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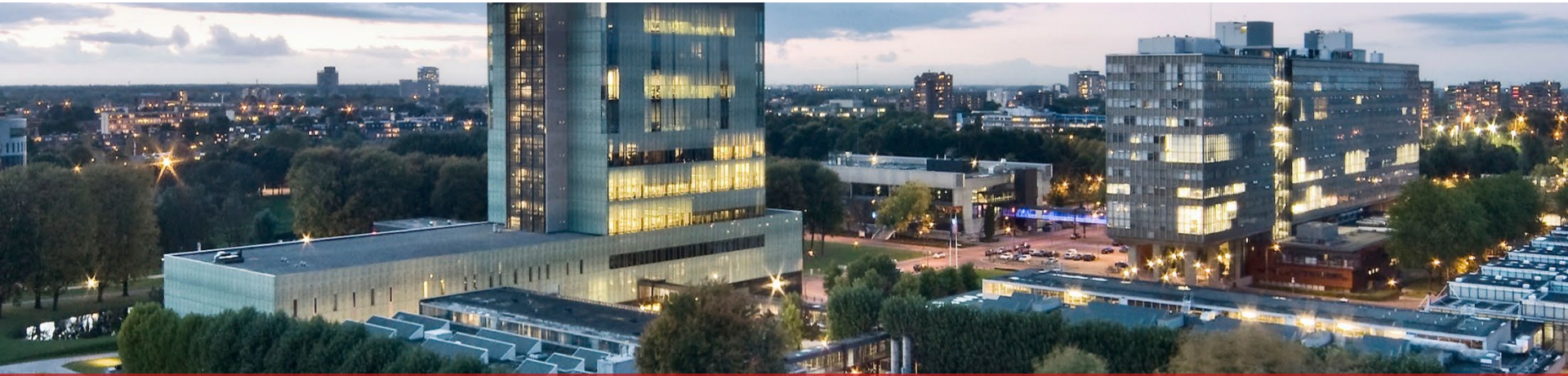


Variational Bayesian inference for system identification

Wouter M. Kouw

Nonlinear System Identification Workshop 2023

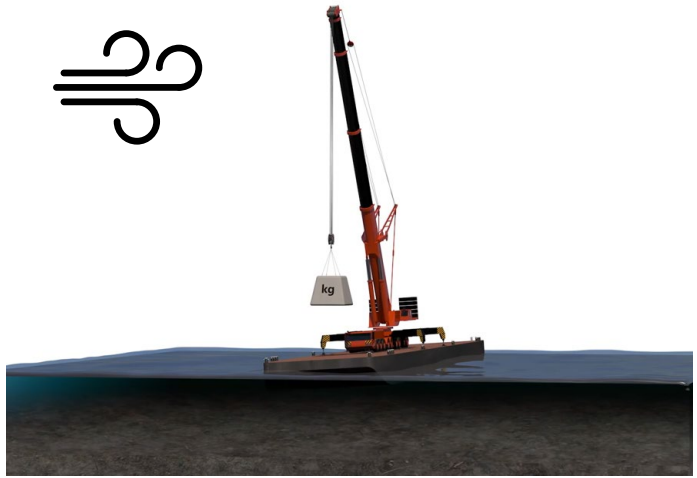
[Code and videos @https://github.com/wmkouw/NSIW2023-keynote](https://github.com/wmkouw/NSIW2023-keynote)

Onderzoeksraad voor veiligheid: https://youtu.be/LJevke4_i5Y

Uncertainty

- What value do parameters in my model have?
- How many parameters affect my system?
- Do my parameters change over time?

- Is my system affected by external disturbances?
- Which model should I select for this system?
- What should I measure to identify my system?



Onderzoeksraad voor veiligheid: https://youtu.be/LJevke4_i5Y

Modelling

Typical models for system identification look something like

$$x_k = f_\theta(x_{k-1}, u_k) + w_k \,,$$
$$y_k = g_\eta(x_k) + v_k \,.$$

or like

$$y_k = f_\theta(u_k, u_{k-1}, \dots, y_{k-1}, \dots) + e_k \,,$$

But where are the uncertainties?

Probabilistic modelling

Probabilistic models aim to include more sources of uncertainty:

$$p(y, u, \theta, \sigma) = p(y|u, \theta, \sigma)\, p(u)\, p(\theta)p(\sigma)$$

generative model **observation model** **input prior** **parameter priors**

Formally, one also conditions on assumptions leading to model design:

$$p(y, u, \theta, \sigma \mid \mathcal{M} = m_1)$$

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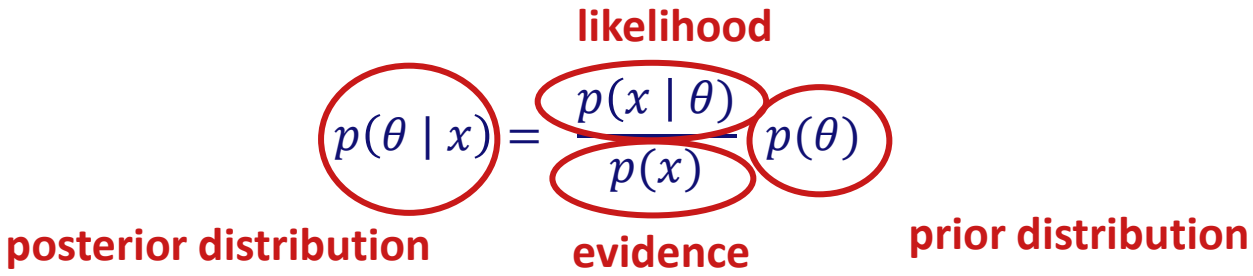
Inference

We can estimate unknowns by inverting the model:



This is known as Bayes' rule.

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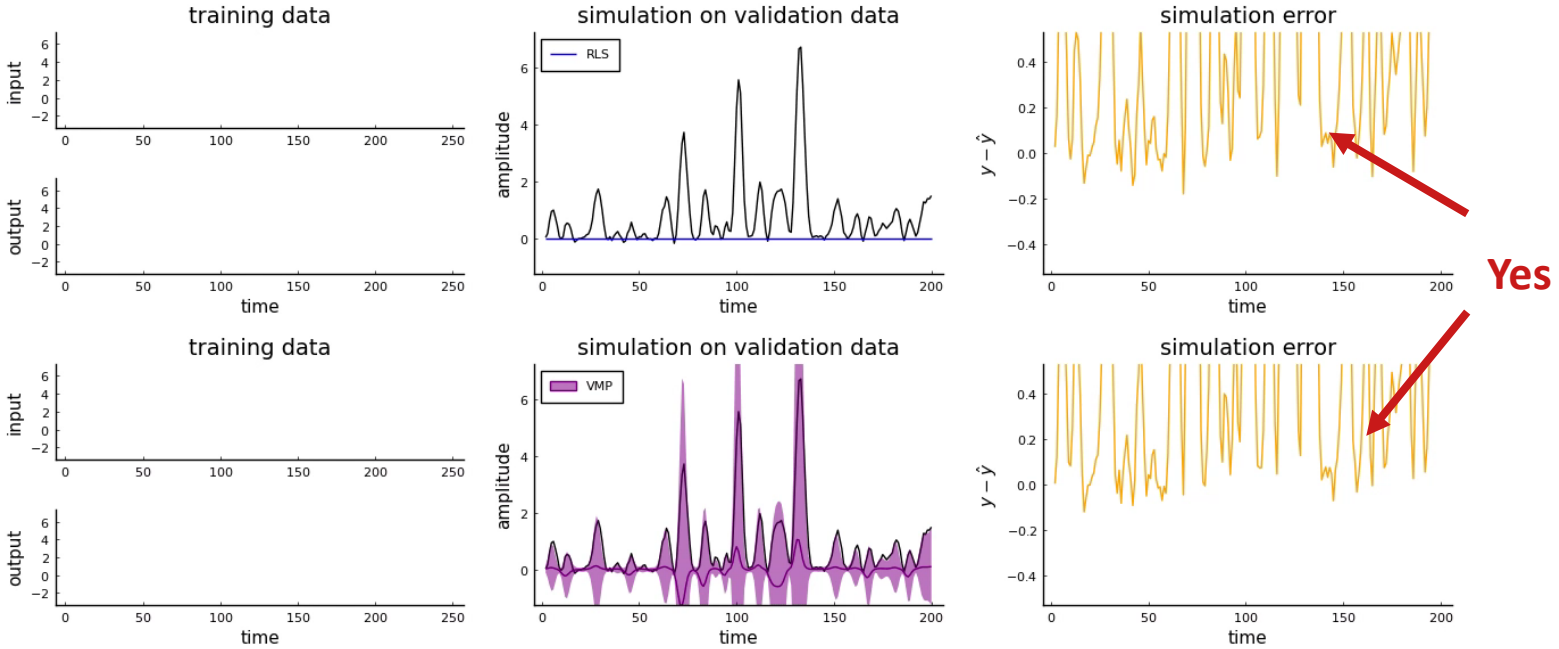
Is all that extra work useful?

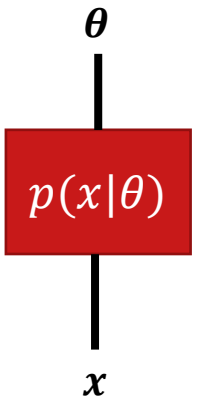


Kouw, Podusenko, Koudahl & Schoukens (2022). Variational message passing for online polynomial NARMAX identification.

Factor graphs

Probabilistic model equations quickly become complex and hard to read.

It helps to adopt a visual language: factor graphs.



Edges represent variables in the model.

Nodes represent relationships between variables.

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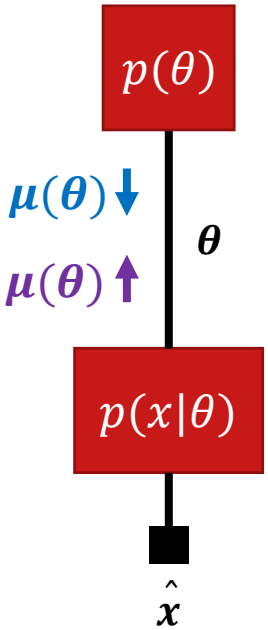
Message passing

The following is a complete factor graph:



Terminal nodes are priors.

The combination of the prior and the likelihood to form the posterior can be expressed as messages passed from nodes.

$$p(\theta|x=\hat{x}) \propto \int \delta(x-\hat{x})p(x|\theta)\,dx\,p(\theta)$$

Black nodes represent observed data.

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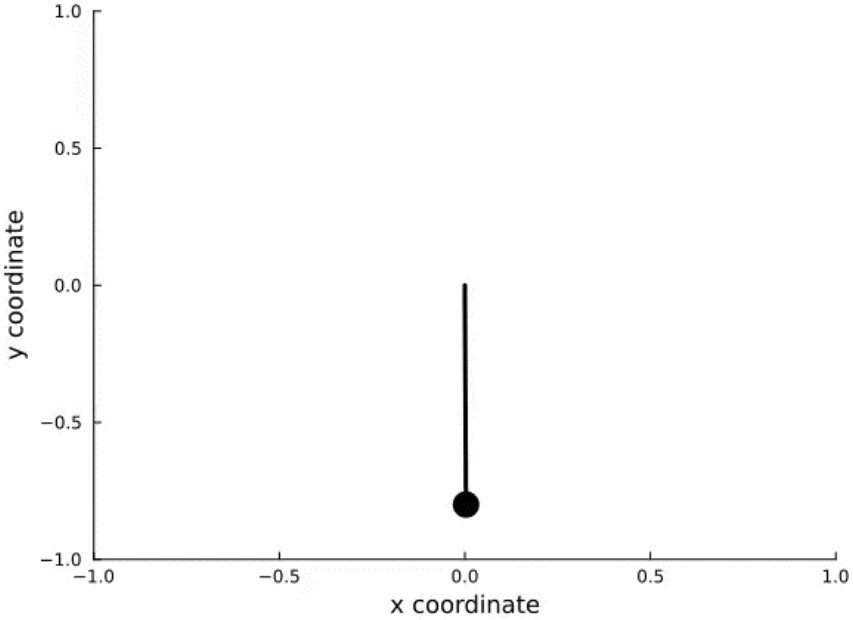
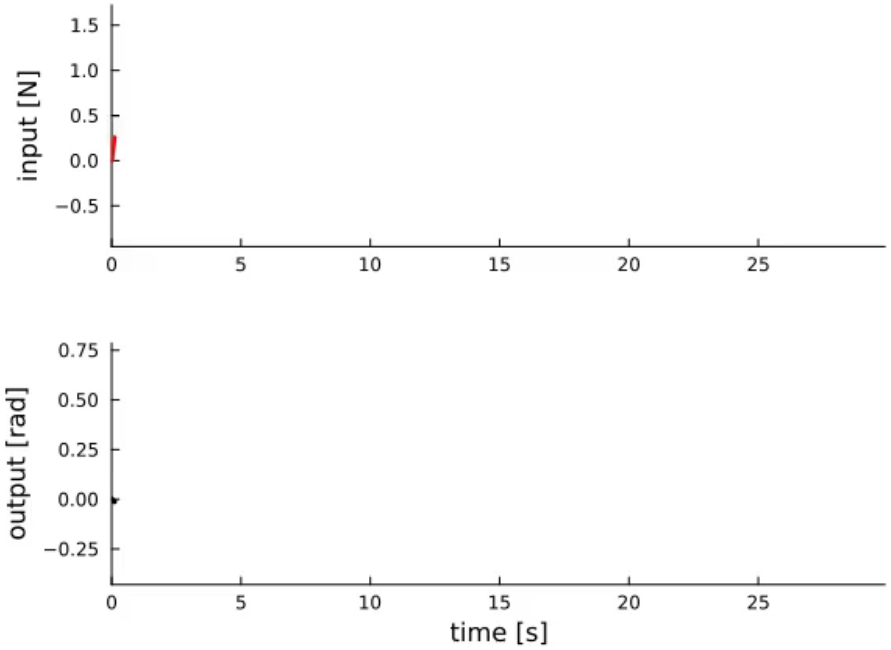
Demonstration system





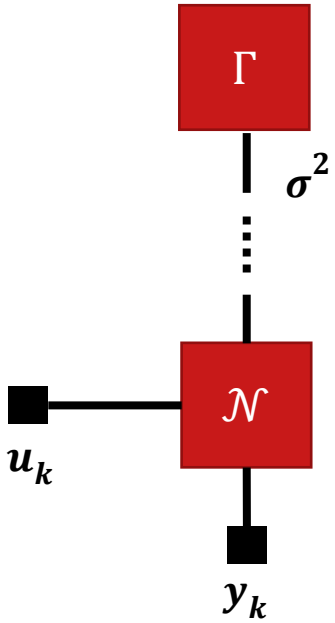
Model 1

Consider a prediction based on an unaltered input u_k with likelihood variance σ^2:

$$y_k = u_k + e_k, \quad \text{with} \quad e_k \sim \mathcal{N}(0, \sigma^2) .$$

In probabilistic model form, this could become:

$$p(y_k, \sigma^2 | u_k) = \mathcal{N}(y_k | u_k, \sigma^2)\Gamma(\sigma^2 | \alpha, \beta) .$$



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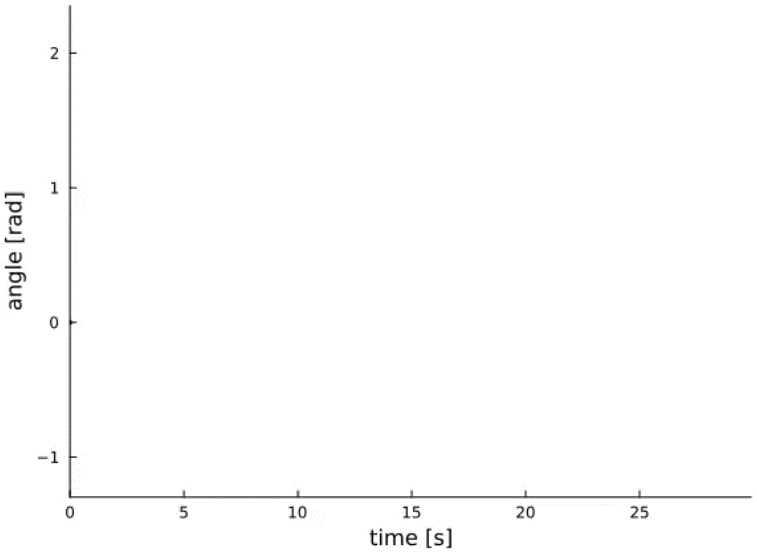
Model 1


angle [rad]
2
1
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−1
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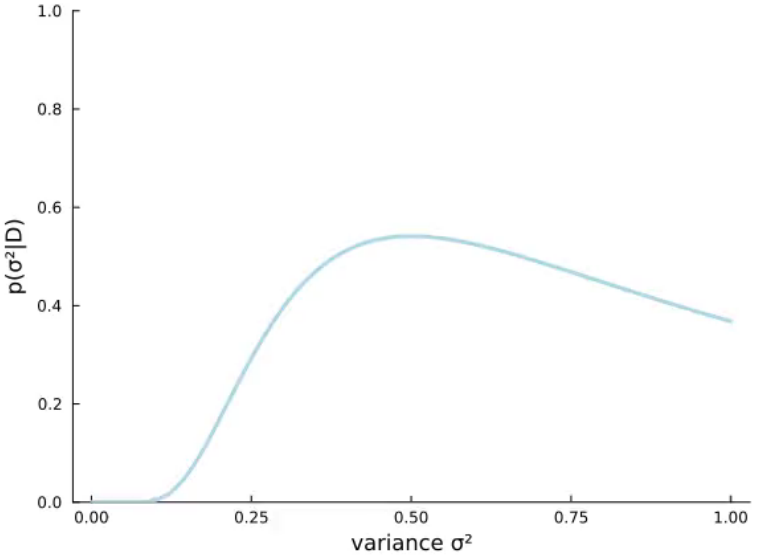
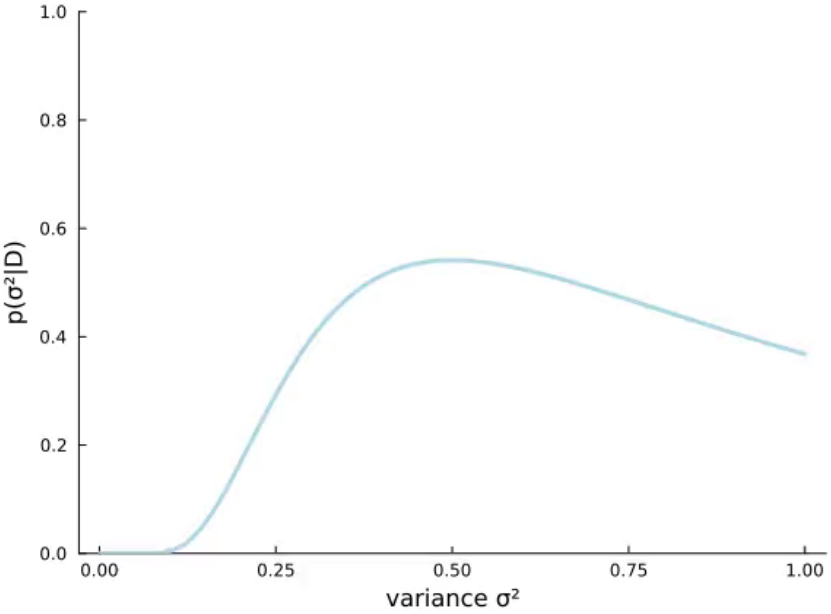
5
10
15
20
25
time [s]


p(σ²|D)
1.0
0.8
0.6
0.4
0.2
0.0
0.00
0.25
0.50
0.75
1.00
variance σ²

Model 1


1.0
0.8
0.6
0.4
0.2
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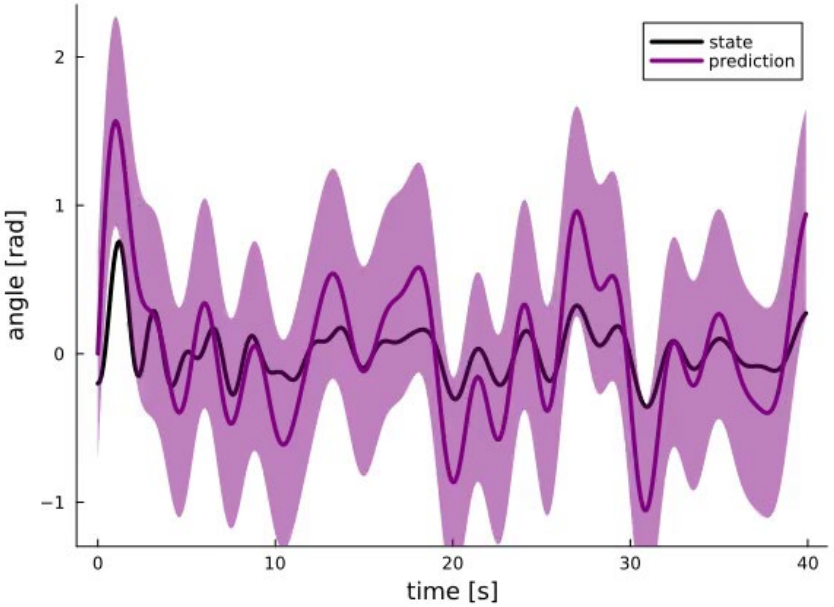
p(σ²|D)
0.00
0.25
0.50
0.75
1.00
variance σ²


state
prediction
2
1
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angle [rad]
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time [s]

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Model 1

This model obviously doesn't work very well.

A straightforward extension is a NARX model:

$$y_k = \theta^\intercal \varphi(u_k, u_{k-1}, \dots, y_{k-1}, \dots) + e_k \,,$$

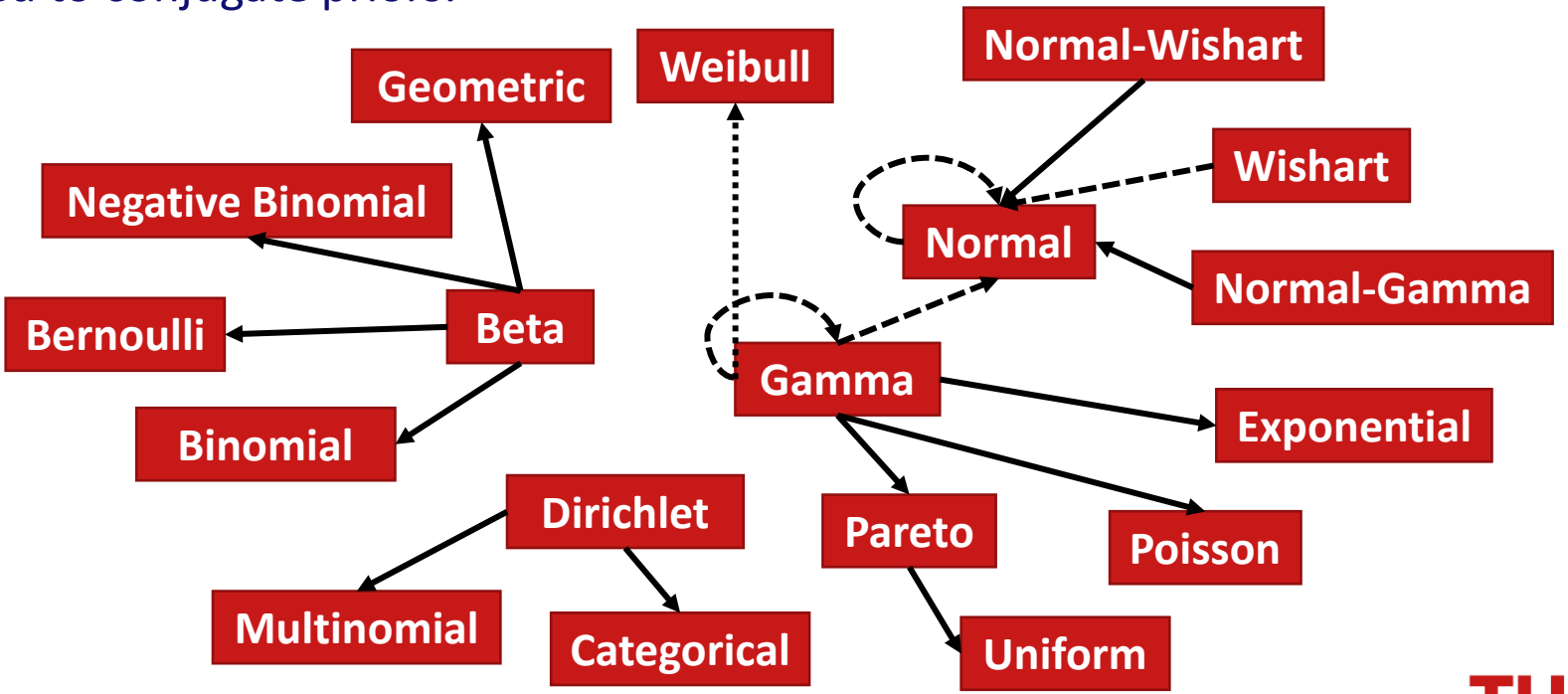
But now we run into a problem: we can't obtain a posterior distribution.

It requires solving an intractable integral:

$$p(y_k|u_k) = \iint p(y_k|u_k, \theta, \sigma^2) p(\theta) p(\sigma^2) d\theta d\sigma^2$$

Exact inference

Limited to conjugate priors:



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Approximate inference

We may approximate the posterior $p(\theta|x)$ with a distribution $q(\theta)$.

To do that, we need an objective characterizing the dissimilarity between q and p.

$$\mathcal{F}[q] = \int_{\Theta} q(\theta) \log \frac{q(\theta)}{p(\theta, x)} d\theta$$

This is known as a "free energy" functional and may be understood through:

$$\mathcal{F}[q] = \int_{\Theta} q(\theta) \log \frac{1}{p(x|\theta)} d\theta + \int_{\Theta} q(\theta) \log \frac{q(\theta)}{p(\theta)} d\theta$$

prediction error **complexity**

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Minimizing free energy

The free energy is a functional, i.e., a function of functions.

We are looking for the *probability distribution function* that minimizes it:

$$q^* = \arg\min_{q \in Q} \mathcal{F}[q]$$

The space Q represents the space of candidate functions.

Possible constraints on Q include:

1. Data, $q(x) = \delta(x - \hat{x})$.
2. Parametrization, $q(\theta) = \mathcal{N}(\theta|m, v)$.
3. Factorization, $q(x, \theta) = q(x)q(\theta)$.
4. Probability mass in a subspace.

Van de Laar (2021). Chance-constrained active inference.

Minimizing free energy

Suppose we have a distribution $p(\theta)$ and we wish to minimize:

$$\mathcal{F}[q] = \int_{\Theta} q(\theta) \log \frac{q(\theta)}{p(\theta)} d\theta$$

The function q is constrained to be a valid probability distribution:

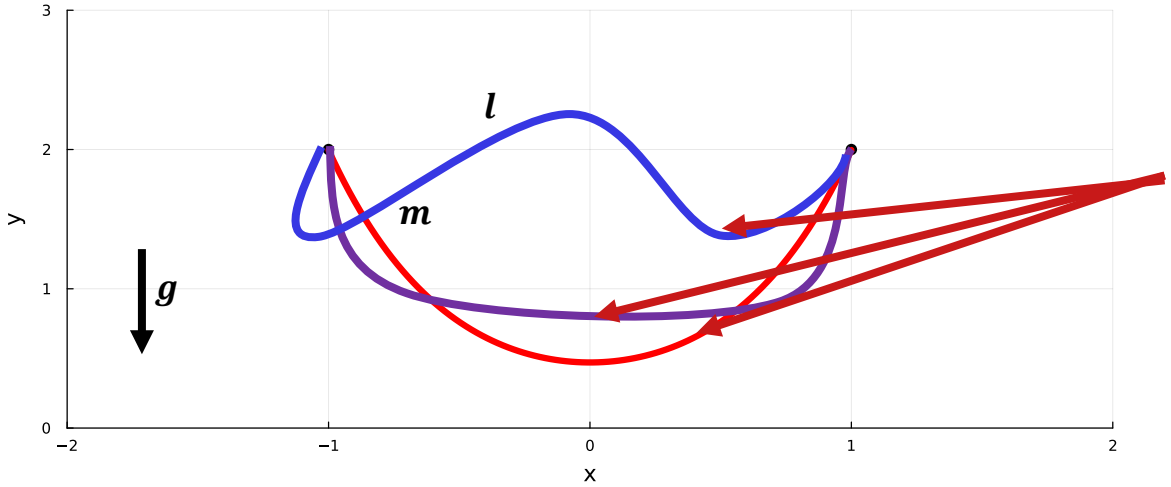
$$\mathcal{L}[q] = \mathcal{F}[q] + \lambda \left(\int_{\Theta} q(\theta)\, d\theta - 1 \right).$$

To find the minimizer, we must find the functional derivative $\frac{\delta}{\delta q} \mathcal{L}[q]$ and set it to 0.

In essence, variational Bayes turns integration into optimization.

Variations on a curve

Consider two fixed anchor points with a chain hanging between them:



The red chain minimizes *potential energy* (from Lagrangian mechanics).

In our probabilistic model, we have variations $q(\theta) = q^*(\theta) + \varepsilon\phi(\theta)$.

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Minimizing free energy

We can find the functional derivative by considering how much the Lagrangian changes as a function of the variation, and setting that to 0;

$$\frac{d}{d\varepsilon}\mathcal{L}[q^* + \varepsilon\phi]\Big|_{\varepsilon=0} = 0$$

Expanding the Lagrangian gives:

$$\int_\Theta \frac{d}{d\varepsilon}(q^* + \varepsilon\phi)\log\frac{q^* + \varepsilon\phi}{p}\Big|_{\varepsilon=0} d\theta + \lambda\int_\Theta \frac{d}{d\varepsilon}(q^* + \varepsilon\phi)\Big|_{\varepsilon=0} d\theta = 0$$

$$\int_\Theta \left(\log\frac{q^*}{p} + 1 + \lambda\right)\phi d\theta = 0$$

Senöz, van de Laar, Bagaev & de Vries (2021). Variational Message Passing and Local Constraint Manipulation in Factor Graphs.

Minimizing free energy

The common term is the functional derivative we were looking for.

$$\int_{\Theta} \left(\log \frac{q^*}{p} + 1 + \lambda \right) \phi d\theta$$

The Lagrangian is 0 when the functional derivative is 0:

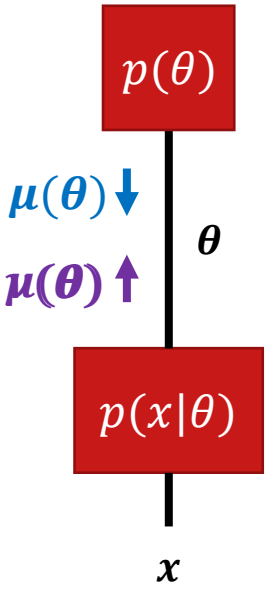
$$\frac{\delta}{\delta q} \mathcal{L}[q] = \log \frac{q^*}{p} + 1 + \lambda = 0$$

$$q^* = \frac{1}{\exp(1+\lambda)} p$$

Senöz, van de Laar, Bagaev & de Vries (2021). Variational Message Passing and Local Constraint Manipulation in Factor Graphs.

Variational message passing

One can distribute the free energy functional over a factor graph.



Variational approximation can be applied to factor nodes locally.

This turns standard messages into “variational messages”.

$$\nu(\theta) \propto \exp\left(\int_x q(x)\log p(x|\theta)\, dx\right)$$

Dauwels (2007). On variational message passing on factor graphs.

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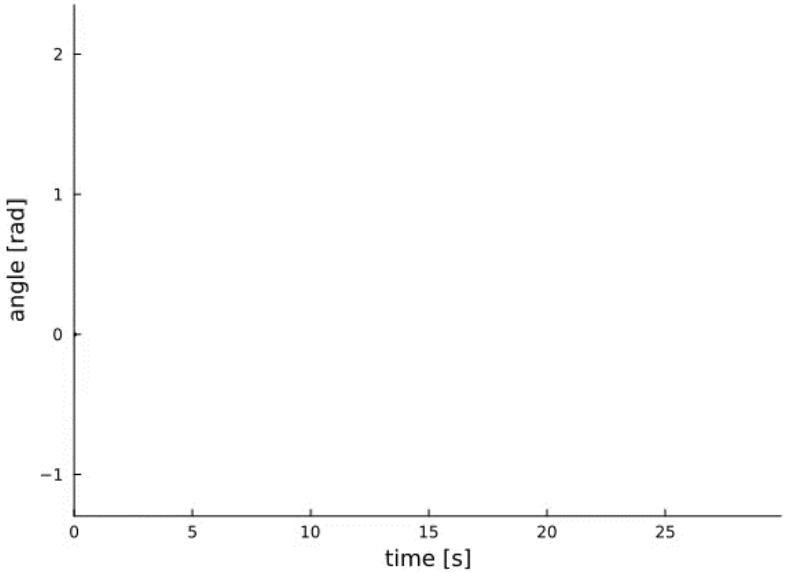
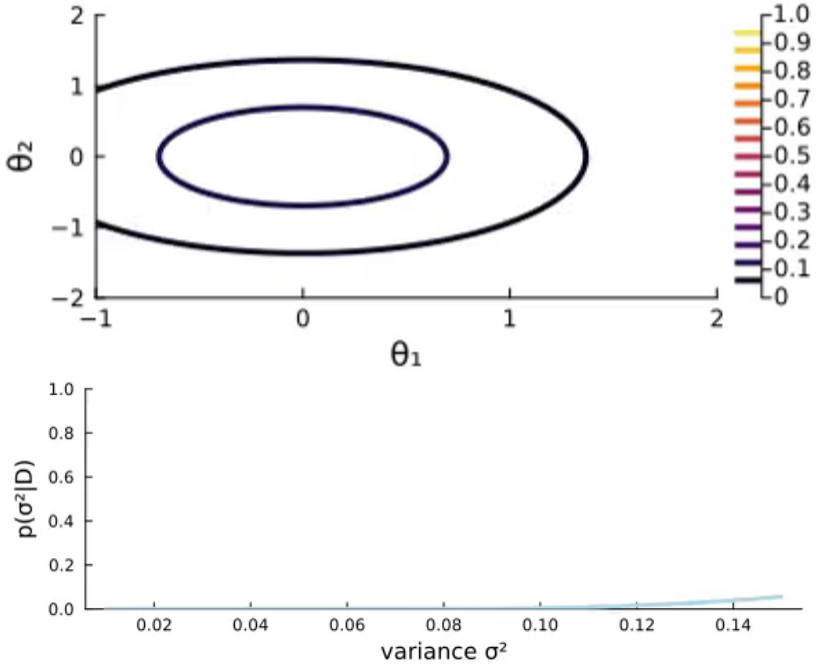
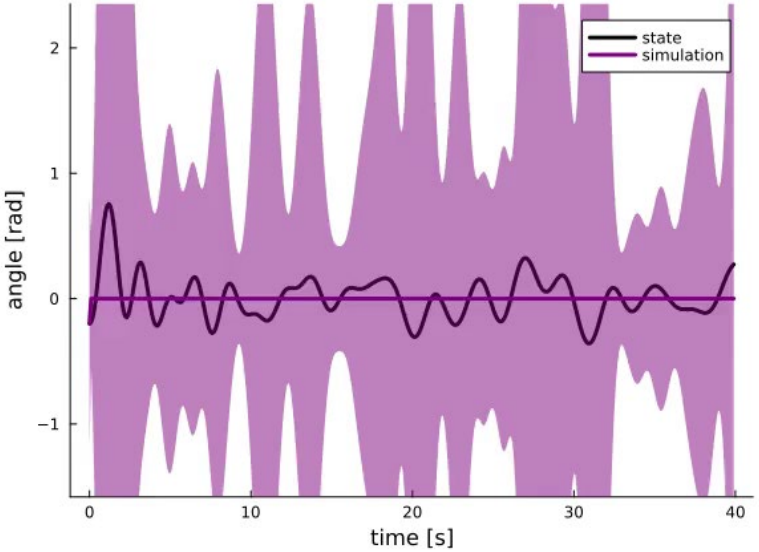
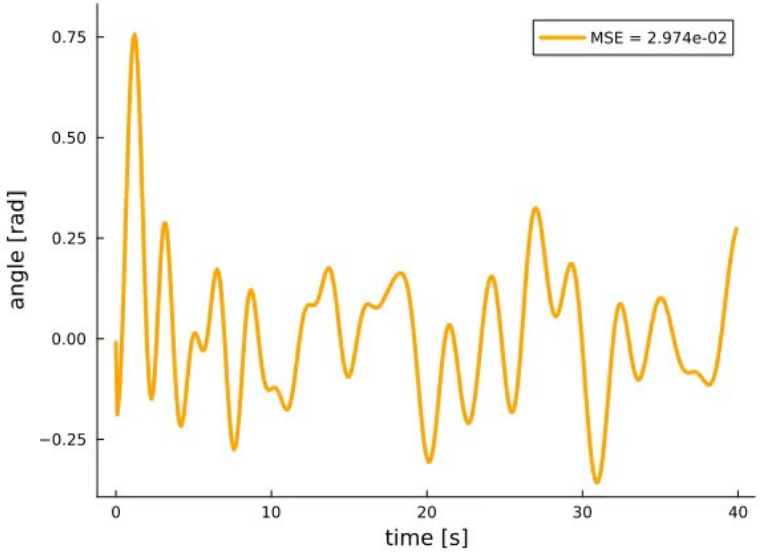
Model 2





Model 2


state
simulation
angle [rad]
time [s]
2
1
0
−1
0
10
20
30
40


MSE = 2.974e-02
angle [rad]
time [s]
0.75
0.50
0.25
0.00
−0.25
0
10
20
30
40

Take-aways

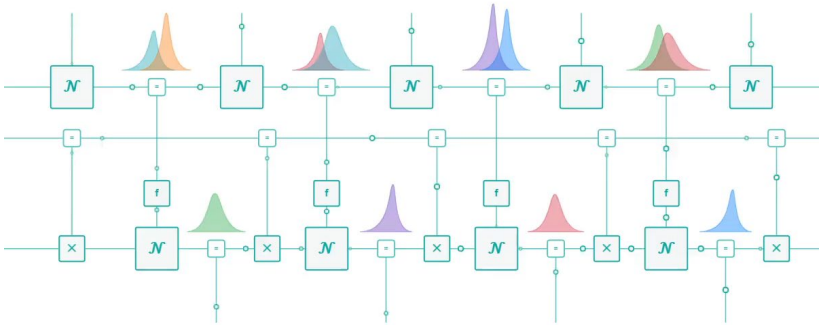
1. Quantified uncertainty should be part of models.

2. Variational Bayes turns integration into optimization.

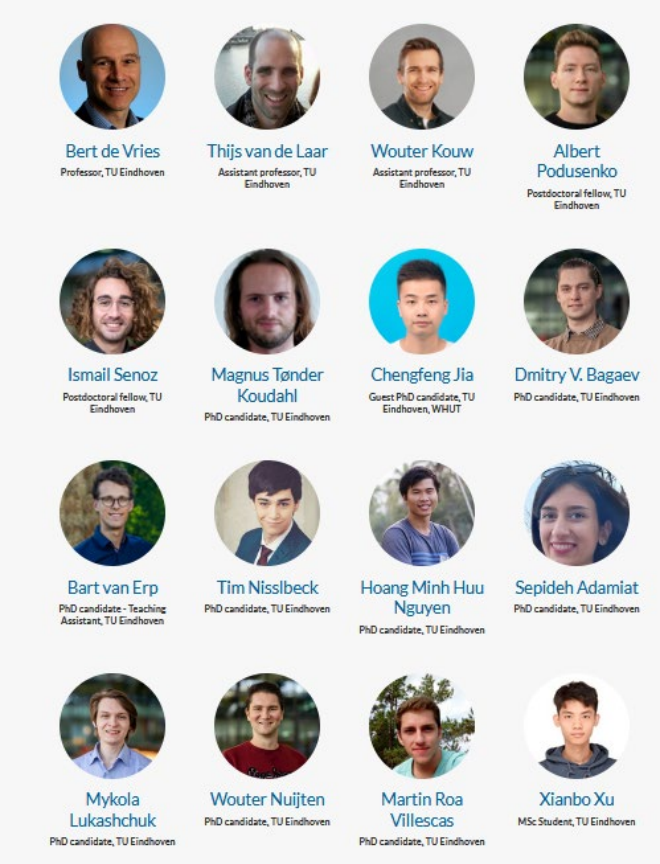
3. Variational message passing is inference distributed over a factor graph.

BIASlab
Signal Processing Systems

Checkout: RxInfer



https://github.com/biaslab/RxInfer.jl



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Weakly informative priors

A common critique is that the act of “choosing priors” leads to non-objective results.

-> One should rely on as generic and uninformative priors as possible.

In the case of polynomial NARX models, I argue that one may use “weak information” in the sense that lower-order terms are more likely to have large coefficients than higher-order terms.

- This may be incorporated by having a zero-mean Gaussian prior with large variances for low-order terms (indicating uncertainty) and small variances for high-order terms (i.e., you are certain that the coefficient is close to 0).

Alternative free energy decomposition

The “free energy” objective decomposes into prediction error and complexity:

$$\mathcal{F}[q] = \int_{\Theta} q(\theta) \log \frac{1}{p(x|\theta)} d\theta + \int_{\Theta} q(\theta) \log \frac{q(\theta)}{p(\theta)} d\theta$$

It can also be decomposed as an upper bound to negative model evidence:

$$\mathcal{F}[q] = \int_{\Theta} q(\theta) \log \frac{q(\theta)}{p(\theta|x)} d\theta - \log p(x)$$

approximation to posterior (≥ 0)

model evidence

In this sense, a smaller free energy means 1) a better approximation of the posterior and/or 2) a better model for the given data.

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Normalization

The solution for q^* led to a mysterious 1 / exp term. Where does that come from?

It comes from the normalization constraint imposed on the Lagrangian.

If we plug the optimal form into the constraint function, we get:

$$\int \frac{1}{\exp(1+\lambda)} p(\theta) d\theta \; - 1 = 0$$

Solving for λ gives:

$$\lambda = \log \int p(\theta) d\theta - 1$$

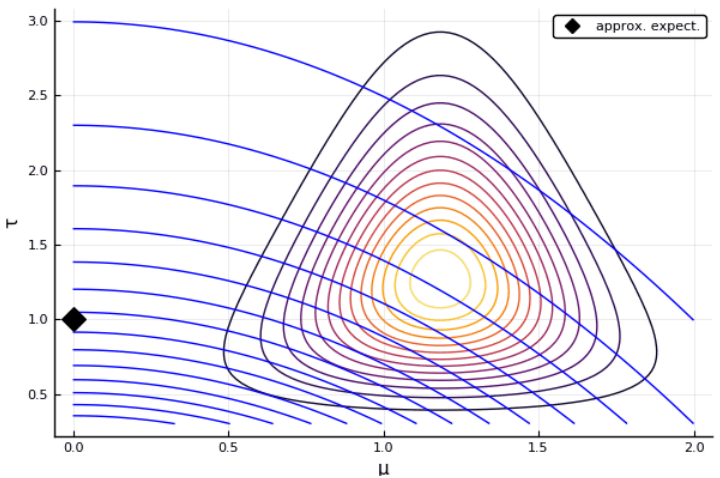
Mean-field

If there are multiple unknowns in the model, then you may choose to factorize q:

$$q(\theta, \sigma^2) := q(\theta)q(\sigma^2)$$

You would have multiple approximations, each dependent on the others.

-> Solutions must be iterated until convergence.



"Mean-field" is a common factorization choice,
but may lead to poor performance.

"Structured" factorizations are richer,
but require more manual derivation work.

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Limitations

Common parametric distributions are not closed under nonlinear transformations.

- A squared Gaussian distributed random variable is not Gaussian distributed.

Typical simplifications of q are based on (in)dependence between variables.

- This may cause under-estimation of variance.

Not much is known about the stability of variational Bayesian estimators and some appear to be (at least numerically) unstable in practice.