Data Science for Actuaries (ACT6100)

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Rappels # 3.4 (Gaussian Vectors)

automne 2020

Random Vectors

Soient X un vecteur aléatoire de dimension d

- L'espérance de X, notée $\mathbb{E}(X)$ est définie (si elle existe) par le vecteur de dimension $d \mathbb{E}(\boldsymbol{X}) = (\mathbb{E}(\boldsymbol{X}_1), \dots, \mathbb{E}(\boldsymbol{X}_d))^{\top}$.
- La matrice de covariance (appelée aussi matrice de variance-covariance de X) est définie (si elle existe) par la matrice de taille (d, d)

$$\mathsf{Var}(oldsymbol{X}) = \mathbb{E}\left((oldsymbol{X} - \mathbb{E}(oldsymbol{X}))(oldsymbol{X} - \mathbb{E}(oldsymbol{X}))^{ op}
ight).$$

Ainsi le terme ij de cette matrice représente la covariance entre X_i et X_i ,

$$Cov(X_i, X_j) = \mathbb{E}\left[(X_i - \mathbb{E}(X_i))(X_j - \mathbb{E}(X_j))\right].$$



Random Vectors

Soit **X** un vecteur aléatoire de dimension d, de moyenne μ et de matrice de covariance Σ .

Soient **A** et **B** deux matrices réeeles de taille (d, p) et (d, q) et enfin soit $\mathbf{a} \in \mathbb{R}^p$ alors

$$lacksquare$$
 $\operatorname{Var}(oldsymbol{X}) = \mathbb{E}\left((oldsymbol{X} - oldsymbol{\mu})(oldsymbol{X} - oldsymbol{\mu})^{ op}
ight) = \mathbb{E}(oldsymbol{X}oldsymbol{X}^{ op}) - oldsymbol{\mu}oldsymbol{\mu}^{ op}.$

$$ightharpoonup \mathbb{E}\left(\mathbf{A}^{\top}\mathbf{X}+\mathbf{a}\right)=\mathbf{A}^{\top}\boldsymbol{\mu}+\mathbf{a}.$$

$$\qquad \qquad \mathsf{Var}\left(\boldsymbol{A}^{\top}\boldsymbol{X} + \mathbf{a}\right) = \boldsymbol{A}^{\top}\boldsymbol{\Sigma}\boldsymbol{A}.$$

$$\triangleright \mathsf{Cov}\left(\mathbf{A}^{\top}\mathbf{X}, \mathbf{B}^{\top}\mathbf{X}\right) = \mathbf{A}^{\top}\mathbf{\Sigma}\mathbf{B}.$$



The Gaussian Distribution

A Gaussian variable, with distribution $\mathcal{N}(\mu, \sigma^2)$, where $\mu \in \mathbb{R}$ and $\sigma > 0$, has density

$$f(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2}\frac{(x-\mu)^2}{\sigma^2}\right), \text{ for all } x \in \mathbb{R}.$$

Then $\mathbb{E}(X) = \mu$ and $Var(X) = \sigma^2$. Observe that if $Z \sim \mathcal{N}(0,1)$, $X = \mu + \sigma Z \sim \mathcal{N}(\mu, \sigma^2)$. The Gaussian vector $\mathcal{N}(\mu, \Sigma)$: $\mathbf{X} = (X_1, ..., X_n)$ is a Gaussian vector with mean $\mathbb{E}(\boldsymbol{X}) = \boldsymbol{\mu}$ and covariance matrix $\mathsf{Var}(oldsymbol{X}) = oldsymbol{\Sigma} = \mathbb{E}\left((oldsymbol{X} - oldsymbol{\mu}) (oldsymbol{X} - oldsymbol{\mu})^ op
ight)$ non-degenerated $(oldsymbol{\Sigma}$ is invertible) if its density is

$$f(\mathbf{x}) = \frac{1}{(2\pi)^{n/2} \sqrt{\det \mathbf{\Sigma}}} \exp\left(-\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^{\top} \mathbf{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right), \ \mathbf{x} \in \mathbb{R}^{n},$$

see multivariate Gaussian distribution

Gaussian (multivariate) distribution

 $X \sim \mathcal{N}(\mu, \Sigma)$, with density

$$f_{\mathbf{X}}(x_1,\ldots,x_k) = \frac{1}{\sqrt{(2\pi)^k |\mathbf{\Sigma}|}} \exp\left(-\frac{1}{2}(\mathbf{x}-\boldsymbol{\mu})^{\top}\mathbf{\Sigma}^{-1}(\mathbf{x}-\boldsymbol{\mu})\right)$$

where $\mathbb{E}(\boldsymbol{X}) = \boldsymbol{\mu}$ and $\text{Var}(\boldsymbol{X}) = \boldsymbol{\Sigma}$.

Estimates are
$$\overline{\mathbf{x}} = \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i}$$
 and $\widehat{\mathbf{\Sigma}} = \frac{1}{n} \sum_{i=1}^{n} (\mathbf{x}_{i} - \overline{\mathbf{x}})(\mathbf{x}_{i} - \overline{\mathbf{x}})^{\top}$

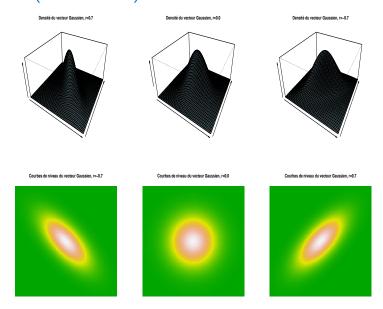
In dimension 2, f(x, y) is proportional to

$$\exp\left(-\frac{1}{2(1-\rho^2)}\left[\frac{(x-\mu_X)^2}{\sigma_X^2}+\frac{(y-\mu_Y)^2}{\sigma_Y^2}-\frac{2\rho(x-\mu_X)(y-\mu_Y)}{\sigma_X\sigma_Y}\right]\right)$$

levels curves (isodensities) are ellipses.



Gaussian (multivariate) distribution



Quadratic Forms

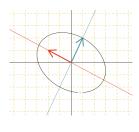
Consider
$$\mathbf{M} = \begin{pmatrix} a & b \\ b & c \end{pmatrix}$$
, and function $\mathbf{z} \mapsto \mathbf{z}^{\top} \mathbf{M} \mathbf{z}$, i.e.

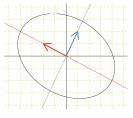
$$f: \begin{pmatrix} x \\ y \end{pmatrix} \mapsto \begin{pmatrix} x & y \end{pmatrix} \begin{pmatrix} a & b \\ b & c \end{pmatrix} \begin{pmatrix} x \\ y \end{pmatrix}$$

or $ax^2 + 2bxy + cy^2$ is a quadratic form.

If M > 0, points z = (x, y) such that $z^{\top}Mz =$ γ , for some $\gamma > 0$, are on an ellipse (centered on **0**)

Let $\lambda_1 > \lambda_2 > 0$ denote the eigenvalues of **M** and $\vec{\mathbf{v}}_1$ and $\vec{\mathbf{v}}_2$ denote the eigenvectors.





Quadratic Forms

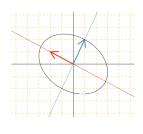
On the picture,
$$\mathbf{M} = \begin{pmatrix} 0.6 & 0.2 \\ 0.2 & 0.9 \end{pmatrix}$$

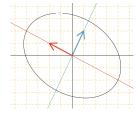
- 1 > M=matrix(c(.6,.2,.2,.9),2,2)
 2 > eigen(M)
 3 eigen() decomposition
 4 \$values
 5 [1] 1.0 0.5
 6 \$vectors
- 7 [,1] [,2]
- 8 [1,] 0.4472136 -0.8944272
- 9 [2,] 0.8944272 0.4472136

i.e.
$$\lambda_1=1$$
 and $\lambda_2=1/2$, and

$$\vec{\mathbf{v}}_1 = \sqrt{5} \begin{pmatrix} 1 \\ 2 \end{pmatrix}, \ \vec{\mathbf{v}}_2 = \sqrt{5} \begin{pmatrix} -2 \\ 1 \end{pmatrix}$$

Note that $\| ec{m{v}}_1 \| = \| ec{m{v}}_2 \| = 1$ and $ec{m{v}}_1 \perp ec{m{v}}_2$





The Gaussian Distribution

If X is a Gaussian vector, then for any i, X_i has a (univariate) Gaussian distribution, but its converse it not necessarily true.

Let $\boldsymbol{X}=(X_1,...,X_n)$ be a random vector with mean $\mathbb{E}(\boldsymbol{X})=\boldsymbol{\mu}$ and with covariance matrix Σ , if \boldsymbol{A} is a $k \times n$ matrix, and $\boldsymbol{b} \in \mathbb{R}^k$. then $\mathbf{Y} = \mathbf{A}\mathbf{X} + \mathbf{b}$ is a Gaussian vector \mathbb{R}^k , with distribution $\mathcal{N}\left(\mathbf{A}\boldsymbol{\mu}+\mathbf{b},\mathbf{A}\boldsymbol{\Sigma}\mathbf{A}^{\top}\right).$

Observe that if (X_1, X_2) is a Gaussian vector, X_1 and X_2 are independent if and only if

$$\mathsf{Cov}\left(X_{1},X_{2}\right)=\mathbb{E}\left(\left(X_{1}-\mathbb{E}\left(X_{1}
ight)
ight)\left(X_{2}-\mathbb{E}\left(X_{2}
ight)
ight)
ight)=0.$$

