

# Econometrics I - TA section

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TEXAS

The University of Texas at Austin

**3.6** For the double exponential distribution show

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$$f(x; \lambda) = \frac{1}{2\lambda} e^{-|x|/\lambda}$$

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$$\int_{-\infty}^{\infty} \frac{1}{2\lambda} e^{-|x|/\lambda} dx =$$

$$\int_{-\infty}^0 \frac{1}{2\lambda} e^{x/\lambda} dx + \int_0^{\infty} \frac{1}{2\lambda} e^{-x/\lambda} dx =$$

$$\left. \frac{1}{2} e^{x/\lambda} \right|_{x=-\infty}^{x=0} + \left. -\frac{1}{2} e^{-x/\lambda} \right|_{x=0}^{x=\infty} = 1$$

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Need to integrate by parts!

$$\frac{1}{2}(x - \lambda)e^{x/\lambda} \Big|_{x=-\infty}^{x=0} + -\frac{1}{2}(x + \lambda)e^{-x/\lambda} \Big|_{x=0}^{x=\infty} = 0$$

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- Similar strategy for next part  $E(x^2) = 2\lambda^2$



## 3.11 For the Pareto distribution show

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$$\frac{\alpha\beta^{\alpha}}{2-\alpha} x^{2-\alpha} \Big|_{\beta}^{\infty} = \frac{\alpha\beta^2}{\alpha-2}$$

$$Var(X) = E(X^2) - E(X)^2 = \frac{\alpha\beta^2}{(\alpha-2)(\alpha-1)^2}$$



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- Independence

$$f(x)f(y) = f(x, y)$$

When integrating multivariate functions. Always write down the set your trying to integrate over ... Double check this equation

- Fubini's theorem

$$\begin{aligned}\int_{(x,y) \in X \times Y} f(x,y) d(x,y) &= \int_{y \in Y} \int_{x \in X} f(x,y) dx dy \\ &= \int_{x \in X} \int_{y \in Y} f(x,y) dy dx\end{aligned}$$

- Will need to use this when  $x$  and  $y$  are not independent

- Law of iterated expectations

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# Chapter 4

**4.1** Let  $f(x, y) = 1/4$  for  $-1 \leq x \leq 1$  and  $-1 \leq y \leq 1$

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Also,  $f(x, y) \geq 0$

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$$\int_{-1}^1 \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{1}{4} dx dy = \int_{-1}^1 \frac{1}{2} \sqrt{1-y^2} dy$$

Using a change of variable let  $y = \sin(\theta)$ ,  $dy = \cos(\theta)d\theta$  i.e.  
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What do we integrate over? Fixing  $y$ ,  $x \in [-2 - y, 2 - y]$ .

And,  $y \in [-1, 1]$

$$\int_{-1}^1 \int_{-1}^1 \frac{1}{4} dx dy = \int_{-1}^1 \frac{1}{2} dy = 1$$

$$4.6 \quad f(x, y) = \begin{cases} cxy & x, y \in [0, 1], x + y \leq 1 \\ 0 & \text{o.w.} \end{cases}$$

1. Find the value of  $c$  such that  $f(x, y)$  is a joint PDF

$$\int_0^1 \int_0^{1-y} f(x, y) dx dy = 1$$

$$\int_0^1 \frac{c(1-y)^2 y}{2} dx dy = 1$$

$$c/24 = 1$$

$$c = 24$$

**4.6**  $f(x, y) = \begin{cases} cxy & x, y \in [0, 1], x + y \leq 1 \\ 0 & \text{o.w.} \end{cases}$

- Find the marginal distribution of  $X$  and  $Y$

$$f(x) = \int_0^{1-x} f(x, y) dy dx = 12(1-x)^2 x$$

$$f(y) = \int_0^{1-y} f(x, y) dy dx = 12(1-y)^2 y$$

- Are  $X$  and  $Y$  independent? No.  $f(x)f(y) \neq f(x, y)$

**4.7** Let  $X$  and  $Y$  have density  $f(x, y) = \exp(-x - y)$  for  $x > 0$  and  $y > 0$ . Find the marginal density of  $X$  and  $Y$ . Are  $X$  and  $Y$  independent or dependent?



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$$f_x(x) = \int_0^{\infty} \exp(-x - y) dy$$

$$f_x(x) = \int_0^{\infty} \exp(-x - y) dy = -\exp(-x - y) \Big|_{y=0}^{y=\infty}$$

$$f_x(x) = \exp(-x)$$

Similarly,  $f_y(y) = \exp(-y)$

$$f(x, y) = f_x(x)f_y(y)$$

so yes. They are independent

## Chapter 4

**4.8** Let  $X$  and  $Y$  have density  $f(x, y) = 1$  on  $0 < x < 1$  and  $0 < y < 1$ . Find the density function of  $Z = XY$ .

A lot of you noticed, that there are “two areas” to integrate over. Comes from LIE.

$$Pr(XY \leq z) = Pr(XY \leq z \cap Y \leq z) + Pr(XY \leq z \cap Y > z)$$

$$Pr(XY \leq z \cap Y \leq z) = \int_0^z \int_0^1 1 dx dy = z$$

$$Pr(XY \leq z \cap Y > z) = \int_z^1 \int_0^{z/y} 1 dx dy = -z \ln(z)$$

Thus, we get  $F(z) = z - z \ln(z)$  and  $f(z) = -\ln(z)$  where  $z \in (0, 1]$

**4.9** Let  $X$  and  $Y$  have density  $f(x, y) = 12xy(1 - y)$  for  $0 < x < 1$  and  $0 < y < 1$ . Are  $X$  and  $Y$  independent or dependent?

$$f(x) = \int_0^1 12xy(1 - y)dy = 2x$$

$$f(y) = \int_0^1 12xy(1 - y)dx = 6(1 - y)y$$

So, yes they are independent

$$f(y)f(x) = 6(1 - y)y2x = 12(1 - y)y2x = f(x, y)$$

**4.14** Let  $X_1 \sim \text{gamma}(r, 1)$  and  $X_2 \sim \text{gamma}(s, 1)$  be independent. Find the distribution of  $Y = X_1 + X_2$ .

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First, derive the moment generating function of the gamma distribution

$$\begin{aligned} M(t; s, 1) &= Ee^{tX} = \int_0^{\infty} e^{tx} f(x; s, 1) dx \\ &= \int_0^{\infty} e^{tx} \frac{1}{\Gamma(s)} x^{s-1} e^{-x} dx \\ &= \frac{1}{\Gamma(s)} \int_0^{\infty} x^{s-1} e^{-(1-t)x} dx = \\ &\quad \frac{1}{\Gamma(s)} \frac{\Gamma(s)}{(1-t)^s} = \frac{1}{(1-t)^s} \end{aligned}$$

**4.14** By using the property of independent random variables, we know

$$M_{X+Y}(t) = M_X(t)M_Y(t)$$

So if  $X_1 \sim \text{gamma}(s, 1)$ ,  $X_2 \sim \text{gamma}(s, 1)$ ,

$$M_{X_1+X_1}(t) = \frac{1}{(1-t)^s} \frac{1}{(1-t)^r} = \frac{1}{(1-t)^{r+s}}$$

You can see the MGF of the product is still in the format of Gamma distribution. Finally we can get

$$X_1 + X_2 \sim \text{gamma}(r + s, 1)$$

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Suppose that the distribution of  $Y$  conditional on  $X = x$  is  $N(x, x^2)$  and the marginal distribution of  $X$  is  $U[0, 1]$

1. Find  $E(Y)$

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$$E(Y) = E[E(Y|X)] = E(x) = \int_0^1 x dx = 1/2$$

2. Find  $Var(Y)$



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1. Find  $E(Y)$

$$E(Y) = E[E(Y|X)] = E(x) = \int_0^1 x dx = 1/2$$

2. Find  $Var(Y)$

$$E(Y^2) = E[E(Y^2|X)] = E[E(Y^2|X) - E(Y|X)^2 + E(Y|X)^2] =$$

$$E[Var(Y|X) + E(Y|X)^2] = E(x^2 + x^2) = \int_0^1 2x^2 dx = 2/3$$

$$E(Y^2) - E(Y)^2 = 2/3 - 1/4 = 5/12$$