

# Diverse Genotype-by-Weather Interactions in Switchgrass

Alice H. MacQueen<sup>a,1,2</sup>, Li Zhang<sup>a,1</sup>, Samuel A. Smith<sup>a,1</sup>, Jason Bonnette<sup>a</sup>, Arvid R. Boe<sup>b</sup>, Phillip A. Fay<sup>c</sup>, Felix B. Fritsch<sup>d</sup>, David B. Lowry<sup>e</sup>, Robert B. Mitchell<sup>f</sup>, Francis M. Rouquette Jr<sup>g</sup>, Yanqi Wu<sup>h</sup>, Arbel Harpak<sup>a</sup>, and Thomas E. Juenger<sup>a,2</sup>

The timing of vegetative and reproductive growth in plants (“phenological timings”) depends on genetic effects, environmental (e.g., weather) cues, and possibly their interaction. Quantifying these GxWeather interactions can aid prediction and manipulation of phenological timings. Here, we map GxWeather effects on phenological timings in two highly divergent switchgrass (*Panicum virgatum*) populations using repeated plantings of cloned individuals from these populations at eight sites across the central United States. Some genetic effects covary with interpretable weather-based cues. For example, in the Gulf population, 65% of genetic effects on the timing of vegetative growth covary with daylength 14 days prior to green-up date, and 33% of genetic effects on the timing of flowering covary with cumulative rainfall in the seven days prior to flowering. However, most variation in genetic effects is site-specific and can not be attributed to variation in the studied weather variables. We demonstrate that we can identify genetic variation with GxWeather and assign these loci to specific weather-based cues or other patterns. Breeding for particular alleles at the GxWeather loci we identified could change flowering responsiveness in a photoperiod or rainfall-specific way. More broadly, our approach proposes a refined, powerful characterization of genotype-by-environment interactions in any species phenotyped in multiple environments.

allele-by-environment effect variation | rank-changing genotype-by-environment interaction | photoperiod | cumulative rainfall | genetic variation

Plant phenological timings are major components of plant fitness affected by multiple external environmental cues (e.g. degree of winter chilling, day length, temperature, and water availability) that signal existing or upcoming growing conditions (1–4). Genetic responses to environmental cues determine the speed, timing, and energy apportioned to vegetative and reproductive growth and shape both the individual’s lifespan and its lifetime production of viable seed. Day length (or photoperiod) is one of the most predictable environmental cues, and genetic sensitivity to photoperiod protects plants from potentially fatal consequences of phenological responses to temperature cues at the “wrong” time of year. However, the usefulness of specific environmental cues depends on both features of the environment, such as cue predictability and relevance, and the species’ adaptive strategies (5). Species with wide natural distributions can have multiple distinct environmentally cued phenological responses: for example, populations of sunflower (*Helianthus annuus*) exhibit day-neutral, facultative short day, and facultative long-day flowering responses, which vary with their environments (6, 7). Distinct genetic responses in different environments are known as genotype-by-environment interactions, or GxE. Flowering time, or the transition from vegetative to reproductive growth, is a common subject of GxE research (6–13), a key output of selection driving adaptation to local environments (2, 14–16), and a selection target for crop improvement to adapt crops to local or future environments (17). Changing flowering responsiveness to photoperiod cues has allowed geographic range expansion and increased yields in several cereal species (18–22) and other crops (23, 24). Recent statistical advances in studying phenological GxE have involved determining critical environmental indices before the phenological event occurs, such as photothermal time within a critical growth window (11). However, most studies of flowering GxE focus on finding a single, best fitting form of genotype-environment covariance, despite the key expectation that different genetic subpopulations, and even different genomic regions, have likely evolved distinct patterns of GxE (25). Additionally, despite theoretical predictions that local adaptation should involve antagonistic pleiotropy, or alleles with effects with opposing fitness outcomes (26–29), previous empirical work has found limited evidence of antagonistic pleiotropy (14, 30, 31). However, this work has been limited by a known statistical bias that reduced detection of

## Significance Statement

The timing of plant seasonal development (phenology) has major impacts on fitness because of the steep fitness cost of plant-environment mismatches. We infer the mixture of ways that genetic effects on phenological traits depend on plant environment, focusing on how effects covary with weather prior to the phenological event (GxWeather). Most effects do not have GxWeather, but a minority do. GxWeather is population-specific. The majority of effects on the timing of vegetative growth in the Gulf population have GxWeather: these effects differ in sign across environments and covary with a photoperiod cue two weeks prior. A minority of effects on flowering date in the Gulf and Midwest populations had GxWeather, covarying with a cumulative rainfall and a photoperiod cue, respectively.

Author affiliations: <sup>a</sup>University of Texas at Austin, Department of Integrative Biology, Austin, 78712; <sup>b</sup>South Dakota State University, Department of Agronomy, Brookings, 57006; <sup>c</sup>USDA-ARS, Grassland, Soil and Water Research Laboratory, Temple, 76702; <sup>d</sup>University of Missouri, Division of Plant Sciences, Columbia, 65211; <sup>e</sup>Michigan State University, Department of Plant Biology, East Lansing, 48824; <sup>f</sup>USDA-ARS, Wheat, Sorghum, and Forage Research Unit, Lincoln, 68583; <sup>g</sup>Texas A&M University, Texas A&M AgriLife Research and Extension Center, Overton, 75684; <sup>h</sup>Oklahoma State University, Department of Plant and Soil Sciences, Stillwater, 74078

T.E.J. designed research. D.B.L. contributed plant material and resources. J.B., D.B.L., and T.E.J. designed and executed field experiments. A.R.B., P.A.F., F.B.F., D.B.L., R.B.M., F.M.R., Y.W., and T.E.J. hosted field experiments. A.H.M., L.Z., and S.A.S. conducted statistical and computational analyses. The manuscript was written by A.H.M. with contributions from all authors.

The authors declare no conflicts of interest.

<sup>1</sup> A.H.M. contributed equally to this work with L.Z. and S.A.S.

<sup>2</sup>To whom correspondence should be addressed. E-mail: tjuenger@utexas.edu

genetic effects that differ in sign (30, 32, 33). Thus, despite substantial interest in the frequencies of various forms of GxE, the frequency of sign-changing GxE relative to other forms of GxE remains unknown.

Previous research suggests that switchgrass phenological timings should have GxWeather and that these timings could differ by genetic subpopulation. Switchgrass is considered a short-day plant with reproductive development strongly linked to day of the year (34). However, as part of its wide environmental adaptation across the eastern half of North America, its photoperiodicity has been predicted to differ by plant latitude of origin (35, 36). We previously found divergent Midwest and Gulf genetic subpopulations of switchgrass with distinct sets of environmental adaptations, in that both populations had distinct genetic variation associated with each of two fitness proxies, biomass and overwinter survival (37). The Midwest genetic subpopulation is primarily composed of individuals from the well-studied upland switchgrass ecotype (38, 39), while the Gulf subpopulation has individuals from the well-studied lowland ecotype and the phenotypically intermediate coastal ecotype (37).

Here, we test for GxWeather for two phenological timings by assigning patterns of genetic effects on phenology at eight common gardens to many patterns of weather covariance at these gardens. To do this, we phenotype a diversity panel of hundreds of switchgrass genotypes from the Midwest and Gulf subpopulations for the timing of vegetative and reproductive development. We do this at eight common garden locations spanning 17 degrees of latitude: these gardens cover the majority of the latitudinal and climatic range of switchgrass and capture the most comprehensive picture to date of the environmental variation this species encounters. We define multiple ways phenological timings might covary with weather (Table 1) and additional ways phenological timings might vary by garden (SI Appendix, Section S1), then jointly re-estimate genetic effects on these timings at all eight common gardens using the set of these covariance matrices that significantly improved the modeled log-likelihood when included (SI Appendix, Section S2-S5) (40). We use the Bayesian framework *mash* (multivariate adaptive shrinkage) developed by (40), to refine effect size estimates from genome-wide association (GWAS). *Mash* allows us to identify and specify multiple covariance structures among genetic effect estimates across sites, including structures that represent covariance in weather variables of interest. Importantly, this method circumvents statistical biases in detecting genetic effects with the same or opposite signs (41). To confirm our genetic mapping of GxWeather, we compare genomic locations of the significant posterior effect estimates from *mash* to mapping results from an outbred mapping population grown at the same sites. Our analyses allow us to describe the weather cues and types of GxE affecting phenology in two divergent natural populations of switchgrass across the species' latitudinal range.

## Results

Genotypes from the Gulf and Midwest subpopulations had distinct phenological timings and distinct patterns of phenological correlations across our eight common garden sites (Figure 1). At the three Texas common gardens (hereafter 'Texas' gardens), located within the natural range of the

Gulf subpopulation, the onset of vegetative growth, or 'green-up date', for Gulf genotypes occurred before Midwestern green-up, and the onset of reproductive growth, or 'flowering date', for Gulf genotypes occurred after Midwestern flowering (Figure 1 A). At the four northernmost common gardens (hereafter 'North' gardens), located within the natural range of the Midwest subpopulation, both green-up and flowering of Gulf genotypes occurred after that of Midwest genotypes. At the Oklahoma common garden, located near the natural range limits of both the Gulf and the Midwest subpopulations, Gulf and Midwest green-up occurred over the same period, and Gulf genotype flowering occurred after Midwestern flowering (Figure 1 A). These patterns led to strong negative phenotypic correlations for the onset of vegetative growth between the North and Texas gardens, particularly in the Gulf and across all individuals from the Gulf and Midwest subpopulations (hereafter, 'Both' subpopulations), and contributed to positive phenotypic correlations for the onset of flowering that had larger magnitudes at more northern gardens (Figure 1 B). Narrow-sense heritabilities ( $h^2$ ) suggested that rank-changing GxE for these phenotypes was present across the common gardens (SI Appendix, Figure S1).  $h^2$  were typically high at individual gardens: 59% on average for green-up date, and 87% for flowering date. However,  $h^2$  were variable across gardens, and green-up dates were uncorrelated ( $r^2 < 0.2$ ) or negatively correlated between pairs of gardens (Figure 1 B). These negative and small correlations undoubtedly contributed to the low  $h^2$  values for green-up and flowering date when estimated jointly at all eight gardens:  $h^2$  was 0.8% for green-up and 23.2% for flowering date (SI Appendix, Figure S1).

**Inference of genome-wide patterns of GxE and GxWeather for vegetative and reproductive timing.** We expected patterns of GxE and GxWeather to vary at the locus, subpopulation, and trait level; thus we were interested in modeling the different ways effects might covary across our eight common gardens. We jointly re-estimated genetic effects of a subset of SNPs across all eight common gardens using *mash* models, which can flexibly capture many modes of covariance across gardens (SI Appendix, Sections S1-S5, Datasets 1-6). To obtain genetic effects, we used subsets of effects from garden- and subpopulation GWAS (SI Appendix, Section S2-S3; Table S1). To obtain covariance structures, we specified both covariance structures introduced in the initial *mash* manuscript, and GxWeather covariance matrices estimated from the covariance of empirical weather patterns at each garden at specific times before the phenological event (Table 1; SI Appendix, Section S1). Then, we used a greedy algorithm to select covariance matrices from each model's set that significantly improved the model likelihood (SI Appendix, Section S4; Table S2). Finally, we fit *mash* models using the selected set of covariance matrices on a subset relatively unlinked ( $r^2 < 0.2$ ), 'strong' genetic effects with low  $p$ -values (SI Appendix, Section S5). The loadings of genetic effects onto the GxE and GxWeather covariance matrices included in each model provide information on genome-wide patterns of SNP-environment interaction. For example, the phenotypic correlations for the onset of vegetative growth had moderate negative correlations between the Texas and North gardens, particularly in the Gulf subpopulation and Both subpopulations (Figure 1 B). If these phenotypic correlations have a genetic component, they

could be partially or completely captured by the covariance structures specified in the *mash* model. In this case, the joint estimate of effects for many SNPs could have high mixture proportions, or mass, on a covariance matrix that captures the negative correlation between Texas and North gardens. In our data, daylength 14 days prior to green-up date has negative correlations between Texas and North gardens (Figure 2 A); if many SNP effects have mass on this GxWeather matrix, we would infer that the effects of those SNPs have that form of GxWeather covariation.

Eight GxWeather covariance structures were selected by the greedy algorithm, one to two per subpopulation: five for green-up date, and three for flowering date (SI Appendix, Table S2). Of these eight, five had mass on them in *mash* models of the strong effects (Figure 2 B,D). Three of these five matrices had negative covariances between sites between Texas and North gardens (Figure 2 A), while two had all positive or near-zero covariances.

For green-up date, SNP-associated phenotypic effects covaried with different weather-based cues in the Gulf & in Both subpopulations (Figure 2 B). In total, 65% of the posterior weight of strong SNP effects in the *mash* model of Gulf green-up fell on a covariance matrix constructed using the covariance of day length 14 days prior to the date of green-up. The covariance matrix for this weather cue was visually similar to the observed pattern of phenotypic correlation for the timing of vegetative growth in the Gulf subpopulation (Figure 1 B; Figure 2 A). *Mash* models of the timing of vegetative growth in the Midwest subpopulation and Both subpopulations did not include this GxWeather covariance; the Midwest had no weight on any GxWeather covariance, while Both subpopulations had non-zero weights on two additional GxWeather covariance types, average temperature one day prior to green-up, and the day length change in seconds in the day prior to green-up (Figure 2 B). The average temperature covariance matrix had negative covariances between Texas and North gardens, though not as strong as the negative phenotypic correlations seen in Both subpopulations (Figure 1 B; Figure 2 A). Only the Gulf subpopulation and Both subpopulations had mass on any GxWeather covariance matrices (Figure 2 C) for vegetative growth.

For flowering date, distinct GxWeather covariance structures captured covariance in effect sizes for SNPs in the Gulf and Midwest subpopulations. 33% of SNP effects on flowering in the Gulf subpopulation covaried with cumulative rainfall in the seven days prior to flowering (Figure 2 D). 22.6% of SNP effects on flowering in the Midwest subpopulation covaried with day length change in the two days prior to flowering (Figure 2 D), which had negative covariances between Texas and North gardens. Neither of these covariance matrices significantly improved the log-likelihood of the *mash* model of Both subpopulations, and no GxWeather covariance matrices had non-zero mass in models of Both subpopulations (Figure 2 E). In five of the six *mash* models of strong effects, the GxWeather covariance matrices captured a minority of the posterior weights of the strong effects (Figure 2 C,E); in these five models, the majority of this mass was on various canonical covariance matrices. These matrices included simple heterozygosity, with intermediate, positive covariances between all gardens, and single effect matrices with garden-specific effects.

**Frequency of rank-changing GxE in significant SNP effects.** We expected to observe common rank-changing GxE at the level of individual loci as this is a key theoretical prediction of local adaptation (26–29). To determine the frequency of rank-changing GxE, we used the local false sign rate (lfsr), an analogue of the lfr that establishes confidence in the effect sign, not the effect's difference from zero, to determine significance. We required lfsr significance ( $p < 0.05$ ) in both gardens to include effects. This means that our tests for a sign change between gardens carry an equal statistical burden to those for effects with the same sign. We separated kinds of effects at the level of individual loci into SNP effects that differ in sign between gardens (effects with rank-changing GxE between gardens), SNP effects that differ in magnitude (effects that are large in one garden or region and smaller in others), and SNP effects that are indistinguishable in two gardens (Figure 3 A).

For green-up date for the Gulf subpopulation, hundreds to thousands of pairwise effects exhibited rank-changing GxE, or a difference in sign between pairs of Texas and North gardens (Figure 3 B; SI Appendix, Fig S2). 78.7% of pairwise comparisons between North and Texas gardens had a difference in sign, while only 28.6% and 0.2% of North-North or Texas-Texas comparisons had a difference in sign, respectively (Figure 3 B). The majority of pairwise effects for greenup for the Midwest (>55%) and Both (>85%) subpopulations were the same sign, and effects most often differed in magnitude between gardens within and between the regions (Figure 3 B; SI Appendix, Fig S2).

For flowering date for the Gulf subpopulation, less than 2% of pairwise effects exhibited rank-changing GxE, or a difference in effect sign within or between regions (Figure 3 B). More effects differed in magnitude between the Texas and North regions than within these regions (42.7% vs <20%; Figure 3 B), while the majority of effects had no GxE. The Midwest population had relatively few significant effects for flowering, but a large proportion of these differed in sign between Texas and North regions (42.7%) or within the North region (65.4%). Finally, in Both subpopulations, less than 20% of pairwise effects differed in sign (Figure 3 B). Most differences in sign were between TX1, the southernmost garden, and all other gardens (SI Appendix, Figure S3). Similarly, more effects that differed in magnitude included gardens in the Texas region (52.3–55.9%), and most effect pairs in the North region were not distinguishable (91.5%).

#### **Confirmation of effects on phenology using an independent mapping population.**

We sought additional experimental support the occurrence and location of our significant SNP effect re-estimates using an independent pseudo-F2 mapping population created from Gulf & Midwest individuals and grown at the same sites (Figure 4 A,B). We conducted quantitative trait loci (QTL) mapping of flowering as functions of five environmental cues that we also used as covariance matrices in *mash*, and identified eight QTL for flowering date, eight QTL for flowering as a function of day length change two days prior, one QTL for the start of vegetative growth, and two QTL for vegetative growth as a function of daylength change one day prior, all of which showed QTL by environment interactions (SI Appendix, Figure S4). All QTL for flowering overlapped one or more homologs from rice or *A. thaliana* with functionally validated roles in flowering (SI Appendix,



Dataset 7); the QTL on Chr04K overlapped a gene recently functionally validated for flowering in switchgrass (42). All flowering and green-up QTL intervals contained at least one SNP significant in at least one *mash* run at a log10-transformed Bayes Factor > 2, or in the 1% tail of significance, whichever was stricter (SI Appendix, Dataset 8). We also looked for enrichments of *mash* SNPs in the 1% tail of significance (the ‘*mash* 1% tail’) within each QTL interval. At the 5% level, three QTL had enrichments of SNPs in the *mash* 1% tail. Overall, there were five significant enrichments ( $p < 0.05$ , hypergeometric test) of SNPs in the *mash* 1% tail in the QTL intervals. Thus, we were able to experimentally support the genomic windows of some re-estimates of significant SNP effects with a QTL mapping experiment using a separate mapping population.

## Discussion

As the climate and the natural environment change, it is increasingly critical to understand how patterns of plant-environment interactions will change in response. To do this, we must understand the current patterns of trait covariation across environments, the genetic underpinnings of these patterns, and the cases where this covariation can be altered. Here, we demonstrate that we can associate multiple patterns of GxWeather with specific genomic regions using a switchgrass diversity panel grown at eight common gardens. We can assign genetic effects to both GxWeather patterns with interpretable weather-based cues, and to unmeasured, site-based patterns. We use this approach to study GxWeather for the timings of vegetative and reproductive development in the deeply genetically diverged Gulf and Midwest subpopulations of switchgrass.

Our analysis of the timing of vegetative growth in the Gulf and in Both subpopulations revealed many alleles with rank-changing GxE, or sign changes in their effects between the Texas and North gardens (Figure 3 B). As phenological timings are major components of plant fitness, this result supports theoretical models that local adaptation should involve antagonistic pleiotropy at the level of individual loci (26–29). In addition, it advances local adaptation research by using more than two field sites and a wide range of genetic variation to determine that rank-changing GxE for phenological timings is common in small genomic regions (14, 43, 44).

Our analysis of the timing of flowering showed that the Gulf and Midwest subpopulations have distinct GxWeather: flowering timing in the Midwest subpopulation has photoperiod-related genetic variation, in that flowering timing covaries with a day length change signal two days before flowering occurs. In contrast, the Gulf subpopulation does not have genetic variation in flowering that covaries with a photoperiod cue. Instead, the Gulf subpopulation has genetic variation in flowering that covaries with the rainfall that occurs in the week prior to flowering. Three genomic regions affecting flowering that we re-estimated across all eight sites were also supported by QTL from an independent mapping population at these sites (Figure 4 B).

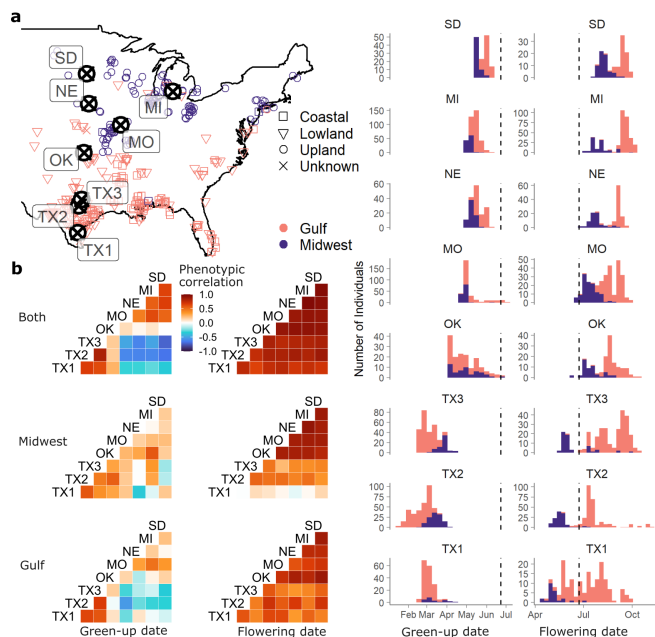
Identifying the environmental cues that are predictive of, or even correlated with, plant phenotypic responses remains a major challenge to studies interrogating gene action across many natural environments. The GxWeather photoperiod

and cumulative rainfall cues we identify here are functions of the genotypes measured and capture only a minority of SNP effects on flowering. We could only assign SNP effects to a GxWeather covariance structures in four of the six phenotype & genetic subpopulations we modeled. It’s likely we did not include some GxWeather important to these phenological cues - for example, overwintering parameters that might cause variation in the start of vegetative growth. More generally, it is difficult to predict the time scales over which individuals may integrate environmental cues, particularly in perennial species which may integrate these cues over longer time scales. If this integration time itself varies between individuals, the covariance structures we modeled cannot reflect this, though these structures would likely be highly correlated with GxWeather structures we did include. Our approach offers an opportunity to specify multiple environmental cues and compete them to explain patterns of genetic effects, allowing us to detect how important these cues are genome-wide, and how strongly each cue influences each SNP. This is a key development to further improve our understanding of genetic variation in GxE.

## Figures and Tables

**Table 1. Weather variables and time frames prior to the start of vegetative and reproductive growth used to construct the hypothesis-based covariance matrices. Covariance matrices were constructed for, and tested on, three subpopulations and two phenological traits. The correlations between values of these weather variables for genetically identical plants grown in different gardens were used to fill off-diagonal cells of the covariance matrices. Narrow-sense heritabilities for these values at each garden were used for the diagonal cells.**

Weather variable	Time Frame	Covariance Matrices
1. cumulative growing degree days at 12C in the time frame	(1-7), 14, 21, 28 days prior to the phenological date	60
2. cumulative rainfall in the time frame	(1-7), 14, 21, 28 days prior to the phenological date	60
3. day length (hours) on a specific day indicated by the time frame	(1-7) and 14 days prior to the phenological date	48
4. day length change (seconds) on a specific day indicated by the time frame	(1-7) and 14 days prior to the phenological date	48
5. average temperature in the time frame	(1-7), 14, 21, 28 days prior to the phenological date	60



**Figure 1.** Characterization of the timing of the onset of vegetative (green-up) and reproductive (flowering) growth in the switchgrass diversity panel. (a) Map and trait histograms of green-up and flowering dates across two genetically distinct switchgrass subpopulations and eight common gardens. Purple represents individuals from the Midwest genetic subpopulation, and pink individuals from the Gulf subpopulation; map positions represent the original collection locations for the genotypes, and shapes represent the ecotype of the genotype. Histogram vertical dashed lines indicate the summer solstice. Common gardens are arranged in latitudinal order. (b) Phenotypic correlations between clonal replicates planted at eight common gardens, within and between two genetic subpopulations.

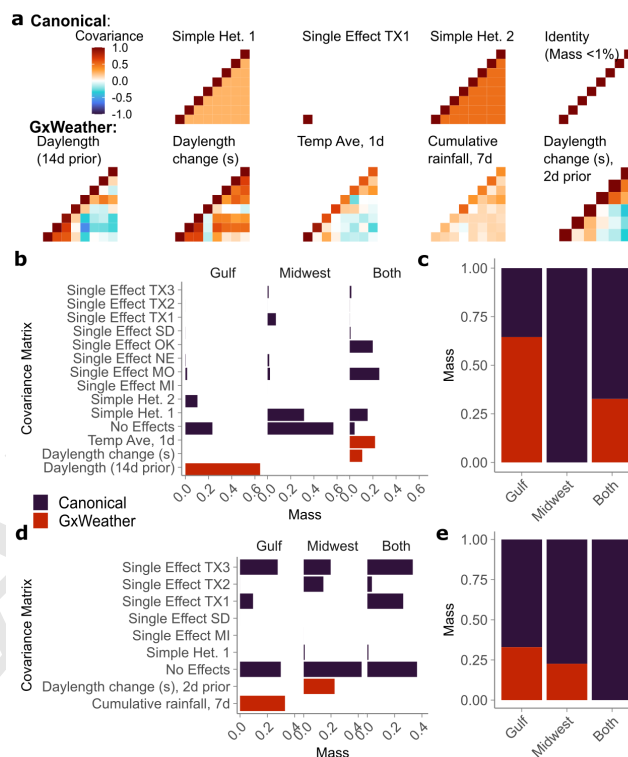
## Materials and Methods

Whenever possible, plant material will be shared upon request. Source data and code to replicate these analyses are available at: <https://github.com/Alice-MacQueen/pvdiv-phenology-gxe.git>. SNP data to replicate these analyses are available from the UT dataverse at <https://doi.org/10.18738/T8/A604BU>.

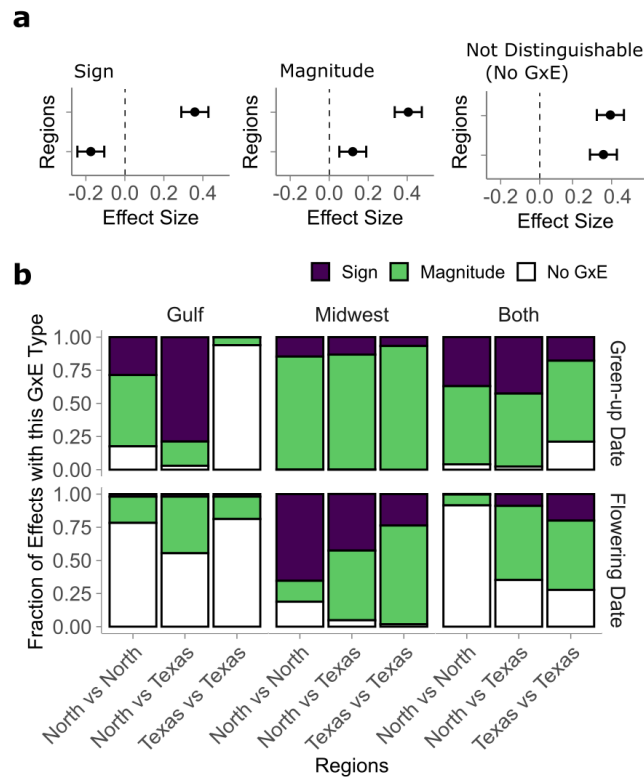
## Scoring onset of vegetative and reproductive development in two mapping panels.

In 2019, we scored two phenological events every two days in two mapping populations of switchgrass, a diversity panel and a pseudo-F<sub>2</sub> cross, planted at eight common garden locations (39, 44, 45). We scored the onset of vegetative growth, or green-up date, as the day of the year when 50% of the tiller area of the crown of the plant cut the previous year had green growth. The onset of reproductive growth, or flowering date, was the day of the year when 50% of the plant tillers had panicles undergoing anthesis.

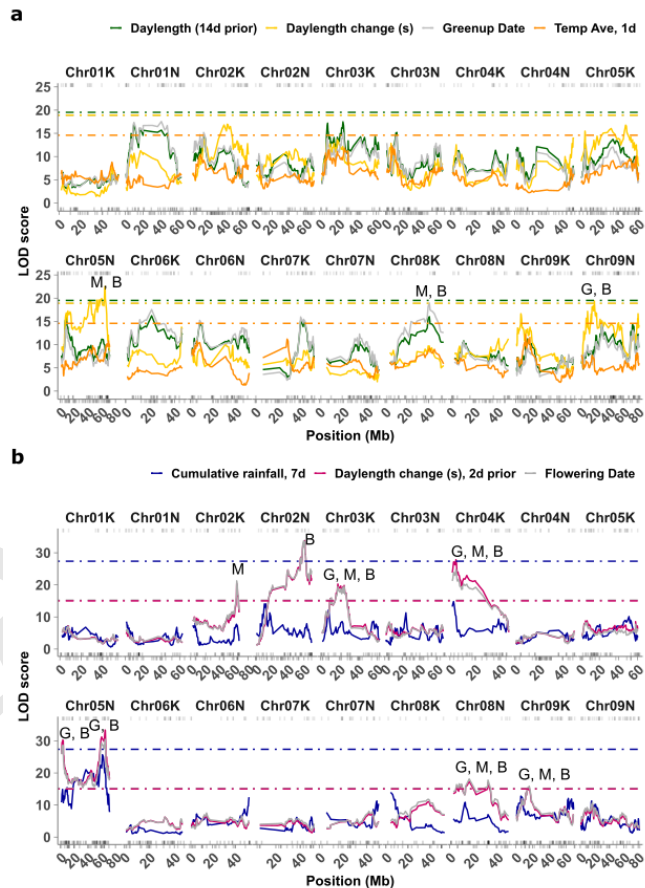
The formation and resequencing of the diversity panel has been described previously (37). The diversity panel contained 134 sequenced, clonally propagated individuals from the Midwest genetic subpopulation, and 229 from the Gulf genetic subpopulation. To allow for the possibility that different subpopulations had different strengths of connection between our phenotypes and genotypes (46), we conducted three sets of genetic analyses: on Gulf and Midwest genotypes separately, and on both subpopulations together ('Both' subpopulations). Analyses to determine narrow-sense heritability ( $h^2$ ) for green-



**Figure 2.** Example hypothesis-driven covariance matrices specified in *mash* and the posterior weights placed on all covariance matrices. (a) Common gardens are arranged in latitudinal order within the matrices. Top row: Four example canonical covariance matrices. Canonical matrices (purple) have simple interpretations, such as equal effects across all common gardens, or effects specific to a single common garden. Bottom row: Five example GxWeather covariance matrices specified for the green-up date or flowering date phenotype; these matrices were created from environment-specific correlations across eight common gardens, and are described in Table 1. (b,d) Total posterior weight placed on each covariance matrix type specified for (b) green-up date and (d) flowering date *mash* models, within and between two genetic subpopulations. Covariance matrices included in *mash* that had zero posterior weight in all three *mash* runs on the genetic subpopulations, such as the identity matrix, are not shown. (c,e) Total posterior weight placed on covariance matrices that were hypothesized or canonical, for the (c) green-up date phenotype and (e) flowering date phenotype.



**Figure 3.** Types of GxE present between pairs of jointly re-estimated effects in eight common gardens, for effects with  $lfsr < 0.05$ . a) Examples of effect patterns at three pairs of sites with three types of GxE. All effects are for an alternate allele, with the reference allele effect defined to be zero at both gardens and represented by the dashed vertical line. Sign: Effects that differ in sign at these pairs of gardens ( $p < 0.05$ ,  $lfsr$ ). Magnitude: Effects identical in sign ( $p < 0.05$ ,  $lfsr$ ) that differ in magnitude by a factor of  $>0.4$ . Not Distinguishable: Effects not distinguishable by magnitude nor sign of the effect, with no measurable GxE. b) The fraction of effects with each GxE type for the onset of vegetative growth (green-up date) and reproductive growth (flowering date), within and between two genetic subpopulations. Common gardens are grouped by the larger region they came from: North gardens are within the natural range of the Midwest subpopulation, and include MO, NE, MI, and SD, while Texas gardens are within the natural range of the Gulf subpopulation, and include TX1, TX2, and TX3.



**Figure 4.** Overlaps of QTL from an outbred pseudo-F2 cross and with jointly re-estimated SNP effects in the 1% tail of significance from a diversity panel. Dotted lines indicate permutation-based significance thresholds for each weather-related function. Stars indicate QTL with significant enrichment for SNPs in the 1% *mash* tail; G, M, and B indicate which subpopulation had enrichment: G - Gulf subpopulation, M - Midwest subpopulation, B - both subpopulations. Rug plots show genomic locations of SNPs in the 1% *mash* tail for flowering date for each subpopulation: Both subpopulations are above the plot panel, the Gulf subpopulation is above the x-axis, and the Midwest subpopulation is below the x-axis. a) QTL mapping for the onset of vegetative growth (green-up date), and three weather-related functions of green-up date. b) QTL mapping for the onset of reproductive growth (flowering date), and two weather-related functions of flowering date.

up and flowering were done using linear mixed models and followed (37) (SI Appendix, Section S6).

To confirm candidate genomic regions found using *mash* on the diversity panel, we analyzed flowering in an outbred pseudo-F2 cross between four individuals, two Midwest and two Gulf individuals. The formation of this mapping population has been described previously (39); additional details on QTL mapping can be found in SI Appendix, Section S7. To be directly comparable to the diversity panel data, only 2019 phenology data from the pseudo-F2 cross from the same eight common garden sites were used.

**Joint re-estimation of SNP effects to assess the frequency of rank-changing GxE and assignment of genome-wide patterns of GxE and GxWeather.** We were interested in specifying genetic models for trait variation that allowed more than one form of GxE, as we reasoned that different loci should display different forms of GxE. In addition, we were interested in an unbiased estimation of the frequency of rank-changing GxE relative to other forms of GxE, as the presence of rank-changing GxE at the level of individual loci is a key theoretical prediction of local adaptation. *Mash* allowed us to both specify multiple forms of GxE and GxWeather and conduct unbiased statistical tests for when SNP effects changed sign between common gardens.

To use *mash* on our diversity panel, we had to specify both a relatively uncorrelated set of covariance matrices, which in our case defined types of GxE and GxWeather between gardens, and we had to specify subsets of SNP effect estimates and standard errors for our traits at each common garden. To specify a set of covariance matrices, we first defined many covariance matrices, including GxWeather matrices that represented the correlation in weather cues between gardens before the phenological event (SI Appendix, Section S1), then implemented a model selection approach that used a greedy algorithm to evaluate if the log likelihood of the *mash* model was significantly improved as additional covariance matrices were included (SI Appendix, Section S4). To specify subsets of SNP effect estimates and standard errors to use for both the greedy algorithm and with the optimal set of covariance matrices, we first calculated best linear unbiased predictors (BLUPs) for each phenological trait in each genetic subpopulation and each common garden (SI Appendix, Section S2). Next, we determined effect estimates for 8.8 to 12.3 million SNPs per subpopulation by conducting garden-specific GWAS on these BLUPs using  $k$  vectors of singular values to correct for population structure, where  $k$  was the smallest integer value that made the genomic control coefficient closest to 1 (SI Appendix, Section S2). Singular values were computed using singular value decomposition of the matrix of all SNPs, with iterative SNP pruning and removal of regions in long-range linkage disequilibrium (47). Third, to make the *mash* models computationally feasible, we extracted two subsets of SNP effect estimates and standard errors from our GWAS effect estimates: (i) effects from a subset of “strong” tests corresponding to stronger effects on our traits; (ii) results from a random subset of all tests to correspond to an unbiased representation of all effects (SI Appendix, Section S3). We used the subset of random effects in our greedy *mash* algorithm (SI Appendix, Section S4) and used the subset of strong effects in our *mash* models with the optimal set of covariance matrices (SI Appendix, Section S5).

The loadings of genetic effects onto the multiple covariance structures specified in our *mash* models provided information on genome-wide patterns of SNP-environment interaction. In addition, the the GxWeather covariance structures allowed hypothesis testing of specific weather variables as cues for the start of vegetative and reproductive growth. Say that our diversity panel contains a SNP in the gene *CONSTANS* (CO), a well-known flowering time regulator, and that only one of the alleles affects the promotion of flowering in a photoperiod-dependent manner. In that case, the joint estimate of effects for that SNP could have a high mixture proportion, or mass, on a covariance matrix created using a photoperiod-based environmental cue, such as day length at some interval prior to flowering. In our data, we would infer that the effect of that SNP on flowering was cued by the weather variable used to create the GxWeather covariance structure.

Our joint re-estimation of SNP effects also allowed us to characterize the overall patterns of GxE in the set of SNPs where there was pairwise significance of effects at pairs of gardens. To do this, we used the ‘get\_GxE’ function of the switchgrassGWAS R package. First, this function determines the set of SNPs with evidence of significant effects in both conditions for all pairs of conditions using local false sign rates (lfsr) as the significance criteria. Then, this function determines if effects significant in both conditions are of opposite sign.

Using the lfsr rather than the local false discovery rate (lfdr) is a critical change in our ability to detect alleles directly contributing to rank-changing GxE between environments. The lfdr, like other measures of FDR, focuses on if we have enough evidence to reject the null hypothesis that an effect  $j$  is 0, or that there is a significant effect. Previous studies of antagonistic pleiotropy (e.g. (44)) have used the lfdr or equivalent statistical tests to detect antagonistic pleiotropy. These tests were conservative, in that they required two non-zero effects of different signs, while tests for differential sensitivity required only one non-zero effect. This previous work recognized that this testing bias could lead to undercounting occurrences of antagonistic pleiotropy (30, 33), and sought to reduce it by permutation (32). However, using the lfsr to test for allelic effects that differ in sign does not undercount these occurrences, as this statistic answers a fundamentally different question. For each effect  $j$ , the  $lfsr_j$  is defined as the probability that we make an error in the sign of effect  $j$  if we were forced to declare the effect positive or negative (41). Thus, rather than asking “Are these two effects different?” - as we reasonably expect two effects to be, even if this difference cannot be measured - the local false sign rate answers a more meaningful question: Can we be confident in the sign of this effect?

In addition, the get\_GxE function also sets an arbitrary threshold to count an effect as changing in magnitude between environments, commonly known as differential sensitivity or a change in amplitude of the effect. For differential sensitivity, this function determines if effects significant in both conditions are of the same sign and of a magnitude (not tested for significance) that differs by a factor of 0.4 or more. The remaining effects that are significant in both conditions have the same effect sign and similar effect magnitudes and we denote these effects as having no GxE. The distinction between effects with different magnitudes is arbitrary but useful to



fully characterize how effects vary across environments to ultimately influence phenotypes. Our use of the lfsr to determine significance and our specification that SNP effects must be significant in both conditions to be included means that our tests for alleles with rank-changing GxE carry an equal

statistical burden to those measuring differential sensitivity and effects without GxE.

# References

1. W. L. Bauerle, *et al.*, [Photoperiodic regulation of the seasonal pattern of photosynthetic capacity and the implications for carbon cycling](#). *Proceedings of the National Academy of Sciences* **109**, 8612–8617 (2012).
2. F. Andrés, G. Coupland, The genetic basis of flowering responses to seasonal cues. *Nature Reviews Genetics* **13**, 627–639 (2012).
3. C. Körner, D. Basler, Phenology under global warming. *Science* **327**, 1461–1462 (2010).
4. R. Maple, P. Zhu, J. Hepworth, J.-W. Wang, C. Dean, [Flowering time: From physiology, through genetics to mechanism](#). *Plant Physiology* **195**, 190–212 (2024).
5. C. A. Botero, F. J. Weissing, J. Wright, D. R. Rubenstein, Evolutionary tipping points in the capacity to adapt to environmental change. *Proceedings of the National Academy of Sciences* **112**, 184–189 (2015).
6. B. K. Blackman, Interacting duplications, fluctuating selection, and convergence: The complex dynamics of flowering time evolution during sunflower domestication. *Journal of experimental botany* **64**, 421–431 (2013).
7. L. P. Henry, R. H. Watson, B. K. Blackman, Transitions in photoperiodic flowering are common and involve few loci in wild sunflowers (helianthus; asteraceae). *American Journal of Botany* **101**, 1748–1758 (2014).
8. J. Ågren, C. G. Oakley, S. Lundemo, D. W. Schemske, Adaptive divergence in flowering time among natural populations of arabidopsis thaliana: Estimates of selection and QTL mapping. *Evolution* **71**, 550–564 (2017).
9. B. Brachi, *et al.*, Linkage and association mapping of arabidopsis thaliana flowering time in nature. *PLoS genetics* **6**, e1000940 (2010).
10. E. L. Dittmar, C. G. Oakley, J. Ågren, D. W. Schemske, Flowering time QTL in natural populations of arabidopsis thaliana and implications for their adaptive value. *Molecular ecology* **23**, 4291–4303 (2014).
11. X. Li, T. Guo, Q. Mu, X. Li, J. Yu, Genomic and environmental determinants and their interplay underlying phenotypic plasticity. *Proceedings of the National Academy of Sciences* **115**, 6679–6684 (2018).
12. J. A. Romero Navarro, *et al.*, A study of allelic diversity underlying flowering-time adaptation in maize landraces. *Nature genetics* **49**, 476–480 (2017).
13. D. Bustos-Korts, *et al.*, [Exome sequences and multi-environment field trials elucidate the genetic basis of adaptation in barley](#). *The Plant Journal* **99**, 1172–1191 (2019).

14. S. M. Wadgymar, *et al.*, Identifying targets and agents of selection: Innovative methods to evaluate the processes that contribute to local adaptation. *Methods in Ecology and Evolution* **8**, 738–749 (2017).
15. M. Blümel, N. Dally, C. Jung, Flowering time regulation in crops—what did we learn from arabidopsis? *Current opinion in biotechnology* **32**, 121–129 (2015).
16. A. Fournier-Level, *et al.*, [Predicting the evolutionary dynamics of seasonal adaptation to novel climates in \*Arabidopsis thaliana\*](#). *Proceedings of the National Academy of Sciences* **113** (2016).
17. C. Jung, A. E. Müller, Flowering time control and applications in plant breeding. *Trends in plant science* **14**, 563–573 (2009).
18. A. Turner, J. Beales, S. Faure, R. P. Dunford, D. A. Laurie, The pseudo-response regulator ppd-H1 provides adaptation to photoperiod in barley. *Science* **310**, 1031–1034 (2005).
19. S. Faure, *et al.*, Mutation at the circadian clock gene EARLY MATURITY 8 adapts domesticated barley (*hordeum vulgare*) to short growing seasons. *Proceedings of the National Academy of Sciences* **109**, 8328–8333 (2012).
20. H.-Y. Hung, *et al.*, ZmCCT and the genetic basis of day-length adaptation underlying the postdomestication spread of maize. *Proceedings of the National Academy of Sciences* **109**, E1913–E1921 (2012).
21. S. Zakhrebekova, *et al.*, Induced mutations in circadian clock regulator mat-a facilitated short-season adaptation and range extension in cultivated barley. *Proceedings of the National Academy of Sciences* **109**, 4326–4331 (2012).
22. Y. Yang, Q. Peng, G.-X. Chen, X.-H. Li, C.-Y. Wu, OsELF3 is involved in circadian clock regulation for promoting flowering under long-day conditions in rice. *Molecular Plant* **6**, 202–215 (2013).
23. P. Pin, O. Nilsson, The multifaceted roles of FLOWERING LOCUS t in plant development. *Plant, cell & environment* **35**, 1742–1755 (2012).
24. J. L. Weller, *et al.*, Parallel origins of photoperiod adaptation following dual domestications of common bean. *Journal of Experimental Botany* **70**, 1209–1219 (2019).
25. E. Weine, S. P. Smith, R. K. Knowlton, A. Harpak, Tradeoffs in modeling context dependency in complex trait genetics. *bioRxiv* (2024) <https://doi.org/10.1101/2023.06.21.545998>.
26. H. Levene, Genetic equilibrium when more than one ecological niche is available. *The American Naturalist* **87**, 331–333 (1953).
27. J. Felsenstein, The theoretical population genetics of variable selection and migration. *Annual review of genetics* **10**, 253–280 (1976).
28. T. J. Kawecki, D. Ebert, Conceptual issues in local adaptation. *Ecology letters* **7**, 1225–1241 (2004).
29. P. W. Hedrick, Genetic polymorphism in heterogeneous environments: A decade later. *Annual review of ecology and systematics* **17**, 535–566 (1986).
30. D. L. Des Marais, K. M. Hernandez, T. E. Juenger, Genotype-by-environment interaction and plasticity: Exploring genomic responses of plants to the abiotic environment. *Annual Review of Ecology, Evolution, and Systematics* **44**, 5–29 (2013).
31. E. Millet, *et al.*, [Genome-wide analysis of yield in Europe: allelic effects as functions of drought and heat scenarios](#). *Plant Physiology*, pp.00621.2016 (2016).
32. J. T. Anderson, C.-R. Lee, C. A. Rushworth, R. I. Colautti, T. Mitchell-Olds, Genetic trade-offs and conditional neutrality contribute to local adaptation. *Molecular ecology* **22**, 699–708 (2013).
33. J. T. Anderson, J. H. Willis, T. Mitchell-Olds, Evolutionary genetics of plant adaptation. *Trends in Genetics* **27**, 258–266 (2011).
34. R. B. Mitchell, K. J. Moore, L. E. Moser, J. O. Fritz, D. D. Redfearn, Predicting developmental morphology in switchgrass and big bluestem. *Agronomy Journal* **89**, 827–832 (1997).
35. D. J. Parrish, J. H. Fike, The biology and agronomy of switchgrass for biofuels. *BPTS* **24**, 423–459 (2005).
36. M. Casler, K. P. Vogel, C. Taliaferro, R. Wynia, Latitudinal adaptation of switchgrass populations. *Crop Science* **44**, 293–303 (2004).
37. J. T. Lovell, *et al.*, Genomic mechanisms of climate adaptation in polyploid bioenergy switchgrass. *Nature* **590**, 438–444 (2021).
38. C. L. Porter Jr, An analysis of variation between upland and lowland switchgrass, *panicum virgatum* L., in central oklahoma. *Ecology* **47**, 980–992 (1966).
39. E. R. Milano, D. B. Lowry, T. E. Juenger, The genetic basis of upland/lowland ecotype divergence in switchgrass (*panicum virgatum*). *G3: Genes, Genomes, Genetics* **6**, 3561–3570 (2016).

1241	40.	S. M. Urbut, G. Wang, P. Carbonetto, M. Stephens,	1303
1242		Flexible statistical methods for estimating and testing	1304
1243		effects in genomic studies with multiple conditions.	1305
1244		<i>Nature genetics</i> <b>51</b> , 187–195 (2019).	1306
1245			1307
1246	41.	M. Stephens, <a href="#">False discovery rates: a new deal</a> . <i>Bio-</i>	1308
1247		<i>statistics</i> <b>18</b> , 275–294 (2016).	1309
1248			1310
1249	42.	S. Choi, <i>et al.</i> , <a href="#">A single amino acid change led to</a>	1311
1250		<a href="#">structural and functional differentiation of <i>PvHd1</i> to</a>	1312
1251		<a href="#">control flowering in switchgrass</a> . <i>Journal of Experimen-</i>	1313
1252		<i>tal Botany</i> <b>74</b> , 5532–5546 (2023).	1314
1253			1315
1254	43.	O. Savolainen, M. Lascoux, J. Merilä, Ecological	1316
1255		genomics of local adaptation. <i>Nature Reviews Genetics</i>	1317
1256		<b>14</b> , 807–820 (2013).	1318
1257			1319
1258	44.	D. B. Lowry, <i>et al.</i> , QTL× environment interactions	1320
1259		underlie adaptive divergence in switchgrass across a	1321
1260		large latitudinal gradient. <i>Proceedings of the National</i>	1322
1261		<i>Academy of Sciences</i> <b>116</b> , 12933–12941 (2019).	1323
1262			1324
1263	45.	J. T. Lovell, <i>et al.</i> , <a href="#">Genomic mechanisms of climate</a>	1325
1264		<a href="#">adaptation in polyploid bioenergy switchgrass</a> . <i>Nature</i> ,	1326
1265		1–7 (2021).	1327
1266			1328
1267	46.	A. Korte, A. Farlow, The advantages and limitations	1329
1268		of trait analysis with GWAS: A review. <i>Plant methods</i>	1330
1269		<b>9</b> , 29 (2013).	1331
1270			1332
1271	47.	F. Privé, H. Aschard, M. G. B. Blum, <a href="#">Efficient</a>	1333
1272		<a href="#">management and analysis of large-scale genome-wide</a>	1334
1273		<a href="#">data with two r packages: Bigstatsr and bigsnpr</a> (2017).	1335
1274			1336
1275			1337
1276	<b>ACKNOWLEDGMENTS.</b> We thank the Brackenridge Field		1338
1277	laboratory, the Ladybird Johnson Wildflower Center, and the		1339
1278	Juenger laboratory for support with plant care and propagation.		1340
1279	This material is based upon work supported in part by the Great		1341
1280	Lakes Bioenergy Research Center, U.S. Department of Energy,		1342
1281	Office of Science, Office of Biological and Environmental Research		1343
1282	under Award Numbers DE-SC0018409 and DE-FC02-07ER64494,		1344
1283	the US Department of Energy Awards DESC0014156 to T.E.J.,		1345
1284	DE-SC0017883 to D.B.L, National Science Foundation PGRP		1346
1285	Awards IOS0922457 and IOS1444533 to T.E.J, and the Long-		1347
1286	term Ecological Research Program (DEB 1832042) at the Kellogg		1348
1287	Biological Station.		1349
1288			1350
1289			1351
1290			1352
1291			1353
1292			1354
1293			1355
1294			1356
1295			1357
1296			1358
1297			1359
1298			1360
1299			1361
1300			1362
1301			1363
1302			1364