

第二次作业

题目 1. 设 $X^{(1)}$ 和 $X^{(2)}$ 均为 p 维随机向量, 已知

$$X = \begin{bmatrix} X^{(1)} \\ X^{(2)} \end{bmatrix} \sim N_{2p} \left(\begin{bmatrix} \mu^{(1)} \\ \mu^{(2)} \end{bmatrix}, \begin{bmatrix} \Sigma_1 & \Sigma_2 \\ \Sigma_2 & \Sigma_1 \end{bmatrix} \right)$$

其中 $\mu^{(i)} (i = 1, 2)$ 为 p 维向量, $\Sigma_i (i = 1, 2)$ 是 p 阶矩阵.

1. 证明 $X^{(1)} + X^{(2)}$ 和 $X^{(1)} - X^{(2)}$ 相互独立;
2. 求 $X^{(1)} + X^{(2)}$ 和 $X^{(1)} - X^{(2)}$ 的分布.

解答. 1. 设 I_p 为 p 阶单位阵, 由于

$$\begin{bmatrix} X^{(1)} + X^{(2)} \\ X^{(1)} - X^{(2)} \end{bmatrix} = \begin{bmatrix} I_p & I_p \\ I_p & -I_p \end{bmatrix} X \sim N_{2p} \left(\begin{bmatrix} \mu_1 + \mu_2 \\ \mu_1 - \mu_2 \end{bmatrix}, \begin{bmatrix} 2(\Sigma_1 + \Sigma_2) & 0 \\ 0 & 2(\Sigma_1 - \Sigma_2) \end{bmatrix} \right)$$

于是 $X^{(1)} + X^{(2)}$ 与 $X^{(1)} - X^{(2)}$ 独立.

2. 由上一问可知

$$\begin{aligned} X^{(1)} + X^{(2)} &= \begin{bmatrix} I_p & I_p \end{bmatrix} X \sim N_p(\mu_1 + \mu_2, 2(\Sigma_1 + \Sigma_2)), \\ X^{(1)} - X^{(2)} &= \begin{bmatrix} I_p & -I_p \end{bmatrix} X \sim N_p(\mu_1 - \mu_2, 2(\Sigma_1 - \Sigma_2)). \end{aligned}$$

题目 2. 设 $X \sim N_3(\mu, \Sigma)$, 其中

$$\mu = (\mu_1, \mu_2, \mu_3)', \quad \Sigma = \begin{bmatrix} 1 & \rho & \rho \\ \rho & 1 & \rho \\ \rho & \rho & 1 \end{bmatrix} \quad (0 < \rho < 1)$$

1. 求条件分布 $(X_1, X_2|X_3)$ 和 $(X_1|X_2, X_3)$.
2. 给定 $X_3 = x_3$ 时, 求出 X_1 和 X_2 的条件协方差.

解答. 1. 由条件分布计算公式可知

$$\begin{aligned} (X_1, X_2|X_3 = x_3) &\sim N_2 \left(\begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix} + \begin{bmatrix} \rho \\ \rho \end{bmatrix} (x_3 - \mu_3), \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix} - \begin{bmatrix} \rho \\ \rho \end{bmatrix} \begin{bmatrix} \rho & \rho \end{bmatrix} \right) \\ &\sim N_2 \left(\begin{bmatrix} \mu_1 + \rho(x_3 - \mu_3) \\ \mu_2 + \rho(x_3 - \mu_3) \end{bmatrix}, \begin{bmatrix} 1 - \rho^2 & \rho - \rho^2 \\ \rho - \rho^2 & 1 - \rho^2 \end{bmatrix} \right) \end{aligned}$$

$$\begin{aligned} (X_1|X_2 = x_2, X_3 = x_3) &\sim N_1 \left(\mu_1 + \begin{bmatrix} \rho & \rho \end{bmatrix} \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}^{-1} \left(\begin{bmatrix} x_2 \\ x_3 \end{bmatrix} - \begin{bmatrix} \mu_2 \\ \mu_3 \end{bmatrix} \right), 1 - \begin{bmatrix} \rho & \rho \end{bmatrix} \begin{bmatrix} 1 & \rho \\ \rho & 1 \end{bmatrix}^{-1} \begin{bmatrix} \rho \\ \rho \end{bmatrix} \right) \\ &\sim N_1 \left(\mu_1 + \frac{\rho}{1 + \rho} (x_2 + x_3 - \mu_2 - \mu_3), 1 - \frac{2\rho^2}{1 + \rho} \right) \end{aligned}$$

2. 由第一问可知, X_1, X_2 在给定 $X_3 = x_3$ 下的条件协方差均为 $1 - \rho^2$.

题目 3. 设 $X_1 \sim N(0, 1)$, $X_2 = \begin{cases} -X_1, & -1 \leq X_1 \leq 1, \\ X_1, & \text{否则.} \end{cases}$

1. 证明: $X_2 \sim N(0, 1)$.

2. 证明 (X_1, X_2) 的联合分布不是正态分布.

证明. 1. 当 $x \in [-1, 1]$ 时, $f_{X_2}(x) = f_{X_1}(-x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$; 当 $x \notin [-1, 1]$ 时, $f_{X_2}(x) = f_{X_1}(x) = \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}}$. 综上 $f_{X_1} = f_{X_2}$, 所以 $X_2 \sim N(0, 1)$.

2. 当 $x_2 \in [-1, 1]$ 时, X_1, X_2 的联合分布函数满足

$$F_{X_1, X_2}(x_1, x_2) = P[X_1 \leq x_1, X_2 \leq x_2] = P[X_1 \leq x_1, -X_1 \leq x_2] = P[-x_2 \leq X_1 \leq x_1]$$

所以 $F_{X_1, X_2}(x_1, x_2)$ 不是正态分布. □

题目 4. 设 $X \sim N_p(\mu, \Sigma)$, A 为对称阵, 证明:

(1). $E(XX') = \Sigma + \mu\mu'$;

(2). $E(X'AX) = \text{tr}(\Sigma A) + \mu' A \mu$;

(3). 当 $\mu = a \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix} =: a \mathbf{1}_p$, $A = I_p - \frac{1}{p} \mathbf{1}_p \mathbf{1}_p'$, $\Sigma = \sigma^2 I_p$ 时, 试利用 (1) 和 (2) 的结果证明

$$E(X'AX) = \sigma^2(p-1).$$

若记 $X = (X_1, \dots, X_p)'$ 此时 $X'AX = \sum_{i=1}^p (X_i - \bar{X})^2$, 则

$$E \left[\sum_{i=1}^p (X_i - \bar{X})^2 \right] = \sigma^2(p-1).$$

解答. (1).

$$\begin{aligned} & \int_{\mathbb{R}^p} \frac{xx'}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \exp \left\{ -\frac{1}{2} (x - \mu)' \Sigma^{-1} (x - \mu) \right\} dx \\ & \stackrel{x \leftarrow x - \mu}{=} \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^p} (x + \mu)(x + \mu)' \exp \left\{ -\frac{1}{2} x' \Sigma^{-1} x \right\} dx \\ & = \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^p} (xx' + 2\mu x' + \mu\mu') \exp \left\{ -\frac{1}{2} x' \Sigma^{-1} x \right\} dx \\ & = \mu\mu' - \frac{\Sigma}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^p} (x + 2\mu) d \exp \left\{ -\frac{1}{2} x' \Sigma^{-1} x \right\} \\ & = \mu\mu' + \frac{\Sigma}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^p} \exp \left\{ -\frac{1}{2} x' \Sigma^{-1} x \right\} dx \\ & = \Sigma + \mu\mu' \end{aligned}$$

(2).

$$\begin{aligned}
& \int_{\mathbb{R}^p} \frac{x'Ax}{(2\pi)^{\frac{p}{2}}|\Sigma|^{\frac{1}{2}}} \exp\left\{-\frac{1}{2}(x-\mu)'\Sigma^{-1}(x-\mu)\right\} dx \\
& \xrightarrow{x \leftarrow x-\mu} \frac{1}{(2\pi)^{\frac{p}{2}}|\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^p} (x+\mu)'A(x+\mu) \exp\left\{-\frac{1}{2}x'\Sigma^{-1}x\right\} dx \\
& \xrightarrow{A \text{ 为对称阵}} \frac{1}{(2\pi)^{\frac{p}{2}}|\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^p} (x'Ax + 2\mu'Ax + \mu'A\mu) \exp\left\{-\frac{1}{2}x'\Sigma^{-1}x\right\} dx \\
& = \mu'A\mu - \frac{\Sigma A}{(2\pi)^{\frac{p}{2}}|\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^p} (x+2\mu) d \exp\left\{-\frac{1}{2}x'\Sigma^{-1}x\right\} \\
& = \mu\mu' + \frac{\Sigma A}{(2\pi)^{\frac{p}{2}}|\Sigma|^{\frac{1}{2}}} \int_{\mathbb{R}^p} \exp\left\{-\frac{1}{2}x'\Sigma^{-1}x\right\} dx \\
& = \text{tr}(\Sigma A) + \mu'A\mu
\end{aligned}$$

(3). 由 (2) 可知: $E(X'AX) = \text{tr}(\Sigma A) + \mu'A\mu = p\sigma^2(1-1/p) + a^2 \begin{bmatrix} 0 & 0 & \cdots & 0 \end{bmatrix} \mathbf{1}_p = \sigma^2(p-1)$.

由于

$$\begin{aligned}
X'AX &= [X_1, \dots, X_p] \begin{bmatrix} 1-1/p & -1/p & \cdots & -1/p \\ -1/p & 1-1/p & \ddots & \vdots \\ \vdots & \ddots & \ddots & -1/p \\ -1/p & \cdots & -1/p & 1-1/p \end{bmatrix} \begin{bmatrix} X_1 \\ \vdots \\ X_p \end{bmatrix} \\
&= [X_1 - \bar{X} \quad X_2 - \bar{X} \quad \cdots \quad X_n - \bar{X}] \begin{bmatrix} X_1 \\ \vdots \\ X_p \end{bmatrix} \\
&= \sum_{i=1}^p X_i^2 - X_i\bar{X} = \sum_{i=1}^p X_i^2 - 2X_i\bar{X} + \sum_{i=1}^p X_i\bar{X} \\
&= \sum_{i=1}^p (X_i^2 - 2X_i\bar{X} + \bar{X}^2) = \sum_{i=1}^p (X_i - \bar{X})^2
\end{aligned}$$

故

$$E \left[\sum_{i=1}^p (X_i - \bar{X})^2 \right] = \sigma^2(p-1).$$