

ECON 0150 | Economic Data Analysis

The economist's data analysis skillset.

Part 3.4 | Hypothesis Testing

A Big Question

How do we learn about the population when we don't know μ or σ ?

- **Part 3.1 | Known Random Variables**
 - *If we know the random variable, we can answer all kinds of probability questions*
- **Part 3.2 | Sampling and Unknown Random Variables**
 - *The sample means of unknown random variables will approximate a normal distribution around the truth*
- **Part 3.3 | Confidence Intervals**
 - *We can use the sampling distribution to know the probability that the sample mean (\bar{x}) will be close to the population mean (μ)*

Sampling Distribution: Unknown μ ; Known σ

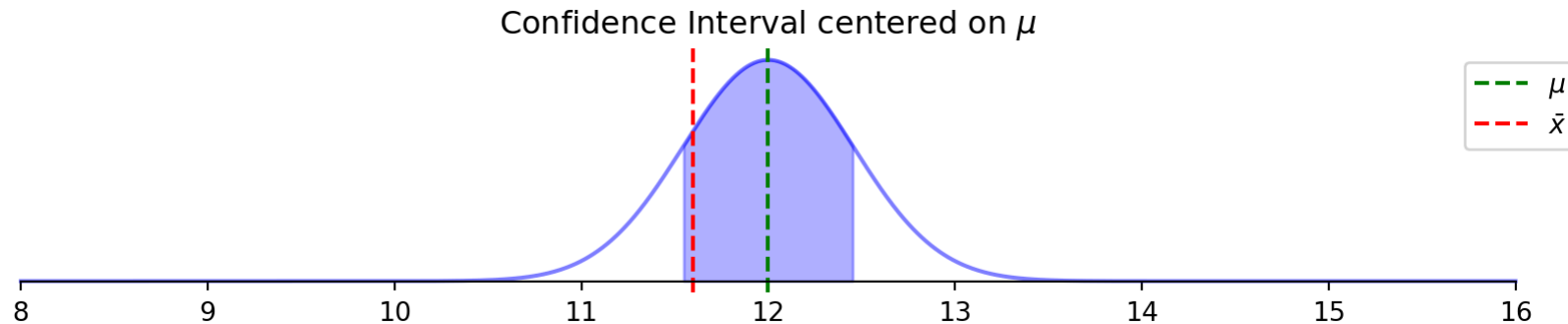
If we know the population mean, we know the sampling distribution is approximately normal.

- *The sample mean is drawn from an approximately normal distribution with mean μ and standard error σ/\sqrt{n} .*
- *Each time we draw a sample we see a different sample mean.*
- *What do we do that we don't observe μ ? We measure 'closeness'.*

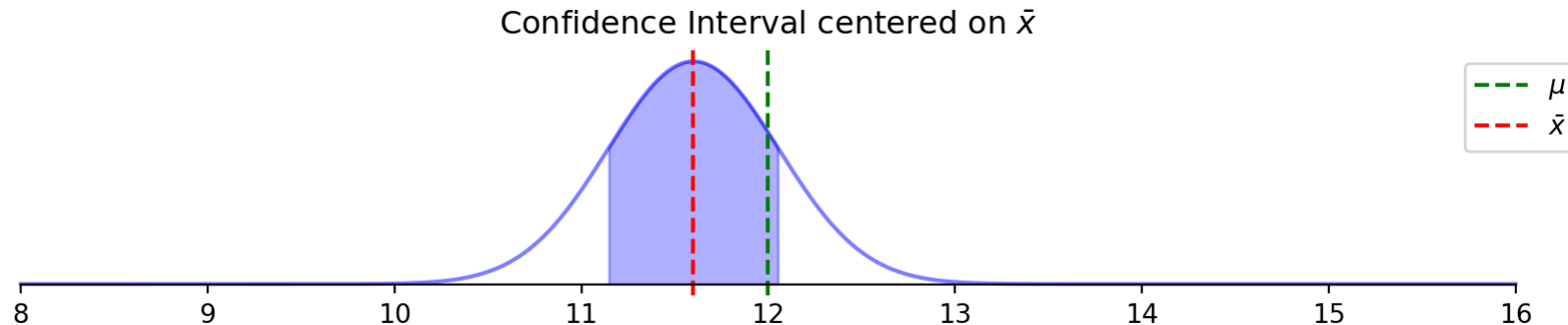
Unknown μ : Two Perspectives

There are two mathematically equivalent perspectives to think about “closeness” between μ and \bar{x} .

Perspective 1: probability \bar{x} is close to μ



Perspective 2: probability μ is close to \bar{x}

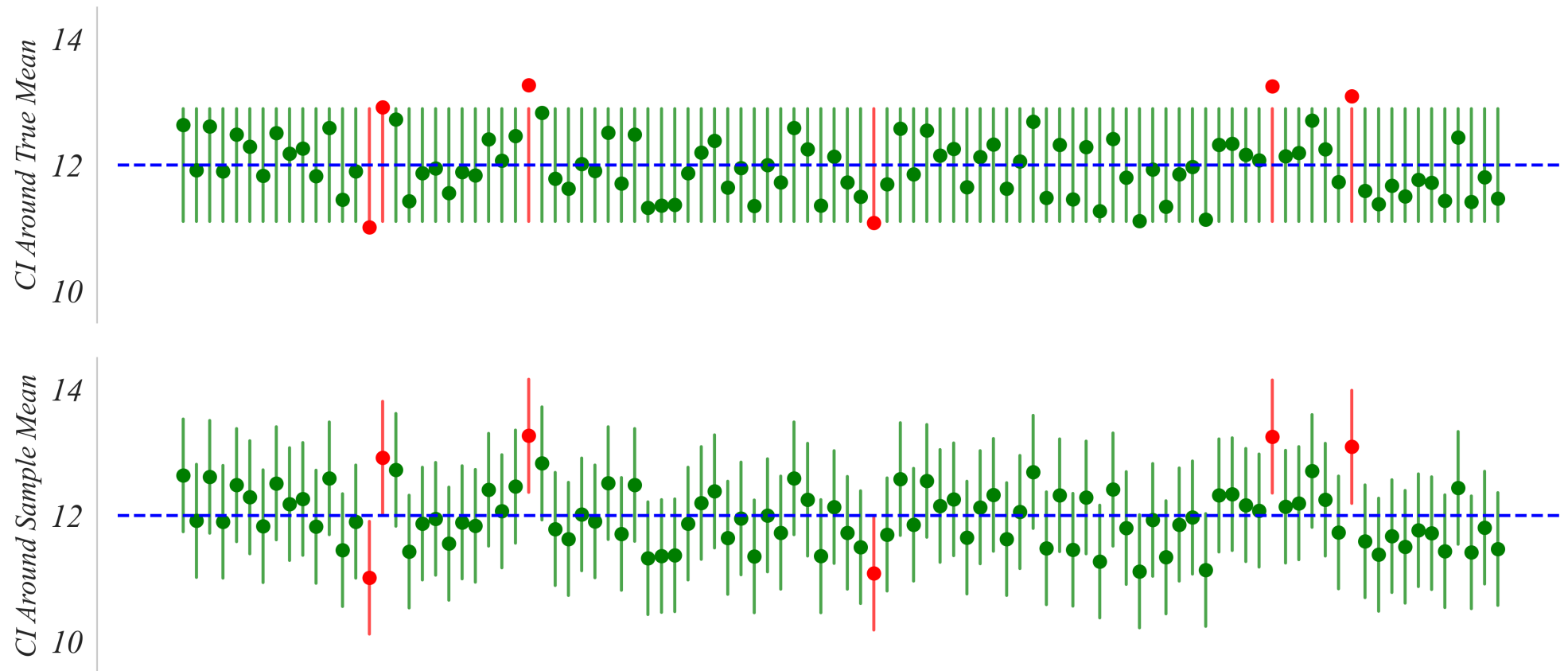


> if \bar{x} is in the CI around μ , then μ will be in the CI around \bar{x} !

Unknown μ : Two Perspectives

There are two mathematically equivalent perspectives to think about “closeness” between μ and \bar{x} .

I repeatedly sampled a distribution and constructed a 95% confidence interval.

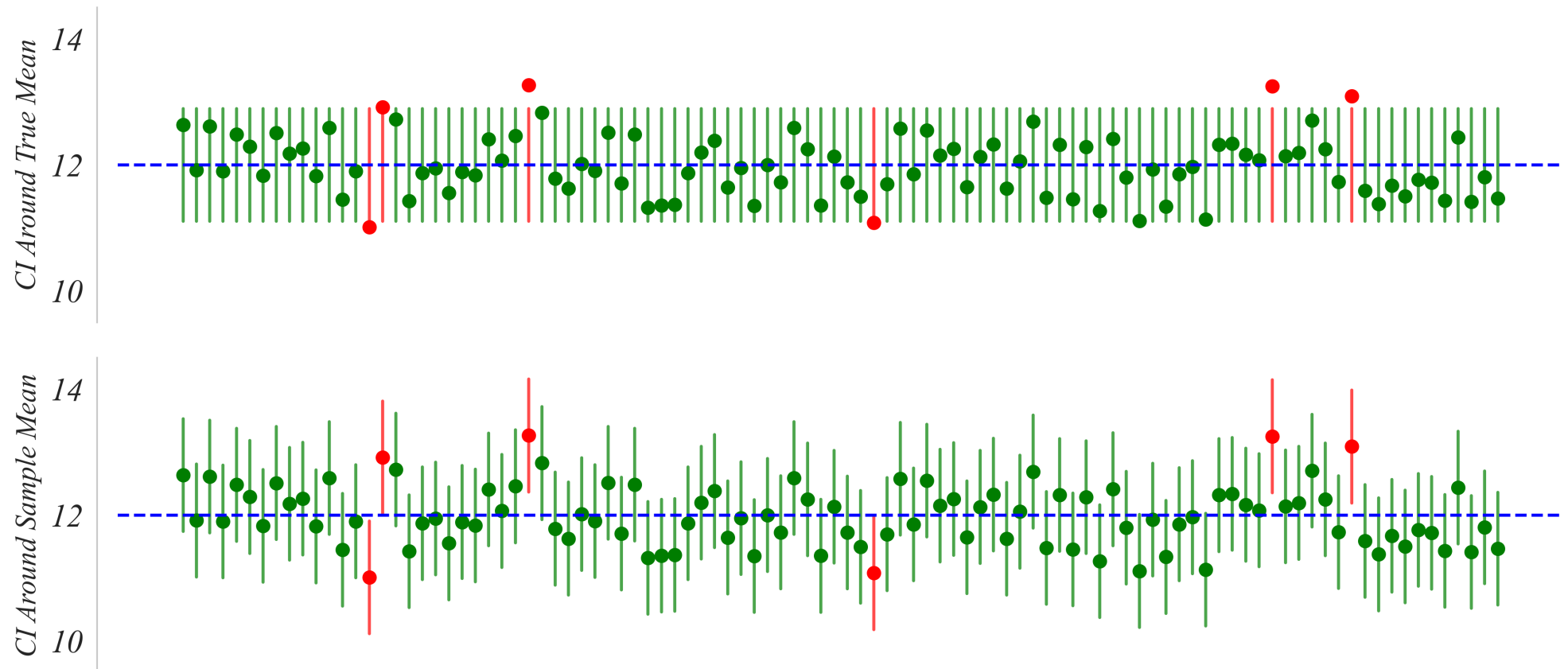


> the samples with \bar{x} in the CI around μ have μ in the CI around \bar{x}

Unknown μ : Two Perspectives

There are two mathematically equivalent perspectives to think about “closeness” between μ and \bar{x} .

I repeatedly sampled a distribution and constructed a 95% confidence interval.

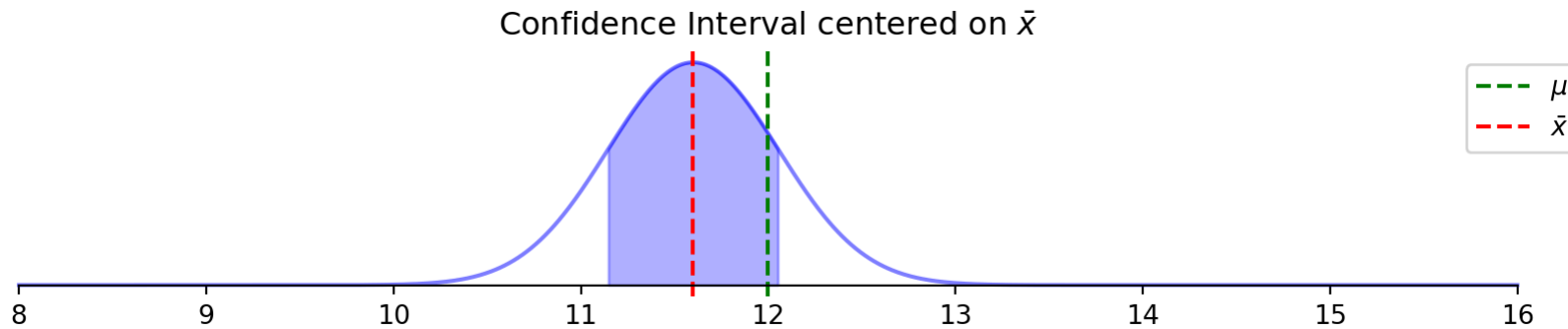


> *it is mathematically equivalent to check whether μ is in the CI around \bar{x} !*

Unknown μ : How ‘close’ is μ to \bar{x} ?

The distance between \bar{x} and μ works both ways.

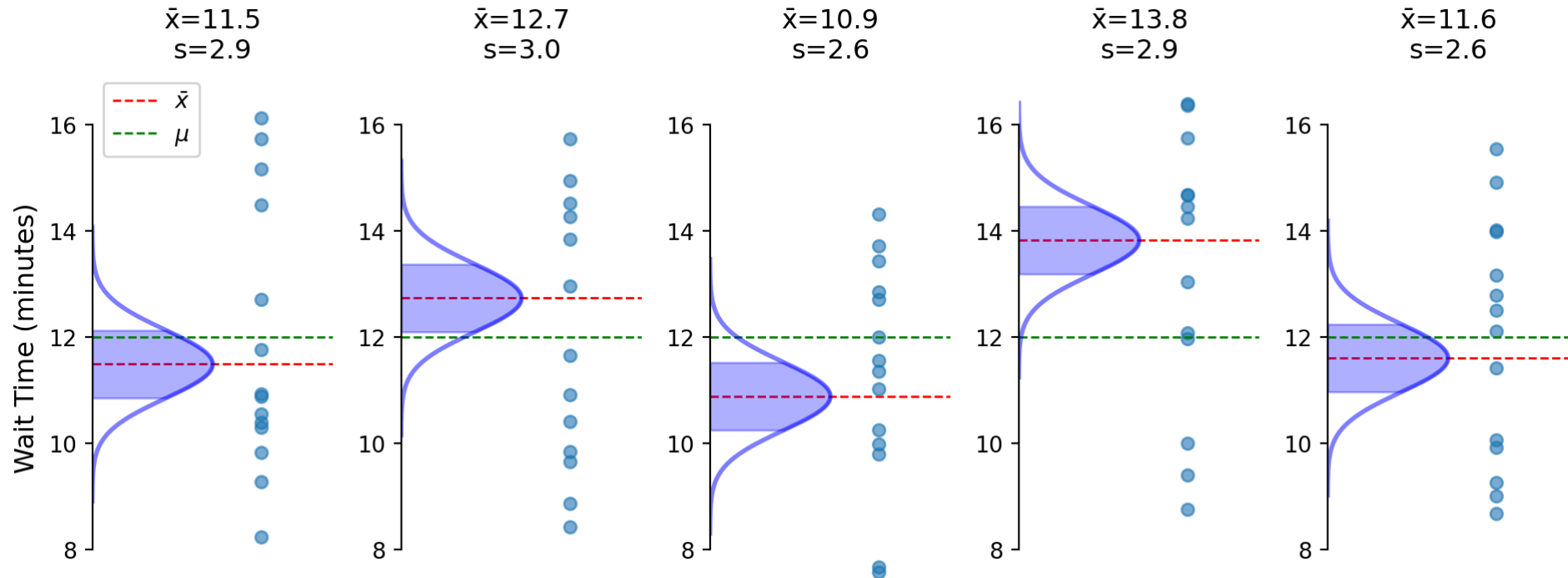
Now we can use the **Sampling Distribution** around \bar{x} to know the probability that μ is any distance from \bar{x} .



> *same distribution shape, just different reference points*

Unknown μ : How ‘close’ is μ to \bar{x} ?

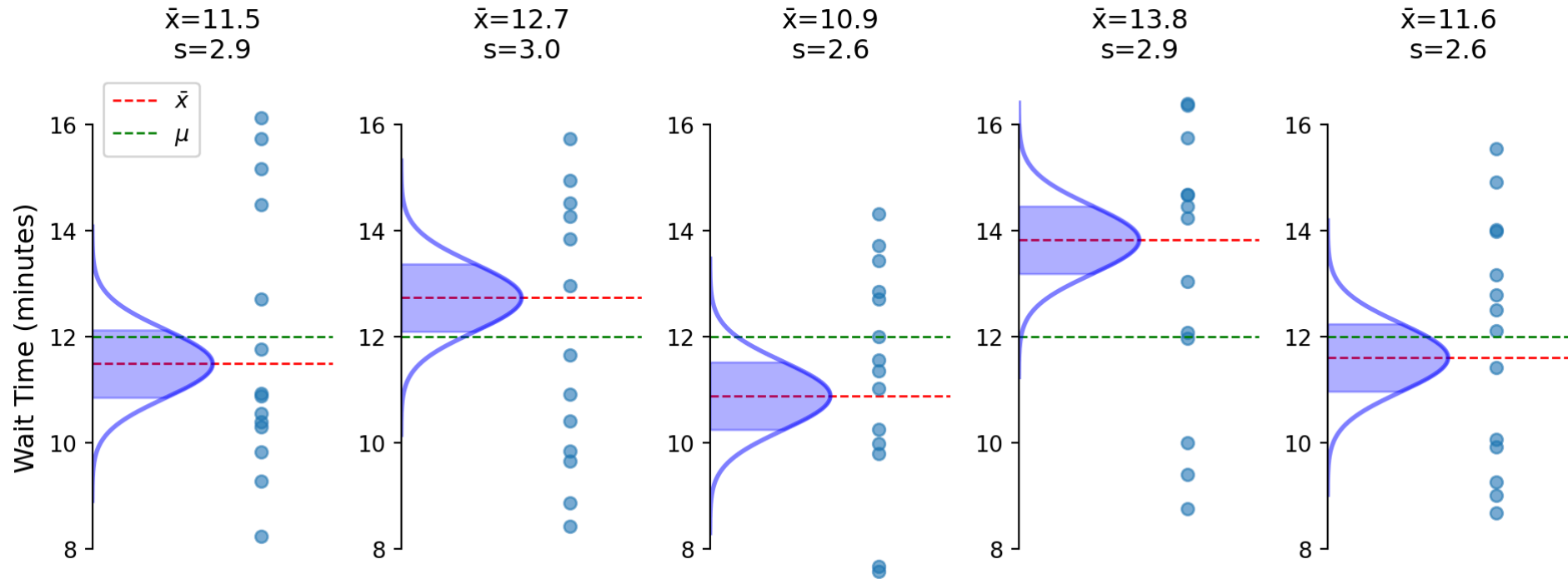
Each sample gives us a different \bar{x} and S .



- > notice both \bar{x} (red lines) and S vary across samples
- > each sample creates its own confidence interval for where μ could be
- > now we know the probability μ is in the CI around \bar{x} !

Unknown σ : How ‘close’ is μ to \bar{x} ?

Each sample gives us a different \bar{x} and S .

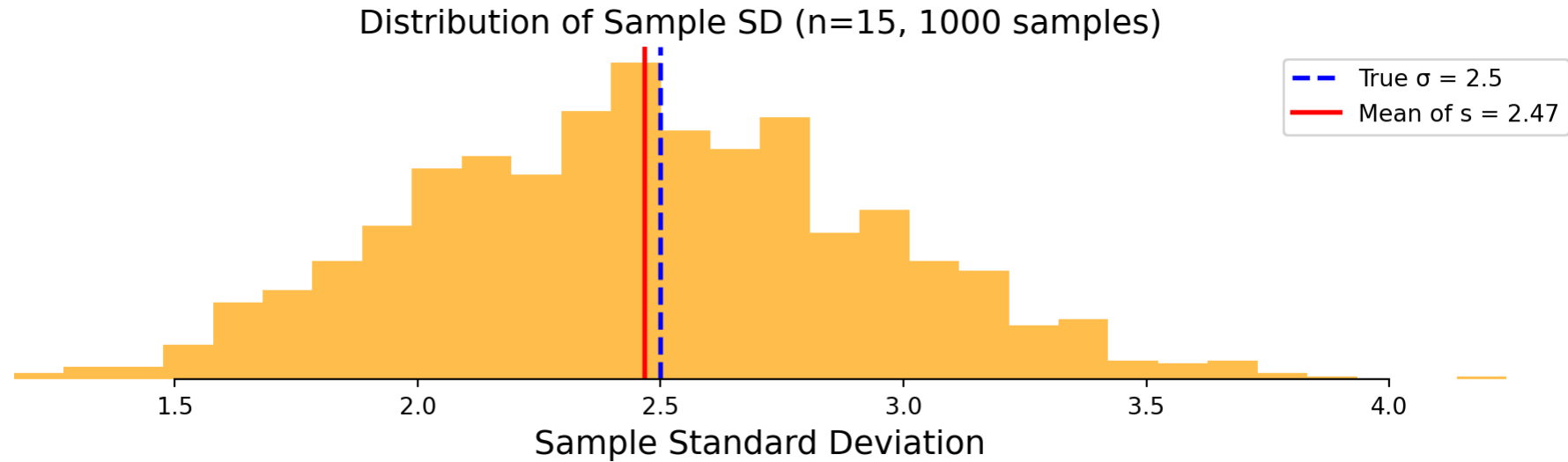


> *but here we're creating the Confidence Intervals using a known σ , which we will never actually observe*

> *each sample has a different S !*

Unknown σ : Variability of S

Just like \bar{x} varies around μ , the S varies around σ .

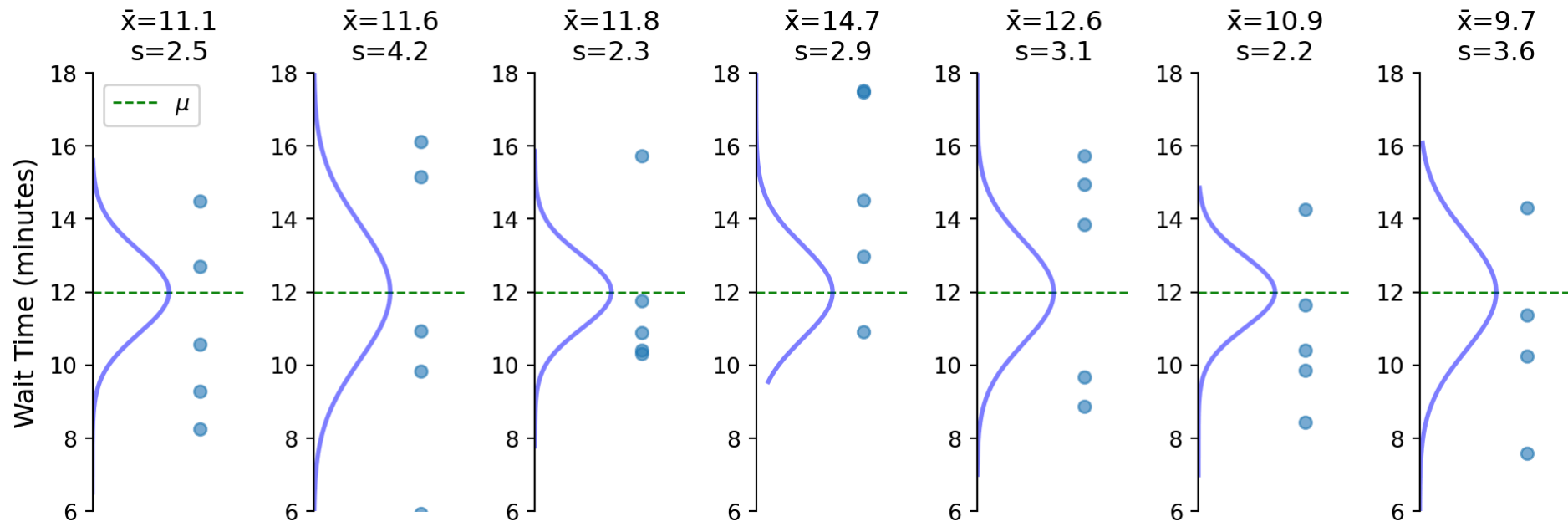


- > *we centered the Sampling Distribution on \bar{x} instead of μ*
- > *what would happen if we used the S in place of σ as a guess?*

Exercise 3.4 | Sampling Variation in S

Will a 90% confidence interval using S in place of σ correctly contain roughly 90% of the population means?

Samples ($n = 5$) with the sampling distribution centered on the population mean to show the differences in each samples' spread.



Exercise 3.4 | Sampling Variation in S

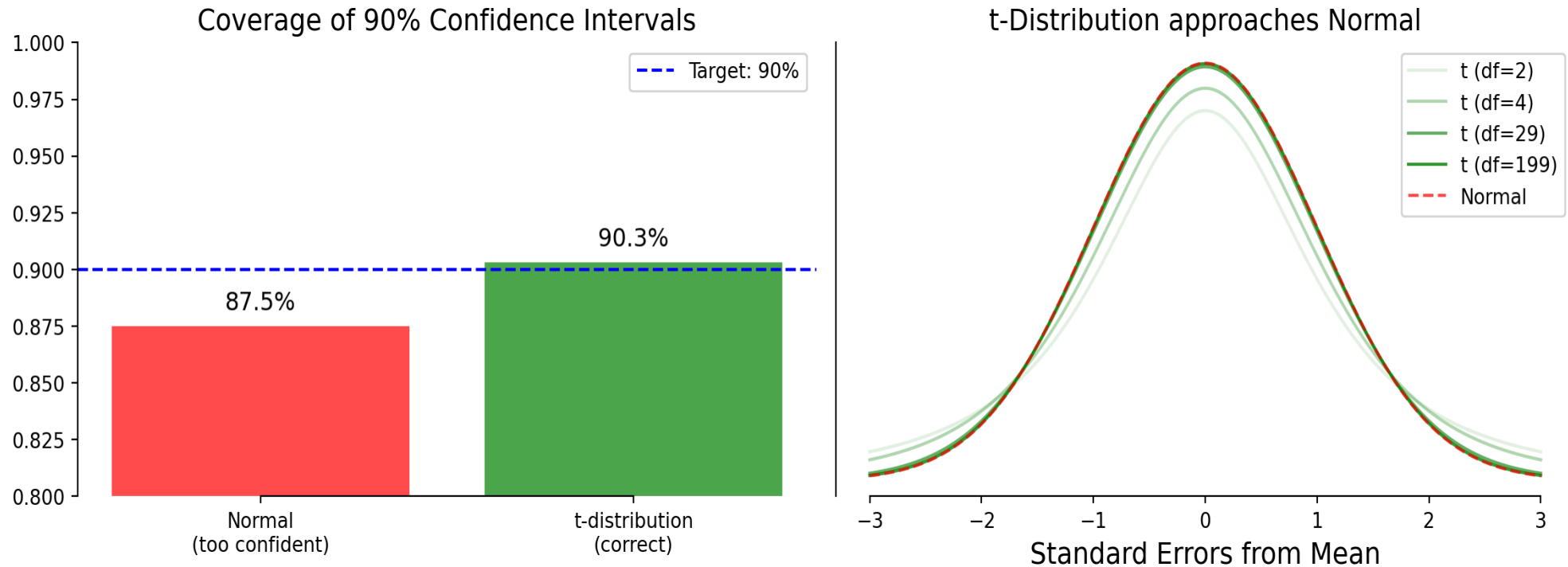
Will a 90% confidence interval using S in place of σ correctly contain roughly 90% of the population means?

Simulate many samples and check how often the 90% confidence interval contains the population mean when we simply swap S for σ .

> theres an additional layer of variability in the sampling distribution coming from the variability in the sample standard deviation (S)

Exercise 3.4 | Sampling Variation in S

Using the normal distribution with S gives wrong coverage rates ($n=15$).



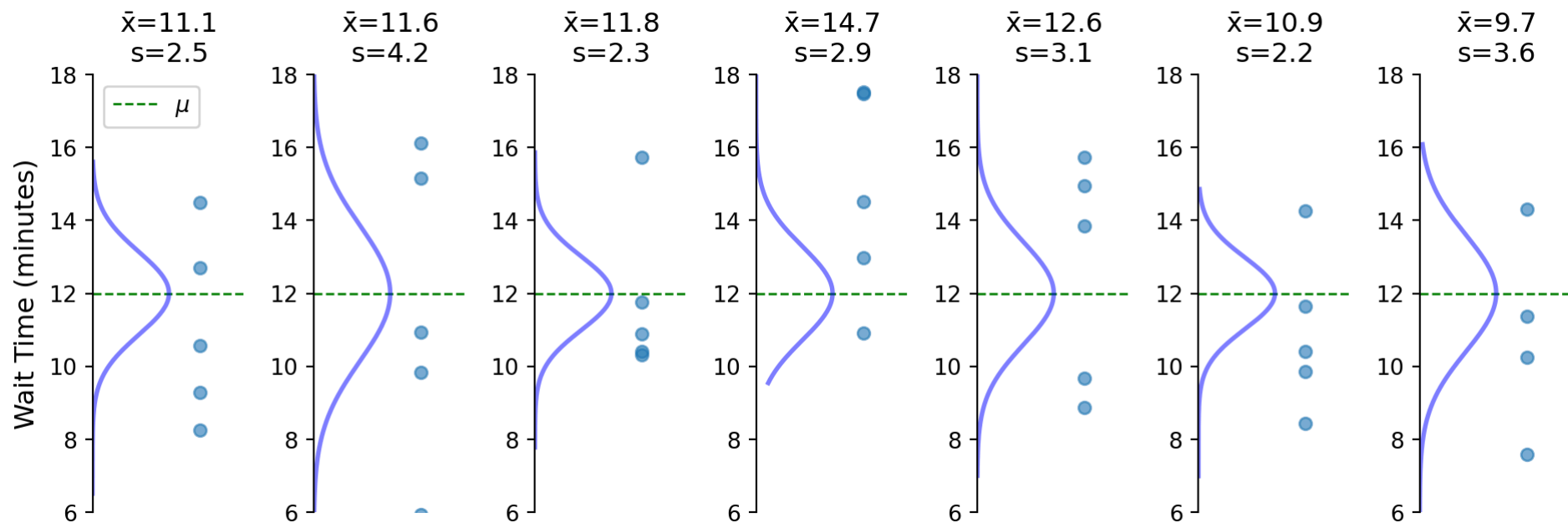
> *we would predict 90% when the actual number is lower (87.5%)*

> *we would be **too confident** if we use the Normal with S/\sqrt{n}*

Exercise 3.4 | Sampling Variation in S

Will a 90% confidence interval using S in place of σ correctly contain roughly 90% of the population means?

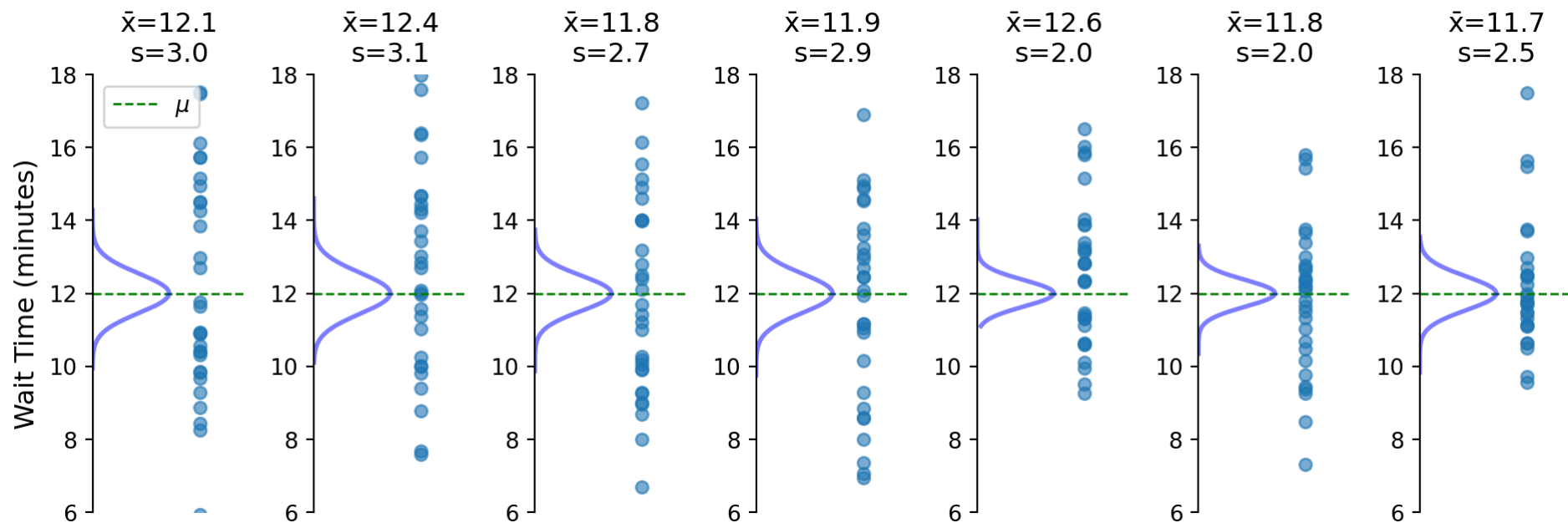
Samples ($n = 5$) with the sampling distribution centered on the population mean to show the differences in each samples' spread.



Exercise 3.4 | Sampling Variation in S

Will a 90% confidence interval using S in place of σ correctly contain roughly 90% of the population means?

Samples ($n = 30$) with the sampling distribution centered on the population mean to show the differences in each samples' spread.



> as the sample size grows (now $n=30$), this variability gets smaller

> but we'll always use a t -Distribution instead of a Normal for testing

Unknown μ and σ : Building Models

What if we want to test a specific claim about the unobserved population mean?

Is our data consistent with the following specific claim?

- *“The mean wait time is 10 minutes.”*

> instead of finding where some μ might be, we're testing a specific value of μ

Example: Wait Times

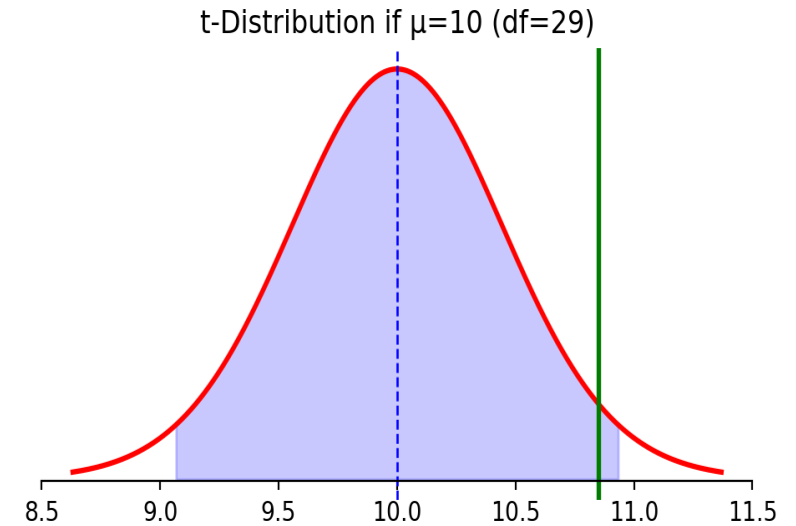
If $\bar{x} = 10.85$, is that consistent with $\mu_0 = 10$?

If sample standard deviation is $s = 2.5$:

$$SE = \frac{s}{\sqrt{n}}$$

$$SE = \frac{2.5}{\sqrt{30}}$$

$$SE = 0.456$$



```
1 s = 2.5
2 n = 30
3 se = s / np.sqrt(30)
```

Example: Wait Times

The math to answer this question is identical to confidence intervals.

If sample standard deviation is $s = 2.5$:

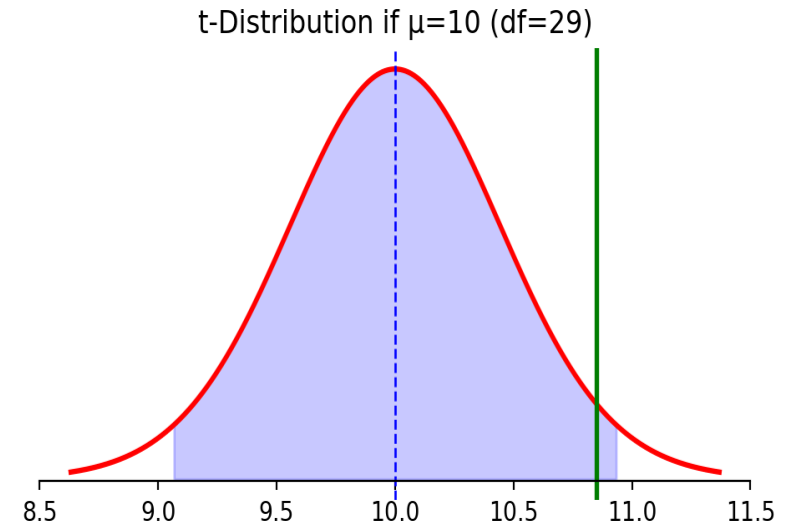
$$SE = 0.456$$

If true mean is $\mu_0 = 10$:

$$\bar{x} \sim t_{29}(10, 0.456)$$

So the critical value for 95%:

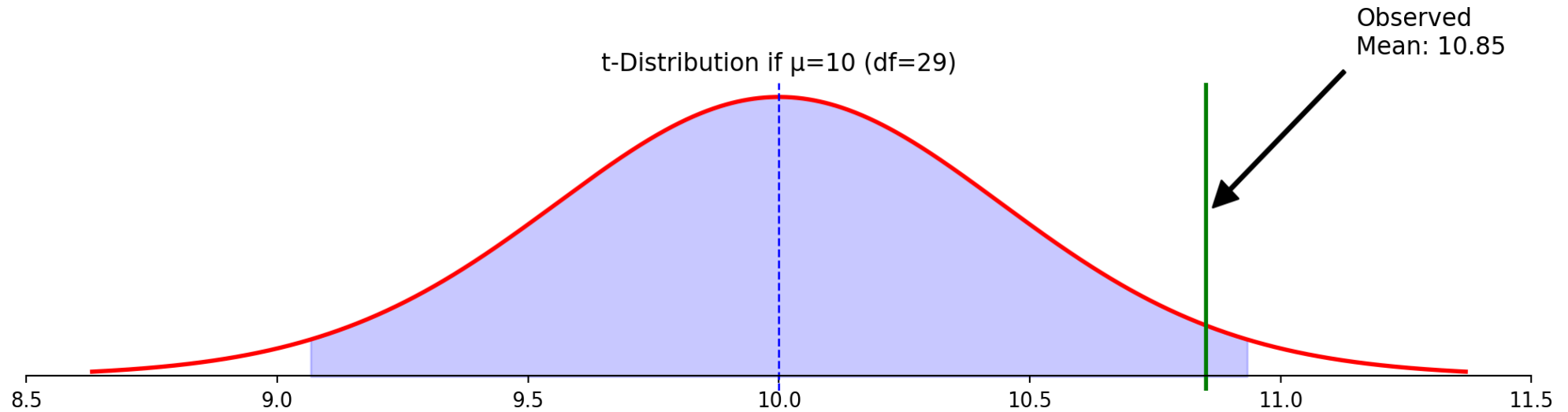
$$t_{crit} = 2.045$$



```
1 stats.t.interval(0.95, df=30)
```

Example: Wait Times

The math to answer this question is identical to confidence intervals.



A 95% confidence interval around μ_0 would be: [9.07, 10.93]

> *our observed mean ($\bar{x} = 10.85$) is within this interval — not surprising if $\mu=10$*

> *but if we observed $\bar{x} = 11.5$, that would be outside the interval — surprising!*

The Null Hypothesis

We formalize this approach by setting up a “null hypothesis”

Null Hypothesis (H_0): *The specific value or claim we’re testing*

- $H_0 : \mu = 10$ (*wait time is 10 minutes*)

Alternative Hypothesis (H_1 or H_a): *What we accept if we reject the null*

- $H_1 : \mu \neq 10$ (*wait time is not 10 minutes*)

Testing Approach:

- *Calculate how “surprising” our data would be if H_0 were true*
- *If sufficiently surprising, we reject H_0*

Quantifying Surprise: p-values

The p-value measures how compatible our data is with the null hypothesis.

p-value: *The probability of observing a test statistic at least as extreme as ours, if the null hypothesis were true*

For our example:

- *Null:* $\mu = 10$

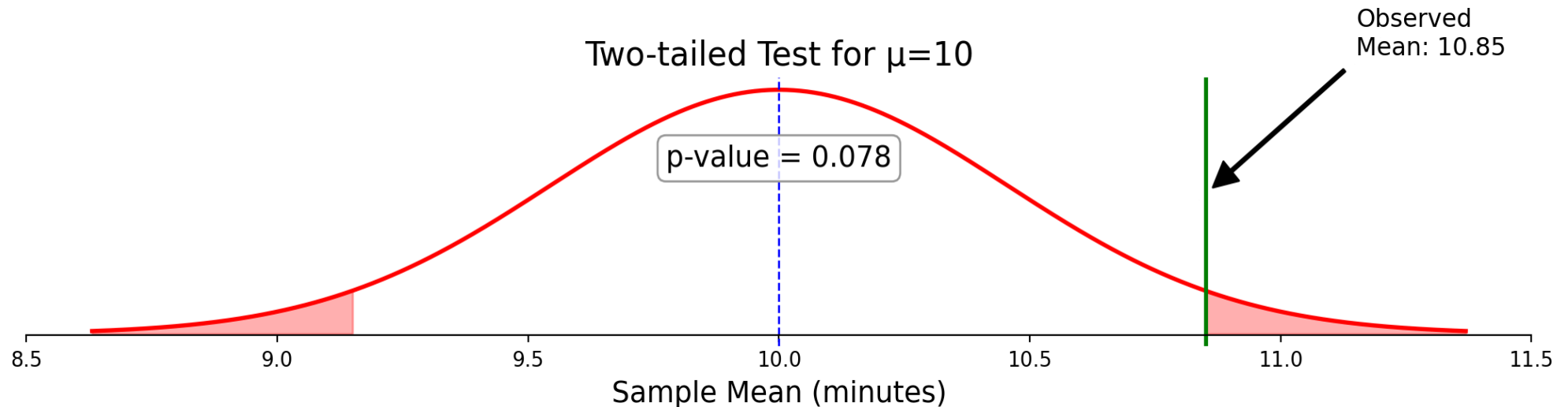
- *Observed:* $\bar{x} = 10.85$

> *How likely is it to get \bar{x} this far or farther from 10, if the true mean is 10?*

Quantifying Surprise: p-values

Example cont.: What is the probability of an error as large as the observed mean?

> *how likely is it to get \bar{x} this far or farther from 10, if the true mean is 10?*



```
1 stats.t.cdf((mu_0-xbar)/se, df=n-1)) * 2
```

> *interpretation: if $\mu=10$, we'd see \bar{x} this far from 10 about 7.2% of the time*

> *often, we reject H_0 if p-value < 0.05 (5%)*

> *here, p-value > 0.05, so we don't reject the claim that $\mu=10$*

Test Statistic: The t-statistic

We can standardize our result for easier interpretation

The **t-statistic** measures how many standard errors our sample mean is from the null value:

$$t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}}$$

Where:

- \bar{x} is our sample mean (10.85)
- μ_0 is our null value (10)
- s is our sample standard deviation (2.5)
- n is our sample size (30)

Test Statistic: The t-statistic

We can standardize our result for easier interpretation

The **t-statistic** measures how many standard errors our sample mean is from the null value:

$$t = \frac{\bar{x} - \mu_0}{s/\sqrt{n}} = \frac{10.85 - 10}{2.5/\sqrt{30}} = \frac{0.85}{0.456} = 1.86$$

Where:

- \bar{x} is our sample mean (10.85)
- μ_0 is our null value (10)
- s is our sample standard deviation (2.5)
- n is our sample size (30)

The t-test

This example has become a formal hypothesis test.

One-sample t-test:

- $H_0 : \mu = 10$
- $H_1 : \mu \neq 10$
- *Test statistic: $t = 1.86$*
- *Degrees of freedom: 29*
- *p-value: 0.072*

```
1 # Imports
2 import numpy as np
3 from scipy import stats
```

```
1 # Sample Data
2 sample_mu = 10.85
3 pop_mu = 10      # null hypothesis
4 std_dev = 2.5
5 n = 30
```

Decision rule:

- *If $p\text{-value} < 0.05$, reject H_0*
- *Otherwise, fail to reject H_0*

```
1 # Calculate t-statistic
2 t_stat = (sample_mu - pop_mu) / (std_dev / np.sqrt(n))
```

```
1 # Calculate p-value
2 p_value = 2 * (1 - stats.t.cdf(abs(t_stat), df=n-1))
```

> t-tests are a univariate version of regression

Statistical vs. Practical Significance

A caution about hypothesis testing

Statistical significance:

- *Formal rejection of the null hypothesis ($p < 0.05$)*
- *Only tells us if the effect is unlikely due to chance*

Practical significance:

- *Whether the effect size matters in the real world*
 - *A statistically significant result can still be tiny*
- > *with large samples, even tiny differences can be statistically significant*
- > *always consider the magnitude of the effect, not just the p-value*

Common Misinterpretations

What a p-value is NOT

✗ Not: The probability that H_0 is true

- *The p-value doesn't tell us if the null hypothesis is correct. It assumes the null is true and then calculates how surprising our result would be under that assumption.*
- *Example: A p-value of 0.04 doesn't mean there's a 4% chance the null hypothesis is true.*

Common Misinterpretations

What a p-value is NOT

✗ Not: The probability that the results occurred by chance

- *All results reflect some combination of real effects and random variation. The p-value doesn't separate these components.*
- *Example: A p-value of 0.04 doesn't mean there's a 4% chance our results are due to chance and 96% chance they're real.*

Common Misinterpretations

What a p-value is NOT

✗ Not: The probability that H_1 is true

- *The p-value doesn't directly address the alternative hypothesis or its likelihood.*
- *Example: A p-value of 0.04 doesn't mean there's a 96% chance the alternative hypothesis is true.*

Common Misinterpretations

What a p-value is NOT

✓ **Correct:** The probability of observing a test statistic at least as extreme as ours, if H_0 were true

- *It measures the compatibility between our data and the null hypothesis.*
- *Example: A p-value of 0.04 means: “If the null hypothesis were true, we’d see results this extreme or more extreme only about 4% of the time.”*

> *think of it like this: The p-value answers “How surprising is this data if the null hypothesis is true?” not “Is the null hypothesis true?”*

Looking Forward: Bivariate GLM

This framework extends directly to regression analysis.

Next time:

- *Bivariate GLM: Comparing means between two groups*
- > *the hypothesis testing framework is foundational for modern science*

Looking Forward: Regression

This framework extends directly to regression analysis.

Today's model: $E[y] = \beta_0$ (just an intercept)

Next: $E[y] = \beta_0 + \beta_1 x$ (intercept and slope)

- *Each β coefficient will have its own t -test*
- *Same framework: estimate $\pm t\text{-critical} \times SE$*
- *The t -distribution accounts for uncertainty in our estimates*

> regression is just an extension of what we learned today

Summary

We've built the foundation for statistical modeling.

- *Flipped perspective: center on what we observe (\bar{x}) not what's unknown (μ)*
- *Sample SD varies, creating need for t -distribution*
- *Built our first model: $E[y] = \beta_0$*
- *Tested hypotheses by shifting data*
- *Connected hypothesis tests to confidence intervals*

> these tools form the foundation of econometric analysis