

among SNPs in different genes can generate dispersed association signal across nearby genes, which is often misinterpreted as evidence for association of specific genes in this set. The same phenomena is relevant for PrediXcan and TWAS approaches [79], arguably amplified due to the methods used to develop the predictive model. Notably, when TWAS methods were contrasted to variant-level colocalization [81], TWAS almost always captures variant-level colocalization; however, in real data it is 100-fold more powerful. We conjecture this power traces, in part, to LD among SNPs across the GWAS locus generating false signals. We propose methods to overcome this challenge, adapting penalized regression approaches from the fairness literature.

B.2a Improving polygenic scores (PGS). PGS, also known as PRS or PGRS (polygenic risk score) when predicting disorder status, predicts phenotypes of individuals by a weighted sum of counts of their associated alleles [82]. The associated allele for each SNP and its weight are estimated from GWAS data that are independent of the samples for which PGS is to be predicted. Many methods have been proposed to construct a PGS [83–89], differing largely by which SNPs to include and how to weight them. Ni et al. [90] review and compare 10 methods popular for psychiatric genetics. Most methods outperform Pruning & Thresholding, the original PGS, and there are modest differences among most methods, with SBayesR [87] showing best overall performance in the authors hands (but see [85]). Of the methods requiring tuning, such as LDpred2 and Lassosum [86] (Lasso with summary statistics), Ni finds them to be sensitive to the tuning population. In Aim 2, we propose methods to reduce this sensitivity.

Most PGS methods use only GWAS and LD information for the score. However, as we noted earlier, functional annotations show strong enrichment for GWAS signals. For this reason, methods that select variation or weight it according to functional annotations should produce a refined score. Indeed, Marquez-Luna et al. [85] show that accounting for function, via LDpred-funct, produces a 4% relative improvement over SBayesR when fitting PGS to UK Biobank phenotypes. We propose to build on this observation. As we noted in Aim 1a, we can model annotations using gradient boosting trees to determine which annotations – and interactions of annotations – are more likely to predict causal variation and whether this varies across the myriad phenotypes evaluated by GWAS. Modeling in LDpred-funct assumes phenotypes matter, we predict that annotations effects are similar across phenotypes and will build on this conjecture if it is supported. By bringing better annotation information in PGS, we will build a more predictive PGS. In the process, we will also optimize other aspects of PGS.

B.2b Portable PGS. A PGS developed from samples of European ancestry (PGS_E) predicts the same phenotype in non-European samples poorly [91,92]. This deficit arises from a multiplicity of causes: (1) many SNPs included in the PGS_E do not have a direct effect on risk, rather they are in LD with causal variants and LD patterns vary across populations; (2) causal variants have somewhat different effects across populations; (3) the genetic architecture of the trait differs somewhat among populations. If (1) causes the greatest lack-of-portability, as we expect, then the solution lies in fixing the PGS_E , as proposed in Aim 2.2a, and is supported by recent studies [46, 93, 94]. A simple solution for (2) is to combine non-European PRS (PRS_N) with a PRS_E in an optimal way. Marquez-Luna et al. [95] develop such an approach, estimating optimal weights for a trans-ethnic pruning and thresholding PGS by training on both populations. Here we propose a solution to (2), building on Lassosum and taking two approaches, one in the spirit of Marquez-Luna et al. and another joint optimization of Lassosum over populations. Because some populations are of admixed ancestry, we will extend these ideas using local ancestry segments [96–98] to obtain portable PGS for a variety of population histories. We will introduce measures and concepts of fairness, a critical feature of PGS non-portability [99], to the human genetic literature. Different genetic architectures (3) creates a difficult hurdle, especially if the populations differ substantially by environmental effects. From a purely genetic perspective, solutions for (1) and (2) are still the best approach for a portable PGS, while incorporating non-genetic terms would be invaluable for prediction.

B.3 Models for rare non-coding variants. To relate clinical outcome to rare coding or copy number variants (CNVs), clinical geneticists typically used a Fisher Exact Test. Recognizing that the FET sacrifices power because it ignores critical information about genes/genomic regions, their mutation rates, and the properties of the mutations, in 2011 we built a more powerful model for CNV association [100]. We then extended the idea to variation in coding sequence [101, 102], and next introduced the Transmission And De novo Association framework [103], which modeled sequence variation of three inheritance types: de novo, inherited, or from case-control samples. The current version of TADA integrates all classes of coding-region variation, including CNVs [30]. Using TADA, we have identified hundreds of genes associated with autism and other neurodevelopmental disorders [20,30,104] and have also illuminated its neurobiology [30,105,106]. Key to TADA's power is prior information about the nature of genes (e.g., level of conservation) and genetic variants (e.g., protein truncating (PTV), conserved missense, silent), as well as gene-level mutation rates. The prior information is critical and revealing. For instance, of sequence variation, de novo PTVs found in conserved genes have the largest effect on risk; yet, PTVs transmitted from unaffected parents to their affected offspring confer far less risk, even when they fall in the

CommonMind and PsychEncode Consortia [21, 51, 52, 74, 148], which have genetic and gene expression data from brain tissue. (3) A new resource (R01MH125235) – partially overlapping with CMC resources – but also producing proteomics data (levels of protein expression, phosphorylation, glycosylation, and ubiquitination) on 400 brain samples. These samples also are, or will be, characterized for gene expression and whole genome sequence (please see Matt MacDonald’s letter of support). (4) BrainVar, a PsychEncode project that will characterize ≈ 100 human brains, performing temporal, regional, and cell-type-specific transcriptome profiling (snRNA-seq) of the developing human brain. The DNA samples will also be characterized for ATAC-seq peaks and whole genome sequence (please see Stephan Sander’s letter of support). And (5), other relevant, new PsychEncode projects. Of these, (2) involves Devlin as an mPI (with Matt MacDonald and Jon Trinidad), thus the data are freely available for validation of methods proposed here and for collaborative research between our groups. (Remark: While colocalization is a goal of R01MH125235, only Ensemble approaches, which combine existing methods, are proposed therein. Thus, there is no methodological overlap with the research proposed here, although we do draw inspiration from it.) For (4) both Devlin and Roeder are involved as collaborators for the BrainVar study (Stephan Sanders and Nenad Sestan, mPIs). Hence the data will be available for validation of methods. (Again, there is no overlap regarding methods development of the BrainVar project and what we propose here.)

AIM 2: Improving polygenic scores and adapting them for diverse ancestries

Background The objective of PGS is to predict a phenotype by a linear model of SNP genotypes. If we knew all of the causal SNPs in a population, the problem would simplify to weighting the SNPs properly. Due to LD and low power, however, we don’t know the causal SNPs. Moreover, even in a GWAS locus, LD among SNPs obscures the causal variant(s), this correlation tends to be blocky, and the size of these blocks varies by ancestry. For these reasons, PGS tend to include many “linked or tag” SNPs, rather than causal variants. And, for this reason, predictors built from Eurocentric samples do not successfully port to other ancestries. To improve a PGS, and make it more portable, we should include functional SNPs rather than tag SNPs. Several efforts in the literature have had some success in this domain [85], but there is room for further improvement. Furthermore, we recognize that some populations are of admixed ancestry. For this reason, we will extend the ideas outlined below to account for it by using local ancestry segments [96–98].

Aim 2a: Building a more predictive PGS There are several methods in the literature for injecting annotations into PGS. Our primary contribution will be in assimilating the information in a powerful and interpretable form. We will investigate available methods for how best to utilize this information in a PGS form. We will apply our more elaborate modeling of functionality, as described in Aim 1a, to identify SNPs that are more likely functional and thereby improve the PGS. However, Aim 1a only targets GWAS loci, whereas a good PGS typically includes a wider set of SNPs. In [7] we lay out an approach using AdaPT for selective inference and XGBoost for model fitting and by which we identify likely causal SNPs by virtue of their annotation and GWAS p-values – even when the p-values are far from GWAS-significant (i.e., $> 5 \times 10^{-8}$). We will explore this approach for selecting SNPs for the PGS. Furthermore, as noted in Aim 1, BSLMM already provides a relatively sparse solution and itself produces an improved PGS [139]. We will seek to hybridize the features of BSLMM and the AdaPT/XGBoost approach to introduce an informative prior and hence produce a refined PGS. (Remark: BSLMM is closely related to SuSiE and therefore we believe this is possible.) Finally, we will explore CVC as a possible replacement for BSLMM and compare results.

Aim 2b: Building a more portable PGS. Here we seek a PGS with maximum portability across ancestries. Typically we have available large training data sets for the majority population, usually Euro-centric (EC), and only modest sized samples for the other population, not EC (nEC). Signals are weak, so large amounts of data are essential for success in the training stage, and yet it will take time to collect such samples from populations whose ancestries are nEC. Ideally we could learn from all the data, but avoid the biases inherent in using primarily the EC sample.

To formulate and validate a PGS model, ideally a large GWAS is available for fitting the model (training), a smaller independent dataset is available from which to choose the tuning parameters (testing) and finally an independent dataset can be utilized for validation. However, commonly no testing data are available, especially for the nEC population. Moreover, for many phenotypes, only summary statistics are available, precluding the option of using cross-validation on the training data. These practical issues create statistical challenges.

We focus on a data structure with $Y \in \{0, 1\}^n$ indicating case-control status, $X \in \mathbb{R}^{n \times (p+1)}$ an intercept and centered variables indicating genotype for p SNPs. (Remark: Methods pertain to continuous phenotypes also.) We want to learn a linear model that predicts Y from X using the lasso estimator, which can be expressed as $\hat{\beta}_\lambda = \arg \min 2R^T \beta + \beta^T C \beta + \lambda \|\beta\|_1$ where $R = \frac{1}{n} X^T Y$ and $C = \frac{1}{n} X^T X$ are the empirical correlation between X and Y , and the empirical covariance of X (i.e., LD), respectively. The Lassosum estimator [86], replaces C with a regularized version of the same obtained from a library such as 1000 genomes. The performance of the

lasso estimator is sensitive to the choice of λ [90], hence reliable methods for selection of the tuning parameter are needed.

When sampling from two ancestries, we anticipate that the signal, β , is similar across ancestries, but that the LD pattern varies. Our proposed method, which we call Joint-Lasso, considers a linear combination of loss functions from the two ancestries, which simplifies to a convenient form. Let $C = \gamma C_1 + (1 - \gamma)C_2$ and $R = \gamma R_1 + (1 - \gamma)R_2$, where C_j, R_j are the sample correlations from population $j \in \{1, 2\}$ and substitute these into the lasso formulation. In this setting, we ultimately need to choose three tuning parameters: the mixing parameter γ and population dependent λ 's to obtain a small risk within each ancestry group. It is challenging to choose the tuning parameters because we only have available summary statistics from a single source. This precludes traditional approaches to cross-validation for training and testing. We focus discussion on λ , but the same ideas apply to choice of γ . We note that all results below can be simplified for the one ancestry problem.

We explored performance of Joint-Lasso and tuning parameters via simulations. Summary statistics were generated for a training and testing sample. For the training data, we using EC and African 1000 Genomes data to generate genotypes with realistic LD structure, then selected p SNPs as causal variants to achieve a certain total h_T^2 , which could be the same or different between populations. Causal variants contributed equally to h_T^2 , but the effects varied as a function of population-based allele frequency. Sample sizes varied, but usually were much larger for the EC sample. Tests of association (GWAS) were then performed to yield summary statistics. We found that Joint-Lasso produces a PGS that performs well for both ancestries and is superior to existing methods targeting portability. Importantly, if we used extreme PGS values from lassosum as a simple classification rule, false positives were far more likely for African than for EC ancestry. For Joint-lasso and the same classification rule, however, the false classification rate was substantially reduced. In addition, from these simulations we obtained preliminary evidence that the optimal choice of λ varies by population, largely due to the information content for the PGS, but that γ is a function of the relative sample sizes and to a far lesser extent, heritability. We will develop these ideas further.

Selecting tuning parameters based on summary statistics The tuning of λ aims at minimizing each predictive risk: $\text{Risk}_j(\hat{\beta}_\lambda) = \mathbb{E}[(Y_{\text{new}}^{(j)} - (X^{(j)})^T \hat{\beta}_\lambda)^2 | \hat{\beta}_\lambda]$ where $Y_{\text{new}}^{(j)}$ is an independent draw from the j th population. We assume fixed $X^{(j)}$ and the expectation is over $Y^{(j)}$ and $Y_{\text{new}}^{(j)}$. Based on a result of Efron [149], the predictive risk can be related to the in-sample predictive risk as follows: $\text{Risk}_j = \|Y^{(j)} - X^{(j)} \hat{\beta}_\lambda\|_2^2 + 2 \sum_{i=1}^{n_j} \text{Cov}(\hat{\mu}_i^{(j)}, Y_i^{(j)})$, where $\hat{\mu}_i^{(j)} = (X_i^{(j)})^T \hat{\beta}$ is the fitted mean value for $\mathbb{E}(Y_i^{(j)} | X_i^{(j)})$. The term $\text{Cov}(\hat{\mu}_i, y_i)$ is called the “optimism” (also known as “covariance penalty”), which quantifies the difference between the actual predictive risk and the in-sample predictive risk. (It is easy to check that the in-sample predictive risk can be computed using only the summary statistics.)

Let $(Y_1^{(1)*}, \dots, Y_{n_1}^{(1)*}), (Y_1^{(2)*}, \dots, Y_{n_2}^{(2)*})$ be a bootstrap sample, then we can approximate the optimism term by $\widehat{\text{Cov}}(\hat{\mu}_i^{(j)}, Y_i^{(j)}) = \text{Cov}_*(\hat{\mu}_i^{(j)*}, y_i^{(j)*})$, where $\mu_i^{(j)*} = (X_i^{(j)})^T \hat{\beta}^*$, and $\hat{\beta}^*$ is the bootstrap version of $\hat{\beta}$ using the bootstrap sample. The bootstrap sample is generated by $y_i^{(j)*} = \hat{\mu}_{i,0}^{(j)} + \epsilon_i^{(j)*}$ where $\epsilon_i^{(j)*}$ is centered Bernoulli noise such that $\mathbb{E}_*(\epsilon_i^{(j)*}) = 0$ and $\mathbb{E}_*(\epsilon_i^{(j)*})^2 = \hat{\mu}_{i,0}^{(j)}(1 - \hat{\mu}_{i,0}^{(j)})$. Here $\hat{\mu}_0$ is a “preliminary estimate”, which is expected to be fairly accurate although not optimal. Such an estimate can usually be obtained using a relatively small λ . The challenge is to approximate the bootstrap procedure when only summary statistics are available.

Now $\sum_{i=1}^{n_j} \text{Cov}_*(\hat{\mu}_i^{(j)*}, Y_i^{(j)*}) = \mathbb{E}_*(\hat{\beta}^*)^T (X^{(j)})^T \epsilon^{(j)*}$. And to obtain $\hat{\beta}^*$, we will need the bootstrap versions of R_1 and R_2 , where $R_j^* = (X^{(j)})^T Y^{(j)*} = (X^{(j)})^T (X^{(j)} \hat{\mu}_0 + \epsilon^{(j)*}) = n_j C_j \hat{\mu}_0 + (X^{(j)})^T \epsilon^{(j)*}$. Therefore, to obtain the optimism estimate, we only need to generate $(X^{(j)})^T \epsilon^{(j)*}$ for $j = 1, 2$, but X is not recorded. To achieve our goal, let $X^{(j)} = U_j D_j V_j^T$ be the singular value decomposition of $X^{(j)}$. If we only have the summary statistic C_j , then U_j is not accessible but (V_j, D_j) is. Because U_j is full rank orthonormal, we only need to generate $\tilde{\epsilon}^*$, where $\tilde{\epsilon}^* = U^T \epsilon^*$. The first and second moments of $\tilde{\epsilon}^{(j)*}$ are: $\mathbb{E}(\tilde{\epsilon}^{(j)*}) = 0$ and $\mathbb{E} \tilde{\epsilon}^{(j)*} (\tilde{\epsilon}^{(j)*})^T = \text{diag}(\hat{\mu}_{1,0}^{(j)} - (\hat{\mu}_{1,0}^{(j)})^2, \dots, \hat{\mu}_{n,0}^{(j)} - (\hat{\mu}_{n,0}^{(j)})^2)$. We can use the following a “partially-second-order” Gaussian approximation $\tilde{\epsilon}^{(j)*} \sim N(0, \tau_j^2 I_{n_j})$ where $\tau_j^2 = \hat{\beta}_{1,0} + \hat{\beta}_0^T C_j \hat{\beta}_0$. Here $\hat{\beta}_{1,0}$ is the first coordinate (i.e., intercept) of the preliminary estimate $\hat{\beta}_0$, and we used the fact that the other columns in $X^{(j)}$ are centered and hence sum to 0.

Finally, given summary statistics R_j, C_j for $j = 1, 2$, corresponding sample sizes n_j , preliminary estimate $\hat{\beta}_0$ and bootstrap sample size B , we implement the following procedure: for $b = 1, \dots, B$, $j = 1, 2$, generate $\tilde{\epsilon}^{(j)*}$; emulate $(X^{(j)})^T \epsilon^{(j)*}$ by $V_j D_j \tilde{\epsilon}^{(j)*}$ where $V_j D_j^2 V_j^T$ is the SVD of $n C_j$; and compute $\hat{\beta}^*$ using (R_1^*, C_1, R_2^*, C_2) . The covariance penalty is approximated by the average of $(\hat{\beta}^*)^T (X^{(j)})^T \epsilon^{(j)*}$ over the bootstrap repetitions.

A and B, there are three possible intersections: $A = 1$ or $B = 1$; $A = 1 \& B = 1$. The considerable overlap between annotation intersections leads to test statistics with substantial levels of correlation, which confounds the interpretability and error rate guarantees of multiple testing procedures. To overcome this challenge we use an agglomerative procedure that was implemented in the context of gene-level testing [11] to cluster highly correlated intersections together for testing. Ultimately our tests can discover clusters of significant annotation intersections. We conjecture these results can be interpreted downstream using new data blurring techniques to enable us to estimate valid post-selection confidence intervals for the effect sizes of the CWAS annotation intersections [150].

Using realistic simulations, our preliminary results show that the ideas described above will work in practice, but we still lack power. To enhance performance we will implement selective inference approaches [191], which we have successfully implemented in the context of SNP and gene-based tests [7, 11]. In this setting, a new twist is needed because the side-information of annotations also defines the categories. Luckily, the correlation structure of intersections can be exploited, allowing us to pool shared information across their intersections as metadata for guiding our multiple testing correction. Specifically, we leverage the overlapping annotation metadata using XGBoost in the AdaPT framework to up-weight hypotheses that are more likely to be non-null. (Due to the discreteness of performing Poisson enrichment tests, we rely on innovations in selective inference masking functions [192, 193] to ensure we retain power in this setting.) Preliminary investigations with simulated data indicate that pooling annotation information in this manner can lead to substantially improved power [194].

Challenges *(1) To evaluating meta-learning, we can exploit the abundance of data and knowledge of variation in the exome. We know MPC scores can roughly separate de novo missense variants into benign and risk variants; can meta-learning match or exceed this success by the evaluation of transmitted missense variation, then applying what we learned to de novo missense variation? (2) We describe our meta-learning approach for parent-offspring trios, from whom we can infer transmitted and de novo variants. Can this type of analysis be modified to incorporate large data sets with standing variation, such as TOPMed, by using a mutational model to predict which annotations and intersections of annotations show a dearth of variation (akin to derivation of LOEUF)? (3) The objectives of meta-learning and selective inference have overlapping goals. We will investigate whether the model learned with meta-learning can bolster the power of CWAS and selective inference by refining and potentially reducing the number of tests performed. (4) We will evaluate whether and how our findings from meta-learning can be used to improve various LR methods for testing non-coding variants. (5) We will pair our meta-learning results with the TADA-A [129] algorithm.*

Data for Aim 3. *Resources for WGS data are expanding and we expect substantial datasets in the near future. For instance, Alzheimer's Disease Sequencing Project [195] recently released WGS data from 16,906 samples, available through NIAGADS, and they anticipate 20,000 additional samples soon (please see Li-San Wang's letter of support). TOPMed Freeze 9 has WGS data from 206,000 individuals, with samples from CCDG, 158,470 TOPMed samples and 2,504 1000 Genomes samples. For mental disorders, Devlin and Roeder are mPIs of the ASC-SSC Whole Genome Consortium (along with Mike Talkowski, Stephan Sanders, and Joe Dougherty; (please see Mike Talkowski's and Stephan Sanders' letters of support)), formed for the analysis of WGS data from ASD families. We anticipate aggregating WGS data from 8,626 ASD probands within 8,189 families this year. In terms of other WGS data for mental health, the Whole Genome Sequencing for Psychiatric Disorders Consortium is also aggregating data. As they note in their publication, "The Whole Genome Sequencing for Psychiatric Disorders Consortium will integrate data for 18,000 individuals with psychiatric disorders, beginning with autism spectrum disorder, schizophrenia, bipolar disorder, and major depressive disorder, along with over 150,000 controls." [196] A data set of special relevance to ASD is congenital heart disease (CHD). Like ASD, CHD is uncommon, occurring in roughly 1% of births. Rare genetic variants, especially de novo variants that damage genes, occur at an elevated rate in CHD individuals. Furthermore, CHD is associated with neurodevelopmental disorders and must be under negative selection. Notably, WGS data are currently available from 763 CHD probands [197] and their families and data from 1812 should be available soon [198]. These data should be an excellent training set from which to learn about important non-coding variation; in turn, one could say the same for the ASD data, it would make an excellent training set of CHD. Clearly both would benefit from the methods we propose here.*

Timeline *We do not anticipate working on the Aims in the order they appear, we will be working on them simultaneously. All are timely and the field is competitive. We expect releasing manuscripts from all three Aims during the first year of the project, with follow-up and additional studies thereafter. During Year 4 we anticipate working on new topics, as well as completing the Aims described herein, in anticipation of a renewal application.*

Sex and other relevant biological variables *Our methods apply regardless of sex.*

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