安徽大学 20 19 -20 20 学年第 - 学期 《 概率论与数理统计 A 》 考试试题 (A 卷) 参考答案及评分标准

一、选择题(每小题2分,共10分)

1. D

2. B

3. A

4. D

5. C

二、填空题 (每小题 2 分, 共 10 分)

7. $\frac{1}{4}$

8. 2 9. μ^2 10. (3.412, 4.588)

三、计算题 (每小题 12 分, 共 72 分)

11. 解:以A和A分别表示订阅了甲报和乙报,B表示第二年续订.则

$$P(A_1\overline{A}_2) = 0.6 - 0.1 = 0.5, P(\overline{A}_1A_2) = 0.5 - 0.1 = 0.4, P(A_1A_2) = 0.1;$$

 $P(B \mid A_1 \overline{A_2}) = 0.7, P(B \mid \overline{A_1} A_2) = 0.6, P(B \mid A_1 A_2) = 0.8.$

(1) 由全概率公式有

$$P(B) = P(\overline{A}_1 A_2) P(B \mid A_1 \overline{A}_2) + P(\overline{A}_1 A_2) P(B \mid \overline{A}_1 A_2) + P(A_1 A_2) P(B \mid A_1 A_2)$$

= 0.5 \times 0.7 + 0.4 \times 0.6 + 0.1 \times 0.8 = 0.67.

(2) 利用贝叶斯公式有

$$P(A_1A_2 \mid B) = \frac{P(A_1A_2)P(B \mid A_1A_2)}{P(B)} = \frac{0.1 \times 0.8}{0.67} = \frac{8}{67}.\dots 12$$

12. 解: (1) 易知 a+b+1/3=1, EX=b+2/3=1, 故 a=b=1/3.

(2) 注意到 $U \ge V$, 故P(U = 0, V = 1) = P(U = 0, V = 2) = P(U = 1, V = 2) = 0, 另 外

$$P(U = 0, V = 0) = P(X = 0, Y = 0) = P(X = 0)P(Y = 0) = 1/9;$$

同理,

$$P(U=1,V=0) = P(X=1,Y=0) + P(X=0,Y=1) = 2/9;$$

 $P(U=1,V=1) = P(X=1,Y=1) = 1/9;$
 $P(U=2,V=0) = P(X=2,Y=0) + P(X=0,Y=2) = 2/9;$
 $P(U=2,V=1) = P(X=2,Y=1) + P(X=1,Y=2) = 2/9;$
 $P(U=2,V=2) = P(X=2,Y=2) = 1/9.$

从而(U,V)的概率分布为

| U | 0 | 1 | 2 |
|---|------------|-----|-----|
| 0 | 1/9 | 0 | 0 |
| 1 | 2/9 | 1/9 | 0 |
| 2 | 2/9 2/9 | 2/9 | 1/9 |

(3) $EUV = 1 \times 1 \times 1/9 + 2 \times 1 \times 2/9 + 2 \times 2 \times 1/9 = 1$,

$$EU = 0 \times 1/9 + 1 \times 3/9 + 2 \times 5/9 = 13/9, EV = 0 \times 5/9 + 1 \times 3/9 + 2 \times 1/9 = 5/9,$$

13. 解: (1) 由
$$f_X(x) = \int_{-\infty}^{+\infty} f(x,y) dy$$
 得, 当 $x > 0$ 时,

$$f_X(x) = \int_{-\infty}^{+\infty} f(x, y) dy = \int_{0}^{+\infty} x e^{-x(1+y)} dy = e^{-x}$$

从而
$$f_X(x) = \begin{cases} e^{-x}, & x > 0, \\ 0, & 其他. \end{cases}$$

同理, $f_{\gamma}(y) = \int_{-\infty}^{+\infty} f(x,y)dx$, 当 y > 0 时,

$$f_Y(y) = \int_{-\infty}^{+\infty} f(x, y) dx = \int_0^{\infty} x e^{-x(1+y)} dx = \frac{1}{(1+y)^2},$$

从而

$$f_{Y}(y) = \begin{cases} \frac{1}{(1+y)^{2}}, & y > 0, \\ 0, & 其他. \end{cases}$$

(2) Z的分布函数

$$F_Z(z) = P(XY \le z) = \begin{cases} 0, & z \le 0, \\ \int_0^\infty \int_0^{z/x} x e^{-x(1+y)} dy dx, & z > 0, \end{cases} = \begin{cases} 0, & z \le 0, \\ 1 - e^{-z}, & z > 0. \end{cases}$$

因此Z的密度函数为

$$f_Z(z) = \begin{cases} e^{-z}, & z > 0, \\ 0, & z \le 0. \end{cases}$$
 12/3

$$EX^2 = \int_{-\infty}^{+\infty} \frac{x^2}{2} e^{-|x|} dx = \int_{0}^{+\infty} x^2 e^{-x} dx = 2,$$

从而 $DX = EX^2 - (EX)^2 = EX^2 = 2$.

(2) $Cov(X, |X|) = EX |X| - EX \cdot E |X| = EX |X| = \int_{-\infty}^{+\infty} \frac{x |x|}{2} e^{-|x|} dx = 0$,故 X 和 |X| 不相关.

(3) 对于给定的 $0 < a < \infty$,由于事件 $(|X| \le a) \subset (X \le a)$,且

 $P(X \le a) < 1, P(|X| \le a) > 0, \text{ in } P(X \le a, |X| \le a) = P(|X| \le a),$

但 $P(X \le a)P(|X| \le a) < P(|X| \le a)$, 所以

 $P(X \le a, |X| \le a) \ne P(X \le a)P(|X| \le a)$, 即 X 和 |X| 不独立.··················12 分

15. 解: (1) 以 \bar{X} 表示该样本均值,则 $\bar{X} \sim N(3.4,36/n)$,故 $\frac{\bar{X}-3.4}{6/\sqrt{n}} \sim N(0,1)$.

从而有

$$P(1.4 < \overline{X} < 5.4) = P(|\overline{X} - 3.4| < 2) = P(\frac{|\overline{X} - 3.4|}{6/\sqrt{n}} < \frac{2}{6/\sqrt{n}})$$
$$= \Phi(\frac{\sqrt{n}}{3}) - \Phi(-\frac{\sqrt{n}}{3}) = 2\Phi(\frac{\sqrt{n}}{3}) - 1 \ge 0.95,$$

故 $\Phi(\frac{\sqrt{n}}{3}) \ge 0.975$,即 $\frac{\sqrt{n}}{3} \ge 1.96$,从而 $n \ge (1.96 \times 3)^2 \approx 34.57$,所以n至少应取 35.

(2) 由于
$$\frac{nS_n^2}{\sigma^2} = \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - \overline{X})^2 \sim \chi^2(n-1)$$
,故

16. 解: (1) 矩估计: 易见 $EX = \int_{c}^{+\infty} \frac{x}{\theta} e^{-(x-c)/\theta} dx = \theta + c$,

$$EX^{2} = \int_{c}^{+\infty} \frac{x^{2}}{\theta} e^{-(x-c)/\theta} dx = \theta^{2} + (\theta + c)^{2}.$$

令 $EX = \overline{x}$, $EX^2 = \frac{1}{n} \sum_{i=1}^{n} x_i^2$, 则 θ 与 c 的矩估计值分别为

$$\hat{\theta} = \sqrt{\frac{1}{n} \sum_{i=1}^{n} (x_i - \overline{x})^2}, \ \hat{c} = \overline{x} - \sqrt{\frac{1}{n} \sum_{i=1}^{n} (x_i - \overline{x})^2}.$$

(2) 极大似然估计: 易见似然函数为

$$L(\theta,c) = \prod_{i=1}^{n} \frac{1}{\theta} e^{-(x_i - c)/\theta} = \frac{1}{\theta^n} e^{-\sum_{i=1}^{n} (x_i - c)/\theta}, \ x_i \ge c, 1 \le i \le n,$$

对数似然函数为

$$\ln L(\theta, c) = -n \ln \theta - \sum_{i=1}^{n} (x_i - c)/\theta, \ x_i \ge c, 1 \le i \le n.$$

故当 $x_i \ge c$, $1 \le i \le n$, 即 $c \le \min(x_1, x_2, \dots, x_n) = x_{(1)}$ 时,

$$\ln L(\theta,c) = -n \ln \theta - \sum_{i=1}^{n} (x_i - c) / \theta,$$

由于

$$\frac{\partial \ln L(\theta, c)}{\partial c} = \frac{n}{\theta} > 0,$$

即 $L(\theta,c)$ 关于 c 单调递增,故为使 $L(\theta,c)$ 取到最大值,c 应取 $\min(x_1,x_2,\cdots x_n)=x_{(1)}$,

即 c 的最大似然估计为 $\hat{c} = x_{(1)}$; 再令

$$\frac{\partial \ln L(\theta,c)}{\partial \theta} = -\frac{n}{\theta} + \sum_{i=1}^{n} (x_i - c) / \theta^2 = 0,$$

可得到θ的极大似然估计值为

四、应用题(每小题8分,共8分)

17. 解: 由题意 $X \sim N(\mu, \sigma^2)$, $\mu = \sigma^2$ 皆未知. 需检验的假设为

$$H_0: \mu = 500 \leftrightarrow H_1: \mu \neq 500.$$

选用统计量

$$t = \frac{\overline{X} - 500}{s / \sqrt{n}},$$

故在 H_0 成立的条件下 $t \sim t(n-1)$,从而得到该假设检验问题的拒绝域为

$$W = \left\{ \left| \frac{\overline{X} - 500}{s / \sqrt{n}} \right| > t_{\frac{\alpha}{2}}(n - 1) \right\}.$$

将 $\bar{x} = 510$, s = 20, n = 9, $\alpha = 0.05$ 代入得,

$$\left| \frac{\overline{X} - 500}{s / \sqrt{n}} \right| = 1.5$$
, $\overline{\text{mi}} t_{0.025}(8) = 2.31$,