Worksheet 3

Foundations of Bayesian Methodology

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Spring Semester 2022

Exercise 3 (Conjugate Bayes: analytical derivation)

Assumptions:

$$y_1, \dots, y_n \mid m \stackrel{i.i.d}{\sim} \mathcal{N}(m, \kappa^{-1})$$

 $m \sim \mathcal{N}(\mu, \lambda^{-1})$

3(a) The prior predictive distribution of one future observation y assuming that no observations have been collected yet.

$$\begin{split} f(y) &= \int_{-\infty}^{\infty} f(y \mid m) f(m) \mathrm{d}m \\ &= \int_{-\infty}^{\infty} \sqrt{\frac{\kappa}{2\pi}} \exp\left(-\frac{\kappa}{2} (y-m)^2\right) \sqrt{\frac{\lambda}{2\pi}} \exp\left(-\frac{\lambda}{2} (m-\mu)^2\right) \mathrm{d}m \\ &= \frac{\sqrt{\kappa\lambda}}{2\pi} \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2} \left(\kappa (y^2 - 2ym + m^2) + \lambda (m^2 - 2m\mu + \mu^2)\right)\right) \mathrm{d}m \\ &= \frac{\sqrt{\kappa\lambda}}{2\pi} \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2} \left((\kappa + \lambda) \left(m - \frac{\kappa y + \lambda \mu}{\kappa + \lambda}\right)^2 - \frac{(\kappa y + \lambda \mu)^2}{\kappa + \lambda} + \kappa y^2 + \lambda \mu^2\right)\right) \mathrm{d}m \\ &= \frac{\sqrt{\kappa\lambda}}{2\pi} \int_{-\infty}^{\infty} \exp\left(-\frac{1}{2} \left((\kappa + \lambda) \left(m - \frac{\kappa y + \lambda \mu}{\kappa + \lambda}\right)^2 + \frac{\kappa \lambda (y - \mu)^2}{\kappa + \lambda}\right)\right) \mathrm{d}m \\ &= \frac{\sqrt{\kappa\lambda}}{2\pi} \exp\left(-\frac{\kappa \lambda (y - \mu)^2}{2(\kappa + \lambda)}\right) \int_{-\infty}^{\infty} \exp\left(-\frac{\kappa + \lambda}{2} \left(m - \frac{\kappa y + \lambda \mu}{\kappa + \lambda}\right)^2\right) \mathrm{d}m \\ &= \frac{\sqrt{\kappa\lambda}}{2\pi} \sqrt{\frac{2\pi}{\kappa + \lambda}} \exp\left(-\frac{\kappa \lambda (y - \mu)^2}{2(\kappa + \lambda)}\right) \int_{-\infty}^{\infty} \sqrt{\frac{\kappa + \lambda}{2\pi}} \exp\left(-\frac{\kappa + \lambda}{2} \left(m - \frac{\kappa y + \lambda \mu}{\kappa + \lambda}\right)^2\right) \mathrm{d}m \\ &= \frac{\sqrt{\kappa\lambda}}{2\pi} \sqrt{\frac{2\pi}{\kappa + \lambda}} \exp\left(-\frac{\kappa \lambda (y - \mu)^2}{2(\kappa + \lambda)}\right) \\ &= \sqrt{\frac{1}{2\pi} \left(\frac{1}{\lambda} + \frac{1}{\kappa}\right)} \exp\left(-\frac{(y - \mu)^2}{2\left(\frac{1}{\lambda} + \frac{1}{\kappa}\right)}\right) \end{split}$$

The prior predictive distribution of one future observation y is

$$\mathcal{N}(\mu, \lambda^{-1} + \kappa^{-1})$$

3(b) The posterior predictive distribution of one future observation y_{n+1} given that y_1, \dots, y_n have been observed.

$$f(y_{n+1} | y_1, \dots, y_n) = \int_{-\infty}^{\infty} f(y_{n+1}, m | y_1, \dots, y_n) dm$$

$$= \int_{-\infty}^{\infty} f(y_{n+1} | m, y_1, \dots, y_n) f(m | y_1, \dots, y_n) dm$$

$$= \int_{-\infty}^{\infty} f(y_{n+1} | m) f(m | y_1, \dots, y_n) dm$$

$$f(m \mid y_1, \dots, y_n) = \sqrt{\frac{n\kappa + \lambda}{2\pi}} \exp\left(-\frac{n\kappa + \lambda}{2} \left(m - \frac{\kappa n\bar{y} + \lambda\mu}{n\kappa + \lambda}\right)^2\right)$$

Denote:

$$\mu_{\text{post}} = \frac{\kappa n \bar{y} + \lambda \mu}{n\kappa + \lambda}$$

$$\lambda_{\text{post}} = n\kappa + \lambda$$

$$f(y_{n+1} \mid y_1, \dots, y_n) = \int_{-\infty}^{\infty} \sqrt{\frac{\kappa}{2\pi}} \exp\left(-\frac{\kappa}{2}(y_{n+1} - m)^2\right) \sqrt{\frac{\lambda_{\text{post}}}{2\pi}} \exp\left(-\frac{\lambda_{\text{post}}}{2}(m - \mu_{\text{post}})^2\right) dm$$

Repeating the same derivation steps as for the prior predictive distribution, we obtain the posterior predictive distribution:

$$y_{n+1} \mid y_1, \dots, y_n \sim \mathcal{N}(\mu_{\text{post}}, \lambda_{\text{post}}^{-1} + \kappa^{-1})$$

Exercise 4 (Conjugate Bayesian analysis in practice)

Prior predictive distribution:

$$y \sim \mathcal{N}(\mu, \lambda^{-1} + \kappa^{-1}) = \mathcal{N}(161, 970)$$

```
## P[y>200]
pnorm(200, mean=161, sd=sqrt(970), lower.tail=F)
```

[1] 0.1052459

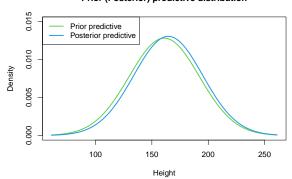
Posterior distribution:

$$m \mid y_1, \dots, y_n \sim \mathcal{N}\left(\frac{\kappa n \bar{y} + \lambda \mu}{n\kappa + \lambda}, (n\kappa + \lambda)^{-1}\right) = \mathcal{N}(164.558, 34.80663)$$

Posterior predictive distribution:

$$y_{n+1} \mid y_1, \dots, y_n \sim \mathcal{N}\left(\frac{\kappa n \bar{y} + \lambda \mu}{n\kappa + \lambda}, (n\kappa + \lambda)^{-1} + \kappa^{-1}\right) = \mathcal{N}(164.558, 934.80663)$$

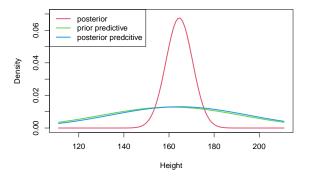
Prior (Posterior) predictive distribution



```
pnorm(200, mean=164.558, sd=sqrt(934.80663), lower.tail=F)
```

[1] 0.1231879

4(c) Comparison between posterior, prior predictive, and posterior distributions



Exercise 5 (The change-of-variables formula)

$$X \sim \text{Gamma}(a, b)$$

$$f(x) = \frac{b^a}{\Gamma(a)} x^{a-1} \exp(-bx)$$

$$P(Y \le y) = P(g(X) \le y)$$
$$= P(X \le g^{-1}(y))$$
$$F_Y(y) = F_X(g^{-1}(y))$$

By differentiating the CDFs on both sides w.r.t. y, we can get the PDF of Y. If the function $g(\cdot)$ is monotonically increasing:

$$f_Y(y) = f_X(g^{-1}(y)) \cdot \frac{d}{dy}g^{-1}(y)$$

If the function $g(\cdot)$ is monotonically decreasing:

$$f_Y(y) = -f_X(g^{-1}(y)) \cdot \frac{d}{dy}g^{-1}(y)$$

Therefore:

$$f_Y(y) = f_X(g^{-1}(y)) \cdot \left| \frac{d}{dy} g^{-1}(y) \right|$$

$$Y = \frac{1}{X} \implies X = \frac{1}{Y}$$

$$f(x) = \frac{b^a}{\Gamma(a)} x^{a-1} \exp(-bx)$$

$$f_Y(y) = f_X(g^{-1}(y)) \cdot \left| \frac{d}{dy} g^{-1}(y) \right|$$

$$= \frac{b^a}{\Gamma(a)} \left(\frac{1}{y} \right)^{a-1} \exp(-\frac{b}{y}) \cdot \left| -\frac{1}{y^2} \right|$$

$$= \frac{b^a}{\Gamma(a)} \left(\frac{1}{y} \right)^{a+1} \exp(-\frac{b}{y})$$

$$Z = \sqrt{\frac{1}{X}} \implies X = \frac{1}{Z^2}$$

$$f(x) = \frac{b^a}{\Gamma(a)} x^{a-1} \exp(-bx)$$

$$f_Z(z) = f_X(g^{-1}(z)) \cdot \left| \frac{d}{dz} g^{-1}(z) \right|$$

$$= \frac{b^a}{\Gamma(a)} \left(\frac{1}{z^2} \right)^{a-1} \exp(-\frac{b}{z^2}) \cdot \left| -\frac{2}{z^3} \right|$$

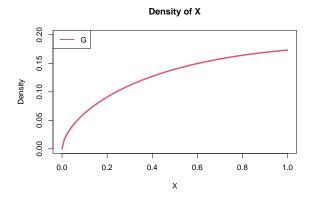
$$= \frac{b^a}{\Gamma(a)} 2 \left(\frac{1}{z} \right)^{2a+1} \exp(-\frac{b}{z^2})$$

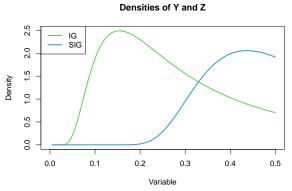
```
## Define inverse-gamma distribution function
dinvgamma <- function(x, a, b) {
   return(
        (b^a)/gamma(a) * (1/x)^(a+1) * exp(-b/x)
   )
}

## Define square root inverse-gamma distribution function</pre>
```

```
dsqrtinvgamma <- function(x, a, b) {
  return(
    2 * (b^a)/gamma(a) * (1/x)^(2*a+1) * exp(-b/x^2)
  )
}</pre>
```

```
a <- 1.6
b <- 0.4
```





Exercise 6 (Monte Carlo: transformations of random variables)

```
## Set seed for reproducible results
set.seed(44566)

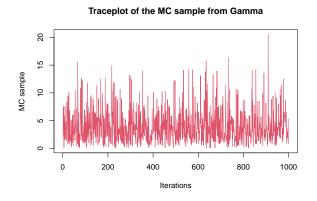
## Parameters for Gamma
a <- 1.6 # shape
b <- 0.4 # rate (inverse of scale)

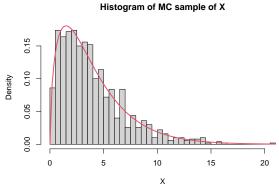
## MC sample size
M <- 1000</pre>
```

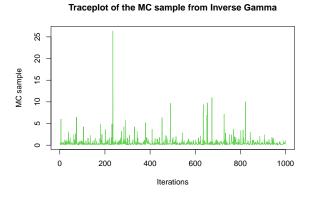
```
## Generate a MC sample of size 1000 from Gamma
mc.G <- rgamma(M, shape=a, rate=b)

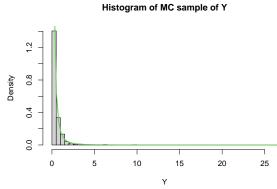
## Generate a MC sample of size 1000 from Inverse Gamma
mc.IG <- 1 / mc.G

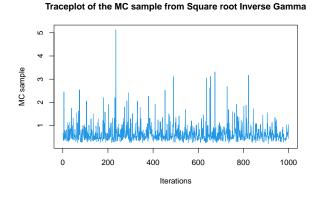
## Generate a MC sample of size 1000 from Square root Inverse Gamma
mc.SIG <- sqrt(1/mc.G)</pre>
```

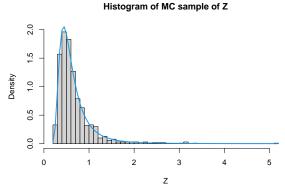












```
hist(mc.G, breaks=50, freq=FALSE)
curve(dgamma(x, a, b), add=TRUE, col=2, lwd=2)
hist(mc.IG, breaks=50, freq=FALSE)
curve(dinvgamma(x, a, b), add=TRUE, col=3, lwd=2)
hist(mc.SIG, breaks=50, freq=FALSE, ylim=c(0, 2))
curve(dsqrtinvgamma(x, a, b), add=TRUE, col=4, lwd=2)
hist(mc.G, breaks=50, freq=FALSE)
curve(dgamma(x, a, b), from=min(mc.G), to=max(mc.G), add=TRUE, col=2, lwd=2)
hist(mc.IG, breaks=50, freq=FALSE)
curve(dinvgamma(x, a, b), from=min(mc.IG), to=max(mc.IG), add=TRUE, col=3, lwd=2)
hist(mc.SIG, breaks=50, freq=FALSE, ylim=c(0, 2))
curve(dsqrtinvgamma(x, a, b), from=min(mc.SIG), to=max(mc.SIG), add=TRUE, col=4, lwd=2)
## Gamma
meanG <- mean(mc.G)</pre>
medG <- median(mc.G)</pre>
## Inverse Gamma
meanIG <- mean(mc.IG)</pre>
medIG <- median(mc.IG)</pre>
## Square root Inverse Gamma
meanSIG <- mean(mc.SIG)</pre>
medSIG <- median(mc.SIG)</pre>
df <- data.frame(</pre>
  c(meanG, meanIG, meanSIG),
  c(medG, medIG, medSIG)
colnames(df) <- c("Sample Mean", "Sample Median")</pre>
rownames(df) <- c("G", "IG", "SIG")</pre>
knitr::kable(df, caption="Summary statistics", align="c")
```

Table 1: Summary statistics

	Sample Mean	Sample Median
G	3.9667004	3.2625686
IG	0.6143637	0.3065072
SIG	0.6672328	0.5536309