# Deep Generative Models Lecture 2

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Spring, 2021

## Recap of previous lecture

We are given i.i.d. samples  $\{\mathbf{x}_i\}_{i=1}^n \in X$  (e.g.  $X = \mathbb{R}^m$ ) from unknown distribution  $\pi(\mathbf{x})$ .

#### Goal

We would like to learn a distribution  $\pi(\mathbf{x})$  for

- evaluating  $\pi(\mathbf{x})$  for new samples (how likely to get object  $\mathbf{x}$ ?);
- ▶ sampling from  $\pi(\mathbf{x})$  (to get new objects  $\mathbf{x} \sim \pi(\mathbf{x})$ ).

## Challenge

Data is complex and high-dimensional.

## MLE problem

Fix probabilistic model  $p(\mathbf{x}|\theta)$  – a set of parameterized distributions, such that  $p(\mathbf{x}|\theta) \approx \pi(\mathbf{x})$ .

$$m{ heta}^* = rg \max_{m{ heta}} p(\mathbf{X}|m{ heta}) = rg \max_{m{ heta}} \prod_{i=1}^n p(\mathbf{x}_i|m{ heta}) = rg \max_{m{ heta}} \sum_{i=1}^n \log p(\mathbf{x}_i|m{ heta}).$$

## Recap of previous lecture

## Likelihood as product of conditionals

Let 
$$\mathbf{x} = (x_1, \dots, x_m)$$
,  $\mathbf{x}_{1:i} = (x_1, \dots, x_i)$ . Then

$$p(\mathbf{x}|\boldsymbol{\theta}) = \prod_{i=1}^{m} p(x_i|\mathbf{x}_{1:i-1}, \boldsymbol{\theta}); \quad \log p(\mathbf{x}|\boldsymbol{\theta}) = \sum_{i=1}^{m} \log p(x_i|\mathbf{x}_{1:i-1}, \boldsymbol{\theta}).$$

## MLE problem for autoregressive model

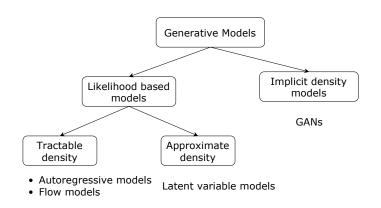
$$\theta^* = \underset{\theta}{\operatorname{arg max}} p(\mathbf{X}|\theta) = \underset{\theta}{\operatorname{arg max}} \sum_{i=1}^n \sum_{j=1}^m \log p(x_{ij}|\mathbf{x}_{i,1:j-1}\theta).$$

## Sampling

$$\hat{x}_1 \sim p(x_1|\boldsymbol{\theta}), \quad \hat{x}_2 \sim p(x_2|\hat{x}_1,\boldsymbol{\theta}), \ldots, \quad \hat{x}_m \sim p(x_m|\hat{\mathbf{x}}_{1:m-1},\boldsymbol{\theta})$$

New generated object is  $\hat{\mathbf{x}} = (\hat{x}_1, \hat{x}_2, \dots, \hat{x}_m)$ .

#### Generative models zoo



### Bayes theorem

$$p(\mathbf{t}|\mathbf{x}) = \frac{p(\mathbf{x}|\mathbf{t})p(\mathbf{t})}{p(\mathbf{x})} = \frac{p(\mathbf{x}|\mathbf{t})p(\mathbf{t})}{\int p(\mathbf{x}|\mathbf{t})p(\mathbf{t})d\mathbf{t}}$$

- x − observed variables, t − unobserved variables (latent variables/parameters);
- $p(\mathbf{x}|\mathbf{t})$  likelihood;
- $p(\mathbf{x}) = \int p(\mathbf{x}|\mathbf{t})p(\mathbf{t})d\mathbf{t}$  evidence;
- ▶  $p(\mathbf{t})$  prior distribution,  $p(\mathbf{t}|\mathbf{x})$  posterior distribution.

## Meaning

We have unobserved variables  $\mathbf{t}$  and some prior knowledge about them  $p(\mathbf{t})$ . Then, the data  $\mathbf{x}$  has been observed. Posterior distribution  $p(\mathbf{t}|\mathbf{x})$  summarizes the knowledge after the obbservations.

Let consider the case, where the unobserved variables  ${\bf t}$  is our model parameters  ${m heta}.$ 

- $\mathbf{X} = \{\mathbf{x}_i\}_{i=1}^n$  observed samples;
- ▶  $p(\theta)$  prior parameters distribution (we treat model parameters  $\theta$  as random variables).

#### Posterior distribution

$$p(\theta|\mathbf{X}) = \frac{p(\mathbf{X}|\theta)p(\theta)}{p(\mathbf{X})} = \frac{p(\mathbf{X}|\theta)p(\theta)}{\int p(\mathbf{X}|\theta)p(\theta)d\theta}$$

## Bayesian inference

$$p(\mathbf{x}|\mathbf{X}) = \int p(\mathbf{x}|\theta)p(\theta|\mathbf{X})d\theta$$

Note the difference from

$$p(\mathbf{x}) = \int p(\mathbf{x}|\boldsymbol{\theta})p(\boldsymbol{\theta})d\boldsymbol{\theta}.$$

#### Posterior distribution

$$p(\boldsymbol{\theta}|\mathbf{X}) = \frac{p(\mathbf{X}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{p(\mathbf{X})} = \frac{p(\mathbf{X}|\boldsymbol{\theta})p(\boldsymbol{\theta})}{\int p(\mathbf{X}|\boldsymbol{\theta})p(\boldsymbol{\theta})d\boldsymbol{\theta}}$$

## Bayesian inference

$$p(\mathbf{x}|\mathbf{X}) = \int p(\mathbf{x}|\theta)p(\theta|\mathbf{X})d\theta$$

If evidence  $p(\mathbf{X})$  is intractable (due to multidimensional integration), we can't get posterior distribution and perform the precise inference.

## Maximum a posteriori (MAP) estimation

$$\boldsymbol{\theta}^* = \argmax_{\boldsymbol{\theta}} p(\boldsymbol{\theta}|\mathbf{X}) = \argmax_{\boldsymbol{\theta}} \left(\log p(\mathbf{X}|\boldsymbol{\theta}) + \log p(\boldsymbol{\theta})\right)$$

#### MAP estimation

$$\boldsymbol{\theta}^* = \argmax_{\boldsymbol{\theta}} p(\boldsymbol{\theta}|\mathbf{X}) = \argmax_{\boldsymbol{\theta}} \left(\log p(\mathbf{X}|\boldsymbol{\theta}) + \log p(\boldsymbol{\theta})\right)$$

Estimated  $\theta^*$  is a deterministic variable, but we could treat it as a random variable with density  $p(\theta|\mathbf{X}) = \delta(\theta - \theta^*)$ .

#### Dirac delta function

$$\delta(x) = \begin{cases} +\infty, & x = 0; \\ 0, & x \neq 0; \end{cases} \qquad \int f(x)\delta(x-y)dx = f(y).$$

#### MAP inference

$$p(\mathbf{x}|\mathbf{X}) = \int p(\mathbf{x}|\theta)p(\theta|\mathbf{X})d\theta \approx p(\mathbf{x}|\theta^*).$$

# Latent variable models (LVM)

#### MLE problem

$$m{ heta}^* = rg \max_{m{ heta}} p(\mathbf{X}|m{ heta}) = rg \max_{m{ heta}} \prod_{i=1}^n p(\mathbf{x}_i|m{ heta}) = rg \max_{m{ heta}} \sum_{i=1}^n \log p(\mathbf{x}_i|m{ heta}).$$

#### Challenge

 $p(\mathbf{x}|\boldsymbol{\theta})$  could be intractable.

#### Extend probabilistic model

Introduce latent variable z for each sample x

$$p(\mathbf{x}, \mathbf{z}|\boldsymbol{\theta}) = p(\mathbf{x}|\mathbf{z}, \boldsymbol{\theta})p(\mathbf{z}); \quad \log p(\mathbf{x}, \mathbf{z}|\boldsymbol{\theta}) = \log p(\mathbf{x}|\mathbf{z}, \boldsymbol{\theta}) + \log p(\mathbf{z}).$$

$$p(\mathbf{x}|\boldsymbol{\theta}) = \int p(\mathbf{x}, \mathbf{z}|\boldsymbol{\theta}) d\mathbf{z} = \int p(\mathbf{x}|\mathbf{z}, \boldsymbol{\theta}) p(\mathbf{z}) d\mathbf{z}.$$

#### Motivation

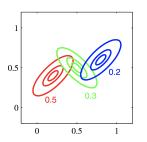
The distributions  $p(\mathbf{x}|\mathbf{z}, \theta)$  and  $p(\mathbf{z})$  could be quite simple.

# Latent variable models (LVM)

$$\log p(\mathbf{x}|oldsymbol{ heta}) = \log \int p(\mathbf{x}|\mathbf{z},oldsymbol{ heta}) p(\mathbf{z}) d\mathbf{z} o \max_{oldsymbol{ heta}}$$

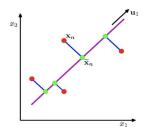
#### **Examples**

Mixture of gaussians



- $ho(\mathbf{x}|\mathbf{z}, oldsymbol{ heta}) = \mathcal{N}(\mathbf{x}|oldsymbol{\mu}_{\mathbf{z}}, oldsymbol{\Sigma}_{\mathbf{z}})$
- $p(z) = \text{Categorical}(z|\pi)$

#### PCA model

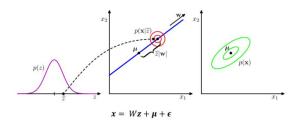


- $\qquad \qquad \rho(\mathsf{x}|\mathsf{z},\theta) = \mathcal{N}(\mathsf{x}|\mathsf{Wz} + \boldsymbol{\mu},\boldsymbol{\Sigma}_{\mathsf{z}})$

# Latent variable models (LVM)

$$\log p(\mathbf{x}|oldsymbol{ heta}) = \log \int p(\mathbf{x}|\mathbf{z},oldsymbol{ heta}) p(\mathbf{z}) d\mathbf{z} 
ightarrow \max_{oldsymbol{ heta}}$$

**PCA goal:** Project original data **X** onto a low dimensional latent space while maximizing the variance of the projected data.



- $p(z) = \mathcal{N}(z|0, I)$

# Incomplete likelihood

**MLE** 

$$egin{aligned} m{ heta}^* &= rg\max_{m{ heta}} p(\mathbf{X}, \mathbf{Z} | m{ heta}) = rg\max_{m{ heta}} \prod_{i=1}^n p(\mathbf{x}_i, \mathbf{z}_i | m{ heta}) = \ &= rg\max_{m{ heta}} \sum_{i=1}^n \log p(\mathbf{x}_i, \mathbf{z}_i | m{ heta}). \end{aligned}$$

Since **Z** is unknown, maximize **incomplete likelihood**.

#### MILE problem

$$\begin{aligned} \boldsymbol{\theta}^* &= \arg\max_{\boldsymbol{\theta}} \log p(\mathbf{X}|\boldsymbol{\theta}) = \arg\max_{\boldsymbol{\theta}} \sum_{i=1}^n \log p(\mathbf{x}_i|\boldsymbol{\theta}) = \\ &= \arg\max_{\boldsymbol{\theta}} \sum_{i=1}^n \log \int p(\mathbf{x}_i, \mathbf{z}_i|\boldsymbol{\theta}) d\mathbf{z}_i = \\ &= \arg\max_{\boldsymbol{\theta}} \log \int p(\mathbf{x}_i|\mathbf{z}_i, \boldsymbol{\theta}) p(\mathbf{z}_i) d\mathbf{z}_i. \end{aligned}$$

## Variational lower bound

$$\log p(\mathbf{x}|\theta) = \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{p(\mathbf{z}|\mathbf{x}, \theta)} =$$

$$= \int q(\mathbf{z}) \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{p(\mathbf{z}|\mathbf{x}, \theta)} d\mathbf{z} = \int q(\mathbf{z}) \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)q(\mathbf{z})}{p(\mathbf{z}|\mathbf{x}, \theta)q(\mathbf{z})} d\mathbf{z} =$$

$$= \int q(\mathbf{z}) \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{q(\mathbf{z})} d\mathbf{z} + \int q(\mathbf{z}) \log \frac{q(\mathbf{z})}{p(\mathbf{z}|\mathbf{x}, \theta)} d\mathbf{z} =$$

$$= \mathcal{L}(q, \theta) + KL(q(\mathbf{z})||p(\mathbf{z}|\mathbf{x}, \theta)) \ge \mathcal{L}(q, \theta).$$

## Kullback-Leibler divergence

- $KL(q||p) = \int q(\mathbf{z}) \log \frac{q(\mathbf{z})}{p(\mathbf{z})} d\mathbf{z};$
- $KL(q||p) \geq 0$ ;
- $\mathsf{KL}(q||p) = 0 \Leftrightarrow q \equiv p.$

#### Variational lower bound

$$\log p(\mathbf{x}|\boldsymbol{\theta}) = \mathcal{L}(q, \boldsymbol{\theta}) + KL(q(\mathbf{z})||p(\mathbf{z}|\mathbf{x}, \boldsymbol{\theta})) \geq \mathcal{L}(q, \boldsymbol{\theta}).$$

## Evidence Lower Bound (ELBO)

$$\mathcal{L}(q, \theta) = \int q(\mathbf{z}) \log \frac{p(\mathbf{x}, \mathbf{z}|\theta)}{q(\mathbf{z})} d\mathbf{z} =$$

$$= \int q(\mathbf{z}) \log p(\mathbf{x}|\mathbf{z}, \theta) d\mathbf{z} + \int q(\mathbf{z}) \log \frac{p(\mathbf{z})}{q(\mathbf{z})} d\mathbf{z}$$

$$= \mathbb{E}_q \log p(\mathbf{x}|\mathbf{z}, \theta) - KL(q(\mathbf{z})||p(\mathbf{z}))$$

Instead of maximizing incomplete likelihood, maximize ELBO (equivalently minimize KL)

$$\max_{\theta} p(\mathbf{x}|\theta) \quad o \quad \max_{q,\theta} \mathcal{L}(q,\theta) \equiv \min_{q,\theta} \mathit{KL}(q(\mathbf{z})||p(\mathbf{z}|\mathbf{x},\theta)).$$

# EM-algorithm

$$\mathcal{L}(q, \theta) = \int q(\mathbf{z}) \log p(\mathbf{x}|\mathbf{z}, \theta) d\mathbf{z} + \int q(\mathbf{z}) \log \frac{p(\mathbf{z})}{q(\mathbf{z})} d\mathbf{z}.$$

#### Block-coordinate optimization

- Initialize θ\*;
- E-step

$$q(\mathbf{z}) = \underset{q}{\operatorname{arg max}} \mathcal{L}(q, \boldsymbol{\theta}^*) = \underset{q}{\operatorname{arg min}} \mathit{KL}(q||p) = p(\mathbf{z}|\mathbf{x}, \boldsymbol{\theta}^*);$$

M-step

$$oldsymbol{ heta}^* = rg \max_{oldsymbol{q}} \mathcal{L}(oldsymbol{q}, oldsymbol{ heta});$$

Repeat E-step and M-step until convergence.

# Ugly pic

#### Amortized variational inference

#### E-step

$$q(\mathbf{z}) = rg \max_{q} \mathcal{L}(q, \boldsymbol{\theta}^*) = rg \min_{q} \mathit{KL}(q||p) = p(\mathbf{z}|\mathbf{x}, \boldsymbol{\theta}^*).$$

- ▶  $p(\mathbf{z}|\mathbf{x}, \boldsymbol{\theta}^*)$  could be **intractable**;
- $ightharpoonup q(\mathbf{z})$  is different for each object  $\mathbf{x}$ .

#### Idea

Restrict a family of all possible distributions  $q(\mathbf{z})$  to a particular parametric class  $q(\mathbf{z}|\mathbf{x}, \phi)$  conditioned on samples  $\mathbf{x}$  with parameters  $\phi$ .

#### Variational Bayes

E-step

$$\phi_k = \phi_{k-1} + \eta \nabla_{\phi} \mathcal{L}(\phi, \theta_{k-1})|_{\phi = \phi_{k-1}}$$

M-step

$$\theta_k = \theta_{k-1} + \eta \nabla_{\theta} \mathcal{L}(\phi_k, \theta)|_{\theta = \theta_{k-1}}$$

# Variational EM-algorithm

#### **ELBO**

$$\log p(\mathbf{x}|\boldsymbol{\theta}) = \mathcal{L}(q,\boldsymbol{\theta}) + KL(q(\mathbf{z})||p(\mathbf{z}|\mathbf{x},\boldsymbol{\theta})) \geq \mathcal{L}(q,\boldsymbol{\theta}).$$

E-step

$$\phi_k = \phi_{k-1} + \eta \nabla_{\phi} \mathcal{L}(\phi, \theta_{k-1})|_{\phi = \phi_{k-1}},$$

where  $\phi$  – parameters of variational distribution  $q(\mathbf{z}|\mathbf{x},\phi)$ .

M-step

$$\theta_k = \theta_{k-1} + \eta \nabla_{\theta} \mathcal{L}(\phi_k, \theta)|_{\theta = \theta_{k-1}},$$

where  $\theta$  – parameters of the generation function  $p(\mathbf{x}|\mathbf{z},\theta)$ .

Now all we have to do is to obtain two gradients  $\nabla_{\phi} \mathcal{L}(\phi, \theta)$ ,  $\nabla_{\theta} \mathcal{L}(\phi, \theta)$ .

**Difficulty:** number of samples n.

$$\sum_{i=1}^{n} \mathcal{L}_{i}(\phi, \theta) = \sum_{i=1}^{n} \mathbb{E}_{q} \log p(\mathbf{x}_{i}|\mathbf{z}_{i}, \theta) - KL(q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi)||p(\mathbf{z}_{i})) \rightarrow \max_{\phi, \theta}.$$

Optimization w.r.t.  $\theta$ : **mini-batching** (1) + **Monte-Carlo** estimation (2)

$$\nabla_{\theta} \sum_{i=1}^{n} \mathcal{L}_{i}(\phi, \theta) = \sum_{i=1}^{n} \int q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi) \nabla_{\theta} \log p(\mathbf{x}_{i}|\mathbf{z}_{i}, \theta) d\mathbf{z}_{i}$$

$$\stackrel{(1)}{\approx} n \int q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi) \nabla_{\theta} \log p(\mathbf{x}_{i}|\mathbf{z}_{i}, \theta) d\mathbf{z}_{i}, \quad i \sim U[1, n]$$

$$\stackrel{(2)}{\approx} n \nabla_{\theta} \log p(\mathbf{x}_{i}|\mathbf{z}_{i}^{*}, \theta), \quad \mathbf{z}_{i}^{*} \sim q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi).$$

Monte-Carlo estimation (2):

$$\mathbb{E}_q f(\mathsf{z}) = \int q(\mathsf{z}) f(\mathsf{z}) d\mathsf{z} pprox f(\mathsf{z}^*), ext{where } \mathsf{z}^* \sim q(\mathsf{z}).$$

$$\sum_{i=1}^{n} \mathcal{L}_{i}(\phi, \theta) = \sum_{i=1}^{n} \mathbb{E}_{q} \log p(\mathbf{x}_{i}|\mathbf{z}_{i}, \theta) - KL(q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi)||p(\mathbf{z}_{i})) \rightarrow \max_{\phi, \theta}.$$

Difference from M-step: density function  $q(\mathbf{z}|\mathbf{x}, \phi)$  depends on the parameters  $\phi$ , it is impossible to use the Monte-Carlo estimation:

$$abla_{m{\phi}} \mathcal{L}(m{\phi}, m{ heta}) = \int 
abla_{m{\phi}} q(\mathbf{z}|\mathbf{x}, m{\phi}) \log p(\mathbf{x}|\mathbf{z}, m{ heta}) d\mathbf{z} - 
abla_{m{\phi}} KL$$

First term is not an expectation due to the gradient.

#### Possible solutions

- log-derivative trick;
- reparametrization trick.

#### Log-derivative trick

$$abla_{\xi} q(\eta|\xi) = q(\eta|\xi) \left( \frac{\nabla_{\xi} q(\eta|\xi)}{q(\eta|\xi)} \right) = q(\eta|\xi) \nabla_{\xi} \log q(\eta|\xi).$$

$$abla_{\phi} q(\mathbf{z}|\mathbf{x}, \phi) = q(\mathbf{z}|\mathbf{x}, \phi) \nabla_{\phi} \log q(\mathbf{z}|\mathbf{x}, \phi).$$

#### **ELBO**

$$\nabla_{\phi} \sum_{i=1}^{n} \mathcal{L}_{i}(\phi, \theta) = \sum_{i=1}^{n} \int \nabla_{\phi} q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi) \log p(\mathbf{x}_{i}|\mathbf{z}_{i}, \theta) d\mathbf{z}_{i} - \nabla_{\phi} KL$$

$$= \sum_{i=1}^{n} \int q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi) \left[ \nabla_{\phi} \log q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi) \log p(\mathbf{x}_{i}|\mathbf{z}_{i}, \theta) \right] d\mathbf{z}_{i} - \nabla_{\phi} KL$$

$$\approx n \nabla_{\phi} \log q(\mathbf{z}_{i}^{*}|\mathbf{x}_{i}, \phi) \log p(\mathbf{x}_{i}|\mathbf{z}_{i}^{*}, \theta) - \nabla_{\phi} KL, \quad \mathbf{z}_{i}^{*} \sim q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi).$$

#### **Problem**

Unstable solution with huge variance.

#### Reparametrization trick

$$f(\xi) = \int q(\eta|\xi)h(\eta)d\eta$$

Let  $\eta = g(\xi, \epsilon)$ , where g is a deterministic function,  $\epsilon$  is a random variable with a density function  $r(\epsilon)$ .

$$egin{aligned} 
abla_{\xi} \int q(\eta|\xi)h(\eta)d\eta &= 
abla_{\xi} \int r(\epsilon)h(g(\xi,\epsilon))d\epsilon \ &pprox 
abla_{\xi}h(g(\xi,\epsilon^*)), \quad \epsilon^* \sim r(\epsilon). \end{aligned}$$

#### Example

$$q(\eta|\xi) = \mathcal{N}(\eta|\mu, \sigma^2), \quad r(\epsilon) = \mathcal{N}(\epsilon|0, 1), \quad \eta = \sigma \cdot \epsilon + \mu, \quad \xi = [\mu, \sigma].$$

$$\nabla_{\phi} \sum_{i=1}^{n} \mathcal{L}_{i}(\phi, \theta) = \sum_{i=1}^{n} \nabla_{\phi} \int q(\mathbf{z}_{i}|\mathbf{x}_{i}, \phi) \log p(\mathbf{x}_{i}|\mathbf{z}_{i}, \theta) d\mathbf{z}_{i} - \nabla_{\phi} KL$$

$$\approx n \nabla_{\phi} \int r(\epsilon) \log p(\mathbf{x}_{i}|g(\mathbf{x}_{i}, \epsilon, \phi), \theta) d\epsilon - \nabla_{\phi} KL, \quad i \sim U[1, n]$$

$$\approx n \nabla_{\phi} \log p(\mathbf{x}_{i}|g(\mathbf{x}_{i}, \epsilon^{*}, \phi), \theta) - \nabla_{\phi} KL, \quad \epsilon^{*} \sim r(\epsilon).$$

#### Variational assumption

$$egin{aligned} q(\mathbf{z}|\mathbf{x},\phi) &= \mathcal{N}(\mu(\mathbf{x}),\sigma(\mathbf{x})). \ \mathbf{z} &= g(\mathbf{x},\epsilon,\phi) &= \sigma(\mathbf{x})\cdot\epsilon + \mu(\mathbf{x}). \end{aligned}$$

 $\nabla_{\phi} KL(q(\mathbf{z}|\mathbf{x},\phi)||p(\mathbf{z}))$  has an analytical solution.

# Variational autoencoder (VAE)

#### Final algorithm

- ▶ pick  $i \sim U[1, n]$ ;
- ightharpoonup compute a stochastic gradient w.r.t.  $\phi$

$$abla_{\phi} \mathcal{L}(\phi, \theta) = n \nabla_{\phi} \log p(\mathbf{x}_{i} | g(\mathbf{x}_{i}, \epsilon^{*}, \phi), \theta) - \\
- \nabla_{\phi} KL(q(\mathbf{z}_{i} | \mathbf{x}_{i}, \phi) || p(\mathbf{z}_{i})), \quad \epsilon^{*} \sim r(\epsilon);$$

ightharpoonup compute a stochastic gradient w.r.t. heta

$$\nabla_{\boldsymbol{\theta}} \mathcal{L}(\boldsymbol{\phi}, \boldsymbol{\theta}) = n \nabla_{\boldsymbol{\theta}} \log p(\mathbf{x}_i | \mathbf{z}_i^*, \boldsymbol{\theta}), \quad \mathbf{z}_i^* \sim q(\mathbf{z}_i | \mathbf{x}_i, \boldsymbol{\phi});$$

• update  $\theta$ ,  $\phi$  according to the selected optimization method (SGD, Adam, RMSProp).

# Variational autoencoder (VAE)

- ▶ Encoder  $q(\mathbf{z}|\mathbf{x}, \phi) = \mathsf{NN}_{\mathsf{e}}(\mathbf{x}, \phi)$  outputs  $\mu(\mathbf{x})$  and  $\sigma(\mathbf{x})$ .
- ▶ Decoder  $p(\mathbf{x}|\mathbf{z}, \boldsymbol{\theta}) = \mathsf{NN}_d(\mathbf{z}, \boldsymbol{\theta})$  outputs parameters of the sample distribution.

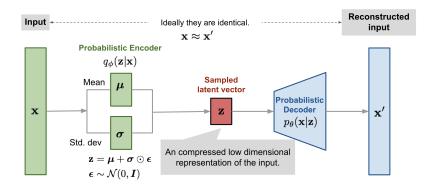
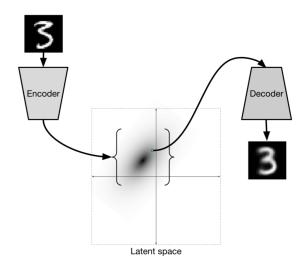


image credit:

## Variational Autoencoder



#### Variational Autoencoder

Generation objects by sampling the latent space  $\mathbf{z} \sim \mathcal{N}(0, \mathbf{I})$ 

## Summary

- Bayesian inference is a generalization of most common machine learning tasks. It allows to construct MLE, MAP and bayesian inference, to compare models complexity and many-many more cool stuff.
- ► LVM introduce latent representation of observed samples to make model more interpretable.
- LVM maximizes variational evidence lower bound to find MLE of model parameters.
- ► ELBO maximization is performed by the general variational EM algorithm.
- Amortized inference allows to efficiently compute stochastic gradients for ELBO and to use deep neural networks for  $q(\mathbf{z}|\mathbf{x}, \phi)$  and  $p(\mathbf{x}|\mathbf{z}, \theta)$ .
- The VAE model is an LVM with an encoder network for  $q(\mathbf{z}|\mathbf{x}, \phi)$  and a decoder network for  $p(\mathbf{x}|\mathbf{z}, \theta)$ .