

Another methodology to find an estimator:-

### Method of Moment Estimation (MME) :-

$X_1, \dots, X_n$   $\stackrel{i.i.d.}{\sim} f(x; \theta_1, \dots, \theta_k)$ .  $\theta = (\theta_1, \dots, \theta_k)$  is a k-dim unknown parameter.

Let us denote

$$m_r' = \frac{1}{n} \sum_{i=1}^n x_i^{(r)} \rightarrow r\text{-th raw moment.}$$

We know, by WLLN,  $\frac{1}{n} \sum_{i=1}^n x_i^{(r)} \xrightarrow{P} E(x_i^{(r)})$ .

The MME ~~est~~ in the sol<sup>n</sup>. of

$$m_r' = m_r'(\theta_1, \dots, \theta_k), r=1, \dots, k.$$

Solve  $\textcircled{*}$ , will have  $\hat{\theta}_{1, \text{MME}}, \dots, \hat{\theta}_{k, \text{MME}}$

Examples!—

1.  $x_1, \dots, x_n \sim \text{Bin}(1, \theta)$ ,  $\theta \in (0, 1)$ .

$$u'_1 = E(x) = \theta$$

$$m'_1 = \frac{1}{n} \sum_{i=1}^n x_i$$

Hence, the MME of  $\theta$  is the root of

$$m'_1 = u'_1 \Rightarrow \frac{1}{n} \sum_{i=1}^n x_i = \theta$$

Hence, the MME of  $\theta$  is  $\frac{1}{n} \sum_{i=1}^n x_i$ .

2.  $x_1, \dots, x_n \stackrel{i.i.d}{\sim} \text{Unif}(-\theta, \theta)$ .

$$E(x) = \mu_1' = \frac{1}{2\theta} \int_{-\theta}^{\theta} x dx = 0$$

$$E(x^2) = \mu_2' = \frac{1}{2\theta} \int_{-\theta}^{\theta} x^2 dx = \frac{\theta^2}{3}$$

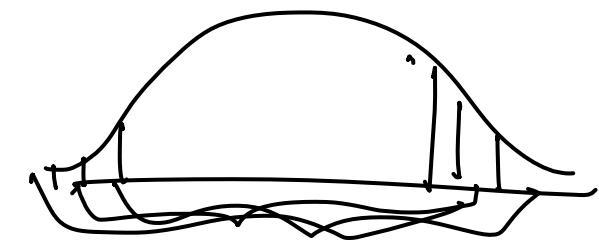
The sol<sup>n</sup> of

Hence,  $\uparrow m_2' = \frac{1}{n} \sum_{i=1}^n x_i^2 = \mu_2' = \frac{\theta^2}{3}$  is the MME

of  $\theta$ . Hence, the MME of  $\theta$  is  $\boxed{\frac{3}{n} \sum_{i=1}^n x_i^2}$

$$\hat{\theta}_{n, \text{MME}} = \sqrt{\frac{3}{n} \sum_{i=1}^n x_i^2}$$

## Interval estimation:-



$x_1, \dots, x_n \stackrel{\text{i.i.d}}{\sim} f(x; \theta), \theta \in \Theta$

Want to find a random interval

$[L(x_1, \dots, x_n), U(x_1, \dots, x_n)]$  of the unknown parameter

$\theta$  such that

$$P[L(x_1, \dots, x_n) < \theta < U(x_1, \dots, x_n)] \geq \underline{1-\alpha}.$$

$\alpha \in (0, 1)$

It is called  $(1-\alpha)$  confidence interval of the unknown parameter  $\theta$ . Here we generally consider a small value of  $\alpha$  such that  $0.01$  or  $0.05$  or  $0.1$  (in practice).

Idea:- For a fixed  $\alpha$ , we try to minimize the length of the interval  $[L(x_1, \dots, x_n), U(x_1, \dots, x_n)]$ .

Example 1:  $x_1, \dots, x_n \sim N(\mu, \sigma^2) \rightarrow \sigma^2 \rightarrow \text{known}$ .  
Want to find  $(1-\alpha)$  CI of  $\mu$ .

Sol :- Since  $\sigma^2$  is known, we have

$$\frac{\sqrt{n} \left( \frac{1}{n} \sum_{i=1}^n x_i - \mu \right)}{\sigma} \sim N(0, 1).$$

Denote  $Y_n = \frac{\sqrt{n} (\bar{x}_n - \mu)}{\sigma}$ , where  $\bar{x}_n = \frac{1}{n} \sum_{i=1}^n x_i$ .

Let  $a$  &  $b$  such that

$$P[a \leq Y_n \leq b] = 1 - \alpha$$

$$\Leftrightarrow \Phi(b) - \Phi(a) = 1 - \alpha \quad \text{--- } ①$$

Since  $Y_n \sim N(\mu, 1)$ ,  $\Phi$  is the CDF of  $N(0, 1)$ .

Now,  $1 - \alpha = P\left[a \leq \frac{\bar{X}_n - \mu}{\sigma} \leq b\right]$

$$\Leftrightarrow 1 - \alpha = P\left[a \frac{\sigma}{\sqrt{n}} \leq \bar{X}_n - \mu \leq b \frac{\sigma}{\sqrt{n}}\right]$$

$$\Leftrightarrow 1 - \alpha = P\left[\bar{X}_n - b \frac{\sigma}{\sqrt{n}} \leq \mu \leq \bar{X}_n - a \frac{\sigma}{\sqrt{n}}\right].$$

Now, the length of the interval is

$$= (b - a) \frac{\sigma}{\sqrt{n}}.$$

Further, observe that  $\sigma$  &  $\sqrt{n}$  are known to us.

Hence, the problem reduces to minimizing  $(b - a)$  with the constraint  $\Phi(b) - \Phi(a) = 1 - \alpha$ .

The forms of the solutions of  $b$  and  $a$  will be

$$b = m \quad \& \quad a = -m \quad \text{---} \quad \textcircled{2}$$

Using  $\textcircled{2}$  in  $\textcircled{1}$ , we have

$$\Phi(m) - \Phi(-m) = 1 - \alpha$$

$$\Leftrightarrow \Phi(m) - \left\{ 1 - \Phi(m) \right\} = 1 - \alpha$$

$$\Leftrightarrow 2\Phi(m) - 1 = 1 - \alpha$$

$$\Leftrightarrow 2\Phi(m) = 2 - \alpha \Leftrightarrow \Phi(m) = 1 - \frac{\alpha}{2}$$

$$\Leftrightarrow m = \Phi^{-1}\left(1 - \frac{\alpha}{2}\right)$$

Hence,  $b = \Phi^{-1}\left(1 - \frac{\alpha}{2}\right)$

$$a = -\Phi^{-1}\left(1 - \frac{\alpha}{2}\right).$$

Since  $\Phi$  is a strictly increasing function.

Therefore, ~~the~~  $(1 - \alpha) \subset I$  of  $m$  is (when  $\sigma$  is known)

$$\left( \frac{1}{n} \sum_{i=1}^n x_i - \Phi^{-1}\left(1 - \frac{\alpha}{2}\right) \frac{\sigma}{\sqrt{n}}, \frac{1}{n} \sum_{i=1}^n x_i + \Phi^{-1}\left(1 - \frac{\alpha}{2}\right) \frac{\sigma}{\sqrt{n}} \right).$$

Fact :- If  $\sigma$  is unknown, one can construct CI based on a consistent estimator of  $\sigma$ . (denoted by  $s$ ).

$$(s = \sqrt{\frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x}_n)^2}) .$$

In this case,  $(1-\alpha)$  - CI of  $\mu$  will be

$$\left( \frac{1}{n} \sum_{i=1}^n x_i - t_{n-1, 1-\frac{\alpha}{2}} \frac{s}{\sqrt{n}}, \frac{1}{n} \sum_{i=1}^n x_i + t_{n-1, 1-\frac{\alpha}{2}} \frac{s}{\sqrt{n}} \right) .$$

## Testing of hypotheses! -

(Neyman-Pearson Approach)

A few facts :-

- \* Assume that the data is a realization of some random variable
- \* \* Assume that the r.v. follows some parametric dist<sup>n</sup>.

$F_{\theta}$ , where  $\theta$  is unknown to us.

\* \* \*

Let  $\theta = 5$  or



One hypothesis

$\theta = 7$ .



Another hypothesis.

Testing of hypotheses will give you an idea, which hypothesis will be accepted and which one will be rejected.