

Econometrics I Empirical Question

Eric Schulman

February 13, 2018

a) Below are the results of the regression described in part (a).

Dep. Variable:	lwage	R-squared:	0.233
Model:	OLS	Adj. R-squared:	0.233
Method:	Least Squares	F-statistic:	428.8
Date:	Sat, 10 Feb 2018	Prob (F-statistic):	3.57e-243
Time:	16:31:36	Log-Likelihood:	-3651.2
No. Observations:	4230	AIC:	7310.
Df Residuals:	4226	BIC:	7336.
Df Model:	3		

	coef	std err	t	P> t	[0.025	0.975]
const	1.1852	0.045	26.488	0.000	1.097	1.273
educ	0.0904	0.003	33.051	0.000	0.085	0.096
exp	0.0354	0.003	14.083	0.000	0.030	0.040
exp_sq	-0.0005	5.03e-05	-9.251	0.000	-0.001	-0.000

Omnibus:	1562.688	Durbin-Watson:	1.770
Prob(Omnibus):	0.000	Jarque-Bera (JB):	34837.483
Skew:	-1.229	Prob(JB):	0.00
Kurtosis:	16.843	Cond. No.	4.31e+03

b) Summing the residuals in Python effectively produces 0 with some floating point error. My last program output was $-1.56521018368e-10$.

Again, they residuals multiplied with X is effectively 0 with some floating point error. $\langle -1.56560986e-10 \quad -1.82083681e-09 \quad -3.00428837e-09 \quad -7.82019924e-08 \rangle$

c) Below are the results of the regression described in part (c).

Dep. Variable:	lwage	R-squared:	0.035
Model:	OLS	Adj. R-squared:	0.035
Method:	Least Squares	F-statistic:	77.07
Date:	Sat, 10 Feb 2018	Prob (F-statistic):	1.33e-33
Time:	16:31:36	Log-Likelihood:	-4137.4
No. Observations:	4230	AIC:	8281.
Df Residuals:	4227	BIC:	8300.
Df Model:	2		

	coef	std err	t	P> t	[0.025	0.975]
const	2.3471	0.031	75.594	0.000	2.286	2.408
exp	0.0344	0.003	12.195	0.000	0.029	0.040
exp_sq	-0.0006	5.62e-05	-10.986	0.000	-0.001	-0.001
Omnibus:		965.638	Durbin-Watson:		1.692	
Prob(Omnibus):		0.000	Jarque-Bera (JB):		16245.042	
Skew:		-0.633	Prob(JB):		0.00	
Kurtosis:		12.517	Cond. No.		2.67e+03	

The average product of the residuals and education is 0.940. Obviously, this is non-zero as education was not included in the regression.

d) Below are the results of the regression described in part (d).

Dep. Variable:	educ	R-squared:	0.112
Model:	OLS	Adj. R-squared:	0.112
Method:	Least Squares	F-statistic:	267.1
Date:	Sat, 10 Feb 2018	Prob (F-statistic):	6.00e-110
Time:	16:31:36	Log-Likelihood:	-10954.
No. Observations:	4230	AIC:	2.191e+04
Df Residuals:	4227	BIC:	2.193e+04
Df Model:	2		

	coef	std err	t	P> t	[0.025	0.975]
const	12.8461	0.156	82.572	0.000	12.541	13.151
exp	-0.0113	0.014	-0.797	0.425	-0.039	0.016
exp_sq	-0.0017	0.000	-5.967	0.000	-0.002	-0.001
Omnibus:		65.618	Durbin-Watson:		1.674	
Prob(Omnibus):		0.000	Jarque-Bera (JB):		96.128	
Skew:		-0.174	Prob(JB):		1.34e-21	
Kurtosis:		3.652	Cond. No.		2.67e+03	

The average product of the new residuals and education is 10.396. One would expect the product to be greater than zero based on the proof below.

$$\begin{aligned}
 & y e' \\
 &= y y' - y \hat{y}' \\
 &= (X \hat{\beta} + \hat{e})(X \hat{\beta} + \hat{e})' - (X \hat{\beta} + \hat{e})(X \hat{\beta})'
 \end{aligned}$$

Since X and \hat{e} are orthogonal,

$$\begin{aligned}
 &= (X \hat{\beta})(X \hat{\beta})' + \hat{e} \hat{e}' - (X \hat{\beta})(X \hat{\beta})' \\
 &= \hat{e} \hat{e}' \geq 0
 \end{aligned}$$

Since the matrix has quadratic form.

- e) Below are the results form (e). As you can see the coefficients on the residuals are the same.

Dep. Variable:	y	R-squared:	0.205
Model:	OLS	Adj. R-squared:	0.205
Method:	Least Squares	F-statistic:	1093.
Date:	Sat, 10 Feb 2018	Prob (F-statistic):	2.01e-213
Time:	16:31:36	Log-Likelihood:	-3651.2
No. Observations:	4230	AIC:	7304.
Df Residuals:	4229	BIC:	7311.
Df Model:	1		

	coef	std err	t	P> t	[0.025	0.975]
x1	0.0904	0.003	33.063	0.000	0.085	0.096

Omnibus:	1562.688	Durbin-Watson:	1.770
Prob(Omnibus):	0.000	Jarque-Bera (JB):	34837.483
Skew:	-1.229	Prob(JB):	0.00
Kurtosis:	16.843	Cond. No.	1.00

- f) In (c) we calculated $M_{X_1}y$.

In (d) we calculated $M_{X_1}X_2$.

In (e) we calculated $((M_{X_1}X_2)'M_{X_1}X_2)^{-1}((M_{X_1}X_2)'M_{X_1}y)$.

Adding a column of ι as an intercept we have:

$$\begin{aligned}
 & \left(\begin{bmatrix} M_{X_1}X_2 & \iota \end{bmatrix}' \begin{bmatrix} M_{X_1}X_2 & \iota \end{bmatrix} \right)^{-1} \left(\begin{bmatrix} M_{X_1}X_2 & \iota \end{bmatrix}' M_{X_1}y \right) \\
 &= \begin{bmatrix} (M_{X_1}X_2)'M_{X_1}X_2 & (M_{X_1}X_2)'\iota \\ \iota'M_{X_1}X_2 & \iota'\iota \end{bmatrix}^{-1} \begin{bmatrix} (M_{X_1}X_2)'M_{X_1}y \\ \iota'M_{X_1}y \end{bmatrix} \\
 & \iota \text{ should be orthogonal to } M_{X_1}X_2 \text{ as it was included in } X_1 \text{ so} \\
 &= \begin{bmatrix} (M_{X_1}X_2)'M_{X_1}X_2 & 0 \\ 0 & \iota'\iota \end{bmatrix}^{-1} \begin{bmatrix} (M_{X_1}X_2)'M_{X_1}y \\ 0 \end{bmatrix} \\
 &= \begin{bmatrix} ((M_{X_1}X_2)'M_{X_1}X_2)^{-1}((M_{X_1}X_2)'M_{X_1}y) \\ 0 \end{bmatrix}
 \end{aligned}$$

Since, M_{X_1} is idempotent and symmetric, this is equivalent to

$$= \begin{bmatrix} (X_2'M_{X_1}X_2)^{-1}X_2'M_{X_1}y \\ 0 \end{bmatrix}$$

Finally, we proved in lecture that the top term is the coefficient on X_2

$$= \begin{bmatrix} \hat{\beta}_2 \\ 0 \end{bmatrix}$$

1 Python Code

```

import pandas
import math
import numpy as np
import matplotlib.pyplot as plt
import statsmodels.api as sm

FNAME = '../cps09mar.dta'

def write_result(model,no):
    result_doc = open('tex/hw2_reg%s.tex'%no,'w+')
    result_doc.write(model.summary().as_latex())
    result_doc.close()

def main():
    #filter matrix and load it into memory
    df = pandas.read_stata(FNAME)
    X = df[(df.female == 0) & (df.hisp == 1) & (df.race == 1)]
    print X['educ'].count()
    y = X['lwage']

    #part a
    X1 = X.loc[:,('educ','exp')]
    X1['exp_sq'] = X1['exp']**2
    X1 = sm.add_constant(X1)
    model1 = sm.OLS(y,X1).fit()
    write_result(model1,1)

    #part b
    e1 = model1.resid
    print sum(e1)
    print np.matmul(e1,X1)

    #part c
    X2 = X1.loc[:,('exp','exp_sq')]
    X2 = sm.add_constant(X2)
    model2 = sm.OLS(y,X2).fit()
    write_result(model2,2)

    #actual test part c
    e2 = model2.resid

    print np.matmul(e2,X1['educ'])/X1['educ'].count()

    #part d
    model3 = sm.OLS(X1['educ'],X2).fit()
    write_result(model3,3)

```

```
e3 = model3.resid
print np.matmul(e3,X1['educ'])/X1['educ'].count()

#part e
model4 = sm.OLS(e2,e3).fit()
write_result(model4,4)

#part f

if __name__ == "__main__":
    main()
```