# Homework 2

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October 24, 2021

#### 1 P1

#### 1.1 a

In this question, the RSS for the linear regression is:

$$RSS = \sum_{i=1}^{n} \left( y_i - \hat{\beta}_0 - \sum_{j=1}^{p} \hat{\beta}_j x_{ij} \right)^2$$

$$= \left( \hat{\beta}_0 + \hat{\beta}_1 x_{11} + \hat{\beta}_2 x_{12} \right)^2 + \left( \hat{\beta}_0 + \hat{\beta}_1 x_{21} + \hat{\beta}_2 x_{22} \right)^2$$
(1.1)

So the Ridge Regression Optimization is:

$$\min RSS + \lambda \sum_{j=1}^{p} \beta_{j}^{2} = \sum_{i=1}^{n} \left( y_{i} - \hat{\beta}_{0} - \sum_{j=1}^{p} \hat{\beta}_{j} x_{ij} \right)^{2} + \lambda \sum_{j=1}^{p} \hat{\beta}_{j}^{2}$$

$$= \left( \hat{\beta}_{0} + \hat{\beta}_{1} x_{11} + \hat{\beta}_{2} x_{12} \right)^{2} + \left( \hat{\beta}_{0} + \hat{\beta}_{1} x_{21} + \hat{\beta}_{2} x_{22} \right)^{2} + \lambda \left( \hat{\beta}_{1}^{2} + \hat{\beta}_{2}^{2} \right)$$

$$(1.2)$$

# 1.2 b

For the optimization result:

$$\frac{\partial f}{\partial \beta_{1}} = -2x_{11} \left( y_{1} - \beta_{1} x_{11} - \beta_{2} x_{12} \right) - 2x_{21} \left( y_{2} - \beta_{1} x_{21} - \beta_{2} x_{22} \right) + 2\lambda \beta_{1} = 0$$

$$\frac{\partial f}{\partial \beta_{2}} = -2x_{12} \left( y_{1} - \beta_{1} x_{11} - \beta_{2} x_{12} \right) - 2x_{22} \left( y_{2} - \beta_{1} x_{21} - \beta_{2} x_{22} \right) + 2\lambda \beta_{2} = 0$$
(1.3)

Denote that  $x_1 = x_{11} = x_{21}$ ,  $x_2 = x_{12} = x_{22}$ . To solve the two equations, we need to make some transform:

$$(x_{1}^{2} + x_{2}^{2} + \lambda) \hat{\beta}_{1} + (x_{1}^{2} + x_{2}^{2}) \hat{\beta}_{2} = x_{1} y_{1} + x_{2} y_{2}$$

$$(x_{1}^{2} + x_{2}^{2}) \hat{\beta}_{1} + (x_{1}^{2} + x_{2}^{2} + \lambda) \hat{\beta}_{2} = x_{1} y_{1} + x_{2} y_{2}$$

$$(x_{1}^{2} + x_{2}^{2} + \lambda) \hat{\beta}_{1} + (x_{1}^{2} + x_{2}^{2}) \hat{\beta}_{2} = (x_{1}^{2} + x_{2}^{2}) \hat{\beta}_{1} + (x_{1}^{2} + x_{2}^{2} + \lambda) \hat{\beta}_{2}$$

$$(x_{1}^{2} + x_{2}^{2} + \lambda) (\hat{\beta}_{1} - \hat{\beta}_{2}) + (x_{1}^{2} + x_{2}^{2}) (\hat{\beta}_{2} - \hat{\beta}_{1}) = 0$$

$$\lambda \hat{\beta}_{1} = \lambda \hat{\beta}_{2}$$

$$(1.4)$$

Since  $\lambda \neq 0$ ,  $\hat{\beta}_1 = \hat{\beta}_2$ 

#### 1.3 c

LASSO Regression Optimization is:

$$\min RSS + \lambda \sum_{j=1}^{p} |\beta_{j}| = \sum_{i=1}^{n} \left( y_{i} - \hat{\beta}_{0} - \sum_{j=1}^{p} \hat{\beta}_{j} x_{ij} \right)^{2} + \lambda \sum_{j=1}^{p} |\hat{\beta}_{j}|$$

$$= \left( \hat{\beta}_{0} + \hat{\beta}_{1} x_{11} + \hat{\beta}_{2} x_{12} \right)^{2} + \left( \hat{\beta}_{0} + \hat{\beta}_{1} x_{21} + \hat{\beta}_{2} x_{22} \right)^{2} + \lambda \left( |\hat{\beta}_{1}| + |\hat{\beta}_{2}| \right)$$

$$(1.5)$$

#### 1.4 d

For the optimization result:

$$\frac{\partial f}{\partial \beta_{1}} = -2x_{11} \left( y_{1} - \beta_{1} x_{11} - \beta_{2} x_{12} \right) - 2x_{21} \left( y_{2} - \beta_{1} x_{21} - \beta_{2} x_{22} \right) \pm \lambda = 0$$

$$\frac{\partial f}{\partial \beta_{2}} = -2x_{12} \left( y_{1} - \beta_{1} x_{11} - \beta_{2} x_{12} \right) - 2x_{22} \left( y_{2} - \beta_{1} x_{21} - \beta_{2} x_{22} \right) \pm \lambda = 0$$
(1.6)

Denote that  $x_1 = x_{11} = x_{21}$ ,  $x_2 = x_{12} = x_{22}$ . To solve the two equations, we need to make some transform:

$$(x_1^2 + x_2^2)\hat{\beta}_1 + (x_1^2 + x_2^2)\hat{\beta}_2 = x_1y_1 + x_2y_2 \pm \lambda$$
  

$$(x_1^2 + x_2^2)\hat{\beta}_1 + (x_1^2 + x_2^2)\hat{\beta}_2 = x_1y_1 + x_2y_2 \pm \lambda$$
(1.7)

Only when  $\hat{\beta_1} > 0$  and  $\hat{\beta_2} > 0$  or  $\hat{\beta_1} < 0$  and  $\hat{\beta_2} < 0$ , there will be solutions for this problem. Due to the two equations are same in form, there will be more than one solution for this problem. if  $\hat{\beta_1} > 0$  and  $\hat{\beta_2} > 0$ , the equation will be:

$$(x_1^2 + x_2^2)\hat{\beta}_1 + (x_1^2 + x_2^2)\hat{\beta}_2 = x_1y_1 + x_2y_2 - \lambda$$
(1.8)

Otherwise, the equation will be:

$$(x_1^2 + x_2^2)\hat{\beta}_1 + (x_1^2 + x_2^2)\hat{\beta}_2 = x_1y_1 + x_2y_2 + \lambda$$
(1.9)

#### 2.1 a

When p=1, the optimization function is  $(y-\beta_1)^2+\lambda\beta_1^2=(1+\lambda)\beta_1^2-2y\beta_1+y^2$ . To find the  $\beta_1$  that minimize the y, we need to calculate  $2(1+\lambda)\beta_1-2y=0$  We can plot the image:

```
1  y = 2
2  lambda = 1
3  beta = seq(-5,5,0.01)
4
5  plot(beta,(y-beta)^2 + lambda*beta^2,type="l",xlab = "beta"
6  ,ylab = "Optimization_function",xlim=c(-5,5))
7
8  min_x = y/(1+lambda)
9  min_y = (y-min_x)^2 + lambda*min_x^2
10  points(min_x,min_y,col="red")
11
12  title("Optimization_Result_for_Ridge_Regression")
```

#### Optimization Result for Ridge Regression

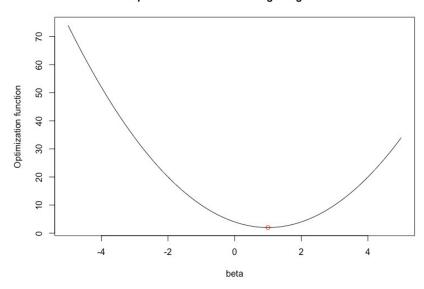


Figure 2.1: Screenshot for 'Ridge Regression'

#### 2.2 b

When p=1, the optimization function is  $(y-\beta_1)^2+\lambda|\beta_1|=\beta_1^2-(2y\pm\lambda)\beta_1+y^2$ . To find the  $\beta_1$  that minimize the y, we need to calculate  $2\beta_1-(2y\pm\lambda)=0$ . If  $\beta_1\geqslant 0$ , that's  $2\beta_1-(2y-\lambda)=0$ . Otherwise, that's  $2\beta_1-(2y+\lambda)=0$ .

We can plot the image for all three conditions:

```
1  y = 2
2  lambda = 3
3  beta = seq(-5,5,0.01)
```

```
4
5 plot(beta,(y-beta)^2 + lambda*abs(beta),type="l",xlab = "beta"
   ,ylab = "Optimization ufunction",xlim=c(-5,5))
  if(y > lambda/2) {
8
     min_x1 = y - lambda/2
9
10 } else {
      if(y < -lambda/2){
11
12
        min_x1 = y + lambda/2
      } else {
13
14
        min_x1 = 0
      }
15
   }
16
17
18 \min_{y_1} = (y_{\min_x_1})^2 + lambda*abs(\min_x_1)
19 points(min_x1,min_y1,col="red")
20
21 title("Optimization_{\square}Result_{\square}for_{\square}LASSO_{\square}Regression")
```

#### Optimization Result for LASSO Regression

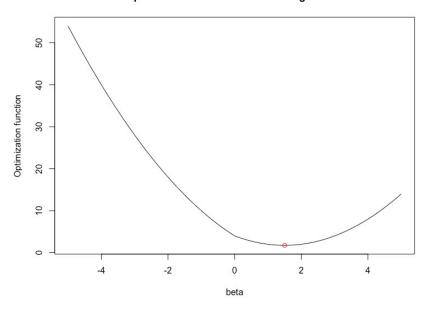


Figure 2.2: Screenshot for 'LASSO Regression'

## Optimization Result for LASSO Regression

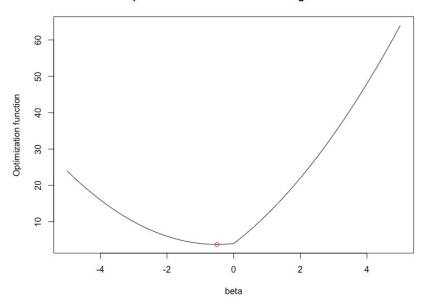


Figure 2.3: Screenshot for 'LASSO Regression'

# Optimization Result for LASSO Regression

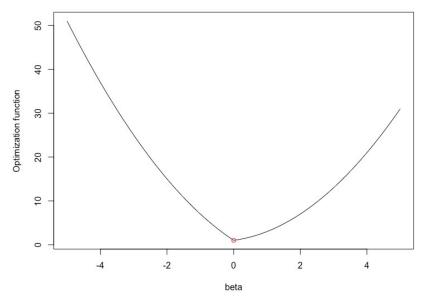


Figure 2.4: Screenshot for 'LASSO Regression'

#### 3.1 a

We know that  $y_i \sim \mathcal{N}(\beta_0 + \sum_{j=1}^p x_{ij}\beta_j, \sigma^2)$ , so the likelihood function is:

$$\mathcal{L}(X, \beta, \sigma \mid Y) = \prod_{i=1}^{n} \frac{1}{\sigma \sqrt{2\pi}} e^{-\frac{\left[y_{i} - \left(\beta_{0} + \sum_{j=1}^{p} \beta_{j} x_{ij}\right)\right]^{2}}{2\sigma^{2}}}$$

$$= \frac{1}{(\sigma \sqrt{2\pi})^{n}} e^{-\frac{1}{2\sigma^{2}} \sum_{i=1}^{n} \left(y_{i} - \left(\beta_{0} + \sum_{j=1}^{p} \beta_{j} x_{ij}\right)\right)^{2}}$$
(3.1)

#### 3.2 b

We know that posterior  $\propto \mathcal{L} \times \text{prior}$ , and  $\text{prior}(\beta) = \frac{1}{2b} e^{(-\frac{\sum_{j=1}^{p} |\beta_j|}{b})}$ , thus:

$$posterior \propto \frac{1}{(\sigma\sqrt{2\pi})^{n}} e^{\left(-\frac{1}{2\sigma^{2}}\sum_{i=1}^{n}(y_{i}-\sum_{j=1}^{p}x_{ij}\beta_{j}-\beta_{0})^{2}\right)} * \frac{1}{2b} e^{\left(-\frac{\sum_{j=1}^{p}|\beta_{j}|}{b}\right)}$$

$$\propto \frac{1}{2b} * \frac{1}{(\sigma\sqrt{2\pi})^{n}} e^{\left(-\frac{1}{2\sigma^{2}}\sum_{i=1}^{n}(y_{i}-\sum_{j=1}^{p}x_{ij}\beta_{j}-\beta_{0})^{2}-\frac{\sum_{j=1}^{p}|\beta_{j}|}{b}\right)}$$
(3.2)

#### 3.3 c

To prove the argument, we need to prove that the maximum of the posterior is given by the LASSO regression with  $\lambda$ :

$$\log(\text{posterior}) = \log(\frac{1}{2b} * \frac{1}{(\sigma\sqrt{2\pi})^n}) + \left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \sum_{j=1}^p x_{ij}\beta_j - \beta_0)^2 - \frac{\sum_{j=1}^p |\beta_j|}{b}\right)$$
(3.3)

To find the maximum for the posterior, we need to find the minimum for the  $\left(\frac{1}{2\sigma^2}\sum_{i=1}^n(y_i-\sum_{j=1}^px_{ij}\beta_j-\beta_0)^2+\frac{\sum_{j=1}^p|\beta_j|}{b}\right):$ 

$$\operatorname{arg min}_{\beta} \left( \frac{1}{2\sigma^{2}} \sum_{i=1}^{n} (y_{i} - \sum_{j=1}^{p} x_{ij}\beta_{j} - \beta_{0})^{2} + \frac{\sum_{j=1}^{p} |\beta_{j}|}{b} \right) = \operatorname{arg min}_{\beta} \left( \sum_{i=1}^{n} (y_{i} - \sum_{j=1}^{p} x_{ij}\beta_{j} - \beta_{0})^{2} + \frac{2\sigma^{2} \sum_{j=1}^{p} |\beta_{j}|}{b} \right) \\
= \operatorname{arg min}_{\beta} \left( \sum_{i=1}^{n} (y_{i} - \sum_{j=1}^{p} x_{ij}\beta_{j} - \beta_{0})^{2} + \frac{2\sigma^{2} \sum_{j=1}^{p} |\beta_{j}|}{b} \right) \\
(3.4)$$

Let  $\lambda = \frac{2\sigma^2}{b}$ , that's the same form with LASSO Regression.

#### 3.4 d

$$\operatorname{prior}(\beta) = \prod_{j=1}^{p} \mathbf{P}(\beta_{j})$$

$$= \prod_{j=1}^{p} \frac{1}{\sqrt{2\pi c}} e^{\left(-\frac{\beta_{j}^{2}}{2c}\right)}$$

$$= \frac{1}{(\sqrt{2\pi c})^{p}} e^{\left(-\frac{1}{2c} \sum_{j=1}^{p} \beta_{j}^{2}\right)}$$
(3.5)

 $\mathsf{posterior} \propto \mathscr{L} \times \mathsf{prior}$ 

$$\propto \frac{1}{(\sigma\sqrt{2\pi})^{n}} e^{\left(-\frac{1}{2\sigma^{2}}\sum_{i=1}^{n}(y_{i}-\sum_{j=1}^{p}x_{ij}\beta_{j}-\beta_{0})^{2}\right)} * \frac{1}{(\sqrt{2\pi}c)^{p}} e^{\left(-\frac{1}{2c}\sum_{j=1}^{p}\beta_{j}^{2}\right)}$$

$$\propto \frac{1}{(\sigma\sqrt{2\pi})^{n}} * \frac{1}{(\sqrt{2\pi}c)^{p}} e^{\left(-\frac{1}{2\sigma^{2}}\sum_{i=1}^{n}(y_{i}-\sum_{j=1}^{p}x_{ij}\beta_{j}-\beta_{0})^{2}-\frac{1}{2c}\sum_{j=1}^{p}\beta_{j}^{2}\right)}$$
(3.6)

#### 3.5 e

To prove the argument, we need to prove that the maximum of the posterior is given by the Ridge Regression with  $\lambda$ :

log(posterior) = log(
$$\frac{1}{2b} * \frac{1}{(\sigma\sqrt{2\pi})^n}$$
) +  $\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (y_i - \sum_{j=1}^p x_{ij}\beta_j - \beta_0)^2 - \frac{1}{2c} \sum_{j=1}^p \beta_j^2\right)$  (3.7)

To find the maximum for the posterior, we need to find the minimum for the  $\left(\frac{1}{2\sigma^2}\sum_{i=1}^n(y_i-\sum_{j=1}^px_{ij}\beta_j-\beta_0)^2+\frac{1}{2c}\sum_{j=1}^p\beta_j^2\right)$ :

$$\operatorname{arg\,min}_{\beta} \left( \frac{1}{2\sigma^{2}} \sum_{i=1}^{n} (y_{i} - \sum_{j=1}^{p} x_{ij}\beta_{j} - \beta_{0})^{2} + \frac{1}{2c} \sum_{j=1}^{p} \beta_{j}^{2} \right) = \operatorname{arg\,min}_{\beta} \left( \sum_{i=1}^{n} (y_{i} - \sum_{j=1}^{p} x_{ij}\beta_{j} - \beta_{0})^{2} + \frac{2\sigma^{2} \sum_{j=1}^{p} |\beta_{j}|}{2c} \right) \\
= \operatorname{arg\,min}_{\beta} \left( \sum_{i=1}^{n} (y_{i} - \sum_{j=1}^{p} x_{ij}\beta_{j} - \beta_{0})^{2} + \frac{\sigma^{2}}{c} \sum_{j=1}^{p} |\beta_{j}| \right) \tag{3.8}$$

Let  $\lambda = \frac{\sigma^2}{c}$ , that's the same form with Ridge Regression.

#### 4.1 a

```
1 set.seed(1)
2 \times = rnorm(100)
3 \text{ eps} = rnorm(100)
   4.2 b
1 beta = c(1, -1, 2, 0.5)
2 y = beta[1] + beta[2]*x + beta[3]*x^2 + beta[4]*x^3 + eps
   4.3 c
1 library(leaps)
2 ans.df = data.frame(y = y, x = x)
3 ans.model = regsubsets(y ~ poly(x, 10, raw = T),
                            data = ans.df, nvmax = 10)
5 ans.summary = summary(ans.model)
7 min_p = which.min(ans.summary$cp)
8 plot(ans.summary$cp, xlab = "Subset_size", ylab = "Cp",
        col = "blue", pch = 20, type = "o")
10 points(min_p, ans.summary$cp[min_p], col = "red", lwd = 3)
11 title("Best_Subset_with_Cp")
12 coefficients(ans.model, id = which.min(ans.summary$cp))
13
14 min_p = which.min(ans.summary$bic)
15 plot(ans.summary$bic, xlab = "Subset_size", ylab = "BIC",
        col = "blue", pch = 20, type = "o")
17 points(min_p, ans.summary$bic[min_p], col = "red", lwd = 3)
18 title("Best_Subset_with_BIC")
19 coefficients(ans.model, id = which.min(ans.summary$bic))
20
21 max_p = which.max(ans.summary$adjr2)
22 plot(ans.summary$adjr2, xlab = "Subsetusize", ylab = "AdjR2",
        col = "blue", pch = 20, type = "o")
24 points(max_p, ans.summary$adjr2[max_p], col = "red", lwd = 3)
25 title("Best_Subset_with_AdjR2")
26 coefficients(ans.model, id = which.max(ans.summary$adjr2))
   The best model uses a X^5 term. |\beta_0 - \hat{\beta}_0| = 0.072, |\beta_1 - \hat{\beta}_1| = 0.445, |\beta_2 - \hat{\beta}_2| = 0.157
```

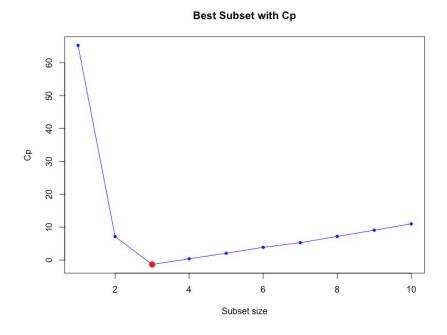


Figure 4.1: Screenshot for *'Best Subset with Cp'* 

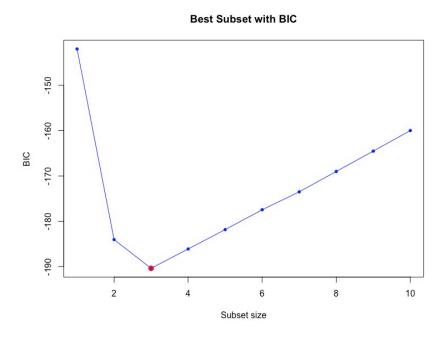


Figure 4.2: Screenshot for 'Best Subset with BIC'

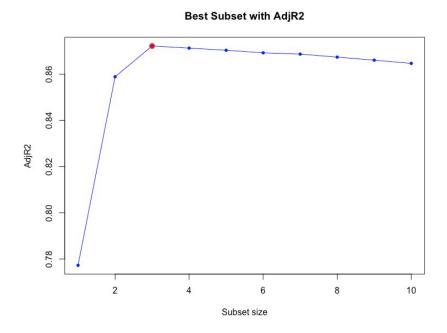


Figure 4.3: Screenshot for 'Best Subset with Adjust  $\mathbb{R}^2$ '

Figure 4.4: Screenshot for 'Coefficients'

#### 4.4 d

If we use forward selection:

```
1 library(leaps)
2 ans.df = data.frame(y = y, x = x)
3 ans.model = regsubsets(y ~ poly(x, 10, raw = T),
                          data = ans.df, nvmax = 10, method = "forward")
5 ans.summary = summary(ans.model)
6
7 min_p = which.min(ans.summary$cp)
8 plot(ans.summary$cp, xlab = "Subset_size", ylab = "Cp",
        col = "blue", pch = 20, type = "o")
10 points(min_p, ans.summary$cp[min_p], col = "red", lwd = 3)
11 title("Best Subset with Cp using 'forward'")
12 coefficients(ans.model, id = which.min(ans.summary$cp))
13
14 min_p = which.min(ans.summary$bic)
15 plot(ans.summary$bic, xlab = "Subset_size", ylab = "BIC",
        col = "blue", pch = 20, type = "o")
16
17 points(min_p, ans.summary$bic[min_p], col = "red", lwd = 3)
18 title("Best_Subset_with_BIC_using_'forward'")
19 coefficients(ans.model, id = which.min(ans.summary$bic))
20
21 max_p = which.max(ans.summary$adjr2)
22 plot(ans.summary$adjr2, xlab = "Subset_size", ylab = "AdjR2",
        col = "blue", pch = 20, type = "o")
24 points(max_p, ans.summary$adjr2[max_p], col = "red", lwd = 3)
25 title("Best Subset with AdjR2 using 'forward'")
26 coefficients(ans.model, id = which.max(ans.summary$adjr2))
```

Forward stepwise uses a  $X^5$  and a  $X^7$  term.  $\hat{\beta}_0$ ,  $\hat{\beta}_1$  and  $\hat{\beta}_2$  do not change a lot.

#### Best Subset with Cp using 'forward'

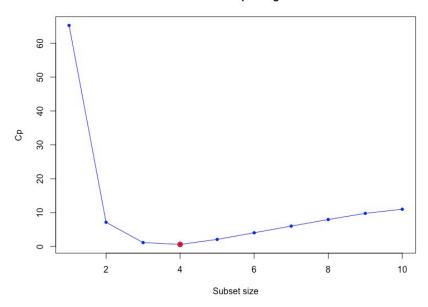


Figure 4.5: Screenshot for 'Best Subset with  $C_p$  using 'forward' '

# Best Subset with BIC using 'forward'

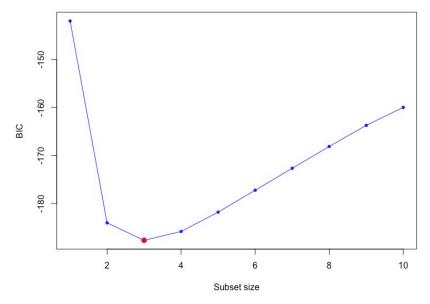


Figure 4.6: Screenshot for 'Best Subset with BIC using 'forward' '

# 

Figure 4.7: Screenshot for 'Best Subset with Adjust R<sup>2</sup> using 'forward' '

Subset size

Figure 4.8: Screenshot for 'Coefficients using 'forward' '

If we use backward selection:

```
1 library(leaps)
2 ans.df = data.frame(y = y, x = x)
3 ans.model = regsubsets(y ~ poly(x, 10, raw = T),
                          data = ans.df, nvmax = 10, method = "backward")
  ans.summary = summary(ans.model)
5
6
7 min_p = which.min(ans.summary$cp)
8
  plot(ans.summary$cp, xlab = "Subsetusize", ylab = "Cp",
        col = "blue", pch = 20, type = "o")
   points(min_p, ans.summary$cp[min_p], col = "red", lwd = 3)
10
  title("Best_Subset_with_Cp_using_'backward'")
12 coefficients(ans.model, id = which.min(ans.summary$cp))
13
14 min_p = which.min(ans.summary$bic)
15 plot(ans.summary$bic, xlab = "Subsetusize", ylab = "BIC",
16
        col = "blue", pch = 20, type = "o")
17 points(min_p, ans.summary$bic[min_p], col = "red", lwd = 3)
18 title("Best Subset with BIC using 'backward'")
19 coefficients(ans.model, id = which.min(ans.summary$bic))
20
21 max_p = which.max(ans.summary$adjr2)
22 plot(ans.summary$adjr2, xlab = "Subsetusize", ylab = "AdjR2",
        col = "blue", pch = 20, type = "o")
23
24 points(max_p, ans.summary$adjr2[max_p], col = "red", lwd = 3)
25 title("Best_Subset_with_AdjR2_using_'backward'")
26 coefficients(ans.model, id = which.max(ans.summary$adjr2))
```

#### Best Subset with Cp using 'backward'

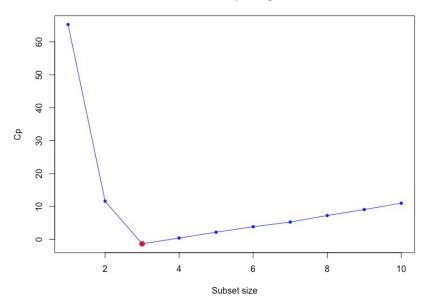


Figure 4.9: Screenshot for 'Best Subset with C<sub>p</sub> using 'backward''

Backward stepwise uses a  $X^5$  term.  $\hat{\beta}_0$ ,  $\hat{\beta}_1$  and  $\hat{\beta}_2$  do not change a lot.

# Best Subset with BIC using 'backward'

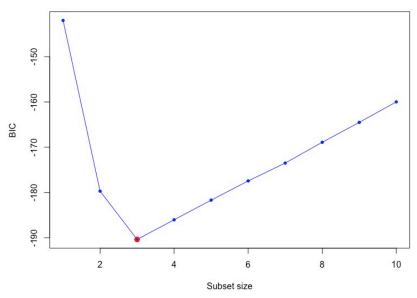


Figure 4.10: Screenshot for 'Best Subset with BIC using 'backward' '

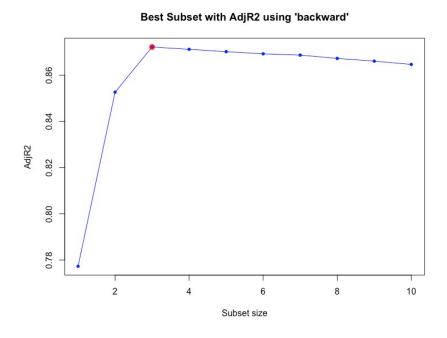


Figure 4.11: Screenshot for 'Best Subset with Adjust  $\mathbb{R}^2$  using 'backward' '

Figure 4.12: Screenshot for 'Coefficients using 'backward''

#### 4.5 e

```
library(glmnet)
xmat = model.matrix(y ~ poly(x, 10, raw = T), data = ans.df)[, -1]
ans.lasso = cv.glmnet(xmat, y, alpha = 1)
ans.lambda = ans.lasso$lambda.min
ans.lambda
plot(ans.lasso)

ans.lasso = glmnet(xmat, y, alpha = 1)
predict(ans.lasso, s = ans.lambda, type = "coefficients")
```

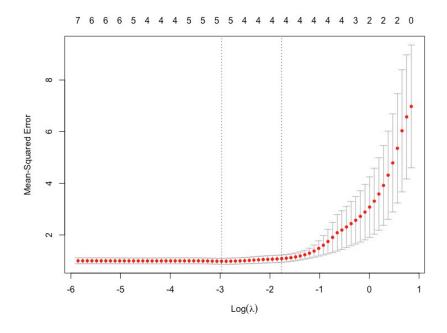


Figure 4.13: Screenshot for 'LASSO Cross-validation'

Figure 4.14: Screenshot for 'Coefficients'

LASSO uses a  $X^4$  and  $X^5$  term.  $\hat{\beta}_0$ ,  $\hat{\beta}_1$  and  $\hat{\beta}_2$  do not change a lot.

#### 4.6 f

```
1 beta 7 = 3
2 yy = beta[1] * beta_7*x^7 + eps
3 \text{ ans.df} = \text{data.frame}(y = yy, x = x)
4 ans.model = regsubsets(y ~ poly(x, 10, raw = T),
                          data = ans.df, nvmax = 10)
6 ans.summary = summary(ans.model)
8 min_p = which.min(ans.summary$cp)
9 plot(ans.summary$cp, xlab = "Subset_size", ylab = "Cp",
        col = "blue", pch = 20, type = "o")
11 points(min_p, ans.summary$cp[min_p], col = "red", lwd = 3)
12 title("Best Subset with Cp")
13 coefficients(ans.model, id = which.min(ans.summary$cp))
14
15 min_p = which.min(ans.summary$bic)
16 plot(ans.summary$bic, xlab = "Subset_size", ylab = "BIC",
17
        col = "blue", pch = 20, type = "o")
18 points(min_p, ans.summary$bic[min_p], col = "red", lwd = 3)
19 title("Best Subset with BIC")
20 coefficients(ans.model, id = which.min(ans.summary$bic))
21
22 max_p = which.max(ans.summary$adjr2)
23 plot(ans.summary$adjr2, xlab = "Subset_size", ylab = "AdjR2",
        col = "blue", pch = 20, type = "o")
25 points(max_p, ans.summary$adjr2[max_p], col = "red", lwd = 3)
26 title("Best_Subset_with_AdjR2")
27 coefficients(ans.model, id = which.max(ans.summary$adjr2))
```

Among all three criterion, BIC picks up the best estimation of the true model.

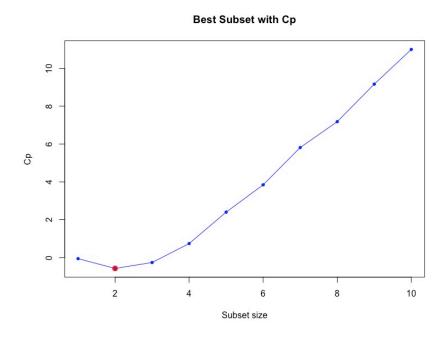


Figure 4.15: Screenshot for 'Best Subset with  $C_p$ '

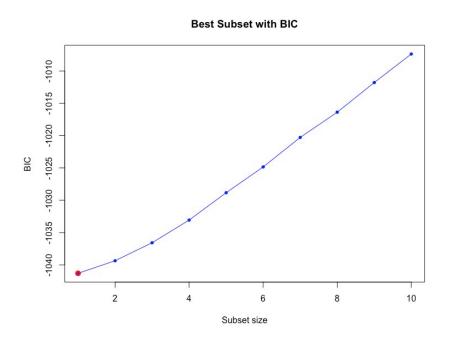


Figure 4.16: Screenshot for 'Best Subset with BIC'

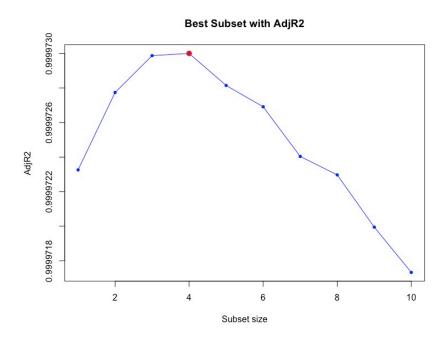


Figure 4.17: Screenshot for 'Best Subset with Adjust  $\mathbb{R}^2$ '

Figure 4.18: Screenshot for 'Coefficients'

```
1 xmat = model.matrix(y ~ poly(x, 10, raw = T), data = ans.df)[, -1]
2 ans.lasso = cv.glmnet(xmat, y, alpha = 1)
3 ans.lambda = ans.lasso$lambda.min
4 ans.lambda
5 plot(ans.lasso)
6
7 ans.lasso = glmnet(xmat, y, alpha = 1)
8 predict(ans.lasso, s = ans.lambda, type = "coefficients")
```

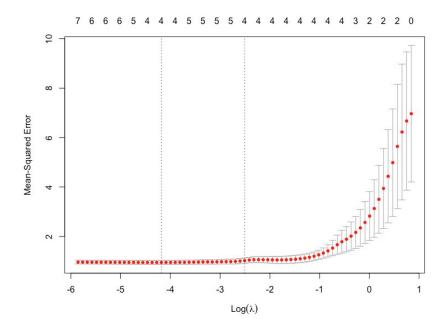


Figure 4.19: Screenshot for 'LASSO Cross-validation'

Figure 4.20: Screenshot for 'Coefficients'

LASSO gives a best estimation of the true model. Only one variable besides the Intercept is given in the estimation.

#### 5.1 a

```
library(ISLR)
set.seed(1)
train.size = nrow(College) / 2
train = sample(1: nrow(College), train.size)
test = -train
College.train = College[train, ]
College.test = College[test, ]

5.2 b

College.lm = lm(Apps~., data = College.train)
summary(College.lm)

College.pred = predict(College.lm, College.test)
College.error = mean((College.test$Apps - College.pred)^2)
```

```
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                                                                                                                                                                                                                                                                                                                                                                                                                                               -\Box

    R 4.1.1 · ~/Columbia/学习资料/E6690-Statistic Learning/HW2/

   > College.test = College[test, ]
> College.lm = lm(Apps~., data = College.train)
   > summary(College.lm)
   lm(formula = Apps ~ ., data = College.train)
   Residuals:
   Min 10 Median 30 Max
-5741.2 -479.5 15.3 359.6 7258.0
Coefficients:

Estimate Std. Error t value Pr(>|t|)
(Intercept) -7.902e+02 6.381e+02 -1.238 0.216410
PrivateYes -3.070e+02 2.006e+02 -1.531 0.126736
Accept 1.779e+00 5.420e-02 32.830 c 2e-16 ***
Enroll -1.470e+00 3.115e-01 -4.720 3.35e-06 ***
Top10perc 6.673e+01 8.310e+00 8.030 1.31e-14 ***
Top25perc -2.231e+01 6.533e+00 -3.415 0.000708 ***
   Top10perc 6.673e+01 8.310e+00
Top25perc -2.231e+01 6.533e+00
F.Undergrad 9.269e-02 5.529e-02
P.Undergrad 9.397e-03 5.493e-02
                                                                                                         1.676 0.094538
0.171 0.864275
  P.Undergrad 9.39/e-03 5.495e-02 0.171 0.8042/5  
Outstate - 1.084e-01 2.700e-02 - 4.014 7.22e-05  
Room. Board 2.115e-01 7.224e-02 2.928 0.003622  
Personal 6.133e-01 3.985e-01 0.731 0.46539  
Personal 6.133e-03 8.803e-02 0.070 0.944497  
PhD -1.548e+01 6.681e+00 -2.316 0.021082  
Terminal 6.415e-00 7.290e+00 0.880 0.379470  
S.F.Ratio 2.283e-01 2.047e-01 1.15 0.265526  
Description 1.134a-06 6.88a-00 0.1860 8.83274
                                                                                                         2.928 0.003622 **
0.731 0.465399

        perc. alumni
        1.134e400
        6.083e+00
        0.186 0.852274

        Expend
        4.857e-02
        1.619e-02
        2.999
        0.002890

        Grad.Rate
        7.490e+00
        4.397e+00
        1.703
        0.089324

   Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' '1
  Residual standard error: 1083 on 370 degrees of freedom
Multiple R-squared: 0.9389, Adjusted R-squared: 0.9361
F-statistic: 334.3 on 17 and 370 DF, p-value: < 2.2e-16
```

Figure 5.1: Screenshot for 'Model summary'

The error for test dataset is 1135758.

#### 5.3 c

```
1 train_mat = model.matrix(Apps ~ ., data = College.train)
2 test_mat = model.matrix(Apps ~ ., data = College.test)
3 College.ridge = cv.glmnet(train_mat, College.train$Apps, alpha = 0)
4 plot(College.ridge)
```

```
5  College.lambda = College.ridge$lambda.min
6  College.pred = predict(College.ridge, newx=test_mat, s=College.lambda)
7  College.error = mean((College.test$Apps - College.pred)^2)
```

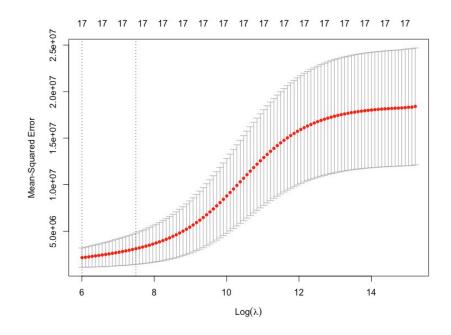


Figure 5.2: Screenshot for 'Ridge Regression'

The error for test dataset is 976261.5. Ridge Regression get a better result than Linear Regression.

#### 5.4 d

```
1 College.lasso = cv.glmnet(train_mat, College.train$Apps, alpha = 1)
2 plot(College.lasso)
3 College.lambda = College.lasso$lambda.min
4 College.pred = predict(College.lasso, newx=test_mat, s=College.lambda)
5 College.error = mean((College.test$Apps - College.pred)^2)
```

The error for test dataset is 1115901. LASSO Regression get a better result than Linear Regression but not as good as Ridge Regression.

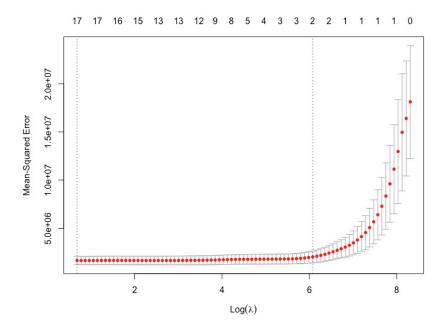


Figure 5.3: Screenshot for 'LASSO Regression'

```
6.1 a
1 set.seed(1)
2 n = 1000
3 p = 20
5 xmat = matrix(rnorm(n*p), n, p)
6 \text{ eps} = rnorm(n)
7
8 beta = rnorm(p)
9 \text{ beta}[3] = 0
10 \text{ beta}[6] = 0
11 \text{ beta}[9] = 0
12 \text{ beta}[12] = 0
13 \text{ beta}[15] = 0
14 \text{ beta}[18] = 0
15
16
17 y = xmat %*% beta + eps
   I set \beta_3, \beta_6, \beta_9, \beta_{12}, \beta_{15}, \beta_{18} to be zero.
   6.2 b
1 train = sample(1: n, 100)
2 \text{ test} = -\text{train}
3 xmat.train = xmat[train, ]
4 xmat.test = xmat[test, ]
5 y.train = y[train, ]
6 y.test = y[test,]
   6.3 c
1 ans.df = data.frame(x = xmat.train, y = y.train)
2 ans.model = regsubsets(y ~ ., data = ans.df, nvmax = p)
3 ans.summary = summary(ans.model)
4 ans.summary
5
6 ans.train_mse = rep(NA, p)
7 ans.xcols = colnames(xmat, do.NULL = F, prefix = "x.")
8 for(i in 1:p){
    c_i = coef(ans.model, id = i)
9
     if(i > 1){
10
        y.train.pred = as.matrix(xmat.train[, ans.xcols %in% names(c_i)] %*%
11
12
                                        c_i[names(c_i) %in% ans.xcols])
     }
13
     else
14
15
        y.train.pred = as.matrix(xmat.train[, ans.xcols %in% names(c_i)] *
16
17
                                        c_i[names(c_i) %in% ans.xcols])
18
     }
```

```
19 ans.train_mse[i] = mean((y.train-y.train.pred)^2)
20 }
21 plot(ans.train_mse, ylab = "Training_MSE",xlab = "Subset_Size",
22  pch = 20, type = "o", col = "blue")
23 title("MSE-Subset_size_Curve")
```

```
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 Call: regsubsets.formula(y \sim ., data = ans.df, nvmax = p)
 20 Variables (and intercept)
Forced in Forced out
           FALSE
          FALSE
                       FALSE
 x.3
x.4
x.5
                       FALSE
FALSE
FALSE
           FALSE
          FALSE
FALSE
 x.6
x.7
x.8
x.9
           FALSE
                       FALSE
          FALSE
FALSE
FALSE
                       FALSE
FALSE
                       FALSE
           FALSE
 x.10
                       FALSE
           FALSE
FALSE
                       FALSE
FALSE
 x.13
           FALSE
                       FALSE
 x 14
           FALSE
                       FALSE
          FALSE
FALSE
FALSE
 x.15
x.16
                       FALSE
FALSE
 x.17
                       FALSE
          FALSE
FALSE
FALSE
                       FALSE
FALSE
FALSE
 x.18
x.19
```

Figure 6.1: Screenshot for 'Model summary'

# Training MSE-Subset size Curve Straining MSE-Subset size Curve Straining MSE-Subset Size Curve

Figure 6.2: Screenshot for 'Training MSE-Subset size Curve'

#### 6.4 d

```
1 ans.test_mse = rep(NA, p)
   ans.xcols = colnames(xmat, do.NULL = F, prefix = "x.")
3
   for(i in 1:p){
4
     c_i = coef(ans.model, id = i)
     if(i > 1){
5
       y.test.pred = as.matrix(xmat.test[, ans.xcols %in% names(c_i)] %*%
6
                                   c_i[names(c_i) %in% ans.xcols])
7
8
     }
9
     else
10
       y.test.pred = as.matrix(xmat.test[, ans.xcols %in% names(c_i)] *
11
12
                                   c_i[names(c_i) %in% ans.xcols])
13
14
     ans.test_mse[i] = mean((y.test-y.test.pred)^2)
15
  plot(ans.test_mse, ylab = "Testing MSE", xlab = "Subset Size",
16
        pch = 20, type = "o", col = "blue")
17
   title("Testing MSE-Subset size Curve")
```

#### Testing MSE-Subset size Curve

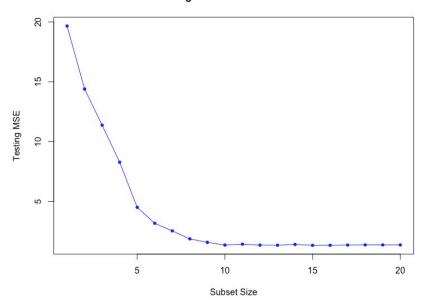


Figure 6.3: Screenshot for 'Testing MSE-Subset size Curve'

#### 6.5 e

#### 1 which.min(ans.test\_mse)

The return of this function is 15, which means when there are 15 predictors in the estimated model, the model reaches the smallest test error.

6.6 f

```
1 coef(ans.model, id = which.min(ans.test_mse))
2 beta
```

Figure 6.4: Screenshot for 'β value'

The model finds that  $\beta_3$ ,  $\beta_6$ ,  $\beta_{12}$ ,  $\beta_{14}$ ,  $\beta_{18}$  are zeros.

#### 6.7 g

```
1 ans.beta_error = rep(NA, p)
2 ans.xcols = colnames(xmat, do.NULL = F, prefix = "x.")
3
 for(i in 1:p){
     c_i = coef(ans.model, id = i)
4
     ans.beta_error[i] = sqrt(sum((beta[ans.xcols %in% names(c_i)] -
5
6
                                      c_i[names(c_i) %in% ans.xcols])^2) +
                                 sum(beta[!(ans.xcols %in% names(c i))]^2))
7
8
  plot(x = 1:p, ans.beta_error, xlab = "Coefficientu#",
9
        ylab = "Beta error")
10
  title("Beta_error_on_each_dimesion")
```

Compared with the result in (d), the model reaches the minimum beta error when there are 15 predictors. That's close to the result of 12 in (d). And when the number of predictors is 15, the  $\beta$  is also very small.

## Beta error on each dimesion

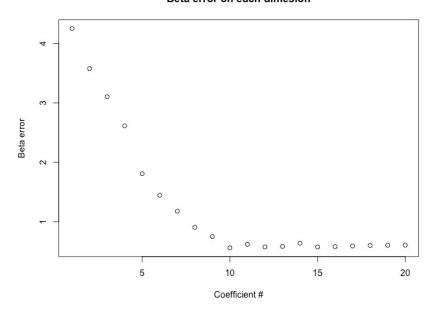


Figure 6.5: Screenshot for ' $\beta$  error'