A Practical Introduction to Propensity Score Analysis using R

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About this event

- ► TI Methods Speaker Series page: Therapeutics Initiative
 - ▶ Dr. Carole Lunny
- SARGC page: Students and Recent Graduates Committee (SARGC) coordinate activities for the Statistical Society of Canada (SSC)'s student and recent graduate members
 - Md. Erfanul Hoque
 - Janie Coulombe

Outline

- 1. Data and Regression
- 2. Exact matching
- 3. Propensity score matching (4 steps)
- 4. Propensity score Reviews in different disease areas

[1] Right Heart Catheterization (RHC) Dataset

The dataset that we will use today is from Connors et al. (1996).

secure | biostat.mc.vanderbilt.edu/wiki/pub/Main/DataSets/rhc.htm

	Swangl	Right Heart Catheterization (RHC)
The effectiveness of right heart catheferization in the initial care of critically III patients. AC Cores, F. Egend, M.V. Dasson, C. Torras,	Sakude	Study Admission Date
	Dfalte	Date of Death
	Locable	Date of Last Contact
	Dschitte	Hospital Discharge Date
	Death	Death at any time up to 190 Days
	Ptid	Patient ID
	Frank E Harrell Jr Last medified: Fri Dec 27 16:14:18 EST 2002	

Notations

- Outcome Death (Y)
 - Death at any time up to 180 Days
- ► Treatment swang1 (A: Swan-Ganz catheter)
 - Whether or not a patient received a RHC
- ► Covariate list: L (age, sex, race, ...)
- Analysis strategy: matching RHC patients with non-RHC patients

[1] Load data

##

##

```
analytic.data <- readRDS("data/RHC.Rds")</pre>
dim(analytic.data)
## [1] 4767
               23
```

names (analytic.data)

[1] "age"

```
[4] "Disease.category" "DNR.status"
##
    [7] "Pr.2mo.survival"
##
```

```
[10] "Temperature"
```

```
"APACHE.II
"No.of.comorbidity" "DASI.2wk.
                     "Blood.pre:
```

"race"

"Heart.rate" "WBC.count"

"SEX"

```
"Pa02.by.F
"Creatinine
```

[13] "Respiratory.rate" ## ## [16] "PaCO2"

"pH"

[1] Inspecting data: Crude

```
##
            Stratified by Death
##
             level
                   Nο
                                 Yes
                                                     test
##
                    2013
                             2754
     n
     RHC (%) No RHC 1315 (65.3) 1268 (46.0)
                                              < 0.001
##
                   698 (34.7) 1486 (54.0)
##
             RHC
```

[1] Inspecting data: Some baseline variables

##

sex = Female (%)

```
baselinevars <- c("age", "sex", "race")</pre>
tab1 <- CreateTableOne(vars = baselinevars,
               data = analytic.data,
                strata = "Death", includeNA = TRUE,
                test = TRUE, smd = FALSE)
print(tab1, showAllLevels = FALSE, smd = FALSE)
##
                      Stratified by Death
##
                       No
                                     Yes
                                                   р
                                                          test
                       2013
                                     2754
##
     n
##
     age (%)
                                                   < 0.001
        [-Inf.50)
                      713 (35.4)
                                      400 (14.5)
##
```

[50,60)351 (17.4) 452 (16.4) ## [60,70)426 (21.2) 789 (28.6) ##

[70,80)382 (19.0) 750 (27.2) ## ## [80, Inf) 141 (7.0) 363 (13.2)

919 (45.7)

865 (31.4)

< 0.001

[1] Crude regression

```
## Variable Units OddsRatio CI.95 p-value
## RHC No RHC Ref
## RHC 2.21 [1.96;2.49] <1e-04</pre>
```

```
##
   Variable
                Units OddsRatio
                                      CI.95
                                                p-value
##
         RHC
               No RHC
                            Ref
                                                < 1e-04
##
                  RHC
                           2.71 [2.38;3.08]
##
         age [-Inf,50)
                            Ref
               [50,60) 3.56 [3.01;4.20]
##
                                                 < 1e-04
##
               [60,70)
                           0.74 [0.64;0.87]
                                              0.0001274
##
               [70,80)
                           1.33 [1.15;1.54]
                                                 < 1e-04
##
             [80. Inf)
                           1.06 [0.93;1.21]
                                              0.3633526
##
                 Male
                            Ref
         sex
               Female
                           0.49 [0.43; 0.56]
                                                < 1e-04
##
##
                white
                            Ref
       race
                           1.10 [0.93;1.31]
##
                black
                                              0.2666157
##
                other
                           0.97 [0.73;1.28]
                                              0.8068800
```

Continuous outcome (Y)

- ▶ treated group A = 1 (RHC)
- ightharpoonup control group A = 0 (no RHC)

Treatment effect = E[Y|A=1] vs. E[Y|A=0]

- Would only work if 2 groups are comparable / exchangeable / ignorable treatment assignment
- Randomization with enough sample size is one

Binary outcome (Y)

Treatment effect = prob[Y = 1|A = 1] vs. prob[Y = 1|A = 0]

In absence of randomization,

$$E[Y|A=1] - E[Y|A=0]$$

includes

- ▶ Treatment effect
- Systematic differences in 2 groups ('confounding')
 - ▶ Doctors may prescribe tx more to frail and older age patients.
 - ▶ In here, L = age is a confounder.

In absence of randomization, if age is a known issue

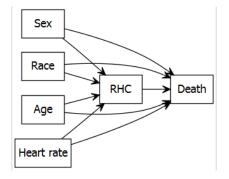
Causal effect for young

▶ E[Y|A = 1, L = younger age] - E[Y|A = 0, L = younger age]

Causal effect for old

ightharpoonup E[Y|A=1, L= older age] - E[Y|A=0, L= older age]

Conditional exchangeability; only works if L is measured



Using a rather inappropriate data as the trial was randomized. But we have modified the analytic data to introduce some bias!!

##

```
baselinevars <- c("age", "sex", "race", "Disease.category",</pre>
                   "Pr.2mo.survival", "No.of.comorbidity",
                   "Creatinine", "Albumin", "GComa.Score")
out.formula <- as.formula(paste("I(Death=='Yes')", "~",
                                paste(baselinevars,
                                       collapse = "+")))
out.formula
## I(Death == "Yes") ~ age + sex + race + Disease.category
```

```
## APACHE.III.score + Pr.2mo.survival + No.of.comorbid:
## DASI.2wk.prior + Temperature + Heart.rate + Blood.pr
```

Respiratory.rate + WBC.count + PaO2.by.FIO2 + PaCO2

##

##

##

##

##

##

##

##

##

```
fit2 <- glm(out.formula,</pre>
             family=binomial, data = analytic.data)
publish(fit2)
```

```
[60,70)
                                        0.73 [0.62;0.87]
##
```

[70,80)

Male

Female

white

black

other

ARF

CHF

MOSF

[80, Inf)

sex

race

Disease.category

0.0

0.

0.

0.0

0.4

1.13 [0.97;1.32]

1.12 [0.97;1.29]

0.48 [0.42; 0.55]

1.12 [0.93;1.36]

1.00 [0.74;1.35]

1.64 [1.24;2.16]

1.06 [0.89;1.26]

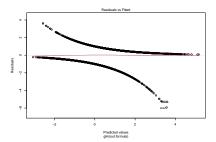
Ref

Ref

Ref

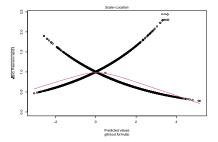
CI.95	${\tt OddsRatio}$	Units	Variable	##
	Ref	[-Inf,50)	age	##
[1.27;1.87]	1.54	[50,60)		##
F0 00 0 0 0 7		Fac = 5)		

plot(fit2, which =1)



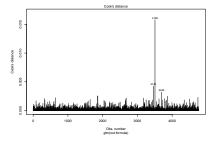
- curvilinear trends?
 - ▶ logistic regression IS curvilinear by nature

plot(fit2, which =3)



- ▶ red line is approximately horizontal?
- points have approximately equal spread around the red line?
 - more about detecting heteroscedasticity?

plot(fit2, which =4)



Cook's D estimates the influence of data points

[2] Alternate to Regression

How sure are you about the model-specification?

- ▶ Interaction?
- ► Polynomial?
- ▶ Potential solution?
 - Exact Matching

[2] Exact Matching: 2 variables

table(var.comb)

[1] 6

##

```
## var.comb
## blackFemale blackMale otherFemale otherMale whiteFemale
```

404

table(analytic.data\$RHC,var.comb)

331

```
## var.comb
## blackFemale blackMale otherFemale otherMale wh:
```

113

140

No RHC 161 239 50 61 ## RHC 170 165 63 79

[2] Exact Matching: 2 variables

##		Control	Treated
##	All	2583	2184
##	Matched	2583	2184
##	${\tt Unmatched}$	0	0
##	Discarded	0	0

[2] Exact Matching: 3 variables

```
var.comb <- do.call('paste0',</pre>
                     analytic.data[, c('race', 'sex', '
                                                          age'
length(table(var.comb))
## [1] 30
table(analytic.data$RHC, var.comb=="otherMale[80, Inf)")
##
##
            FALSE TRUE
##
     No RHC
            2580
```

```
##
     RHC
          2183
table(analytic.data$RHC, var.comb=="otherFemale[80, Inf)")
```

FALSE TRUE

No RHC 2581

[2] Exact Matching: 3 variables

##		Control	Treated
##	All	2583	2184
##	Matched	2581	2184
##	${\tt Unmatched}$	0	0
##	Discarded	0	0

[2] Exact Matching: 3 variables

```
matched.data <- match.data(m.out)
dim(matched.data)

## [1] 4765 25

nrow(analytic.data)-nrow(matched.data) # subjects deleted</pre>
```

```
## [1] 2
```

```
## Variable Units OddsRatio CI.95 p-value
## RHC No RHC Ref
## RHC 2.21 [1.96;2.49] <1e-04
```

[2] Exact Matching: many categorical variables

##		Control	Treated
##	All	2583	2184
##	Matched	2524	2150
##	${\tt Unmatched}$	0	0
##	Discarded	0	0

[2] Exact Matching: many categorical variables

[2] Exact Matching: including a continuous variable

##		Control	Treated
##	All	2583	2184
##	Matched	929	947
##	${\tt Unmatched}$	0	0
##	Discarded	0	0

[2] Exact Matching: including more continuous variables

##		Control	Treated
##	All	2583	2184
##	Matched	3	3
##	${\tt Unmatched}$	0	0
##	Discarded	0	0

[2] Exact Matching: including more continuous variables

```
matched.data <- match.data(m.out)</pre>
dim(matched.data)
## [1] 6 25
nrow(analytic.data) -nrow(matched.data) # subjects deleted
## [1] 4761
fit3m <- glm(I(Death=="Yes")~RHC,
```

family=binomial, data = matched.data)

```
## Variable Units OddsRatio CI.95 p-value
## RHC No RHC Ref
## RHC 1.00 [0.03;29.81] 1
```

publish(fit3m)

Defining Propensity score (PS)

- Conditional Probability of getting treatment, given the observed covariates
- \blacktriangleright Prob(treatment: A = 1 | baseline or pre-treatment covariates: L)
 - Prob(RHC = treated/RHC group | age, sex, race, etc.)
 - ► f(L) = Prob(A=1|L)

baselinevars

[16] "PaCO2"

[19] "Albumin"

##	ΓŦJ	age
##	[4]	"Disease.category"
##	[7]	"Pr.2mo.survival"
##	[10]	"Temperature"
##	[13]	"Respiratory.rate"

```
"sex" "race"
"DNR.status" "APACHE.II
```

"No.of.comorbidity" "DASI.2wk.]
"Heart.rate" "Blood.pre

"pH"
"GComa.Score"

"WBC.count"

"PaO2.by.F]

The central role of the propensity score in observational studies for causal effects <u>PR Rosenbaum. DR Rubin</u> - Biometrika, 1983. - academic oup.com.

The propensity score is the conditional probability of assignment to a particular treatment.

The propersity score is the conditional probability of assignment to a particular treatment given a vector of observed covariates. Both large and small sample theory show that adjustment for the solar propersity score is sufficient to remove bias due to all observed covariates. Applications included (in patched sampling on the univariate propersity score, which is a generalization of discriminant matching (i) multivariate absignment by which is a generalization of discriminant matching (ii) multivariate absignment by which is a generalization of discriminant matching (ii) multivariate absignment by which is a generalization of discriminant matching (iii) multivariate absignment by which is a present that the contract of the contract

Theoretical result

Rosenbaum, Rubin (1983) showed:

- For potential outcomes (Y^0, Y^1) , if you have sufficient observed covariate list L to reduce confounding ('strong ignoribility'): A being treatment assignment here:
 - ▶ i.e., if $(Y^0, Y^1) \perp \!\!\! \perp A|L$ (Note that is this NOT $Y \perp \!\!\! \perp A|L$)
- then
 - \blacktriangleright $(Y^0, Y^1) \perp \!\!\!\perp A|PS$ and
 - A ⊥⊥ L|PS

Assumptions

- no unmeasured confounding
- ▶ positivity (\$ 0 < PS < 1 \$)</p>
- well-defined treatment
- sufficient overlap
- model-specification

Variable selection for propensity score models

MA Brookhart. S Schneeweiss... - American journal of ..., 2006 - academic oup com Despite the growing popularity of propensity score (PS) methods in epidemiology, relatively little has been written in the epidemiologic literature about the problem of variable selection for PS models. The authors present the results of two simulation studies designed to help ... 25 99 Cited by 1550 Related articles All 18 versions Immort into BibTeX

The design *versus* the analysis of observational studies for causal effects: parallels with the design of randomized trials

DB Rubin - Statistics in medicine, 2007 - Wiley Online Library
For estimating causal effects of treatments, randomized experiments are generally
considered the gold standard. Nevertheless, they are often infeasible to conduct for a variety
of reasons, such as ethical concerns, excessive expense, or timeliness. Consequently, much
of our knowledge of causal effects must come from non-randomized observational studies.
This article will advocate the position that observational studies can and should be designed
to approximate randomized experiments as closely as possible. In particular, observational ...

**y 90 Cited by 980 Related articles All 8 versions import into Bibl's.

- Observed covariates are used to fix design
- ▶ Which covariates should be selected:
 - known to be a confounder (causes of Death and RHC)
 - ▶ known to be a cause of the outcome (risk factors of Death)
 - avoid known instruments or noise variables: SE suffers
 - mediating factors should be avoided (total effect = goal)
- Stepwise (p-value or criterion based) not recommended
 - depending on sample size, different values can get selected
 - may select variables highly associated with A
- Don't look at the outcome (Death) in your data to select covariates

Many ways to use propensity scores (PS) in the analysis

- ▶ **PS matching** [our focus today]
- ► PS weighting
- ▶ PS stratification
- ▶ PS used as a covariate

[3] Propensity Score Matching

```
A tutorial and case study in propensity score analysis: an application to estimating the effect of in-hospital smoking cessation counseling on mortality PC_Austin-Multivariate behavioral research, 2011 - 13yor & Francis
Propensity score methods allow investigators to estimate causal treatment effects using observational or connadomized data in this article very provide a practical illustration of the appropriate steps in conducting propensity score analyses. For illustrative purposes, we use ... $\frac{4}{2}$ 99 Cited by 204 Related articles All 13 versions import in biblifs*.
```

Propensity score matching has 4 steps

- ▶ Stage 1: exposure modelling: PS = Prob(A = 1|L)
- ► Stage 2: Match by *PS*
- ▶ Stage 2: Assess balance and overlap (PS and L)
- ▶ phase 4: outcome modelling: Prob(Y = 1|A = 1)

An introduction to **propensity score** methods for reducing the effects of confounding in observational studies

PC Austin - Multivariate behavioral research, 2011 - Taylor & Francis

The propensity score is the probability of treatment assignment conditional on observed baseline characteristics. The propensity score allows one to design and analyze an observational (nonrandomized) study so that it mimics some of the particular characteristics ...

\$\delta\$ 90 Cited by 5235 Related articles & All 15 versions Import into BibTeX

- Assessment of Balance in the whole data
 - balance = similarity of the covariate distributions
 - ightharpoonup d or SMD > 0.1 can be considered as imbalance

$$\begin{split} d &= \frac{(\overline{x}_{treatment} - \overline{x}_{control})}{\sqrt{s_{treatment}^2 + s_{control}^2}} \\ d &= \frac{(\hat{p}_{treatment} - \hat{p}_{control})}{\sqrt{\hat{p}_{treatment}(1 - \hat{p}_{treatment} - \hat{p}_{control})}} \end{split}$$

print(tab1e, smd = TRUE)

black

other

ARF

Disease.category (%)

##

##

##

##

##

** **				
##		No RHC		RHC
##	n	2583		2184
##	age (%)			
##	[-Inf,50)	573	(22.2)	540
##	[50,60)	432	(16.7)	371
##	[60,70)	638	(24.7)	577
##	[70,80)	603	(23.3)	529
##	[80, Inf)	337	(13.0)	167
##	sex = Female (%)	878	(34.0)	906
##	race (%)			
##	white	2072	(80.2)	1707

Stratified by RHC

400 (15.5)

111 (4.3)

1206 (46.7)

335

142

909

|3| Propensity Score Matching Step 1: PS estimation

Specify the propensity score model to estimate propensity scores, and fit the model

```
ps.formula <- as.formula(paste("I(RHC == 'RHC')", "~"</pre>
                 paste(baselinevars, collapse = "+")))
ps.formula
```

```
##
       APACHE.III.score + Pr.2mo.survival + No.of.comorbid
##
       DASI.2wk.prior + Temperature + Heart.rate + Blood.pr
```

I(RHC == "RHC") ~ age + sex + race + Disease.category +

Respiratory.rate + WBC.count + PaO2.by.FIO2 + PaCO2 ## Creatinine + Albumin + GComa.Score

- Coef of PS model fit is not of concern
- Model can be rich: to the extent that prediction is better
- But look for multi-collinearity issues
 - SE too high?

##

While PS has balancing property, PS is unknown and needs to be estimated:

- Other machine learning alternatives are possible to use instead of logistic regression.
 - tree based methods have better ability to detect non-linearity / non-additivity (model-specification aspect)
 - shrinkage methods lasso / elastic net may better deal with multi-collinearity
 - ensemble learners / super learners were successfully used
 - shallow/deep learning!

- Don't loose sight that better balance is the ultimate goal for propensity score
- Prediction of A is just a means to that end (as true PS is unknown).
- May attract variables highly associated with A

Improving propensity score estimators' robustness to model misspecification using super learner

R Efranchio. ML. Patersen... - American journal of ..., 2015 - scademic outp com The consistency of propensity score (Fg) estimators relies on correct specification of the PS model. The PS is frequently estimated using main-effects logistic regression. However, the underlying model assumptions may not hold. Machine learning methods provide an alternative nonparametric approach to PS estimation. In this simulation study, we evaluated the benefit of using Super-Learner (SL) for PS estimation. We created 1,000 simulated data sets (m = 500) under 4 different scenarios characterized by various degrees of deviance from ... \$\frac{1}{2}\$ \$\frac{

Should a propensity score model be super? The utility of ensemble procedures for causal adjustment

SAlam EDM Models: DA Stephenas - Statistics in medicine, 2019 - Wiley Online Library in investigations of the effect of treatment on outcome, the repressily score is a told reliable invaluation of the effect of treatment on outcome, the repressily score is a told eliminate invaluation in the distribution of confounding variables between treatment projectic regression in nominear settings, however, experience with real data analyses tends to show workfilling of the propentily score model using this approach. We investigated at wide range of simulated settings of varying complexities including simulations based on real data to ...
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ртмы Propensity score estimation: machine learning and classification methods as alternatives to logistic regression

<u>D.Westreich. J.Lessler</u>, N.J. Funk. - Journal of clinical epidemiology, 2010 - nebn.lm.nih.gov. Objective Propensity scores for the analysis of observational data are typically estimated using logistic regression. Our objective in this Review was to assess machine learning alternatives to logistic regression which may accomplish the same goals but with fewer assumptions or greater accuracy.

☆ 99 Cited by 47 Related articles All 4 versions Import into BibTeX

Can we train machine learning methods to outperform the high-dimensional propensity score algorithm?

ME Xaim. M Pang, RW-Platt - Epidemiology, 2018 - ingentaconnect.com The use of retrospective health care claims datasets is frequently criticized for the lack of complete information on potential conflounders. Ultizing patient is health status-related information from claims datasets as surrogates or proxies for mismeasured and unobserved conflounders, the high-dimensional propensity score algorithm enables us to reduce bias.

[3] Propensity score Matching Step 1

```
# summarize propensity scores
summary(analytic.data$PS)
```

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 0.009182 0.268112 0.454924 0.458150 0.640362 0.975476
```

```
# summarize propensity scores by exposure group
tapply(analytic.data$PS, analytic.data$RHC, summary)
```

```
## $`No RHC`
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 0.009182 0.184909 0.330687 0.357838 0.504012 0.974095
##
## $RHC
## Min. 1st Qu. Median Mean 3rd Qu. Max.
```

0.05156 0.42874 0.59400 0.57679 0.74044 0.97548

[3] Propensity Score Matching Step 2: PS matching

```
logitPS <- -log(1/analytic.data$PS - 1)
# logit of the propensity score
.2*sd(logitPS) # suggested in the literature</pre>
```

[1] 0.2382708

```
0.1*sd(logitPS) # we are using this
```

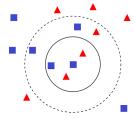
```
## [1] 0.1191354
```

choosing too strict PS has unintended consequences

Step 2

Match using estimates propensity scores

- nearest-neighbor (NN) matching
- without replacement
- ▶ with caliper = .1*SD of logit of propensity score
- ▶ with 1:1 ratio (pair-matching)



Step 2

Match using estimates propensity scores

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 0.009182 0.268112 0.454924 0.458150 0.640362 0.975476
```

Step 2

match.obj\$nn

```
## Control Treated

## All 2583 2184

## Matched 1519 1519

## Unmatched 1064 665

## Discarded 0 0
```

Step 2
Step 1 and 2 can be done together by specifying distance

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 0.009182 0.268112 0.454924 0.458150 0.640362 0.975476
```

Step 2: Taking a closer look at the matches

```
# Ref: https://lists.gking.harvard.edu/pipermail/matchit/2
matches <- as.data.frame(match.obj$match.matrix)
colnames(matches)<-c("matched_unit")
matches$matched_unit<-as.numeric(
    as.character(matches$matched_unit))
matches$treated_unit<-as.numeric(rownames(matches))
matches.only<-matches[!is.na(matches$matched_unit),]
head(matches.only)</pre>
```

##		matched_unit	treated_unit	
##	5	438	5	
##	10	2385	10	
##	12	4177	12	
##	13	4429	13	
##	17	5228	17	
##	22	1009	22	

[3] Propensity Score Matching Step 2: Taking a closer look at the matches (1st pair)

3.599609 3.500000

438 3.500000 2.799805

5

```
##
            age sex race Disease.category DNR.status APA
## 5
         [60,70) Male white
                                       MOSF
                                                   Yes
## 438 [80, Inf) Male white
                                        ARF
                                                   Yes
      Pr.2mo.survival No.of.comorbidity DASI.2wk.prior Ter
##
## 5
           0.43699980
                                              21.05078
## 438
           0.01399994
                                              15.95312
      Blood.pressure Respiratory.rate WBC.count Pa02.by.F.
##
## 5
                  65
                                   27 29.699219
                                                    478.00
## 438
                  46
                                       8.699219 138.09
      Creatinine Albumin GComa.Score RHC Death ID
##
```

41

RHC Yes 5 0.4

100 No RHC Yes 438 0.4

[3] Propensity Score Matching Step 2: Taking a closer look at the matches (2nd pair)

##

10

0.500000

2385 1.599854 3.5

2.5

```
##
             age sex race Disease.category DNR.status
## 10 [-Inf,50) Female white
                                            ARF
                                                       No
## 2385
          [70,80) Female white
                                            ARF
                                                       No
       Pr.2mo.survival No.of.comorbidity DASI.2wk.prior To
##
## 10
             0.6689453
                                               23.25781
## 2385
             0.6219997
                                                18.35156
##
        Blood.pressure Respiratory.rate WBC.count Pa02.by.1
## 10
                   73
                                    40 20.597656
                                                    68.0
## 2385
                   67
                                       3.199707
                                                     168.
```

Creatinine Albumin GComa.Score RHC Death ID

R.H.C

No

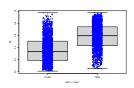
0 No RHC No 2385 0

10 0

Step 3: Assessing balance and overlap

Balance is more important than prediction!

- Criteria to assess success of step 2: PS estimation
 - better balance
 - better overlap
 - ightharpoonup PS = 0 or PS = 1 needs close inspection



Step 3

Vizualization

```
plot(match.obj, type = "jitter")
```

```
Distribution of Propensity Scores

Unmatched Teatment Units

Matched Teatment Units

Matched Teatment Units

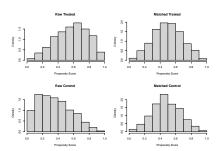
Unmatched Cores Units

Unmatched Cores Units
```

```
## [1] "To identify the units, use first mouse button; to s
## integer(0)
```

Step 3 Vizualization for assessing overlap issues

plot(match.obj, type = "hist")



Assessment of Balance

Compare the similarity of baseline characteristics between treated and untreated subjects in a the propensity score-matched sample.

- ▶ In this case, we will compare SMD < 0.1 or not.
- ▶ In some literature, other generous values (0.25) are proposed.

```
An introduction to propensity score methods for reducing the effects of confounding in observational studies 
EQ. Austin - Multivariate behavioral research, 2011 - Taylor & Francis 
The propensity score is the probability of retainment assignment conditional on observed baseline characteristics. The propensity score allows one to design and analyze an observational normationized study so that it minrics some of the particular characteristics ... 
$\frac{1}{2}$ TO Cited by $232$ Related articles All 15 versions import into Biblick 
Using propensity sorviers to help design observational studies: application to the tobacco litigation 
BR Bubin - Health Services and Outcomes Research Methodoloxy. 2001 - Sormorr
```

Propensity score methodology can be used to help design observational studies in a way

Step 3: Variance ratio

- ▶ Variance ratios ~ 1 means:
- equal variances in groups
- group balance
- could vary from 1/2 to 2
- other cut-points are suggested as well (0.8 to 1.2)

Balance diagnostics for comparing the distribution of baseline covariates between treatment groups in propensity-score matched samples

PC Austin - Statistics in medicine, 2009 - Wiley Online Library

The propensity score is a subject's probability of treatment, conditional on observed baseline covariates. Conditional on the true propensity score, treated and untreated subjects have similar distributions of observed baseline covariates. Propensity-score matching is a popular method of using the propensity score in the medical literature. Using this approach. matched sets of treated and untreated subjects with similar values of the propensity score are formed. Inferences about treatment effect made using propensity-score matching are ...

☆ 99 Cited by 2214 Related articles All 11 versions Import into BibTeX

гнтмы Matching methods for causal inference: A review and a look forward EA Stuart - Statistical science: a review journal of the Institute of 2010 - ncbi.nlm.nih.gov

When estimating causal effects using observational data, it is desirable to replicate a randomized experiment as closely as possible by obtaining treated and control groups with similar covariate distributions. This goal can often be achieved by choosing well-matched samples of the original treated and control groups, thereby reducing bias due to the covariates. Since the 1970's, work on matching methods has examined how to best choose treated and control subjects for comparison. Matching methods are gaining popularity in ...

☆ 99 Cited by 3066 Related articles All 22 versions Import into BibTeX ≫

Step 3: Variance ratio

Note: 's.d.denom' not specified; assuming pooled.

```
baltab.res$Balance$V.Ratio.Adj
```

```
## [1] 1.0990553
                     NA
                              NA
                                       NA
                                                NA
##
   [8]
             NΑ
                  NA
                           NA
                                       NA
                                                NA
             NA 1.0867497 0.9714495 0.9605864 0.8305596 (
## [15]
## [22] 0.9841995 1.0655834 1.0262382 0.9733399 1.0919443
  [29] 2.0325397 0.9847091
```

[60,70)

[70,80)

race (%) white

black

other

[80, Inf)

sex = Female (%)

##

##

##

##

##

prin	nt(tab1m, <u>showAllLeve</u>	els = FALSE, smd = TRUE,	test = FAL
##		Stratified by I	RHC
##		No RHC	RHC
##	n	1519	1519
##	age (%)		
##	[-Inf,50)	367 (24.2)	365
##	[50,60)	255 (16.8)	272

385 (25.3)

368 (24.2)

144 (9.5)

560 (36.9)

1212 (79.8)

229 (15.1)

78 (5.1)

395

351

136

572

1212

234

73

Possible to get p-values to check balance: but strongly discouraged

► P-value based balance assessment can be influenced by sample size

```
A tutorial and case study in propensity score analysis: an application to estimating the effect of in-hospital smoking cessation counseling on mortality PZ-Asstin-Nutherate behavioral several, 2011—19y9 of Francis

PZ-Asstin-Nutherated behavioral several, 2011—19y9 of Francis
observational or nonrandomized data in this article we provide a practical illustration of the appropriate lessely in conducting propensity score analyses. For illustrative purpose, we use ...
```

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[50,60)

44

##

print(tab1m, showAllLevels = FALSE, smd = FALSE, test =

##		Stratilied by F	(HC
##		No RHC	RHC
##	n	1519	1519
##	age (%)		
##	[-Inf,50)	367 (24.2)	365

CT---T:C:-3 L-- DIIC

(16.8)

272

255

[60,70) 385 (25.3) 395 ## [70,80] 368 (24.2) 351

Step 3

Assessment of balance in the matched data

```
smd.res <- ExtractSmd(tab1m)
t(round(smd.res,2))</pre>
```

```
##
           age sex race Disease.category DNR.status APACHI
## 1 vs 2 0.04 0.02 0.02
                                      0.04
                                                 0.01
##
          Pr.2mo.survival No.of.comorbidity DASI.2wk.prior
                     0.06
## 1 vs 2
                                        0.01
          Blood.pressure Respiratory.rate WBC.count PaO2.by
##
                    0.05
                                     0.04
                                                0.01
## 1 vs 2
##
          Creatinine Albumin GComa. Score
## 1 vs 2
                0.02
                        0.03
                                    0.01
```

Step 4: Outcome modelling

- Some flexibility in choosing outcome model
 - considered independent of exposure modelling
 - some propose double robust approach
 - adjusting imbalanced covariates only?

Estimate the effect of treatment on outcomes using propensity score-matched sample

[3] Propensity Score Matching Step 4: Outcome modelling

out.formula

##

```
## I(Death == "Yes") ~ age + sex + race + Disease.category
## APACHE.III.score + Pr.2mo.survival + No.of.comorbid:
## DASI.2wk.prior + Temperature + Heart.rate + Blood.pr
```

Respiratory.rate + WBC.count + PaO2.by.FIO2 + PaCO2
Creatinine + Albumin + GComa.Score

Variable

```
## age [-Inf,50) Ref
## [50,60) 1.62 [1.26;2.09] 0.0
## [60,70) 0.76 [0.61;0.94] 0.0
```

Units OddsRatio

CI.95

Step 4: Other cosiderations for outcome model

The above analysis do not take matched pair into consideration while regressing. Literature proposes different strategies:

- do not control for pairs / clusters
 - use glm as is
- control for pairs / clusters
 - use cluster option or GEE or conditional logistic
- Bootstrap for matched pairfor WOR
 - may not be appropriate for WR

The use of bootstrapping when using propensity-score matching without replacement: a simulation study

PCAustin, DS Small - Statistics in medicine, 2014 - Wiley Online Library Propestily-score matching is frequently used to estimate the effect of treatments, exposures, and interventions when using observational data. An important issue when using propensity-score matching is how to estimate the satinadard error of the estimated terror characteristic variance estimation permits construction of confidence intervals that have the advertised coverage rates and tests of statistical significance that have the correct type I error rates. There is disagreement in the literature as to how standard errors should be ... \$\frac{1}{2}\$ 99 Cited by 139 Related articles All II typescrips import this Districts.

Step 4

- ► The example compared RHC (a treated group; target) vs No RHC (untreated).
- ▶ The corresponding treatment effect estimate is known as
 - Average Treatment Effects on the Treated (ATT)
- Other estimates from PS analysis are possible that compared the whole population
 - what if everyone treated vs. what if nobody was treated (ATE)

Other matching algorithms

- Optimal
- genetic matching
- ► CEM
- variable ratio NN

- Matchlt
- Matching

Other useful packages

- cobalt
- twang

Outdated package

nonrandom

[4] Discipline-specific PS Systematic Reviews

- Propensity score matching most popular
 - Cardiovascular / Infective endocarditis / Intensive care
 - Critical care / anesthesiology / Sepsis / Psychology
 - Cancer / Multiple sclerosis
- Not meta-analysis; but reviews of usage of PS methods in different disciplines

imme. Propensity-scene matching in the cardiovascular surgery literature from 2004 to 2006; a significant crowled and suggestions for improval. If Edigith. — The Assmit of throate and activiousals supery, 2007. Eleverir Opensian Conscious a Systematic review of the use of grouped spice matting in the superposits additional review of the superposits and surgery and superposits additional matchines were used. Marchail consensed 50 orders published in the Asset of Phance Superposits point of Conscious Superposits and Confedential Superposits and Confedential Superposits and Confedential Superposits. 10 October 2004 Telescopies and of Confedential Superposits and Confedential Superposits.

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A systematic review of properatylly score methods in the social sciences. EICTROMERS, EICT, withhousite between 2011-Taylor & International the search properaty is cores in psychological and electation invested has been steady make a properaty in the search properaty in a series controm rescriptions about a consequent plant and 10 years. Invested these resists controlled the properaty is core adults in addition, recording proposed by properaty score adults on the search controlled type the properaty is considerable and proposed proposed

Reporting and guidelines in propensity score analysis: a systematic review of canner and canner surgical studies.

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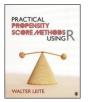
pmay Do the observational studies using propensity score analysis agree with randomized controlled trials in the area of sepsis? Z.Zang, H. N., X.V. - Journal of critical ear., 2014 - Bevier Background and objectives Sepsis in a leading cause of mortality and morbidity in the

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Reference

For newbie:



[воок] Practical propensity score methods using R

W Leite - 2016 - books.google.com

Practical Propensity Score Methods Using R by Walter Leite is a practical book that uses a step-by-step analysis of realistic examples to help students understand the theory and code for implementing propensity score analysis with the R statistical language. With a ...

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Companion site: study.sagepub.com/leite

Thank you!

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