Appendix H

Inverse-Gaussian Distribution

The normalized inverse-Gaussian distribution is a probability distribution on the unit simplex, and can be viewed as an analog of the finite-dimensional Dirichlet distribution. It arises from normalizing a vector of independent inverse-Gaussian variables, which replace the gamma variables that would yield the Dirichlet distribution.

Recall that the *inverse-Gaussian distribution* IGau (α, γ) with shape parameter $\alpha > 0$ and scale parameter $\gamma > 0$ is the probability distribution on $(0, \infty)$ with density

$$y \mapsto \frac{\alpha e^{\alpha \gamma}}{\sqrt{2\pi}} \frac{1}{y^{3/2}} e^{-\frac{1}{2}(\alpha^2/y + \gamma^2 y)}.$$

The IGau $(0, \gamma)$ distribution is understood to be the distribution degenerate at 0. (The distribution takes its name from the distribution of the time until a Brownian motion $B = (B_t: t \ge 0)$ with linear drift reaches a given level: the hitting time $\inf\{t > 0: B_t + t = \alpha\}$ possesses the IGau $(\alpha, 1)$ -distribution. This hitting time can be considered to be "inverse" to the (Gaussian) distribution of this process itself.)

Definition H.1 (Normalized inverse-Gaussian distribution) Given $Y_i \stackrel{\text{ind}}{\sim} \text{IGau}(\alpha_i, 1)$ and $Y = \sum_{i=1}^k Y_i$, the vector $(Y_1, \dots, Y_k)/Y$ is said to possess the *normalized inverse-Gaussian distribution* NIGau $(k; \alpha)$ with parameters $k \in \mathbb{N}$ and $\alpha = (\alpha_1, \dots, \alpha_k) \geq 0$.

As the second parameter γ of an inverse-Gaussian distribution is a scale parameter, and one degree of freedom is lost by normalization, taking this parameter in the distribution of the variables Y_i in the preceding definition to be 1 is not a loss of generality.

The following propositions give elementary properties of inverse-Gaussian distributions; these resemble properties of the Dirichlet distribution.

Proposition H.2 (Aggregation)

- (i) If $Y_i \stackrel{ind}{\sim} IGau(\alpha_i, \gamma)$, then $\sum_i Y_i \sim IGau(\sum_i \alpha_i, \gamma)$.
- (ii) If $(W_1, ..., W_k) \sim \text{NIGau}(k; \alpha_1, ..., \alpha_k)$ and $Z_j = \sum_{i \in I_j} W_i$ for a given partition $I_1, ..., I_m$ of $\{1, ..., k\}$, then $(Z_1, ..., Z_m) \sim \text{NIGau}(m; \beta_1, ..., \beta_m)$, where $\beta_j = \sum_{i \in I_i} \alpha_i$, for j = 1, ..., m.

Proposition H.3 (Density) If $\alpha_i > 0$ for i = 1, ..., k, then the vector $(W_1, ..., W_k) \sim \text{NIGau}(k; \alpha_1, ..., \alpha_k)$ possesses a probability density function, given by

$$w = (w_1, \dots, w_k) \mapsto \frac{e^{\sum_{i=1}^k \alpha_i} \prod_{i=1}^k \alpha_i}{2^{k/2-1} \pi^{k/2}} \prod_{i=1}^k \frac{1}{w_i^{3/2}} \frac{K_{-k/2} \left(\left(\sum_{i=1}^k \alpha_i^2 / w_i \right)^{1/2} \right)}{\left(\sum_{i=1}^k \alpha_i^2 / w_i \right)^{k/4}}, \quad w \in \mathbb{S}_k,$$

where K_m is the modified Bessel function of the third kind of order m (see Watson (1995)). In particular, the marginal density of W_1 is

$$w \mapsto \frac{e^{\alpha_1 + \alpha_2} \alpha_1 \alpha_2}{\pi} \frac{1}{w^{3/2} (1 - w)^{3/2}} \frac{K_{-1} \left((\alpha_1^2 / w + \alpha_2^2 / (1 - w))^{1/2} \right)}{(\alpha_1^2 / w + \alpha_2^2 / (1 - w))^{1/2}}, \qquad w \in (0, 1).$$

Proposition H.4 (Laplace transform) The Laplace transform of $Y \sim \text{IGau}(\alpha, \gamma)$ is given by $\text{E}(e^{-\lambda Y}) = \exp\left[-\alpha(\sqrt{2\lambda + \gamma^2} - \gamma)\right]$, for $\lambda \geq 0$.

Proposition H.5 (Mean) If $(W_1, ..., W_k) \sim \text{NIGau}(k; \alpha_1, ..., \alpha_k)$, then, $E(W_i) = \alpha_i / \sum_{j=1}^k \alpha_j$.

Proof Write $W_i = Y_i/Y$, where $Y_j \stackrel{\text{ind}}{\sim} \text{IGau}(\alpha_i, 1)$ and $Y = \sum_{j=1}^k Y_j$. Also let $Y_{-i} = \sum_{j \neq i} Y_j$ and $|\alpha| = \sum_{j=1}^k \alpha_j$. Then

$$E(W_i) = \int_0^\infty E(Y_i e^{-uY_i}) du = \int_0^\infty E(Y_i e^{-uY_i}) E(e^{-uY_{-i}}) du$$
$$= -\int_0^\infty \frac{d}{du} E(e^{-uY_i}) E(e^{-uY_{-i}}) du.$$

Now $Y_{-i} \sim \text{IGau}(|\alpha| - \alpha_i, \gamma)$, by Proposition H.2. Therefore, by Proposition H.4 the preceding display is equal to

$$-\int_0^\infty \frac{d}{du} e^{-\alpha_i(\sqrt{2u+1}-1)} e^{-(|\alpha|-\alpha_i)(\sqrt{2u+1}-1)} du = \alpha_i \int_0^\infty (2u+1)^{-1/2} e^{-|\alpha|(\sqrt{2u+1}-1)} du.$$

We conclude the proof by the change of variables $y = \sqrt{2u + 1} - 1$.

The "generalized" inverse-Gaussian distribution replaces the power $y^{-3/2}$ in the density of an inverse-Gaussian distribution by a general one.

Definition H.6 (Generalized inverse-Gaussian distribution) The *generalized inverse-Gaussian distribution* GIGau(a, b, p) with parameters $\alpha > 0$, $\gamma > 0$ and $p \in \mathbb{R}$ is the probability distribution on $(0, \infty)$ with density

$$y \mapsto \frac{(\alpha/\gamma)^p}{2K_p(\alpha\gamma)} y^{p-1} e^{-\frac{1}{2}(\alpha^2/y + \gamma^2 y)},$$

where K_p is the modified Bessel function of the third kind.