

The Hyperedge Event Model

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Abstract. We introduce the hyperedge event model (HEM)—a generative model for events that can be represented as directed edges with one sender and one or more receivers or one receiver and one or more senders. To define the model, we integrate a dynamic version of the exponential random graph model (ERGM) of edge structure with a survival model for event timing to jointly understand who interacts with whom, and when. The HEM offers three innovations with respect to the literature—first, it extends a growing class of dynamic network models to model hyperedges. Second, we derive a receiver-selection distribution that forces the sender to select at least one receiver. Third, we incorporate both a timing equation and an edge-formation equation into the model specification. We use the HEM to analyze emails sent among department managers in Montgomery County government in North Carolina. Our application demonstrates that the model is effective at predicting and explaining time-stamped network data involving edges with multiple receivers. We present a out-of-sample prediction experiment to illustrate how researchers can select between different specifications of the model.

MSC 2010 subject classifications: Primary 60K35, 60K35; secondary 60K35.

Keywords: continuous time network model, email data analysis.

1 Introduction

Processes that arise as time-stamped directed interactions are common in the social, natural, and physical sciences. The data produced by such processes can be represented as dynamic directed networks—a data structure that has given rise to the development of several statistical model families. For example, stochastic actor-oriented models (SAOMs) (Snijders, 1996; Snijders et al., 2007) characterize networks as evolving as the senders of edges decide to add an edge to or remove an edge from a network one edge at a time, one sender at a time. Event-based network models (Butts, 2008; Vu et al., 2011; Hunter et al., 2011; Perry and Wolfe, 2013) provide a general framework for modeling the realization of edges that occur as instances in continuous time streams of events. These models are flexible enough to specify a generative model that accounts for the ways in which past network structure begets future network structure (e.g., nodes that sent a lot of edges in the past will send a lot in the future, if node i sent a tie to node j recently, then node j will send one to node i in the near future). Thus, dynamic network models are very useful to understand which traits and behaviours are predictive of interactions.

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A major limitation of existing dynamic network models is that they apply to edges with one sender and one receiver. Many sources of dynamic network data arise as “hyperedges” (Karypis et al., 1999; Ghoshal et al., 2009; Zlatić et al., 2009; Zhang and Liu, 2010) that include one sender and multiple receivers or one receiver and multiple senders. For example, networks formed through the analysis of email (Newman et al., 2002) arise with one sender and possibly multiple receivers; networks formed between neurons via axons and dendrites have one sender and multiple receivers (axons) or one receiver and multiple senders (dendrites) (Partzsch and Schüffny, 2012). Networks formed through the analysis of legislative bill cosponsorship (Fowler, 2006) arise with multiple senders (cosponsors) and one receiver (sponsor); and networks formed through the analysis of economic sanctions (Crammer et al., 2014) between countries arise with multiple sending countries and one target country. Existing models require that researchers alter hyperedge data to fit with the pairwise edge structure of the model. For instance, Perry and Wolfe (2013) treat multicast interactions—one type of directed hyperedge which involves one sender and one or more receivers—via duplication (i.e., obtain pairwise interactions from the original multicast), to construct approximate likelihood function in their inferential framework for model parameters. Similarly, Fan and Shelton (2009) duplicates emails sent from one sender to one or more receivers and randomly jitter the sent times, in order to avoid the violation of the assumption that two events cannot occur at the exact same time.

We develop a statistical dynamic network model, which we term the hyperedge event model (HEM), that integrates the two components that govern time-stamped event formation: 1) the formation of the vertices that are incident to the hyperedge, and 2) the timing of the event in which the edge manifests. In what follows, we introduce the HEM by defining a generative process for time and hyperedge attributed event data (Section 2), derive the conditional posteriors for Bayesian inference and present tests of our software implementation (Section 3). After defining the structure of and deriving Bayesian inference methods for HEM, we demonstrate the model’s applicability by analyzing data from a corpus of internal county government emails. In the application we illustrate model selection, posterior predictive checks, and exploratory analysis (Section 4). Finally, the paper finishes in Section 5 with conclusion and discussion.

2 The Hyperedge Event Model

In this section we define the HEM by describing a process according to which hyperedge events are generated in continuous time. Data generated under the HEM consists of D unique edges. A single edge, indexed by $d \in [D]$, is represented by the three components: the sender $a_d \in [A]$, an indicator vector of receivers $\mathbf{r}_d = \{u_{dr}\}_{r=1}^A$, and the timestamp $t_d \in (0, \infty)$. For simplicity, we assume that edges are ordered by time such that $t_d \leq t_{d+1}$. While the model can be applied for two types of hyperedges—edges with (1) one sender and one or more receivers, and (2) one or more senders and one receiver—here we only present the generative process for those involving one sender and one or more receivers (i.e., multicast). As a part of the generating process, we draw candidate receivers and candidate timestamps for each sender, which are the auxiliary variables

in the HEM that serve as candidate data. Finally, we generate the observable data from the candidate data by taking the sender who generated the minimum candidate timestamp as an observed sender, the corresponding receivers as observed receivers, and the observed time as the timestamp of previous edge plus the smallest time generated. We explain these steps in more detail below. For the latter type of hyperedges—one receiver and one or more senders, we treat a_d to be an indicator vector of senders $\mathbf{a}_d = \{u_{dr}\}_{r=1}^A$ and r_d to be the single receiver, and then follow the alternative generative process provided in Appendix A, which is analogous to the process used for multicasts.

2.1 Candidate Receivers

For every possible sender-receiver pair $(a, r)_{a \neq r}$, we define the “receiver intensity”—an approximate inverse logit of the probability that edge d is being sent from sender a to receiver r —as a linear combination of statistics relevant to the receiver selection process:

$$\lambda_{adr} = \mathbf{b}^\top \mathbf{x}_{adr}, \quad (2.1)$$

where \mathbf{b} is a P -dimensional vector of coefficients and \mathbf{x}_{adr} is a set of receiver selection features. Such features can capture common network processes that include, but are not limited to, popularity, reciprocity, and transitivity, as well as the effects of attribute features defined on the sender, receivers, and/or sender-receiver pairs (e.g., the gender of the sender and/or receiver, an indicator of whether the sender is a supervisor of the receiver’s). In addition, we include an intercept term to account for the average (or baseline) number of receivers. We place a Normal prior $\mathbf{b} \sim N(\boldsymbol{\mu}_b, \Sigma_b)$.

A basic assumption of the HEM is that, at any given time, everyone is planning to send a document to one or more receivers. So for each edge d and each sender a , the HEM draws a list of candidate receivers that would be receivers on edge d if a were the sender. For an edge d , we first define an $A \times A$ matrix \mathbf{u}_d where the a^{th} row denotes sender a ’s receiver vector of zeros and 1’s—i.e., 1’s imply the willingness of sender a to send an edge to whom. We then assume that each receiver vector \mathbf{u}_{ad} comes from a modification of the multivariate Bernoulli (MB) distribution (Dai et al., 2013)—a model that has been used to model graphs in which the state of each edge indicator is drawn independently from an edge-specific Bernoulli distribution. In order to avoid drawing edges with no receivers, we define a probability measure “ MB_G ” motivated by the Gibbs measure (Fellows and Handcock, 2017). The probability measure we define amounts to a non-empty Gibbs measure, in which the all-zero vector is excluded from the support of the multivariate Bernoulli distribution. As a result, this measure helps us to 1) allow a sender to select multiple receiver for a single edge, 2) force the sender to select at least one receiver, and 3) ensure a tractable normalizing constant. To be specific, we draw a binary vector $\mathbf{u}_{ad} = (u_{ad1}, \dots, u_{adA})$

$$\mathbf{u}_{ad} \sim \text{MB}_G(\boldsymbol{\lambda}_{ad}), \quad (2.2)$$

where $\boldsymbol{\lambda}_{ad} = \{\lambda_{adr}\}_{r=1}^A$. In particular, we define $\text{MB}_G(\boldsymbol{\lambda}_{ad})$ as

$$\Pr(\mathbf{u}_{ad} | \mathbf{b}, \mathbf{x}_{ad}) = \frac{1}{Z(\boldsymbol{\lambda}_{ad})} \exp \left(\log(I(\|\mathbf{u}_{ad}\|_1 > 0)) + \sum_{r \neq a} \lambda_{adr} u_{adr} \right), \quad (2.3)$$

where $Z(\boldsymbol{\lambda}_{ad}) = \prod_{r \neq a} (\exp(\lambda_{adr}) + 1) - 1$ is the normalizing constant and $\|\cdot\|_1$ is the l_1 -norm. Again, this is equivalent to assuming that each u_{adr} is drawn with the probability of 1 being $\text{logit}(\lambda_{adr})$, excluding the case when $u_{adr} = 0$ for all $r \in [A]$. We provide detailed derivation steps for the normalizing constant in Appendix B.

2.2 Candidate Timestamps

Similarly, for each sender and edge combination, the HEM draws a candidate time at which the edge would be generated. The timing rate for sender a and edge d is

$$\mu_{ad} = g^{-1}(\boldsymbol{\eta}^\top \mathbf{y}_{ad}), \quad (2.4)$$

where $\boldsymbol{\eta}$ is a Q -dimensional vector of coefficients with a Normal prior $\boldsymbol{\eta} \sim N(\boldsymbol{\mu}_\eta, \Sigma_\eta)$, \mathbf{y}_{ad} is a set of event timing features—covariates that could affect timestamps of the edge, and $g(\cdot)$ is the appropriate link function such as identity, log, or inverse.

In modeling “when,” we do not directly model the timestamp t_d . Instead, we assume that each sender’s “time increment”—i.e., waiting time to next interaction since t_{d-1} —is drawn from a specific distribution in the exponential family. We define the time increment from edge $d-1$ to edge d as τ_d (i.e., $\tau_d = t_d - t_{d-1}$) and specify the distribution of candidate timestamps with sender-specific mean μ_{ad} . Following the generalized linear model (GLM) framework (Nelder and Baker, 1972), we assume the mean and variance of the τ_{ad} satisfy

$$\begin{aligned} E(\tau_{ad}) &= \mu_{ad}, \\ V(\tau_{ad}) &= V(\mu_{ad}), \end{aligned} \quad (2.5)$$

where τ_{ad} here is a positive real number. Possible choices of distribution include exponential, Weibull, gamma, and log-normal¹ distributions, which are commonly used in time-to-event modeling (Rao, 2000; Rizopoulos, 2012). Based on the specific distribution, we may need another latent variable to draw the time increment beyond the features used to model the rate to account for the variance of time increments $V(\mu)$ —e.g., shape parameter k for Weibull, shape parameter θ for gamma, and variance parameter σ_τ^2 for log-normal. We use $f_\tau(\cdot; \mu, V(\mu))$ and $F_\tau(\cdot; \mu, V(\mu))$ to denote the probability density function (p.d.f) and cumulative density function (c.d.f), respectively, with mean μ and variance $V(\mu)$.

2.3 Senders, Receivers, and Timestamps

Finally, we generate the actual sender, receivers, and timestamp by selecting the sender-receiver-set pair with the smallest time increment (Snijders, 1996):

$$\begin{aligned} a_d &= \operatorname{argmin}_a(\tau_{ad}), \\ \mathbf{r}_d &= \mathbf{u}_{a_d d}, \\ t_d &= t_{d-1} + \tau_{a_d d}. \end{aligned} \quad (2.6)$$

¹The log-normal distribution is not exponential family but can be used via modeling of $\log(\tau_d)$.

Algorithm 1 Generative Process: one sender and one or more receivers

Input: number of edges and nodes (D, A), covariates (\mathbf{x}, \mathbf{y}), and the coefficients ($\mathbf{b}, \boldsymbol{\eta}$)

```

for d=1 to D do
  for a=1 to A do
    for r=1 to A ( $r \neq a$ ) do
      set  $\lambda_{adr} = \mathbf{b}^\top \mathbf{x}_{adr}$ 
    end for
    draw  $\mathbf{u}_{ad} \sim \text{MB}_G(\boldsymbol{\lambda}_{ad})$ 
    set  $\mu_{ad} = g^{-1}(\boldsymbol{\eta}^\top \mathbf{y}_{ad})$ 
    draw  $\tau_{ad} \sim f_\tau(\mu_{ad}, V(\mu_{ad}))$ 
  end for
  if  $n \geq 2$  tied events then
    set  $a_d, \dots, a_{d+n-1} = \text{argmin}_a(\tau_{ad})$ 
    set  $\mathbf{r}_d = \mathbf{u}_{a_dd}, \dots, \mathbf{r}_{d+n-1} = \mathbf{u}_{a_{d+n-1}d}$ 
    set  $t_d, \dots, t_{d+n-1} = t_{d-1} + \min_a \tau_{ad}$ 
    jump to  $d = d + n$ 
  else
    set  $a_d = \text{argmin}_a(\tau_{ad})$ 
    set  $\mathbf{r}_d = \mathbf{u}_{a_dd}$ 
    set  $t_d = t_{d-1} + \min_a \tau_{ad}$ 
  end if
end for

```

Therefore, it is a sender-driven process in that the receivers and timestamp of an edge is jointly determined by the sender's urgency to send the edge to chosen receivers. Note that our generative process accounts for tied events such that in case of tied events (i.e., multiple senders draw exactly same candidate timestamps), we observe all of the tied events without assigning the orders of tied events. Algorithm 1 summarizes the entire generative process for directed edges with one sender and one or more receivers, and Figure 1 presents an illustrative example on how the d th edge is generated, assuming $t_{d-1} = 0$ for simplicity.

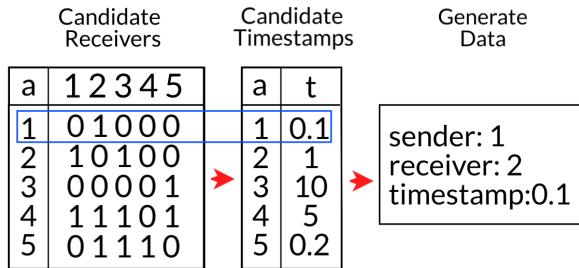


Figure 1: An illustrative example of the generative process of the HEM.

3 Posterior Inference

Our inference goal is to invert the generative process to obtain the posterior distribution over the latent variables—candidate receivers $\{\mathbf{u}_d\}_{d=1}^D$, coefficients for edge covariates \mathbf{b} , and coefficients for timestamp covariates $\boldsymbol{\eta}$ —conditioned on the observed data $\{(a_d, \mathbf{r}_d, t_d)\}_{d=1}^D$, covariates $\{(\mathbf{x}_d, \mathbf{y}_d)\}_{d=1}^D$, and hyperparameters $(\boldsymbol{\mu}_b, \Sigma_b, \boldsymbol{\mu}_\eta, \Sigma_\eta)$. We draw the samples using Markov chain Monte Carlo (MCMC) methods, repeatedly resampling the value of each latent variable from its conditional posterior via a Metropolis-within-Gibbs sampling algorithm. In this section, we provide each latent variable’s conditional posterior, and demonstrate how we perform the prior-posterior simulator test of [Geweke \(2004\)](#) for the HEM. At the end of this section, we provide the pseudocode of our MCMC algorithm in Algorithm 2.

3.1 Conditional Posteriors

Candidate Receivers

In the HEM, direct computation of the posterior densities for the latent variables \mathbf{b} and $\boldsymbol{\eta}$ —i.e., $P(\mathbf{b}|\mathbf{x}, \mathbf{a}, \mathbf{r}, \mathbf{t})$ and $P(\boldsymbol{\eta}|\mathbf{y}, \mathbf{a}, \mathbf{r}, \mathbf{t})$ —are not possible. However, it is possible to augment the data by candidate receivers \mathbf{u} such that we can obtain their conditional posterior by collapsing the known distributions— $P(\mathbf{b}, \mathbf{u}|\mathbf{x}, \mathbf{a}, \mathbf{r}, \mathbf{t})$ and $P(\boldsymbol{\eta}, \mathbf{u}|\mathbf{y}, \mathbf{a}, \mathbf{r}, \mathbf{t})$ —through integrating out \mathbf{u} . We adapt this common tool in Bayesian statistics called “data augmentation” ([Tanner and Wong, 1987](#); [Neal and Kypraios, 2015](#)). Since u_{adr} is a binary random variable, its new value may be sampled directly from a multinomial distribution with probabilities

$$\begin{aligned}\Pr(u_{adr} = 1 | \mathbf{u}_{ad \setminus r}, \mathbf{b}, \mathbf{x}, \mathbf{a}, \mathbf{r}, \mathbf{t}) &\propto \exp(\lambda_{adr}); \\ \Pr(u_{adr} = 0 | \mathbf{u}_{ad \setminus r}, \mathbf{b}, \mathbf{x}, \mathbf{a}, \mathbf{r}, \mathbf{t}) &\propto I(\|\mathbf{u}_{ad \setminus r}\|_1 > 0),\end{aligned}\tag{3.1}$$

where $I(\cdot)$ is the indicator function that is used to prevent from the instances where a sender chooses zero number of receivers.

Coefficients for edge covariates

Unlike the candidate receivers above, new values for \mathbf{b} cannot be sampled directly from its conditional posterior, but may instead be obtained using the Metropolis–Hastings (M-H) algorithm. Assuming an uninformative prior (i.e., $N(0, \infty)$), the conditional posterior over \mathbf{b} is

$$\Pr(\mathbf{b} | \mathbf{u}, \mathbf{x}, \mathbf{a}, \mathbf{r}, \mathbf{t}) \propto \prod_{d=1}^D \prod_{a=1}^A \frac{1}{Z(\boldsymbol{\lambda}_{ad})} \exp \left(\log(I(\|\mathbf{u}_{ad}\|_1 > 0)) + \sum_{r \neq a} \lambda_{adr} u_{adr} \right).\tag{3.2}$$

Algorithm 2 MCMC Algorithm

Input: number of outer and inner iterations (O, I_1, I_2) and initial values of $(\mathbf{u}, \mathbf{b}, \boldsymbol{\eta})$

```

for o=1 to O do
  for d=1 to D do
    for a = 1 to A do
      for r = 1 to A (r ≠ a) do
        update  $u_{adr}$  using Gibbs update—equation (3.1)
      end for
    end for
    end for
    for n=1 to  $I_1$  do
      update  $\mathbf{b}$  using M-H algorithm—equation (3.2)
    end for
    for n=1 to  $I_2$  do
      update  $\boldsymbol{\eta}$  using M-H algorithm—equation (3.3) or (3.4)
    end for
    if extra parameter for  $V(\mu)$  then
      update the variance parameter using M-H algorithm—equation (3.3) or (3.4)
    end if
  end for
  summarize the results with the last chain of  $\mathbf{b}$  and  $\boldsymbol{\eta}$ 

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Coefficients for timestamp covariates

Likewise, we use the M-H algorithm to update the latent variable $\boldsymbol{\eta}$. Assuming an uninformative prior $\boldsymbol{\eta}$ (i.e., $N(0, \infty)$), the conditional posterior for no-tied event case is

$$\Pr(\boldsymbol{\eta}|\mathbf{u}, \mathbf{y}, \mathbf{a}, \mathbf{r}, \mathbf{t}) \propto \prod_{d=1}^D \left(f_\tau(\tau_d; \mu_{ad}, V(\mu_{ad})) \times \prod_{a \neq ad} (1 - F_\tau(\tau_d; \mu_{ad}, V(\mu_{ad}))) \right), \quad (3.3)$$

where $f_\tau(\tau_d; \mu_{ad}, V(\mu_{ad}))$ is the probability that the d^{th} observed time increment comes from the specified distribution $f_\tau(\cdot)$ with the observed sender's mean μ_{ad} , and $\prod_{a \neq ad} (1 - F_\tau(\tau_d; \mu_{ad}, V(\mu_{ad})))$ is the probability that the rest of (unobserved) senders for event d all draw time increments greater than τ_d . Moreover, under the existence of tied-event, the conditional posterior of $\boldsymbol{\eta}$ is written as

$$\begin{aligned} \Pr(\boldsymbol{\eta}|\mathbf{u}, \mathbf{y}, \mathbf{a}, \mathbf{r}, \mathbf{t}) &\propto \prod_{m=1}^M \left(\prod_{d: t_d = t_m^*} f_\tau(t_m^* - t_{m-1}^*; \mu_{ad}, V(\mu_{ad})) \right. \\ &\quad \times \left. \prod_{a \notin \{ad\}_{d: t_d = t_m^*}} (1 - F_\tau(t_m^* - t_{m-1}^*; \mu_{ad}, V(\mu_{ad}))) \right), \end{aligned} \quad (3.4)$$

where t_1^*, \dots, t_M^* are the unique timepoints across D events ($M \leq D$). If $M = D$ (i.e., no tied events), equation (3.4) reduces to equation (3.3). Note that when we have the

latent variable to quantify the variance in time increments $V(\mu)$ (based on the choice of timestamp distribution in Section 2.2), we also use equation (3.3) (or equation (3.4) in case there exist tied events) for the additional M-H update—e.g., $\Pr(k|\boldsymbol{\eta}, \mathbf{u}, \mathbf{y}, \mathbf{a}, \mathbf{r}, \mathbf{t})$ for Weibull, $\Pr(\theta|\boldsymbol{\eta}, \mathbf{u}, \mathbf{y}, \mathbf{a}, \mathbf{r}, \mathbf{t})$ for gamma, and $\Pr(\sigma_\tau^2|\boldsymbol{\eta}, \mathbf{u}, \mathbf{y}, \mathbf{a}, \mathbf{r}, \mathbf{t})$ for log-normal distribution.

3.2 Getting It Right (GiR) Test

Software development is integral to the objective of applying the HEM to real world data. Code review is a valuable process in any research computing context, and the prevalence of software bugs in statistical software is well documented (e.g., Altman et al., 2004; McCullough, 2009). With highly complex models such as the HEM, there are many ways in which software bugs can be introduced and go unnoticed. As such, we present a joint analysis of the integrity of our generative model, sampling equations, and software implementation.

Geweke (2004) introduced the “Getting it Right” (GiR) test—a joint distribution test of posterior simulators which can detect errors in sampling equations as well as coding errors—and it has been used to test the implementation of Bayesian inference algorithms (Zhao et al., 2016). The test involves comparing the distributions of variables simulated from two joint distribution samplers, which we call “forward” and “backward” samples. The forward sampler draws unobservable variables from the prior and then generates the observable data conditional on the unobservables. The backward sampler alternates between the inference and an observables simulator, by running the inference code on observable data to obtain posterior estimates of the unobservable variables and then re-generating the observables given the inferred unobservables. The backward sampler is initialized by running an iteration of inference on observables drawn directly from the prior. Since the only information on which both the forward and backward samplers are based is the prior, if the sampling equations are correct and the code is implemented without bugs, each variable should have the same distribution in the forward and backward samples.

In the forward samples, both observable and unobservable variables are generated using Algorithm 1. In the backward samples, unobservable variables are generated using the sampling equations for inference, which we derived in Section 3.1. For each forward and backward sample that consists of D number of edges, we save these statistics:

1. Mean of observed receiver sizes $\|\mathbf{r}_d\|_1$ across $d = 1, \dots, D$,
2. Variance of observed receiver sizes $\|\mathbf{r}_d\|_1$ across $d = 1, \dots, D$,
3. Mean of time increments τ_d across $d = 1, \dots, D$,
4. Variance of time increments τ_d across $d = 1, \dots, D$,
5. b_p value used to generate the samples $p = 1, \dots, P$,
6. η_q value used to generate the samples $q = 1, \dots, Q$,

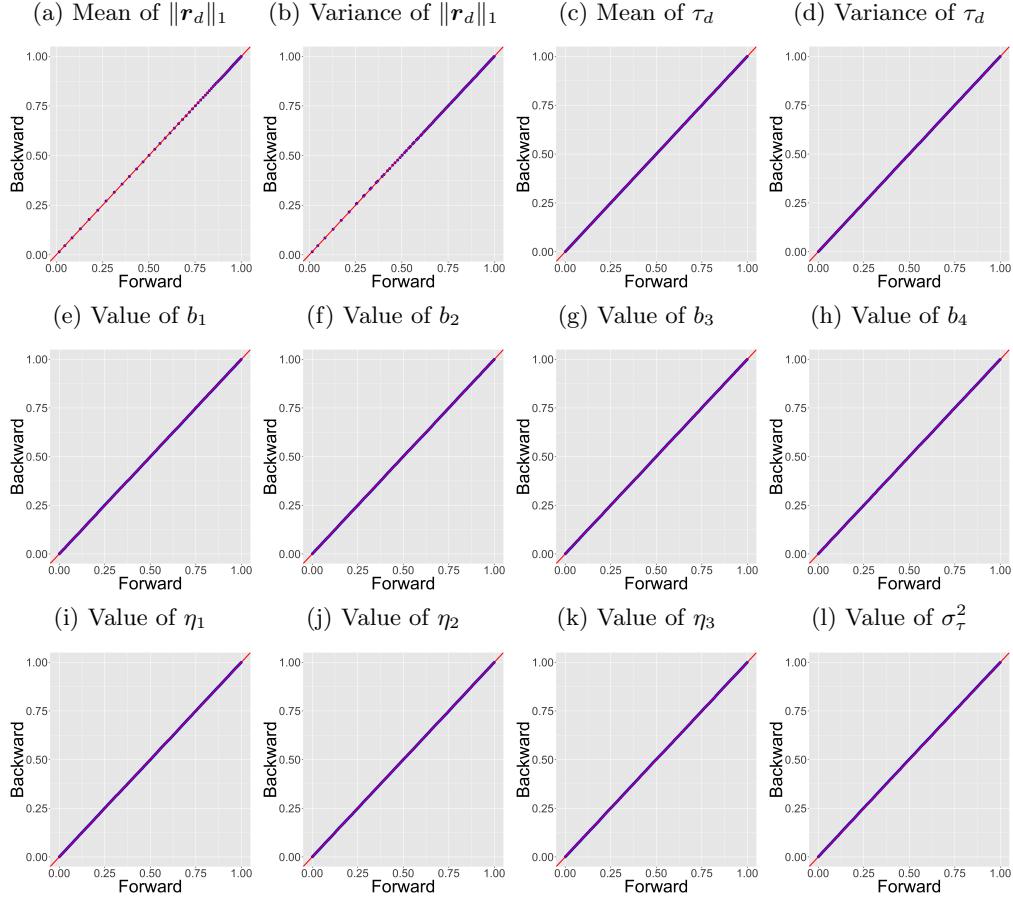


Figure 2: Probability-probability (PP) plots for the GiR test statistics.

7. σ_τ^2 value used to generate the samples in log-normal distribution

To keep the computational burden of re-running thousands of rounds of inference manageable, we run the GiR using a relatively small artificial sample, consisting of $D = 100$ edges, $A = 5$ actors, $P = 4$ number of edge covariates, and $Q = 3$ number of timestamp covariates per each forward or backward samples, using log-normal distribution for the time increments f_τ . We generated 10^5 sets of forward and backward samples, and then calculated 1,000 quantiles for each of the statistics. We also calculated t-test and Mann-Whitney test p-values in order to test for differences in the distributions generated in the forward and backward samples. Before we calculated these statistics, we thinned our samples by taking every 9th sample starting at the 10,000th sample for a resulting sample size of 10,000, in order to reduce the autocorrelation in the Markov chains. In each case, if we observe a large p-value, this gives us evidence that the distributions

generated under forward and backward sampling have the same locations. We depict the GiR results using probability-probability (PP) plots. To compare two samples with PP plots we calculate the empirical quantile in each sample of a set of values observed across the two samples, then plot the sets of quantiles in the two samples against each other. If the two samples are from equivalent distributions, the quantiles should line up on a line with zero y -intercept, and unit slope (i.e., a 45-degree line). The GiR test results are depicted in Figure 2, which show that we pass the test on every statistic.

4 Application to Email Data

We now present a case study applying our method to Montgomery county government email data. Our data come from the North Carolina county government email dataset collected by [ben Aaron et al. \(2017\)](#) that includes internal email corpora covering the inboxes and outboxes of managerial-level employees of North Carolina county governments. Out of over twenty counties, we chose Montgomery County to 1) test our model using data with large proportion of hyperedges (16.76%), all of which are emails sent from one sender to two or more receivers, and 2) limit the scope of this initial application. To summarize, Montgomery County email network contains 680 emails, sent and received by 18 department managers over a period of 3 months (March–May) in 2012. For this case study, we formulate the edge covariates \mathbf{x} and timestamp covariates \mathbf{y} and ground them in illustrative examples. We then report a suite of experiments—out-of-sample prediction for model selection and posterior predictive checks—that test the HEM’s ability to form the posterior distribution over its latent variables. Finally, we demonstrate an exploratory analysis of Montgomery County email data using the model estimates to discover interesting patterns in organizational communication networks.

4.1 Covariates

Edge covariates

Given a set of network data, a primary modeling goal of the generative process for edges lies in determining which characteristics and behaviors are predictive of the receiver selection. This email application specifically give rise to the following question: “To what extent are nodal, dyadic or triadic network effects relevant to predicting future emails?” As an illustrative example, we form the receiver selection features \mathbf{x} for Montgomery County email data using some nodal, dyadic, and triadic covariates. First, as we want to test whether gender plays a role in receiver selection process, we include three nodal covariates—the gender information of sender and receiver, and their homophily indicator. Additionally, we include four interval-based nodal network covariates—outdegree of sender, indegree of receiver, hyperedge size of sender, and the interaction between outdegree and hyperedge size—to study the effect of nodal behaviors on future interactions. For dyadic and triadic network effects, we employ the network statistics in [Perry and Wolfe \(2013\)](#) and summarize past interaction behaviors based on the time interval prior to and including t_{d-1} . Specifically, our time interval tracks 7 days prior to the last email was sent $l_d = (t_{d-1} - 7 \text{ days}, t_{d-1}]$. For $a \in [A], r \in [A]$, and $d \in [D]$, we define 14

covariates for \mathbf{x}_{adr} :

1. intercept: $x_{adr1} = 1$;
2. gender of sender: $x_{adr2} = I(\text{gender of sender } a = \text{female})$;
3. gender of receiver: $x_{adr3} = I(\text{gender of receiver } r = \text{female})$;
4. gender homophily: $x_{adr4} = I(x_{adr2} = x_{adr3})$;
5. outdegree of sender: $x_{adr5} = \sum_{d':t_{d'} \in l_d} I(a_{d'} = a)$;
6. indegree of receiver: $x_{adr6} = \sum_{d':t_{d'} \in l_d} I(u_{d'r} = 1)$;
7. hyperedge size of sender: $x_{adr7} = \sum_{d':t_{d'} \in l_d} \sum_{r=1}^A I(a_{d'} = a) I(u_{d'r} = 1)$;
8. interaction between outdegree and hyperedge size: $x_{adr8} = x_{adr5} \times x_{adr7}$;
9. send: $x_{adr9} = \sum_{d':t_{d'} \in l_d} I(a_{d'} = a) I(u_{d'r} = 1)$;
10. receive: $x_{adr10} = \text{send}(r, a)$;
11. 2-send: $x_{adr11} = \sum_{h \neq a, r} \text{send}(a, h) \text{send}(h, r)$;
12. 2-receive: $x_{adr12} = \sum_{h \neq a, r} \text{send}(h, a) \text{send}(r, h)$;
13. sibling: $x_{adr13} = \sum_{h \neq a, r} \text{send}(h, a) \text{send}(h, r)$;
14. cosibling: $x_{adr14} = \sum_{h \neq a, r} \text{send}(a, h) \text{send}(r, h)$;

where $I(\cdot)$ is an indicator function. The network statistics (5–14) are designed so that their coefficient have a straightforward interpretation. The statistics “outdegree” and “indegree” measures the gregariousness and popularity effect of the node by counting the number of emails sent from a and received by r , respectively, within the last 7 days. Moreover, in order to capture individual tendency of send emails to two or more receivers, we include the statistic “hyperedge size”—the number of emails sent from a within last 7 days where emails with k number of receivers are counted as k separate emails—as a variant of outdegree statistic, accounting for hyperedges. We also include the interaction term between outdegree and hyperedge size. Dyadic statistics “send” and “receive” are defined as above such that these covariates measure the number of emails sent from a to r and r to a , respectively, within the last 7 days. In the example of triadic statistics, the covariate “2-send” counts the pairs of emails involving some node h distinct from a and r such that emails from a to h and h to r are both observed within the last 7 days. Other triadic covariates behave similarly, and their interpretations are also analogous, which are illustrated in Figure 3.

$$\begin{array}{ll}
 \textbf{2-send} & \sum_h (a \rightarrow h \rightarrow r) \\
 \textbf{2-receive} & \sum_h (a \leftarrow h \leftarrow r) \\
 \textbf{sibling} & \sum_h (h \xrightarrow{a} r) \\
 \textbf{cosibling} & \sum_h (h \xleftarrow{a} r)
 \end{array}$$

Figure 3: Visualization of triadic statistics: 2-send, 2-receive, sibling, and cosibling.

Timestamp covariates

For the event timing features \mathbf{y} introduced in Section 2.2, we identify a set of covariates which may possibly effect “time to the next email.” Similar to the edge covariates, we include nodal statistics which are time-invariant (such as gender or manager status) or time-dependent (such as the network statistics used for \mathbf{x}). In addition, we select some edge-specific covariates based on the temporal aspect of the $(d - 1)^{th}$ email—e.g., whether the previous email was sent (1) during the weekend and (2) before or past midday (AM/PM)—since we expect the email interactions within county government to be less active during the weekend and in the evening. To be specific, the timestamp statistics are defined as

1. intercept: $y_{ad1} = 1$;
2. gender of sender: $y_{ad2} = I(\text{gender of sender } a = \text{female})$;
3. manager status of sender: $y_{ad3} = I(\text{sender } a \text{ is the County Manager})$;
4. outdegree of sender: $y_{ad4} = \sum_{d':t_{d'} \in l_d} I(a_{d'} = a)$;
5. indegree of sender: $y_{ad5} = \sum_{d':t_{d'} \in l_d} I(u_{d'a} = 1)$;
6. weekend indicator of previous email: $y_{ad6} = I(t_{d-1} \text{ is during the weekend})$;
7. PM indicator of previous email: $y_{ad7} = I(t_{d-1} \text{ in PM})$.

Note that our generative process for timestamps in Section 2.2 is sender-oriented process where the sender determines “when to send the email,” thus we incorporate network statistics that depends on a only—outdegree of sender a and indegree of sender a (not r this time).

4.2 Model Selection

The HEM has many component parts that need to be specified by the user (i.e., the receiver selection features \mathbf{x} , the selection of the event timing features \mathbf{y} , and the distribution of time increments f). Many of these components will be specified based on user expertise (e.g., regarding which features would drive receiver selection), but some decisions may require a data-driven approach to model specification. For example, though

theoretical considerations may inform the specification of features, subject-matter expertise is unlikely to inform the decision regarding the family of the event time distribution. Furthermore, since different distribution families (and model specifications more generally) may involve different size parameter spaces, any data-driven approach to model comparison must guard against over-fitting the data. In this section we present a general-purpose approach to evaluating the HEM specification using out-of-sample prediction. We illustrate this approach by comparing alternative distributional families for the event timing component of the model. Here, we specifically compare the predictive performance from two distributions—log-normal and exponential. We particularly choose the log-normal distribution based on some exploratory analysis (e.g., histogram and simple regressions) on raw time increments data, and take exponential distribution as an alternative since exponential is the most commonly specified distribution for time-to-event data which is also used in the stochastic actor-oriented models (SAOMs) (Snijders, 1996) as well as their extensions (Snijders et al., 2007).

We evaluate the model's ability to predict edges and timestamps from Montgomery County email data, conditioned on their “training” part of the data. To perform the experiment, we separately formed a test split of each three model components—sender, receivers, and timestamps—by randomly selecting “test” data with probability $p = 0.1$. Any missing variables were imputed by drawing samples from their conditional posterior distributions, given the observed data, model estimates, and current values of test data. We then run inference to update the latent variables given the imputed and observed data. We iterate the two steps—imputation and inference—multiple times to obtain enough number of estimates for test data. Algorithm 3 outlines this procedure in detail. We run the experiment and measure the predictive performance of two separate time distributions using $N = 500$ predicted samples, by comparing the predictions in terms of classification accuracy in predicting the senders and receivers, as well as prediction error in the timestamps. We summarize the results of prediction experiments for missing senders, receivers, and timestamps in Figure 4. First, we compare the posterior probability of correct senders for each of missing emails $\{d : a_d = \text{NA}\}$, which corresponds to $\pi_{a_d} = P(a_d = a_d^{\text{obs}} | \cdot)$ in Algorithm 3. We call this measure as “correct sender posterior probability.” On Figure 4 (a), we draw boxplots for the distribution of mean correct sender posterior probability—i.e., $\hat{\pi}_{a_d} = \frac{1}{N} \sum_{n=1}^N \pi_{a_d}^n$ —across the missing emails. The results show that both log-normal and exponential distributions achieve better predictive performance for missing senders compared to what is expected under random guess (i.e., choose one out of A possible senders = 1/18), with log-normal distribution showing higher performance than exponential distribution. Secondly, since the receiver vector is binary, we compute F_1 scores for missing receiver indicators (i.e., all d and r with $r_{dr} = \text{NA}$) by taking the harmonic mean of precision and recall:

$$F_1 = 2 \cdot \frac{\text{precision} \cdot \text{recall}}{\text{precision} + \text{recall}}, \text{ where} \\ \text{recall} = \frac{\text{TP}}{\text{TP} + \text{FN}} \text{ and precision} = \frac{\text{TP}}{\text{TP} + \text{FP}}, \quad (4.1)$$

with TP denoting true positive (i.e., $r_{dr}^{\text{obs}} = r_{dr}^{\text{pred}} = 1$), FN denoting false negative (i.e., $r_{dr}^{\text{obs}} = 1$ but $r_{dr}^{\text{pred}} = 0$), and FP denoting false positive (i.e., $r_{dr}^{\text{obs}} = 0$ but $r_{dr}^{\text{pred}} = 1$).

Algorithm 3 Out-of-Sample Predictions

Input: data $\{(a_d, \mathbf{r}_d, t_d)\}_{d=1}^D$, number of new data to generate R , and initial values of $(\mathbf{b}, \boldsymbol{\eta}, \mathbf{u}, \sigma_\tau^2)$

Test splits (with $p = 0.1$):
 draw test senders (out of D senders)
 draw test receivers (out of $D \times (A - 1)$ receiver indicators $\{\{\mathbf{r}_{dr}\}_{r \in [A] \setminus a_d}\}_{d=1}^D$)
 draw test timestamps (out of D timestamps)
 set the test data as “missing” (NA)

Imputation and inference:

```

for  $r = 1$  to  $R$  do
  for  $d = 1$  to  $D$  do
    if  $a_d = \text{NA}$  then
      for  $a = 1$  to  $A$  do
        compute  $\pi_a = P(a_d = a | \cdot)$  using equation (3.3) (without the product term)
      end for
      draw  $a_d \sim \text{multinomial}(\pi_a)$ 
    end if
    for  $r \in [A] \setminus a_d$  do
      if  $r_{dr} = \text{NA}$  then
        draw  $r_{dr}$  from multinomial with probability  $P(r_{dr} = 1 | \cdot)$  and  $P(r_{dr} = 0 | \cdot)$ 
        using equation (3.1)
      end if
    end for
    if  $t_d = \text{NA}$  then
      draw  $\tau_d^{new}$  from equation (3.3) (without the product term) via importance sampling2
    end if
    run inference and update  $(\mathbf{u}, \mathbf{b}, \boldsymbol{\eta})$  given the imputed and observed data
  end for
  store the estimates for test data
end for

```

Although the generative process for edges (Section 2.1) is not directly affected by the choice of timestamp distribution, Figure 4 (b) reveals slight difference between log-normal and exponential in their performance in predicting missing receiver indicators, where log-normal on average outperforms exponential. Finally, the prediction error for d^{th} missing timestamp is estimated using the median of absolute relative errors³ across $N = 500$ predictions:

$$e_{\tau_d} = \text{median} \left(\left| \frac{\tau_d^{obs} - \tau_d^{pred_1}}{\tau_d^{obs}} \right|, \dots, \left| \frac{\tau_d^{obs} - \tau_d^{pred_N}}{\tau_d^{obs}} \right| \right). \quad (4.2)$$

³Hyndman and Koehler (2006) refer to this as median absolute percentage error (MdAPE).

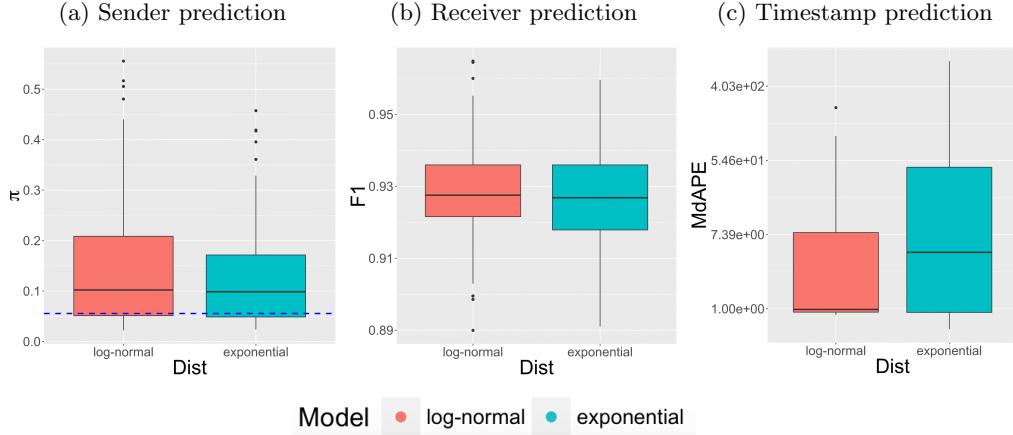


Figure 4: Comparison of predictive performance between log-normal and exponential distributions: (a) correct sender posterior probability from sender predictions, (b) F_1 scores from receiver predictions, and (c) median absolute relative error from timestamp predictions. Blue line in (a) represents the correct sender probability expected by random guess—i.e., $1/A = 1/18 \approx 0.056$.

Figure 4 (c) presents boxplots for the median absolute relative errors, where plot the estimates in a log-scale. Surprisingly, we have a huge benefit in the performance of timestamp prediction when we assume log-normal distribution for time-increments compared to exponential distribution. This difference can be explained by overdispersion in exponential distribution, because there exists greater variability in the time increments of emails than would be expected under exponential distribution. As illustrated above, we can use this out-of-sample prediction task for two usage—1) provide an effective answer to the question “how are we filling in the missing components of time-stamped network data?” and 2) offer one standard way to determine the distribution of time increments in Section 2.2.

4.3 Posterior Predictive Checks

In this section, we perform posterior predictive checks (PPC) (Rubin et al., 1984) to evaluate the appropriateness of our model specification for Montgomery County email data. We formally generated entirely new data, by simulating $N = 500$ synthetic email data $\{(a_d, r_d, t_d)\}_{d=1}^D$ from the generative process in Section 2, conditional upon a set of inferred latent variables from the inference in Section 4.4. For the test of goodness-of-fit in terms of network dynamics, we use multiple statistics that summarize meaningful aspects of the data: outdegree distribution, indegree distribution, receiver size distribution, and probability-probability (PP) plot for time increments.

Figure 5 illustrates the results of posterior predictive checks using log-normal distribu-

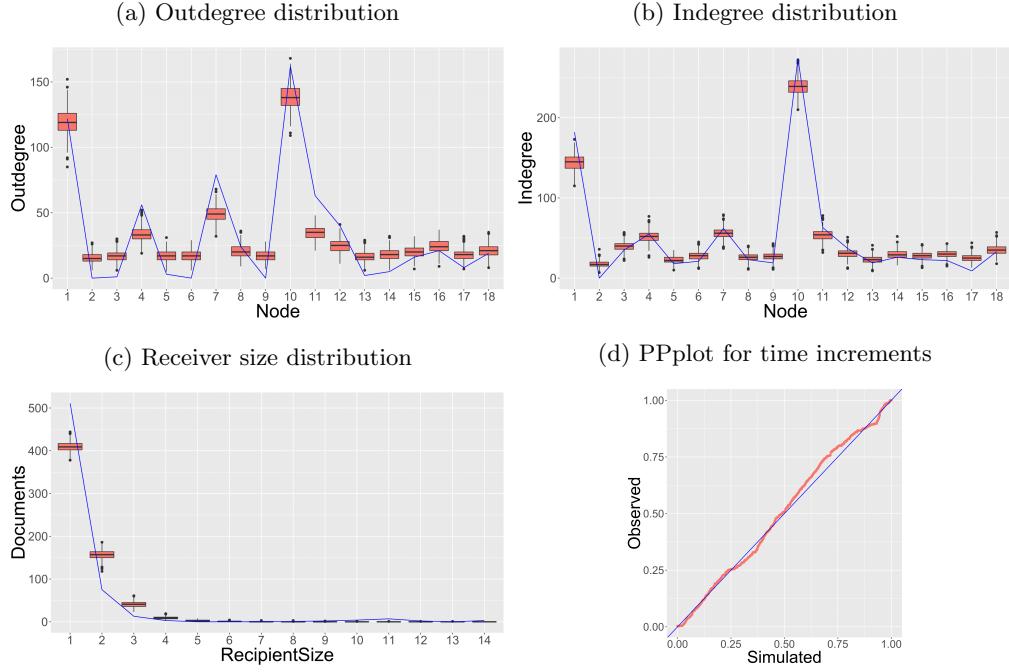


Figure 5: PPC results from log-normal distribution. Blue lines denote the observed statistics in (a)–(c) and denotes the diagonal line in (d).

tion, which shows better performance in Section 4.2. Upper two plots show node-specific posterior predictive degree distributions across $N = 500$ synthetic data, where the left one for outdegree statistic and the right plot is for indegree statistic. For both plots, the x-axis represents the nodes ($A = 1, \dots, 18$), and the y-axis represents the number of emails sent or received by the node. When compared with the observed outdegree and indegree statistics (red lines), our model recovers the overall distribution of sending and receiving activities across the nodes. For example, node 1 and 10 show significantly higher level of both sending and receiving activities relative to the rest, and the model-simulated data captures those big jumps, showing acceptable fit to the data. Outdegree distribution of some low-activity nodes are not precisely recovered, however, indegree distribution looks much better. Since we use more information in the receiver selection process (i.e., network effects) while we rely solely on minimum time increments when choosing the observed sender, these results are expected. Lower left plot is the distribution of receiver sizes, where the x-axis spans over the size of receivers 1 to 14 (which is the maximum size of observed receivers) and the y-axis denotes the number of emails with x-number of receivers. The result shows that our model is underestimating emails with one receiver while overestimating emails with two, three, and four receivers. One explanation behind what we observe is that the model is trying to recover so-called “broadcast” emails, which are the emails with ≥ 10 number of receivers, so that the

intercept estimate b_1 is slightly moved toward right. To improve the goodness of fit for receiver sizes, we can add more covariates (e.g., sender indicators) or assign strong prior structure on b_1 . In the end, the plot on the lower right is the PP plot for time increments, which depicts the two cumulative distribution functions—one for simulated time increments and another for observed time increments—against each other in order to assess how closely two data sets agree. Here, the closeness to the diagonal line connecting $(0, 0)$ and $(1, 1)$ gives a measure of difference between the simulated and observed time increments, and our PP plot shows that we have great performance in reproducing the observed time distribution. All our findings from predictive experiments in Section 4.2 are further revealed in the PPC from exponential distribution, where the PPC plots comparing log-normal and exponential distributions are attached in Appendix C.

4.4 Exploratory Analysis

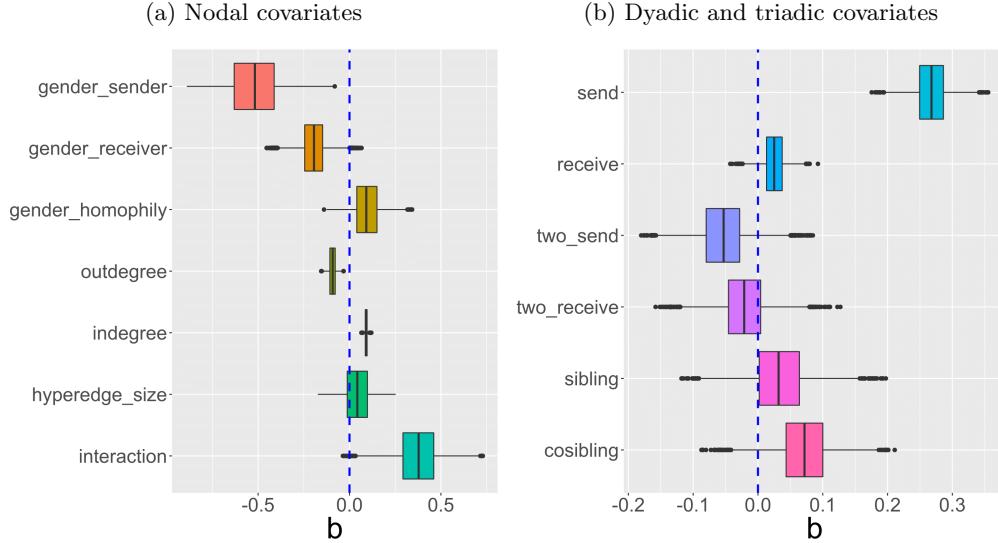
Based on the prediction experiments in Section 4.2, we choose to analyze Montgomery county email data and interpret the results using log-normal distribution. We assume weakly informative priors for latent variables such as $\boldsymbol{b} \sim N(\boldsymbol{\mu}_b = \mathbf{0}, \Sigma_b = 2 \times I_P)$, $\boldsymbol{\eta} \sim N(\boldsymbol{\mu}_\eta = \mathbf{0}, \Sigma_\eta = 2 \times I_Q)$, and $\sigma_\tau^2 \sim \text{inverse-Gamma}(a = 2, b = 1)$, and run MCMC algorithm in Algorithm 2 with $O = 55,000$ outer iterations with a burn-in of 15,000, where we thin by keeping every 40th sample. While the inner iterations for σ_τ^2 is fixed as 1, we specify the inner iterations $I_1 = 20$ for \boldsymbol{b} and $I_2 = 10$ for $\boldsymbol{\eta}$ to adjust for slower convergence rates. Convergence diagnostics including the traceplots and Geweke diagnostics (Geweke et al., 1991) are provided in Appendix D.

Coefficients for edge covariates

Figure 6 shows the boxplots summarizing posterior samples of \boldsymbol{b} , where Figure 6 (a) displays the coefficients for nodal covariates and 6 (b) displays the coefficients for dyadic and triadic covariates. Since we use the logit probability of λ_{adr} when generating candidate receivers in Section 2.1, we assume that

$$\text{logit}(\lambda_{adr}) = \log\left(\frac{\lambda_{adr}}{1 - \lambda_{adr}}\right) = b_1 + b_2 x_{adr2} \dots + b_{14} x_{adr14},$$

and can interpret the \boldsymbol{b} estimates in terms of odds ratio $\frac{\lambda_{adr}}{1 - \lambda_{adr}} = \exp(b_1 + b_2 x_{adr2} \dots + b_{14} x_{adr14})$. To begin with, we find out that the effect of nodal covariates “gender of sender” and “gender of receiver” are both significant in negative direction. **substantive interpretation on gender effects.** Next, the statistic “outdegree” also has significantly negative effect—i.e., if sender a sent n number of emails to anyone last week, then sender a is approximately $\exp(-0.109 \times n) \approx (0.897)^n$ times less likely to send an email to r —possibly due to its multicollinearity with the statistics “hyperedge size” where the two terms’ show significant interaction effect. This result is likely to appear under the following scenario: if sender a sent large number of emails over the last week but 1) not to r (excluding the effect of “send”) and the email of interest is not a 2) broadcast email which is originally aimed at large number of receivers (excluding the effect of “hyperedge size”), then sender a is less likely to send an email to receiver r .

Figure 6: Posterior distribution of \mathbf{b} estimates.

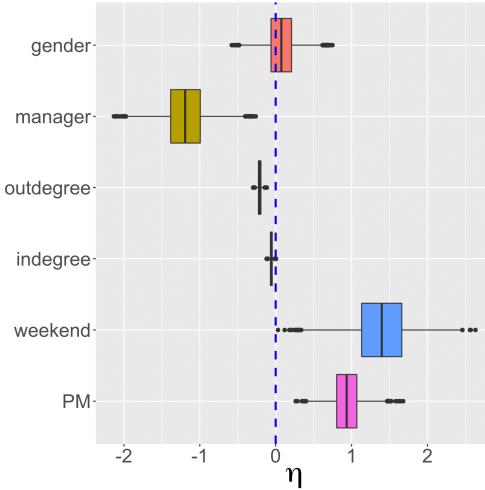
Similar interpretation can be applied for the covariates “indegree,” but in the opposite direction. For example, if the receiver r received n number of emails over the last week, sender a is $\exp(0.086 \times n) \approx (1.091)^n$ times more likely to send an email to r .

When we look at the effects of dyadic and triadic covariates, one thing we can check is that the effect of statistic “send” (i.e., number of times sender a sent emails to receiver r over the last week) is significant and positive with $\hat{b}_4 = 0.274$, implying that if sender a sent n number of emails to r last week, then sender a is approximately $\exp(0.274 \times n) \approx (1.315)^n$ times more likely to send an email to r . Another dyadic statistic “receive” is not statistically significant in that its 95% credible interval do not include 0. Surprisingly, none of the covariates for triadic network effects seem to have significant effect on the generative process of edges.

Coefficients for timestamp covariates

For timestamp covariates, Figure 7 shows the boxplots summarizing posterior samples of $\boldsymbol{\eta}$. Note that interpretations of the estimated coefficients for $\hat{\boldsymbol{\eta}}$ should be based on the specified timeunit of the dataset, which we use “hour” for Montgomery county email data. Moreover, since we assume log-normal distribution for time increments, we need careful interpretation regarding the timing equations

$$\log(\tau_{ad}) \sim N(\mu_{ad}, \sigma_\tau^2), \text{ with} \\ \mu_{ad} = \eta_1 + \eta_2 y_{ad2} + \dots + \eta_7 y_{ad7}.$$

Figure 7: Posterior distribution of η estimates.

To begin with, the posterior estimates of two temporal effects—“weekend” and “PM”—indicate that if the $(d - 1)^{th}$ email was sent during the weekend or after midday, then the time to d^{th} email is expected to take $\exp(1.552) \approx 4.722$ hours and $\exp(0.980) \approx 2.665$ hours longer, respectively, compared to their counterparts (i.e., weekdays and am). On the contrary, the covariates “manager”, “outdegree”, and “indegree” turn out to shorten the amount of time to next email. For example, being a county manager (i.e., the manager of department managers) lowers the expected value of $\log(\tau_{ad})$ by $\hat{\eta}_3 = -1.070$, since the manager in general sends or receives a lot more emails which may shorten the response time and many of those emails. This argument is also supported by the posterior estimates for “outdegree” and “indegree” statistics, where the estimated coefficients are approximately $\hat{\eta}_4 = -0.206$ and $\hat{\eta}_5 = -0.060$, respectively. Gender of the department manager is shown to have no significant effect on time increments. In addition, the posterior mean estimates for variance parameter σ_τ^2 in lognormal distribution is approximately $\hat{\sigma}_\tau^2 = 14.093$ with its 95% credible interval (12.709, 15.555), indicating that there exists large variability in the time increments of emails.

5 Conclusion

Motivated by a growing class of dynamic network models which deal with edges exchanged in continuous time, the hyperedge event model (HEM) can effectively learn the underlying dynamics in edge and timestamp formations, providing novel insights to the literature. First, we explicitly model hyperedges by deriving a receiver-selection distribution that forces the sender to select at least one receiver, which is more realistic approach for events with one sender and one or more receivers and one or more sender and one receiver compared to treating them as pure duplicates. In modeling the

timestamps (more precisely time increments) of events, our generalized linear model (GLM) based formulation offers new innovations by eliminating the need to stick with one parameter distribution (e.g., exponential distribution). To make better use of the proposed model, we provide an algorithm for predictive experiment that helps to learn which distribution better fits the data. To sum, the HEM's flexible generative process thus leads to more accurate and precise inference on model parameters.

We have demonstrated effectiveness of our model by analyzing Montgomery County government emails, where emails serve as the canonical example of directed edges with one sender and one or more receivers. The estimated effects for edge covariates reveal that the HEM is able to understand the structural dynamics similar to those used in the exponential random graph model (ERGM), but we also learn about the effect of timestamp covariates on waiting time to next event, by integrating a survival model for event timing. Although we illustrate the entire framework and application in the context of one type of hyperedges, one sender and one or more receivers, our model can be easily extended to allow the opposite case, one or more sender and one receiver, by slight modification of the generative process (shown in Appendix A). This extension involves promising applications including international sanctions and co-sponsorship of bills. Furthermore, while currently the model's applicability is limited to small-sized networks since we assume fixed set of nodes across entire timepoints, we can adjust the model to allow time-varying nodes or improve its computational efficiency such that the model is well applicable to large-scale networks. Considering the recent explosion of network dataset with huge number of nodes (e.g., online communications), the development of model adjustments for better applicability are the object of future work.

Acknowledgments

This work was supported in part by the University of Massachusetts Amherst Center for Intelligent Information Retrieval and in part by National Science Foundation grants DGE-1144860, SES-1619644, and CISE-1320219. Any opinions, findings, and conclusions or recommendations are those of the authors and do not necessarily reflect those of the sponsors.

Appendix

Appendix A: Alternative Generative Process

Algorithm 4 Generative Process: one receiver and one or more senders

```

Input: number of edges and nodes ( $D, A$ ), covariates ( $\mathbf{x}, \mathbf{y}$ ), and the coefficients ( $\mathbf{b}, \boldsymbol{\eta}$ )
for  $d=1$  to  $D$  do
  for  $r=1$  to  $A$  do
    for  $a=1$  to  $A$  ( $a \neq r$ ) do
      set  $\lambda_{adr} = \mathbf{b}^\top \mathbf{x}_{adr}$ 
    end for
    draw  $\mathbf{u}_{rd} \sim \text{MB}_G(\boldsymbol{\lambda}_{rd})$ 
    set  $\mu_{rd} = g^{-1}(\boldsymbol{\eta}^\top \mathbf{y}_{rd})$ 
    draw  $\tau_{rd} \sim f_\tau(\mu_{rd}, \sigma_\tau^2)$ 
  end for
  if  $n \geq 2$  tied events then
    set  $r_d = \text{argmin}_r(\tau_{rd})$ 
    set  $\mathbf{a}_d = \mathbf{u}_{rdd}, \dots, \mathbf{a}_{d+n-1} = \mathbf{u}_{r_{d+n-1}d}$ 
    set  $t_d, \dots, t_{d+n-1} = t_{d-1} + \min_r \tau_{rd}$ 
    jump to  $d = d + n$ 
  else
    set  $r_d = \text{argmin}_r(\tau_{rd})$ 
    set  $\mathbf{a}_d = \mathbf{u}_{rdd}$ 
    set  $t_d = t_{d-1} + \min_r \tau_{rd}$ 
  end if
end for

```

Appendix B: Normalizing Constant of MB_G

Our probability measure “ MB_G ”—the multivariate Bernoulli distribution with non-empty Gibbs measure—defines the probability of sender a selecting the binary receiver vector \mathbf{u}_{ad} as

$$\Pr(\mathbf{u}_{ad} | \mathbf{b}, \mathbf{x}_{ad}) = \frac{1}{Z(\boldsymbol{\lambda}_{ad})} \exp \left(\log(I(\|\mathbf{u}_{ad}\|_1 > 0)) + \sum_{r \neq a} \lambda_{adr} u_{adr} \right),$$

where the receiver intensity is a linear combination of edge covariates—i.e., $\lambda_{adr} = \mathbf{b}^\top \mathbf{x}_{adr}$ —as defined in Section 2.1.

To use this distribution efficiently, we derive a closed-form expression for $Z(\boldsymbol{\lambda}_{ad})$ that does not require brute-force summation over the support of \mathbf{u}_{ad} (i.e. $\forall \mathbf{u}_{ad} \in [0, 1]^A$). We recognize that if \mathbf{u}_{ad} were drawn via independent Bernoulli distributions in which $\Pr(u_{adr} = 1 | \mathbf{b}, \mathbf{x}_{ad})$ was given by $\text{logit}(\lambda_{adr})$, then

$$\Pr(\mathbf{u}_{ad} | \mathbf{b}, \mathbf{x}_{ad}) \propto \exp \left(\sum_{r \neq a} \lambda_{adr} u_{adr} \right).$$

This is straightforward to verify by looking at

$$\Pr(u_{adr} = 1 | \mathbf{u}_{ad[-r]}, \mathbf{b}, \mathbf{x}_{ad}) = \frac{\exp(\lambda_{adr})}{\exp(\lambda_{adr}) + 1}.$$

We denote the logistic-Bernoulli normalizing constant as $Z^l(\boldsymbol{\lambda}_{ad})$, which is defined as

$$Z^l(\boldsymbol{\lambda}_{ad}) = \sum_{\mathbf{u}_{ad} \in [0,1]^A} \exp\left(\sum_{r \neq a} \lambda_{adr} u_{adr}\right).$$

Now, since

$$\exp\left(\log\left(\mathbb{I}(\|\mathbf{u}_{ad}\|_1 > 0)\right) + \sum_{r \neq a} \lambda_{adr} u_{adr}\right) = \exp\left(\sum_{r \neq a} \lambda_{adr} u_{adr}\right),$$

except when $\|\mathbf{u}_{ad}\|_1 = 0$, we note that

$$\begin{aligned} Z(\boldsymbol{\lambda}_{ad}) &= Z^l(\boldsymbol{\lambda}_{ad}) - \exp\left(\sum_{\forall u_{adr}=0} \lambda_{adr} u_{adr}\right) \\ &= Z^l(\boldsymbol{\lambda}_{ad}) - 1. \end{aligned}$$

We can therefore derive a closed form expression for $Z(\boldsymbol{\lambda}_{ad})$ via a closed form expression for $Z^l(\boldsymbol{\lambda}_{ad})$. This can be done by looking at the probability of the zero vector under the logistic-Bernoulli model:

$$\frac{1}{Z^l(\boldsymbol{\lambda}_{ad})} \exp\left(\sum_{\forall u_{adr}=0} \lambda_{adr} u_{adr}\right) = \prod_{r \neq a} \left(1 - \frac{\exp(\lambda_{adr})}{\exp(\lambda_{adr}) + 1}\right).$$

Then, we have

$$\frac{1}{Z^l(\boldsymbol{\lambda}_{ad})} = \prod_{r \neq a} \frac{1}{\exp(\lambda_{adr}) + 1}.$$

Finally, the closed form expression for the normalizing constant is

$$Z(\boldsymbol{\lambda}_{ad}) = \prod_{r \neq a} (\exp(\lambda_{adr}) + 1) - 1.$$

Appendix C: Comparison of PPC Results: log-normal vs. exponential

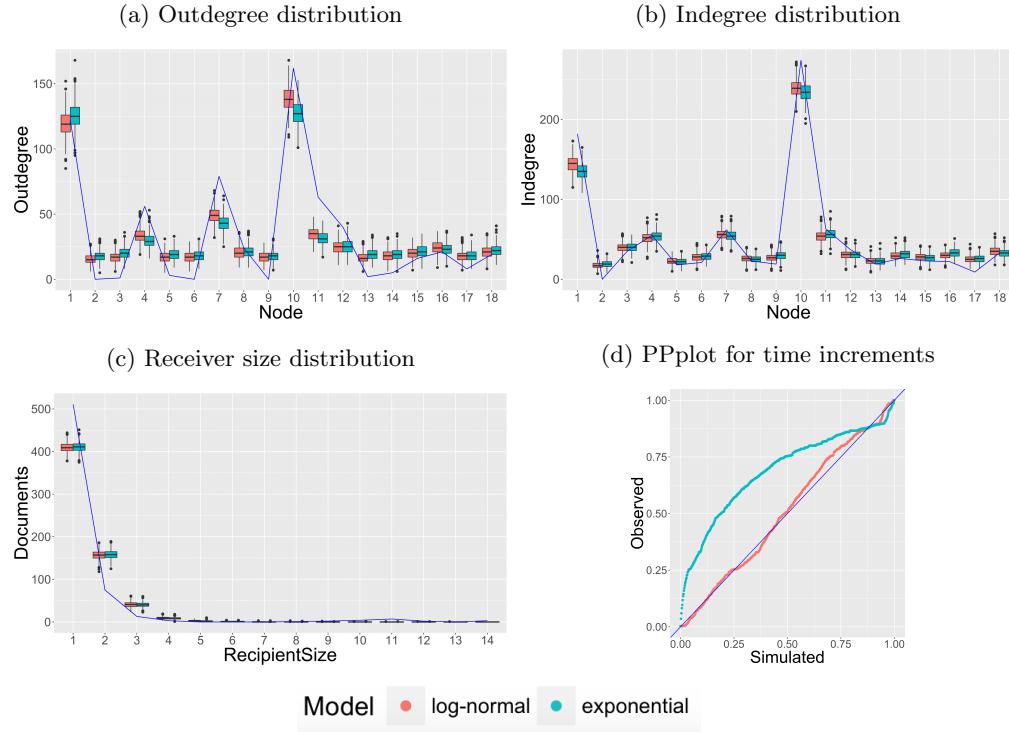


Figure 8: Comparison of PPC results between log-normal (red) and exponential (green) distributions. Blue lines denote the observed statistics in (a)–(c) and denotes the diagonal line in (d).

Appendix D: Convergence Diagnostics

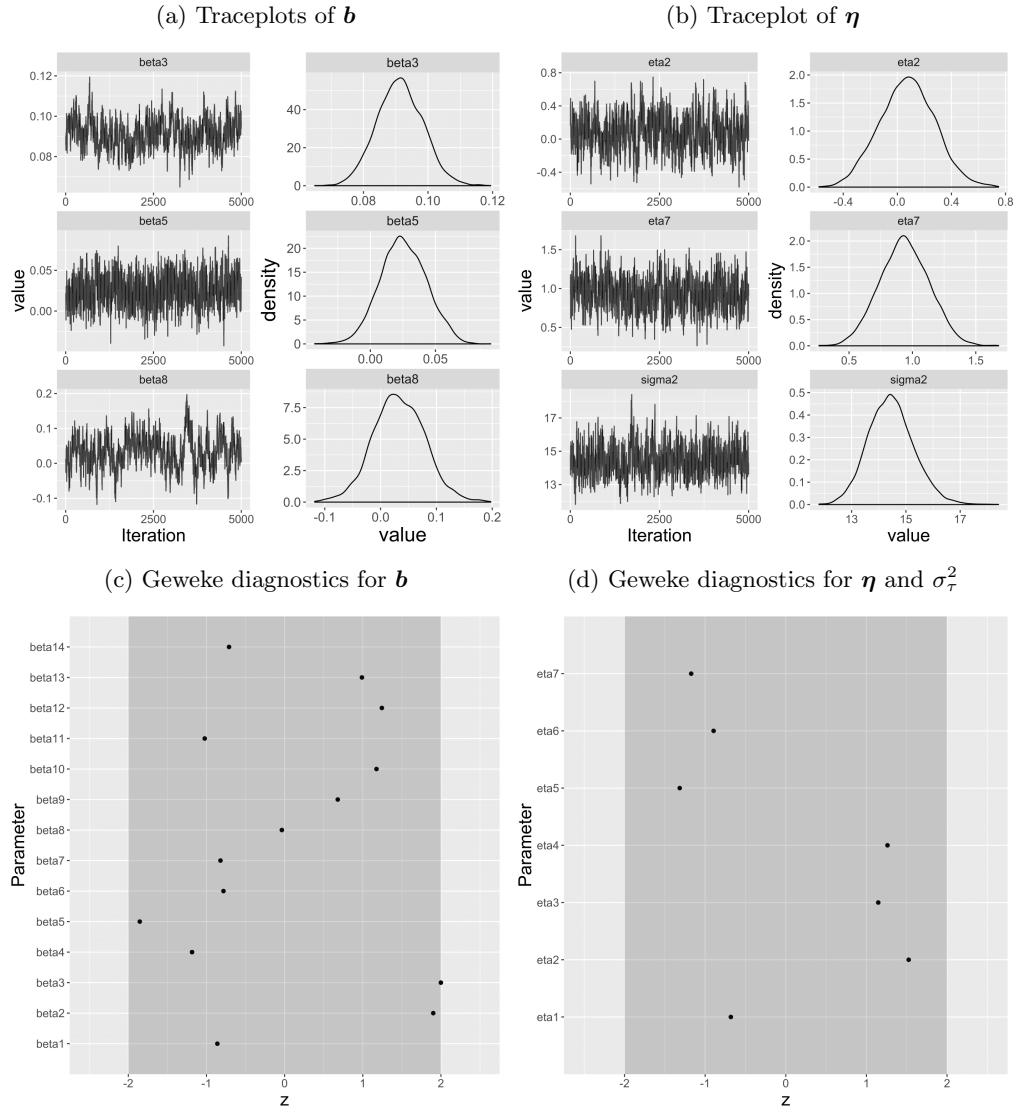


Figure 9: Convergence diagnostics from log-normal distribution.

Supplementary Material

Title of the Supplement A (<http://www.some-url-address.org/dowload/0000.zip>). Add description for supplement material.

References

- Altman, M., Gill, J., and McDonald, M. P. (2004). *Numerical issues in statistical computing for the social scientist*, volume 508. John Wiley & Sons. 8
- ben Aaron, J., Denny, M., Desmarais, B., and Wallach, H. (2017). “Transparency by Conformity: A Field Experiment Evaluating Openness in Local Governments.” *Public Administration Review*, 77(1): 68–77. 10
- Butts, C. T. (2008). “A RELATIONAL EVENT FRAMEWORK FOR SOCIAL ACTION.” *Sociological Methodology*, 38(1): 155–200. 1
- Cranmer, S. J., Heinrich, T., and Desmarais, B. A. (2014). “Reciprocity and the structural determinants of the international sanctions network.” *Social Networks*, 36: 5–22. 2
- Dai, B., Ding, S., Wahba, G., et al. (2013). “Multivariate bernoulli distribution.” *Bernoulli*, 19(4): 1465–1483. 3
- Fan, Y. and Shelton, C. R. (2009). “Learning continuous-time social network dynamics.” In *Proceedings of the Twenty-Fifth Conference on Uncertainty in Artificial Intelligence*, 161–168. AUAI Press. 2
- Fellows, I. and Handcock, M. (2017). “Removing Phase Transitions from Gibbs Measures.” In *Artificial Intelligence and Statistics*, 289–297. 3
- Fowler, J. H. (2006). “Legislative cosponsorship networks in the US House and Senate.” *Social Networks*, 28(4): 454–465. 2
- Geweke, J. (2004). “Getting it right: Joint distribution tests of posterior simulators.” *Journal of the American Statistical Association*, 99(467): 799–804. 6, 8
- Geweke, J. et al. (1991). *Evaluating the accuracy of sampling-based approaches to the calculation of posterior moments*, volume 196. Federal Reserve Bank of Minneapolis, Research Department Minneapolis, MN, USA. 17
- Ghoshal, G., Zlatić, V., Caldarelli, G., and Newman, M. (2009). “Random hypergraphs and their applications.” *Physical Review E*, 79(6): 066118. 2
- Hunter, D., Smyth, P., Vu, D. Q., and Asuncion, A. U. (2011). “Dynamic egocentric models for citation networks.” In *Proceedings of the 28th International Conference on Machine Learning (ICML-11)*, 857–864. 1
- Hyndman, R. J. and Koehler, A. B. (2006). “Another look at measures of forecast accuracy.” *International journal of forecasting*, 22(4): 679–688. 14
- Karypis, G., Aggarwal, R., Kumar, V., and Shekhar, S. (1999). “Multilevel hypergraph partitioning: applications in VLSI domain.” *IEEE Transactions on Very Large Scale Integration (VLSI) Systems*, 7(1): 69–79. 2
- McCullough, B. D. (2009). “The accuracy of econometric software.” *Handbook of computational econometrics*, 55–79. 8

- Neal, P. and Kypraios, T. (2015). “Exact Bayesian inference via data augmentation.” *Statistics and Computing*, 25(2): 333–347. [6](#)
- Nelder, J. A. and Baker, R. J. (1972). *Generalized linear models*. Wiley Online Library. [4](#)
- Newman, M. E., Forrest, S., and Balthrop, J. (2002). “Email networks and the spread of computer viruses.” *Physical Review E*, 66(3): 035101. [2](#)
- Partzsch, J. and Schüffny, R. (2012). “Developing structural constraints on connectivity for biologically embedded neural networks.” *Biological cybernetics*, 106(3): 191–200. [2](#)
- Perry, P. O. and Wolfe, P. J. (2013). “Point process modelling for directed interaction networks.” *Journal of the Royal Statistical Society: Series B (Statistical Methodology)*, 75(5): 821–849. [1](#), [2](#), [10](#)
- Rao, P. (2000). “Applied survival analysis: regression modeling of time to event data.” *Journal of the American Statistical Association*, 95(450): 681–681. [4](#)
- Rizopoulos, D. (2012). *Joint models for longitudinal and time-to-event data: With applications in R*. CRC Press. [4](#)
- Rubin, D. B. et al. (1984). “Bayesianly justifiable and relevant frequency calculations for the applied statistician.” *The Annals of Statistics*, 12(4): 1151–1172. [15](#)
- Snijders, T., Steglich, C., and Schweinberger, M. (2007). *Modeling the coevolution of networks and behavior*. na. [1](#), [13](#)
- Snijders, T. A. (1996). “Stochastic actor-oriented models for network change.” *Journal of mathematical sociology*, 21(1-2): 149–172. [1](#), [4](#), [13](#)
- Tanner, M. A. and Wong, W. H. (1987). “The calculation of posterior distributions by data augmentation.” *Journal of the American statistical Association*, 82(398): 528–540. [6](#)
- Vu, D. Q., Hunter, D., Smyth, P., and Asuncion, A. U. (2011). “Continuous-Time Regression Models for Longitudinal Networks.” In Shawe-Taylor, J., Zemel, R., Bartlett, P., Pereira, F., and Weinberger, K. (eds.), *Advances in Neural Information Processing Systems 24*, 2492–2500. Curran Associates, Inc. [1](#)
- Zhang, Z.-K. and Liu, C. (2010). “A hypergraph model of social tagging networks.” *Journal of Statistical Mechanics: Theory and Experiment*, 2010(10): P10005. [2](#)
- Zhao, T., Wang, Z., Cumberworth, A., Gsponer, J., de Freitas, N., Bouchard-Côté, A., et al. (2016). “Bayesian analysis of continuous time Markov chains with application to phylogenetic modelling.” *Bayesian Analysis*, 11(4): 1203–1237. [8](#)
- Zlatić, V., Ghoshal, G., and Caldarelli, G. (2009). “Hypergraph topological quantities for tagged social networks.” *Physical Review E*, 80(3): 036118. [2](#)