

$$1. (a) P(N=1) = P(X_1 < X_2)$$

$$= \int_0^{\infty} \lambda_1 e^{-\lambda_1 x_1} P(X_2 > x_1) dx_1$$

$$= \int_0^{\infty} \lambda_1 e^{-\lambda_1 x_1} e^{-\lambda_2 x_1} dx_1$$

$$= \frac{\lambda_1}{\lambda_1 + \lambda_2}.$$

$$P(N=2) = 1 - P(N=1) = \frac{\lambda_2}{\lambda_1 + \lambda_2}.$$

$$(b) P(U > t) = P(X_1 > t, X_2 > t)$$

$$= P(X_1 > t) P(X_2 > t)$$

$$= e^{-(\lambda_1 + \lambda_2)t}$$

$$(c) P(U > t | N=1) = P(U > t | X_1 < X_2)$$

$$= P(X_1 > t | X_1 < X_2)$$

$$= \frac{P(t < X_1 < X_2)}{P(X_1 < X_2)}$$

$$= \frac{\lambda_1 + \lambda_2}{\lambda_1} \cdot \int_t^{\infty} \lambda_1 e^{-\lambda_1 x_1} P(X_2 > x_1) dx_1$$

$$= e^{-(\lambda_1 + \lambda_2)t}$$

$$= P(U > t)$$

$$\text{同理 } P(U > t | N=2) = P(U > t).$$

由定义知 U, N 独立.

$$(d) P(W > t | N=1) = P(X_2 > X_1 + t | X_1 < X_2)$$

$$= \frac{P(X_2 > X_1 + t)}{P(X_2 > X_1)}$$

$$= \frac{\lambda_1 + \lambda_2}{\lambda_1} \int_0^{\infty} \lambda_1 e^{-\lambda_1 x_1} P(X_2 > x_1 + t) dx_1$$

$$= (\lambda_1 + \lambda_2) \int_0^{\infty} e^{-\lambda_1 x_1} e^{-\lambda_2 (x_1 + t)} dx_1$$

$$= e^{-\lambda_2 t}$$

$$\text{同理 } P(W > t | N=2) = e^{-\lambda_1 t}$$

$$\begin{aligned}
 (e) \quad X_1 < X_2 \text{ 时}, P(W > t, U > u) &= P(X_1 > u, X_2 > X_1 + t) \\
 &= \int_u^\infty \lambda e^{-\lambda_1 x_1} P(X_2 > x_1 + t) dx_1 \\
 &= e^{-\lambda_1 u} e^{-\lambda_2 (u+t)} \\
 &= P(U > u) P(W > t).
 \end{aligned}$$

同理 $X_1 \geq X_2$ 时也成立 $P(W > t, U > u) = P(W > t) P(U > u)$.

由定义知 W, U 独立

2. 由于震动次数 $\sim P(\lambda)$, 知震动间隔 $\sim \text{Exp}(\lambda)$
 k 个独立 $\text{Exp}(\lambda)$ 的和 $\sim \Gamma(k, \lambda)$, 即 $T \sim \Gamma(k, \lambda)$.

$$f_T(t) = \begin{cases} \frac{1}{\Gamma(k)} \lambda^k t^{k-1} e^{-\lambda t}, & t > 0, \\ 0, & t \leq 0. \end{cases}$$

$$\begin{aligned}
 3. \quad P(Y(t) = y) &= \sum_{x=y}^{\infty} C_x^y p^y (1-p)^{x-y} P(X(t) = x) \\
 &= \sum_{x=y}^{\infty} C_x^y p^y (1-p)^{x-y} \frac{(\lambda t)^x}{x!} e^{-\lambda t} \\
 &= \sum_{x=y}^{\infty} p^y (1-p)^{x-y} \frac{(\lambda t)^x}{y! (x-y)!} e^{-\lambda t} \\
 &= p^y \frac{(\lambda t)^y}{y!} e^{-\lambda t} \sum_{x=y}^{\infty} \frac{((1-p)\lambda t)^{x-y}}{(x-y)!} \\
 &= p^y \frac{(\lambda t)^y}{y!} e^{-\lambda t} e^{(1-p)\lambda t} \\
 &= \frac{(p\lambda t)^y}{y!} e^{-p\lambda t}
 \end{aligned}$$

故 $Y(t)$ 是参数为 $p\lambda$ 的泊松过程.

$$\begin{aligned}
 4 \quad P(Z_1(t)=z) &= \sum_{x=0}^z P(X(t)=x) P(Y(t)=z-x) \\
 &= \sum_{x=0}^z \frac{(\lambda_1 t)^x}{x!} e^{-\lambda_1 t} \frac{(\lambda_2 t)^{z-x}}{(z-x)!} e^{-\lambda_2 t} \\
 &= \frac{((\lambda_1 + \lambda_2)t)^z}{z!} e^{-(\lambda_1 + \lambda_2)t} \sum_{x=0}^z C_z^x \left(\frac{\lambda_1}{\lambda_1 + \lambda_2}\right)^x \left(\frac{\lambda_2}{\lambda_1 + \lambda_2}\right)^{z-x} \\
 &= \frac{((\lambda_1 + \lambda_2)t)^z}{z!} e^{-(\lambda_1 + \lambda_2)t}
 \end{aligned}$$

故 $Z_1(t)$ 是泊松过程, 参数为 $\lambda_1 + \lambda_2$.

$Z_2(t)$ 一定不是泊松过程, 因为 $Z_2(t) < 0$ 不是零概率的.

$Z_3(t)$ 一定不是泊松过程, 因为 $\lim_{h \rightarrow 0} \frac{P(Z_3(t+h)=1)}{h}$ 不是常数.

5. 设 $X(t)$ 为参数 $\lambda=3$ 的泊松过程.

$$(a) P(X(4)=0) = \frac{(3 \times 4)^0}{0!} e^{-3 \times 4} = e^{-12}$$

(b) 设 T 为首次时刻 (从 12 时开始).

$$P(T \leq t) = P(X(t) > 0) = 1 - P(X(t) = 0) = 1 - e^{-3t}.$$

$$f_T(t) = P'_t(T \leq t) = 3e^{-3t}.$$

$$6 \quad \text{cov}(X(t), X(t+\tau)) = E(X(t)X(t+\tau)) - E(X(t))E(X(t+\tau))$$

$$= E(X(t)X(t+\tau)) - \lambda t \cdot \lambda(t+\tau)$$

$$E(X(t)X(t+\tau) | X(t)=n) = n E(X(t+\tau) | X(t)=n)$$

$$= n E(n + X(\tau))$$

$$= n(n + \lambda\tau)$$

$$= n^2 + n\lambda\tau.$$

$$E(X(t)X(t+\tau)) = E(E(X(t)X(t+\tau) | X(t)))$$

$$= \sum_{n=0}^{\infty} (n^2 + n\lambda\tau) P(X(t)=n)$$

$$= \sum_{n=0}^{\infty} (n^2 + n\lambda\tau) \frac{(\lambda t)^n}{n!} e^{-\lambda t}$$

$$= e^{-\lambda t} \lambda t \sum_{n=0}^{\infty} (n + \lambda\tau + 1) \frac{(\lambda t)^n}{n!}$$

$$= \lambda t (\lambda t + \lambda\tau + 1).$$

$$\text{故 } \text{cov}(X(t), X(t+\tau)) = \lambda t.$$

$$7. P_{00}(t+h) = P_{00}(t)P_{00}(h) + P_{01}(t)P_{10}(h).$$

$$\text{令 } h \rightarrow 0, P_{00}(t+h) = P_{00}(t)(1-\lambda h) + P_{01}(t)\mu h.$$

$$\begin{aligned} \mathbb{E}P'_{00}(t) &= -\lambda P_{00}(t) + \mu P_{01}(t) \\ &= -\lambda P_{00}(t) + \mu(1-P_{00}(t)) \\ &= \mu - (\lambda + \mu)P_{00}(t). \end{aligned}$$

$$\begin{aligned} (e^{(\lambda+\mu)t}P_{00}(t))' &= e^{(\lambda+\mu)t}(P'_{00}(t) + (\lambda+\mu)P_{00}(t)) \\ &= \mu e^{(\lambda+\mu)t}. \end{aligned}$$

$$\begin{aligned} P_{00}(t) &= e^{-(\lambda+\mu)t} \int \mu e^{(\lambda+\mu)t} dt \\ &= \frac{\mu}{\lambda+\mu} + C e^{-(\lambda+\mu)t} \\ &= \frac{\mu}{\lambda+\mu} + \frac{\lambda}{\lambda+\mu} e^{-(\lambda+\mu)t} \quad (P_{00}(0)=1) \end{aligned}$$

8. 由于状态改变概率与所处状态无关,

每次改变时间 $\sim \text{Exp}(\lambda)$, i.i.d.

故 $X(t)$ 为参数 λ 的泊松过程, 即 $P_n(t) = \frac{(\lambda t)^n}{n!} e^{-\lambda t}$.

$$9. P_1(t+h) = P_1(t)(1-\lambda_1 h) + P_2(t)\lambda_2 h, h \rightarrow 0.$$

$$\begin{aligned} \Rightarrow P_1'(t) &= -\lambda_1 P_1(t) + \lambda_2 P_2(t) \\ &= -\lambda_1 P_1(t) + \lambda_2(1-P_1(t)) \\ &= \lambda_2 - (\lambda_1 + \lambda_2)P_1(t). \end{aligned}$$

$$\begin{aligned} \text{同题7, 易得 } P_1(t) &= \frac{\lambda_2}{\lambda_1 + \lambda_2} + C e^{-(\lambda_1 + \lambda_2)t} \\ &= \frac{\lambda_2}{\lambda_1 + \lambda_2} - \frac{\lambda_1}{\lambda_1 + \lambda_2} e^{-(\lambda_1 + \lambda_2)t} \quad (P_1(0)=0) \end{aligned}$$

$$\begin{aligned} P_2(t) &= 1 - P_1(t) \\ &= \frac{\lambda_1}{\lambda_1 + \lambda_2} + \frac{\lambda_2}{\lambda_1 + \lambda_2} e^{-(\lambda_1 + \lambda_2)t}. \end{aligned}$$

$$10. p'_{00}(t) = -\lambda_2 p_{00}(t)$$

$$p'_{01}(t) = \lambda_2 p_{00}(t) - \lambda_1 p_{01}(t)$$

$$p'_{02}(t) = \lambda_1 p_{01}(t) - \lambda_2 p_{02}(t)$$

⋮

$$\text{求和, 有 } \sum_{i=0}^{\infty} i p'_{0i}(t) = \sum_{i=0}^{\infty} \lambda_1 p_{0i}(t) + \sum_{i=0}^{\infty} \lambda_2 p_{0i}(t)$$

$$= \lambda_1 p_1(t) + \lambda_2 p_2(t)$$

$$= \frac{\lambda_1 \lambda_2}{\lambda_1 + \lambda_2} (1 - e^{-(\lambda_1 + \lambda_2)t}) + \frac{\lambda_1 \lambda_2}{\lambda_1 + \lambda_2} + \frac{\lambda_2^2}{\lambda_1 + \lambda_2} e^{-(\lambda_1 + \lambda_2)t}$$

$$= \frac{2\lambda_1 \lambda_2}{\lambda_1 + \lambda_2} + \frac{\lambda_2(\lambda_2 - \lambda_1)}{\lambda_1 + \lambda_2} e^{-(\lambda_1 + \lambda_2)t}$$

$$\triangleq f(t).$$

$$E(X(t)) = \sum_{i=0}^{\infty} i p_{0i}(t)$$

$$= \int_0^t f(t) dt$$

$$= \frac{2\lambda_1 \lambda_2}{\lambda_1 + \lambda_2} t + \frac{\lambda_2(\lambda_1 - \lambda_2)}{(\lambda_1 + \lambda_2)^2} e^{-(\lambda_1 + \lambda_2)t} + \frac{\lambda_2(\lambda_2 - \lambda_1)}{(\lambda_1 + \lambda_2)^2} \quad (E(X(0)) = 0).$$

$$11 \quad F(t+h) = F(t) + (1-F(t))g(t)h, \quad h \rightarrow 0.$$

$$\Rightarrow F'(t) = (1-F(t))g(t).$$

$$\text{设 } G(x) = \int_0^x g(x) dx$$

$$(e^{G(t)} F(t))' = e^{G(t)} g(t)$$

$$\Rightarrow F(t) = e^{-G(t)} \int e^{G(t)} g(t) dt$$

$$= e^{-G(t)} \cdot (e^{G(t)} + C)$$

$$= 1 - e^{-G(t)}. \quad (F(0) = 0)$$

12 1a) 设 $f(s) = P(E(s)=0)$

$$f(s+h) = f(s) \cdot (1 - \lambda(s)h), h \rightarrow 0.$$

$$\Rightarrow f'(s) = -\lambda(s)f(s)$$

$$\Rightarrow (\ln f(s))' = -\lambda(s)$$

$$\Rightarrow f(s) = e^{-\int_0^s \lambda(t) dt} \quad (f(0)=1).$$

(b) $P_k(s+h) = P_{k-1}(s)\lambda(s)h + P_k(s)(1 - \lambda(s)h), h \rightarrow 0, k > 0.$

$$\Rightarrow P'_k(s) = P_{k-1}(s)\lambda(s) - P_k(s)\lambda(s)$$

$$\text{设 } G(s) = \int_0^s \lambda(t) dt.$$

$$(e^{G(s)} P_k(s))' = e^{G(s)} P_{k-1}(s) \lambda(s)$$

$$P_k(s) = e^{-G(s)} \int_0^s e^{G(t)} P_{k-1}(t) \lambda(t) dt, k > 0.$$

$$P_0(s) = f(s) = e^{-G(s)}.$$

$$P_1(s) = e^{-G(s)} \int_0^s \lambda(t) dt$$

$$= G(s) e^{-G(s)}.$$

$$P_2(s) = e^{-G(s)} \int_0^s G(t) \lambda(t) dt$$

$$= \frac{1}{2} G^2(s) e^{-G(s)}.$$

⋮

$$1) \text{归纳假设 } P_k(s) = \frac{1}{k!} G^k(s) e^{-G(s)}, k \geq 0.$$

13. 先考虑一条电缆, 设 $f(t)$ 为其在 t 时刻工作的概率.

$$f(t+h) = f(t)(1-\lambda h) + (1-f(t))\mu h, h \rightarrow 0.$$

$$\Rightarrow f'(t) = \mu - (\lambda + \mu)f(t)$$

$$\text{问题7可得 } f(t) = \frac{1}{\lambda + \mu} (\mu + \lambda e^{-(\lambda + \mu)t}) \quad (f(0) = 1).$$

由于两条电缆独立, 所求即为

$$f^2(t) = \frac{1}{(\lambda + \mu)^2} (\mu + \lambda e^{-(\lambda + \mu)t})^2.$$

$$14. P(\text{首现时数为 } k) = \left(\frac{\lambda}{\lambda + \mu}\right)^{k-1} \frac{\mu}{\lambda + \mu}.$$

$$15. \lambda_i = i\lambda + a, \mu_i = i\mu.$$

$$\tau_j = \frac{\lambda_0 \cdots \lambda_{j-1}}{\mu_1 \cdots \mu_j} = \frac{1}{j! \mu^j} \prod_{i=0}^{j-1} (i\lambda + a) = C_{j-1+a/\lambda}^j \left(\frac{\lambda}{\mu}\right)^j.$$

$$\sum_{j=0}^{\infty} \tau_j = \sum_{j=0}^{\infty} C_{j+a/\lambda-1}^j \left(\frac{\lambda}{\mu}\right)^j = \left(1 - \frac{\lambda}{\mu}\right)^{-a/\lambda}. \quad \left(\sum_{x=0}^{\infty} C_{x+r-1}^x p^r (1-p)^x = 1\right)$$

$$\sum_{j=0}^{\infty} \tau_j < \infty, \text{ 故 } p_n = \tau_n / \sum_{j=0}^{\infty} \tau_j = C_{n+a/\lambda-1}^n \left(\frac{\lambda}{\mu}\right)^n \left(1 - \frac{\lambda}{\mu}\right)^{a/\lambda}$$

16. 显然这是生灭过程.

$$\lambda_i = \begin{cases} \lambda, & 0 \leq i < m, \\ 0, & i \geq m. \end{cases}$$

$$\mu_i = \begin{cases} i\mu, & 1 \leq i \leq m, \\ 0, & i=0 \vee i > m. \end{cases}$$

$$\tau_j = \frac{\lambda_0 \cdots \lambda_{j-1}}{\mu_1 \cdots \mu_j} = \frac{\lambda^j}{j! \mu^j}, \quad 0 \leq j \leq m, \quad \tau_j = 0, j > m.$$

$$\sum_{j=0}^{\infty} \tau_j = \sum_{j=0}^m \frac{\lambda^j}{j! \mu^j} < \infty.$$

$$p_n = \tau_n / \sum_{j=0}^m \tau_j = \frac{\lambda^n}{n! \mu^n} / \sum_{j=0}^m \frac{\lambda^j}{j! \mu^j}.$$

17. 取微元 h .

$$\begin{aligned} P(X_{t+h} - X_t = 1 | X_t = m) &= C_{N-m}^1 (\lambda h + o(h)) (1 - \lambda h + o(h))^{N-m-1} \\ &= (N-m) (\lambda h + o(h)) (1 + o(1)) \\ &= (N-m) \lambda h + o(h) \\ &\approx N \lambda h + o(h) \end{aligned}$$

$$\begin{aligned} P(X_{t+h} - X_t = 0 | X_t = m) &= (1 - \lambda h + o(h))^{N-m} \\ &= 1 - (N-m) \lambda h + o(h) \\ &\approx 1 - N \lambda h + o(h). \end{aligned}$$

根据定义, X_t 近似于参数为 $N\lambda$ 的泊松过程.

18. (i) 不具有. t 较大时不能将 $N-m$ 近似为 N

(ii) 不同 (i).

(iii) 有. 转移概率与 t 无关

(iv) 是. 未来与过去无关.

19. $\sum_{i=1}^{\infty} \pi_{j=1}^i \frac{\mu_j}{\lambda_j} = \infty$, 故必定灭亡.

$$\rho_i = \pi_{j=1}^{i-1} \lambda_j / \pi_{j=1}^i \mu_j = \begin{cases} \frac{1}{\lambda_i} \left(\frac{\lambda}{\mu} \right)^i, & 1 \leq i \leq k, \\ 0, & i > k \end{cases}$$

由于 $\sum_{i=1}^{\infty} \rho_i < \infty$, 故平均灭亡时间为

$$\sum_{i=1}^{\infty} \rho_i = \sum_{i=1}^k \frac{1}{\lambda_i} \left(\frac{\lambda}{\mu} \right)^i$$

20. 这是一个有限状态连续时间 Markov 链,

具有三个状态: 工作, 失灵-, 失灵+.

$$\text{无限矩阵 } A = \begin{pmatrix} -(\lambda_1 + \mu_2) & \lambda_1 & \lambda_2 \\ \mu_1 & -\mu_1 & 0 \\ \mu_2 & 0 & -\mu_2 \end{pmatrix}.$$

在 $P(0) = I$ 下有 $P(t) = e^{At}$. $P(t)$ 的 $(0,0)$ 元即为所求.

21. 排队长度和服务忙碌期相同, 因为显然与服务顺序无关.
等待时间不同, 因为显然取决于服务顺序.

22 设 $X(t)$ 为 t 时刻的队长
已知 $X(t)$ 是生灭过程, 参数为

$$\lambda_i = \rho\lambda, i \geq 0,$$

$$\mu_i = \begin{cases} \mu, i > 0, \\ 0, i = 0. \end{cases}$$

23 由题21知队长与服务顺序无关, 故 $X(t)$ 是生灭过程,

参数为 $\lambda_i = \lambda, i \geq 0,$

$$\mu_i = \begin{cases} \mu, i > 0, \\ 0, i = 0. \end{cases}$$

$$24. P_n(t) = e^{-\beta t} (1 - e^{-\beta t})^{n-1} \sim GE(e^{-\beta t}), n \geq 1.$$

我们知道对于 $X \sim GE(p), EX = \frac{1}{p}, \text{Var}(X) = \frac{1-p}{p^2}.$

$$\text{故 } E(X(t)) = e^{\beta t}, \text{Var}(X(t)) = e^{2\beta t} (1 - e^{-\beta t})$$