

# MATH 5473 Homework 2 LUO Yuanhui

1. *Phase transition in PCA "spike" model*: Consider a finite sample of  $n$  i.i.d vectors  $x_1, x_2, \dots, x_n$  drawn from the  $p$ -dimensional Gaussian distribution  $\mathcal{N}(0, \sigma^2 I_{p \times p} + \lambda_0 u u^T)$ , where  $\lambda_0/\sigma^2$  is the signal-to-noise ratio (SNR) and  $u \in \mathbb{R}^p$ . In class we showed that the largest eigenvalue  $\lambda$  of the sample covariance matrix  $S_n$

$$S_n = \frac{1}{n} \sum_{i=1}^n x_i x_i^T$$

pops outside the support of the Marcenko-Pastur distribution if

$$\frac{\lambda_0}{\sigma^2} > \sqrt{\gamma},$$

or equivalently, if

$$\text{SNR} > \sqrt{\frac{p}{n}}.$$

(Notice that  $\sqrt{\gamma} < (1 + \sqrt{\gamma})^2$ , that is,  $\lambda_0$  can be "buried" well inside the support Marcenko-Pastur distribution and still the largest eigenvalue pops outside its support). All the following questions refer to the limit  $n \rightarrow \infty$  and to almost surely values:

- Find  $\lambda$  given  $\text{SNR} > \sqrt{\gamma}$ .
- Use your previous answer to explain how the SNR can be estimated from the eigenvalues of the sample covariance matrix.
- Find the squared correlation between the eigenvector  $v$  of the sample covariance matrix (corresponding to the largest eigenvalue  $\lambda$ ) and the "true" signal component  $u$ , as a function of the SNR,  $p$  and  $n$ . That is, find  $|\langle u, v \rangle|^2$ .
- Confirm your result using MATLAB, Python, or R simulations (e.g. set  $u = e$ ; and choose  $\sigma = 1$  and  $\lambda_0$  in different levels. Compute the largest eigenvalue and its associated eigenvector, with a comparison to the true ones.)

Solution: (a) Let the corresponding eigenvector be  $v$ , then

$$S_n v = \lambda v$$

Let  $\Sigma = \sigma^2 I_{p \times p} + \lambda_0 u u^T$ ,  $y_v \triangleq \Sigma^{-\frac{1}{2}} X v$ , then  $Y = \Sigma^{-\frac{1}{2}} X \sim \mathcal{N}(0, I_p)$

$T_n = \frac{1}{n} Y Y^T$  is a Wishart Matrix, then the limit distribution of

$T_n$ 's eigenvalues follows a MP distribution.

Notice that  $T_n = \frac{1}{n} Y Y^T = \frac{1}{n} (\Sigma^{-\frac{1}{2}} X) (\Sigma^{-\frac{1}{2}} X)^T = \Sigma^{-\frac{1}{2}} S_n \Sigma^{-\frac{1}{2}}$ , then

$$S_n = \Sigma^{\frac{1}{2}} T_n \Sigma^{\frac{1}{2}}$$

$$S_n v = \Sigma^{\frac{1}{2}} T_n (\Sigma \Sigma^{-\frac{1}{2}}) v = \lambda v \Leftrightarrow T_n \Sigma (\Sigma^{-\frac{1}{2}} v) = \lambda (\Sigma^{-\frac{1}{2}} v)$$

Then  $\lambda$ ,  $\Sigma^{-\frac{1}{2}} v$  is the eigenvalue and corresponding eigenvector of  $T_n \Sigma$

Let  $v^* = c(\Sigma^{-\frac{1}{2}} v)$  s.t.  $v^{*T} v^* = 1$ , we have  $c^2 = (R(u^T \delta) + 1) \delta^2$ .

then  $v^*$  is a normalized eigenvector of  $T_n \Sigma$

$$T_n \Sigma v^* = T_n (\delta^2 I_p + \lambda_0 u u^T) v^* = \lambda v^* \Leftrightarrow \lambda_0 T_n u u^T v^* = (\lambda I_p - T_n \delta^2 I_p) v^*$$

$$\Leftrightarrow u^T v^* = u^T (\lambda I_p - T_n \delta^2 I_p)^{-1} \lambda_0 T_n u u^T v^* \quad (*)$$

Suppose  $u^T v^* \neq 0$ ,  $T_n = W \Lambda W^T$ ,  $W W^T = I_p$ ,  $\Lambda = \text{diag} \{ \lambda_1, \dots, \lambda_p \}$ , then

$$1 = \lambda_0 \sum_{i=1}^p u_i^2 \frac{\lambda_i}{\lambda - \delta^2 \lambda_i} = \lambda_0 \int_a^b \frac{t}{\lambda - \delta^2 t} d\mu_{u_p}, \text{ by Stieltjes transform,}$$

$$1 = \frac{\lambda_0}{4\delta} [2\lambda - (a+b) - 2\sqrt{(\lambda-a)(b-\lambda)}] \text{ for } \lambda > (1+\sqrt{r})^2 \triangleq b \text{ and } \text{SNR} > \sqrt{r}$$

$$\text{Then given } \text{SNR} > \sqrt{r}, \lambda = \lambda_0 + \frac{r}{\lambda_0} + 1 + r = (1+\lambda_0)(1+\frac{r}{\lambda_0})$$

$$\text{Therefore, } \lambda_{\max}(S_n) = \begin{cases} (1+\sqrt{r})^2 = b, & \delta_x^2 \leq \sqrt{r} \\ (1+\delta_x^2)(1+\frac{r}{\delta_x^2}), & \delta_x^2 > \sqrt{r} \end{cases}$$

1b) For  $S_n = \frac{1}{n} X X^T$ ,  $b = (1+\sqrt{r})^2$ , we can estimate SNR

by  $\begin{cases} \text{SNR} \leq \sqrt{r} & \text{if } \lambda_{\max}(S_n) = b \\ \text{SNR} > \sqrt{r} & \text{if } \lambda_{\max}(S_n) = (1 + \sigma_x^2)(1 + \frac{r}{\sigma_x^2}) \end{cases}$ . Here SNR

$= \sigma_x^2$  by assuming  $\sigma_e = 1$  WLOA.

(c) By (\*) we have  $u^T v^* = u^T (\lambda I_p - T_n \sigma^2 I_p)^T \lambda_0 T_n u u^T v^*$ ,

then  $(u^T v^*)^T (u^T v^*) = \lambda_0^2 (u^T v^*)^T u^T T_n (\lambda I_p - T_n \sigma^2 I_p)^2 T_n u (u^T v^*)$ .

Then  $|u^T v^*|^2 \sim \lambda_0^2 \int_a^b \frac{t^2}{(\lambda - \sigma^2 t^2)^2} d\mu_{mp}(t) = \frac{\lambda_0^2}{4r} (-4r + (a+b) +$

$2\sqrt{(\lambda-a)(\lambda-b)} + \frac{\lambda(2\lambda - (a+b))}{\sqrt{(\lambda-a)(\lambda-b)}}) = \frac{1 - \frac{r}{R}}{1 + \sigma + \frac{2r}{R}}$

Then  $(u^T v)^2 = (\frac{1}{c} u^T \Sigma^{\frac{1}{2}} v^*)^2 = \frac{1}{c^2} ((\sqrt{1+R})u)^T v^*)^2 =$

$\frac{(1+R)(u^T v^*)^2}{R(u^T v)^2 + 1} = \frac{1+R - \frac{r}{R} - \frac{r}{R^2}}{1+R + r + \frac{r}{R}} = \frac{1 - \frac{r}{R^2}}{1 + \frac{r}{R}}$

(d) The code can be seen in the Ex1. ipynb.