Overview

## Estimation of parameters for stochastic dynamic models

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#### Outline

- 1 Overview
- 2 Stochastic simulation
  - Discrete time
  - Continuous time
- 3 Fitting: simple approaches
  - Trajectory matching
  - Gradient matching
  - Comparison
- 4 Fancier methods
  - SIMEX
  - Kalman filter



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Timber marks

#### Modeling

rypicai stats	rypicai matn
stochastic	deterministic
static	dynamic
phenomenological	mechanistic

Timber state

- Time-series models: mostly phenomenological and linear (e.g. ARIMA, spectral/wavelet analyses)
- Biomath models: mostly mechanistic and nonlinear (e.g. Lotka-Volterra, SIR, Fitzhugh-Nagumo)

Time	Population state	Randomness	Model
discrete	discrete	deterministic	*Henson et al. (2001)
discrete	discrete	stochastic	
discrete	continuous	deterministic	
discrete	continuous	stochastic	random walk
continuous	discrete	deterministic	weird
continuous	discrete	stochastic	branching process /
			cont. Markov
continuous	continuous	deterministic	ODE
continuous	continuous	stochastic	stochastic ODE

Overview

#### Process and measurement error

- For stochastic models need to define both a process model and an observation model (= measurement model)

  Process model  $Y(t+1) \sim F(Y(t))$ Measurement model  $Y_{\text{obs}}(t) \sim Y(t)$
- Only process error affects the future dynamics of the process (usually)
- Might decompose process model into a deterministic model fo the expectation and (additive?) noise around the expectation: e.g.  $Y(t) = \mu + \epsilon$ ,  $Y(t) \sim \text{Poisson}(\exp(\eta))$

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- Process error induces dynamic changes in variance
- Process+observation error induce correlations between subsequent observations
- Observation at next time step depends on unobserved value at current time step
- Simple statistical methods

   (i.e. uncorrelated, equal variance)

   are incorrect

#### Consequences

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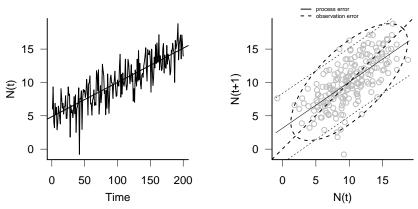
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#### Linear example



How should we interpret this single realization?

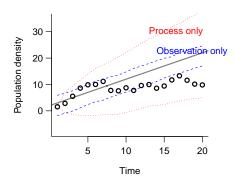
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#### Why simulate?

- understand dynamics
- test methods in best-case scenario
- explore precision/power
- quantify properties of statistical estimators
- evaluate robustness

#### Linear model



$$egin{aligned} extstyle N(1) &= a \ extstyle N(t+1) &\sim extstyle extstyle Normal(N(t)+b, \sigma_{ extstyle proc}^2) \ extstyle N_{ extstyle obs}(t) &\sim extstyle extstyle Normal(N(t), \sigma_{ extstyle obs}^2) \end{aligned}$$

#### R code (version 1)

```
## set up parameters etc.
nt <- 20: a <- 6: b <- 1
sd_proc <- sqrt(2)
sd_obs <- sqrt(2)
N <- Nobs <- numeric(nt)
set.seed(101) ## for reproducibility
## actual model
N[1] \leftarrow a
Nobs[1] \leftarrow rnorm(1,N[1],sd_obs)
for (i in 1:nt) {
  N[i+1] <- rnorm(1,N[i]+b,sd_proc)</pre>
  Nobs[i+1] \leftarrow rnorm(1,N[i+1],sd_proc)
```

00000000

```
library(deSolve)
linfun <- function(t,y,parms) {</pre>
  g <- with(as.list(c(y,parms)), { ## magic to use param n
     N_new <- rnorm(1,N+b,sd_proc)</pre>
     c(N=N_new, Nobs=rnorm(1, N_new, sd_obs))
  })
  return(list(g)) ## deSolve needs this format
}
set.seed(101)
NO \leftarrow c(N=a, Nobs=rnorm(1, a, sd_obs))
linparms <- c(a=6,b=1,sd_proc=sd_proc,sd_obs=sd_obs)</pre>
ode(N0,1:nt,linfun,linparms,method="iteration")
```

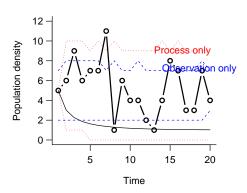
#### R code (version 3)

Overview

For this particular example, we can cheat because the process error doesn't affect the future dynamics — it just accumulates:

```
N_det <- a+b*(0:(nt-1))
set.seed(101) ## for reproducibility
proc_noise <- rnorm(nt-1,mean=0,sd=sd_proc)
N <- N_det+cumsum(c(0,proc_noise))
N_obs <- rnorm(nt,mean=N,sd=sd_obs)</pre>
```

#### Hyperbolic nonlinear model



$$egin{aligned} & extstyle N(1) = extstyle N_0 \ & N(t+1) \sim \mathsf{Poisson}\left(rac{aN(t)}{b+N(t)}
ight) \ & N_{\mathsf{obs}}(t) \sim \mathsf{Binomial}(N(t), p) \end{aligned}$$

time N N\_obs

#### R code

```
hypfun <- function(t,y,parms) {</pre>
    g <- with(as.list(c(y,parms)), { ## magic to use param
         N \det <- a*N/(b+N)
         N <- rpois(1,lambda=N_det)</pre>
         N_obs <- rbinom(1,size=N,prob=prob_obs)</pre>
         c(N=N.Nobs=N obs)
  })
  return(list(as.numeric(g))) ## deSolve needs this format
}
set.seed(101)
NO \leftarrow c(N=4, N \text{ obs}=4)
hypparms \leftarrow c(a=6,b=1,prob_obs=0.9)
ode(NO, times=1:nt, func=hypfun, parms=hypparms, method="iterat:
```

References

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Overview

#### tochastic ODES

- continuous-time, continuous-state
- ordinary differential equations plus a Wiener process
   (= derivative of a Brownian motion)
- delicate analysis (For biologists: Turelli (1977); Roughgarden (1995). For mathematicians: Øksendal (2003))
- More common for cellular/physiological than population models
- Solve via Euler-Maruyama
   (= Euler + appropriately scaled Gaussian noise)

#### Markov processes

- continuous-time, discrete-state
- specify (limits of) probabilities of transitions per unit time, e.g.  $P(N \to N+1)$  in the interval (t, t+dt) is rN(t) dt
- Even harder than SDEs to analyze rigorously . . .
- But computationally straightforward: Gillespie algorithm and variations (Gillespie, 2007): exponentially distributed time between transitions

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#### Trajectory matching

- Easiest: simulate the deterministic version of the model (i.e., with neither observation nor process error) and compare
- Because measurement error is (typically) independent at each observation, overall log-likelihood is sum of log-likelihood
- i.e.
- for Normally distributed, equal-variance error, maximum likelihood estimation equivalent to least-squares
- Very common for ODE models, e.g. Gani and Leach (2001);
   van Veen et al. (2005)
- Brute force can be slow/unstable: use sensitivity equations
   Raue et al. (2013)

#### Pseudo-code

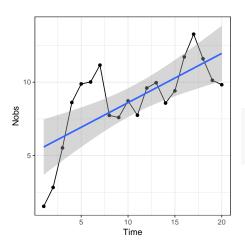
```
## deterministic dynamics:
## function of parameters, possibly including ICs
determ_fun <- function(determ_params) {</pre>
    ## code...
## objective function (neg. log-likelihood, SSQ, ...)
## 'params' includes process and observation parameters
obj_fun <- function(params,data) {
  estimate <- determ_fun(params[determ_params]))</pre>
  obj <- likfun(estimate,data,params[obs_params])
  return(obj)
find_minimum(obj_fun,starting_params,...)
```

#### Real code (using for loops)

```
determ_fun <- function(p,nt) {
   with(as.list(p),a+b*(1:nt))
}
obj_fun <- function(p,nt,Nobs) {
   estimate <- determ_fun(p[c("a","b")],nt)
   ## negative log-lik. of Normal
   obj <- -sum(dnorm(Nobs,estimate,p["sd"],log=TRUE))
   return(obj)
}
optim(fn=obj_fun,par=c(a=5,b=2,sd=1),nt=20,Nobs=linN)</pre>
```

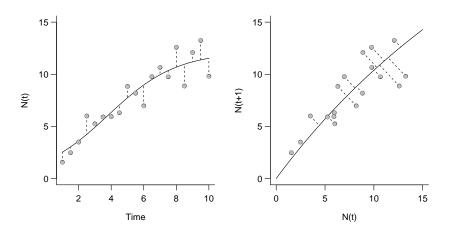
#### Real code (using mle2())

mle2() simplifies computation of confidence intervals, likelihood profiles, predicted values, etc..



linDF <- data.frame(Time=1:nt,</pre> Nobs=linN) lm(Nobs~Time,data=linDF)

#### Logistic model fit



- Next-easiest approach: assume only process error (no measurement error)
- N(t+1) depends only on N(t) (which we know exactly): conditional independence
- One-step-ahead prediction
- Simple for discrete-time models (we need to specify  $N(t+1) \sim N(t)$  anyway)
- More complicated for continuous-time models (Ellner et al., 2002): fit a smooth curve to data, then fit to derivatives of the curve

find\_minimum(obj\_fun,starting\_params,...)

}

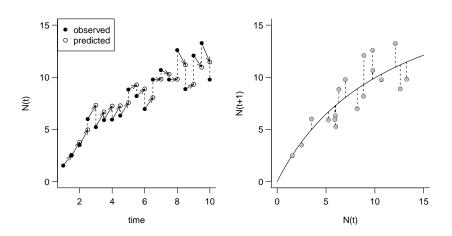
return(sum(obj))

Overview

```
## deterministic dynamics:
## function of parameters and previous values
onestep_fun <- function(determ_params,Nt) { ... }</pre>
## objective function (neg. log-likelihood, SSQ, ...)
obj_fun <- function(params,data) {</pre>
  obj <- ... ## numeric vector of length (nt-1)
 for (i in 1:(nt-1)) {
     estimate <- onestep_fun(N[i],params[determ_params])</pre>
```

obj[i] <- fun(estimate, N[i+1], params[obs\_params])

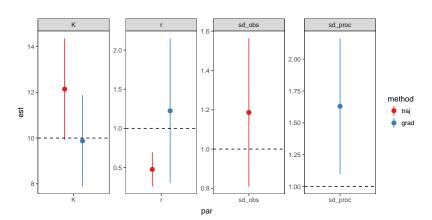
#### Logistic growth fit



#### How can we use these?

- Try both and hope the answers are not importantly different . . .
- Use biological knowledge of whether process ≫ observation error or vice versa

#### Logistic fit comparisons



#### Outline

- Fancier methods
  - SIMEX
  - Kalman filter

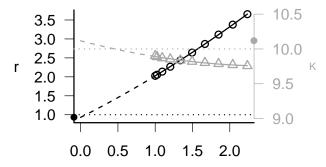
#### SIMulation-EXtrapolation method

- Requires (1) an independent estimate of the observation error;
   (2) that we can sensibly add additional observation error to the data
- Slightly easier for Normal errors
- Probably most sensible for experimental data?
- Examples: Ellner et al. (2002); Melbourne and Chesson (2006)

## Procedure

- based on estimated observation error, pick a range of increased error values, e.g. tripling the existing observation variance in 4–8 steps
- for each error magnitude, generate a data set with that increased error (more stable to inflate a single set of errors)
- estimate parameters for each set using gradient matching (i.e. assume  $\sigma_{\rm obs}^2=0$ )
- fit a linear or quadratic regression model for parameter = f(total error)
- extrapolate the fit to zero

# Logistic fit



Total observation error

## Kalman filter

- General approach to account for dynamic variance, expected population state
- Works for linear (typically Normal) models; can be extended to nonlinear models
- Natural multivariate extensions: include bias, external shocks, etc. (Schnute, 1994)

#### Concept

- Variance increases with process error; decreases with (accurate) observations
- Expected population state follows expected dynamics; drawn toward (accurate) observations
- Procedure (pseudo-pseudo-code)
  - Run KF for specified values of parameters,  $\sigma^2_{\rm obs}$ ,  $\sigma^2_{\rm proc}$  to compute  $\hat{N}(t)$ ,  $\sigma^2_N(t)$
  - Estimate objective function (SSQ) for  $N_{\text{obs}}|\hat{N}, \sigma_{N}^{2}$
  - Minimize over {parameters,  $\sigma_{\rm obs}^2, \sigma_{\rm proc}^2$ }

## Concept and implementation

#### Concept

- Variance increases with process error; decreases with (accurate) observations
- Expected population state follows expected dynamics; drawn toward (accurate) observations
- Procedure (pseudo-pseudo-code)
  - Run KF for specified values of parameters,  $\sigma_{\rm obs}^2$ ,  $\sigma_{\rm proc}^2$  to compute  $\hat{N}(t)$ ,  $\sigma_N^2(t)$
  - Estimate objective function (SSQ) for  $N_{\text{obs}}|\hat{N}, \sigma_N^2$
  - Minimize over {parameters,  $\sigma_{obs}^2$ ,  $\sigma_{proc}^2$ }

$$extstyle{N(t)} \sim extstyle{Normal(a+bN(t-1), \sigma^2_{ extstyle{proc}})} \ extstyle{N_{ extstyle{Obs}}(t)} \sim extstyle{Normal(N((t), \sigma^2_{ extstyle{obs}}))}$$

- $b < 1, a > 0 \rightarrow \text{stable dynamics}$
- ullet  $b>1 
  ightarrow {
  m exponential}$  growth

#### Procedure

① Update mean, variance of true density according to previous expected mean and variance:

$$ext{mean}(N(t)|N_{ ext{obs}}(t-1)) \equiv \mu_1 = a + b\mu_0$$
 $ext{Var}(N(t)|N_{ ext{obs}}(t-1)) \equiv \sigma_1^2 = b^2\sigma_0^2 + \sigma_{ ext{proc}}^2$ 

$$\begin{split} \text{mean}(\textit{N}_{\text{obs}}(t)|\textit{N}_{\text{obs}}(t-1)) &\equiv \mu_2 = \mu_1 \\ \text{Var}(\textit{N}_{\text{obs}}(t)|\textit{N}_{\text{obs}}(t-1)) &\equiv \sigma_2^2 = \sigma_1^2 + \sigma_{\text{obs}}^2 \end{split}$$

Now update true (expected) mean and variance to account for current observation:

$$\begin{split} \text{mean}(\textit{N}|\textit{N}_{\text{obs}}(t)) &\equiv \mu_3 = \mu_1 + \frac{\sigma_1^2}{\sigma_2^2}(\textit{N}_{\text{obs}}(t) - \mu_2) \\ \text{Var}(\textit{N}(t)|\textit{N}_{\text{obs}}(t)) &\equiv \sigma_3^2 = \sigma_1^2 \left(1 - \frac{\sigma_1^2}{\sigma_2^2}\right) \end{split}$$

#### Pseudo-code

```
KFpred <- function(params, var_proc, var_obs, init) {</pre>
  set_initial_values
  for (i in 2:nt) {
     ## ... calculate mu\{1-3\}, sigma^2\{1-3\} as above
     N[i] <- mu_3; Var[i] <- sigmasq_3</pre>
  }
  return(list(N=N, Var=Var))
KFobj <- function(params, var_proc, var_obs, init, Nobs) {</pre>
    pred <- KFpred(params, var_proc, var_obs, init)</pre>
    obj_fun(Nobs,mean=pred$N,sd=sqrt(pred$Var))
}
minimize(KFobj,start_values,Nobs)
```

## Extended Kalman filter

Overview

To fit (mildly) nonlinear models with the deterministic skeleton

$$N(t+1)=f(N(t)),$$

we just replace a and b in the autoregressive model N(t+1)=a+bN(t) with the coefficients of the first two terms of the Taylor expansion of f():

$$f(N(\tau)) \approx f(N(t)) + \frac{df}{dN}(N(\tau) - N(t)) + \dots$$

# Multivariate extension (Schnute, 1994)

process: 
$$m{X}_t = m{A}_t + m{B}_t m{X}_{t-1} + m{\delta}_t$$
 observation:  $m{Y}_t = m{C}_t + m{D}_t m{X}_t + m{\epsilon}_t$ 

Allows for bias, cross-species effects in both process and observation, correlation in process and observation noise . . .

### References

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