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# MLE from a Competing Risks Model of an Exponential Failure and a Lognormal Survival

Nikola Chochkov, MSc Statistics, Humboldt University Berlin

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## 1 Introduction.

### 1.1 Assumptions.

Two usage-measured lifetime random variables are discussed in a fixed-time life test. The items under study are considered to accumulate usage independently and with a different rate. The survival ssspopulation is considered *Lognormally distributed* (with parameters  $\mu$  and  $\sigma$ ), while the failure popoulation - *Exponential* (with parameter  $\lambda$ ).

An estimation of the parameter vector  $(\lambda, \mu, \sigma)'$  is sought using all usage data collected (from both survival and failure cases)

- $\eta \sim \text{Lognormal}$
- $\psi \sim \text{Exponential}$
- $f_\eta(x) = \frac{1}{x\sigma\sqrt{2\pi}}e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}} = \frac{1}{x\sigma}\phi\left(\frac{\ln x - \mu}{\sigma}\right)$  is the *probability densitiy function* of the Lognormal distribution, with  $\phi$  being the density function of the Standard Normal Distribution
- $\bar{F}_\eta(x) = 1 - \frac{1}{\sqrt{2\pi}}e^{-\frac{(\ln x - \mu)^2}{2\sigma^2}} = \bar{\Phi}\left(\frac{\ln x - \mu}{\sigma}\right)$  is its *survival function*, with  $\bar{\Phi}$  being the survival function of the Standard Normal Distribution
- $f_\psi(x) = \lambda e^{-\lambda x}$  is the *pdf* of the Exponential distribution
- $\bar{F}_\psi(x) = e^{-\lambda x}$  is its *survival function*
- $N_f = \{i : x_i \text{ is failed}\}$  and  $n_f = \#N_f$  (i.e. number of failures)
- $N_s = \{i : x_i \text{ is survived}\}$  and  $n_s = \#N_s$  (i.e. number of survivals)
- $n$  is the total number of units under test ( $n = n_f + n_s$ )
- $\mathbf{x}$  is the data vector and  $\mathbf{x} = [\mathbf{x}^f, \mathbf{x}^s] = [x_1^f, \dots, x_{n_f}^f, x_1^s, \dots, x_{n_s}^s]$ , where:
- $\mathbf{x}^f$  is the data for failed items,  $\mathbf{x}^f = [x_1^f, \dots, x_{n_f}^f]$
- $\mathbf{x}^s$  is the data for survived items,  $\mathbf{x}^s = [x_1^s, \dots, x_{n_s}^s]$
- $\lambda, \mu, \sigma, x > 0$

## 1.2 Maximum Likelihood Method.

The Likelihood function is given by (since  $\eta$  and  $\psi$  are independent):

$$L(\lambda, \mu, \sigma | \mathbf{x}) = \prod_{i \in N_f} [f_\psi(x_i^f) \bar{F}_\eta(x_i^f)] \prod_{i \in N_s} [f_\eta(x_i^s) \bar{F}_\psi(x_i^s)] \quad (1)$$

Now from the above we can derive the Log Likelihood:

$$\text{Log} L = L^*(\lambda, \mu, \sigma | \mathbf{x}) = \sum_{i \in N_f} \ln [f_\psi(x_i^f)] + \sum_{i \in N_f} \ln [\bar{F}_\eta(x_i^f)] + \sum_{i \in N_s} \ln [f_\eta(x_i^s)] + \sum_{i \in N_s} \ln [\bar{F}_\psi(x_i^s)] \quad (2)$$

And if we apply the above notation and rework:

$$L^*(\lambda, \mu, \sigma | \mathbf{x}) = N_f \ln \lambda - \lambda \sum_{i=1}^n x_i + \sum_{i \in N_f} \ln \bar{\Phi} \left( \frac{\ln x_i^f - \mu}{\sigma} \right) - \sum_{i \in N_s} \ln \sigma x_i^s + \sum_{i \in N_s} \ln \phi \left( \frac{\ln x_i^s - \mu}{\sigma} \right) \quad (3)$$

The MLE estimators  $(\hat{\lambda}, \hat{\mu}, \hat{\sigma})$  would be the ones that turn  $L^*$  into a minimum, so now we need to compute the *Gradient* vector and *Hessian* matrix in order to find them. Furthermore we would like to show that the estimators have the *consistency* property.

## 1.3 First order derivatives of $L^*$ w.r.t $(\lambda, \mu, \sigma)$

Let's denote:  $z_i^{f,s} = \frac{(\ln x_i^{f,s} - \mu)}{\sigma}$ . Then we derive:

$$\frac{\partial z}{\partial \mu} = -\frac{1}{\sigma}, \frac{\partial z}{\partial \sigma} = -\frac{1}{\sigma} z, \frac{\partial \phi}{\partial \sigma} = \frac{1}{\sigma} \phi z^2, \frac{\partial \phi}{\partial \mu} = \frac{1}{\sigma} \phi z \quad (4)$$

$$\frac{\partial L^*(\lambda, \mu, \sigma | \mathbf{x})}{\partial \lambda} = \frac{n_f}{\lambda} - \sum_{i=1}^n x_i \quad (5)$$

$$\frac{\partial L^*(\lambda, \mu, \sigma | \mathbf{x})}{\partial \mu} = \frac{1}{\sigma} \sum_{i \in N_f} \frac{\phi(z_i^f)}{\bar{\Phi}(z_i^f)} + \frac{1}{\sigma} \sum_{i \in N_s} z_i^s \quad (6)$$

$$\frac{\partial L^*(\lambda, \mu, \sigma | \mathbf{x})}{\partial \sigma} = \frac{1}{\sigma} \sum_{i \in N_f} \frac{\phi(z_i^f) z_i^f}{\bar{\Phi}(z_i^f)} + \frac{1}{\sigma} \sum_{i \in N_s} (z_i^s)^2 - \frac{n_s}{\sigma} \quad (7)$$

## 1.4 Second order derivatives of $L^*$ w.r.t $(\lambda, \mu, \sigma)$

Let's denote again:  $z_i^{f,s} = \frac{(\ln x_i^{f,s} - \mu)}{\sigma}$ . Then we derive:

$$\frac{\partial^2 L^*(\lambda, \mu, \sigma | \mathbf{x})}{\partial \lambda^2} = -\frac{n_f}{\lambda^2} \quad (8)$$

$$\frac{\partial^2 L^*(\lambda, \mu, \sigma | \mathbf{x})}{\partial \mu^2} = \frac{1}{\sigma^2} \sum_{i \in N_f} \left[ \frac{\phi(z_i^f)}{\bar{\Phi}(z_i^f)} \left( 1 + \frac{\phi(z_i^f)}{\bar{\Phi}(z_i^f)} \right) z_i^f \right] - \frac{1}{\sigma^2} \quad (9)$$

$$\frac{\partial^2 L^*(\lambda, \mu, \sigma | \mathbf{x})}{\partial \mu \partial \sigma} = \frac{1}{\sigma^2} \sum_{i \in N_f} \frac{\phi(z_i^f)}{\bar{\Phi}(z_i^f)} + \frac{1}{\sigma^2} \sum_{i \in N_f} \left[ \frac{\phi(z_i^f)}{\bar{\Phi}(z_i^f)} \left( 1 + \frac{\phi(z_i^f)}{\bar{\Phi}(z_i^f)} \right) z_i^f \right] - \frac{2}{\sigma^2} \sum_{i \in N_s} z_i^s \quad (10)$$

$$\frac{\partial^2 L^*(\lambda, \mu, \sigma | \mathbf{x})}{\partial \sigma^2} = \frac{1}{\sigma^2} \sum_{i \in N_f} \left[ \frac{\phi(z_i) z_i^3}{\bar{\Phi}} - 2 \frac{\phi(z_i) z_i}{\bar{\Phi}} - \frac{\phi^2(z_i) z_i^2}{\bar{\Phi}^2} \right] - \frac{3}{\sigma^2} \sum_{i \in N_s} z_i^2 + \frac{n_s}{\sigma^2} \quad (11)$$

## **2 Maximum Likelihood Estimators' properties.**

### **2.1 Normality of the estimators.**

This can be seen from the results from 4 simulation sets of true value parameters. 1000 items under study, 150 repetitions for each parameter set. Diagram 1.

### **2.2 Consistency of the estimators.**

Simulations were done with the same parameter sets but with 2000 items under study, 150 simulation repetitions. Diagram 2.

## **3 Next steps.**

Check the properties of the Variance Covariance matrix, Score, and perform more simulations on other true parameter sets and with more repetitions.