Deep Generative Models Lecture 2

Roman Isachenko



Ozon Masters

Spring, 2021

Recap of previous lecture

We are given i.i.d. samples $\{\mathbf{x}_i\}_{i=1}^n \in X$ (e.g. $X = \mathbb{R}^m$) from unknown distribution $\pi(\mathbf{x})$.

Goal

We would like to learn a distribution $\pi(\mathbf{x})$ for

- evaluating $\pi(\mathbf{x})$ for new samples (how likely to get object \mathbf{x} ?);
- ▶ sampling from $\pi(\mathbf{x})$ (to get new objects $\mathbf{x} \sim \pi(\mathbf{x})$).

Challenge

Data is complex and high-dimensional.

MLE problem

Fix probabilistic model $p(\mathbf{x}|\boldsymbol{\theta})$ – a set of parameterized distributions, such that $p(\mathbf{x}|\boldsymbol{\theta}) \approx \pi(\mathbf{x})$.

$$m{ heta}^* = rg \max_{m{ heta}} p(\mathbf{X}|m{ heta}) = rg \max_{m{ heta}} \prod_{i=1}^n p(\mathbf{x}_i|m{ heta}) = rg \max_{m{ heta}} \sum_{i=1}^n \log p(\mathbf{x}_i|m{ heta}).$$

Recap of previous lecture

Likelihood as product of conditionals

Let
$$\mathbf{x} = (x_1, \dots, x_m)$$
, $\mathbf{x}_{1:i} = (x_1, \dots, x_i)$. Then

$$p(\mathbf{x}|\boldsymbol{\theta}) = \prod_{i=1}^{m} p(x_i|\mathbf{x}_{1:i-1}, \boldsymbol{\theta}); \quad \log p(\mathbf{x}|\boldsymbol{\theta}) = \sum_{i=1}^{m} \log p(x_i|\mathbf{x}_{1:i-1}, \boldsymbol{\theta}).$$

MLE problem for autoregressive model

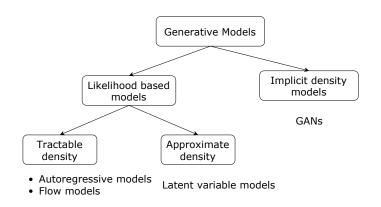
$$\theta^* = \underset{\theta}{\operatorname{arg max}} p(\mathbf{X}|\theta) = \underset{\theta}{\operatorname{arg max}} \sum_{i=1}^n \sum_{j=1}^m \log p(x_{ij}|\mathbf{x}_{i,1:j-1}\theta).$$

Sampling

$$\hat{x}_1 \sim p(x_1|\boldsymbol{\theta}), \quad \hat{x}_2 \sim p(x_2|\hat{x}_1,\boldsymbol{\theta}), \ldots, \quad \hat{x}_m \sim p(x_m|\hat{\mathbf{x}}_{1:m-1},\boldsymbol{\theta})$$

New generated object is $\hat{\mathbf{x}} = (\hat{x}_1, \hat{x}_2, \dots, \hat{x}_m)$.

Generative models zoo



Bayes theorem

$$p(\mathbf{t}|\mathbf{x}) = \frac{p(\mathbf{x}|\mathbf{t})p(\mathbf{t})}{p(\mathbf{x})} = \frac{p(\mathbf{x}|\mathbf{t})p(\mathbf{t})}{\int p(\mathbf{x}|\mathbf{t})p(\mathbf{t})d\mathbf{t}}$$

- x − observed variables, t − unobserved variables (latent variables/parameters);
- $p(\mathbf{x}|\mathbf{t})$ likelihood;
- $p(\mathbf{x}) = \int p(\mathbf{x}|\mathbf{t})p(\mathbf{t})d\mathbf{t}$ evidence;
- ▶ $p(\mathbf{t})$ prior distribution, $p(\mathbf{t}|\mathbf{x})$ posterior distribution.

Meaning

We have unobserved variables \mathbf{t} and some prior knowledge about them $p(\mathbf{t})$. Then, the data \mathbf{x} has been observed. Posterior distribution $p(\mathbf{t}|\mathbf{x})$ summarizes the knowledge after the obbservations.

Let consider the case, where the unobserved variables ${\bf t}$ is our model parameters ${m heta}.$

- $\mathbf{X} = \{\mathbf{x}_i\}_{i=1}^n$ observed samples;
- ▶ $p(\theta)$ prior parameters distribution (we treat model parameters θ as random variables).

Posterior distribution

$$p(\theta|\mathbf{X}) = \frac{p(\mathbf{X}|\theta)p(\theta)}{p(\mathbf{X})} = \frac{p(\mathbf{X}|\theta)p(\theta)}{\int p(\mathbf{X}|\theta)p(\theta)d\theta}$$

Bayesian inference

$$p(\mathbf{x}|\mathbf{X}) = \int p(\mathbf{x}|\theta)p(\theta|\mathbf{X})d\theta$$

Note the difference from

$$p(\mathbf{x}) = \int p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})d\boldsymbol{\theta}.$$

Posterior distribution

$$p(\boldsymbol{\theta}|\mathbf{X}) = \frac{p(\mathbf{X}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{p(\mathbf{X})} = \frac{p(\mathbf{X}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{\int p(\mathbf{X}|\boldsymbol{\theta})p(\boldsymbol{\theta})d\boldsymbol{\theta}}$$

Bayesian inference

$$p(\mathbf{x}|\mathbf{X}) = \int p(\mathbf{x}|\theta)p(\theta|\mathbf{X})d\theta$$

If evidence $p(\mathbf{X})$ is intractable (due to multidimensional integration), we can't get posterior distribution and perform the precise inference.

Maximum a posteriori (MAP) estimation

$$\boldsymbol{\theta}^* = \argmax_{\boldsymbol{\theta}} p(\boldsymbol{\theta}|\mathbf{X}) = \argmax_{\boldsymbol{\theta}} \left(\log p(\mathbf{X}|\boldsymbol{\theta}) + \log p(\boldsymbol{\theta})\right)$$

MAP estimation

$$\boldsymbol{\theta}^* = \argmax_{\boldsymbol{\theta}} p(\boldsymbol{\theta}|\mathbf{X}) = \argmax_{\boldsymbol{\theta}} \left(\log p(\mathbf{X}|\boldsymbol{\theta}) + \log p(\boldsymbol{\theta})\right)$$

Estimated θ^* is a deterministic variable, but we could treat it as a random variable with density $p(\theta|\mathbf{X}) = \delta(\theta - \theta^*)$.

Dirac delta function

$$\delta(x) = \begin{cases} +\infty, & x = 0; \\ 0, & x \neq 0; \end{cases} \qquad \int f(x)\delta(x-y)dx = f(y).$$

MAP inference

$$p(\mathbf{x}|\mathbf{X}) = \int p(\mathbf{x}|\theta)p(\theta|\mathbf{X})d\theta \approx p(\mathbf{x}|\theta^*).$$

Latent variable models (LVM)

MLE problem

$$m{ heta}^* = rg \max_{m{ heta}} p(\mathbf{X}|m{ heta}) = rg \max_{m{ heta}} \prod_{i=1}^n p(\mathbf{x}_i|m{ heta}) = rg \max_{m{ heta}} \sum_{i=1}^n \log p(\mathbf{x}_i|m{ heta}).$$

Challenge

 $p(\mathbf{x}|\boldsymbol{\theta})$ could be intractable.

Extend probabilistic model

Introduce latent variable z for each sample x

$$p(\mathbf{x}, \mathbf{z}|\theta) = p(\mathbf{x}|\mathbf{z}, \theta)p(\mathbf{z}); \quad \log p(\mathbf{x}, \mathbf{z}|\theta) = \log p(\mathbf{x}|\mathbf{z}, \theta) + \log p(\mathbf{z}).$$

$$p(\mathbf{x}|\boldsymbol{\theta}) = \int p(\mathbf{x}, \mathbf{z}|\boldsymbol{\theta}) d\mathbf{z} = \int p(\mathbf{x}|\mathbf{z}, \boldsymbol{\theta}) p(\mathbf{z}) d\mathbf{z}.$$

Motivation

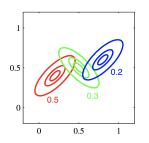
The distributions $p(\mathbf{x}|\mathbf{z}, \theta)$ and $p(\mathbf{z})$ could be quite simple.

Latent variable models (LVM)

$$\log p(\mathbf{x}|\boldsymbol{ heta}) = \log \int p(\mathbf{x}|\mathbf{z}, \boldsymbol{ heta}) p(\mathbf{z}) d\mathbf{z} o \max_{oldsymbol{ heta}}$$

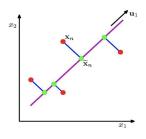
Examples

Mixture of gaussians



- $ho(\mathbf{x}|\mathbf{z}, oldsymbol{ heta}) = \mathcal{N}(\mathbf{x}|oldsymbol{\mu}_{\mathbf{z}}, oldsymbol{\Sigma}_{\mathbf{z}})$
- $ho(z) = Categorical(z|\pi)$

PCA model

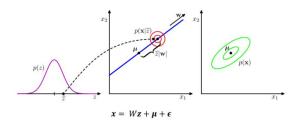


- $\qquad \qquad \rho(\mathsf{x}|\mathsf{z},\theta) = \mathcal{N}(\mathsf{x}|\mathsf{Wz} + \boldsymbol{\mu},\boldsymbol{\Sigma}_{\mathsf{z}})$
- $p(\mathbf{z}) = \mathcal{N}(\mathbf{z}|0,\mathbf{I})$

Latent variable models (LVM)

$$\log p(\mathbf{x}|oldsymbol{ heta}) = \log \int p(\mathbf{x}|\mathbf{z},oldsymbol{ heta}) p(\mathbf{z}) d\mathbf{z}
ightarrow \max_{oldsymbol{ heta}}$$

PCA goal: Project original data **X** onto a low dimensional latent space while maximizing the variance of the projected data.



- $p(\mathbf{z}) = \mathcal{N}(\mathbf{z}|0,\mathbf{I})$

Incomplete likelihood

MLE

$$egin{aligned} oldsymbol{ heta}^* &= rg\max_{oldsymbol{ heta}} p(\mathbf{X}, \mathbf{Z} | oldsymbol{ heta}) = rg\max_{oldsymbol{ heta}} \sum_{i=1}^n \log p(\mathbf{x}_i, \mathbf{z}_i | oldsymbol{ heta}). \end{aligned}$$

Since **Z** is unknown, maximize **incomplete likelihood**.

MILE problem

$$\begin{split} \boldsymbol{\theta}^* &= \arg\max_{\boldsymbol{\theta}} \log p(\mathbf{X}|\boldsymbol{\theta}) = \arg\max_{\boldsymbol{\theta}} \sum_{i=1}^n \log p(\mathbf{x}_i|\boldsymbol{\theta}) = \\ &= \arg\max_{\boldsymbol{\theta}} \sum_{i=1}^n \log \int p(\mathbf{x}_i, \mathbf{z}_i|\boldsymbol{\theta}) d\mathbf{z}_i = \\ &= \arg\max_{\boldsymbol{\theta}} \log \sum_{i=1}^n \int p(\mathbf{x}_i|\mathbf{z}_i, \boldsymbol{\theta}) p(\mathbf{z}_i) d\mathbf{z}_i. \end{split}$$

Variational lower bound

$$\log p(\mathbf{x}|\theta) = \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{p(\mathbf{z}|\mathbf{x}, \theta)} =$$

$$= \int q(\mathbf{z}) \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{p(\mathbf{z}|\mathbf{x}, \theta)} d\mathbf{z} = \int q(\mathbf{z}) \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)q(\mathbf{z})}{p(\mathbf{z}|\mathbf{x}, \theta)q(\mathbf{z})} d\mathbf{z} =$$

$$= \int q(\mathbf{z}) \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{q(\mathbf{z})} d\mathbf{z} + \int q(\mathbf{z}) \log \frac{q(\mathbf{z})}{p(\mathbf{z}|\mathbf{x}, \theta)} d\mathbf{z} =$$

$$= \mathcal{L}(q, \theta) + KL(q(\mathbf{z})||p(\mathbf{z}|\mathbf{x}, \theta)) \ge \mathcal{L}(q, \theta).$$

Kullback-Leibler divergence

- $KL(q||p) = \int q(\mathbf{z}) \log \frac{q(\mathbf{z})}{p(\mathbf{z})} d\mathbf{z};$
- $KL(q||p) \geq 0$;
- $\mathsf{KL}(q||p) = 0 \Leftrightarrow q \equiv p.$

Variational lower bound

$$\log p(\mathbf{x}|\boldsymbol{ heta}) = \mathcal{L}(q, \boldsymbol{ heta}) + \mathit{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{x}, \boldsymbol{ heta})) \geq \mathcal{L}(q, \boldsymbol{ heta}).$$

Evidence Lower Bound (ELBO)

$$\mathcal{L}(q, \theta) = \int q(\mathbf{z}) \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{q(\mathbf{z})} d\mathbf{z} =$$

$$= \int q(\mathbf{z}) \log p(\mathbf{x}|\mathbf{z}, \theta) d\mathbf{z} + \int q(\mathbf{z}) \log \frac{p(\mathbf{z})}{q(\mathbf{z})} d\mathbf{z}$$

$$= \mathbb{E}_q \log p(\mathbf{x}|\mathbf{z}, \theta) - KL(q(\mathbf{z})||p(\mathbf{z}))$$

Instead of maximizing incomplete likelihood, maximize ELBO (equivalently minimize KL)

$$\max_{m{ heta}} p(\mathbf{x}|m{ heta}) \quad o \quad \max_{q,m{ heta}} \mathcal{L}(q,m{ heta}) \equiv \min_{q,m{ heta}} \mathit{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{x},m{ heta})).$$

EM-algorithm

$$\mathcal{L}(q, \theta) = \int q(\mathbf{z}) \log p(\mathbf{x}|\mathbf{z}, \theta) d\mathbf{z} + \int q(\mathbf{z}) \log \frac{p(\mathbf{z})}{q(\mathbf{z})} d\mathbf{z}.$$

Block-coordinate optimization

- Initialize θ*;
- E-step

$$q(\mathbf{z}) = \underset{q}{\operatorname{arg max}} \mathcal{L}(q, \boldsymbol{\theta}^*) = \underset{q}{\operatorname{arg min}} \mathit{KL}(q||p) = p(\mathbf{z}|\mathbf{x}, \boldsymbol{\theta}^*);$$

M-step

$$oldsymbol{ heta}^* = rg \max_{oldsymbol{ heta}} \mathcal{L}(oldsymbol{q}, oldsymbol{ heta});$$

Repeat E-step and M-step until convergence.

Ugly pic

Amortized variational inference

E-step

$$q(\mathbf{z}) = rg \max_{q} \mathcal{L}(q, \boldsymbol{\theta}^*) = rg \min_{q} \mathit{KL}(q||p) = p(\mathbf{z}|\mathbf{x}, \boldsymbol{\theta}^*).$$

- \triangleright $p(\mathbf{z}|\mathbf{x}, \boldsymbol{\theta}^*)$ could be intractable;
- $ightharpoonup q(\mathbf{z})$ is different for each object \mathbf{x} .

Idea

Restrict a family of all possible distributions $q(\mathbf{z})$ to a particular parametric class $q(\mathbf{z}|\mathbf{x}, \phi)$ conditioned on samples \mathbf{x} with parameters ϕ .

Variational Bayes

E-step

$$\phi_k = \phi_{k-1} + \eta \nabla_{\phi} \mathcal{L}(\phi, \theta_{k-1})|_{\phi = \phi_{k-1}}$$

M-step

$$\theta_k = \theta_{k-1} + \eta \nabla_{\theta} \mathcal{L}(\phi_k, \theta)|_{\theta = \theta_{k-1}}$$

Variational EM-algorithm

ELBO

$$\log p(\mathbf{x}|\boldsymbol{\theta}) = \mathcal{L}(q,\boldsymbol{\theta}) + KL(q(\mathbf{z})||p(\mathbf{z}|\mathbf{x},\boldsymbol{\theta})) \geq \mathcal{L}(q,\boldsymbol{\theta}).$$

E-step

$$\phi_k = \phi_{k-1} + \eta \nabla_{\phi} \mathcal{L}(\phi, \theta_{k-1})|_{\phi = \phi_{k-1}},$$

where ϕ – parameters of variational distribution $q(\mathbf{z}|\mathbf{x},\phi)$.

M-step

$$\theta_k = \theta_{k-1} + \eta \nabla_{\theta} \mathcal{L}(\phi_k, \theta)|_{\theta = \theta_{k-1}},$$

where θ – parameters of the generation function $p(\mathbf{x}|\mathbf{z},\theta)$.

Now all we have to do is to obtain two gradients $\nabla_{\phi} \mathcal{L}(\phi, \theta)$, $\nabla_{\theta} \mathcal{L}(\phi, \theta)$.

Difficulty: number of samples n.

Summary

- Bayesian inference is a generalization of most common machine learning tasks. It allows to construct MLE, MAP and bayesian inference, to compare models complexity and many-many more cool stuff.
- ► LVM introduce latent representation of observed samples to make model more interpretable.
- LVM maximizes variational evidence lower bound to find MLE of model parameters.
- ► ELBO maximization is performed by the general variational EM algorithm.