PCA and MDS A Geometric View

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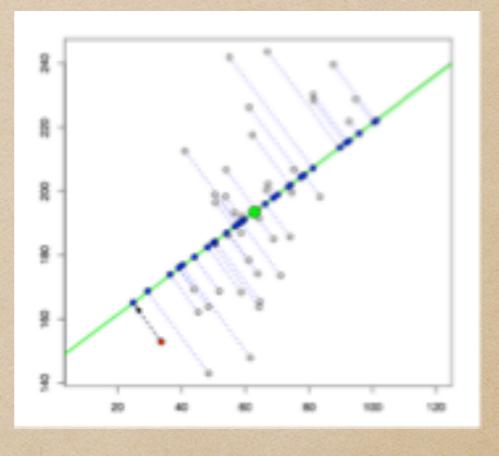
Geometric Embedding

- . A Fundamental Problem in Data Representation
- · Unstructured data -> Euclidean Space
 - . PCA: high dim -> low dim affine space
 - . MDS: metric -> Euclidean space
- . a.k.a. 'feature' learning (e.g. deep learning)
- . speech, text, image, video...

Principal Component Analysis (PCA)

Let $x_i \in \mathbb{R}^p$, i = 1, ..., n, be n samples in \mathbb{R}^p .

· Can you find a low dimensional affine representation?



Principal Component Analysis (PCA)

Best k-affine space approximation of data:

Let
$$X = [X_1|X_2|\cdots|X_n] \in \mathbb{R}^{p \times n}$$
.

(2)
$$\min_{\beta,\mu,U} I := \sum_{i=1}^{n} \|X_i - (\mu + U\beta_i)\|^2$$

where $U \in \mathbb{R}^{p \times k}$, $U^T U = I_p$, and $\sum_{i=1}^n \beta_i = 0$ (nonzero sum of β_i can be repre-

Plug in the expression of $\hat{\mu}_n$ and β_i

$$I = \sum_{i=1}^{n} \|X_i - \hat{\mu}_n - UU^T (X_i - \hat{\mu}_n)\|^2$$

$$= \sum_{i=1}^{n} \|X_i - \hat{\mu}_n - P_k (X_i - \hat{\mu}_n)\|^2$$

$$= \sum_{i=1}^{n} \|Y_i - P_k (y_i)\|^2, \quad Y_i := X_i - \hat{\mu}_n$$

where $P_k = UU^T$ is a projection operator satisfying the idempotent property $P_k^2 = P_k$.

Denote $Y = [Y_1|Y_2|\cdots|Y_n] \in \mathbb{R}^{p \times n}$, whence the original problem turns into

$$\begin{aligned} \min_{U} \sum_{i=1}^{n} \|Y_i - P_k(Y_i)\|^2 &= \min \operatorname{trace}[(Y - P_k Y)^T (Y - P_k Y)] \\ &= \min \operatorname{trace}[Y^T (I - P_k) (I - P_k) Y] \\ &= \min \operatorname{trace}[YY^T (I - P_k)^2] \\ &= \min \operatorname{trace}[YY^T (I - P_k)] \\ &= \min[\operatorname{trace}(YY^T) - \operatorname{trace}(YY^T UU^T)] \\ &= \min[\operatorname{trace}(YY^T) - \operatorname{trace}(U^T YY^T U)]. \end{aligned}$$

Above we use cyclic property of trace and idempotent property of projection.

Since Y does not depend on U, the problem above is equivalent to

(3)
$$\max_{UU^T=I_k} Var(U^TY) = \max_{UU^T=I_k} \frac{1}{n} \operatorname{trace}(U^TYY^TU) = \max_{UU^T=I_k} \operatorname{trace}(U^T\hat{\Sigma}_n U)$$

where $\hat{\Sigma}_n = \frac{1}{n} Y Y^T = \frac{1}{n} (X - \hat{\mu}_n \mathbf{1}^T) (X - \hat{\mu}_n \mathbf{1}^T)^T$ is the sample variance. Assume

$$\frac{\partial I}{\partial \mu} = -2\sum_{i=1}^{n} (X_i - \mu - U\beta_i) = 0 \Rightarrow \hat{\mu}_n = \frac{1}{n} \sum_{i=1}^{n} X_i$$

$$\frac{\partial I}{\partial \beta_i} = (x_i - \mu - U\beta_i)^T U = 0 \Rightarrow \beta_i = U^T (X_i - \mu)$$

Principal Component Analysis Summary

. PCA is given by the topk eigenvector of covariance matrix

$$\widehat{\Sigma}_n = \frac{1}{n-1} \widehat{X} \cdot \widehat{X}^T$$

$$\widetilde{X} = XH = X - \frac{1}{n}X \cdot \mathbf{1}\mathbf{1}^T = \widetilde{U}\widetilde{S}\widetilde{V}^T, \quad H = I - \frac{1}{n}\mathbf{1}\mathbf{1}^T, \ \mathbf{1} = (1, \dots, 1)^T \in \mathbb{R}^n$$

(left) singular vectors here gives PCA

How much variances in data explained by PCA?

total variance:

$$\operatorname{trace}(\hat{\Sigma}_n) = \sum_{i=1}^{p} \hat{\lambda}_i;$$

percentage of variance explained by top-k principal components:

$$\sum_{i=1}^k \hat{\lambda}_i/\mathrm{trace}(\hat{\Sigma}_n);$$

· generalized variance as total volume:

$$\det(\hat{\Sigma}_n) = \prod_{i=1}^p \hat{\lambda}_i.$$

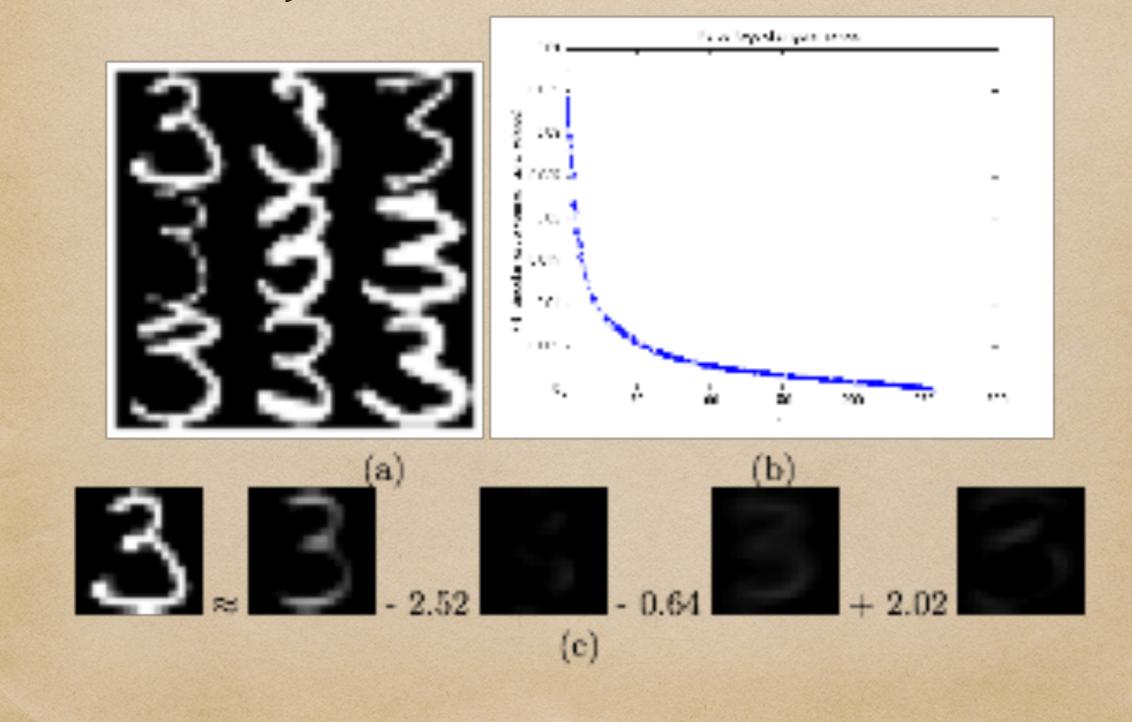
Example:

. Choose k such that

$$\sum_{i=1}^{k} \hat{\lambda}_i / \operatorname{trace}(\hat{\Sigma}_n) > 0.95.$$

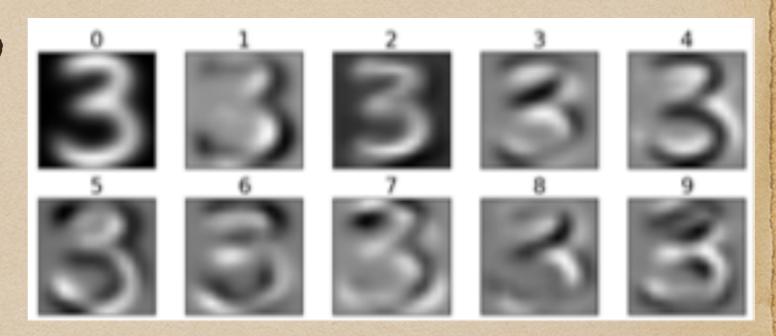
explains 95% total variations in data

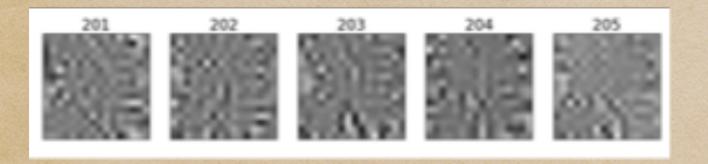
Example of PCA



Mean and PCs

Mean and top 9
PCs that
explains > 98%
total variations





PC 201-205 are less informative

Random Permutation Test

- . a.k.a. Horn's Parallel Analysis
- randomly permute samples for decorrelation
- · compute eigenvalues of random matrices

$$X = \begin{bmatrix} X_{1,1} & X_{1,2} & \cdots & X_{1,n} \\ X_{2,1} & X_{2,2} & \cdots & X_{2,n} \\ \vdots & \vdots & \ddots & \vdots \\ X_{p,1} & X_{p,2} & \cdots & X_{p,n} \end{bmatrix} \longrightarrow X^{1} = \begin{bmatrix} X_{1,\pi_{1}(1)} & X_{1,\pi_{1}(2)} & \cdots & X_{1,\pi_{1}(n)} \\ X_{2,\pi_{2}(1)} & X_{2,\pi_{2}(2)} & \cdots & X_{2,\pi_{2}(n)} \\ \vdots & \vdots & \ddots & \vdots \\ X_{p,\pi_{p}(1)} & X_{p,\pi_{p}(2)} & \cdots & X_{p,\pi_{p}(n)} \end{bmatrix}$$

$$\{\hat{\lambda}_i^1\}_{i=1,\dots,p}$$

Horn's Parallel Analysis

 Repeat this for R times, obtain a R-by-p eigenvalue matrix

 $\begin{bmatrix} \lambda_1^1 & \lambda_2^1 & \cdots & \lambda_p^1 \\ \hat{\lambda}_1^2 & \hat{\lambda}_2^2 & \cdots & \hat{\lambda}_p^2 \end{bmatrix}$

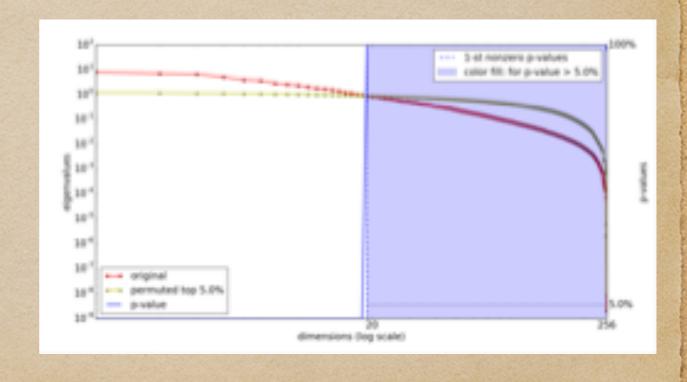
Define the p-value for the i-th eigenvalue, and only keep eigenvalues whose pval is smaller than a threshold, e.g.

$$\mathrm{pval}_i = \frac{1}{R} \# \{ \hat{\lambda}_i^r > \hat{\lambda}_i \} \qquad \mathrm{pval}_i < 0.05$$

Example

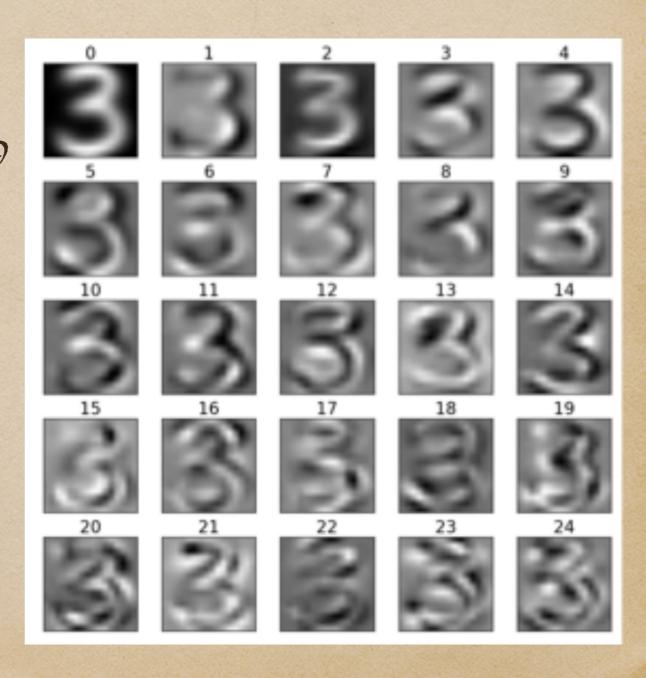
- Fig 1. Examples of randomly permuted data
- Fig 2. Results of parallel analysis on digit 3, with R≈100.
 There are 19 PCs whose pval's <5%





Example

· Top 24 PCs and top 19 is suggested by parallel analysis at 5% level, which might be conservative as the images lie near a submanifold in the space



Summary: PCA=SVD

(SVD) of $X = [x_1, \dots, x_n]^T \in \mathbb{R}^{n \times p}$ in the following sense,

$$Y = X - \frac{1}{n} \mathbf{1} \mathbf{1}^T X = \hat{U} \hat{S} \hat{V}^T, \quad \mathbf{1} = (1, ..., 1)^T \in \mathbb{R}^n$$

²In statistics, data matrix is often written as samples row-wise and variables column-wise, i.e. n-by-p matrix. So be careful on your way of writing the data matrix!

- top k right singular vectors give PCA (Covariance spectrum)
- top k left singular vectors? It gives
 MDS (Kernel spectrum)