



# Chapter 4

**Further development and analysis of  
the classical linear regression model**

# Generalising the Simple Model to Multiple Linear Regression

- Before, we have used the model

$$y_t = \alpha + \beta x_t + u_t, \quad t = 1, 2, \dots, T$$

- But what if our dependent variable ( $y$ ) depends on more than one independent variable?

For example the **number of cars sold** might plausibly depend on

1. the price of cars
2. the price of public transport
3. the price of petrol
4. the extent of the public's concern about global warming

- Similarly, **stock returns** might depend on several factors: inflation, the difference in returns on short and long dated bonds, industrial production, default risks.
- Having just one independent variable is no good in this case - we want to have more than one  $x$  variable. It is very easy to **generalise the simple model to the one with  $k$  regressors.**

# Multiple Regression and the Constant Term

- Now we write

$$y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + \dots + \beta_k x_{kt} + u_t, t=1,2,\dots,T$$

- Where is  $x_1$ ? By convention it is the constant column of length  $T$ :

$$x_1 = \begin{bmatrix} 1 \\ 1 \\ \vdots \\ 1 \end{bmatrix}$$

- $k$ : **the number of regressors** (including  $x_1$ ), which is the number of parameters  $\beta_i$

# Different Ways of Expressing the Multiple Linear Regression Model

- We could write out a separate equation for every value of  $t$ :

$$y_1 = \beta_1 + \beta_2 x_{21} + \beta_3 x_{31} + \dots + \beta_k x_{k1} + u_1$$

$$y_2 = \beta_1 + \beta_2 x_{22} + \beta_3 x_{32} + \dots + \beta_k x_{k2} + u_2$$

$$\vdots \qquad \qquad \qquad \vdots \qquad \qquad \qquad \vdots$$

$$y_T = \beta_1 + \beta_2 x_{2T} + \beta_3 x_{3T} + \dots + \beta_k x_{kT} + u_T$$

- We can write this in matrix form

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{u}$$

where

$\mathbf{y}$  is  $T \times 1$

$\mathbf{X}$  is  $T \times k$

$\boldsymbol{\beta}$  is  $k \times 1$

$\mathbf{u}$  is  $T \times 1$

# Inside the Matrices of the Multiple Linear Regression Model

- Example: if  $k$  is 2, we have 2 regressors, one of which is a column of ones:

$$\begin{array}{ccccccc} \left[ \begin{array}{c} y_1 \\ y_2 \\ \vdots \\ y_T \end{array} \right] & = & \left[ \begin{array}{cc} 1 & x_{21} \\ 1 & x_{22} \\ \vdots & \vdots \\ 1 & x_{2T} \end{array} \right] & \left[ \begin{array}{c} \beta_1 \\ \beta_2 \end{array} \right] & + & \left[ \begin{array}{c} u_1 \\ u_2 \\ \vdots \\ u_T \end{array} \right] \\ T \times 1 & & T \times 2 & 2 \times 1 & & T \times 1 \end{array}$$

- Notice that the matrices written this way are comfortable,  *$x_{it}$  is the element in the  $t^{\text{th}}$  row and  $i^{\text{th}}$  column*

## How Do We Calculate the Parameters (the $\beta$ ) in this Generalised Case?

- Previously, we took the residual sum of squares, and minimised it w.r.t.  $\alpha$  and  $\beta$ .
- Similarly, with the matrix notation, we have

$$\hat{u} = \begin{bmatrix} \hat{u}_1 \\ \hat{u}_2 \\ \vdots \\ \hat{u}_T \end{bmatrix}$$

- The **Residual Sum Square (RSS)** would be given by

$$\hat{u}'\hat{u} = \begin{bmatrix} \hat{u}_1 & \hat{u}_2 & \dots & \hat{u}_T \end{bmatrix} \begin{bmatrix} \hat{u}_1 \\ \hat{u}_2 \\ \vdots \\ \hat{u}_T \end{bmatrix} = \hat{u}_1^2 + \hat{u}_2^2 + \dots + \hat{u}_T^2 = \sum \hat{u}_t^2$$

# The OLS Estimator for the Multiple Regression Model

- In order to obtain the parameter estimates,  $\beta_1, \beta_2, \dots, \beta_k$ , we would minimise the RSS with respect to all the  $\beta$ .
- It can be shown that

$$\hat{\beta} = \begin{bmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \\ \vdots \\ \hat{\beta}_k \end{bmatrix} = (X'X)^{-1} X' y$$

## Calculating the **Standard Errors** for the Multiple Regression Model

- Check the dimensions:  $\hat{\beta}$  is  $k \times 1$  as required.
- But how do we calculate the standard errors of the coefficient estimates?
- Previously, to estimate the variance of the errors,  $\sigma^2$ , we use  $s^2 = \frac{\hat{u}'\hat{u}}{T-k}$

where  $k$  = number of parameters.

- Parameter Variance-Covariance matrix:  $Var(\hat{\beta}) \approx s^2 (X'X)^{-1}$
- Then  $SE(\hat{\beta})$  is obtained by taking square root of the leading diagonal of the Parameter Variance-Covariance matrix



## Calculating Parameter and Standard Error Estimates for Multiple Regression Models: An Example

- **Example:** The following model with  $k=3$  is estimated over 15 observations:

$$y = \beta_1 + \beta_2 x_2 + \beta_3 x_3 + u$$

and the following data have been calculated from the original  $X$ 's.

$$(X'X)^{-1} = \begin{bmatrix} 2.0 & 3.5 & -1.0 \\ 3.5 & 1.0 & 6.5 \\ -1.0 & 6.5 & 4.3 \end{bmatrix}, (X'y) = \begin{bmatrix} -3.0 \\ 2.2 \\ 0.6 \end{bmatrix}, \hat{u}'\hat{u} = 10.96$$

Calculate the coefficient estimates and their standard errors.

- To calculate the coefficients, just multiply the matrix by the vector to obtain

$$\hat{\beta} = (X'X)^{-1} X'y$$

- To calculate the standard errors, we need an estimate of  $\sigma^2$ .

$$s^2 = \frac{RSS}{T-k} = \frac{10.96}{15-3} = 0.91$$

## Calculating Parameter and Standard Error Estimates for Multiple Regression Models: An Example (cont'd)

- The variance-covariance matrix of  $\hat{\beta}$  is given by

$$s^2(X'X)^{-1} = 0.91(X'X)^{-1} = \begin{bmatrix} 1.83 & 3.20 & -0.91 \\ 3.20 & 0.91 & 5.94 \\ -0.91 & 5.94 & 3.93 \end{bmatrix}$$

- The variances are on the leading diagonal:

$$\text{Var}(\hat{\beta}_1) = 1.83 \quad SE(\hat{\beta}_1) = 1.35$$

$$\text{Var}(\hat{\beta}_2) = 0.91 \Leftrightarrow SE(\hat{\beta}_2) = 0.96$$

$$\text{Var}(\hat{\beta}_3) = 3.93 \quad SE(\hat{\beta}_3) = 1.98$$

- We write:  $\hat{y} = 1.10 - 4.40x_{2t} + 19.88x_{3t}$   
(1.35) (0.96) (1.98)

## Testing Multiple Hypotheses: The $F$ -test

- We used the  $t$ -test to test single hypotheses, i.e. hypotheses involving only one coefficient. But what if we want to test more than one coefficient simultaneously?
- We do this using the  $F$ -test. The  $F$ -test involves estimating 2 regressions.
- The unrestricted regression is the one in which the coefficients are freely determined by the data, as we have done before, with residual sum of squares **URSS**
- The restricted regression is the one in which the coefficients are restricted, i.e. the restrictions are imposed on some  $\beta$ s, with residual sum of squares **RRSS**

# The $F$ -test:

## Restricted and Unrestricted Regressions

- **Example:**

The general regression is

$$y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + \beta_4 x_{4t} + u_t \quad (1)$$

- We want to test the restriction that  $\beta_3 + \beta_4 = 1$  (we have some hypothesis from theory which suggests that this would be an interesting hypothesis to study). The **unrestricted regression is (1) above**, but what is the restricted regression?

$$y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + \beta_4 x_{4t} + u_t \quad s.t. \quad \beta_3 + \beta_4 = 1$$

- We substitute the restriction ( $\beta_3 + \beta_4 = 1$ ) into the regression so that it is automatically imposed on the data.

$$\beta_3 + \beta_4 = 1 \Rightarrow \beta_4 = 1 - \beta_3$$

# The $F$ -test: Forming the Restricted Regression

$$y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + (1 - \beta_3)x_{4t} + u_t$$

$$y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + x_{4t} - \beta_3 x_{4t} + u_t$$

- Gather terms in  $\beta$ 's together and rearrange

$$(y_t - x_{4t}) = \beta_1 + \beta_2 x_{2t} + \beta_3 (x_{3t} - x_{4t}) + u_t$$

- This is the restricted regression. We actually estimate it by creating two new variables, call them, say,  $P_t$  and  $Q_t$ .

$$P_t = y_t - x_{4t}$$

$$Q_t = x_{3t} - x_{4t}$$

so

$P_t = \beta_1 + \beta_2 x_{2t} + \beta_3 Q_t + u_t$  is the restricted regression we actually estimate.

# Calculating the F-Test Statistic

- The **test statistic** is given by

$$\text{test statistic} = \frac{RRSS - URSS}{URSS} \times \frac{T - k}{m}$$

where  $URSS$  = RSS from unrestricted regression

$RRSS$  = RSS from restricted regression

$m$  = number of restrictions

$T$  = number of observations

$k$  = number of regressors in unrestricted regression including the constant (or the total number of parameters to be estimated).

# The $F$ -Distribution

- The test statistic follows the  $F$ -distribution, which has 2 d.f. parameters.
- The value of the **degrees of freedom parameters are  $(m, T-k)$**  respectively (the order of the d.f. parameters is important).
- The appropriate critical value will be in column  $m$ , row  $(T-k)$ .
- The  $F$ -distribution has only positive values and is not symmetrical. **We therefore only reject the null if the test statistic  $>$  critical  $F$ -value**

## Determining the **Number of Restrictions m** in an $F$ -test

- Examples:

$H_0$ : hypothesis	No. of restrictions, $m$
$\beta_1 + \beta_2 = 2$	1
$\beta_2 = 1$ and $\beta_3 = -1$	2
$\beta_2 = 0$ , $\beta_3 = 0$ and $\beta_4 = 0$	3

- If the model is  $y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + \beta_4 x_{4t} + u_t$ ,  
then the null hypothesis

**$H_0$ :  $\beta_2 = 0$ , and  $\beta_3 = 0$  and  $\beta_4 = 0$**  is tested by “**THE regression  $F$ -statistic**”, or “**junk regression**”. It tests the null hypothesis that all of the coefficients except the intercept coefficient are zero.

- Note the form of the alternative hypothesis for all tests when more than one restriction is involved:  **$H_1$ :  $\beta_2 \neq 0$ , or  $\beta_3 \neq 0$  or  $\beta_4 \neq 0$**



# What we Cannot Test with Either an $F$ or a $t$ -test

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- We cannot test using this framework hypotheses which are not linear or which are multiplicative, e.g.

$$H_0: \beta_2 \beta_3 = 2 \text{ or } H_0: \beta_2^2 = 1$$

cannot be tested.

# The Relationship between the $t$ and the $F$ -Distributions

- Any hypothesis which could be tested with a  $t$ -test could have been tested using an  $F$ -test, but not the other way around.

For example, consider the hypothesis

$$H_0: \beta_2 = 0.5$$

$$H_1: \beta_2 \neq 0.5$$

We could have tested this using the usual  $t$ -test:  $test\ stat = \frac{\hat{\beta}_2 - 0.5}{SE(\hat{\beta}_2)}$

or it could be tested in the framework above for the  $F$ -test.

- Note that the two tests always give the same result since the  $t$ -distribution is just a special case of the  $F$ -distribution.
- If we have some random variable  $Z \sim t(T-k)$  then  $Z^2 \sim F(1, T-k)$**

## ***F*-test Example**

- **Question**: Suppose a researcher wants to test whether the returns on a company stock ( $y$ ) show unit sensitivity to two factors (factor  $x_2$  and factor  $x_3$ ) among three considered. The regression is carried out on 144 monthly observations. The regression is  $y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + \beta_4 x_{4t} + u_t$ 
  - What are the restricted and unrestricted regressions?
  - If the two RSS are 436.1 and 397.2 respectively, perform the test.
- **Solution**:

Unit sensitivity implies  $H_0: \beta_2=1$  and  $\beta_3=1$ . The unrestricted regression is the one in the question. The restricted regression is  $(y_t - x_{2t} - x_{3t}) = \beta_1 + \beta_4 x_{4t} + u_t$  or letting  $z_t = y_t - x_{2t} - x_{3t}$ , the restricted regression is  $z_t = \beta_1 + \beta_4 x_{4t} + u_t$

In the  $F$ -test formula,  $T=144$ ,  $k=4$ ,  $m=2$ ,  $RRSS=436.1$ ,  $URSS=397.2$

$F$ -test statistic = 6.68. Critical value is an  $F(2,140) = 3.07$  (5%) and 4.79 (1%).

Conclusion: Reject  $H_0$ .

## Goodness of Fit Statistics

- We would like some measure of how well our regression model actually fits the data.
- We have goodness of fit statistics to test this: i.e. how well the sample regression function (SRF) fits the data.
- The most common goodness of fit statistic is known as  $R^2$ . One way to define  $R^2$  is to say that it is the square of the correlation coefficient between  $y$  and  $\hat{y}$ .
- For another explanation, recall that what we are interested in doing is explaining the variability of  $y$  about its unconditional mean  $\bar{y}$ , i.e. the **total sum of squares  $TSS$** :

$$TSS = \sum_t (y_t - \bar{y})^2$$

- We can split the  $TSS$  into two parts: the part explained by the model (**explained sum of squares  $ESS$** ), and the part not explained by the model (**residual sum of squares  $RSS$** ).

## Defining $R^2$

- That is,  $TSS = ESS + RSS$

$$\sum_t (y_t - \bar{y})^2 = \sum_t (\hat{y}_t - \bar{y})^2 + \sum_t \hat{u}_t^2$$

- Our **goodness of fit statistic** is

$$R^2 = \frac{ESS}{TSS}$$

- But since  $TSS = ESS + RSS$ , we can also write

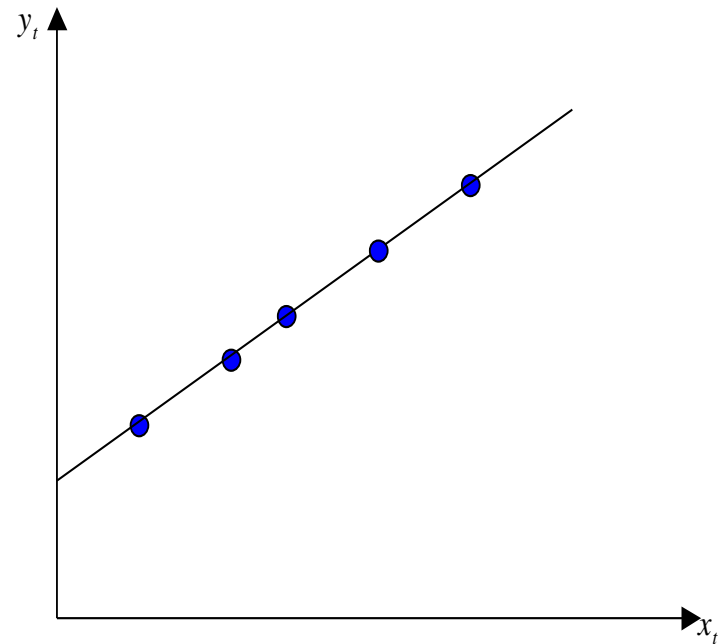
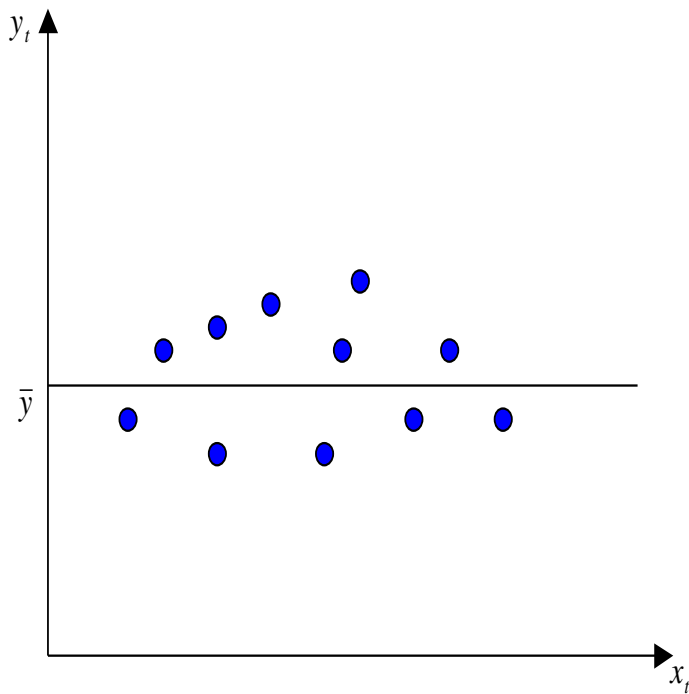
$$R^2 = \frac{ESS}{TSS} = \frac{TSS - RSS}{TSS} = 1 - \frac{RSS}{TSS}$$

- $R^2$  must always lie between zero and one. To understand this, consider two extremes

$$RSS = TSS \quad \text{i.e.} \quad ESS = 0 \quad \text{so} \quad R^2 = ESS/TSS = 0$$

$$ESS = TSS \quad \text{i.e.} \quad RSS = 0 \quad \text{so} \quad R^2 = ESS/TSS = 1$$

## The Limit Cases: $R^2 = 0$ and $R^2 = 1$



## Problems with $R^2$ as a Goodness of Fit Measure

- There are a number of them:

**1.  $R^2$  is defined in terms of variation about the mean of  $y$**  so that if a model is reparameterised (rearranged) and the dependent variable changes,  $R^2$  will change.

**2.  $R^2$  never falls if more regressors are added to the regression**, e.g. consider:

$$\text{Regression 1: } y_t = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + u_t$$

$$\text{Regression 2: } y = \beta_1 + \beta_2 x_{2t} + \beta_3 x_{3t} + \beta_4 x_{4t} + u_t$$

$R^2$  will always be at least as high for regression 2 relative to regression 1.

## Adjusted $R^2$

- In order to get around these problems, a modification is often made which takes into account the loss of degrees of freedom associated with adding extra variables. This is known as  $\bar{R}^2$ , or **adjusted  $R^2$** :

$$\bar{R}^2 = 1 - \left[ \frac{T-1}{T-k} (1 - R^2) \right]$$

- So if we add an extra regressor,  $k$  increases and unless  $R^2$  increases by a more than offsetting amount,  $\bar{R}^2$  will actually fall.
- **Example:** If **adjusted  $R^2$**  is 0.65 then 65% of the total variability of the dependent variable about the mean is explained by the model.



## A Regression Example: Hedonic House Pricing Models

- **Hedonic models** are used to value real assets, especially housing, and view the asset as representing a bundle of characteristics.
- Des Rosiers and Thériault (1996) consider the effect of various amenities on rental values for buildings and apartments 5 sub-markets in the Quebec area of Canada.
- The **rental value** in Canadian Dollars per month (the dependent variable) **is a function of 9 to 14 variables** (depending on the area under consideration). The paper employs 1990 data, and for the Quebec City region, there are 13,378 observations, and the 12 explanatory variables are:

LnAGE            - log of the apparent age of the property

NBROOMS       - number of bedrooms

AREABYRM      - area per room (in square metres)

ELEVATOR      - a dummy variable = 1 if the building has an elevator; 0 otherwise

BASEMENT      - a dummy variable = 1 if the unit is located in a basement; 0 otherwise

# Hedonic House Pricing Models: Variable Definitions

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OUTPARK	- number of outdoor parking spaces
INDPARK	- number of indoor parking spaces
NOLEASE	- a dummy variable = 1 if the unit has no lease attached to it; 0 otherwise
LnDISTCBD	- log of the distance in kilometres to the central business district
SINGLPAR	- percentage of single parent families in the area where the building stands
DSHOPCNTR	- distance in kilometres to the nearest shopping centre
VACDIFF1	- vacancy difference between the building and the census figure

- Examine the signs and sizes of the coefficients.
  - The coefficient estimates themselves show the Canadian dollar rental price per month of each feature of the dwelling.

# Hedonic House Price Results

## Dependent Variable: Canadian Dollars per Month

Variable	Coefficient	<i>t</i> -ratio	<i>A priori</i> sign expected
Intercept	282.21	56.09	+
LnAGE	-53.10	-59.71	-
NBROOMS	48.47	104.81	+
AREABYRM	3.97	29.99	+
ELEVATOR	88.51	45.04	+
BASEMENT	-15.90	-11.32	-
OUTPARK	7.17	7.07	+
INDPARK	73.76	31.25	+
NOLEASE	-16.99	-7.62	-
LnDISTCBD	5.84	4.60	-
SINGLPAR	-4.27	-38.88	-
DSHOPCNTR	-10.04	-5.97	-
VACDIFF1	0.29	5.98	-

Notes: Adjusted  $R^2 = 0.651$ ; regression  $F$ -statistic = 2082.27. Source: Des Rosiers and Thériault (1996). Reprinted with permission of the American Real Estate Society.

## Tests of Non-nested Hypotheses

- All of the hypothesis tests concluded thus far have been in the context of “**nested**” models.

- But what if we wanted to compare between the following models?

$$\text{Model 1: } y_t = \alpha_1 + \alpha_2 x_{2t} + u_t$$

$$\text{Model 2: } y_t = \beta_1 + \beta_2 x_{3t} + v_t$$

- We could use  $R^2$  or adjusted  $R^2$ , but what if the number of explanatory variables were different across the 2 models?
- An alternative approach is an encompassing test, based on examination of the **hybrid model**:  
Model 3:  $y_t = \gamma_1 + \gamma_2 x_{2t} + \gamma_3 x_{3t} + w_t$

## Tests of Non-nested Hypotheses (cont'd)

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- There are 4 possible outcomes when Model 3 is estimated:
  - $\gamma_2$  is significant but  $\gamma_3$  is not
  - $\gamma_3$  is significant but  $\gamma_2$  is not
  - $\gamma_2$  and  $\gamma_3$  are both statistically significant
  - Neither  $\gamma_2$  nor  $\gamma_3$  are significant
- Problems with encompassing approach
  - Hybrid model may be meaningless
  - Possible high correlation between  $x_2$  and  $x_3$ .