### **Preliminares**

**Definition 0.1.** The normal distribution  $N(\mu, \sigma^2)$  has the probability density function

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}}\exp(-\frac{1}{2}(\frac{x-\mu}{\sigma})^2)$$

**Definition 0.2.** The joint normal distribution  $N(\mu, K)$  is defined by probability density function:

$$f(\mathbf{x}) = \frac{1}{\sqrt{(2\pi)^n \det(K)}} \exp(-\frac{1}{2}(\mathbf{x} - \mu)^{\mathsf{T}} K^{-1}(\mathbf{x} - \mu))$$

**Theorem 0.1** (Joint normality implies marginal normality). If  $\mathbf{X} = (X_1, X_2, \dots, X_n)$  follows a joint normal distribution, then any linear combination of  $\mathbf{X}$  follows normal distribution.

## 1 Entropy, mutual information, divergence

**Definition 1.1** (Entropy). The entropy of a random variable  $X \in \mathcal{X}$  is defined by

$$H(X) = -\sum_{x \in \mathcal{X}} p_X(x) \log p_X(x)$$

**Definition 1.2** (KL Divergence). The **KL Divergence** between two probability distributions p, q on the same support  $\mathcal{X}$  is defined by

$$D(p||q) = \sum_{x \in \mathcal{X}} p(x) \log \frac{p(x)}{q(x)}$$

**Definition 1.3** (Mutual information). The mutual information between two random variables X, Y is defined by

$$I(X;Y) = H(X) - H(X \mid Y)$$

It can also be expressed as the KL divergence between the joint distribution  $p_{X,Y}$  and the product of marginal distribution  $p_X p_Y$ :

$$I(X;Y) = D(p_{X,Y} || p_X \cdot p_Y)$$

**Proposition 1.1.** Conditioning does not increase entropy

$$H(X \mid Y) \leq H(X)$$

A nice trick for reasoning about the entropy of "sum of random variables":

$$H(X + Y \mid X) = \sum_{x \in \mathcal{X}} p_X(x)H(X + Y \mid X = x)$$

$$= \sum_{x \in \mathcal{X}} p_X(x)H(x + Y \mid X = x)$$

$$= \sum_{x \in \mathcal{X}} p_X(x)H(Y \mid X = x)$$

$$= H(Y \mid X)$$

### 1.1 Convexity

**Definition 1.4.** A function f is convex over the interval  $x_1 \le x \le x_2$  if for all  $0 \le t \le 1$ :

$$(1-t)f(x_1) + tf(x_2) \ge f((1-t)x_1 + tx_2)$$

Some notable functions' convexity:

- 1. Entropy is a concave function with respect to the probability distribution
- 2. Mutual information I(X;Y) is concave with respect to  $p_X$  and is convex with respect to  $p_{Y|X}$  for some fixed  $p_X$
- 3. log is a concave function
- 4. The negative of a convex function is concave
- 5. The sum or product of a linear function and a concave function is concave
- 6. The composition of a linear function and a concave function is concave

**Theorem 1.1** (Jensen's Inequality). If f is a convex function and X is a random variable, then

$$E[f(X)] \ge f(E[X])$$

#### 1.2 Markov chain

**Definition 1.5.** Random variables X, Y, Z form a Markove chain (denoted by  $X \to Y \to Z$ ) if  $p(z \mid x, y) = p(z \mid y)$ 

**Theorem 1.2** (Data processing inequality). If  $X \to Y \to Z$ , then

$$I(X;Z) \le I(Y;Z)$$

Proposition 1.2.

$$I(X; Y \mid Z) = 0 \Leftrightarrow X \to Z \to Y$$

**Proposition 1.3.** If  $X \to Y \to Z$ , then  $Z \to Y \to X$ 

**Theorem 1.3.** Let X, Y be random variables. I(X; Y) is concave with respect to the probability distribution of X. For a fixed marignal distribution of X, I(X; Y) is convex with respect to  $f_{Y|X}$ 

**Theorem 1.4** (Fano's inequality). Let  $X \to Y \to \hat{X}$  represent an encode-decode process, where  $X, \hat{X} \in \mathcal{X}$  have the same support. Let e denote decoding error  $\hat{X} \neq X$ , then:

$$H(X \mid Y) \le h(P_e) + P_e \log(|\mathcal{X}|)$$

## 2 Entropy rate

**Definition 2.1.** A stochastic process is stationary if for all integer  $n \ge 1$  and integer time shift (possibly negative) l:

$$P[X_1^n] = P[X_{1+l}^{n+l}]$$

**Definition 2.2.** A Markov chain is time invariant if for all  $n \geq 1$ :

$$P[X_{n+1} \mid X_n] = P[X_2 \mid X_1]$$

**Definition 2.3.** There are two definitions of entropy rate:

1. 
$$H(\mathcal{X}) = \lim_{n \to \infty} \frac{1}{n} H(X_1^n)$$

2. 
$$H'(\mathcal{X}) = \lim_{n \to \infty} H(X_n \mid X_1^{n-1})$$

**Theorem 2.1.** If  $X_i$  is a stationary process, then the two entropy rates both exist and are equal

**Lemma 2.1.1.** If  $X_i$  is a stationary process, then  $H(X_n \mid X_1^{n-1})$  is non-increasing with respect to n and has a limit

## 3 Asymptotic equipartition property

**Theorem 3.1** (Weak law of large numbers (WLLN)). Let  $X \sim p_X$  be a random variable, and let  $X_l \sim p_X$  be i.i.d. random variables following the same distributions, then:

$$\lim_{n \to \infty} \frac{1}{n} \sum_{i=1}^{n} X_i = E[X]$$

In other words, for any  $\delta, \epsilon > 0$ , there exists N > 0 such that for n > N:

$$P[|\frac{1}{n}\sum_{i=1}^{n}X_{i} - E[X]| < \delta] < \epsilon$$

Recall that the entropy of a random variable can be defined as  $H(X) = -E[\log(P_X)]$ , so the weak law of large numbers also applies: the joint distribution of i.i.d.  $X_i$  converges to the probability  $2^{-nH(X)}$ .

**Definition 3.1** (Typical sequence). For random variable  $X \sim p_X$ , the set of typical sequence  $A_{\epsilon}^n$  denotes the set of sequences whose probability is close to the entropy of the random variable

$$A_{\epsilon}^{(n)} = \{ \mathbf{x} \in \mathcal{X}^n \mid |-\frac{1}{n} \log p_{\mathbf{X}}(\mathbf{x}) - H(X)| < \epsilon \}$$

**Theorem 3.2** (Asymptotic equipartition property). Let  $X \sim p_X$  and let  $\mathbf{X} \in \mathcal{X}^n$  denote i.i.d. sequence following the same distribution, then:

- 1.  $\lim_{n\to\infty} P[\mathbf{X} \in A_{\epsilon}^{(n)}] = 1$
- 2.  $|A_{\epsilon}^{(n)}| < 2^{n(H-\epsilon)}$
- 3. For sufficiently large n,  $|A_{\epsilon}^{(n)}| \ge (1 \epsilon)2^{n(H + \epsilon)}$

Proof. A sketch of the proof (1) is the direct consequence of the weak law of large numbers

- (2): begin with the sum of probability of all possible sequence is 1
- (3): begin with for sufficiently large n,  $P[\mathbf{x} \in A_{\epsilon}^{(n)}] \geq 1 \epsilon$

## 4 Data compressions

**Theorem 4.1** (Kraft inequality). A D-ary prefix code for a finite set of m symbols  $\mathcal{X}$  exists if and only if the lengths of the code words  $l_1, l_2, \ldots, l_m$  satisfy:

$$\sum_{x \in \mathcal{X}} D^{-l(x)} \le 1$$

Using the method of Lagrange multiplier, we can optimize the expected codeword length E[L] under the constraint fo Kraft inequality:

$$l_i = \lceil \log_D \frac{1}{p_i} \rceil$$

**Theorem 4.2** (Optimal code length).  $H_D(X) \leq E[L] \leq H_D(X) + 1$ 

**Theorem 4.3** (McMillan inequality). A D-ary uniquely decodable code exists if and only if

$$\sum_{i} D^{-l_i} \le 1$$

In practice, **Huffman code** offers the optimal noiseless compression:

Requires knowledge of source probabilities.

$$(x_3, p_1 = 0.4, p_2 = 0.2, p_3 = 0.15, p_7 = 0.1)$$
 $(x_3, p_1 = 0.4, p_2 = 0.2, p_3 = 0.15, p_7 = 0.1)$ 
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 $(x_4, p_1 = 0.15, p_1 = 0.15, p_1 = 0.15, p_1 = 0.15)$ 

to dewords are defined by poth from root to corresponding leaf.

## 5 Channel capacity

In this section we discuss how much information can be transmitted through a noisy channel

**Definition 5.1.** A discrete channel consists of a discrete set of words W, two sets of symbols X, Y, an encoder  $W \to X$ , a decoder  $Y \to W$ , and a channel described by the conditional distribution  $p_{Y|X}$ :

$$W \to X \to Y \to \hat{W}$$

For this section we are only concerned with **memoryless** channel: given  $X_n$ ,  $Y_n$  is independent of all other variables:

$$p_{\mathbf{Y}|\mathbf{X}} = \prod_{l=1}^{n} p_{Y_l|X_l}$$

For a fixed  $p_{Y|X}$ , information channel capacity is defined by

$$C^I = \max_{p_X} I(p_X; p_{Y|X})$$

#### 5.1 Channel rate

**Definition 5.2** ((M, n) code). An (M,n)-code consists of a collection of M words W: |W| = M, two sets of symbols X, Y, an encoder  $f: W \to X^n$ , and a decoder  $g: Y^n \to W$ 

**Definition 5.3** (Decoding error). For a single word  $w \in W$ , the **conditional probability of error** is defined by

$$\lambda_l = P[g(\mathbf{Y}) \neq l \mid \mathbf{X} = f(w)]$$

Denote the maximal probability of error across all words by

$$\lambda^{(n)} = \max_{w \in \mathcal{W}} \lambda_w$$

Denote the average probability of error across all words by

$$P_e^{(n)} = \frac{1}{M} \sum_{w \in \mathcal{W}} \lambda_w$$

The rate of an (M,n)-code denotes the number of bits of information transmitted per channel use:

$$R = \frac{\log(M)}{n}$$

**Definition 5.4** (Achievable rate). A rate R > 0 is **achievable** if there exists a sequence of  $(2^{nR}, n)$ -code such that:

$$\lim_{n \to \infty} \lambda^{(n)} = 0$$

The capacity of a channel is the supremum of all achievable rates

### 5.2 Jointly typical sequence

Let  $(\mathbf{X}, \mathbf{Y}) \in \mathcal{X}^n \times \mathcal{Y}^n$  be i.i.d. sequences drawn according to the joint distribution  $P_{X,Y}$ , then by the weak law of large numbers,  $\lim_{n\to\infty} -\frac{1}{n} \log(p_{X,Y}(\mathbf{x},\mathbf{y})) = H(X,Y)$ . We can define notion of typicality on the joint sequence similar to the notion of typicality on individual sequences

**Definition 5.5** (Jointly typical sequence). The set of jointly typical sequence  $A_{\epsilon}^{(n)}$  is the set of  $(\mathbf{x}, \mathbf{y})$  such that:

- 1.  $\left|-\frac{1}{n}\log(p_X(\mathbf{x})) H(X)\right| \le \epsilon$
- 2.  $\left|-\frac{1}{n}\log(p_Y(\mathbf{y})) H(Y)\right| \le \epsilon$
- 3.  $\left|-\frac{1}{n}\log(p_{X,Y}(\mathbf{x},\mathbf{y})) H(X,Y)\right| \le \epsilon$

In other words,  $(\mathbf{x}, \mathbf{y})$  is jointly typical if  $\mathbf{x}$  and  $\mathbf{y}$  are each individually typical according to their marginal distribution, and  $(\mathbf{x}, \mathbf{y})$  is typical according to the joint distribution

**Theorem 5.1** (Joint asymptotic equipartition property). Let  $(\mathbf{X}, \mathbf{Y}) \in \mathcal{X}^n \times \mathcal{Y}^n$  be i.i.d. sequences drawn according to  $p_{X,Y}$ , then

- 1.  $\lim_{n\to\infty} P[(\mathbf{X}, \mathbf{Y}) \in A_{\epsilon}^{(n)}] = 1$
- 2.  $|A_{\epsilon}^{(n)}| \leq 2^{n(H(X,Y)+\epsilon)}$
- 3. For sufficiently large n:

$$|A_{\epsilon}^{(n)}| \ge (1 - \epsilon)2^{n(H(X,Y) - \epsilon)}$$

4. Let  $\tilde{\mathbf{X}}$  and  $\tilde{\mathbf{Y}}$  be drawn independently from  $p_X$  and  $p_Y$ , then

$$P[(\tilde{\mathbf{X}}, \tilde{\mathbf{Y}}) \in A_{\epsilon}^n] \le 2^{-n(I(X;Y) - 3\epsilon)}$$

5. Let  $\tilde{\mathbf{X}}$  and  $\tilde{\mathbf{Y}}$  be drawn independently from  $p_X$  and  $p_Y$ , then for sufficiently large n:

$$P[(\tilde{\mathbf{X}}, \tilde{\mathbf{Y}}) \in A_{\epsilon}^n] \ge (1 - \epsilon)2^{-n(I(X;Y) + 3\epsilon)}$$

*Proof.* (1) is the direct consequence of the weak law of large numbers (2):

$$1 = \sum_{(\mathbf{x}, \mathbf{y}) \in \mathcal{X}^n \times \mathcal{Y}^n} p(\mathbf{x}, \mathbf{y})$$

$$\geq \sum_{(\mathbf{x}, \mathbf{y}) \in A_{\epsilon}^{(n)}} p(\mathbf{x}, \mathbf{y})$$

$$\geq \sum_{(\mathbf{x}, \mathbf{y}) \in A_{\epsilon}^{(n)}} 2^{-n(H(X, Y) + \epsilon)}$$

$$= |A_{\epsilon}^{(n)}| 2^{-n(H(X, Y) + \epsilon)}$$

(3): for sufficiently large n,  $P[(\mathbf{X}, \mathbf{Y}) \in A_{\epsilon}^{(n)}] \geq 1 - \epsilon$ . On the other hand:

$$\begin{split} P[(\mathbf{X}, \mathbf{Y}) \in A_{\epsilon}^{(n)}] &= \sum_{(\mathbf{x}, \mathbf{y}) \in A_{\epsilon}} P[(\mathbf{X}, \mathbf{Y}) = (\mathbf{x}, \mathbf{y})] \\ &\leq \sum_{(\mathbf{x}, \mathbf{y}) \in A_{\epsilon}} 2^{-n(H_{X,Y} - \epsilon)} \\ &= |A_{\epsilon}| \cdot 2^{-n(H_{X,Y} - \epsilon)} \end{split}$$

Putting the two inequalities above together:

$$|A_{\epsilon}| \cdot 2^{-n(H_{X,Y} - \epsilon)} \ge 1 - \epsilon$$

(4): consider the probability of  $(\tilde{\mathbf{X}}, \tilde{\mathbf{Y}})$  falling into the jointly typical set:

$$\begin{split} P[(\tilde{\mathbf{X}}, \tilde{\mathbf{Y}}) \in A_{\epsilon}^{(n)}] &= \sum_{(\tilde{\mathbf{x}}, \tilde{\mathbf{y}}) \in A_{\epsilon}^{(n)}} P[\tilde{\mathbf{X}} = \tilde{\mathbf{x}}, \tilde{\mathbf{Y}} = \tilde{\mathbf{y}}] \\ &= \sum_{(\tilde{\mathbf{x}}, \tilde{\mathbf{y}}) \in A_{\epsilon}^{(n)}} P[\tilde{\mathbf{X}} = \tilde{\mathbf{x}}] \cdot P[\tilde{\mathbf{Y}} = \tilde{\mathbf{y}}] \\ &\leq \sum_{(\tilde{\mathbf{x}}, \tilde{\mathbf{y}}) \in A_{\epsilon}^{(n)}} 2^{-n(H_X - \epsilon)} \cdot 2^{-n(H_Y - \epsilon)} \\ &= |A| \cdot 2^{-n(H_X - \epsilon)} \cdot 2^{-n(H_Y - \epsilon)} \\ &< 2^{n(H(X, Y) + \epsilon)} \cdot 2^{-n(H_X - \epsilon)} \cdot 2^{-n(H_Y - \epsilon)} \end{split}$$

(5): Using similar logic as shown above:

$$P[(\tilde{\mathbf{X}}, \tilde{\mathbf{Y}}) \in A_{\epsilon}^{(n)}] = \sum_{(\tilde{\mathbf{x}}, \tilde{\mathbf{y}}) \in A_{\epsilon}^{(n)}} P[\tilde{\mathbf{X}} = \tilde{\mathbf{x}}, \tilde{\mathbf{Y}} = \tilde{\mathbf{y}}]$$

$$= \sum_{(\tilde{\mathbf{x}}, \tilde{\mathbf{y}}) \in A_{\epsilon}^{(n)}} P[\tilde{\mathbf{X}} = \tilde{\mathbf{x}}] \cdot P[\tilde{\mathbf{Y}} = \tilde{\mathbf{y}}]$$

$$\geq \sum_{(\tilde{\mathbf{x}}, \tilde{\mathbf{y}}) \in A_{\epsilon}^{(n)}} 2^{-n(H_X + \epsilon)} \cdot 2^{-n(H_Y + \epsilon)}$$

$$= |A_e| 2^{-n(H_X + \epsilon)} \cdot 2^{-n(H_Y + \epsilon)}$$

$$\geq (1 - \epsilon) 2^{n(H_X, Y - \epsilon)} \cdot 2^{-n(H_X + \epsilon)} \cdot 2^{-n(H_Y + \epsilon)}$$

#### 5.3 Channel coding theorem

**Theorem 5.2.** A rate R is achievable if and only if it is below the information channel capacity

Proof. This is a sketch of proof

First we show that rate below information capacity is achievable: Let R be a rate such that R < C, then we can construct a  $(2^{nR}, n)$ -code through the following procedure:

- 1. Find some distribution of  $p_X$  and  $\epsilon > 0$  such that  $R = I(p_X; p_{Y|X}) 4\epsilon$
- 2. For each  $w \in \{1, \ldots, 2^{nR}\}$ , the encoding is  $f(w) = (x_1(w), x_2(w), \ldots, x_n(w))$  where  $x_i(w)$  is i.i.d. sampled from  $p_X$
- 3. Transmit each symbol  $x_i(w)$  according to  $p_{Y|X}$
- 4. Upon receiving  $\mathbf{y}$ , find some  $\hat{w} \in \{1, \dots, 2^{nR}\}$  such that  $(\mathbf{x}, \mathbf{y})$  is jointly typical according to  $p_X$  and  $p_{Y|X}$  (which can be used to compute  $p_{X,Y}$  and  $p_Y$ )

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For each word w, decoding error occurs if and only if one of two scenarios occurs:

- 1.  $(\mathbf{x}(w), \mathbf{y})$  is not jointly typical, in which case no appropriate  $\hat{w}$  can be found in the decoding step
- 2. there is another  $w' \neq w$  such that  $(\mathbf{x}(w'), \mathbf{y})$  is also jointly typical

The probability of (1) converges to 0 because  $\lim_{n\to\infty} P[(\mathbf{x},\mathbf{y}) \notin A_{\epsilon}^{(n)}] = 0$ .

The probability of (2) converges to 0 because  $\mathbf{y}$  is independent of  $\mathbf{x}(w')$ . By joint AEP, the probability of independently sampled sequences being jointly typical is at most  $2^{-n(I(X;Y)-3\epsilon)}$ , hence:

$$\begin{split} \sum_{w' \neq w} P[(\mathbf{x}(w'), \mathbf{y}) \in A_{\epsilon}^{(n)}] &= (2^{nR} - 1) P[(\mathbf{x}(w'), \mathbf{y}) \in A_{\epsilon}^{(n)}] \\ &\leq (2^{nR} - 1) 2^{-n(I(X;Y) - 3\epsilon)} \\ &\leq 2^{nR - n(I_{X;Y} - 3\epsilon)} \\ &= 2^{n(I_{X;Y} - 4\epsilon) - n(I_{X;Y} - 3\epsilon)} \\ &= 2^{-n\epsilon} \end{split}$$

Which also converges to 0 as  $n \to \infty$ 

Converse: achievable rates are below information capacity

Recall that the channel represents a Markov chain  $W \to \mathbf{X} \to \mathbf{Y} \to \hat{W}$ . By Fano's inequality, we know that:

$$H(W \mid \hat{W}) \le h(p_e) + p_e \cdot \log(|\mathcal{W}|)$$

Consider the mutual information between W and  $\hat{W}$ :

$$I(W; \hat{W}) = H(W) - H(W \mid \hat{W})$$

Assuming that W is uniformly distributed,  $H(W) = \log(|\mathcal{W}|) = nR$ .

By the data-processing inequality  $I(W; \hat{W}) \leq I(\mathbf{X}; \mathbf{Y})$ . By the chain rule of mutual information:  $I(\mathbf{X}; \mathbf{Y}) \leq nI(X; Y)$ . By the definition of information channel capacity  $I(X; Y) \leq C$ . Putting everything together, we have

$$nR - nC \le p_e$$

Assuming the rate is achievable, then  $\lim_{n\to\infty} p_e = 0$ , which implies  $R \leq C$ .

## 6 Differential entropy

**Definition 6.1.** Let  $X \in \mathbb{R}$  be a random variable with probability density function  $f_X$ . The differential entropy of X is defined by

$$h(X) = -\int_{x \in \mathbb{R}} f_X(x) \ln(f_X(x)) dx = -E[\ln(f_X(X))]$$

Some notable results:

- For uniform distribution over [0, a],  $f_X(x) = \frac{1}{a}$ ,  $h(X) = \ln(a)$
- For  $X \sim N(0, \sigma^2), h(X) = \frac{1}{2} \ln(2\pi e \sigma^2)$
- For multivariate normal  $\mathbf{X} \sim N(\mathbf{0}, K), h(\mathbf{X}) = \frac{1}{2} \log((2\pi e)^n \det K)$
- For some constant a, h(X+a) = h(X)
- For some constant a,  $h(aX) = \ln(|a|) + h(X)$

**Theorem 6.1.** Let X be continuous random variable with 0 mean and  $\sigma^2$  variance, then

$$h(X) \le \frac{1}{2} \ln(2\pi e \sigma^2)$$

Equality is reached if and only if X is Gaussian

#### 7 Gaussian channel

**Definition 7.1** (Information channel capacity). Let Y = X + Z, where  $Z \stackrel{\$}{\leftarrow} N(0, \sigma^2)$  and  $E[X^2] \leq P$  for some power level constraint P. The **information channel capacity** is defined by

$$C^{I} = \max_{f_X: E[X^2] \le P} I(X;Y)$$

Theorem 7.1. The information channel capacity of a Gaussian channel is

$$\max_{f_X: E[X^2] < P} I(X;Y) = \frac{1}{2} \log \left(1 + \frac{P}{\sigma^2}\right) \tag{1}$$

Where P is the power constraint, and  $\sigma^2$  is the variance of the Gaussian noise. The maximum is achieved when X follows Gaussian distribution  $X \sim N(0, P)$ 

#### 7.1 Parallel Gaussian Channel

Suppose for  $1 \le l \le n$ ,  $Y_l = X_l + Z_l$ , where  $Z_l \sim N(0, N_l)$  is independent Gaussian noise, then  $\mathbf{Y} = (Y_l)_{l=1}^n$  can be seen as the output of n parallel Gaussian channels combined into a single channel. The capacity of the combined channel is

$$C = \max_{f_{\mathbf{X}}} I(\mathbf{X}; \mathbf{Y})$$

The optimal strategy for maximizing the capacity is to make  $X_l$  independent and individually Gaussian. If there is a combined power constraint  $\sum_{l=1}^{n} E[X_l^2] \leq P$ , then power should be first given to the channel with the lowest amount of noise until the combined noise + power exceeds the next lowest noise (waterfilling).

# 8 Rate distortion theory

**Definition 8.1.** For some fixed  $p_X$ , the information rate distortion is defined by

$$R^I(D) = \min_{p_{\hat{X}|X}: E[d(X,\hat{X})] \le D} I(X; \hat{X})$$

**Theorem 8.1.** Let X follow Bernoulli(p), then

$$R(D) = \begin{cases} h(p) - h(D) & When \ D < \min(p, 1 - p) \\ 0 & otherwise \end{cases}$$

**Theorem 8.2.** For  $X \stackrel{\$}{\leftarrow} N(0, \sigma^2)$ :

$$R(D) = \begin{cases} \frac{1}{2} \log(\frac{\sigma^2}{D}) & (D \le \sigma^2) \\ 0 & otherwise \end{cases}$$