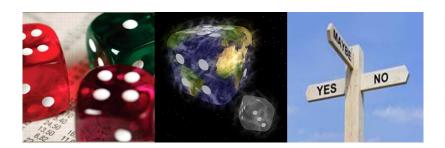
CHAPTER 6: Jointly Distributed Random Variables

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Introduction

- We want to study two or more random variables at same time.
- We want to describe the relation between them.
- We need to look at their joint distribution.

Joint Distribution Function

 Let X and Y be random variables. The joint cumulative distribution function (joint cdf) of X and Y is

$$F_{X,Y}(x,y) = P(X \le x, Y \le y).$$

- We can use $F_{X,Y}$ to calculate probability involve X and Y.
- Obtain single cdf of X from joint cdf:

$$F_X(x) = P(X \le x)$$

$$= P(X \le x, Y \le \infty)$$

$$= \lim_{y \to \infty} F_{X,Y}(x, y)$$

• Same for Y: $F_Y(y) = \lim_{x \to \infty} F_{X,Y}(x,y)$

Single cumulative distribution function

Example 1a

Evaluate P(X > x, Y > y).

$$P(X > x, Y > y) =$$

$$= 1 - P[(X > x, Y > y)^{c}]$$

$$= 1 - P[(X > x)^{c} \cup (Y > y)^{c}]$$

$$= 1 - P[(X \le x) \cup (Y \le y)]$$

$$= 1 - P(X \le x) - P(Y \le y + P(X \le x, Y \le y))$$

$$= 1 - F_{X}(x) - F_{Y}(y) + F_{X,Y}(x, y).$$

General case:

$$P(x_1 < X \le x_2, y_1 < Y \le y_2) = F(x_2, y_2) + F(x_1, y_1) - F(x_1, y_2) - F(x_2, y_1)$$

Prove it?

Discrete joint probability mass function (pmf)

- X and Y are discrete RVs
- The joint probability mass function (pmf) of X and Y p(x,y) = P(X = x, Y = y)
- Single pmf

$$p_X(x) = P(X = x) = \sum_{v} p(x, y)$$

$$p_Y(y) = P(Y = y) = \sum_{x} p(x, y)$$

Joint probability mass function (pmf)

Example 1b

15% of the families in a community have no children, 20% have 1 child, 35% have 2 children, and 30% have 3. In each family each child is equally likely (independently) to be a boy or a girl. If a family is chosen at random, find the joint pmf of B (the number of boys) and G (the number of girls) in this family.

$$P\{B = 0, G = 0\} = P\{\text{no children}\} = .15$$

$$P\{B = 0, G = 1\} = P\{1 \text{ girl and total of 1 child}\}$$

$$= P\{1 \text{ child}\}P\{1 \text{ girl}|1 \text{ child}\} = (.20)\left(\frac{1}{2}\right)$$

$$P\{B = 0, G = 2\} = P\{2 \text{ girls and total of 2 children}\}$$

$$= P\{2 \text{ children}\}P\{2 \text{ girls}|2 \text{ children}\} = (.35)\left(\frac{1}{2}\right)^2$$

TABLE 6.2: $P\{B = i, G = j\}$

j	0	1	2	3	Row sum = $P\{B = i\}$
0	.15	.10	.0875	.0375	.3750
1	.10	.175	.1125	0	.3875
2	.0875	.1125	0	0	.2000
3	.0375	0	0	0	.0375
$Columnsum = P\{G = j\}$.3750	.3875	.2000	.0375	

Joint probability mass function (pmf)

Example 1c

Suppose that 3 balls are randomly selected from an urn containing 3 red, 4 white, and 5 blue balls. If we let X and Y denote, respectively, the number of red and white balls chosen, find the joint probability mass function of X and Y.

$$p(0,1) = \binom{4}{1} \binom{5}{2} / \binom{12}{3} = \frac{40}{220}$$

$$p(0,2) = \binom{4}{2} \binom{5}{1} / \binom{12}{3} = \frac{30}{220}$$

$$p(0,3) = \binom{4}{3} / \binom{12}{3} = \frac{4}{220}$$

$$p(1,0) = \binom{3}{1} \binom{5}{2} / \binom{12}{3} = \frac{30}{220}$$

 $p(0,0) = \binom{5}{3} / \binom{12}{3} = \frac{10}{220}$

$$p(1,1) = \binom{3}{1} \binom{4}{1} \binom{5}{1} / \binom{12}{3} = \frac{60}{220}$$

$$p(1,2) = \binom{3}{1} \binom{4}{2} / \binom{12}{3} = \frac{18}{220}$$

$$p(2,0) = \binom{3}{2} \binom{5}{1} / \binom{12}{3} = \frac{15}{220}$$

$$p(2,1) = \binom{3}{2} \binom{4}{1} / \binom{12}{3} = \frac{12}{220}$$

$$p(3,0) = \binom{3}{3} / \binom{12}{3} = \frac{1}{220}$$

Joint probability mass function (pmf)

TABLE 6.1: $P\{X = i, Y = j\}$

j	0	1	2	3	Row sum = $P{X = i}$
0	$\frac{10}{220}$	$\frac{40}{220}$	$\frac{30}{220}$	$\frac{4}{220}$	$\frac{84}{220}$
1	$\frac{30}{220}$	$\frac{60}{220}$	$\frac{18}{220}$	0	$\frac{108}{220}$
2	$\frac{15}{220}$	$\frac{12}{220}$	0	0	$\frac{27}{220}$
3	$\frac{1}{220}$	0	0	0	$\frac{1}{220}$
Column sum = $P{Y = j}$	$\frac{56}{220}$	$\frac{112}{220}$	$\frac{48}{220}$	$\frac{4}{220}$	

Joint probability density function (pdf)

Definition

X and Y are *jointly continuous* if there exists function f(x,y) defined on \mathbb{R}^2 so that for any Borel subset B

$$P[(X,Y) \in B] = \iint_B f(x,y) dxdy$$

f(x, y) is the joint pdf of X and Y

If A and B are Borel subset of \mathbb{R} then

$$P[(X, Y) \in (A \times B)] = \int_{A} \int_{B} f(x, y) dx dy$$
$$F(a, b) = \int_{-\infty}^{a} \int_{-\infty}^{b} f(x, y) dx dy$$
$$f(x, y) = \frac{\partial^{2}}{\partial x \partial y} F(x, y)$$

Single pdf

- Q: How to get a single PDF from joint PDF?
- If X and Y are jointly continuous then they are continuous with pdf.
 The single PDFs are:

$$f_X(x) = \int_{-\infty}^{\infty} f(x, y) dy$$

$$f_Y(y) = \int_{-\infty}^{\infty} f(x, y) dx$$

Joint PDF and single PDF

Example 1d

The joint pdf of X and Y is

$$f(x,y) = \begin{cases} 2e^{-x}e^{-2y} & 0 < x, y < \infty \\ 0 & \text{otherwise} \end{cases}$$

Compute:

(a)
$$P(X > 1, Y < 1)$$
; (b) $P(X < Y)$; (c) $P(X < a)$.

Solution: (a)

$$P\{X > 1, Y < 1\} = \int_0^1 \int_1^\infty 2e^{-x}e^{-2y} dx dy$$
$$= \int_0^1 2e^{-2y} \left(-e^{-x} \Big|_1^\infty \right) dy$$
$$= e^{-1} \int_0^1 2e^{-2y} dy$$
$$= e^{-1} (1 - e^{-2})$$

(b)

$$P(X < Y) = \iint_{x < y} 2e^{-x}e^{-2y}dxdy$$
$$= \int_0^\infty \int_0^y 2e^{-x}e^{-2y}dxdy$$
$$= 1/3$$

(c)

$$P(X < a) = \int_0^a \int_0^\infty 2e^{-x} e^{-2y} dx dy$$

= 1 - e^{-a}

Joint pdf and single PDF

Example 1e

Joint pdf of X and Y

$$f(x,y) = \begin{cases} e^{-(x+y)} & 0 < x, y < \infty \\ 0 & \text{otherwise} \end{cases}$$

Find the pdf of X/Y

Compute the cdf of X/Y:

$$F_{X/Y}(z) = P(X/Y \le z = \iint_{x/y \le z} e^{-(x+y)} dx dy$$
$$= \int_0^\infty \int_0^{zy} e^{-(x+y)} dx dy = 1 - \frac{1}{z+1}$$

Taking derivative of the cdf, we get the pdf of X/Y

$$f(z) = \frac{1}{(z+1)^2}, \quad 0 < z < \infty$$

Independent Random Variables

Independent Random Variables

Definition

Two random variables X and Y are *independent* if for any Borel sets A and B, the events $(X \in A)$ and $(Y \in B)$ are independent.

$$P(X \in A, Y \in B) = P(X \in A)P(Y \in B)$$

It's very easy to find the joint cdf of independent RVs

$$F(x,y) = F_X(x)F_Y(y)$$

Independent Random Variables

Example 2f

The number of people who enter a post office on a given day is a Poisson RV with parameter λ . Show that if each person is a male with probability p and a female with probability 1-p, then the number of males and females entering the post office are independent Poisson RVs with parameters λp and $\lambda(1-p)$.

Solution

- Let X = number of males, Y = number of females
- Z = X + Y: total number of people entering the post office then $Z = Poisson(\lambda)$
- We have

$$P(X = i, Y = j) =$$

 $P(X = i, Y = j | Z = i + j)P(Z = i + j)$

• Need to compute P(X = i, Y = j | Z = i + j) and P(Z = i + j).

Example 2f (Cont.)

• Given (i + j) people entering the post office, probability that exactly i males and j females is Binomial(i + j, p)

$$P(X = i, Y = j | Z = i + j) =$$

$$= {i + j \choose i} p^{i} (1 - p)^{j}$$

• Z is Poisson with λ , so $P(Z = i + j) = e^{-\lambda} \frac{\lambda^{i+j}}{(i+j)!}$.

$$P\{X = i, Y = j\} = \binom{i+j}{i} p^{i} (1-p)^{j} e^{-\lambda} \frac{\lambda^{i+j}}{(i+j)!}$$

$$= e^{-\lambda} \frac{(\lambda p)^{i}}{i!j!} [\lambda (1-p)]^{j}$$

$$= \frac{e^{-\lambda p} (\lambda p)^{i}}{i!} e^{-\lambda (1-p)} \frac{[\lambda (1-p)]^{j}}{i!}$$

Compute the single pmf

On the other hand:

$$P(X = i) = \sum_{j} P(X = i, Y = j) = e^{-\lambda p} \frac{(\lambda p)^{i}}{i!}$$

so X is Poisson(λp). Also,

$$P(Y = j) = \sum_{i} P(X = i, Y = j) = e^{-\lambda(1-p)} \frac{(\lambda(1-p))^{j}}{j!}$$

So *Y* is Poisson($\lambda(1-p)$).

Therefore

$$P(X = i, Y = j) = P(X = i)P(Y = j)$$

then X and Y are independent.

Separable joint pdf

- If X and Y are independent then the joint pdf is product of single pdfs.
- Is the reverse true?
- YES!

Proposition

Two continuous (discrete) RVs X and Y are independent if and only if the joint pdf (pmf) can be expressed as

$$f_{X,Y}(x,y) = h(x)g(y)$$

for all x, y.

Separable joint pdf

Example 2g

Given the joint pdf of X and Y is

$$f(x,y)=24xy$$

for 0 < x < 1, 0 < y < 1, 0 < x + y < 1 and zero otherwise.

Are X and Y independent?

Solution: Let

$$A = \{0 < x < 1, 0 < y < 1,$$

$$0 < x + y < 1$$
.

Define the indicator function $I_A(x,y) = 1$ if $(x,y) \in A$ and zero otherwise.

Then $f(x, y) = 24xyI_A(x, y)$ not separable. Therefore, X and Y are not independent (or, dependent).

Sum of independent Random Variables

Sum of independent Random Variables

- X and Y continuous and independent
- We can find the cdf of X + Y

$$\begin{split} F_{X+Y}(a) &= P\{X + Y \le a\} \\ &= \iint_{x+y \le a} f_X(x) f_Y(y) \, dx \, dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_X(x) f_Y(y) \, dx \, dy \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{a-y} f_X(x) \, dx f_Y(y) \, dy \\ &= \int_{-\infty}^{\infty} F_X(a - y) f_Y(y) \, dy \end{split}$$

pdf of sum

$$f_{X+Y}(a) = \frac{d}{da} \int_{-\infty}^{\infty} F_X(a - y) f_Y(y) \, dy$$
$$= \int_{-\infty}^{\infty} \frac{d}{da} F_X(a - y) f_Y(y) \, dy$$
$$= \int_{-\infty}^{\infty} f_X(a - y) f_Y(y) \, dy$$

 f_{X+Y} is called the *convolution* of f_X and f_Y .

Sum of normal RVs and Poisson RVs

Sum of normal RVs

lf

$$X_i = N(\mu_i, \sigma_i^2), i = 1, \ldots, n$$

are independent, then

$$\sum_{i} X_{i} = N(\sum_{i} \mu_{i}, \sum_{i} \sigma_{i}^{2})$$

Sum of Poisson RVs

If X and Y are independent Poisson RVs with λ_1 and λ_2 , then X+Y has Poisson distribution with $\lambda_1 + \lambda_2$.

Conditional Distributions: Discrete case and continuous case

Conditional Distributions: Discrete case

The conditional pmf

If X and Y are discrete random variables, we define the conditional probability mass function of X given that Y = y, by

$$p_{X|Y}(x|y) = P(X = x|Y = y) = \frac{p(x,y)}{p_Y(y)}$$

for all values of y such that $p_Y(y) > 0$.

The conditional pdf

the conditional probability distribution function of X given that Y = y is defined, for all y such that $p_Y(y) > 0$, by

$$F_{X|Y}(x|y) = P(X \le x|Y = y) = \sum_{a \le x} p_{X|Y}(a|y).$$

Conditional Distributions: Discrete case

Example

Suppose that p(x, y), the joint probability mass function of X and Y, is given by p(0,0) = 0.4, p(0,1) = 0.2, p(1,0) = 0.1, p(1,1) = 0.3. Calculate the conditional probability mass function of X given that Y = 1.

Solution

Note that

$$p_Y(1) = \sum_{x} p(x, 1) = p(0, 1) + p(1, 1) = 0.5$$

Therefore,

$$p_{X|Y}(0|1) = \frac{p(0,1)}{p_Y(1)} = \frac{2}{5}$$

$$p_{X|Y}(1|1) = \frac{p(1,1)}{p_Y(1)} = \frac{3}{5}$$

Conditional Distributions: Discrete case

Example

If X and Y are independent Poisson random variables with respective parameters λ_1 and λ_2 , calculate the conditional distribution of X given that X+Y=n.

Hint:

$$P(X = k | X + Y = n) = \frac{P(X = k) P(Y = n - k)}{P(X + Y = n)}$$

$$P(X = k | X + Y = n) = {n \choose k} \left(\frac{\lambda_1}{\lambda_1 + \lambda_2}\right)^k \left(\frac{\lambda_2}{\lambda_1 + \lambda_2}\right)^{n - k}$$

Conditional Distributions: Continuous case

Definition

If X and Y have a joint probability density function f(x, y), then the conditional probability density function of X given that Y = y is defined, for all values of y such that $f_Y(y) > 0$, by

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)}$$

If X and Y are jointly continuous, then, for any set A

$$P(X \in A|Y = y) = \int_{A} f_{X|Y}(x|y) dx$$

Conditional Distributions: Continuous case

We define the conditional cumulative distribution function of X given that Y = y by

$$F_{X|Y}(a|y) = P(X \le a|Y = y) = \int_{-\infty}^{a} f_{X|Y}(x|y) dx$$

Example

The joint density of X and Y is given by

$$f(x,y) = \begin{cases} \frac{12}{5}x(2-x-y), 0 < x < 1, 0 < y < 1\\ 0, \text{ otherwise} \end{cases}$$

Compute the conditional density of X given that Y = y, where 0 < y < 1.

Solution For 0 < x < 1, 0 < y < 1, we have

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{f(x,y)}{\int_0^1 \frac{12}{5} x(2-x-y) dx} = \frac{6x(2-x-y)}{4-3y}$$

Conditional Distributions: Continuous case

Example

The joint density of X and Y is given by

$$f(x,y) = \begin{cases} \frac{e^{-x/y}e^{-y}}{y}, & 0 < x < \infty, 0 < y < \infty \\ 0, & \text{otherwise} \end{cases}$$

Find P(X > 1 | Y = y).

Solution

We first obtain the conditional density of X given that Y = y

$$f_{X|Y}(x|y) = \frac{f(x,y)}{f_Y(y)} = \frac{e^{-x/y}e^{-y}/y}{e^{-y}\int_0^\infty (1/y)e^{-x/y}dx} = \frac{e^{-x/y}}{y}$$
$$P(X > 1|Y = y) = \int_1^\infty (1/y)e^{-x/y}dx = e^{-1/y}$$

- Let X_1 and X_2 be jointly continuous random variables with joint probability density function f_{X_1,X_2} .
- Suppose that $Y_1 = g_1(X_1, X_2)$ and $Y_2 = g_2(X_1, X_2)$ for some functions g_1 and g_2 . We want to obtain the joint distribution of the random variables Y_1 and Y_2 .
- Assume:
 - 1. The equations $y_1 = g_1(x_1, x_2)$ and $y_2 = g_2(x_1, x_2)$ can be uniquely solved for x_1 and x_2 : $x_1 = h_1(y_1, y_2)$, $x_2 = h_2(y_1, y_2)$.
 - 2. The functions g_1 and g_2 have continuous partial derivatives at all points (x_1, x_2) and are such that

$$J(x_1, x_2) = \begin{vmatrix} \frac{\partial g_1}{\partial x_1} & \frac{\partial g_1}{\partial x_2} \\ \frac{\partial g_2}{\partial x_1} & \frac{\partial g_2}{\partial x_2} \end{vmatrix} \neq 0$$

- Let X_1 and X_2 be jointly continuous random variables with joint probability density function f_{X_1,X_2} .
- Suppose that $Y_1 = g_1(X_1, X_2)$ and $Y_2 = g_2(X_1, X_2)$ for some functions g_1 and g_2 . We want to obtain the joint distribution of the random variables Y_1 and Y_2 .
- Conditions:
 - 1. The equations $y_1 = g_1(x_1, x_2)$ and $y_2 = g_2(x_1, x_2)$ can be uniquely solved for x_1 and x_2 : $x_1 = h_1(y_1, y_2)$, $x_2 = h_2(y_1, y_2)$.
 - 2. The functions g_1 and g_2 have continuous partial derivatives at all points (x_1, x_2) and are such that

$$J(x_1, x_2) = \begin{vmatrix} \frac{\partial g_1}{\partial x_1} & \frac{\partial g_1}{\partial x_2} \\ \frac{\partial g_2}{\partial x_1} & \frac{\partial g_2}{\partial x_2} \end{vmatrix} \neq 0$$

Theorem

Under these two conditions, it can be shown that the random variables Y_1 and Y_2 are jointly continuous with joint density function given by

$$f_{Y_1Y_2}(y_1, y_2) = f_{X_1, X_2}(x_1, x_2)|J(x_1, x_2)|^{-1},$$

where $x_1 = h_1(y_1, y_2)$, $x_2 = h_2(y_1, y_2)$.

Example

Let X_1 and X_2 be jointly continuous random variables with probability density function f_{X_1,X_2} . Let $Y_1 = X_1 + X_2$, $Y_2 = X_1 - X_2$. Find the joint density function of Y_1 and Y_2 in terms of f_{X_1,X_2} .

We have

$$J(x_1,x_2) = \left| \begin{array}{cc} 1 & 1 \\ 1 & -1 \end{array} \right| = -2$$

and $x_1 = (y_1 + y_2)/2$, $x_2 = (y_1 - y_2)/2$. It follows

$$f_{Y_1,Y_2}(y_1,y_2) = \frac{1}{2} f_{X_1,X_2}(\frac{y_1+y_2}{2},\frac{y_1-y_2}{2})$$

-END OF CHAPTER 6-