MT.Zunqiu.Wang

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```
library(reshape2)
```

Q1

```
prime <- function(num) {</pre>
  lst <- c(2) # initialize a vector containing 2</pre>
  sq \leftarrow seq(2, num)
  for (i in sq) {
    div <- seq(2, sq[i-1]) # since starting 2 so the biggest divider less than
    # i is i-1 position of sq
    rem <- sq[i] %% div # test reminder to loop every sq elements
    if ((0 %in% rem) == FALSE) { # check reminder if 0
      lst[i] <- sq[i] # added to lst if it has reminder after division</pre>
    } else {
      next
    }
    lst <- lst[!is.na(lst)] # get rid of NAs where reminder is 0</pre>
  return(lst)
  }
prime(100)
```

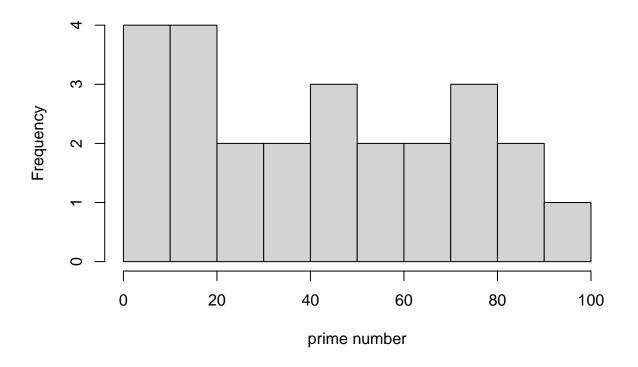
[1] 2 3 5 7 11 13 17 19 23 29 31 37 41 43 47 53 59 61 67 71 73 79 83 89 97

```
prime.list <- prime(100)</pre>
```

Q2

```
# distribution of first 100 prime numbers
hist(prime.list, breaks = 10, xlab = "prime number", main = "histogram of first 100 prime numbers")
```

histogram of first 100 prime numbers



among which given first is head at least 3 or more heads in a row are: (HHHTT), (HHHHTH), (HHHHHH), (HTHHHH)

so prob(more than 3 heads in a row given first is head)=(5/32)/(1/2)=5/16

Q4

$$P(A|+) = \frac{P(+|A)P(A)}{P(+|A)P(A) + P(+|A')P(A')}$$

```
# A = hit
sensitivity <- 0.99
fp <- 0.01
pA <- 1/100000

prob <- function(sensitivity, fp, pA) {
   prb <- (sensitivity *pA)/(sensitivity *pA + fp * (1-pA))
   return(prb)
}
prob(sensitivity, fp, pA)</pre>
```

```
## [1] 0.0009890307
Q_5
# for 5 or more we want upper tail
ppois(5,1,F)
## [1] 0.0005941848
Q6
H_0 =the average college student get 7 hours of sleep a night H_a =the average college student does not get
7 hours of sleep a night
# create a function to calculate all statistics
sample \leftarrow c(7,6,5,8,6,6,4,5,8,7)
x.bar <- mean(sample)</pre>
mu <- 7
sd <- sd(sample)</pre>
CI <- 0.95
n <- 10
desc.val.table <- function(x.bar, mu, sd, n, CI) {</pre>
  se <- sd/sqrt(n)
 t.stat <- (x.bar-mu)/se # t stat
 t.crit.perc.low <- (1-CI)/2 # % for lower bound 2 tail
 t.crit.perc.high <- (1-CI)/2 + CI # % for higher bound 2 tail
  t.crit.2.tail.low <- qt(t.crit.perc.low , n-1) # t critical for lower bound 2 tail
  t.crit.2.tail.high <- qt(t.crit.perc.high, n-1) # t critical for higher bound 2 tail
  CI.low.2.tail <- x.bar - t.crit.2.tail.high*se # CI lower bound for 2 tail
  CI.high.2.tail <- x.bar + t.crit.2.tail.high*se # CI higher bound for 2 tail
  p.val.2.tail \leftarrow 2*(1- pt(-1 * t.stat, n-1)) # negative t stat has to be
  # converted to positive for two tail
  df <- data.frame(t.stat=t.stat,</pre>
                    t.crit.2.tail.low=t.crit.2.tail.low,
                    t.crit.2.tail.high=t.crit.2.tail.high, CI.low.2.tail=CI.low.2.tail,
                    CI.high.2.tail=CI.high.2.tail,
                    p.val.2.tail=p.val.2.tail)
  df <- melt(df)
 return(df)
}
desc.val.table(x.bar, mu, sd, n, CI)
## No id variables; using all as measure variables
##
               variable
                               value
                 t.stat -1.92153785
## 2 t.crit.2.tail.low -2.26215716
## 3 t.crit.2.tail.high 2.26215716
## 4
          CI.low.2.tail 5.25818887
## 5
         CI.high.2.tail 7.14181113
```

p.val.2.tail 0.08684229

6

at 95% CI, we fail to reject null hypothesis since p val(0.08) > p val threshold(0.05)

```
Q7
t_{stat} = \frac{\bar{x} - \mu}{\frac{s}{\sqrt{n}}} < -2.26215716 -2.26215716 = \frac{6.2 - 7}{\frac{1.316561}{1.316561}}  n = 13.8595
so 14-10=4, we need 4 more people
Q8
# paired t test
control \leftarrow c(7,6,5,8,6,6,4,5,8,7)
\exp < -c(5,4,5,7,5,4,5,4,6,5)
n <- 10
mean.diff <- mean(control- exp)</pre>
sd.diff <- sd(control- exp)</pre>
se.diff <- sd.diff/sqrt(n)</pre>
t.stat <- mean.diff/se.diff</pre>
t.stat
## [1] 3.674235
t.crit \leftarrow qt(c(0.025,0.975), n-1)
## [1] -2.262157 2.262157
p.val <- 2*pt(t.stat, n-1, lower.tail = FALSE)</pre>
p.val
## [1] 0.005121073
t.test(control, exp, paired = TRUE, alternative = "two.sided")
##
## Paired t-test
##
## data: control and exp
## t = 3.6742, df = 9, p-value = 0.005121
\#\# alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## 0.4611826 1.9388174
## sample estimates:
## mean of the differences
##
                           1.2
```

Q9

t stat is greater than t crit, we reject null hypothesis

```
# chi sq calculations
df9 <- data.frame(condition=c("treatment", "control"), live=c(4,8), die=c(11,7))</pre>
sum.live <- colSums(df9[,2:3])[1]</pre>
sum.die <- colSums(df9[,2:3])[2]</pre>
sum.trt <- rowSums(df9[,2:3])[1]</pre>
sum.ctr <- rowSums(df9[,2:3])[2]</pre>
total <- sum.trt + sum.ctr</pre>
e.liv.trt <- sum.live * sum.trt/total
e.liv.ctr <- sum.live * sum.ctr/total
e.die.trt <- sum.die * sum.trt/total
e.die.ctr <- sum.die * sum.ctr/total</pre>
chisq.test(df9[,2:3])
##
    Pearson's Chi-squared test with Yates' continuity correction
##
## data: df9[, 2:3]
## X-squared = 1.25, df = 1, p-value = 0.2636
qchisq(0.95,1)
## [1] 3.841459
1-pchisq(1.25,1)
## [1] 0.2635525
```

Based on 95% CI, chi² stat(1.25) < chi crit(3.84) so we conclude that plants live or die is not dependent on what they are watered with.

$$\chi^2 = \sum \frac{(f_o - f_e)^2}{f_e}$$

$$\chi^2 = \frac{(4-6)^2}{6} + \frac{(8-6)^2}{6} + \frac{(11-9)^2}{9} + \frac{(7-9)^2}{9} = 1.25$$

$$df = (r-1)(c-1) = 1 * 1 = 1$$
Q10 $H_0 = \mu_w = \mu_v = \mu_c$
 H_a =at least one is different

F-stat= $\frac{\text{average variance between groups}}{\text{average variance within groups}}$
N=total number of observations=40
G=Groups=3

Between variance = $\frac{n_1(\bar{y}_1 - \bar{y})^2 + \dots + n_G(\bar{y}_G - \bar{y})^2}{G-1}$
Within variance = $\frac{(n_1 - 1)s_1^2 + \dots + (n_G - 1)s_G^2}{N-G}$

$$df_1 = G - 1 \ df_2 = N - G$$

```
\#create\ a\ function\ to\ calculate\ F\ stat,\ p\ val\ based\ on\ formula
y \leftarrow c(50, 45, 55)
s \leftarrow c(10, 7, 4)
n \leftarrow c(20, 10, 10)
mu <- 50
anov <- function(y, mu, s, n) {</pre>
  bvec <- numeric()</pre>
  wvec <- numeric()</pre>
  for (i in 1:length(n)) {
    bvec[i] \leftarrow n[i] * (y[i]-mu)^2
    wvec[i] \leftarrow (n[i]-1) * s[i]^2
  df1 \leftarrow length(n) - 1
  df2 \leftarrow sum(n) - length(n)
  BV <- sum(bvec)/df1
  WV <- sum(wvec)/df2
  fstat <- BV/WV
  fcrit \leftarrow qf(0.95, df1, df2)
  pval <- 1 - pf(fstat, df1, df2)</pre>
  df <- data.frame(param = c("f_stat", "f_crit", "p_val"), values = c(fstat, fcrit, pval))</pre>
  return(df)
anov(y, mu, s, n)
      param
                   values
```

```
## param values
## 1 f_stat 3.72233400
## 2 f_crit 3.25192385
## 3 p_val 0.03365643
```

at 95% CI, f stat(3.7223) > f crit(2.352), so we reject null hypothesis

and accept that there is significant difference among 3 groups