

Sparse Learning via Inverse Scale Space Method

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Sparsity lies in everywhere ...

In “compressive sensing”,

- signals are sparse w.r.t. certain basis (e.g. Fourier, Wavelet)
- signals are piecewise constant or smooth
- signals might be of low rank matrix or tensor
- sparse graphical models, e.g. Ising, Potts

In “deep learning”,

- sparse (local) inputs in convolutional networks ([Minsky et al.](#))
- symmetry: translation invariance, weight sharing ([Fukushima, LeCun](#))
- Compositional Sparsity ([Poggio](#))
- pruning: Lottery Ticket Hypothesis ([Frankle & Carbin](#))

1 DessiLBI (Python)

- The Lottery Ticket Hypothesis

2 From LASSO to Inverse Scale Space

- The LASSO
- LASSO and Bias
- Differential Inclusion of Inverse Scale Space
- Statistical Path Consistency with Early Stopping
- Large Scale Algorithm: Linearized Bregman Iteration (LBI)

3 Variable Splitting: Split LBI

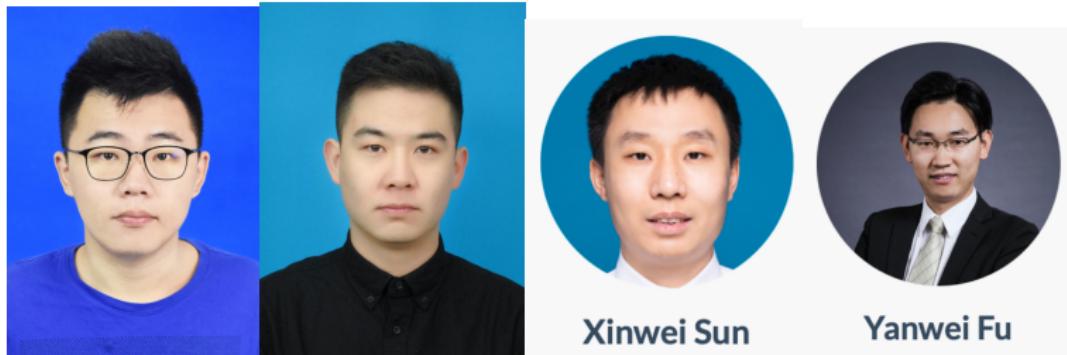
- A Weaker Irrepresentable/Incoherence Condition
- Application: Finding Sparse Subnets in Deep Learning

4 Summary

5 *Appendix: Libra (R)

DessiLBI: Deep structurally splitting Linearized Bregman Iteration

<https://github.com/DessiLBI2020/DessiLBI>



Courtesy of Chen Liu/Donghao Li (HKUST-Fudan), Xinwei Sun/Yanwei Fu (Fudan)

Finding the “winning tickets”

- Existing approaches typically adopt the following 2-stage algorithms:
 - [(i)] Train dense overparameterized networks (overparameterization helps optimization landscape without hurting the generalization, yet expensive)
 - [(ii)] Prune+retrain to produce sparse subnets
- **DessiLBI** gives us 1-stage procedure (end-to-end) to render sparse subnet structures with *Early Stopping*
 - [Reference] Yanwei Fu, et al. ICML 2020 and TPAMI 2022, with a CVPR 2022 tutorial at
<https://sparse-learning.github.io/>

DessiLBI: Sparse Filters Learned on MNIST

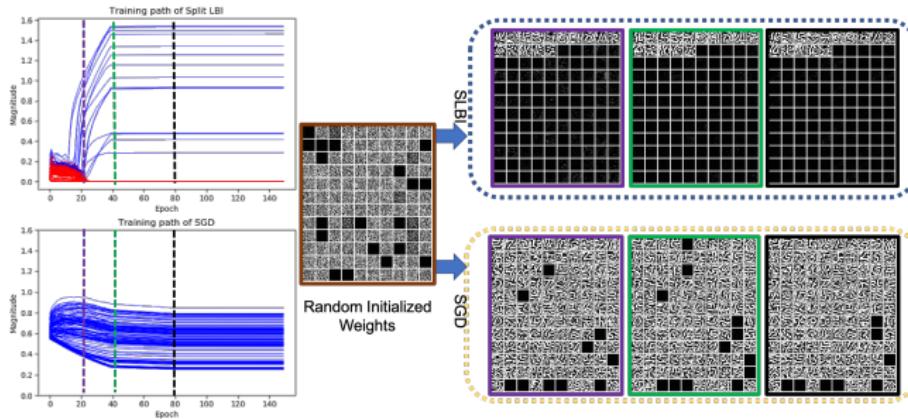


Figure: [Fu et al. ICML 2020/TPAMI 2022] Visualization of solution path and filter patterns in the third convolutional layer (i.e., conv.c5) of LeNet-5, trained on MNIST, showing a sparse selection of filters without sacrificing accuracy.

Sparse Neural Nets in Early Stopping (Lottery Tickets)

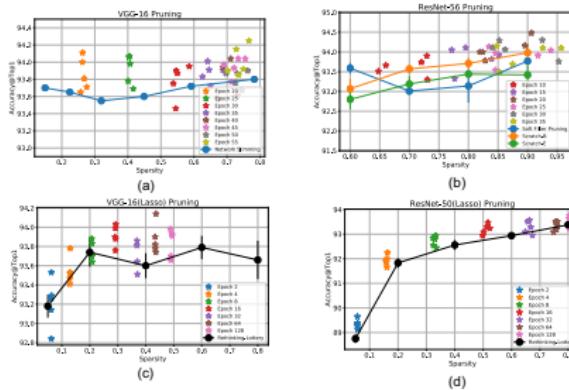


Figure: SplitLBI with early stopping finds sparse subnets whose test accuracies (stars) after retrain are comparable or even better than the baselines (Network Slimming, Soft-Filter Pruning, Scratch-B, Scratch-E, and “Rethinking-Lottery” as reported in Rethink the Value of Pruning. Sparse filters of VGG-16 and ResNet-56 are show in (a) and (b), while sparse weights of VGG-16 and ResNet-50 are shown in (c) and (d).

DessiLBI: Non-semantic Features Learned on ImageNet

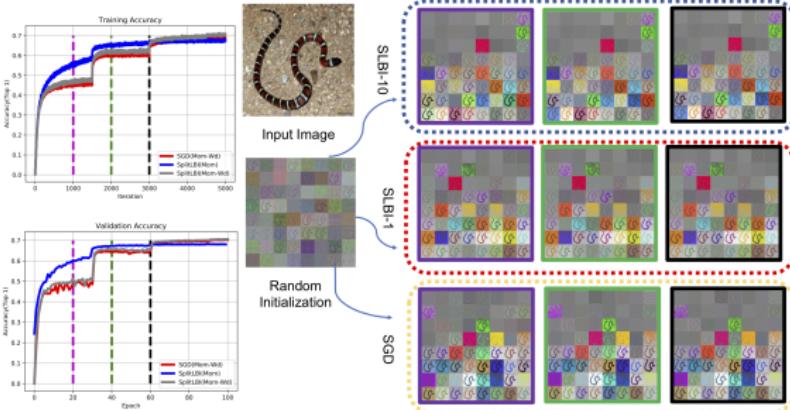


Figure: Visualization of the first convolutional layer filters of ResNet-18 trained on ImageNet-2012, where **texture** features are more important than **colour/shapes**. Given the input image and initial weights visualized in the middle, filter response gradients at 20 (purple), 40 (green), and 60 (black) epochs are visualized. SGD with Momentum (Mom) and Weight Decay (WD), is compared with SLBI.

How does it work?

In the sequel, we shall see a story on [statistical model selection consistency with early stopping](#):

- A simple iterative algorithm shadows a particular kind of dynamics: [differential inclusions of inverse scale spaces](#), as special cases of [Mirror Descent](#), where important features are learned fast
- Simple discretized algorithm, amenable for parallel implementation
- Under nearly the same condition as LASSO, it reaches [model selection consistency via early stopping](#)
- but may incur [less bias](#) than LASSO
- Equipped with [variable splitting](#), it [weakens](#) the conditions of generalized LASSO in feature selection

Sparse Linear Regression

Assume that $\beta^* \in \mathbb{R}^p$ is sparse and unknown. Consider recovering β^* from n linear measurements

$$y = X\beta^* + \epsilon, \quad y \in \mathbb{R}^n$$

where $\epsilon \sim \mathcal{N}(0, \sigma^2)$ is **noise**.

- **Basic Sparsity:** $S := \text{supp}(\beta^*)$ ($s = |S|$) and T be its complement.
 - X_S (X_T) be the columns of X with indices restricted on S (T)
 - X is n -by- p , with $p \gg n \geq s$.
- Generalized **Structural/Transformational Sparsity:** $\gamma^* = D\beta^*$ is sparse, where D is a linear transform (wavelet, gradient, etc.), $S = \text{supp}(\gamma^*)$
- *How to recover β^* (or γ^*) sparsity pattern (**sparsistency**) and estimate values with variations (**consistency**)?*

Best Possible in Basic Setting: The Oracle Estimator

Had God revealed S to us, the *oracle estimator* was the subset least square solution (MLE) with $\tilde{\beta}_T^* = 0$ and

$$\tilde{\beta}_S^* = \beta_S^* + \frac{1}{n} \Sigma_n^{-1} X_S^T \epsilon, \quad \text{where } \Sigma_n = \frac{1}{n} X_S^T X_S \quad (1)$$

“Oracle properties”

- **Model selection consistency:** $\text{supp}(\tilde{\beta}^*) = S$;
- **Normality:** $\tilde{\beta}_S^* \sim \mathcal{N}(\beta^*, \frac{\sigma^2}{n} \Sigma_n^{-1})$.

So $\tilde{\beta}^*$ is **unbiased**, i.e. $\mathbb{E}[\tilde{\beta}^*] = \beta^*$.

The Lasso

- Lasso stands for **Least Absolute Shrinkage and Selection Operator**.
 - Chen-Donoho-Saunders'1996 (BPDN)
 - Tibshirani'1996 (LASSO)
- The Lasso estimator $\hat{\beta}_\lambda^L$ is the minimizer of

$$\sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2 + \lambda \sum_{j=1}^p |\beta_j|$$

- We may use $\|\beta\|_1 = \sum_{j=1}^p |\beta_j|$, which is the l_1 norm.
- LASSO often shrinks coefficients to be identically 0. (This is not the case for ridge)
- Hence it performs variable selection, and yields sparse models.

Example: Lasso Regularization Path in Credit data.

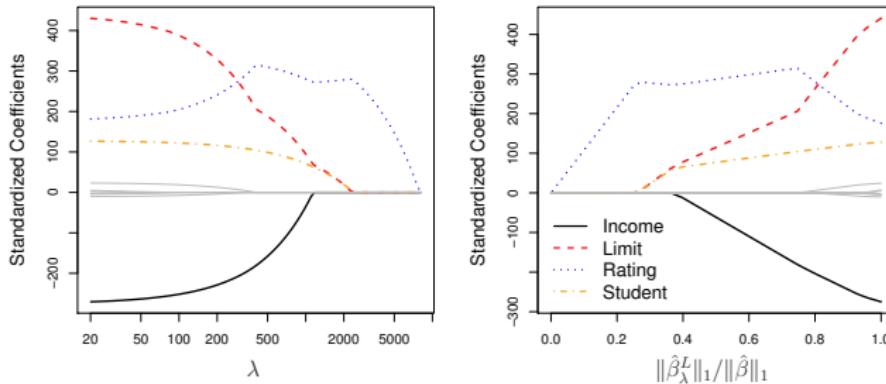


Figure: The standardized lasso coefficients on the Credit data (ISLR Chapter 6) set are shown as a function of λ and $\|\hat{\beta}_\lambda^L\|_1 / \|\hat{\beta}\|_1$.

Another formulation

- For Lasso: Minimize

$$\sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2 \text{ subject to } \sum_{j=1}^p |\beta_j| \leq s$$

- For Ridge: Minimize

$$\sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2 \text{ subject to } \sum_{j=1}^p \beta_j^2 \leq s$$

- For l_0 : Minimize

$$\sum_{i=1}^n (y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij})^2 \text{ subject to } \sum_{j=1}^p I(\beta_j \neq 0) \leq s$$

l_0 method penalizes number of non-zero coefficients. A difficult (NP-hard) problem for optimization.

Variable selection property for Lasso

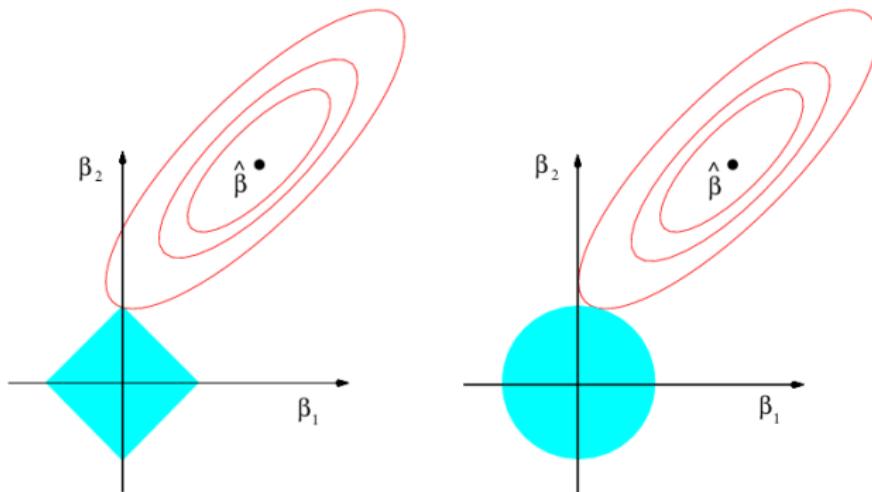


Figure: Contours of the error and constraint functions for the lasso (left) and ridge regression (right). The solid blue areas are the constraint regions, $|\beta_1| + |\beta_2| \leq s$ and $\beta_1^2 + \beta_2^2 \leq s$, while the red ellipses are the contours of the RSS.

Simple cases

- Consider the simple model $y_i = \beta_i + \epsilon_i$, $i = 1, \dots, n$ and $n = p$. Then,
 - The least squares: $\hat{\beta}_j^{\text{LSE}} = y_j$;
 - The ridge: $\hat{\beta}_j^R = y_j/(1 + \lambda)$;
 - The Lasso: $\hat{\beta}_j^L = \text{sign}(y_j)(|y_j| - \lambda/2)_+$.
- Slightly more generally, suppose input columns of the X are standardized to be mean 0 and variance 1 and are orthogonal.

$$\hat{\beta}_j^R = \hat{\beta}_j^{\text{LSE}} / (1 + \lambda)$$

$$\hat{\beta}_j^L = \text{sign}(\hat{\beta}_j^{\text{LSE}}) (|\hat{\beta}_j^{\text{LSE}}| - \lambda/2)_+$$

for $j = 1, \dots, p$.

Example

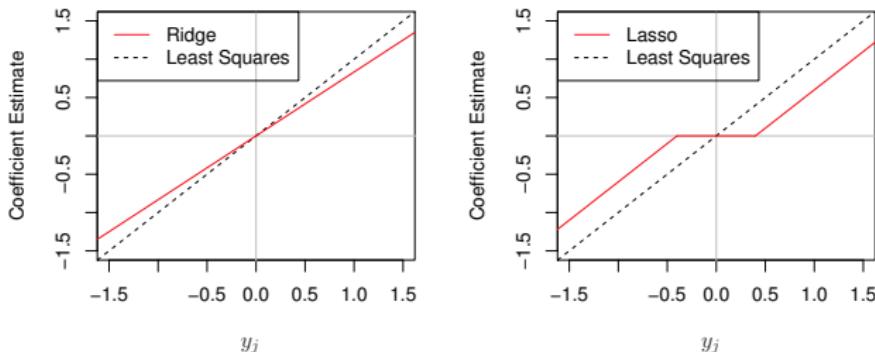


Figure: The ridge regression and lasso coefficient estimates for a simple setting with $n = p$ and X a diagonal matrix with 1 on the diagonal. Left: The ridge regression coefficient estimates are shrunken proportionally towards zero, relative to the least squares estimates. Right: The lasso coefficient estimates are soft-thresholded towards zero.

The First Order KKT Condition of LASSO

Rewrite **LASSO** as:

$$\min_{\beta} \|\beta\|_1 + \frac{t}{2n} \|y - X\beta\|_2^2.$$

1st order optimality condition:

$$\frac{\rho_t}{t} = \frac{1}{n} X^T (y - X\beta_t), \quad (2a)$$

$$\rho_t \in \partial \|\beta_t\|_1, \quad (2b)$$

where we set $t = 2n/\lambda$ for the later convenience.

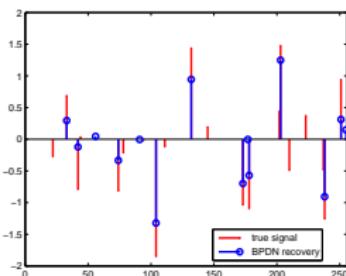
The Bias of LASSO

LASSO is **biased**, i.e. $\mathbb{E}(\hat{\beta}) \neq \beta^*$

- e.g. $X = Id$, $n = p = 1$, LASSO is soft-thresholding

$$\hat{\beta}_\tau = \begin{cases} 0, & \text{if } \tau < 1/\tilde{\beta}^*; \\ \tilde{\beta}^* - \frac{1}{\tau}, & \text{otherwise,} \end{cases}$$

- e.g. $n = 100$, $p = 256$, $X_{ij} \sim \mathcal{N}(0, 1)$, $\epsilon_i \sim \mathcal{N}(0, 0.1)$



True vs LASSO (t hand-tuned,
courtesy of Wotao Yin)

LASSO Estimator is Biased at Path Consistency

Even when the following **path consistency** (conditions given by [Zhao-Yu'06](#), [Zou'06](#), [Yuan-Lin'07](#), [Wainwright'09](#), etc.) is reached at τ_n :

$$\exists \tau_n \in (0, \infty) \text{ s.t. } \text{supp}(\hat{\beta}_{\tau_n}) = S,$$

LASSO estimate is biased away from the oracle estimator

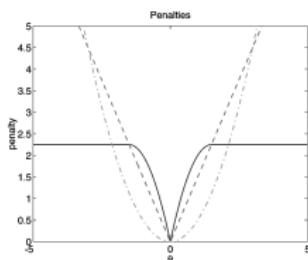
$$(\hat{\beta}_{\tau_n})_S = \tilde{\beta}_S^* - \frac{1}{\tau_n} \Sigma_{n,S}^{-1} \text{sign}(\beta_S^*), \quad \tau_n > 0.$$

How to remove the bias and return the Oracle Estimator?

Nonconvex Regularization?

- To reduce bias, **non-convex** regularization was proposed ([Fan-Li's SCAD](#), [Zhang's MPLUS](#), [Zou's Adaptive LASSO](#), l_q ($q < 1$), etc.)

$$\min_{\beta} \sum_i p(|\beta_i|) + \frac{t}{2n} \|y - X\beta\|_2^2.$$



- Yet it is generally hard to locate the **global optimizer**
- Any other simple scheme?*

New Idea

- LASSO:

$$\min_{\beta} \|\beta\|_1 + \frac{t}{2n} \|y - X\beta\|_2^2.$$

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$$\Rightarrow \rho_t = \frac{1}{n} X^T (y - X\beta_t) t$$

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- Taking derivative (assuming differentiability) w.r.t. t

$$\Rightarrow \dot{\rho}_t = \frac{1}{n} X^T (y - X(\dot{\beta}_t t + \beta_t)), \quad \rho_t \in \partial \|\beta_t\|_1$$

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- Assuming sign-consistency in a neighborhood of τ_n ,

for $i \in S$, $\rho_{\tau_n}(i) = \text{sign}(\beta^*(i)) \in \pm 1 \Rightarrow \dot{\rho}_{\tau_n}(i) = 0$,

$$\Rightarrow \dot{\beta}_{\tau_n} \tau_n + \beta_{\tau_n} = \tilde{\beta}^*$$

New Idea

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for $i \in S$, $\rho_{\tau_n}(i) = \text{sign}(\beta^*(i)) \in \pm 1 \Rightarrow \dot{\rho}_{\tau_n}(i) = 0$,

$$\Rightarrow \dot{\beta}_{\tau_n} \tau_n + \beta_{\tau_n} = \tilde{\beta}^*$$

- Equivalently, the blue part removes bias of LASSO automatically

$$\beta_{\tau_n}^{lasso} = \tilde{\beta}^* - \frac{1}{\tau_n} \Sigma_n^{-1} \text{sign}(\beta^*) \Rightarrow \dot{\beta}_{\tau_n}^{lasso} \tau_n + \beta_{\tau_n}^{lasso} = \tilde{\beta}^* (\text{oracle})!$$

Differential Inclusion: Inverse Scaled Spaces (ISS)

Differential inclusion replacing $\dot{\beta}_{\tau_n}^{lasso} \tau_n + \beta_{\tau_n}^{lasso}$ by β_t

$$\dot{\rho}_t = \frac{1}{n} X^T (y - X \beta_t), \quad (3a)$$

$$\rho_t \in \partial \|\beta_t\|_1. \quad (3b)$$

starting at $t = 0$ and $\rho(0) = \beta(0) = \mathbf{0}$.

- Replace ρ/t in LASSO KKT by $d\rho/dt$

$$\frac{\rho_t}{t} = \frac{1}{n} X^T (y - X \beta_t)$$

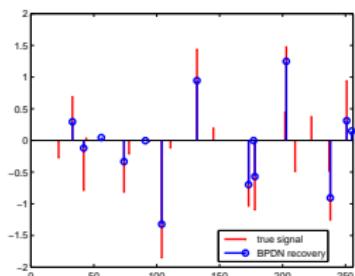
- Burger-Gilboa-Osher-Xu'06 (in image recovery it recovers the objects in an inverse-scale order as t increases (larger objects appear in β_t first))

Examples

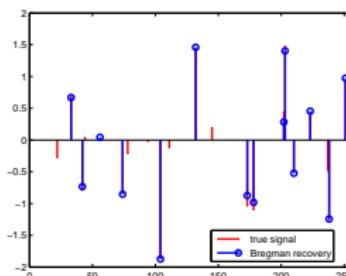
- e.g. $X = Id$, $n = p = 1$, hard-thresholding

$$\beta_\tau = \begin{cases} 0, & \text{if } \tau < 1/(\tilde{\beta}^*); \\ \tilde{\beta}^*, & \text{otherwise,} \end{cases}$$

- the same example shown before (figures by courtesy of Wotao Yin)



True vs LASSO



True vs ISS

Solution Path: Sequential Restricted Maximum Likelihood Estimate

- ρ_t is piece-wise linear in t ,

$$\rho_t = \rho_{t_k} + \frac{t - t_k}{n} X^T (y - X\beta_{t_k}), \quad t \in [t_k, t_{k+1})$$

where $t_{k+1} = \sup\{t > t_k : \rho_{t_k} + \frac{t-t_k}{n} X^T (y - X\beta_{t_k}) \in \partial \|\beta_{t_k}\|_1\}$

- β_t is piece-wise constant in t : $\beta_t = \beta_{t_k}$ for $t \in [t_k, t_{k+1})$ and $\beta_{t_{k+1}}$ is the sequential restricted Maximum Likelihood Estimate by solving nonnegative least square (Burger et al.'13; Osher et al.'16)

$$\begin{aligned} \beta_{t_{k+1}} &= \arg \min_{\beta} \|y - X\beta\|_2^2 \\ \text{subject to } & (\rho_{t_{k+1}})_i \beta_i \geq 0 \quad \forall i \in S_{k+1}, \\ & \beta_j = 0 \quad \forall j \in T_{k+1}. \end{aligned} \tag{4}$$

- Note: Sign consistency $\rho_t = \text{sign}(\beta^*) \Rightarrow \beta_t = \tilde{\beta}^*$ the oracle estimator

Example: Regularization Paths of LASSO vs. ISS

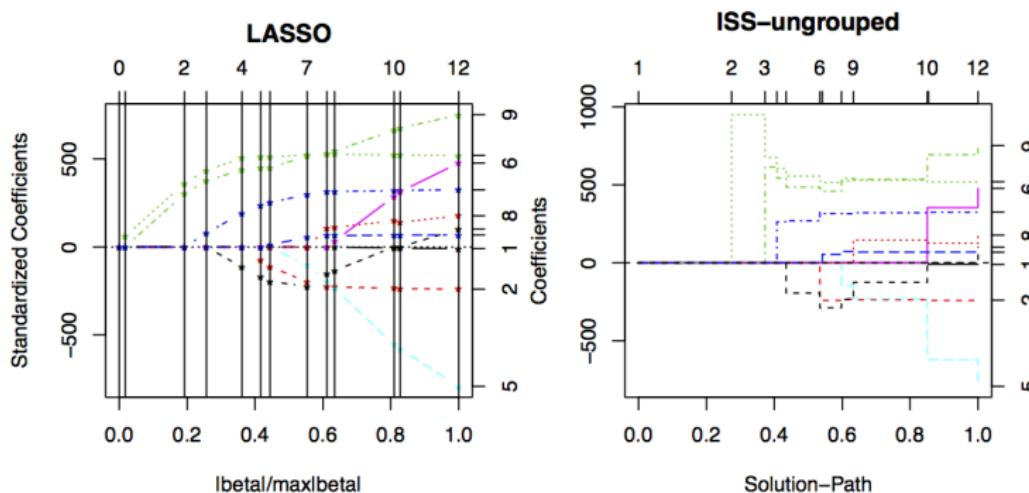


Figure: Diabetes data (Efron et al.'04) and regularization paths are different, yet bearing similarities on the order of parameters being nonzero

Why? A Path Consistency Theory

Our aim is to show that under nearly the **same** conditions for sign-consistency of LASSO, there exists points on their paths $(\beta(t), \rho(t))_{t \geq 0}$, which are

- **sparse**
- **sign-consistent** (the same sparsity pattern of nonzeros as true signal)
- **the oracle estimator** which is unbiased, better than the LASSO estimate.
- **Early stopping** regularization is necessary to prevent overfitting noise!

Assumptions

(A1) **Restricted Strongly Convex**: $\exists \gamma \in (0, 1]$,

$$\frac{1}{n} X_S^T X_S \geq \gamma I$$

(A2) **Incoherence/Irrepresentable Condition**: $\exists \eta \in (0, 1)$,

$$\left\| \frac{1}{n} X_T^T X_S^\dagger \right\|_\infty = \left\| \frac{1}{n} X_T^T X_S \left(\frac{1}{n} X_S^T X_S \right)^{-1} \right\|_\infty \leq 1 - \eta$$

- "Irrepresentable" means that one can not represent (regress) column vectors in X_T by covariates in X_S .
- The incoherence/irrepresentable condition is used independently in [Tropp'04](#), [Yuan-Lin'05](#), [Zhao-Yu'06](#), [Zou'06](#), [Wainwright'09](#), etc.

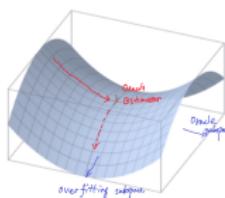
Understanding the Dynamics

ISS as **restricted gradient descent**:

$$\dot{\beta}_t = -\nabla L(\beta_t) = \frac{1}{n} X^T (y - X\beta_t), \quad \rho_t \in \partial \|\beta_t\|_1$$

such that

- **incoherence condition** and **strong signals** ensure it firstly evolves on index set S (Oracle Subspace) to reduce the loss
- **strongly convex** in subspace restricted on index set $S \Rightarrow$ fast decay in loss
- **early stopping** after all strong signals are detected, before overfitting noise



Path Consistency

Theorem (Osher-Ruan-Xiong-Y.-Yin'2016)

Assume (A1) and (A2). Define an early stopping time

$$\bar{\tau} := \frac{\eta}{2\sigma} \sqrt{\frac{n}{\log p}} \left(\max_{j \in T} \|X_j\| \right)^{-1},$$

and the smallest magnitude $\beta_{\min}^* = \min(|\beta_i^*| : i \in S)$. Then

- **No-false-positive:** for all $t \leq \bar{\tau}$, the path has no-false-positive with high probability, $\text{supp}(\beta(t)) \subseteq S$;
- **Consistency:** moreover if the signal is strong enough such that

$$\beta_{\min}^* \geq \left(\frac{4\sigma}{\gamma^{1/2}} \vee \frac{8\sigma(2 + \log s)(\max_{j \in T} \|X_j\|)}{\gamma\eta} \right) \sqrt{\frac{\log p}{n}},$$

there is $\tau \leq \bar{\tau}$ such that solution path $\beta(t) = \tilde{\beta}^*$ for every $t \in [\tau, \bar{\tau}]$.

Note: equivalent to LASSO with $\lambda^* = 1/\bar{\tau}$ (Wainwright'09) up to $\log s$.

Large scale algorithm: Linearized Bregman Iteration

Damped Dynamics: continuous solution path

$$\dot{\rho}_t + \frac{1}{\kappa} \dot{\beta}_t = \frac{1}{n} X^T (y - X\beta_t), \quad \rho_t \in \partial \|\beta_t\|_1. \quad (5)$$

Linearized Bregman Iteration as forward Euler discretization proposed even earlier than ISS dynamics (Osher-Burger-Goldfarb-Xu-Yin'05, Yin-Osher-Goldfarb-Darbon'08): for $\rho_k \in \partial \|\beta_k\|_1$,

$$\rho_{k+1} + \frac{1}{\kappa} \beta_{k+1} = \rho_k + \frac{1}{\kappa} \beta_k + \frac{\alpha_k}{n} X^T (y - X\beta_k), \quad (6)$$

where

- Damping factor: $\kappa > 0$
- Step size: $\alpha_k > 0$ s.t. $\alpha_k \kappa \|\Sigma_n\| \leq 2$
- Moreau Decomposition: $z_k := \rho_k + \frac{1}{\kappa} \beta_k \Leftrightarrow \beta_k = \kappa \cdot \text{Shrink}(z_k, 1)$

Comparison with ISTA

Linearized Bregman (LB) iteration:

$$z_{t+1} = z_t - \alpha_t X^T (\kappa X \textcolor{red}{Shrink}(z_t, 1) - y)$$

which is not **ISTA**:

$$z_{t+1} = \textcolor{red}{Shrink}(z_t - \alpha_t X^T (Xz_t - y), \lambda).$$

Comparison:

- **ISTA:**

- as $t \rightarrow \infty$ solves **LASSO**: $\frac{1}{n} \|y - X\beta\|_2^2 + \lambda \|\beta\|_1$
- **parallel run** ISTA with $\{\lambda_k\}$ for LASSO regularization paths

- **LB: a single run** generates the whole regularization path at same cost of ISTA-LASSO estimator for a fixed regularization

LBI generates regularization paths

$n = 200, p = 100, S = \{1, \dots, 30\}, x_i \sim N(0, \Sigma_p)$ ($\sigma_{ij} = 1/(3p)$ for $i \neq j$ and 1 otherwise)

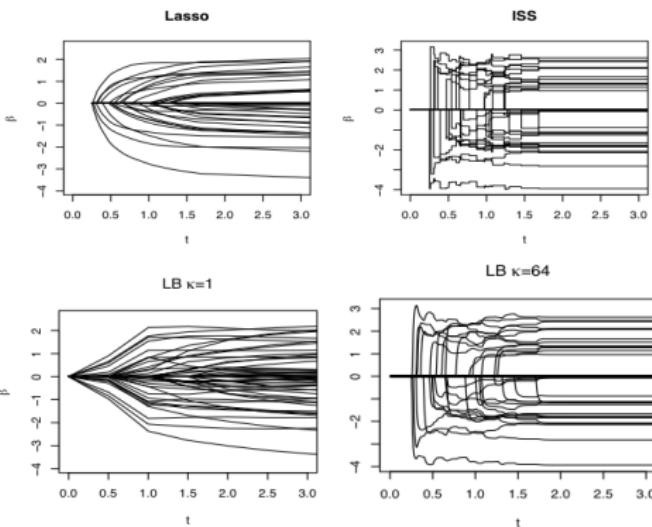


Figure: As $\kappa \rightarrow \infty$, LB paths have a limit as piecewise-constant ISS path

General Loss and Regularizer

$$\dot{\eta}_t = -\frac{\kappa_0}{n} \sum_{i=1}^n \nabla_{\eta} \ell(x_i, \theta_t, \eta_t) \quad (7a)$$

$$\dot{\rho}_t + \frac{\dot{\theta}_t}{\kappa_1} = -\frac{1}{n} \sum_{i=1}^n \nabla_{\theta} \ell(x_i, \theta_t, \eta_t) \quad (7b)$$

$$\rho_t \in \partial \|\theta_t\|_* \quad (7c)$$

where

- $\ell(x_i, \theta)$ is a loss function: negative logarithmic likelihood, non-convex loss (neural networks), etc.
- $\|\theta_t\|_*$ is the Minkowski-functional (gauge) of dictionary convex hulls:

$$\|\theta\|_* := \inf\{\lambda \geq 0 : \theta \in \lambda K\}, \quad K \text{ is a symmetric convex hull of } \{a_i\}$$

- it can be generalized to non-convex regularizers

More reference on generalizations

- **Logistic Regression:** loss – conditional likelihood, regularizer – ℓ_1
([Shi-Yin-Osher-Sajda'10, Huang-Yao'18](#))
- **Graphical Models** (Gaussian/Ising/Potts Model): loss – likelihood, composite conditional likelihood, regularizer – ℓ_1 and group ℓ_1
([Huang-Yao'18](#))
- **Fused LASSO/TV:** split Bregman with composite ℓ_2 loss and ℓ_1 gauge
([Osher-Burger-Goldfarb-Xu-Yin'06, Burger-Gilboa-Osher-Xu'06, Yin-Osher-Goldfarb-Darbon'08, Huang-Sun-Xiong-Yao'16](#))
- **Matrix Completion/Regression:** gauge – the matrix nuclear norm
([Cai-Candès-Shen'10](#))

Structural or Transformational Sparsity

Structural/Transformational Sparse Regression:

$$y = X\beta^* + \epsilon, \quad (8a)$$

$$\gamma^* = D\beta^*, \quad (8b)$$

where

$$S = \text{supp}(\gamma^*), \quad s := |S| \ll p.$$

Split LBI vs. Generalized LASSO

- Generalized LASSO (genlasso):

$$\arg \min_{\beta} \left(\frac{1}{2n} \|y - X\beta\|_2^2 + \lambda \|\mathcal{D}\beta\|_1 \right). \quad (9)$$

- Split LBI: Loss that splits prediction vs. sparsity control

$$\ell(\beta, \gamma) := \frac{1}{2n} \|y - X\beta\|_2^2 + \frac{1}{2\nu} \|\gamma - \mathcal{D}\beta\|_2^2 \quad (\nu > 0). \quad (10)$$

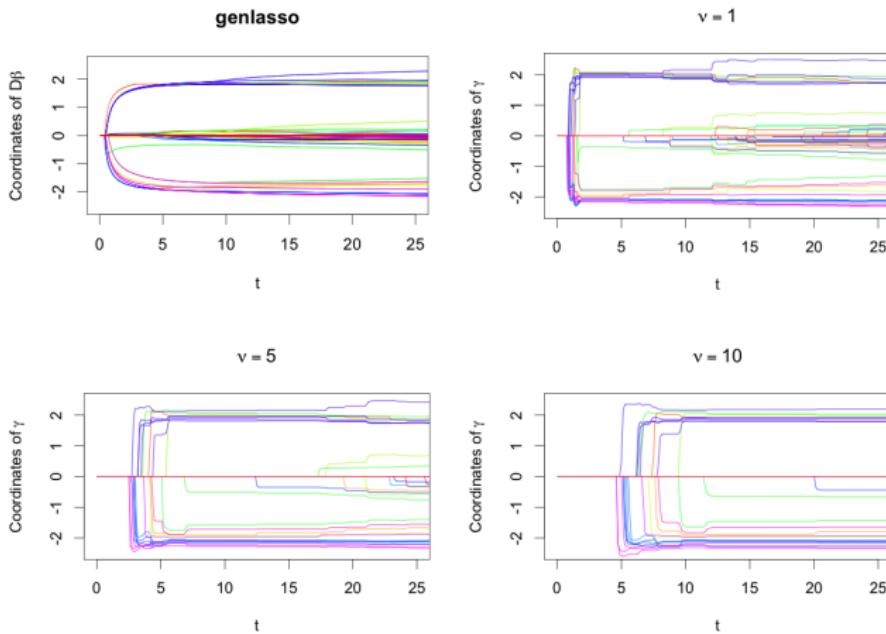
Algorithm [Huang-Sun-Xiong-Y. 2016] :

$$\beta_{k+1} = \beta_k - \kappa \alpha \nabla_{\beta} \ell(\beta_k, \gamma_k), \quad (11a)$$

$$z_{k+1} = z_k - \alpha \nabla_{\gamma} \ell(\beta_k, \gamma_k), \quad (11b)$$

$$\gamma_{k+1} = \kappa \cdot \text{prox}_{\|\cdot\|_1}(z_{k+1}), \quad (11c)$$

Split LBI vs. Generalized LASSO paths



Split LBI may beat Generalized LASSO in Model Selection

genlasso	Split LBI			genlasso	Split LBI		
	$\nu = 1$	$\nu = 5$	$\nu = 10$		$\nu = 1$	$\nu = 5$	$\nu = 10$
.9426 (.0390)	.9845 (.0185)	.9969 (.0065)	.9982 (.0043)	.9705 (.0212)	.9955 (.0056)	.9996 (.0014)	.9998 (.0009)

- Example: $n = p = 50$, $X \in \mathbb{R}^{n \times p}$ with $X_j \sim N(0, I_p)$, $\epsilon \sim N(0, I_n)$
- (Left) $D = I$ (LASSO vs. Split LBI)
- (Right) 1-D fused (generalized) LASSO vs. **Split LBI**
- In terms of Area Under the ROC Curve (AUC), Split LBI has less false discoveries than genlasso
- *Why?* Split LBI may need **weaker** irrepresentable conditions than generalized LASSO...

Remark: Identifiable Conditions (IC)

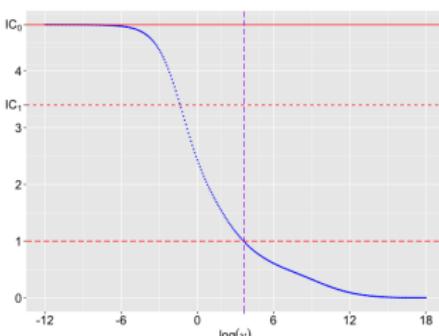
- Let the columns of W form an orthogonal basis of $\ker(D_{S^c})$.

$$\Omega^S := \left(D_{S^c}^\dagger\right)^T \left(X^* X W \left(W^T X^* X W\right)^\dagger W^T - I\right) D_S^T, \quad (14)$$

$$\mathbf{IC_0} := \left\| \Omega^S \right\|_\infty, \quad \mathbf{IC_1} := \min_{u \in \ker(D_{S^c})} \left\| \Omega^S \text{sign}(D_S \beta^*) - u \right\|_\infty. \quad (15)$$

- The sign consistency of genlasso has been proved, under $\mathbf{IC_1} < 1$ [Vaiter et al. 2013].
- The sign consistency of Split LBI is proved under $\text{IRR}(\nu) < 1$ [Huang-Sun-Xiong-Y.'2016].
- As $\text{IRR}(\nu) < \mathbf{IC_1}$ when ν grows, our IRR is easier to meet.

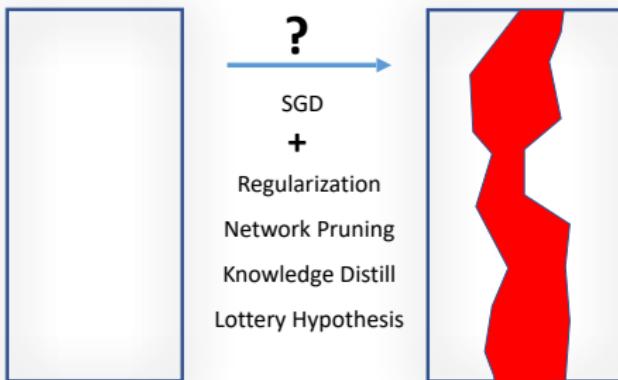
Split LBI improves Irrepresentable Condition



Theorem ([Huang-Sun-Xiong-Y.'2020](#))

- $IC_0 \geq IC_1$.
- $IRR(\nu) \rightarrow IC_0 (\nu \rightarrow 0)$.
- $IRR(\nu) \rightarrow C (\nu \rightarrow \infty)$ with $C = 0 \iff \ker(X) \subseteq \ker(D_S)$.

Lottery Ticket Hypothesis: Learning Sparse Subnetworks



Existing 2-Stage Approaches:

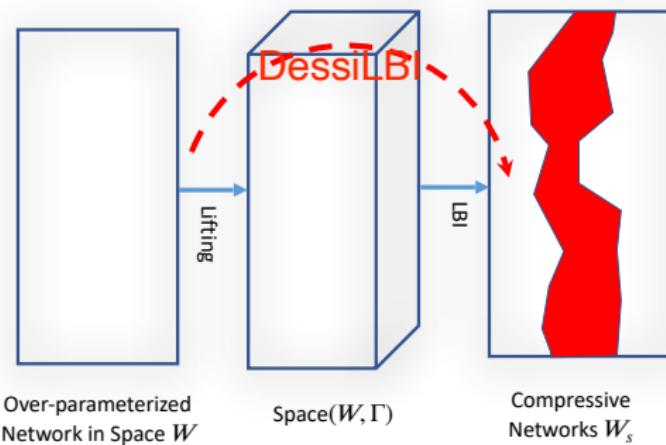
- Fully training dense network,
- Finding good sparse subnets.

Our method: 1-Stage Approach (end-to-end)
Without fully training a dense model.

Over-parameterized
Networks W

Compressive
Networks W_s

The key idea of DessiLBI: Lifting Parameters



- 1, Lifting parameter space W to (W, Γ) coupling the **inverse scale space**.
- 2, Γ learns **structural sparsity** in inverse scale space.
- 3, Network's solution path in (W, Γ) as the discretization of dynamics, and solved by LBI.
- 4, Our optimizer enjoys a provable global convergence guarantee.

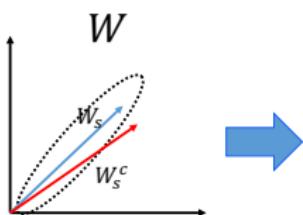
LOSS

W_s : wining ticket

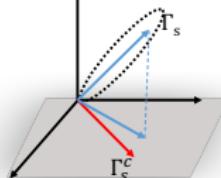
W : dense parameters

$$W_s \cup W_{s^c} = W$$

W_s and W_{s^c} correlated



$$(W, \Gamma)$$



$$\Gamma_s \cup \Gamma_{s^c} = \Gamma$$

Γ_s and Γ_{s^c} orthogonal

Differential
Inclusion of Inverse
Scale Space

GD of Weight Space

$$\dot{W}_t = -\nabla_W \hat{\mathcal{L}}_n(W)$$



$$\bar{\mathcal{L}}(W, \Gamma) = \hat{\mathcal{L}}_n(W) + \boxed{\frac{1}{2\nu} \|W - \Gamma\|_F^2}$$

DessiLBI Algorithm

$$\frac{\dot{W}_t}{\kappa} = -\nabla_W \bar{\mathcal{L}}(W_t, \Gamma_t)$$

$$\dot{V}_t = -\nabla_{\Gamma} \bar{\mathcal{L}}(W_t, \Gamma_t)$$

$$V_t \in \partial \left(\Omega(\Gamma) + \frac{1}{2\kappa} \|\Gamma\|^2 \right)$$

$$\bar{\mathcal{L}}(W, \Gamma) = \hat{\mathcal{L}}_n(W) + \frac{1}{2\nu} \|W - \Gamma\|_F^2$$

Differential Inclusion of Inverse Scale Space

$$\begin{aligned} W_{k+1} &= W_k - \kappa \alpha_k \cdot \nabla_W \bar{\mathcal{L}}(W_k, \Gamma_k), \\ V_{k+1} &= V_k - \alpha_k \cdot \nabla_{\Gamma} \bar{\mathcal{L}}(W_k, \Gamma_k), \\ \Gamma_{k+1} &= \kappa \cdot \text{Prox}_{\Omega_\lambda}(V_{k+1}) \end{aligned}$$

The Simple Discretization - DessLBI

$$\begin{aligned} W_{k+1} &= W_k - \kappa \alpha_k \cdot \nabla_W \bar{\mathcal{L}}(W_k, \Gamma_k), \\ V_{k+1} &= V_k - \alpha_k \cdot \nabla_{\Gamma} \bar{\mathcal{L}}(W_k, \Gamma_k) \end{aligned}$$

Gradient Descent

$$\begin{aligned} \Gamma_{k+1} &= \kappa \cdot \text{Prox}_{\Omega}(V_{k+1}) \\ \text{Prox}_{\Omega}(V) &= \arg \min_{\Gamma} \left\{ \frac{1}{2} \|\Gamma - V\|_2^2 + \Omega(\Gamma) \right\} \end{aligned}$$

Proximal Mapping

V is the sub-gradient for some sparsity-enforced, often non-differentiable regularization $\Omega_\lambda(\Gamma) = \lambda \Omega_1(\Gamma)$, ($\lambda \in \mathbb{R}_+$) such as Lasso or group Lasso penalties for $\Omega_1(\Gamma)$

- W_t follows the gradient descent with ℓ_2 -regularization
- Important Features of Γ_t are first selected: *Inverse Scale Space*

Proximal Map for Sparse Structure

$$\text{Prox}_{\Omega}(V) = \arg \min_{\Gamma} \left\{ \frac{1}{2} \|\Gamma - V\|_2^2 + \Omega(\Gamma) \right\}$$

DessLBI enforce structural sparsity by Group lasso penalty,

$$\Omega(\Gamma) = \sum_g \|\Gamma^g\|_2 = \sum_g \sqrt{\sum_{i=1}^{|\Gamma^g|} (\Gamma_i^g)^2}$$

A close form solution:

$$\Gamma^g = \kappa \cdot \max(0, 1 - 1/\|V^g\|_2) V^g$$

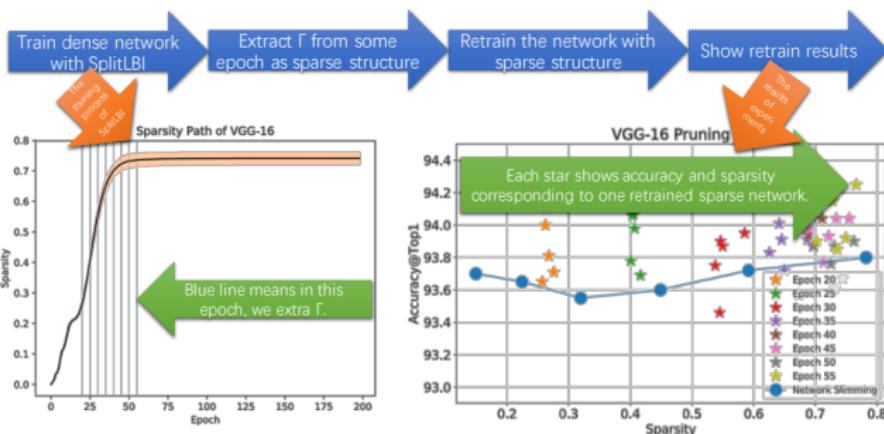
Convolutional layer, $\Gamma^g = \Gamma^g(c_{in}, c_{out}, \text{size})$ c_{in} : No. of input channel

Fully connected layer $\Gamma = \Gamma(c_{in}, c_{out})$ c_{out} : No. of output channel
size: kernel size

- (Batch) DessLBI w./o. Momentum and Weight-decay (Mom-Wd)
- We have a theorem that guarantees the **global convergence** of DessLBI: **from any initialization**, DessLBI sequence converges to a critical point.

A Flow Chart for Finding Sparse Structure

Experiments Design



Summary

- The limit of Linearized Bregman iteration follows **differential inclusion of inverse scale space**, where significant features emerge earlier on solution paths
- It renders the **unbiased Oracle Estimator** under sign-consistency
- Sign consistency under nearly the **same** condition as LASSO
 - Restricted Strongly Convex + Irrepresentable Condition
- **Split** extension: sign consistency under a **weaker** condition than generalized LASSO
 - under a provably weaker Irrepresentable Condition
- **Early stopping** regularization is exploited against overfitting noise

A Renaissance of Boosting as restricted gradient descent ...

Some Reference

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- R package: <http://cran.r-project.org/web/packages/Libra/index.html>

Cran R package: Libra

<http://cran.r-project.org/web/packages/Libra/>



CRAN - Package Libra

<https://cran.r-project.org/web/packages/Libra/index.html>

Libra: Linearized Bregman Algorithms for Generalized Linear Models

Efficient procedures for fitting the regularization path for linear, binomial, multinomial, Ising and Potts models with lasso, group lasso or column lasso(only for multinomial) penalty. The package uses Linearized Bregman Algorithm to solve the regularization path through iterations. Bregman Inverse Scale Space Differential Inclusion solver is also provided for linear model with lasso penalty.

Version: 1.5
Depends: R (\geq 3.0), [mnl](#)
Suggests: [lars](#), [MASS](#), [igraph](#)
Published: 2016-02-17
Author: Feng Ruan, Jiechao Xiong and Yuan Yao
Maintainer: Jiechao Xiong <xiongjiechao@pku.edu.cn>
License: [GPL-2](#)
URL: <http://arxiv.org/abs/1406.7728>
NeedsCompilation: yes
SystemRequirements: GNU Scientific Library (GSL)
CRAN checks: [Libra results](#)

Downloads:

Reference manual: [Libra.pdf](#)
Package source: [Libra_1.5.tar.gz](#)
Windows binaries: r-devel: [Libra_1.5.zip](#), r-release: [Libra_1.5.zip](#), r-oldrel: [Libra_1.5.zip](#)
OS X Snow Leopard binaries: r-release: [Libra_1.5.tgz](#), r-oldrel: not available
OS X Mavericks binaries: r-release: [Libra_1.5.tgz](#)
Old sources: [Libra archive](#)



Libra (1.6) currently includes

Sparse statistical models:

- linear regression: ISS (differential inclusion), LBI
- logistic regression (binomial, multinomial): LBI
- graphical models (Gaussian, Ising, Potts): LBI

Two types of regularization:

- LASSO: l_1 -norm penalty
- Group LASSO: $l_2 - l_1$ penalty

Libra computes regularization paths via Linearized Bregman Iteration (LBI)

for $\theta_0 = z_0 = \mathbf{0}$ and $k \in \mathbb{N}$,

$$z_{k+1} = z_k - \frac{\alpha_k}{n} \sum_{i=1}^n \nabla_{\theta} \ell(x_i, \theta_k) \quad (16a)$$

$$\theta_{k+1} = \kappa \cdot \text{prox}_{\|\cdot\|_*}(z_{k+1}) \quad (16b)$$

where

- $\ell(x, \theta)$ is the *loss* function to minimize
- $\text{prox}_{\|\cdot\|_*}(z) := \arg \min_u \left(\frac{1}{2} \|u - z\|^2 + \|u\|_* \right)$
- $\alpha_k > 0$ is step-size
- $\kappa > 0$ while $\alpha_k \kappa \|\nabla_{\theta}^2 \hat{\mathbb{E}} \ell(x, \theta)\| < 2$
- as simple as ISTA (easy to parallel implementation), yet different limit dynamics

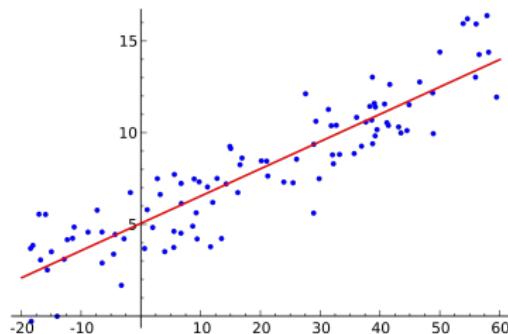
Linear Regression

Linear Regression:

$$y = X\beta + \epsilon$$

β is sparse or group sparse, with two types of penalty:

- "ungrouped": $\sum_i |\beta_i|$
- "grouped": $\sum_g \sqrt{\sum_{g_i=g} \beta_i^2}$

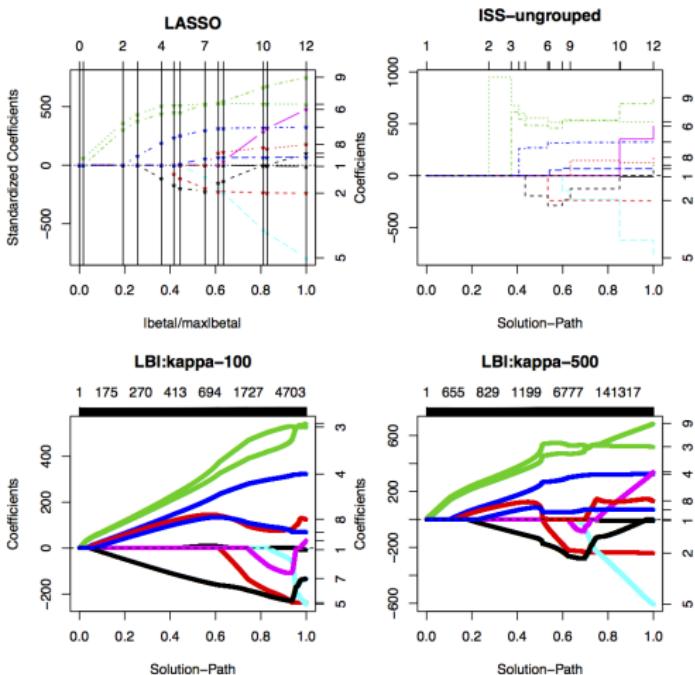


Linear Regression Example: Diabetes Data

```
data('diabetes')
attributes(x)
##$dim
# [1] 442   10
##$dimnames[[2]]
# [1] "age" "sex" "bmi" "map" "tc"   "ldl" "hdl" "tch" "ltg" "glu"

lassopath = lars(x,y)
isspath = iss(x,y)
lb(x,y,kappa=100,alpha=0.005,family="gaussian",group="ungrouped",
    intercept=FALSE,normalize=FALSE)
lb(x,y,kappa=500,alpha=0.001,family="gaussian",group="ungrouped",
    intercept=FALSE,normalize=FALSE)
```

LBI generates iterative regularization paths



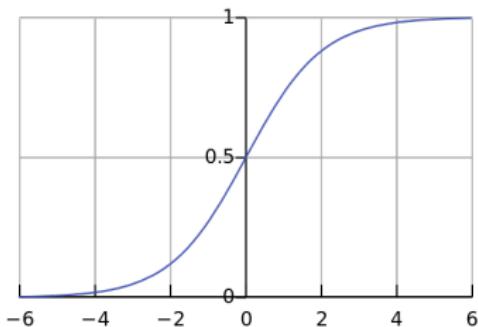
Logistic Regression

Logistic Regression:

$$\log \frac{P(y = 1|X)}{P(y = -1|X)} = X\beta \Leftrightarrow P(y = 1|X) = \frac{e^{X\beta}}{1 + e^{X\beta}} =: \sigma(X\beta)$$

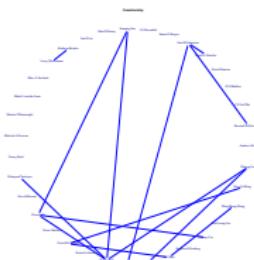
β is sparse or group sparse, with two types of penalty:

- "ungrouped": $\sum_i |\beta_i|$
- "grouped": $\sum_g \sqrt{\sum_{g_i=g} \beta_i^2}$



Example: Publications of COPSS Award Winners

- dataset is provided by Prof. [Jiashun Jin](#) @CMU
- 3248 papers by 3607 authors between 2003 and the first quarter of 2012 from:
 - the Annals of Statistics, Journal of the American Statistical Association, Biometrika and Journal of the Royal Statistical Society Series B
- a subset of 382 papers by 35 COPSS award winners
- Question: can we **model the coauthorship structure to predict the out-of-sample behavior?**



A logistic regression path with early stopping regularization

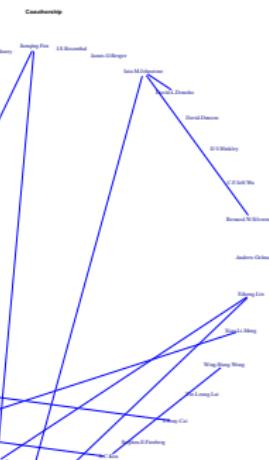
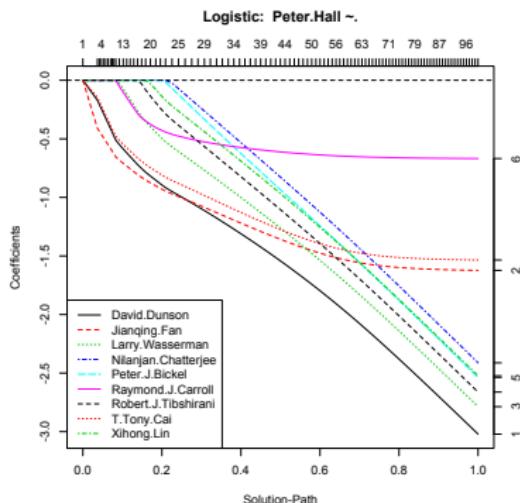


Figure: Peter Hall vs. other COPSS award winners in sparse logistic regression [papers from AoS/JASA/Biometrika/JRSSB, 2003-2012]: true coauthors are merely Tony Cai, R.J. Carroll, and J. Fan

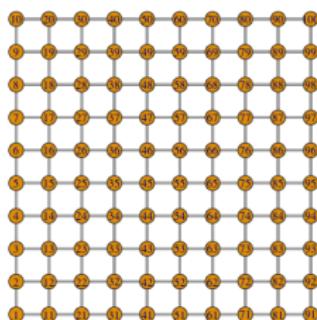
Sparse Ising Model

All models are wrong, but some are useful (George Box):

$$P(x_1, \dots, x_p) \sim \exp \left(\sum_i H_i x_i + \sum_{i,j} J_{ij} x_i x_j \right)$$

- Ising model: $x_i = 1$ if author i appears in a paper, otherwise 0
- H_i describes the mean publication rate of author i
- J_{ij} describes the interactions between author i and j
 - $J_{ij} > 0$: author i and j collaborate more often than others
 - $J_{ij} < 0$: author i and j collaborate less frequently than others
 - sparsity: $J_{ij} = 0$ mostly, a model of collaboration network
 - learned by maximum composite conditional likelihood with LB

Early stopping against overfitting in sparse Ising model learning



a true Ising model of 2-D grid

a movie of LB path

Application: Sparse Ising Model of COPSS Award Winners

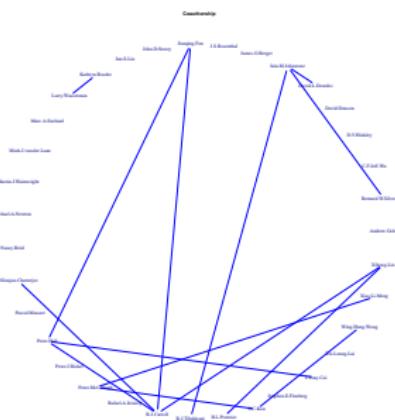


Figure: Left: LB path of Ising Model learning; Right: coauthorship network of existing data. Typically COPSS winners do not like working together; Peter Hall (1951-2016) is the hub of statisticians, like Erdős for mathematicians