

DD2434 – Advanced Machine Learning

Lecture 6: Representation Learning

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Computational Science and Technology (CST) KTH Royal Institute of Technology

- Graphical models, basics
- Factor analysis

Short outline for today

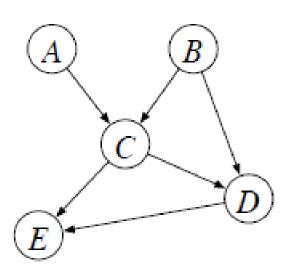
- 1. Basics of graphical models
- Factor analysis with Probabilistic Principal Component Analysis (PPCA)

- · Graphical models, basics
- Factor analysis

- Graphical models
- · Latent variable models

"Marriage between probability theory and graph theory"

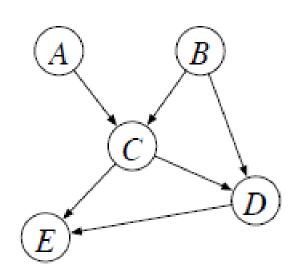
- tool for dealing with uncertainty, independence, and complexity
- > an intuitive way of representing and visualising the relationships between many variables.



- · Graphical models
- · Latent variable models

"Marriage between probability theory and graph theory"

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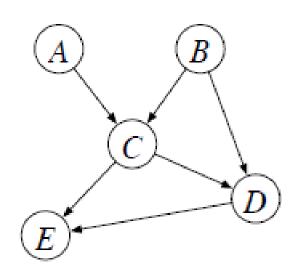
Focus on directed acyclic graphs (DAGs)

factorisation of the joint probability distribution

$$p(A, B, C, D, E) = p(A)p(B)p(C|A, B)p(D|B, C)p(E|C, D)$$

"Marriage between probability theory and graph theory"

- tool for dealing with uncertainty, independence, and complexity
- > an intuitive way of representing and visualising the relationships between many variables.



Focus on directed acyclic graphs (DAGs)

> factorisation of the joint probability distribution (associated with the whole graph)

$$p(X_1, ..., X_n) = \prod_{i=1}^n p(X_i | X_{pa(i)})$$

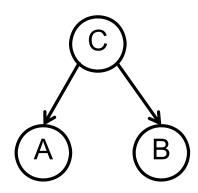
- Graphical models, basics
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Conditional independence relationships

way to abstract out such relationships between variables from the details of their parametric forms

"Is A dependent on B given that we know the value of C?"



- · Graphical models
- · Latent variable models

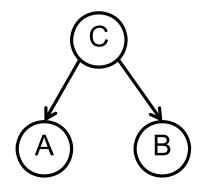
Conditional independence relationships

way to abstract out such relationships between variables from the details of their parametric forms

"Is A dependent on B given that we know the value of C?"

A and B are conditionally probabilistically independent given C if and only if

$$p(A,B|C) = p(A|C) * p(B|C)$$



No direct link between children nodes, A and B

Conditional independence relationships

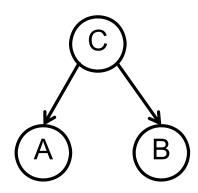
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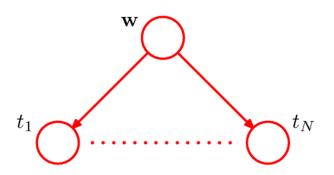
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$$p(A,B,C) = p(A|C) * p(B|C)*p(C)$$



- Graphical models, basics
- Factor analysis

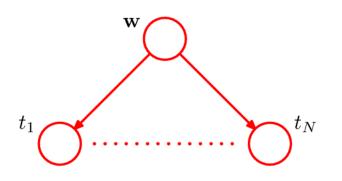
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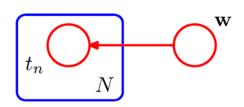


$$p(\mathbf{t}, \mathbf{w}) = p(\mathbf{w}) \prod_{n=1}^{N} p(t_n | \mathbf{w})$$

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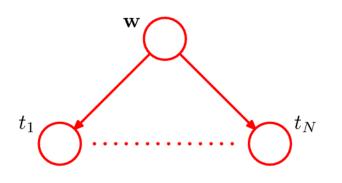
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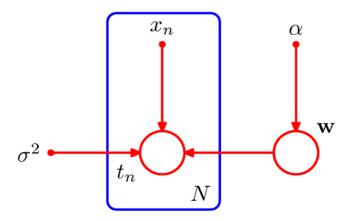




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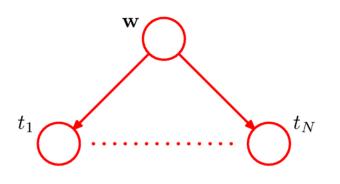
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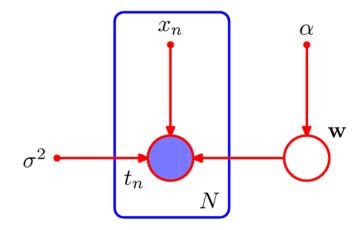




$$p(\mathbf{t}, \mathbf{w} | \mathbf{X}, \alpha, \sigma^2) = p(\mathbf{w} | \alpha) \prod_{n=1}^{N} p(t_n | \mathbf{w}, x_n, \sigma^2)$$

- Graphical models
- Latent variable models





$$p(\mathbf{t}, \mathbf{w} | \mathbf{X}, \alpha, \sigma^2) = p(\mathbf{w} | \alpha) \prod_{n=1}^{N} p(t_n | \mathbf{w}, x_n, \sigma^2)$$

Solid Blue Circle means observable values.

- · Graphical models, basics
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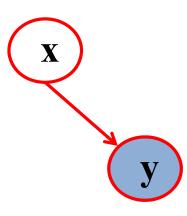
- observable data
- task to model p(y)



- Graphical models, basics
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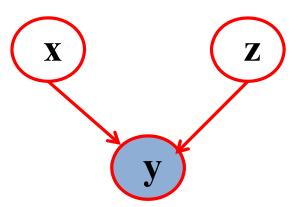
- observable data
- task to model p(y)
- unobservable: latent variables



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- observable data
- task to model p(y)
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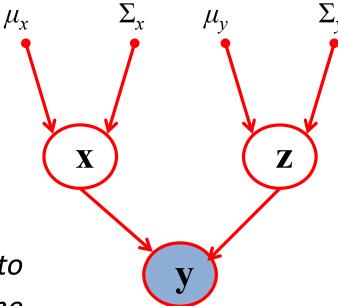
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- observable data
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- unobservable: *latent* variables

"The primary role of the **latent variables** is to allow a complicated distribution over the observed variables to be represented in terms of a model constructed from simpler (typically exponential family) conditional distributions."

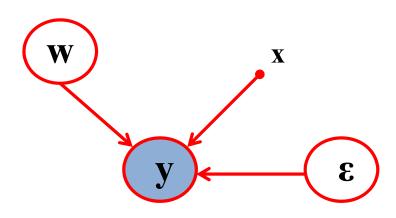
Bishop 2006



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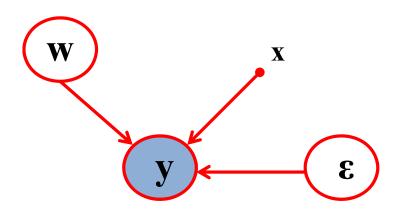


$$y_i = \mathbf{w}^{\mathrm{T}} \mathbf{x}_i + \varepsilon$$

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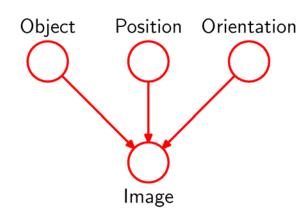
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- observable data
- task to model p(y)
- unobservable: *latent* variables
- "explaining away" phenomenon



$$y_i = \mathbf{w}^{\mathrm{T}} \mathbf{x}_i + \varepsilon$$

- observable data
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Bishop 2006

- generative models for conditional distributions being

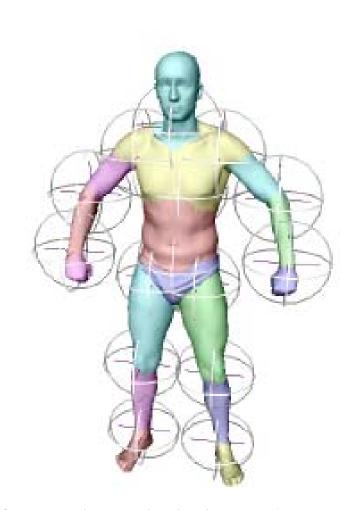
parameterised by a latent random variable - input: p(y|x)

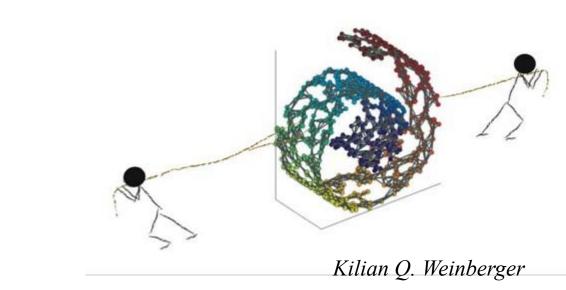
(the concept of ancestral sampling -> creation of the "fantasy" observed data)

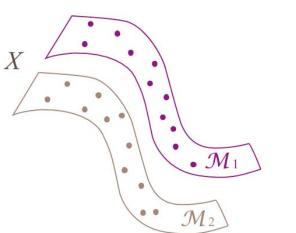
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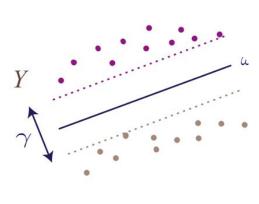
- · Data dimensionality, degrees of freedom
- Latent linear models
- Factor analysis general formulation
- PCA and other models

Data dimensionality









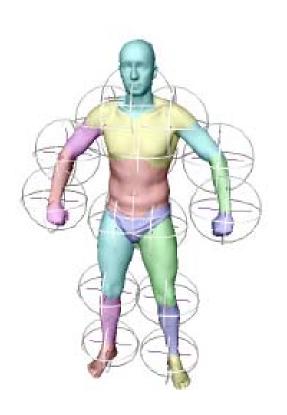
adapted from Carl Henrik Ek's lecture (DD2434, 2015)

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Intrinsic dimensionality

- Data dimensionality vs intrinsic dimensionality
 - fewer "data-driven" degrees of freedom
 - discovering lower-dimensional manifold
 - re-parameterisation of data
 - to understand/interpret relevant aspects of data
 - to boost generalisation
 - for computational efficiency



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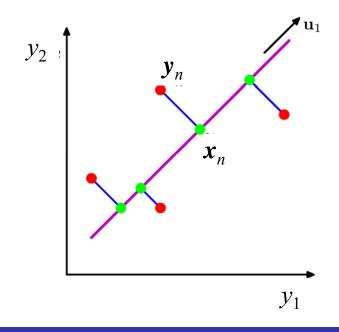
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Popular PCA

- Principal component analysis (PCA)
 - > Project D-dimensional data $\mathbf{Y} = \{\mathbf{y}_n\}$, on to the principal subspace $\mathbf{X} = \{\mathbf{x}_n\}$ of lower-dimensionality, M < D

$$X = Yu$$

Variance of the projected data should be maximised



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$$\max_{\mathbf{u}} \arg \left\{ \operatorname{var}(\mathbf{X}) : \mathbf{X} = \mathbf{Y}\mathbf{u} \right\} = \max_{\mathbf{u}} \arg \left\{ \left(\mathbf{Y}\mathbf{u}\right)^{\mathsf{T}} \mathbf{Y}\mathbf{u} \right\}, \quad \text{subject to } \mathbf{u}^{\mathsf{T}}\mathbf{u} = 1$$

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$$\mathbf{u}_{k}^{\mathrm{T}}\operatorname{cov}(\mathbf{Y})\ \mathbf{u}_{k}=\lambda_{k}$$

eigenvector formulation

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Data observed Y

$$p(\mathbf{Y})$$

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Data observed Y

$$p(\mathbf{Y})$$

• Assumption: $\mathbf{Y} \in \mathbb{R}^{N \times D}$ have been generated from $\mathbf{X} \in \mathbb{R}^{N \times M}$ by means of some generative mapping, \mathbf{f}

$$p(\mathbf{Y} | \mathbf{f}, \mathbf{X})$$
, where $\mathbf{f}: \mathbf{X} \to \mathbf{Y}$

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• X is then considered as *latent* variable

$$\mathbf{X}_n \xrightarrow{\mathrm{f}} \mathbf{y}_n$$

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Linear latent variable model

$$p(\mathbf{Y} | \mathbf{W}, \mathbf{X}) = \prod_{n=1}^{N} p(\mathbf{y}_n | \mathbf{W}, \mathbf{x}_n)$$

$$p(\mathbf{y} \mid \mathbf{W}, \mathbf{x}) = \mathcal{N}(\mathbf{W}\mathbf{x} + \boldsymbol{\mu}, \boldsymbol{\Psi})$$

- It looks like a regression problem but without inputs (we observe Y and want to infer X)
- We need to find suitable data representation
- Can we learn the underlying (generative) mapping?

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Linear latent variable model

$$p(\mathbf{Y} | \mathbf{W}, \mathbf{X}) = \prod_{n=1}^{N} p(\mathbf{y}_n | \mathbf{W}, \mathbf{x}_n)$$

The power of priors, $p(\mathbf{x})$

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- encoding prior belief
- preference (choice out of many options)
- regularisation of the solution space
- Can v

- Graphical models, basics
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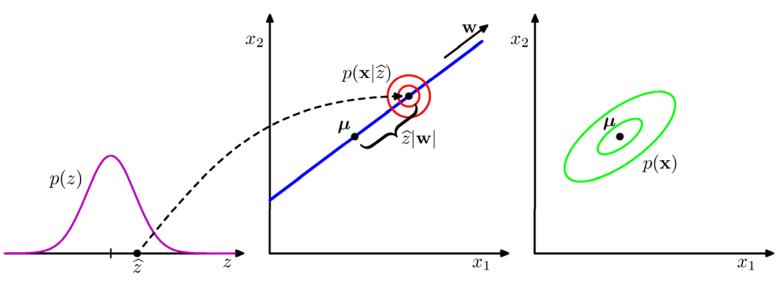
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Factor analysis – general formulation

Generative mapping from the latent to the observed variable

$$y = Wx + \mu + \varepsilon$$

$$\mathbf{y} \in \mathbb{R}^D, \ \mathbf{x} \in \mathbb{R}^M, \ \mathbf{W} \in \mathbb{R}^{D \times M}$$



$$p(x) = \mathcal{N}\left(x \mid \mu_0, \sigma_0\right)$$

Bishop 2006 (variable notation: z is x, x is y)

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factor loading matrix

Gaussian prior distribution over latent variables

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0)$$

Conditional distribution of the observed y conditioned on latent x

$$p(\mathbf{y} \mid \mathbf{W}, \mathbf{x}) = \mathcal{N}(\mathbf{W}\mathbf{x} + \boldsymbol{\mu}, \boldsymbol{\Psi})$$

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Generative mapping from the latent to the observed variable

$$y = Wx + \mu + \varepsilon$$

 $\mathbf{y} \in \mathbb{R}^D$, $\mathbf{x} \in \mathbb{R}^M$, $\mathbf{W} \in \mathbb{R}^{D \times M}$

Gaussian

- The generative mapping is linear
- lacktriangle Assumption: Ψ is diagonal (to explain the correlation)
- Looks like regression but we have no inputs so there is need to specify prior

Condition

 $p(\mathbf{y} | \mathbf{W}, \mathbf{x}) = \mathcal{N}(\mathbf{W}\mathbf{x} + \boldsymbol{\mu}, \boldsymbol{\Psi})$

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Marginal distribution

$$p(\mathbf{y}_{n} \mid \boldsymbol{\theta}) = \int p(\mathbf{y}_{n} \mid \mathbf{x}_{n}, \boldsymbol{\theta}) p(\mathbf{x}_{n}) d\mathbf{x}_{n} = \int \mathcal{N}(\mathbf{W}\mathbf{x}_{n} + \boldsymbol{\mu}, \boldsymbol{\Psi}) \mathcal{N}(\boldsymbol{\mu}_{0}, \boldsymbol{\Sigma}_{0}) d\mathbf{x}_{n}$$
conditional prior distr.

Murphy (2002), ch.12

adapted from Carl Henrik Ek's lecture (DD2434, 2015)

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Marginal distribution

distr.

$$p(\mathbf{y}_n \mid \boldsymbol{\theta}) = \int p(\mathbf{y}_n \mid \mathbf{x}_n, \boldsymbol{\theta}) p(\mathbf{x}_n) d\mathbf{x}_n = \int \mathcal{N}(\mathbf{W} \mathbf{x}_n + \boldsymbol{\mu}, \boldsymbol{\Psi}) \, \mathcal{N}(\boldsymbol{\mu}_0, \boldsymbol{\Sigma}_0) d\mathbf{x}_n$$
conditional prior

- X and W are related
- here, we integrate out X
 (marginalisation)

Murphy (2002), ch.12

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conditional prior distr.

$$p(\mathbf{y}_i \mid \boldsymbol{\theta}) = \mathcal{N} (\mathbf{W} \boldsymbol{\mu}_0 + \boldsymbol{\mu}, \boldsymbol{\Psi} + \mathbf{W} \boldsymbol{\Sigma}_0 \mathbf{W}^{\mathrm{T}})$$

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• if $\Sigma_0 = \mathbf{I}$ and $\mu_0 = \mathbf{0}$, i.e. $p(\mathbf{x}) = \mathcal{N}(0, \mathbf{I})$, then $p(\mathbf{y}_i) = \mathcal{N}(\mu, \Psi + \mathbf{W}\mathbf{W}^T)$

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Factor analysis – parameter learning

Marginal distribution

$$p(\mathbf{y}_{n} \mid \boldsymbol{\theta}) = \int p(\mathbf{y}_{n} \mid \mathbf{x}_{n}, \boldsymbol{\theta}) p(\mathbf{x}_{n}) d\mathbf{x}_{n} = \int \mathcal{N}(\mathbf{W}\mathbf{x}_{n} + \boldsymbol{\mu}, \boldsymbol{\Psi}) \, \mathcal{N}(\boldsymbol{\mu}_{0}, \boldsymbol{\Sigma}_{0}) d\mathbf{x}_{n}$$
conditional prior distr.

$$p(\mathbf{y}_i \mid \boldsymbol{\theta}) = \mathcal{N} (\mathbf{W} \boldsymbol{\mu}_0 + \boldsymbol{\mu}, \boldsymbol{\Psi} + \mathbf{W} \boldsymbol{\Sigma}_0 \mathbf{W}^{\mathrm{T}})$$

• if
$$\Sigma_0 = \mathbf{I}$$
 and $\mu_0 = \mathbf{0}$, then $p(\mathbf{y}_i) = \mathcal{N}(\mu, \Psi + \mathbf{W}\mathbf{W}^T)$ rank of dim \mathbf{X}

A low rank parameterisation (density model) of the original ${f Y}$

- · Graphical models, basics
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Factor analysis – unidentifiability

If we try to rotate our weight matrix W by an orthogonal R:

$$\tilde{\mathbf{W}} = \mathbf{W}\mathbf{R}$$

$$p(\mathbf{y}_n \mid \boldsymbol{\theta}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Psi} + \mathbf{W}\mathbf{R}\mathbf{R}^{\mathrm{T}}\mathbf{W}^{\mathrm{T}})$$

$$= \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Psi} + \mathbf{W}\mathbf{W}^{\mathrm{T}})$$

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$$= \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\Psi} + \mathbf{W}\mathbf{W}^{\mathrm{T}})$$

The marginal likelihood is invariant to a rotation

- \succ no unique solution for $\mathbf{W} o$ we cannot uniquely identify latent factors
- > interpretation becomes somewhat tricky
- > it is possible to impose extra constraints on W

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Probabilistic PCA (PPCA)

Probabilistic formulation of PCA stems from factor analysis, if

- 1) Ψ is isotropic, i.e. $\Psi = \sigma^2 \mathbf{I}$ (the same noise variances for each variable)
- 2) factor loadings are orthogonal, i.e. $W^TW=I$

$$p(\mathbf{y} \mid \boldsymbol{\theta}) = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\sigma}^2 \mathbf{I} + \mathbf{W} \mathbf{W}^{\mathrm{T}})$$

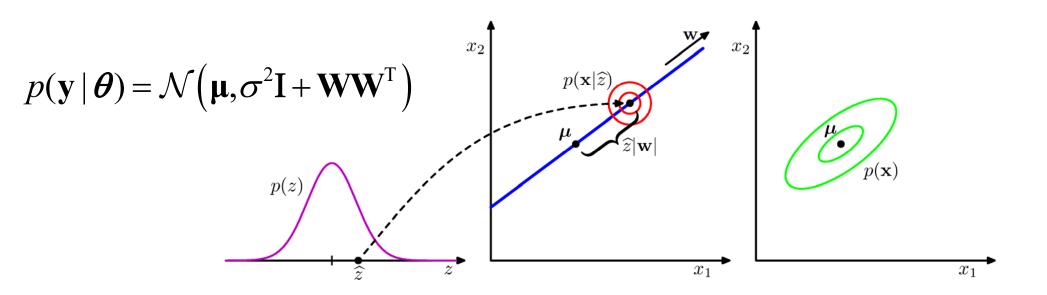
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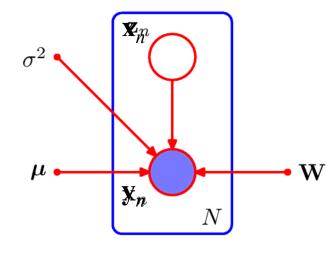
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We first obtained marginal distribution by integrating out latent variable $oldsymbol{x}$

$$p(\mathbf{y}_n \mid \boldsymbol{\theta}) = \int p(\mathbf{y}_n \mid \mathbf{x}_n, \boldsymbol{\theta}) p(\mathbf{x}_n) d\mathbf{x}_n = \mathcal{N}(\boldsymbol{\mu}, \boldsymbol{\sigma}^2 \mathbf{I} + \mathbf{W} \mathbf{W}^{\mathrm{T}})$$



Bishop 2006

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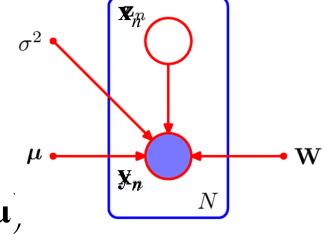
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$$p(\mathbf{y}_n \mid \boldsymbol{\theta}) = \int p(\mathbf{y}_n \mid \mathbf{x}_n, \boldsymbol{\theta}) p(\mathbf{x}_n) d\mathbf{x}_n = \mathcal{N}(\boldsymbol{\mu}, \underline{\sigma^2 \mathbf{I} + \mathbf{W} \mathbf{W}^{\mathrm{T}}})$$

Log-likelihood:

$$\log p(\mathbf{Y}|\mathbf{\mu}, \mathbf{W}, \sigma^2) = \sum_{n=1}^{N} \log p(\mathbf{y}_n | \mathbf{\mu}, \mathbf{W}, \sigma^2) =$$

$$= -\frac{ND}{2} \log(2\pi) - \frac{N}{2} \log|\mathbf{C}| - \frac{1}{2} \sum_{n=1}^{N} (\mathbf{y}_n - \mathbf{\mu})^{\mathrm{T}} \mathbf{C}^{-1} (\mathbf{y}_n - \mathbf{\mu})$$



Bishop 2006

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- Factor analysis

- Data dimensionality, degrees of freedom
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- Factor analysis general formulation
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Now we can select parameters based on the maximum likelihood principle: $\mathbf{W}_{\text{ML}} = \underset{\mathbf{W}}{\operatorname{arg\,max}} \left\{ \log p(\mathbf{Y} | \boldsymbol{\mu}, \mathbf{W}, \sigma^2) \right\}$

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$$\mathbf{W}_{\mathrm{ML}} = \mathbf{U}_{M} (\mathbf{\Lambda}_{M} - \boldsymbol{\sigma}^{2} \mathbf{I})^{1/2}$$

$$\boldsymbol{\mu} = \overline{\mathbf{y}}$$

$$\boldsymbol{\sigma}_{\mathrm{ML}}^{2} = \frac{1}{D - M} \sum_{i=M+1}^{D} \lambda_{i}$$

$$\mathbf{U}_{\scriptscriptstyle M}:D\times M,\quad \mathbf{\Lambda}_{\scriptscriptstyle M}:M\times M$$

Tipping and Bishop, 1999

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Interpretations

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- Columns of \mathbf{W} are PC eigenvec scaled by the variances $\lambda_i \sigma^2$
- what is the variance in the direction v (unit vector), v^TCv ?

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- Columns of \mathbf{W} are PC eigenvec scaled by the variances $\lambda_i \sigma^2$
- what is the variance in the direction v (unit vector), v^TCv ?
 - if v is orthogonal to the principal space, then $v^TU=0$ and $v^TCv=\sigma^2$
 - if $\mathbf{v} = \mathbf{u}_i$ then $\mathbf{v}^T \mathbf{C} \mathbf{v}$ (variance) is $\lambda_i \sigma^2 + \sigma^2 = \lambda_i$.

Bishop, 2006

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Posterior distribution ("reverse" mapping) can be obtained by applying Bayes' theorem and making calculus on Gaussian distr.

It is easier to make inverse if $\Psi = \sigma^2 I$ (PPCA):

$$p(\mathbf{x} \mid \mathbf{y}) = \mathcal{N}\left(\left(\mathbf{W}^{\mathsf{T}}\mathbf{W} + \sigma^{2}\mathbf{I}\right)^{-1}\mathbf{W}^{\mathsf{T}}\left(\mathbf{y} - \boldsymbol{\mu}\right), \sigma^{2}\left(\mathbf{W}^{\mathsf{T}}\mathbf{W} + \sigma^{2}\mathbf{I}\right)^{-1}\right)$$

posterior mean

posterior covariance

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So, the point targeted (projected to) in data from latent space is

$$\mathbf{W}\mathbb{E}[\mathbf{x} | \mathbf{y}] + \mathbf{\mu}$$

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For $\mathbf{W}=\mathbf{W}_{\mathrm{ML}}$ and in the limit $\sigma^2 \to 0$, the posterior mean becomes

$$\left(\mathbf{W}_{\mathrm{ML}}^{\mathrm{T}}\mathbf{W}_{\mathrm{ML}}\right)^{-1}\mathbf{W}_{\mathrm{ML}}^{\mathrm{T}}\left(\mathbf{y}-\overline{\mathbf{y}}\right)$$

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$$\left(\mathbf{W}_{\mathrm{ML}}^{\mathrm{T}}\mathbf{W}_{\mathrm{ML}}\right)^{-1}\mathbf{W}_{\mathrm{ML}}^{\mathrm{T}}\left(\mathbf{y}-\overline{\mathbf{y}}\right) \to \mathbf{W}_{\mathrm{ML}}^{\mathrm{T}}\left(\mathbf{y}-\overline{\mathbf{y}}\right)$$

if ${f W}$ is orthogonal

..... and the posterior covariance = 0.

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Posterior distribution ("reverse" mapping) can be obtained by applying Bayes' theorem and making calculus on Gaussian distr.

It is easier to make inverse if $\Psi = \sigma^2 I$ (PPCA):

• For $\sigma^2 \to 0$ we recover the *classical PCA* formulation!

• For $\sigma^2 > 0$ we obtain PPCA (sensible PCA)

the posterior mean is no longer an orthogonal projection

it is shrunk towards the prior mean

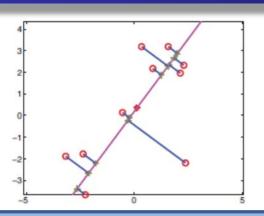
the reconstruction is closer to the overall data mean $\overline{y} = \mu$

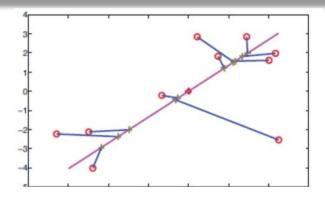
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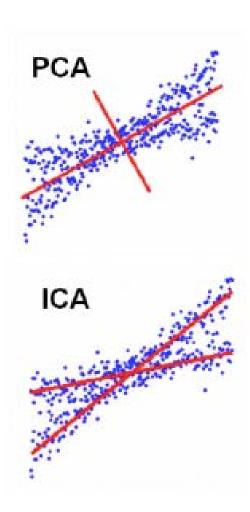
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Independent component analysis (ICA)

- Cocktail party problem (blind source separation)
- Factorisation of the latent variables distribution

$$p(\mathbf{x}) = \prod_{j=1}^{M} p(x_j)$$

- Non-gaussian distributions of different sources
- The prior determines different models



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Summary

- Graphical models an intuitive way of representing and visualising the relationships between many variables
- Implicit representation and degrees of freedom
- Generative models
- Priors as preference
- Factor analysis is a linear continuous latent variable model
- PPCA is a factor analysis with special conditions: isotropic covariance matrix and orthogonality of W
- PPCA reduces to classical PCA formulation with $\sigma^2 \rightarrow 0$

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Recommended additional reading

On PPCA

M.E. Tipping and C.M. Bishop. (1999) Probabilistic principal component analysis.

Journal of the Royal Statistical Society, Series B 21(3), 611–622.6.

K.P. Murphy (2012) Machine Learning: Probabilistic Perspective. Chapter 12. MIT Press.

On GP-LVMs

Neil D Lawrence. (2005) Probabilistic non-linear principal component analysis with Gaussian process latent variable models". *The Journal of Machine Learning Research* 6