final.Zunqiu.Wang

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```
set.seed(1)
library(caret)
## Loading required package: ggplot2
## Loading required package: lattice
library(psych)
##
## Attaching package: 'psych'
## The following objects are masked from 'package:ggplot2':
##
##
       %+%, alpha
library(glmnet)
## Loading required package: Matrix
## Loaded glmnet 4.1-2
library(e1071)
Q1
# assume i rolled 10^5 times
t <- 10<sup>5</sup>
# create sum of 5 sided dice funtion
count <- 0
dice <- function(num, side, times) {</pre>
  results <- replicate(times, sum(sample(1:side, num, replace = TRUE)))</pre>
  for (i in 1:t) {
  if (results[i] >= 15 & results[i] <= 20) {</pre>
    count = count + 1
  }
  return(count)
}
dice(5,6,t)
## [1] 55551
dice(5,6,t)/10^5
## [1] 0.5595
```

```
Q2
x \leftarrow rnorm(100)
ep <- rnorm(100)
y < -0.1 + 2 * x + ep
t.test(y, x)
##
     Welch Two Sample t-test
##
##
## data: y and x
## t = 0.28704, df = 135.71, p-value = 0.7745
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -0.4433402 0.5938877
## sample estimates:
##
      mean of x mean of y
## 0.03956547 -0.03570831
from t test, mean of y and mean of x dont differ significantly at 95% level
- calculate mean of y \bar{y} and x \bar{x}
y.bar \leftarrow mean(y[1:5])
y.bar
## [1] -0.1447467
x.bar \leftarrow mean(x[1:5])
x.bar
## [1] -0.01137821
   • calculate se_{diff} = \sqrt{\frac{s_y^2}{n_y} + \frac{s_x^2}{n_x}}
      s_y is standard deviation of y and s_x is standard deviation of x
      n_y is number of observations from y and n_x is number of observations from x
se.diff \leftarrow  sqrt(sd(y[1:5])^2/5 + sd(x[1:5])^2/5)
se.diff
## [1] 1.281371
   • calculate t value = \frac{\bar{x} - \bar{y}}{se_{diff}}
t.val <- (x.bar - y.bar)/se.diff
t.val
## [1] 0.1040826

    calculate

                                               df = \frac{(s_y^2/n_y + s_x^2/n_x)^2}{\frac{(s_y^2/n_y)^2}{n_y - 1} + \frac{(s_x^2/n_x)^2}{n_y - 1}}
sy < - sd(y[1:5])
sx < - sd(x[1:5])
```

```
df \leftarrow (sy^2/5 + sx^2/5)^2 / ((sy^2/5)^2/(5-1) + (sx^2/5)^2/(5-1))
## [1] 5.500271
qt(0.975,df)
## [1] 2.501826
Since t value=0.1040826 and t crit=2.501826, t value < t crit. Therefore, we fail to reject the null hypothesis
mean of y is equal to mean of x for the first 5 observations.
\mathbf{c}
y.bar
## [1] -0.1447467
sy
## [1] 2.621451
for (i in 6:100000) {
# Recalculate the standard error and CI
stand_err <- sy / sqrt(i)</pre>
ci \leftarrow y.bar + c(qt(0.01/2, i-1), qt(1-0.01/2, i-1))*stand_err
if (ci[2] < 0)
break # condition met, leave the for loop
}
i # minimum observations needed
## [1] 2181
ci
## [1] -2.894610e-01 -3.242557e-05
new_t <- (y.bar - 0) / stand_err</pre>
new_p_val \leftarrow pt(new_t, i-1)*2
new_p_val # confirm p=0.01 level
## [1] 0.009983347
i-5 # additional observations needed
## [1] 2176
Q3
y \leftarrow 0.1 + 0.2 * x + ep
lm.3 < - lm(y ~ x)
summary(lm.3)
##
## Call:
## lm(formula = y \sim x)
```

```
##
## Residuals:
##
       Min
                  1Q
                       Median
  -1.93732 -0.70038 -0.01064
                                        1.87553
##
                              0.61814
##
##
  Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
## (Intercept) 0.11465
                           0.09356
                                     1.225
                                           0.22337
## x
                0.30259
                           0.08924
                                     3.391 0.00101 **
##
## Signif. codes:
                   0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.935 on 98 degrees of freedom
## Multiple R-squared: 0.105, Adjusted R-squared: 0.09587
## F-statistic: 11.5 on 1 and 98 DF, p-value: 0.001006
qt(0.975, 98)
## [1] 1.984467
# 95% CI lower bound
lm.3$coefficients[2] - qt(0.975, 98) * coef(summary(lm.3))[4]
##
## 0.1254995
# 95% CI upper bound
lm.3$coefficients[2] + qt(0.975, 98) * coef(summary(lm.3))[4]
##
           Х
## 0.4796795
```

the coefficient on x is the slope of the line of best fit and implies that we are 95% confindence that an increase in x by 1.0 will result in an increase in y by the estimated slope. since there is uncertainty, the standard error in slope will help determine the 95% CI

b

```
# retrieve t value of coefficient of x
coef(summary(lm.3))[[6]]

## [1] 3.390812
pt(1.847, lower.tail = FALSE, 98) * 2
```

[1] 0.06776439

It confirms that it matches regression output and it indicates that at p=0.05 threhold level, we fail to reject the null hypothesis that coefficient is 0, which means change in x wont have effect on y

 \mathbf{c}

```
pf(3.41, 1, 98, lower.tail = F)
```

```
## [1] 0.06782021
```

F test is to measure overall significance of model as well by asking whether any coefficients are significantly different from 0. Since at p=0.05 level, we fail to reject the null hypothesis that none of coefficients are

d

```
x5 \leftarrow x[1:5]
y5 <- y[1:5]
x5.bar \leftarrow mean(x5)
y5.bar <- mean(y5)
# calculate slope
slope \leftarrow sum((x5-x5.bar)*(y5-y5.bar))/sum((x5-x5.bar)^2)
slope
## [1] 0.242416
intercept <- y5.bar - slope * x5.bar
intercept
## [1] -0.1215077
#calculate se
y5.pred <- intercept + slope * x5
error <- y5 - y5.pred
se.slope \leftarrow sqrt(sum(error^2)/(5-2)) * (1/sqrt(sum(x5-x5.bar)^2))
se.slope
## [1] 2.3648e+16
# calculate adjusted R^2
tss <- sum((y5-y5.bar)^2)
sse <- sum((y5-y5.pred)^2)</pre>
\#dft=n-1
dft <- 5-1
#df2=n-k-1
dfe < -5-1-1
adj.r2 \leftarrow (tss/dft-sse/dfe) / (tss/dft)
adj.r2
## [1] -0.2568931
Q4
y \leftarrow 0.1 + 0.2 * x - 0.5 * x^2 + ep
fit.r4 \leftarrow lm(y \sim x + I(x^2))
summary(fit.r4)
##
## Call:
## lm(formula = y \sim x + I(x^2))
##
## Residuals:
        Min
                   1Q Median
                                       ЗQ
                                                 Max
## -1.97988 -0.61633 -0.01342 0.61830 1.92586
```

```
##
## Coefficients:
##
               Estimate Std. Error t value Pr(>|t|)
                                     1.581 0.11709
## (Intercept) 0.18306
                           0.11577
## x
                0.29551
                            0.08951
                                      3.301 0.00135 **
               -0.56247
                            0.06227 -9.032 1.65e-14 ***
## I(x^2)
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.935 on 97 degrees of freedom
## Multiple R-squared: 0.502, Adjusted R-squared: 0.4918
## F-statistic: 48.9 on 2 and 97 DF, p-value: 2.061e-15
They are both significant
(0.1 + 0.2 * 2 - 0.5 * 2^2) - (0.1 + 0.2 * 1 - 0.5 * 1^2)
## [1] -1.3
\mathbf{c}
predict(fit.r4, data.frame(x=-0.7)) - predict(fit.r4, data.frame(x=-0.5))
## -0.1940939
Q_5
x2 \leftarrow rnorm(100, mean=-1, sd=1)
y \leftarrow 0.1 + 0.2 * x - 0.5 * x * x^2 + ep
x.mean <- mean(x)</pre>
(0.1 + 0.2 * x.mean - 0.5 * x.mean * 1) - (0.1 + 0.2 * x.mean - 0.5 * x.mean * 0)
## [1] 0.01785416
b
lm.complete \leftarrow lm(y \sim x + x2 + x:x2)
predict(lm.complete, data.frame(x=-0.7, x2=1)) - predict(lm.complete, data.frame(x=-0.5, x2=1))
##
## 0.07929373
lm.reduced <- lm(y~x)</pre>
r2.complete <- summary(lm.complete)$r.squared
r2.reduced <- summary(lm.reduced)$r.squared</pre>
\#df1=2, which is additional variables in complete model
\#n=100, k=3, so df2=100-3-1
fstat \leftarrow ((r2.complete - r2.reduced) / 2) / ((1 - r2.complete) / (100 - 3 - 1))
```

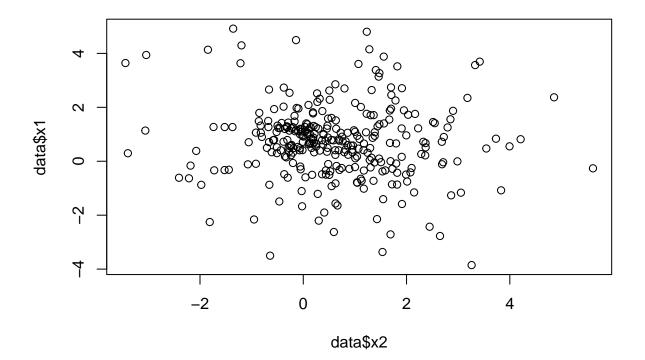
```
# p value
pf(fstat,2,(100-3-1),lower.tail=F)
```

[1] 1.073654e-08

the p value is way less than threshold p=0.05, so the F test reject the null hypothesis that complete regression model do no better than reduced regression model.

Q6 a

```
n <- 100
f <- c(rep("a", n), rep("b", n), rep("c", n))
x1 <- c(rnorm(n, 1, 2), rnorm(n, 0, 1), rnorm(n, 1, 0.5))
x2 <- c(rnorm(n, 1, 2), rnorm(n, 1, 1), rnorm(n, 0, 0.5))
data <- as.data.frame(cbind(x1, x2, f))
data$f <- as.factor(data$f)
plot(data$x2, data$x1)</pre>
```



```
kout <- kmeans(data[,1:2], 3, 1000)
kout$centers</pre>
```

```
## x1 x2
## 1 1.8660783 1.971075
## 2 -0.6663173 1.128106
## 3 1.0869306 -0.299753
```

```
data$cluster <- kout$cluster
table(data$f, data$cluster)</pre>
```

```
## ## 1 2 3 ## a 41 21 38 ## b 17 64 19 ## c 5 6 89
```

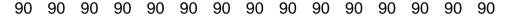
compare centroids from kmeans output with true means from normal distribution, I see they are not similar. And it is quite challenging to distinguish clusters of the plot from normal distribution.

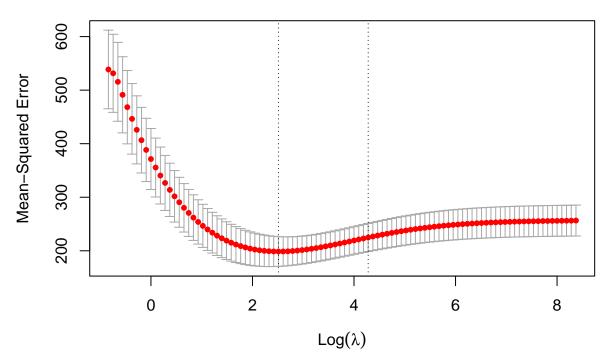
However, centroids from kmeans is pretty obvious: cluster c/3 has a positive on x1 around 1 and negative or around 0 on x2, cluster b/2 has a negative on x1 around -1 and positive around 1 on x2, cluster a/1 has both positive and around 2.

Regarding similarity between original cluster and cluster after kmeans, c cluster has very high accuracy, while b cluster is mostly classified as correct b or incorrectly a or b with each around 50%, a cluster is poorly classified and distributed all over 3 cluster with 30%.

Q7

```
X <- matrix(rnorm(200*90, 0, 1), nrow=200)</pre>
beta \leftarrow c(rep(2, 30), rep(-1, 30), rep(0, 30))
episilon \leftarrow rnorm(200, 0, 10)
y <- X %*% beta + episilon
#ridge, lasso, elastic net model with cv
ridge.fit <- cv.glmnet(X[1:100,], y[1:100], alpha=0)
lasso.fit <- cv.glmnet(X[1:100,], y[1:100], alpha=1)</pre>
elnet.fit <- cv.glmnet(X[1:100,], y[1:100], alpha=0.5)</pre>
ridge.fit
##
## Call: cv.glmnet(x = X[1:100, ], y = y[1:100], alpha = 0)
## Measure: Mean-Squared Error
##
##
       Lambda Index Measure
                                 SE Nonzero
## min 12.32
                  64
                       198.6 27.77
                                          90
## 1se 72.18
                       224.8 26.11
                  45
                                          90
plot(ridge.fit)
```





min(ridge.fit\$cvm)# lowest error for ridge

```
## [1] 198.5938
```

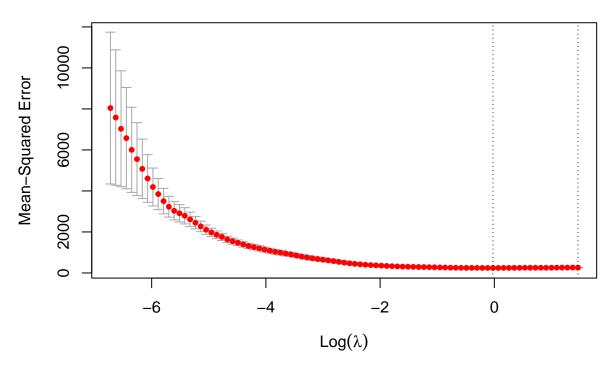
ridge.fit\$lambda.min

[1] 12.32324

lasso.fit # find min lmbda and its lowest MSE

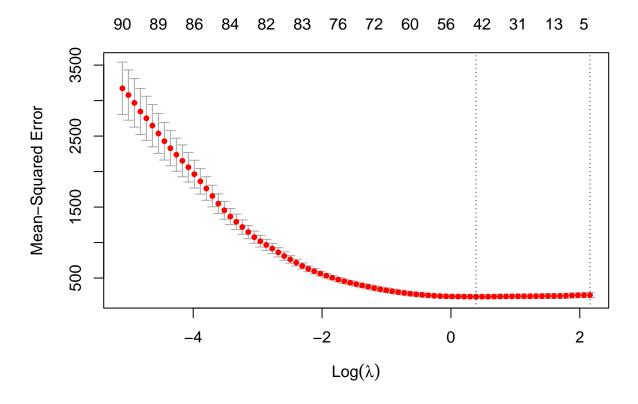
```
##
## Call: cv.glmnet(x = X[1:100, ], y = y[1:100], alpha = 1)
##
## Measure: Mean-Squared Error
##
## Lambda Index Measure SE Nonzero
## min 0.977 17 245.2 33.23 37
## 1se 4.327 1 260.1 14.39 0
plot(lasso.fit)
```

90 90 90 88 85 84 82 82 73 66 58 53 37 24 9



```
min(lasso.fit$cvm) # lowest error for lasso
```

```
## [1] 245.251
elnet.fit # find min lmbda and its lowest MSE
##
## Call: cv.glmnet(x = X[1:100, ], y = y[1:100], alpha = 0.5)
## Measure: Mean-Squared Error
##
##
       Lambda Index Measure
                               SE Nonzero
## min 1.478
                 20
                      237.0 27.04
## 1se 8.654
                      258.6 33.38
                  1
plot(elnet.fit)
```



```
min(elnet.fit$cvm) # lowest error for elastic net
```

[1] 237.0356

Ridge regression with alpha=0, lambda=12.32 produces lowest MSE so i choose this model for testing.

b

```
fit.coef <- coef(ridge.fit)[, "s1"]
fit.30.acc <- sum(round(fit.coef[2:31]) == 2)
fit.60.acc <- sum(round(fit.coef[32:61]) == -1)
fit.90.acc <- sum(round(fit.coef[62:91]) == 0)
sum(fit.30.acc, fit.60.acc, fit.90.acc)</pre>
```

[1] 31

Overall, after rounding to leave only the digit to compare with the simulated data, about 31/90 is approximately matched.

 \mathbf{c}

```
beta.fit <- coef(ridge.fit)[, "s1"]
X200 <- X[101:200, ]
y200 <- y[101:200]
#predicted y
yhat.200 <- cbind(1,X200) %*% beta.fit</pre>
```

```
# mse
mean((y200 - yhat.200)^2)
## [1] 171.4463
d
# multiple regression
lm.100 \leftarrow lm(y[1:100] \sim X[1:100,])
# beta
lm.beta.fit <- coef(lm.100)</pre>
# predicted
yhat.lm <- cbind(1,X200) %*% lm.beta.fit</pre>
# mse
mean((y200 - yhat.lm)^2)
## [1] 759.6831
The MSE from regular regression is about 5 times than ridge regression. One possible reason is that regular
regression is overfitting in sample data and predicted poorly on out of sample data. The ridge regression
will drop any variable due to having a high number of independent variables.
Q8
y2 <- cut(y,breaks=c(-Inf,0,Inf),labels=c("0","1"))</pre>
\mathbf{a}
df \leftarrow data.frame(X 100=X[1:100,], y2 100=y2[1:100])
costvalues \leftarrow 10^{\circ}seq(-3,2,1)
# linear kernel
svm.l <- tune(svm, y2_100 ~ ., data=df, ranges=list(cost=costvalues), kernel="linear")</pre>
\# best cost = 0.01
svm.l$best.model
##
## Call:
## best.tune(method = svm, train.x = y2_100 ~ ., data = df, ranges = list(cost = costvalues),
       kernel = "linear")
##
##
##
## Parameters:
##
      SVM-Type: C-classification
##
    SVM-Kernel: linear
##
           cost: 0.1
##
## Number of Support Vectors: 67
# linear kernel in smaple accuracy
yhat.in.l <- predict(svm.l$best.model, df)</pre>
sum(yhat.in.l == df$y2_100)/length(df$y2_100) # 0.99 accuracy
```

[1] 0.99

```
# radial kernel
svm.r <- tune(svm, y2_100 ~ ., data=df, ranges=list(cost=costvalues), kernel="radial")</pre>
# best cost = 10
svm.r$best.model
##
## Call:
## best.tune(method = svm, train.x = y2_100 ~ ., data = df, ranges = list(cost = costvalues),
##
       kernel = "radial")
##
##
## Parameters:
      SVM-Type: C-classification
##
## SVM-Kernel: radial
          cost: 10
##
##
## Number of Support Vectors: 100
# linear kernel in smaple accuracy
yhat.in.r <- predict(svm.r$best.model, df)</pre>
sum(yhat.in.r == df$y2_100)/length(df$y2_100) # 1 accuracy
## [1] 1
b
y2_200 <- y2[101:200]
# choose radial kernel predict with test, accuracy since it has higher in sample accuracy
yhat.out.r <- predict(svm.r$best.model, X[101:200,])</pre>
table(yhat.out.r, y2_200)
             y2_200
## yhat.out.r 0 1
##
            0 31 11
            1 18 40
##
sum(yhat.out.r == y2_200)/length(y2_200) #0.68 accuracy
## [1] 0.71
```