

Lecture 6 Fisher Information

IEMS 402 Statistical Learning

Northwestern

Asymptotic Normality

Asymptotic Theory for ERM?

what is the asymptotic distribution of $\hat{\theta}_n := \arg \min \mathbb{E}_{P_n} l_\theta(x)$

For example: maximum likelihood $l_\theta(x) := \log P_\theta(x)$

Optimality Condition:

$$\mathbb{E}_{P_n} \nabla l_{\hat{\theta}_n}(x) = 0$$

$$\mathbb{E}_{P_n} \nabla l_{\theta^*}(x) + \nabla^2 l_{\theta^*}(x)(\hat{\theta}_n - \theta) + \text{h.o.t.}$$

$$\mathbb{E}_{P_n} \nabla l_{\theta^*}(x) - \mathbb{E}_{P_n} \nabla l_{\theta^*}(x) \xrightarrow{\text{D}, \text{ optimality condition}} \mathbb{E}_{P_n} \nabla l_{\theta^*}(x)$$

"", optimality condition

θ^* is the optimal one

$$\text{Today's AIM: } \sqrt{n}(\hat{\theta}_n - \theta^*) \xrightarrow{} N(0, e'(\theta^*)^{-1} \mathbb{E}_{P_\theta^*} (\nabla l \nabla l^\top) (\theta^*)^{-1}) \text{ where } e(\theta) = \mathbb{E}_{P_\theta^*} \nabla^2 l_\theta$$

inverse the Hessian

Asymptotic theory

Theorem

Let $X_i \stackrel{\text{iid}}{\sim} P_{\theta_0}$ and assume $\hat{\theta}_n = \operatorname{argmax}_{\theta} P_n \ell_{\theta}(X)$ is consistent.

Define the covariance

$$\Sigma_{\theta} := (P_{\theta} \nabla^2 \ell_{\theta}(X))^{-1} \operatorname{Cov}_{\theta}(\nabla \ell_{\theta}(X)) (P_{\theta} \nabla^2 \ell_{\theta}(X))^{-1}$$

Under the previous assumptions,

$$\sqrt{n}(\hat{\theta}_n - \theta_0) \xrightarrow{d} \mathcal{N}(0, \Sigma_{\theta_0})$$

- ▶ “typically” $\Sigma_{\theta} = -(P_{\theta} \nabla^2 \ell_{\theta}(X))^{-1} = \operatorname{Cov}_{\theta}(\dot{\ell}_{\theta})$

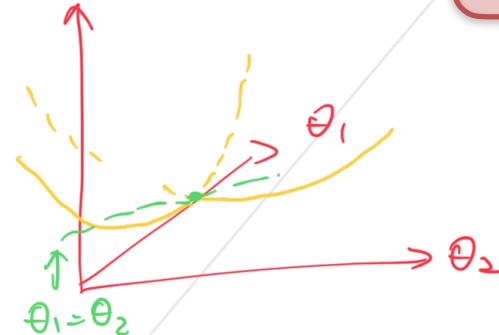
Proof

Bias-variance trade-off in Asymptotic?

- $\theta_1 = \theta_2$
- $Y = \theta_1 X + \theta_2 X$ learn θ_1, θ_2
- $Y = \theta_1 X + \theta_1 X$ learn θ_1



Var = Var of Estimation
Project to $\theta_1 = \theta_2$.



Not Required

Duchi J, Ruan F. Asymptotic optimality in stochastic optimization. arXiv preprint
arXiv:1612.05612, 2016.

Moment Estimator

if we know $e(\theta) = \mathbb{E}_{X \sim P_\theta}[F(X)]$, we define $e(\hat{\theta}_n) = \mathbb{E}_{\mathbb{P}_n} f(X)$

Inverse Function Theorem

$$(F^{-1})'(t) = \frac{\partial}{\partial t} F^{-1}(t) = (F'(F^{-1}(t)))^{-1}.$$

Hints for future research

$f(\theta) = \arg \min_f F_\theta(f)$, What is $f'(\theta)$?

Not Required

Exponential Family

Definition 3.1. $\{P_\theta\}_{\theta \in \Theta}$ is a regular exponential family if there is a sufficient statistic $T : \mathcal{X} \rightarrow \mathbb{R}^d$ such that P_θ has density

$$P_\theta = \exp(\theta^T T(x) - A(\theta))$$

with respect to μ , where $A(\theta) = \log \int e^{\theta^T T(x)} d\mu(x)$.

Fact: Moment estimator for exp family using moment T equals to ERM estimator

Fisher Information

Asymptotic Theory for

Max like-lihood

what is the asymptotic distribution of $\hat{\theta}_n := \arg \min \mathbb{E}_{P_n} l_\theta(x)$

For example: maximum likelihood $l_\theta(x) := \log P_\theta(x)$

Thm. $\sqrt{n}(\hat{\theta}_n - \theta^*) \rightarrow N(0, I_{\theta^*}^{-1})$ I_{θ^*} is the Fisher Information

Claim. $\mathbb{E}_{P_\theta} [\sigma^2 l_\theta] = -\mathbb{E} [\nabla l_\theta \nabla l_\theta^\top]$ $\nabla l_\theta = \frac{\nabla P}{P}$

$$\frac{\nabla^2 P_\theta}{P_\theta} - \frac{\mathbb{E} P \nabla P^\top}{P_\theta^2}$$

Today's AIM: $\sqrt{n}(\hat{\theta}_n - \theta^*) \rightarrow N(0, e'(\theta^*)^{-1} e' \mathbb{E}_{P_\theta^*} (\nabla l \nabla l^\top) \theta^*)^{-1}$ where $e(\theta) = \mathbb{E}_{P_\theta^*} \nabla^2 l_\theta$

$$\mathbb{E} \frac{\nabla^2 P_\theta}{P_\theta} = \int \nabla^2 P_\theta = \int \nabla P \nabla P^\top = 0$$

Fisher Information

Definition (Fisher information)

For a model family $\{P_\theta\}$ on \mathcal{X} , the *Fisher information* is

$$I(\theta) := \mathbb{E}_\theta[\nabla \ell_\theta(X) \nabla \ell_\theta(X)^\top]$$

- when \mathbb{E} and ∇ are interchangeable, then $I(\theta) = -\mathbb{E}[\nabla^2 \ell_\theta(X)]$

Claim. $\mathbb{E} \nabla \ell(\theta) = 0$ ① optimality condition of max-likelihood

$$\mathbb{E}_P[\nabla \log p] = \mathbb{E}_P\left[\frac{\partial P}{P}\right] = \int \partial P = \int P \cdot \partial 1 = 0$$

$\frac{1}{x} \ln(x) = \frac{1}{x}$ is a key structure!!

Score function!

$$\nabla \ell_\theta(x) = \left[\frac{\partial}{\partial \theta_j} \log p_\theta(x) \right]_{j=1}^d \in \mathbb{R}^d$$
$$\nabla^2 \ell_\theta(x) = \left[\frac{\partial^2}{\partial \theta_i \partial \theta_j} \log p_\theta(x) \right]_{i,j=1}^d \in \mathbb{R}^{d \times d},$$

Cramér–Rao lower bound

Thm. If $\mathbb{E}[\hat{\theta}_n] = \theta^*$ \Leftrightarrow unbiaj.

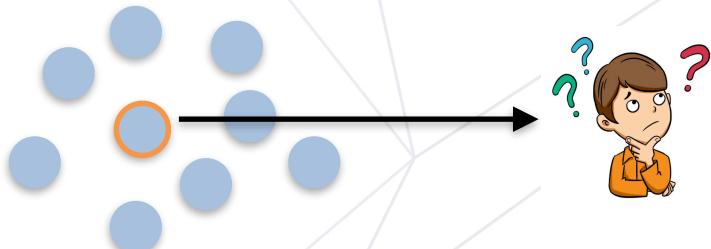
$$\text{Var}(\hat{\theta}_n) \geq I_{\theta^*}^{-1}$$

Pf. $\mathbb{E}[\hat{\theta} \circ \log P(\omega)] = 1$ then use $\text{Cov}(X, Y) \geq \text{Var}(X)\text{Var}(Y)$

$$\mathbb{E} \int \hat{\theta} \frac{dP}{P} = \mathbb{E} \hat{\theta} P = \mathbb{E} \theta^* = I.$$

Influence Function

influence function



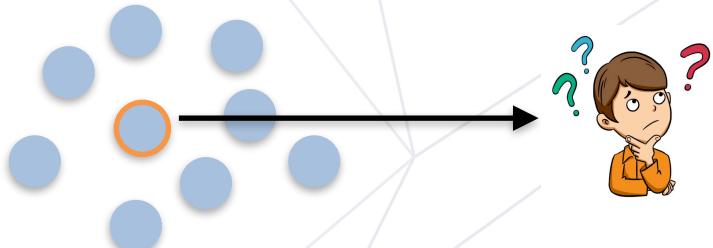
What is the influence that
we delete the data?

$$\hat{\theta} = \arg \min_{\theta} \frac{1}{n} \left(\sum_{i=1}^n l(x_i, y_i; \theta) + l(x_n, y_n; \theta) \right)$$



$$\hat{\theta}_- = \arg \min_{\theta} \frac{1}{n} \sum_{i=1}^n l(x_i, y_i; \theta)$$

influence function



What is the influence that we delete the data?

$$\hat{\theta} = \arg \min_{\theta} \frac{1}{n} \left(\sum_{i=1}^n l(x_i, y_i; \theta) + l(x_n, y_n; \theta) \right)$$



$$\hat{\theta}_- = \arg \min_{\theta} \frac{1}{n} \sum_{i=1}^n l(x_i, y_i; \theta)$$

$$\hat{\theta}_{\epsilon} = \arg \min_{\theta} \frac{1}{n} \left(\sum_{i=1}^n l(x_i, y_i; \theta) + \epsilon l(x_n, y_n; \theta) \right)$$



Influence function: $\frac{d\hat{\theta}_{\epsilon}}{d\epsilon}$



How to compute that?

Influence function

$$\hat{\theta}_{\epsilon} = \arg \min_{\theta} \frac{1}{n} \left(\sum_{i=1}^n l(x_i, y_i; \theta) + \epsilon l(x_n, y_n; \theta) \right)$$

AIM: $\frac{d\hat{\theta}_{\epsilon}}{d\epsilon}$

$$\left(\sum_{i=1}^n \nabla_{\theta} l(x_i, y_i; \hat{\theta}_{\epsilon}) + \epsilon \nabla_{\theta} l(x_n, y_n; \hat{\theta}_{\epsilon}) \right) = 0$$

Influence function

$$\hat{\theta}_{\epsilon} = \arg \min_{\theta} \frac{1}{n} \left(\sum_{i=1}^n l(x_i, y_i; \theta) + \epsilon l(x_n, y_n; \theta) \right)$$

$$\left(\sum_{i=1}^n \nabla_{\theta} l(x_i, y_i; \hat{\theta}_{\epsilon}) + \epsilon \nabla_{\theta} l(x_n, y_n; \hat{\theta}_{\epsilon}) \right) = 0$$

Take gradient respect to ϵ

$$\sum_{i=1}^n H_{\theta} l(x_i, y_i; \hat{\theta}_{\epsilon}) \frac{d\hat{\theta}_{\epsilon}}{d\epsilon} + \epsilon H_{\theta} l(x_n, y_n; \hat{\theta}_{\epsilon}) \frac{d\hat{\theta}_{\epsilon}}{d\epsilon} + \nabla_{\theta} l(x_n, y_n; \hat{\theta}_{\epsilon}) = 0$$

$$\mathcal{I}_{\text{up, params}}(z) \stackrel{\text{def}}{=} \frac{d\hat{\theta}_{\epsilon, z}}{d\epsilon} \Big|_{\epsilon=0} = -[H_{\hat{\theta}}^{-1}] \nabla_{\theta} L(z, \hat{\theta}),$$

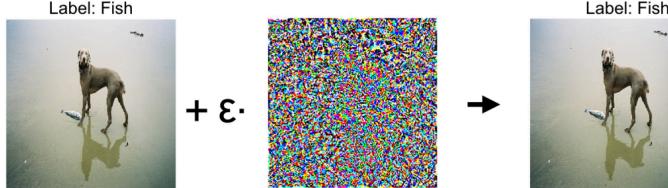
Hessian of all data
Gradient of the data of interest



How to compute this?

Applications

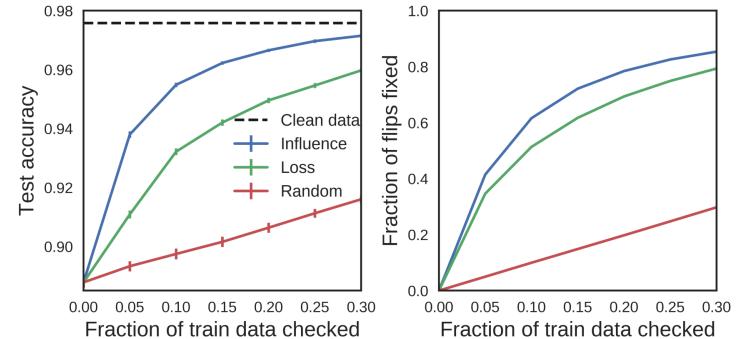
A small perturbation to one training example:



Can change multiple test predictions:



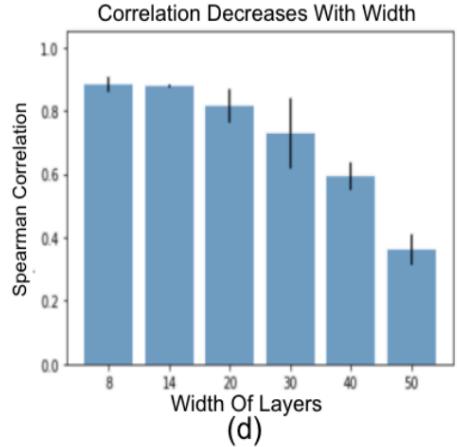
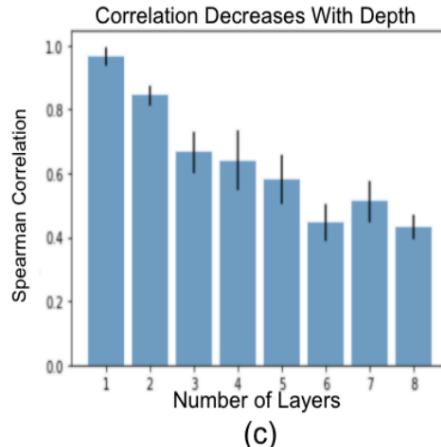
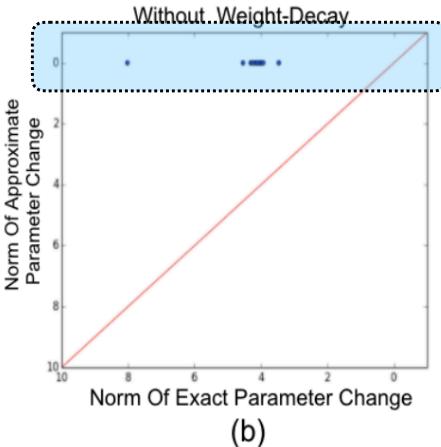
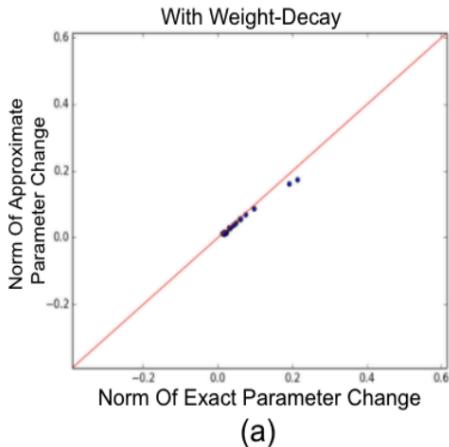
Figure 5. Training-set attacks. We targeted a set of 30 test images featuring the first author's dog in a variety of poses and backgrounds. By maximizing the average loss over these 30 images, we found a visually-imperceptible change to the particular training image (shown on top) that flipped 1 test image



Checking mislabeled data

<https://arxiv.org/pdf/1703.04730>

However



<https://arxiv.org/pdf/2006.14651>

Overparameterize: SVM example

Conjecture: $\text{IF} = 0$ for overparameterized margin

Open!

Margin (gap between decision boundary and hyperplanes)

