A Reasoned Basis for Inference: Fun with Fisher

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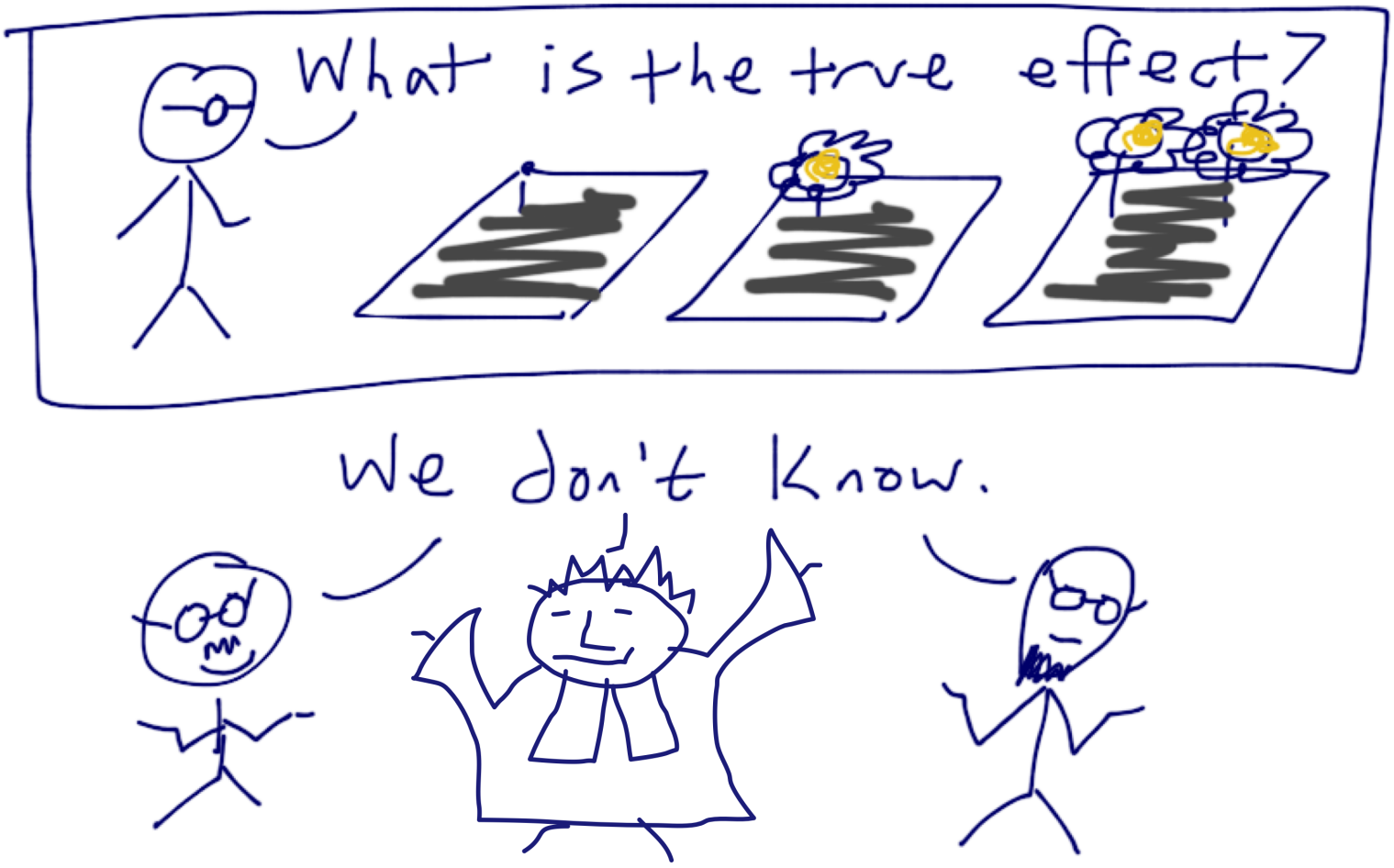
Overview Statistical Inference for Causal Quantities

What is the true effect of the treatment assignment?



What is the true effect of the treatment assignment?

I don't know the truth, but I can provide a good guess of the average causal effect.

i	Z_i	Y_i	y_{i1}	y_{i0}
A	0	16	?	16
B	1	22	22	?
C	0	7	?	7
D	1	14	14	?
			$\bar{y}_{i1}$	$\bar{y}_{i0}$

$$ATE = \bar{Y}_i|Z_i=1 - \bar{Y}_i|Z_i=0$$

$$= \frac{22+14}{2} - \frac{16+7}{2} = 6.5$$

What is the true effect of the treatment assignment?

I dew nut knew thee truth,
but, given pryors, I cane
predikte itf
probabeeleetee.

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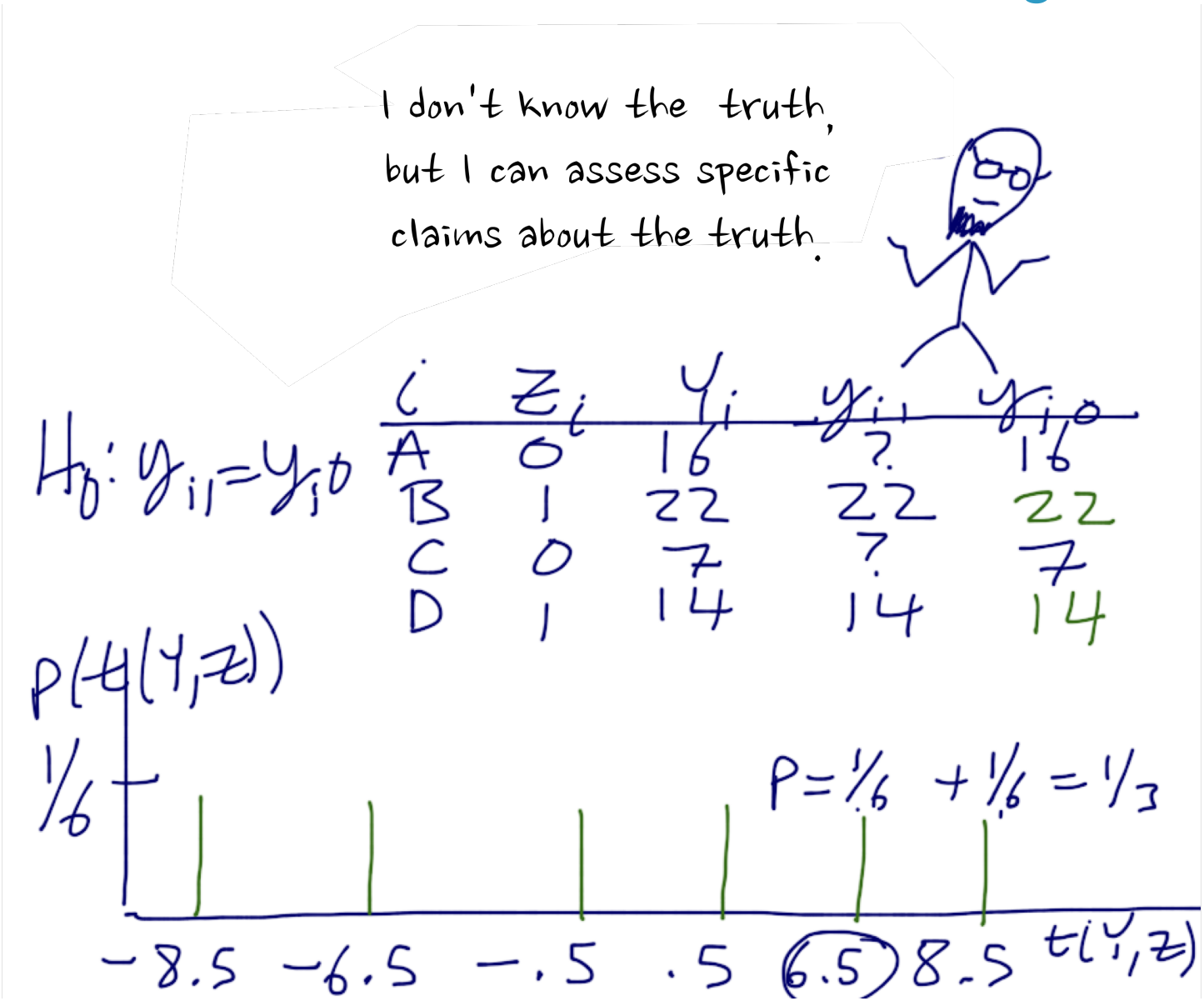


Testing the Sharp Null of No Effects

```
Z <- c(0,1,0,1)
Y <- c(16,22,7,14)
Om <- matrix(0,ncol=choose(4,2),nrow=length(Z))
whotrted <- combn(1:4,2)
for(i in 1:choose(4,2)){ Om[cbind(whotrted[,i],i)]<-1 }
meandifftz <- function(y,z){ mean(y[z==1]) - mean(y[z==0]) }
thedist<-apply(Om,2, function(z){ meandifftz(Y,z) })
rbind(Om,thedist)

        [,1] [,2] [,3] [,4] [,5] [,6]
         1.0  1.0  1.0  0.0  0.0  0.0
         1.0  0.0  0.0  1.0  1.0  0.0
         0.0  1.0  0.0  1.0  0.0  1.0
         0.0  0.0  1.0  0.0  1.0  1.0
thedist  8.5 -6.5  0.5 -0.5  6.5 -8.5

table(thedist)

thedist
-8.5 -6.5 -0.5  0.5  6.5  8.5
   1    1    1    1    1    1

theobs <- meandifftz(Y,Z)
mean(thedist >= theobs)

[1] 0.3333333
```

The Newspapers Study

City	Pair	Treatment	Turnout Baseline	Turnout Outcome	Newspaper
Saginaw	1	0	17	16	
Sioux City	1	1	21	22	Sioux City Journal
Battle Creek	2	0	13	14	
Midland	2	1	12	7	Midland Daily News
Oxford	3	0	26	23	
Lowell	3	1	25	27	Lowell Sun
Yakima	4	0	48	58	
Richland	4	1	41	61	Tri-City Herald

Table 1: Design and outcomes in the Newspapers Experiment. The Treatment column shows treatment with the newspaper ads as 1 and lack of treatment as 0. Panagopoulos (2006) provides more detail on the design of the experiment.

The Newspapers Study: $H_0 : y_{i1} = y_{i0}, p = 6/16.$

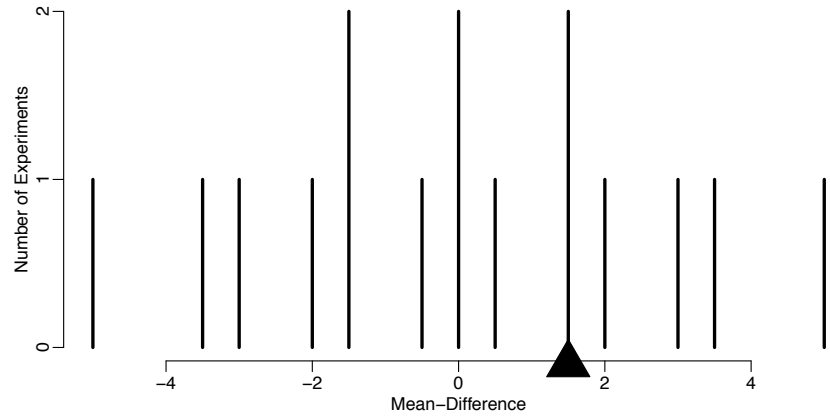


Figure 2: The randomization distribution of the mean-difference test statistic under the null hypothesis of no effects is shown by the tall black lines. The black triangle shows the value of the mean-difference observed in the actually fielded experiment.

What range of effects might be surprising?

Hypothesize a model of potential outcomes For $H_0 : y_{i1} = y_{i0} + \tau$, what τ might be surprising?

Map the model to observation via design What would $\tau = 6$ imply for what we observe? If $\tau = 6$ and $Y_i = Z_i y_{i1} + (1 - Z_i) y_{i0}$ then $Y_i - Z_i 6 = y_{i0}$.

Generate the randomization distribution of this hypothesis As before

Summarize information against the hypothesis For hypotheses of $\tau \geq 6$ we have $p \leq .125$ using a mean difference test statistic.

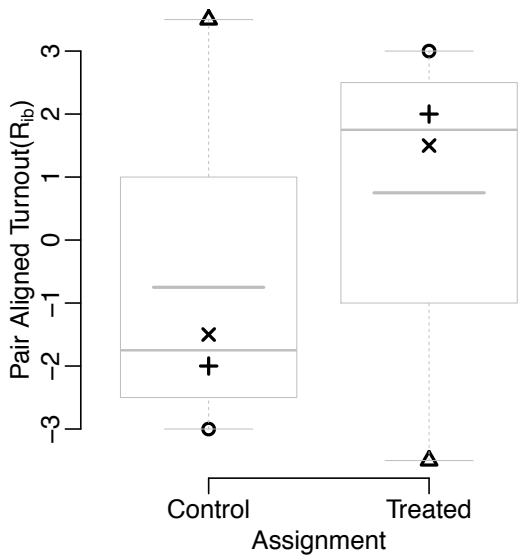


Figure 3: The distribution of outcomes in the control and treatment groups. The values of turnout are "pair aligned" or "pair mean centered" so that we can focus attention on the paired differences rather than on the differences in turnout levels between pairs.

What effects might be least surprising?

The idea of a "best guess" maps onto the Hodges-Lehmann point estimate: the hypotheses for which $E(t(Z, Y) = 0$: ex. the difference of means is zero. Here $\tau = 1.5$ for the mean difference and $\tau = 3.25$ for the rank-based test (which equalizes medians).

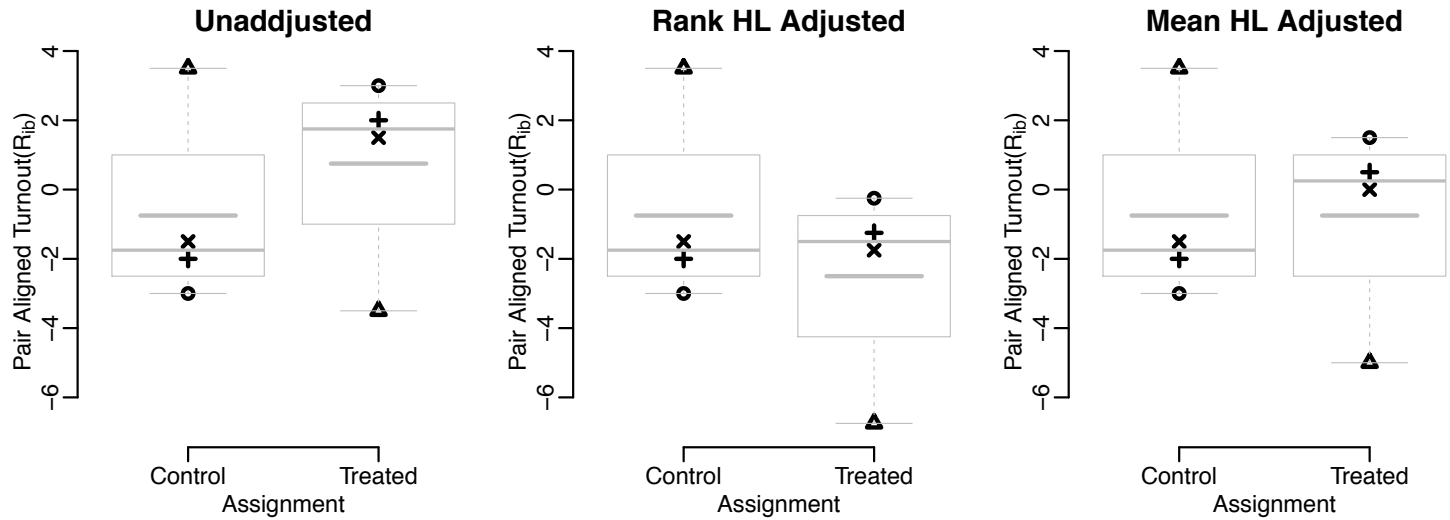
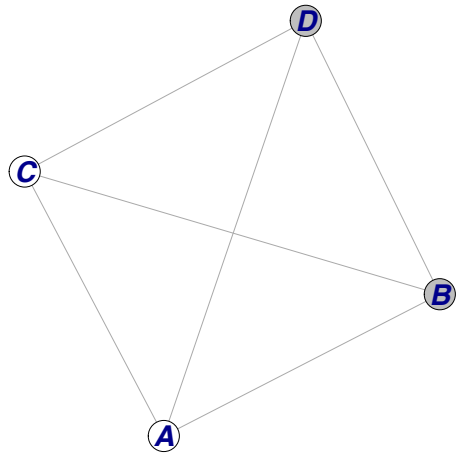


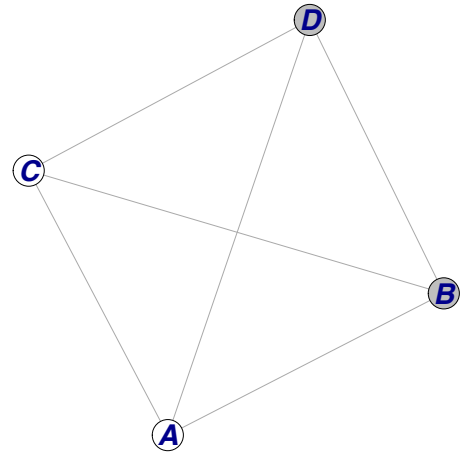
Figure 4: The unadjusted comparison of controls versus treated (left most plot) as compared to the results of adjusting the treatment group based on the Hodges–Lehmann (HL) point estimates derived from two different test statistics (Wilcoxon paired signed ranks and the mean difference). Notice that the control group remains the same and it is the hypotheses which imply changes in the distribution of the treatment group compared to the unadjusted, observed, outcomes.

Statistical inference for counterfactual quantities with interference?



i	Z_i	Y_i	$y_{i,1100}$	$y_{i,0101}$	$y_{i,1001}$	$y_{i,0110}$	$y_{i,1010}$	$y_{i,0011}$
A	0	16	?	16	?	?	?	?
B	1	22	?	22	?	?	?	?
C	0	7	?	7	?	?	?	?
D	1	14	?	14	?	?	?	?

Statistical inference for counterfactual quantities with interference?



i	Z_i	Y_i	$y_{i,1100}$	$y_{i,0101}$	$y_{i,1001}$	$y_{i,0110}$	$y_{i,1010}$	$y_{i,0011}$	$y_{i,0000} \equiv y_{i,\mathbf{0}}$
A	0	16	?	16	?	?	?	?	16
B	1	22	?	22	?	?	?	?	22
C	0	7	?	7	?	?	?	?	7
D	1	14	?	14	?	?	?	?	14

The sharp null of no effects is a model of no interference:

$H_0 : y_{i,1100} = y_{i,0101} = y_{i,1001} = y_{i,0110} = y_{i,1010} = y_{i,0011} = y_{i,0000}$,
$y_{i,\mathbf{0}} = \mathcal{H}(y_{i,\mathbf{z}}, \mathbf{0}) = y_{i,\mathbf{z}}$, $p = 0.33$.

Introducing the **uniformity trial** $\equiv \mathbf{y}_{i,0000}$ (Rosenbaum, 2007).

Theoretical models of potential outcomes can produce sharp hypotheses

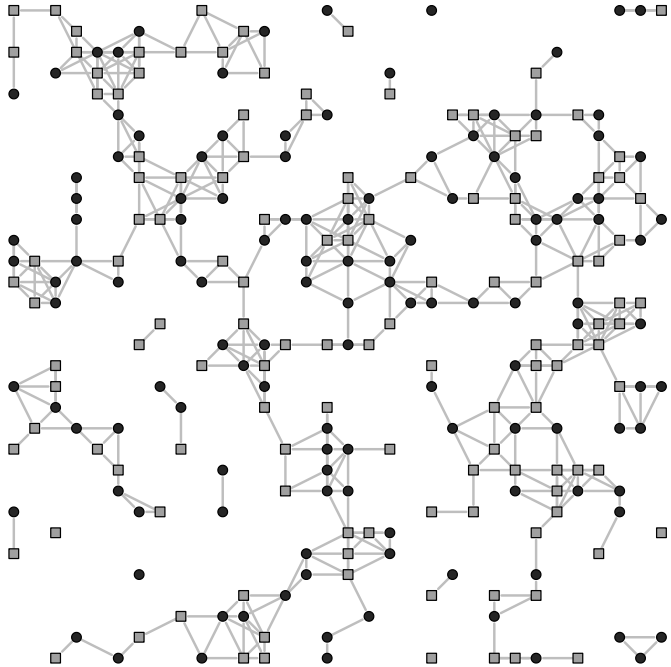


Figure: A simulated data set with 256 units and 512 connections. The $256/2 = 128$ treated units ($Z_i = 1$) are shown as filled circles and an equal number of control units ($Z_i = 0$) are shown as as gray squares.

Theoretical models of potential outcomes can produce sharp hypotheses

$$\mathcal{H}(\mathbf{y_0}, \mathbf{z}, \beta, \tau) = \left[\beta + (1 - z_i)(1 - \beta) \exp\left(-\tau^2 \mathbf{z}^T \mathbf{S}\right)\right] \mathbf{y_0} \quad (1)$$

$$\mathcal{H}(\mathbf{y_z}, \mathbf{0}, \beta, \tau) = \left[\beta + (1 - z_i)(1 - \beta) \exp\left(-\tau^2 \mathbf{z}^T \mathbf{S}\right)\right]^{-1} \mathbf{y_z} \equiv \mathbf{y_0} \quad (2)$$

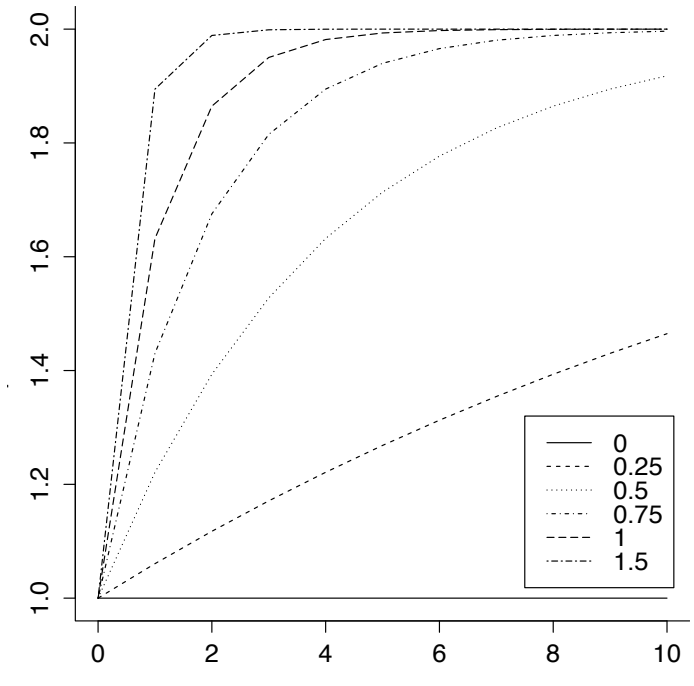


Figure: Growth curve of spillover effects for the expression $\beta + (1 - \beta) \exp\left(-\tau^2 \mathbf{z}^T \mathbf{S}\right)$ as the number of treated neighbors, $\mathbf{z}^T \mathbf{S}$, increases for $\beta = 2$ and a selection of τ values.

Theoretical models of potential outcomes can produce sharp hypotheses

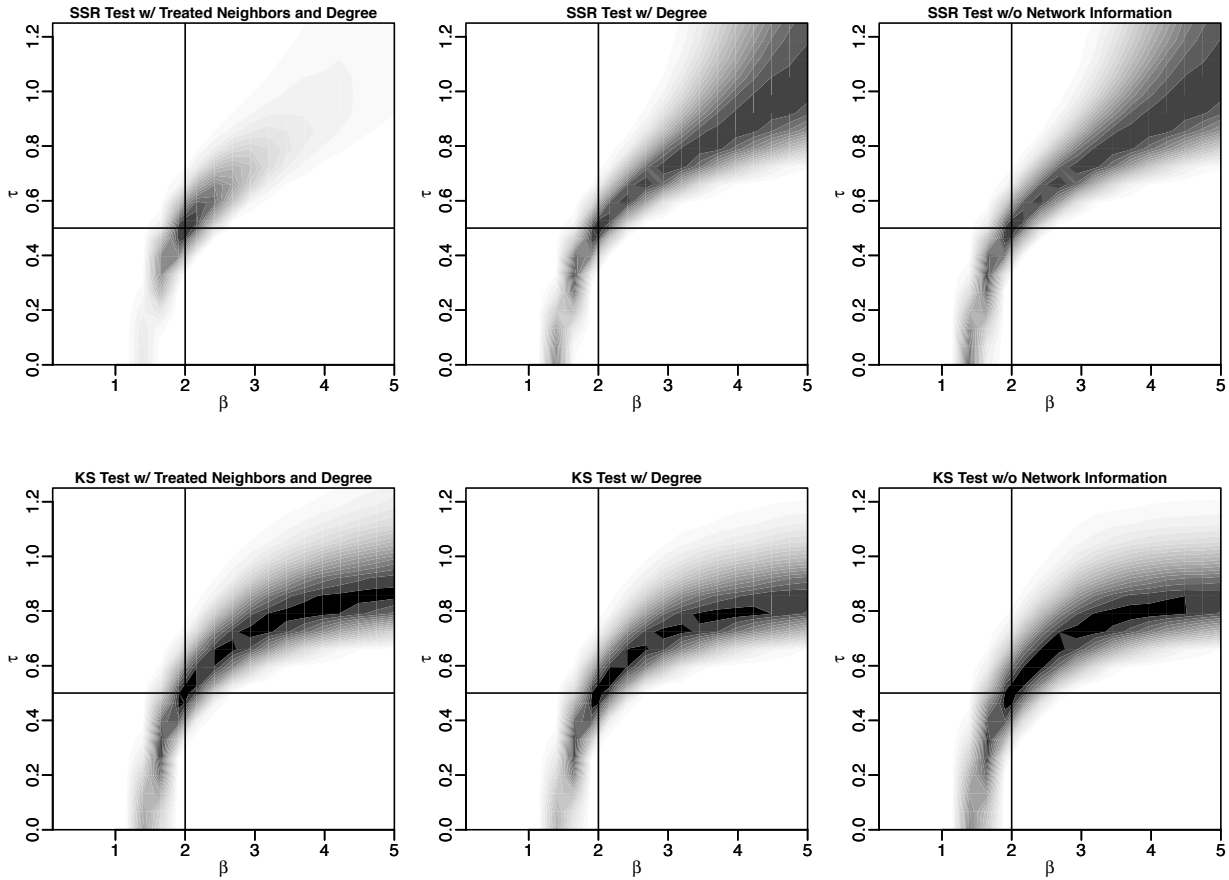


Figure: Proportion of *p*-values less than .05 for randomization tests of joint hypotheses about τ and β. Darker values mean less rejection. Truth is at $\tau = .5, \beta = 2$. All tests reject the truth no more than 5% of the time at $\alpha = .05$. All simulations, not Normal approximations.

A General Fisherian Inference Algorithm

1. Write a model ($\mathcal{H}(\mathbf{y_0}, \mathbf{z}, \theta)$) converting uniformity trial into observed data (i.e. a causal model).
2. Solve for $\mathbf{y_0}$: $\mathcal{H}(\mathbf{y_z}, \mathbf{0}, \theta_0) = \mathbf{y_0}$
3. Select a test statistic that is effect increasing in all relevant dimensions.
4. Compute p-values for substantively meaningful range of θ. Or calculate boundaries of regions.

Fisherian Tests work when asymptopia is out of reach

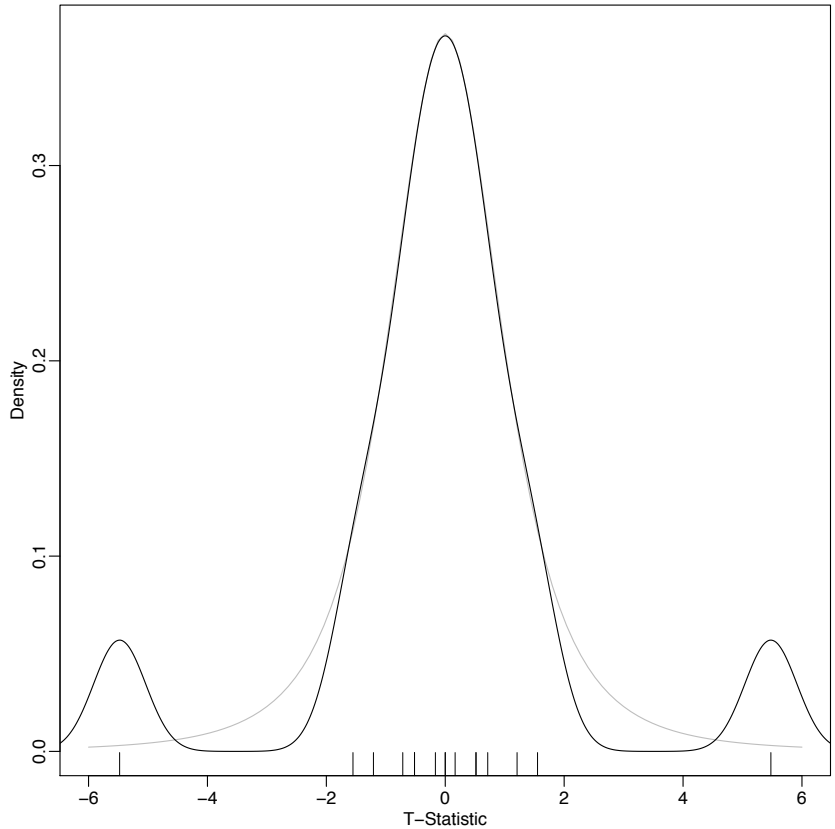


Figure 5: Under the null hypothesis of no effects, t-statistics calculated after repeating the experimental assignment process should be distributed around zero. The exact randomization distribution is tri-modal for the Newspapers study and does not match the t-distribution that we would expect under a larger sample.

Fisherian Tests work when asymptopia is out of reach

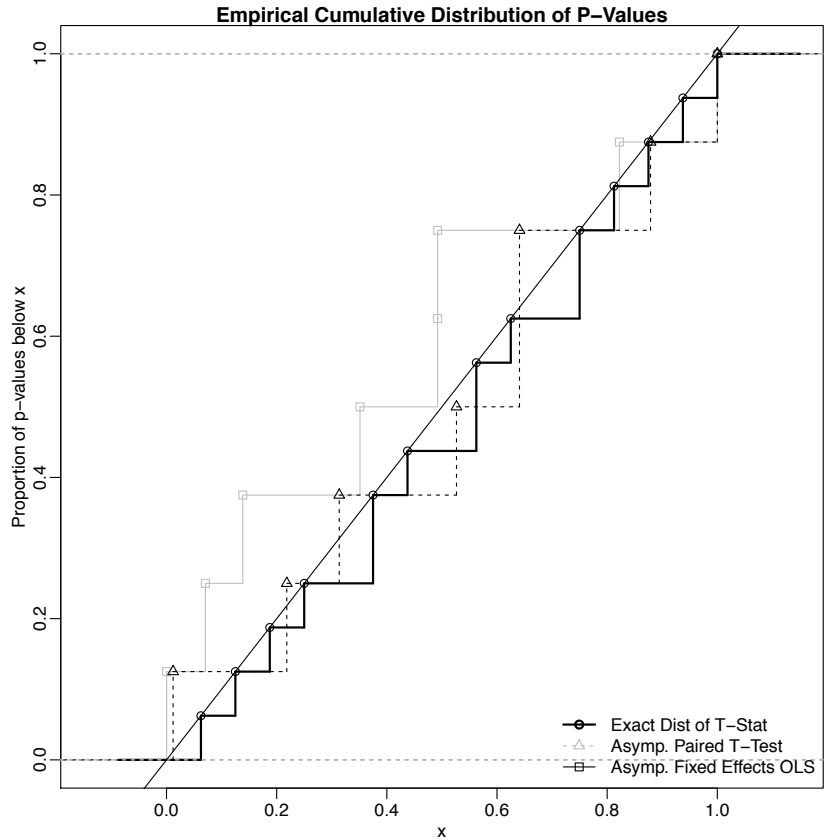


Figure 6: Cumulative distributions of p-values for testing the null hypothesis of no effects. A valid test would be at or below the diagonal line. Tests which would encourage rejection too quickly have points above the line.

Robust test statistics can increase power.

For the simple mean difference test statistic, we have $p = .375$, for a rank sum test (the sum of the ranks of the treated units), we have a $p = .4375$, and for an M-estimator based test (like mean-differences but with weights roughly inversely proportional to the Cook's D influence measure) we have $p = .3125$.

Cook's D	MM-Weights	City	Pair	Treatment	Turnout
0.13	1.00	Saginaw	1	0	16
0.13	1.00	Sioux City	1	1	22
0.48	0.94	Battle Creek	2	0	14
0.48	0.94	Midland	2	1	7
0.04	1.00	Oxford	3	0	23
0.04	1.00	Lowell	3	1	27
0.01	1.00	Yakima	4	0	58
0.01	1.00	Richland	4	1	61

Table 2: Not all cases have the same influence on the mean difference. Cook's D summarizes this influence — Battle Creek and Midland have disproportionate influence. A robust fit (details in the text) downweights Battle Creek and Midland.

Covariance adjusted tests can increase power.

Using the difference pre-vs-post as the outcome (comparing treated pre-vs-post with paired control pre-vs-post), and using the robust test statistic for $H_0 : y_{i1} = y_{i0}$ $p = 4/16 = .25$.
Using $e_i = (Y_i - Y_{i,t-1}) - (\hat{\beta}_0 + \hat{\beta}_1 \text{pop} + \hat{\beta}_2 \text{num candidates})$ and the robust test statistic to test $H_0 : y_{i1} = y_{i0}$ we have $p = 2/16 = .125$.

Key features of Fisher's approach

Flexible Any scientific model than can generate implications for all units' potential outcomes can, in principle, produce testable parameters.

Design based Requires knowledge of probability of Z not Y or $Y|X$ or $\beta|\gamma$.

Finite Sample Oriented Does not require asymptopia. Can use asymptopia when there for a visit.

Can be slow In between 8 cities and asymptopia is a land of many permutations.

Probably conservative Uses relatively little of the total information we have available about the science.

If you want to know more read Paul Rosenbaum's work The version of Fisher's approach I discuss here is built on work by Paul Rosenbaum. Read his work if you want to learn more.

New Questions

- How to choose test statistics for multidimensional sharp-hypothesis testing? Are there multi-dimensional “effect increasing” characteristics that we can assess for a given model?
- Are there general classes of scientific/counterfactual models?
- How should we interpret and display results?

On models

- Models are mathematical functions, multiple functions can have similar adjustments to the data.
- Assessing more than one model may enhance insight (Rosenbaum, 2010).
- When more than one model is plausible, what should you do?
- Our method can help eliminate the implausible, not accept the plausible.

New Questions

- Where do models of counterfactuals come from? Do we have advice about going from words to math?
- Math has its own logic. Some expressions for models may not be sensitive to changes in parameters. How can we assess what a given model is telling? How can we go from math to words before testing hypotheses?
- The KS-statistic is low powered for tail-differences. Recall that we are testing $t(\mathcal{H}(), \mathbf{z})$ not just $\mathcal{H}()$. Some results might tell us that our test is low powered against certain alternatives more than that we have identified a region of plausibility. How to find a better test statistic?
- How can this work learn from other modes of statistical inference and other representations of causal inference? What are the connections to ATE and other estimation frameworks (Spatial Econometrics, Network Analysis (ERGMs), etc...)?

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