

Q.1

Posterior distribution

$$\underline{p(w | X, t)}$$

$$t = y + \varepsilon = Xw + \varepsilon$$

$$X = \begin{bmatrix} x_1^{(1)} & \dots & x_1^{(D)} \\ x_2^{(1)} & \dots & x_2^{(D)} \\ \vdots & & \vdots \\ x_N^{(1)} & \dots & x_N^{(D)} \end{bmatrix} \Rightarrow \begin{bmatrix} x_1^T \\ x_2^T \\ \vdots \\ x_N^T \end{bmatrix} \quad X: N \times D$$

$$x_i = [x_i^{(1)}, \dots, x_i^{(D)}]^T : D\text{-dimensional input}$$

$$w = [w^{(1)}, \dots, w^{(D)}]^T : D\text{-dimensional parameter vector}$$

$$t = [t_1, \dots, t_N]^T : \begin{array}{l} N \text{ data samples (here: output),} \\ \text{each output } t_n \text{ is one-dimensional,} \\ N \text{ samples of one-dim output} \\ \text{forms a } N\text{-dimensional column vector} \end{array}$$

We expect our posterior to follow a Gaussian distribution:

$$p(w | X, t) = \mathcal{N}(\mu_w, \Sigma_w) \quad \begin{array}{l} \mu_w - \text{mean} \\ \Sigma_w - \text{covariance} \\ \text{of the } w \text{ posterior} \end{array}$$

$$\frac{1}{(2\pi)^{\frac{D}{2}} |\Sigma|^{1/2}} p(w | X, t) \sim e^{-\frac{1}{2}(w - \mu_w)^T \Sigma_w^{-1} (w - \mu_w)} =$$

$$= \underbrace{e^{-\frac{1}{2} w^T \Sigma_w^{-1} w}}_{\text{quadratic}} \cdot \underbrace{e^{w^T \Sigma_w^{-1} \mu_w}}_{\text{linear term}} \cdot \underbrace{e^{-\frac{1}{2} \mu_w^T \Sigma_w^{-1} \mu_w}}_{\text{const.}}$$

We are going to "complete the square" soon to identify μ_w and Σ_w .

(p. 2)

Likelihood

$$\underline{p(t|w, X)}$$

$$t = y + \varepsilon = Xw + \varepsilon$$

$$\underline{p(t|w, X) = \mathcal{N}(Xw, \sigma^2 I_N)}$$

N-dimensional Gaussian distr.
with a diagonal covariance
(conditional independence of t_n)

Prior over w

$$p(w) = \mathcal{N}(0, \Sigma^{-1})$$

So, posterior $p(w|t, X) \propto p(t|w, X) \cdot p(w)$

\Downarrow

$$p(w|t, X) \propto e^{-\frac{1}{2\sigma^2}(t-Xw)^T(t-Xw)} \cdot e^{-\frac{1}{2}w^T \Sigma^{-1} w}$$

If we focus now on the exponent:

$$-\frac{1}{2\sigma^2}(t-Xw)^T(t-Xw) - \frac{1}{2}w^T \Sigma^{-1} w =$$

$$= -\frac{1}{2\sigma^2} t^T t + \frac{1}{\sigma^2} t^T Xw - \frac{1}{2\sigma^2} (Xw)^T (Xw) - \frac{1}{2} w^T \Sigma^{-1} w$$

$\underbrace{\hspace{10em}}$
const. (from w perspective)

p.3

We have then 3 terms:

$$\textcircled{\text{I}} \quad -\frac{1}{2\sigma^2} t^T t \rightarrow \text{const. for } w$$

$$\textcircled{\text{II}} \quad \frac{1}{\sigma^2} t^T X w = \frac{1}{\sigma^2} w^T X^T t \rightarrow \text{term linear in } w$$

$$\textcircled{\text{III}} \quad -\frac{1}{2\sigma^2} (Xw)^T (Xw) - \frac{1}{2} w^T \Sigma_w^{-1} w =$$

$$= -\frac{1}{2\sigma^2} w^T X^T X w - \frac{1}{2} w^T \Sigma_w^{-1} w =$$

$$= -\frac{1}{2} w^T \left(\frac{1}{\sigma^2} X^T X \right) w - \frac{1}{2} w^T \Sigma_w^{-1} w =$$

$$= -\frac{1}{2} w^T \left(\frac{1}{\sigma^2} X^T X + \Sigma_w^{-1} \right) w$$

quadratic term in w

If we look back at our posterior $p(w|t, X)$
we can identify corresponding terms.

Let's complete the square:

$$\textcircled{\text{I}} \quad \text{let's leave out corresponding const. terms } -\frac{1}{2\sigma^2} t^T t$$

$$\textcircled{\text{III}} \quad \text{quadratic terms:}$$

$$-\frac{1}{2} w^T \left(\frac{1}{\sigma^2} X^T X + \Sigma_w^{-1} \right) w$$

vs.

$$-\frac{1}{2} w^T \Sigma_w^{-1} w - \frac{1}{2} w^T \Sigma_w^{-1} w$$

$$\text{So, } \underline{\underline{\Sigma_w^{-1} = \frac{1}{\sigma^2} X^T X + \Sigma_w^{-1}}}$$

p.4

II

linear terms

$$\frac{1}{\sigma^2} \mathbf{w}^T \mathbf{X}^T \mathbf{t} \quad \text{vs.} \quad \mathbf{w}^T \Sigma_w^{-1} \mu_w$$

$$\Downarrow$$
$$\mathbf{w}^T \Sigma_w^{-1} \mu_w = \mathbf{w}^T \left(\frac{1}{\sigma^2} \mathbf{X}^T \mathbf{X} + \Sigma^{-1} \right) \mu_w$$

$$\Downarrow$$
$$\cancel{\mathbf{w}^T} \left(\frac{1}{\sigma^2} \mathbf{X}^T \mathbf{X} + \Sigma^{-1} \right) \mu_w = \frac{1}{\sigma^2} \cancel{\mathbf{w}^T} \mathbf{X}^T \mathbf{t}$$

$$\underline{\underline{\mu_w = \frac{1}{\sigma^2} \left(\frac{1}{\sigma^2} \mathbf{X}^T \mathbf{X} + \Sigma^{-1} \right)^{-1} \mathbf{X}^T \mathbf{t}}}}$$

So, let's summarise the posterior

$$p(\mathbf{w} | \mathbf{t}, \mathbf{X}) = \mathcal{N} \left(\frac{1}{\sigma^2} \left(\frac{1}{\sigma^2} \mathbf{X}^T \mathbf{X} + \Sigma^{-1} \right)^{-1} \mathbf{X}^T \mathbf{t}, \frac{1}{\sigma^2} \mathbf{X}^T \mathbf{X} + \Sigma^{-1} \right)$$

σ comes from the likelihood

Σ comes from the prior (~~matrix~~ covariance)