Reference Control (MRC):

$$\mathbf{u}_{1:K} = \operatorname{argmin}_{\mathbf{u}} \left\{ \sum_{k} (y_k - \hat{y}_k)^2 \right\}$$

with \hat{y}_k being a desired trajectory

• Consider a parametric structure for the controller: $\mathbf{u}_k = h_{\theta^u}(\mathbf{x}_{k-1})$ and re-construct the NSSS:

$$\mathbf{x}_k = f_{\theta^{\times}}(\mathbf{x}_{k-1}, h_{\theta^u}(\mathbf{x}_{k-1})) + \mathbf{v}_k$$

- Jointly estimate states and parameters to minimize MRC objective
- ullet Generate control command using the learned $h_{ heta^u}$



NSSS in Observer Trajectory Planning

- Consider a moving object with motion state \mathbf{x}_k , and a moving observer with state \mathbf{x}_k^o that gathers measurements y_k
- NSSS in dynamical target tracking:

$$\mathbf{x}_k = A\mathbf{x}_{k-1} + B\mathbf{v}_k$$

 $y_k = g_{\mathbf{x}_k^o}(\mathbf{x}_k) + \mathbf{w}_k$

with A, B being known matrices of state-transition model, and g as a nonlinear measurement model such as bearing

 The goal is to estimate observer locations to minimize variance of target state estimation

State and Parameter Estimation in NSSS

- Three main sources in NSSS:
 - Observations: $\{\mathbf{y}_{0:k}, \mathbf{u}_{0:k}\}$ • Parameters: $\{\theta^{\mathsf{x}}, \theta^{\mathsf{y}}\}$
 - Hidden states: x_{0·k}
- State estimation: estimate posterior density of hidden states given observations:
 - Filtering: $p(\mathbf{x}_{0:k}|\mathbf{y}_{0:k},\mathbf{u}_{0:k})$
 - Smoothing: $p(\mathbf{x}_{0:k'}|\mathbf{y}_{0:k},\mathbf{u}_{0:k})$, k' < k
- Parameter estimation: obtain an estimate of model parameters by optimizing a proper objective such as likelihood

$$\theta^* = \operatorname{argmax}_{\theta} \left\{ p(\mathbf{y}_{0:k}; \theta) \right\}$$



Maximum Likelihood Estimator

Parameters can be obtained by maximizing the log-likelihood function

$$\theta^* = \operatorname{argmax}_{\theta} \left(J = \sum_{k} \log \left\{ p(\mathbf{y}_k; \theta) \right\} \right)$$

- Incomplete likelihood: the likelihood cannot be calculated since hidden states are missing
- Main trick: given the concavity of log-function and using Jensen inequality:

$$\log \left\{ \underbrace{p(\mathbf{y}_k; \theta)}_{E_q(\frac{p(\mathbf{y}_k; \theta)}{q(\mathbf{x})})} \right\} \ge E_{q(\mathbf{x})} \left(\underbrace{p(\mathbf{y}_k | \mathbf{x}_k; \theta)}_{\text{Complete Likelihood}} \right) - \mathcal{H}(\mathbf{x}_k)$$

with q(.) being a distribution over states and $\mathcal{H}(.)$ as the entropy

Expectation Maximization (EM) Algorithm

- Instead of maximizing the likelihood, we aim to maximize the lower-bound
- EM algorithm consists of two main phases:
 - E-phase: assuming parameters of the model are fixed, minimize the bound-gap through minimizing the following kullback leibler divergence:

$$D_{KL}(q||p) = -E_q\left(\log\left(\frac{p(\mathbf{x}_k|\mathbf{y}_k;\theta)}{q(\mathbf{x}_k)}\right)\right)$$

 M-phase: given estimates of hidden-states, maximize the expectation over the complete likelihood:

$$\theta^* = \operatorname{argmax}_{\theta} E_q \left(\log \left(p(\mathbf{y}_k | \mathbf{x}_k; \theta) \right) \right)$$



Iterative State/Parameter Estimation Using EM

- Initialize model parameters, training iteration n=0 and model parameters $\theta[0]$
- Do until convergence:
 - E-phase: by estimating $q(\mathbf{x}_k) = p(\mathbf{x}_k | \mathbf{y}_k; \theta[n]), k \in \{0, \dots, K\}$
 - M-phase: update parameters by maximizing $\sum_{k=0}^{K} E_q \left(\log \left(p(\mathbf{y}_k | \mathbf{x}_k; \theta) \right) \right) \text{ with } q \text{ being used from the E-phase}$
 - Calculate log-likelihood using estimated states and new parameter values
 - If no significant change in log-likelihood **break**, otherwise set $n \leftarrow n+1$
- Theorem: Under exact state estimation, EM guarantees log-likelihood increases at each step: $\mathcal{L}(\theta[n+1]) \geq \mathcal{L}(\theta[n])$ with $\mathcal{L}(\theta)$ being the complete likelihood

Nonlinearity and State/parameter Estimation

- Under additive Gaussian noises and linear state-transition and output models:
 - Optimal filtering such as **Kalman filter** and Gaussian distributed posterior $p(\mathbf{x}_k|\mathbf{y}_k;\theta) = \mathcal{N}\left(\mathbf{x}_{k|k},P_{k|k}\right)$ with $\mathbf{x}_{k|k}$ and $P_{k|k}$ being mean and covariance of estimates, respectively
 - Obtain a **closed-form** for $\sum_{k=1}^{K} E_q \left(\log \left(p(\mathbf{y}_k | \mathbf{x}_k; \theta) \right) \right)$ and find the optimal parameter value
- Under general nonlinear state-transition and/or ourput models:
 - Local linearization and Extended Kalman Filter for approximate E-phase and Gradient updates for M-phase
 - Approximate the posterior distribution of states using samples and find an unbiased estimate of the log-likelihood that could be maximized against parameters of the model

Approximate State Estimation Using Sequential Monte Carlo (SMC)

 Main idea is to represent the posterior distribution as a weighted summation of samples:

$$p(\mathbf{x}_k|\mathbf{y}_{1:k}) \sim \sum_{m=1}^{M} \tilde{w}_k^m \delta(\mathbf{x}_k - \mathbf{x}_k^m)$$

with \tilde{w}_k^m being normalized importance weights, $\delta(.)$ as the delta-Dirac function and M as the total number of particles

• Approximate M-phase with SMC sampler:

$$\theta^* = \operatorname{argmax}_{\theta} \sum_{k=1}^{K} \sum_{m=1}^{M} \tilde{w}_k^m \log (\mathbf{y}_k | \mathbf{x}_k^m; \theta)$$

Sequential Importance Sampling

• Draw samples from a proposal distribution $s(\mathbf{x}_{0:k}|\mathbf{y}_{0:k})$ and update importance weights recursively:

$$w_k^m = w_{k-1}^m \frac{p(\mathbf{x}_k^m | \mathbf{x}_{k-1}^m) p(\mathbf{y}_k | \mathbf{x}_k^m)}{s(\mathbf{x}_k^m | \mathbf{x}_{k-1}^m, \mathbf{y}_k)}$$

and normalize the weights $\tilde{w}_k^m = \frac{w_k^m}{\sum_{m=1}^M w_k^m}$

- Choice of proposal density function:
 - State-transition: $s(\mathbf{x}_k^m | \mathbf{x}_{k-1}^m, \mathbf{y}_k) = p(\mathbf{x}_k^m | \mathbf{x}_{k-1}^m)$
 - Using the outputs and sample from $p(\mathbf{x}_k^m|\mathbf{x}_{k-1}^m)p(\mathbf{y}_k|\mathbf{x}_k^m)$
 - Use sub-optimal Extended Kalman Filter (EKF) to obtain an estimate of the posterior $p(\mathbf{x}_k|\mathbf{y}_k)$ and assign the distribution to the proposal

Sequential Importance Re-sampling

- Importance sampling provides estimates whose variance increases exponentially over time
 - Only a few particles survive with a high weight with other particles will be assigned to a zero-weight
- We define re-sampling as an operation that maps the set $\{\tilde{w}_k^m, \mathbf{x}_k^m\}$ to a set of equally-weighted states $\{\frac{1}{M}, \bar{\mathbf{x}}_k^m\}$
- Idea: samples with higher weights will be replicated, but lower-weighted samples also survive
- Different re-sampling strategies:
 - Residual re-sampling, systematic re-sampling and multinomial re-sampling
 - Adaptive re-sampling: re-sample only if the effective sample-size of weights is below a threshold:

$$\left(\sum_{m=1}^{M} (\tilde{w}_k^m)^2\right)^{-1} \le \tau$$

with
$$\tau = \frac{M}{2}$$

General SMC Sampler for Filtering

- Sample initial particles \mathbf{x}_0^m from an arbitrary distribution
- Compute initial normalized weights \tilde{w}_0^m
- Check the re-sampling criterion and if needed, run re-sampling and generate $\{\frac{1}{N}, \bar{\mathbf{x}}_0^m\}$
- For k > 0:
 - Draw new samples from the proposal distribution $s(\mathbf{x}_k^m | \mathbf{\bar{x}}_{k-1}^m, \mathbf{y}_k)$
 - Compute the normalized weights \tilde{w}_k^m
 - Check for the re-sampling criterion:
 - If effective sample-size is lower than the threshold, run re-sampling and $\{\tilde{w}_k^m, \mathbf{x}_k^m\} \to \{\frac{1}{M}, \bar{\mathbf{x}}_k^m\}$
 - If effective sample-size is above the threshold, use $\{\tilde{w}_k^m, \mathbf{x}_k^m\}$ as the set of particles and associated weights

Summary: EM with SMC

- **Initialization**: randomly initialize parameters $\theta[0]$ and M particles with associated weights $\{\frac{1}{M}, \mathbf{x}_0^m\}$
- Do until convergence:
 - E-phase: run one step of SMC sampler and obtain sampled states and associated weights $\{\tilde{w}_k^m, \bar{\mathbf{x}}_k^m\}, m \in \{1, \dots, M\}, k \in \{0, \dots, K\}$
 - M-phase: approximate the complete log-likelihood and obtain new parameters through $\theta[n+1] = argmax_{\theta} \sum_{k=1}^{K} \sum_{m=1}^{M} \tilde{w}_{k}^{m} \log \left(\mathbf{y}_{k} | \mathbf{\bar{x}}_{k}^{m} ; \theta \right)$
 - Stop if the likelihood does not change significantly $|\mathcal{L}(\theta[n+1]) \mathcal{L}(\theta[n])| \le \epsilon$ or $n \ge n_{max}$, otherwise $n \leftarrow n+1$

State Estimation: Smoothing

- Given a batch of observations $(\mathbf{y}_{0:k})$, estimate $p(\mathbf{x}_k|\mathbf{y}_{0:K})$ instead of filtering distribution:
 - Makes state estimation more robust to the process or observation noise
 - Allows the training procedure to have a better estimate of model parameters
- Different perspectives:
 - Forward-backward: Use estimated states at the filtering stage (forward path) to approximate smoothing distribution
 - Two-filter smoothing: Break down the smoothing distribution into two separate filtering stages and estimate each stage through a single filter
 - Maximum a Posteriori smoothing: Run the forward path and estimate states. Then, obtain smoothed estimates by maximizing the posterior distribution of hidden states over measurements (through Dynamic Programming)

Forward-backward Particle Smoothing

- Modeling the smoothing distribution as $p(\mathbf{x}_k|\mathbf{y}_{0:K}) \sim \sum_{m=1}^{M} \tilde{w}_{k|K}^m \delta(\mathbf{x} \mathbf{x}_k^m)$
- Updating weights using the following forward-backward rule:

$$\tilde{w}_{k|K}^{m} = \tilde{w}_{k}^{m} \left\{ \sum_{m=1}^{M} \tilde{w}_{k+1|K}^{m} \frac{p(\mathbf{x}_{k+1}^{m}|\mathbf{x}_{k}^{m})}{\sum_{m'=1}^{M} \tilde{w}_{k}^{m'} p(\mathbf{x}_{k+1}^{m'}|\mathbf{x}_{k}^{m'})} \right\}$$

- Some considerations: the computational complexity is of order M^3
- There are approximations to reduce to $M^2 \log(M)$ by partitioning particles
- EM Particle Smoothing: calculate the complete likelihood using samples generated by smoothing distribution



Summary

- Family of discrete-time nonlinear state space systems was discussed
- An NSSS is fully characterized through hidden states and its model parameters
- EM algorithm is used to estimate hidden states and learn parameters of the model
- SMC sampler (particle filtering) aims to provide an approximate low-variance estimate of the posterior distribution of states
- EM with particle filtering/smoothing forms a universal framework for jointly estimating hidden states and learning parameters of a general nonlinear state space system

References

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Questions

