

Ex6-1

Bayes' theorem: $p(A|B) = \frac{p(A,B)}{p(B)} = \frac{p(A,B)}{\int p(B|A)p(A)}$

probability: $\theta \rightarrow p(A)$

$$p(B) = \sum_j p(B|A_j) p(A_j)$$

likelihood: $p(A|\theta)$

$$p(A_i|B) = \frac{p(B|A_i) p(A_i)}{\sum_j p(B|A_j) p(A_j)}$$

θ : continuous

$$p(\theta|D) = \frac{p(D|\theta) p(\theta) \leftarrow \text{prior}}{\int p(D|\theta) p(\theta) d\theta} = \frac{p(D,\theta)}{\int p(D,\theta) d\theta}$$

\uparrow posterior \nwarrow likelihood \nearrow normalization

(a) $KL(Q(\theta) || p(\theta|D)) + 1$

$$= \int Q(\theta) \log \frac{Q(\theta)}{p(\theta|D)} d\theta + \int Q(\theta) \log \frac{p(\theta,D)}{Q(\theta)} d\theta$$

$$= \int Q(\theta) \cdot \log \frac{Q(\theta)}{p(\theta|D)} \cdot \frac{p(\theta|D) p(D)}{Q(\theta)} d\theta$$

$$= \int Q(\theta) \cdot \log p(D) d\theta$$

$$= \log p(D) \int Q(\theta) d\theta = \log p(D)$$

(b) lower bound:

Let (A, \leq) is a partial ordered set.

z is a lower bound if $\exists z \in A$.

s.t. $\forall x \in B \subseteq A, x \geq z$.

$$KL(Q(\theta) \parallel P(\theta|D)) = \int Q(\theta) \cdot \log \frac{Q(\theta)}{P(\theta|D)} d\theta$$

Gibbs' inequality:

$$\begin{aligned} P &= \{p_1, \dots, p_n\} \\ Q &= \{q_1, \dots, q_n\} \end{aligned} \quad - \sum_{i=1}^n p_i \log_2 p_i \leq - \sum_{i=1}^n p_i \log_2 q_i$$

$$KL(Q(\theta) \parallel P(\theta|D)) \geq 0$$

$$KL + L = \log p(D)$$

$$L = \log p(D) - KL \leq \log p(D) \Rightarrow L \text{ is a lower bound.}$$

with $KL(p \parallel Q)$ zero ~~iff~~ iff $p = Q$ almost everywhere.
a.e.

Ex 6-2

$$(a) \quad p(D|\theta) = \prod_{i=1}^n p(d_i | \theta, \sigma^2)$$

$$= \prod_{i=1}^n \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{(d_i - \theta)^2}{2\sigma^2} \right\}$$

$$= \frac{1}{(2\pi\sigma^2)^{n/2}} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{i=1}^n (d_i - \theta)^2 \right\}$$

$$(b) \quad p(\theta|D) = N(\mu_m, \sigma_m^2)$$

$$\mu_m = \frac{\sigma^2}{n\sigma_p^2 + \sigma^2} \mu_p + \frac{n\sigma_p^2}{n\sigma_p^2 + \sigma^2} \left(\frac{1}{n} \sum_{i=1}^n d_i \right)$$

$$\frac{1}{\sigma_m^2} = \frac{1}{\sigma_p^2} + \frac{n}{\sigma^2}$$

$$\mu_p = 0, \sigma_p = 1. \Rightarrow \frac{1}{\sigma_m^2} = \frac{1}{1} + \frac{n}{1} = n+1$$

$$\sigma^2 = 1$$

$$\sigma_m^2 = \frac{1}{n+1}$$

$$\mu_m = \frac{1}{1+1} \cdot 0 + \frac{n}{n+1} \left(\frac{1}{n} \sum_{i=1}^n d_i \right) = \frac{1}{n+1} \sum_{i=1}^n d_i$$

$$p(\theta|D) = N \left(\frac{1}{n+1} \sum_{i=1}^n d_i, \frac{1}{n+1} \right)$$

(c)

$$p(\theta|D) = \frac{p(D|\theta) p(\theta)}{p(D)}$$

$$= \frac{\prod_{i=1}^n N(\theta, \sigma^2) \cdot N(\mu_p, \sigma_p^2)}{\int \prod_{i=1}^n N(\theta, \sigma^2) \cdot N(\mu_p, \sigma_p^2) d\theta}.$$

If $p(\theta)$ changes, then $p(\theta)$'s conjugated prior might not have an analytical solution.

Gaussian likelihood is self conjugated.

conjugated distribution: prior dist as same as posterior dist.

conjugated priors leads to closed-form expression for the posterior.