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- If $X(\Omega)$ is infinite uncountable X is said to be a continuous random variable, denoted c.r.v.

Moreover a c.r.v. is said to be absolutely continuous if it admits a continuous and derivable distribution function (except possibly at some points).

Discrete random variables

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Probability distribution of a discrete random variable

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Definition

Let X be a d.r.v. one calls probability distribution or mass function of the r.v. X the application

$$\begin{aligned} p &: \mathbb{R} \longrightarrow [0, 1] \\ x &\longmapsto p(x) = \mathbb{P}(X = x). \end{aligned}$$

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Properties:

1. $\forall x \in \mathbb{R}, p(x) \geq 0$;
2. $\sum_{x \in \mathbb{R}} p(x) = 1$.

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The distribution is usually written in the following form

x	0	1	2	$\sum_{x=0}^2 p(x)$
$\mathbb{P}(X = x)$	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$	1

Discrete random variables

Cumulative distribution function

1. If X is a discrete r.v. then

$$F_X(x) = \sum_{x_i \leq x} \mathbb{P}(X = x_i) = \sum_{x_i \leq x} p(x_i).$$

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Indeed, $\forall x_j \in X(\Omega)$

$$\mathbb{P}(X = x_j) = \sum_{i=1}^j \mathbb{P}(X \leq x_i) - \sum_{i=1}^{j-1} \mathbb{P}(X \leq x_i) = F_X(x_j) - F_X(x_{j-1}).$$

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The probability law of X is

x	0	1	2	3	$\sum_{x=0}^3 p(x)$
$\mathbb{P}(X = x)$	$\frac{5^3}{6^3}$	$\frac{3 \cdot 5^2}{6^3}$	$\frac{3 \cdot 5}{6^3}$	$\frac{1}{6^3}$	1

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We throw 3 dice and define the r.v. X as the number of 6 obtained.

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x	0	1	2	3	$\sum_{x=0}^x p(x)$
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$$F_X(x) = \begin{cases} 0 & \text{if } x < 0 \\ \frac{125}{216} & \text{if } 0 \leq x < 1 \\ \frac{200}{216} & \text{if } 1 \leq x < 2 \\ \frac{215}{216} & \text{if } 2 \leq x < 3 \\ 1 & \text{if } 3 \leq x \end{cases}$$

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1. F_X is continuous on \mathbb{R} ;
2. F_X is derivable in every point $x \in \mathbb{R}$ except perhaps on a finite set D .

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Theorem

Let X be an absolutely continuous random variable, with cumulative distribution function F_X , then for any pair $(a, b) \in \mathbb{R}^2$ such that $a < b$, we have

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1. $\mathbb{P}(X = a) = 0$.
2. $\mathbb{P}(X \in]a, b]) = \mathbb{P}(X \in]a, b[) = \mathbb{P}(X \in [a, b[) = \mathbb{P}(X \in [a, b]) = F_X(b) - F_X(a)$.

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3. $\mathbb{P}(X \in]a, \infty[) = \mathbb{P}(X \in [a, \infty[) = 1 - F_X(a).$
4. $\mathbb{P}(X \in]-\infty, b]) = \mathbb{P}(X \in]-\infty, b[) = F_X(b).$

Continuous random variables

Definition

A real random variable X defined on a probability space $(\Omega, \mathcal{A}, \mathbb{P})$ with cumulative distribution function F_X is said to be absolutely continuous random variable, if there exists a real function f_X satisfying the following conditions:

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1. $f_X(x) \geq 0; \forall x \in \mathbb{R};$
2. f_X is continuous on \mathbb{R} , except perhaps on a finite number of points where it has a finite left limit and finite right limit.

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3. The integral $\int_{-\infty}^{+\infty} f_X(x) dx$ exists and is equal to 1.

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4. The cumulative distribution function F_X can be written, for all $x \in \mathbb{R}$ in the form

$$F_X(x) = \int_{-\infty}^x f_X(s) ds.$$

Continuous random variables

Definition

A function f that satisfies the four previous conditions is called a probability density function or distribution function of an absolutely continuous random variable X .

Continuous random variables

Example

Let X be a random variable with cumulative distribution function F_X given by

$$F_X(x) = \begin{cases} 0 & \text{if } x < 0 \\ 1 - \frac{1}{2}(x+2)e^{-\frac{x}{2}} & \text{if } x \geq 0 \end{cases}$$

1. Show that the random variable X is absolutely continuous.
2. Find the constant C such that the function f defined by

$$f(x) = \begin{cases} Cxe^{-\frac{x}{2}} & \text{if } x \geq 0 \\ 0 & \text{elsewhere} \end{cases}$$

be the probability density of the random variable X .

3. Verify that

$$F_X(x) = \int_{-\infty}^x f(s) ds.$$

Solution

1. F_X is continuous on $] -\infty, 0[$ and on $] 0, +\infty[$ show that it is continuous in 0. We have $\lim_{x \rightarrow 0} \left(1 - \frac{1}{2} (x + 2) e^{-\frac{x}{2}} \right) = 0$ hence F_X is continuous in 0.

F_X is derivable on $] -\infty, 0[$ and on $] 0, +\infty[$ show that it is derivable in 0. We have $\lim_{x \rightarrow 0} \left(\frac{F_X(x) - F_X(0)}{x} \right) = 0$

F_X is derivable on \mathbb{R} , hence X is an absolutely continuous variable.

Solution

2. To show that f is a density function we determine first the constant C using the condition 3 of the definition i.e.

$\int_{-\infty}^{+\infty} f(x) dx = 1$, then we verify the other conditions.

We have

$$\begin{aligned}\int_{-\infty}^{+\infty} f(x) dx &= \int_0^{+\infty} Cxe^{-\frac{x}{2}} dx = C \left(\left[-2xe^{-\frac{x}{2}} \right]_0^{\infty} + 2 \int_0^{+\infty} e^{-\frac{x}{2}} dx \right) \\ &= C \left[-4e^{-\frac{x}{2}} \right]_0^{\infty} = 4C = 1.\end{aligned}$$

hence $C = \frac{1}{4}$ and

$$f(x) = \begin{cases} 0 & \text{if } x < 0 \\ \frac{x}{4}e^{-\frac{x}{2}} & \text{if } x \geq 0 \end{cases}$$

We have $f_X(x) \geq 0; \forall x \in \mathbb{R}$.

It is a continuous function in 0 and then continuous on \mathbb{R} .

Then f is a probability density function of the random variable X .

Solution

3. If $x < 0$, $\int_{-\infty}^x f(s) ds = 0$ since on $]-\infty, 0[$, $F_X(x) = 0$
If $x \geq 0$,

$$\begin{aligned}\int_{-\infty}^x f(s) ds &= \int_{-\infty}^0 f(s) ds + \int_0^x f(s) ds = 0 + \int_0^x \frac{s}{4} e^{-\frac{s}{2}} ds \\&= \frac{1}{4} \left(\left[-2se^{-\frac{s}{2}} \right]_0^x + 2 \int_0^x e^{-\frac{s}{2}} ds \right) \\&= \frac{1}{4} \left(-2xe^{-\frac{x}{2}} - 4 \left[e^{-\frac{s}{2}} \right]_0^x \right) \\&= \frac{1}{4} \left(-2xe^{-\frac{x}{2}} - 4e^{-\frac{x}{2}} + 4 \right) = 1 - \frac{1}{2} (x+2) e^{-\frac{x}{2}}\end{aligned}$$

hence $F_X(x) = \int_{-\infty}^x f(s) ds$.

Mathematical expectation and variance

Definition

Let X be a d.r.v. with possible values x_1, x_2, \dots and mass function $p(x)$.

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Let X be a d.r.v. with possible values x_1, x_2, \dots and mass function $p(x)$. The mathematical expectation of X is

$$\mathbb{E}[X] = \sum_{i \geq 1} x_i p(x_i) = \sum_{i \geq 1} x_i \mathbb{P}(X = x_i)$$

provided that the above serie is absolutely convergent, otherwise we will say that X does not have a mathematical expectation.

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Remark

If X has a finite number of values then $\mathbb{E}[X]$ exists.

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Example

Let T be a c.r.v. with distribution function f defined by

$$f(t) = \begin{cases} \frac{1}{t^2} & \text{if } t > 1 \\ 0 & \text{elsewhere} \end{cases}$$

Determine $\mathbb{E}[T]$.

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$$\begin{aligned}\int_{-\infty}^{+\infty} |tf(t)| dt &= \int_1^{+\infty} \left| \frac{1}{t} \right| dt \\ &= \lim_{x \rightarrow \infty} \int_1^x \left| \frac{1}{t} \right| dt = \lim_{x \rightarrow \infty} \log x - \log 1 = +\infty\end{aligned}$$

hence the expectation doesn't exist.

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Definition

Let G be a function of a random variable X , the expectation of $G(X)$ is given by

$$\mathbb{E}[G(X)] = \begin{cases} \sum_{x \in \mathbb{R}} G(x) p(x) & \text{if } X \text{ is discrete} \\ \int_{-\infty}^{+\infty} G(x) f(x) dx & \text{if } X \text{ is continuous} \end{cases}$$

provided that the above serie and integral are absolutely convergent.

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2. $\mathbb{E}[\alpha H(X) + \beta G(X)] = \alpha \mathbb{E}[H(X)] + \beta \mathbb{E}[G(X)]$ where H and G are functions of X and α, β are reals. Provided that the different expectations exist.

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Definition

Let X be a random variable, we call moment of order k ($k \in \mathbb{N}$) the following value

$$\mathbb{E}[X^k] = \begin{cases} \sum_{x \in \mathbb{R}} x^k p(x) & \text{if } X \text{ is discrete} \\ \int_{-\infty}^{+\infty} x^k f(x) dx & \text{if } X \text{ is continuous} \end{cases}$$

provided that the above serie and integral are absolutely convergent.

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Let X be a random variable, the variance of X , noted σ_X^2 or $\text{Var}(X)$ is

$$\sigma_X^2 = \mathbb{E} \left[(X - \mathbb{E}[X])^2 \right] = \mathbb{E}[X^2] - \mathbb{E}[X]^2.$$

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We call standard deviation of X the number

$$\sigma_X = \sqrt{\text{Var}(X)}.$$

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If $\mathbb{E}[X] = 0$ we say that the random variable is centred.

If $\text{Var}(X) = 1$ we say that the random variable is reduced.

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Theorem

Let X be a random variable with expectation $\mathbb{E}[X]$ and variance σ_X^2 . If $Y = aX + b$ where a and b are real constants, then

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$$\mathbb{E}[Y] = a\mathbb{E}[X] + b \text{ and } \sigma_Y^2 = a^2\sigma_X^2.$$

Discrete probability distribution (finite case)

Discrete uniform distribution

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The r.v. X has the discrete uniform distribution on the set of real numbers $\{x_1, \dots, x_n\}$

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The r.v. X has the discrete uniform distribution on the set of real numbers $\{x_1, \dots, x_n\}$ if \mathbb{P}_X is the equiprobability on this set i.e.:

$X \in X(\Omega) = \{x_1, \dots, x_n\}$ and $\forall k \in \Omega, \mathbb{P}(X = k) = \frac{1}{n}$

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We note $X \rightsquigarrow \mathcal{U}(n)$.

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$$\mathbb{E}(X) = \frac{n+1}{2}; \text{Var}(X) = \frac{n^2-1}{12}.$$

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Example

When we throw a dice, the number obtained follow the uniforme distribution on $\{1, \dots, 6\}$ with $\mathbb{P}_X(x) = \frac{1}{6}, \forall x \in \{1, \dots, 6\}$.

Discrete probability distribution (finite case)

Bernoulli distribution

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Example

In the toss of an unbalanced coin, the probability of getting "heads" is $p \neq \frac{1}{2}$. X the r.v. defined by $X = 1$ if we get "heads" and $X = 0$ if we get "tails". $X \rightsquigarrow \mathcal{B}(p)$ with the probability distribution

$$\mathbb{P}(X = x) = \begin{cases} p, & \text{if } x = 1 \\ q, & \text{if } x = 0 \end{cases}$$

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Remark

The r.v. X can be defined as a sum of n independent Bernoulli r.v. X_1, X_2, \dots, X_n ($X = X_1 + X_2 + \dots + X_n$). Such that $\mathbb{P}(X_i = 1) = p$.

Discrete probability distribution (finite case)

Binomiale distribution

Definition

A r.v. X follows a binomial distribution of parameters (n, p) where $n \geq 0$ and $(p \in [0, 1])$ if $X(\Omega) = \{0, 1, \dots, n\}$ and $\mathbb{P}(X = k) = C_n^k p^k (1 - p)^{n-k}, \forall k = 0, 1, \dots, n$ (with: $p + q = 1$).

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Example

Let $X \rightsquigarrow \mathcal{B}(n, p)$.

1. Determine n such that $\mathbb{P}(X = 0) \leq 0,01$;
2. Determine n such that $\mathbb{P}(X \geq 1) \geq 0,90$.

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Hypergeometric distribution

One carries out n successive drawings of a ball, without handing-over, which is the same as when one takes a sample of n balls in only one blow, in an urn containing N balls of two categories:

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Remark

The possible values of X are $\max(0, n - N_q) \leq k \leq \min(n, N_p)$

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We note $X \rightsquigarrow \mathcal{H}(N, n, p)$, with $p = \frac{N_p}{N}$, $p + q = 1$. The a.v. X follows the hypergeometric law of parameters

$$\mathbb{E}(X) = np; \text{Var}(X) = npq \frac{N-n}{N-1}.$$

Discrete probability distribution (finite case)

Geometric distribution

The geometric distribution is the law of expectation of the first success of a sequence of independent trials each of which has a probability p of success, i.e. $\mathbb{P}(X = k)$ is the probability that the k^{th} trial is the first success.

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Example

We play heads or tails with a rigged coin such that the probability of getting tails is $\frac{1}{3}$. Let X be the r.v. representing the number of

Discrete probability distribution (finite case)

Pascal (Negative Binomial or Polya) distribution

If a r.v. represents the number of fails before the r^{th} success of a sequence of independent Bernoulli trials each of which has a probability p of success.

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