

Use & Value of Unusual Data

Arthur Charpentier

UQAM

Actuarial Summer School 2019

The Team for the Week

Jean-Philippe Boucher UQAM (Montréal, Canada)
Professor, holder of The Co-operators Chair in Actuarial Risk Analysis, author of several research articles in actuarial science.

✉@J_P_Boucher Ⓜ jpboucher.uqam.ca



Arthur Charpentier UQAM (Montréal, Canada)
Professor, editor of *Computational Actuarial Science with R*. Former director of the Data Science for Actuaries program of the French Institute of Actuaries

✉@freakonometrics Ⓜfreakonometrics Ⓜ freakonometrics.hypotheses.org



Ewen Galic AMSE (Aix-Marseille, France)
Assistant professor, teaching computational econometrics, data science and machine learning.

✉@3wen Ⓜ3wen Ⓜ egallic.fr



Practical Issues

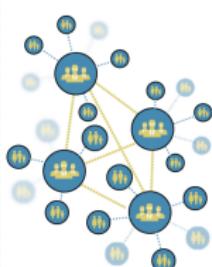
General Context (in Insurance)

The "insurance company" node at the top holds all the power with one central balance sheet.



Individual risk is transferred to a central entity that has all the power. This one balance sheet gives insurance company immense leverage.

The "insurance company" node still holds the majority of power, but also share risk with each other.



Distributed risk from individuals is transferred to a central entity, but risk is also spread amongst central entities to further reduce risk. Similarly to regular insurance, all the leverage for who receives care and how much is centralized.



About Friendsurance

In 2010, the founders of the FinTech company, Friendsurance, realized that insurance is expensive and lacks transparency. People pay high premiums each year and get nothing in return.

The vision of Friendsurance is to make insurance easier and more affordable for customers, helping to save customers time and money that can better be spent on the finer side of life. Fifty

Networks and P2P Insurance (via <https://sharable.life> and source)

General Context (in Insurance)

FEATURES

PEER-TO-PEER INSURANCE

GOING BACK TO BASICS

Innovators have set their sights on simplifying insurance by adopting new peer-to-peer business models. Who are the key players and does the Australian industry need to pay attention?

DISCUSSION OF the sharing economy has taken up considerable column space in recent times.

PricewaterhouseCoopers estimates the five main sharing sectors (peer-to-peer finance, online staffing, peer-to-peer accommodation, car sharing and music video streaming) will continue general global revenue of US\$830bn by 2016. Today, PwC says revenue generated is around US\$10bn (A\$15.67bn).

One sharing economy sector still very much in its infancy is peer-to-peer insurance. But Amy Gibbs, digital communications and content strategy manager at ANZIIF, says popularity of the concept is increasing.

"When it takes off, it will likely happen much quicker than we expect," she says.

"P2P insurance is not about a new technology that replaces the old. It's about people demanding an industry that gives them what they believe they deserve. Consumers expect different things in 2016, and technology now allows them the power to get what they want or go elsewhere."

The current crop in P2P
On 24 March, Germany's high-profile P2P player, Friendinsurance, announced it had collected US\$81.3m (A\$82.60m) from investors in its latest round of financing.

Tim Kundi, Friendinsurance's co-founder and managing director, says the organisation intends to use that fresh capital to grow further in

Germany and expand internationally. He says the first expansion target for 2016 is Australia, and expansion opportunities for other markets are currently being considered.

Friendinsurance is one of the players Gibbs has been keeping a close eye on.

"Friendinsurance and Guevara are ones that I watch closely," she says.

"They have interesting models and appear to have put a lot of thought into them, plus they get the marketing/consumer angle, which is crucial."

Friendinsurance, founded in Germany in 2010, operates as an independent insurance broker. It describes its mission as to make insurance easier and more affordable for customers, and to reduce the number of fraudulent claims.

"Our idea is inspired by insurance in its original form, when people got together in small groups... and supported each other in case of damage with their own resources," says Tim Kundi, friendinsurance's director of insurance business.

"This was easy and efficient but also limited in the extent of coverage. Today, big insurance

companies can carry claims of any amount, but marketing, administration and fraud cause remarkable costs.

Against this background, we developed an insurance concept that again [creates] smaller groups within bigger insurance societies and rewards remaining claimless within [these groups] with annual cashbacks."

Friendinsurance customers with the same insurance type form small groups online. Part of their insurance premium is paid into a cashback pool, and part of it is provided to the group insurer (or reinsurer). When small claims are made, customers are reimbursed from the general fund. Larger claims still go through the insurer. Groups that have no claims during a year receive a cashback bonus the following January. Friendinsurance says its claims-free bonus is available for private liability, home contents and legal expense insurance.

"So far, more than 80% of users received some of their insurance fees back," says Kundi.

"In the property insurance line, the average cashback was 33% of the paid insurance fees."

Kundi reports that, in 2015, Friendinsurance engaged 75,000 new customers.

"Today, we have a six-digit number of customers, 70 insurance partners, 15 corporates and 60 employees."

Over in the UK, start-up Guevara has received



"P2P insurance is not about a new technology threatening an industry but about people demanding an industry that gives them what they believe they deserve"

Amy Gibbs, ANZIIF



Networks and P2P Insurance, via Insurance Business (2016)

General Context (in Insurance)

Telematics Usage-Based Insurance (UBI)

During the first quarter of 2014, LexisNexis commissioned an independent firm to conduct two studies. The first was a study of drivers who were using telematics services to reduce their risk of being involved in a motor vehicle accident. The second was a study of 4000 small fleet managers who were using telematics devices to coordinate fleets of 10 to 20 vehicles.

Consumer Insights

Overall consumer awareness has plateaued however it's growing among younger drivers



Consumers are now as comfortable sharing UBI driving data as they are sharing many other types of information



And their interest in UBI at lower discount levels is rising



Offering popular value-added services in addition to the discount increases demand



**Decline Dynamics
IN AUTO
INSURANCE**

Technology will bring huge changes to automobiles and personal auto insurance over the next decade. Assuming there will be no changes is a risky choice.

By Dr. Michael Rydell

In the US, the auto-insurance industry is significant, generating \$20 billion in annual revenue and insuring 270 million vehicles. It makes up more than 10 percent of the total auto-industry. It is predicted that by 2020, the number of connected cars will have a substantially adverse impact on the auto-insurance industry, leading to the eventual decline of the current way of doing business.

While this might be true, the more important question is the potential for insurance companies to change the decline—in decline dynamics. Many in the auto-insurance industry have predicted a slow decline of profitability over the next decade. However, the introduction of telematics has created a whole new dynamic, one that could lead to a much faster decline in a shorter period of time.

It is important to understand how these dynamics work. In this article, we will introduce you to the concept of decline dynamics and its knowledge of the concepts of industrial decline

case by case. Cognitive biases! We humans are creatures of habit and tend to stick to the same patterns and behaviors. This is why we often make the same mistakes over and over again. It is also why we tend to ignore new information that contradicts our beliefs.

Technology and Turnover Rates

The lack of concern for money in the insurance industry is the primary driver for the relatively slow introduction of technology in the industry. Creating a \$10 billion automotive driving system—which can compete with losses like the one above—

30 YEARS // June 2017 // www.actuarymagazine.org

TheActuary
The magazine of the Institute and Faculty of Actuaries

INTERVIEW
Michael Green
PwC: Is the IoT revolution can benefit society

INNOVATION
Leadership and pricing long-term data

TECHNOLOGY
Tough industrial challenges in the insurance landscape

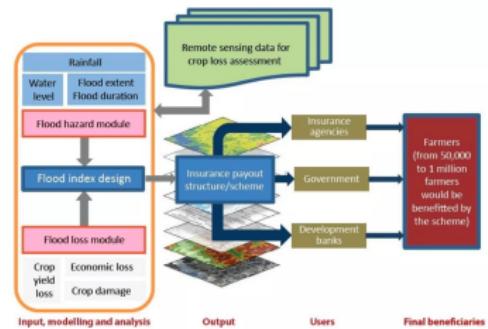
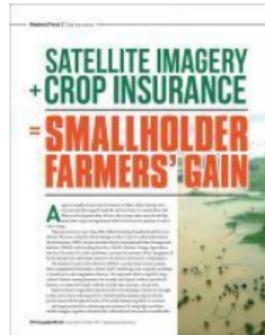
ENVIRONMENT
Environmental extremes in sustainable investment

Who's in the driving seat?

How telematics could transform motor insurance

Telematics and Usage-Based-Insurance ([source](#), [source](#), [The Actuary](#))

General Context (in Insurance)



Satellite pictures ([source](#) and Index-based Flood Insurance (IBFI),
<http://ibfi.iwmi.org/>)

General Context (in Insurance)



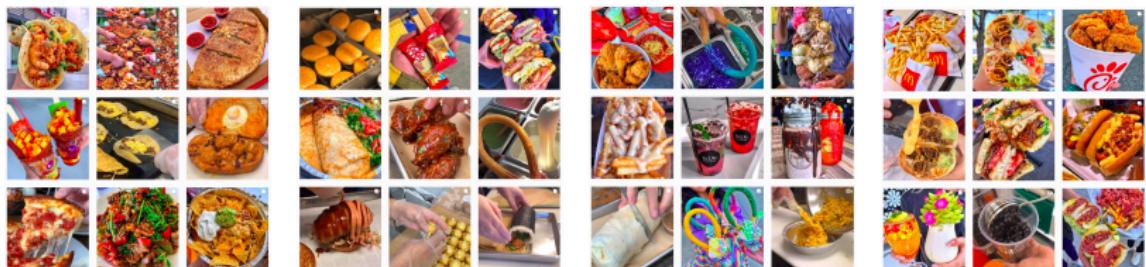
Very small granularity (1 pixel is $15 \times 15\text{cm}^2$, see [detail](#))

More and More (exponentially) Data

1969, (legend says) 5 photographs were taken by Neil Armstrong



More than 800,000,000 pictures are uploaded on every day, on Facebook (2 billion photos per day with [instagram](#), Messenger and Whatsapp)



Data Complexity

Pictures, drawings, videos, handwriting text, ECG, medical imagery, satellite pictures, tweets, phone calls, etc.

Hard to distinguish structured and un-structured data...

Data Complexity



Insurance Data

From questionnaires
to data frames $x_{i,j}$
row i : insured, policy, claim
column j : variables (features)

HEALTH INSURANCE QUESTIONNAIRE					
<small>Please provide all the information requested and return this form to your eligibility worker. Use and attach a copy of your insurance policy, membership card, or any other aid to help complete this questionnaire. PLEASE TYPE OR PRINT. DO NOT ABBREVIATE. Additional instructions and information collection and access are on the reverse. If you have any questions about completing this form or require Spanish translation, call toll-free 1-800 952-5294 (7:30 a.m. to 5:00 p.m.).</small>					
<small>COMPLETE THIS FORM FOR ANY MEDICAL INSURANCE, INCLUDING MEDICARE SUPPLEMENT, PREPAID HEALTH PLANS/HEALTH MAINTENANCE ORGANIZATIONS, AND OTHER PRIVATE HEALTH INSURANCE. THIS FORM DOES NOT AFFECT YOUR MEDICAL ELIGIBILITY; HOWEVER, FAILURE TO REPORT OTHER HEALTH INSURANCE MAY BE A CAUSE FOR TERMINATION OF YOUR MEDICAL ELIGIBILITY.</small>					
Case name		FOR COUNTY USE ONLY		STATE USE ONLY	
Worker name		Verifier by			
Case address		Date	Date		Initials
		Worker telephone number ()	Date		Initials
<input type="checkbox"/> Initial intake <input type="checkbox"/> Retermination <input type="checkbox"/> NIPR		Optional District number		Scope	CC number
SECTION I: Beneficiary Information. LIST ALL PERSONS, INCLUDING UNBORN, ON MEDICAL AND COVERED BY HEALTH INSURANCE POLICY					
14-DIGIT MEDICAL NUMBER					
OHC	Beneficiary Name (First, Middle, Last)	Social Security Number	Sex	Date of Birth	Code Code
	— — —	— — —	—	— — —	— — —
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	— — —	— — —	—	— — —	— — —
	— — —	— — —	—	— — —	— — —
SECTION II: Health Insurance Information					
1. What is the name and address of your health insurance company? Include street number, city, state, and ZIP. Do not use abbreviations. Name: _____ Address: _____ City, State, ZIP: _____					
2. Do you have to obtain medical services from a specific facility or a group of providers? <input type="checkbox"/> Yes <input type="checkbox"/> No					
3. Where do you send your claims? Name: _____ Address: _____ City, State, ZIP: _____					
4. What is the full name, address, phone number, and SSA number of individual, employee, union/member, or person to whom the insurance policy was issued? Name: _____ Address: _____ City, State, ZIP: _____					
Social Security Number: _____ Telephone number: _____ Absent parent? <input type="checkbox"/> Yes <input type="checkbox"/> No					
5. What is the policy number?					
6. When were the dates of your policy? Beginning date: _____ Ending date (if applicable): _____ <input type="checkbox"/> Medical coverage available through employer, but has not been applied for					
7. Premium amount \$ _____ <input type="checkbox"/> Monthly <input type="checkbox"/> Quarterly <input type="checkbox"/> Yearly How are premiums paid? <input type="checkbox"/> By insured <input type="checkbox"/> By employer <input type="checkbox"/> By payroll deduction					
8. Give name, address, and telephone number of urban, employee, group, organization, or school. Name: _____ Address: _____ City, State, ZIP: _____ Telephone number: _____					
9. Does anyone covered beneficiary have an adult, chronic, or pre-existing illness that requires him/her to use a physician? <input type="checkbox"/> Yes <input type="checkbox"/> No If yes, please specify the illness:					
<input type="checkbox"/> Hospital outpatient (i.e., lab work/physical therapy) <input type="checkbox"/> Prescription drugs <input type="checkbox"/> Long term care/nursing home <input type="checkbox"/> Hospital stays <input type="checkbox"/> Dental care <input type="checkbox"/> Only specific illnesses (i.e., cancer) <input type="checkbox"/> Doctor visits <input type="checkbox"/> Vision care <input type="checkbox"/> Type of illness					
10. Is the policy a Medicare Supplement? <input type="checkbox"/> Yes <input type="checkbox"/> No					
Premium: _____					

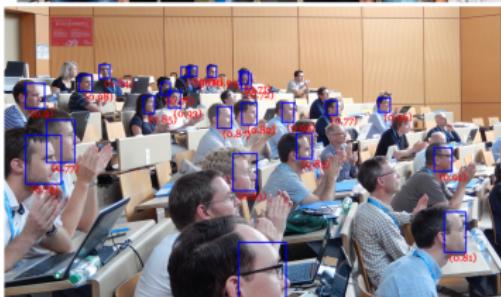
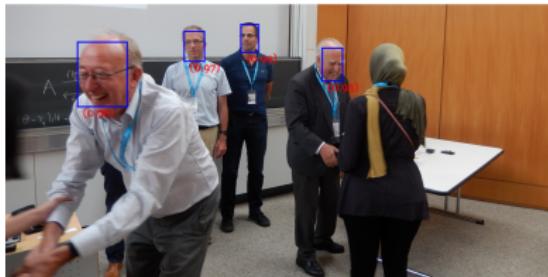
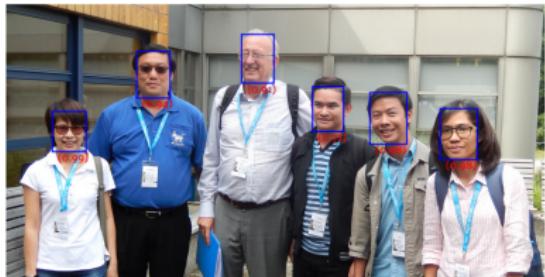
From Pictures to Data

Consider pictures saa-iss.ch/testalbum/nggallery/all-old-galleries/



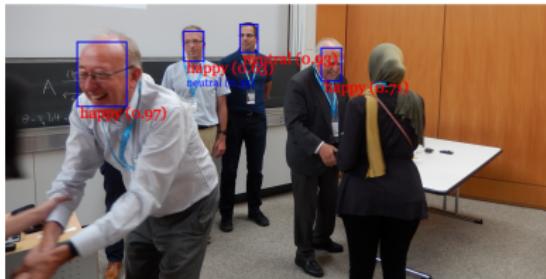
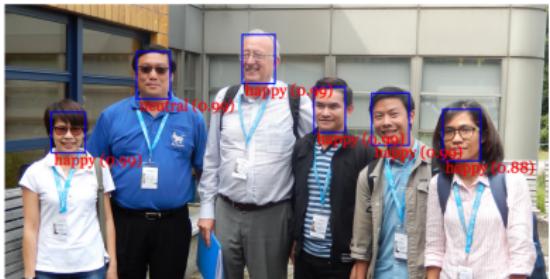
From Pictures to Data

Face detection face-api.js (justadudewhohacks.github.io)



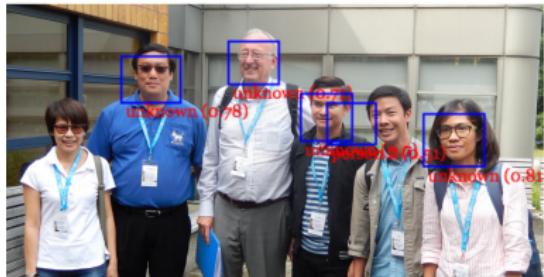
From Pictures to Data

Expression recognition face-api.js (justadudewhohacks.github.io)



From Pictures to Data

Face recognition face-api.js (justadudewhohacks.github.io)



Pictures

Extracting information from a picture

What is on the picture ?

What is the cost of such a claim ?

High dimensional data :

$h \times d$ color picture, e.g. 850×1280

$\mathbf{x}_i \in \mathbb{R}^d$ with $d > 3$ million

Need new statistical tools

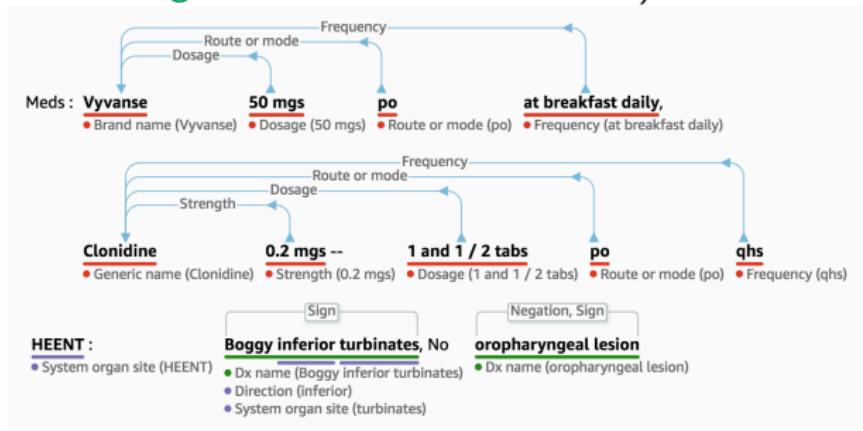
(deep) neural networks perform well
or reduce dimension

Pictures Odermatts (2003, [Karambolage](#))



Text

Kaufman *et al.* (2016, Natural Language Processing-Enabled and Conventional Data Capture Methods for Input to Electronic Health Records), Chen *et al.* (2018, A Natural Language Processing System That Links Medical Terms in Electronic Health Record Notes to Lay Definitions) or Simon (2018, Natural Language Processing for Healthcare Customers)

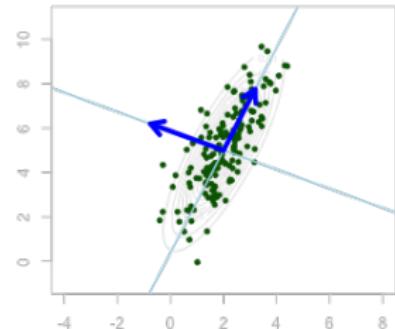


Statistics & ML Toolbox : PCA

Projection method, used to reduce dimensionality

Principal Component Analysis (PCA)

find a small number of directions in input space that explain variation in input data represent data by projecting along those directions



n observations in dimension d , stored in matrix $\mathbf{X} \in \mathcal{M}_{n \times d}$

Natural ideal: linearly project (multiply by a projection matrix) to much lower dimensional space $k < d$

Search for orthogonal directions in space with highest variance

Information is stored in the covariance matrix $\Sigma = \mathbf{X}^\top \mathbf{X}$ when variables are centered

Statistics & ML Toolbox : PCA

Find the k largest eigenvectors of Σ : principal components

Assemble these eigenvectors into a $d \times k$ matrix $\tilde{\Sigma}$

Express d -dimensional vectors x by projecting them to
 m -dimensional \tilde{x} , $\tilde{x} = \tilde{\Sigma}^\top x$

We have data $X \in \mathcal{M}_{n \times d}$, i.e. n vectors $x \in \mathbb{R}^d$,

Let $\omega_1 \in \mathbb{R}^d$ denote weights.

We want to maximize the variance of $z = \omega_1^\top x$

$$\max \left\{ \frac{1}{n} \sum_{i=1}^n (\omega_1^\top x_i - \omega_1^\top \bar{x})^2 \right\} = \max \left\{ \omega_1^\top \Sigma \omega_1 \right\} \text{ s.t. } \|\omega_1\| = 1$$

Statistics & ML Toolbox : PCA

To solve

$$\max \left\{ \frac{1}{n} \sum_{i=1}^n (\omega_1^\top \mathbf{x}_i - \omega_1^\top \bar{\mathbf{x}})^2 \right\} = \max \left\{ \omega_1^\top \boldsymbol{\Sigma} \omega_1 \right\} \text{ s.t. } \|\omega_1\| = 1$$

we can use Lagrange multiplier to solve this constraint optimization problem

$$\max_{\omega_1, \lambda_1} \left\{ \omega_1^\top \boldsymbol{\Sigma} \omega_1 + \lambda (1 - \omega_1^\top \omega_1) \right\}$$

If we differentiate $\boldsymbol{\Sigma} \omega_1 - \lambda_1 \omega_1 = \mathbf{0}$, i.e. $\boldsymbol{\Sigma} \omega_1 = \lambda_1 \omega_1$ i.e. ω_1 is an eigenvector of $\boldsymbol{\Sigma}$, with eigenvalue the Lagrange multiplier. It should be the one with the largest eigenvalue.

Statistics & ML Toolbox : PCA

$$\begin{aligned} & \max \left\{ \omega_2^\top \Sigma \omega_2 \right\} \\ & \text{subject to } \|\omega_2\| = 1 \\ & \quad \omega_2^\top \omega_1 = 0 \end{aligned}$$

The Lagrangian is $\omega_2^\top \Sigma \omega_2 + \lambda_2(1 - \omega_2^\top \omega_2) - \lambda_1 \omega_2^\top \omega_1$

when differentiating, we get $\Sigma \omega_2 = \lambda_2 \omega_2$

λ_2 is the second largest eigenvalue

Set $\mathbf{z} = W\mathbf{x}$ and $\tilde{\mathbf{x}} = M\mathbf{z}$. We want to solve (where ℓ is the standard ℓ_2 norm)

$$\min_{W,M} \left\{ \sum_{i=1}^n \ell(x_i, \tilde{x}_i) \right\}$$

i.e.

$$\min_{W,M} \left\{ \sum_{i=1}^n \ell(x_i, MWx_i) \right\}$$

Statistics & ML Toolbox : GLM

Generalized Linear Model

Flexible generalization of linear regression that allows for y to have distribution other than a normal distribution

The *classical* regression is (matrix form)

$$\underset{n \times 1}{\mathbf{y}} = \underset{n \times k}{\mathbf{X}} \underset{k \times 1}{\beta} + \underset{n \times 1}{\varepsilon} \text{ or } y_i = \mathbf{x}_i^T \beta + \varepsilon_i$$

$$Y_i \sim \mathcal{N}(\mu_i, \sigma^2) \quad , \text{ with } \mu_i = \mathbb{E}[Y_i | \mathbf{x}_i] = \mathbf{x}_i^T \beta$$

It is possible to consider a more general regression.

For a classification $\mathbf{y} \in \{0, 1\}^n$,

$$Y_i \sim \mathcal{B}(p_i) \quad , \text{ with } p_i = \mathbb{E}[Y_i | \mathbf{x}_i] = \frac{e^{\mathbf{x}_i^T \beta}}{1 + e^{\mathbf{x}_i^T \beta}}$$

Recall that we assumed that (Y_i, \mathbf{X}_i) were i.i.d. with unknown distribution \mathbb{P} .

Statistics & ML Toolbox : GLM

In the (standard) linear model, we assume that
 $(Y|\mathbf{X} = \mathbf{x}) \sim \mathcal{N}(\mathbf{x}^\top \boldsymbol{\beta}, \sigma^2)$.

The maximum-likelihood estimator of $\boldsymbol{\beta}$ is then

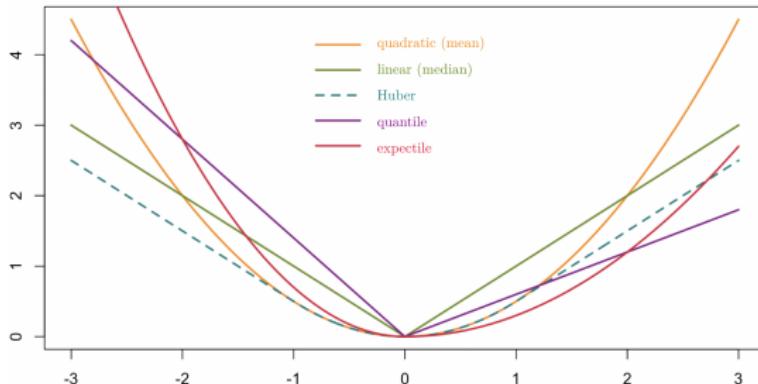
$$\hat{\boldsymbol{\beta}} = \underset{\boldsymbol{\beta} \in \mathbb{R}^{p+1}}{\operatorname{argmax}} \left\{ - \sum_{i=1}^n (y_i - \mathbf{x}_i^\top \boldsymbol{\beta})^2 \right\}$$

For the logistic regression, $(Y|\mathbf{X} = \mathbf{x}) \sim \mathcal{B}(\text{logit}^{-1}(\mathbf{x}^\top \boldsymbol{\beta}))$.
The maximum-likelihood estimator of $\boldsymbol{\beta}$ is then

$$\underset{\boldsymbol{\beta} \in \mathbb{R}^{p+1}}{\operatorname{argmin}} \left\{ \sum_{i=1}^n (y_i \log(\text{logit}^{-1}(\mathbf{x}_i^\top \boldsymbol{\beta})) + (1 - y_i) \log(1 - \text{logit}^{-1}(\mathbf{x}_i^\top \boldsymbol{\beta}))) \right\}$$

Statistics & ML Toolbox : loss and risk

In a general setting, consider some **loss function** $\ell : \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}_+$,
e.g. $\ell(y, m) = (y - m)^2$ (quadratic ℓ_2 norm)

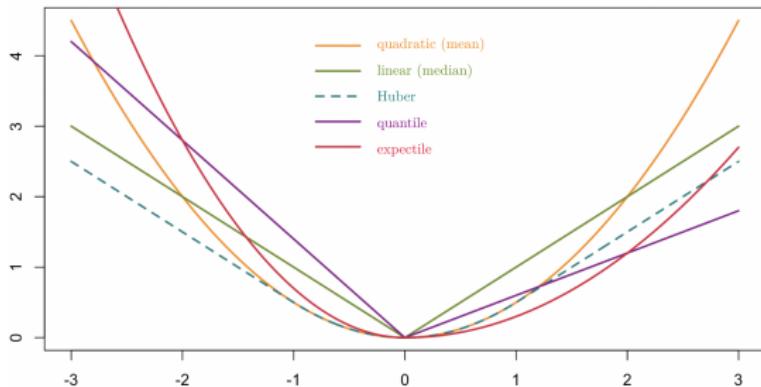


Average Risk

The **average risk** associated with \hat{m}_n is
$$\mathcal{R}_{\mathbb{P}}(\hat{m}_n) = \mathbb{E}_{\mathbb{P}}[\ell(Y, \hat{m}_n(\mathbf{X}))]$$

Statistics & ML Toolbox : loss and risk

In a general setting, consider some **loss function** $\ell : \mathbb{R} \times \mathbb{R} \rightarrow \mathbb{R}_+$,
e.g. $\ell(y, m) = (y - m)^2$ (quadratic ℓ_2 norm)



Empirical Risk

The **empirical risk** associated with \hat{m}_n is

$$\hat{\mathcal{R}}_n(\hat{m}_n) = \frac{1}{n} \sum_{i=1}^n \ell(y_i, \hat{m}_n(x_i)).$$

Statistics & ML Toolbox : loss and risk

If we minimize the average risk, we overfit...

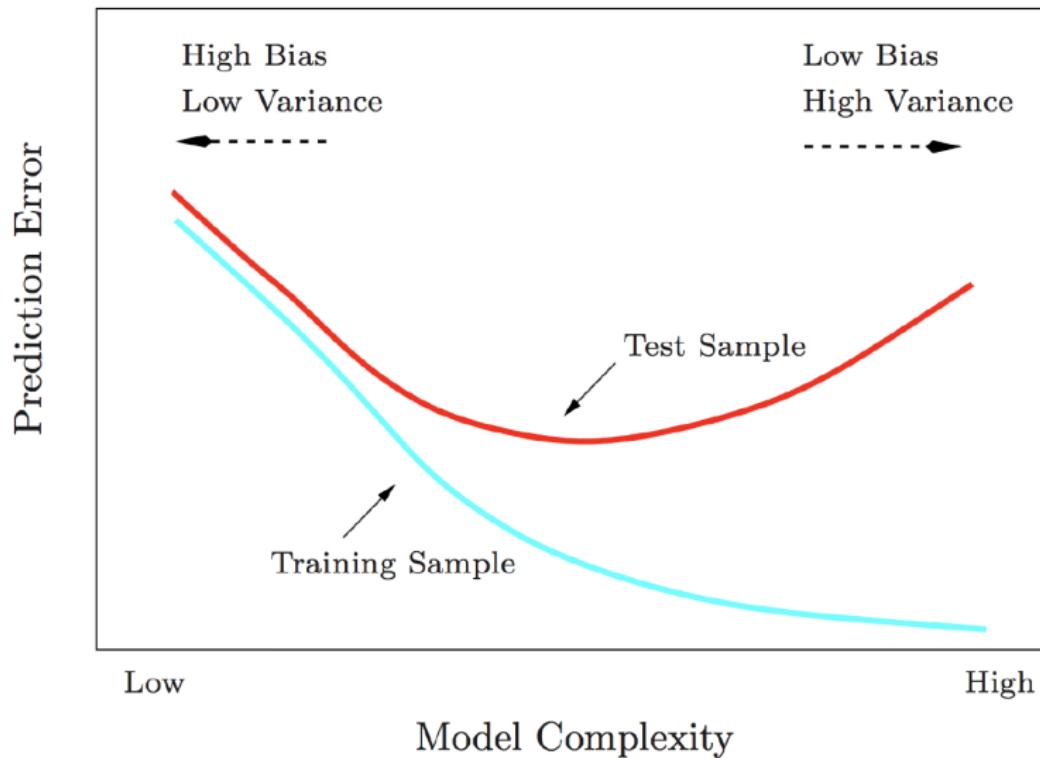
Hold-Out Cross Validation

1. Split $\{1, 2, \dots, n\}$ in T (training) and V (validation)
- 2 . Estimate \hat{m} on sample $(y_i, \mathbf{x}_i), i \in T$:

$$\hat{m}_T = \operatorname{argmin}_{m \in \mathcal{M}} \frac{1}{|T|} \sum_{i \in V} \ell(y_i, m(x_{1,i}, \dots, x_{p,i}))$$

3. Compute $\frac{1}{|V|} \sum_{i \in V} \ell(y'_i, \hat{m}_T(x_{1,i''}, \dots, x_{p,i'}))$

Statistics & ML Toolbox : loss and risk



source: Hastie *et al.* (2009, [The Elements of Statistical Learning](#))

Statistics & ML Toolbox : loss and risk

Leave-one-Out Cross Validation

1. Estimate n models : estimate \hat{m}_{-j} on sample (y_i, \mathbf{x}_i) ,
 $i \in \{1, \dots, n\} \setminus \{j\}$

$$\hat{m}_{-j} = \operatorname{argmin}_{m \in \mathcal{M}} \frac{1}{n-1} \sum_{i \neq j} \ell(y_i, m(x_{1,i}, \dots, x_{p,i}))$$

2. Compute $\frac{1}{n} \sum_{i=1}^n \ell(y_i, \hat{m}_{-i}(\mathbf{x}_i))$

Statistics & ML Toolbox : loss and risk

K-Fold Cross Validation

1. Split $\{1, 2, \dots, n\}$ in K groups V_1, \dots, V_K
2. Estimate K models : estimate \hat{m}_k on sample (y_i, \mathbf{x}_i) , $i \in \{1, \dots, n\} \setminus V_k$
3. Compute $\frac{1}{K} \sum_{k=1}^K \frac{1}{|V_k|} \sum_{i \in V_k} \ell(y_i, \hat{m}_k(x_{1,i}, \dots, X_{p,i}))$

Statistics & ML Toolbox : loss and risk

Bootstrap Cross Validation

1. Generate B bootstrap samples from $\{1, 2, \dots, n\}$, I_1, \dots, I_B
2. Estimate B models : \hat{m}_b on sample (y_i, \mathbf{x}_i) , $i \in I_b$
3. Compute $\frac{1}{B} \sum_{b=1}^B \frac{1}{n - |I_b|} \sum_{i \notin I_b} \ell(y_i, \hat{m}_b(x_{1,i}, \dots, X_{p,i}))$

The probability that $i \notin I_b$ is $\left(1 - \frac{1}{n}\right)^n \sim e^{-1} (= 36.78\%)$

At stage b , we validate on $\sim 36.78\%$ of the dataset.

Statistics & ML Toolbox : loss and risk

$\mathbf{y} \in \mathbb{R}^d$, $\bar{y} = \operatorname{argmin}_{m \in \mathbb{R}} \left\{ \sum_{i=1}^n \frac{1}{n} [\underbrace{y_i - m}_{\varepsilon_i}]^2 \right\}$ is the empirical version of

$$\mathbb{E}[Y] = \operatorname{argmin}_{m \in \mathbb{R}} \left\{ \int [\underbrace{y - m}_{\varepsilon}]^2 dF(y) \right\} = \operatorname{argmin}_{m \in \mathbb{R}} \left\{ \mathbb{E}[\|\underbrace{Y - m}\|_{\ell_2}] \right\}$$

where Y is a random variable.

Thus, $\operatorname{argmin}_{m: \mathbb{R}^k \rightarrow \mathbb{R}} \left\{ \sum_{i=1}^n \frac{1}{n} [\underbrace{y_i - m(\mathbf{x}_i)}_{\varepsilon_i}]^2 \right\}$ is the empirical version of $\mathbb{E}[Y | \mathbf{X} = \mathbf{x}]$.

See Legendre (1805, *Nouvelles méthodes pour la détermination des orbites des comètes*) and Gauß (1809, *Theoria motus corporum coelestium in sectionibus conicis solem ambientium*).

Statistics & ML Toolbox : loss and risk

$\text{med}[\mathbf{y}] \in \operatorname{argmin}_{m \in \mathbb{R}} \left\{ \sum_{i=1}^n \frac{1}{n} \underbrace{|y_i - m|}_{\varepsilon_i} \right\}$ is the empirical version of

$$\text{med}[Y] \in \operatorname{argmin}_{m \in \mathbb{R}} \left\{ \int \underbrace{|y - m|}_{\varepsilon} dF(y) \right\} = \operatorname{argmin}_{m \in \mathbb{R}} \left\{ \mathbb{E}[\underbrace{\|Y - m\|_{\ell_1}}_{\varepsilon}] \right\}$$

where $\mathbb{P}[Y \leq \text{med}[Y]] \geq \frac{1}{2}$ and $\mathbb{P}[Y \geq \text{med}[Y]] \geq \frac{1}{2}$.

$\operatorname{argmin}_{m: \mathbb{R}^k \rightarrow \mathbb{R}} \left\{ \sum_{i=1}^n \frac{1}{n} \underbrace{|y_i - m(\mathbf{x}_i)|}_{\varepsilon_i} \right\}$ is the empirical version of
 $\text{med}[Y | \mathbf{X} = \mathbf{x}]$.

See Boscovich (1757, *De Litteraria expeditione per pontificiam ditionem ad dimetiendo duos meridiani*) and Laplace (1793, *Sur quelques points du système du monde*).

Statistics & ML Toolbox : loss and risk

For least-squares (ℓ_2 norm), we have a strictly convex problem...
classical optimization routines can be used (e.g. gradient descent)

More complicated with the ℓ_1 norm

Consider a sample $\{y_1, \dots, y_n\}$.

To compute the median, solve $\min_{\mu} \left\{ \sum_{i=1}^n |y_i - \mu| \right\}$ which can be solved using linear programming techniques., see Dantzig (1963, [Linear programming](#)).

More precisely, this problem is equivalent to $\min_{\mu, \mathbf{a}, \mathbf{b}} \left\{ \sum_{i=1}^n a_i + b_i \right\}$ with $a_i, b_i \geq 0$ and $y_i - \mu = a_i - b_i, \forall i = 1, \dots, n$.

Statistics & ML Toolbox : loss and risk

For a quantile, the linear program is $\min_{\mu, \mathbf{a}, \mathbf{b}} \left\{ \sum_{i=1}^n \tau a_i + (1 - \tau) b_i \right\}$

with $a_i, b_i \geq 0$ and $y_i - \mu = a_i - b_i, \forall i = 1, \dots, n.$

This is extended to the quantile regression,

$$\min_{\beta^\tau, \mathbf{a}, \mathbf{b}} \left\{ \sum_{i=1}^n \tau a_i + (1 - \tau) b_i \right\}$$

with $a_i, b_i \geq 0$ and $y_i - \mathbf{x}_i^\top \beta^\tau = a_i - b_i, \forall i = 1, \dots, n.$

But is the following problem the one we should really care about ?

$$m^* = \operatorname{argmin}_{m \in \mathcal{M}} \left\{ \sum_{i=1}^n \ell(y_i, m(\mathbf{x}_i)) \right\}$$

Statistics & ML Toolbox : Penalization

Usually, there are **tuning** parameters p . To avoid overfit,

1. use cross-validation

$$m_{p^*}^* = \operatorname{argmin}_p \left\{ \sum_{i \in \text{validation}} \ell(y_i, m_p^*(x_i)) \right\}$$

$$\text{where } m_p^* = \operatorname{argmin}_{m \in \mathcal{M}_p} \left\{ \sum_{i \in \text{training}} \ell(y_i, m(x_i)) \right\}$$

2. use penalization

$$m^* = \operatorname{argmin}_{m \in \mathcal{M}} \left\{ \sum_{i=1}^n \ell(y_i, m(x_i)) + \text{penalization}(m) \right\}$$

Statistics & ML Toolbox : Penalization

Penalization often yield bias... but it might interesting if the risk is the mean squared error...

Consider a parametric model, with true (unknown) parameter θ , then

$$\text{mse}(\hat{\theta}) = \mathbb{E}[(\hat{\theta} - \theta)^2] = \underbrace{\mathbb{E}[(\hat{\theta} - \mathbb{E}[\hat{\theta}])^2]}_{\text{variance}} + \underbrace{\mathbb{E}[(\mathbb{E}[\hat{\theta}] - \theta)^2]}_{\text{bias}^2}$$

One can think of a **shrinkage** of an unbiased estimator,

Let $\tilde{\theta}$ denote an unbiased estimator of θ ,

$$\hat{\theta} = \frac{\theta^2}{\theta^2 + \text{mse}(\tilde{\theta})} \cdot \tilde{\theta} \leq \tilde{\theta}$$

satisfies $\text{mse}(\hat{\theta}) \leq \text{mse}(\tilde{\theta})$.

Statistics & ML Toolbox : Ridge

as in Wieringen (2018 [Lecture notes on ridge regression](#))

Ridge Estimator (OLS)

$$\hat{\beta}_\lambda^{\text{ridge}} = \operatorname{argmin} \left\{ \sum_{i=1}^n (y_i - \mathbf{x}_i^\top \boldsymbol{\beta})^2 + \lambda \sum_{j=1}^p \beta_j^2 \right\}$$
$$\hat{\beta}_\lambda^{\text{ridge}} = (\mathbf{X}^\top \mathbf{X} + \lambda \mathbb{I})^{-1} \mathbf{X}^\top \mathbf{y}$$

Ridge Estimator (GLM)

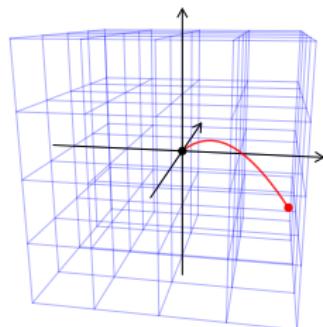
$$\hat{\beta}_\lambda^{\text{ridge}} = \operatorname{argmin} \left\{ - \sum_{i=1}^n \log f(y_i | \mu_i = g^{-1}(\mathbf{x}_i^\top \boldsymbol{\beta})) + \frac{\lambda}{2} \sum_{j=1}^p \beta_j^2 \right\}$$

Statistics & ML Toolbox : Ridge

Observe that if $\mathbf{x}_{j_1} \perp \mathbf{x}_{j_2}$, then

$$\hat{\boldsymbol{\beta}}_{\lambda}^{\text{ridge}} = [1 + \lambda]^{-1} \hat{\boldsymbol{\beta}}^{\text{ols}}$$

which explain relationship with shrinkage.
But not in the general case...



Smaller mse

There exists λ such that $\text{mse}[\hat{\boldsymbol{\beta}}_{\lambda}^{\text{ridge}}] \leq \text{mse}[\hat{\boldsymbol{\beta}}^{\text{ols}}]$

Statistics & ML Toolbox : Ridge

From a Bayesian perspective,

$$\underbrace{\mathbb{P}[\theta|y]}_{\text{posterior}} \propto \underbrace{\mathbb{P}[y|\theta]}_{\text{likelihood}} \cdot \underbrace{\mathbb{P}[\theta]}_{\text{prior}} \quad \text{i.e.} \quad \log \mathbb{P}[\theta|y] = \underbrace{\log \mathbb{P}[y|\theta]}_{\text{log likelihood}} + \underbrace{\log \mathbb{P}[\theta]}_{\text{penalty}}$$

If β has a prior $\mathcal{N}(\mathbf{0}, \tau^2 \mathbb{I})$ distribution, then its posterior distribution has mean

$$\mathbb{E}[\beta|y, X] = \left(X^T X + \frac{\sigma^2}{\tau^2} \mathbb{I} \right)^{-1} X^T y.$$

Statistics & ML Toolbox : LASSO

LASSO Estimator (OLS)

$$\hat{\beta}_\lambda^{\text{lasso}} = \operatorname{argmin} \left\{ \sum_{i=1}^n (y_i - \mathbf{x}_i^\top \boldsymbol{\beta})^2 + \lambda \sum_{j=1}^p |\beta_j| \right\}$$

LASSO Estimator (GLM)

$$\hat{\beta}_\lambda^{\text{lasso}} = \operatorname{argmin} \left\{ - \sum_{i=1}^n \log f(y_i | \mu_i = g^{-1}(\mathbf{x}_i^\top \boldsymbol{\beta})) + \frac{\lambda}{2} \sum_{j=1}^p |\beta_j| \right\}$$

Statistics & ML Toolbox : sparsity

In several applications, k can be (very) large, but a lot of features are just noise: $\beta_j = 0$ for many j 's. Let s denote the number of relevant features, with $s \ll k$, cf Hastie, Tibshirani & Wainwright (2015, *Statistical Learning with Sparsity*),

$$s = \text{card}\{\mathcal{S}\} \text{ where } \mathcal{S} = \{j; \beta_j \neq 0\}$$

The model is now $y = \mathbf{X}_{\mathcal{S}}^T \boldsymbol{\beta}_{\mathcal{S}} + \varepsilon$, where $\mathbf{X}_{\mathcal{S}}^T \mathbf{X}_{\mathcal{S}}$ is a full rank matrix.

Statistics & ML Toolbox : sparsity

ℓ_0 -regularization Estimator (OLS)

$$\hat{\beta}_\lambda^{\text{lasso}} = \operatorname{argmin} \left\{ \sum_{i=1}^n (y_i - \mathbf{x}_i^\top \boldsymbol{\beta})^2 + \lambda \|\boldsymbol{\beta}\|_{\ell_0} \right\}$$

where $\|\boldsymbol{a}\|_{\ell_0} = \sum \mathbf{1}(|a_j| > 0) = \dim(\boldsymbol{\beta})$.

More generally

ℓ_p -regularization Estimator (OLS)

$$\hat{\beta}_\lambda^{\text{lasso}} = \operatorname{argmin} \left\{ \sum_{i=1}^n (y_i - \mathbf{x}_i^\top \boldsymbol{\beta})^2 + \lambda \|\boldsymbol{\beta}\|_{\ell_p} \right\}$$

- **sparsity** is obtained when $p \leq 1$
- **convexity** is obtained when $p \geq 1$

Statistics & ML Toolbox : sparcity

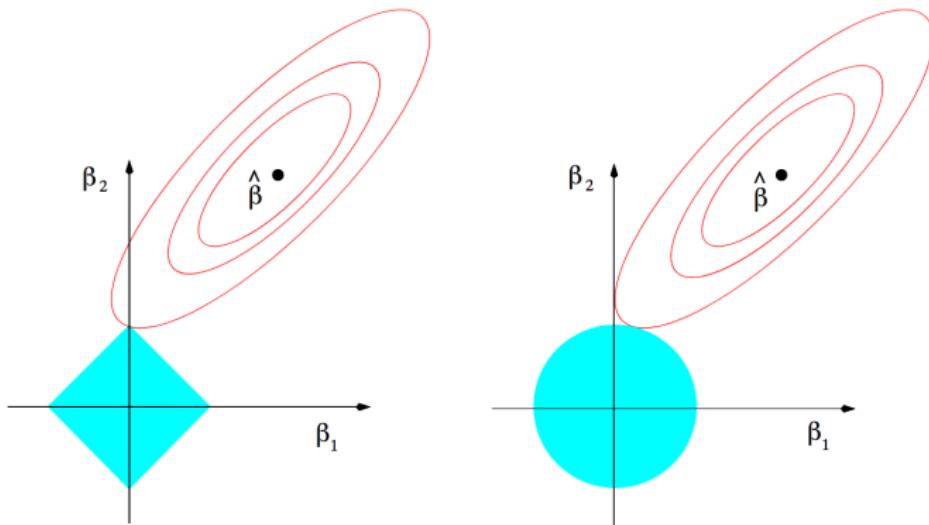


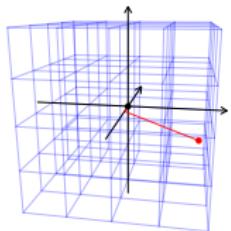
FIGURE 3.11. Estimation picture for the lasso (left) and ridge regression (right). Shown are contours of the error and constraint functions. The solid blue areas are the constraint regions $|\beta_1| + |\beta_2| \leq t$ and $\beta_1^2 + \beta_2^2 \leq t^2$, respectively, while the red ellipses are the contours of the least squares error function.

source: Hastie et al. (2009, [The Elements of Statistical Learning](#))

Statistics & ML Toolbox : LASSO

In the orthogonal case, $\mathbf{X}^T \mathbf{X} = \mathbb{I}$,

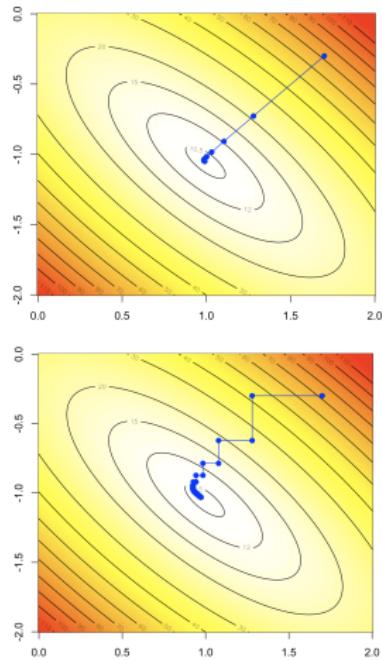
$$\hat{\beta}_{k,\lambda}^{\text{lasso}} = \text{sign}(\hat{\beta}_k^{\text{ols}}) \left(|\hat{\beta}_k^{\text{ols}}| - \frac{\lambda}{2} \right)$$



i.e. the LASSO estimate is related to the soft threshold function...

LASSO Coordinate Descent Algorithm

1. Set $\beta_0 = \hat{\beta}$
- 2 . For $k = 1, \dots$
for $j = 1, \dots, p$
 - (i) compute $R_j = \mathbf{x}_j^\top (\mathbf{y} - \mathbf{X}_{-j} \beta_{k-1(-j)})$
 - (ii) set $\beta_{k,j} = R_j \cdot \left(1 - \frac{\lambda}{2|R_j|}\right)_+$
3. The final estimate β_k is $\hat{\beta}_\lambda$



The **softmax** function $\log(1 + e^{x-s})$ can be used as a smooth version of $(x - s)_+$

Statistics & ML Toolbox : trees

Use recursive binary splitting to grow a tree

Trees (CART)

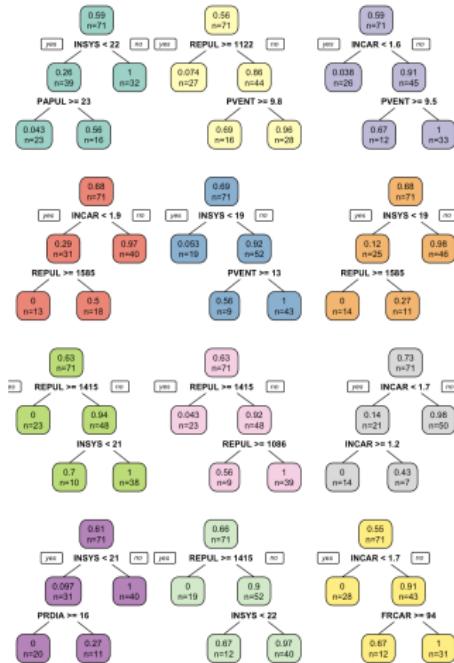
Each interior node corresponds to one of the input variables; there are edges to children for each of the possible values of that input variable. Each leaf represents a value of the target variable given the values of the input variables represented by the path from the root to the leaf.

Statistics & ML Toolbox : forests

Bagging (Bootstrap + Aggregation)

1. For $k = 1, \dots$
 - (i) draw a bootstrap sample from (y_i, \mathbf{x}_i) 's
 - (ii) estimate a model \hat{m}_k on that sample
2. The final model is

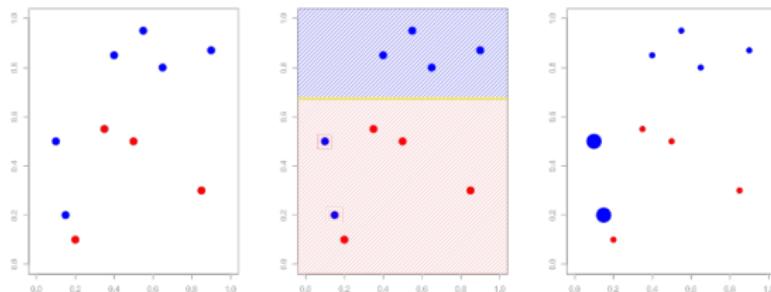
$$m^*(\cdot) = \frac{1}{\kappa} \sum_{k=1}^{\kappa} \hat{m}_k(\cdot)$$



Statistics & ML Toolbox : boosting

Bosting & Sequential Learning

$$m_k(\cdot) = m_{k-1}(\cdot) + \operatorname{argmin}_{h \in \mathcal{H}} \left\{ \sum_{i=1}^n \ell(\underbrace{y_i - m_{k-1}(\mathbf{x}_i)}_{\varepsilon_i}, h(\mathbf{x}_i)) \right\}$$



Statistics & ML Toolbox : Neural Nets

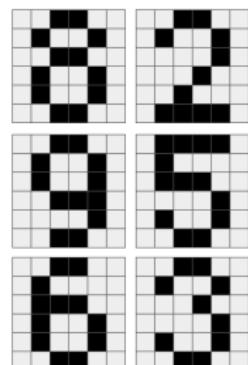
Very popular for pictures...

Picture x_i is

- a $n \times n$ matrix in $\{0, 1\}^{n^2}$ for black & white
- a $n \times n$ matrix in $[0, 1]^{n^2}$ for grey-scale
- a $3 \times n \times n$ array in $([0, 1]^3)^{n^2}$ for color
- a $T \times 3 \times n \times n$ tensor in $(([0, 1]^3)^T)^{n^2}$ for video

y here is the label ("8", "9", "6", etc)

Suppose we want to recognize a "6" on a picture



$$m(x) = \begin{cases} +1 & \text{if } x \text{ is a "6"} \\ -1 & \text{otherwise} \end{cases}$$

Statistics & ML Toolbox : Neural Nets

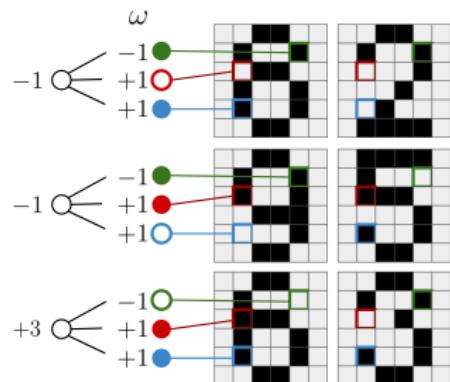
Consider some specific pixels, and associate weights ω such that

$$\hat{m}(\mathbf{x}) = \text{sign} \left(\sum_{i,j} \omega_{i,j} x_{i,j} \right)$$

where

$$x_{i,j} = \begin{cases} +1 & \text{if pixel } x_{i,j} \text{ is black} \quad \blacksquare \\ -1 & \text{if pixel } x_{i,j} \text{ is white} \quad \square \end{cases}$$

for some weights $\omega_{i,j}$ (that can be negative...)



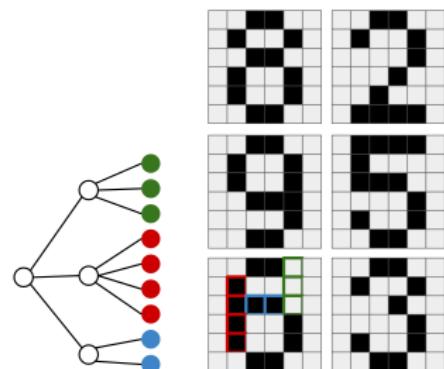
Statistics & ML Toolbox : Neural Nets

A deep network is a network with a lot of layers

$$\hat{m}(\mathbf{x}) = \text{sign} \left(\sum_i \omega_i \hat{m}_i(\mathbf{x}) \right)$$

where \hat{m}_i 's are outputs of previous neural nets

Those layers can capture shapes in some areas
nonlinearities, cross-dependence, etc



Binary Threshold Neuron & Perceptron

If $x \in \{0, 1\}^P$, McCulloch & Pitts (1943, [A logical calculus of the ideas immanent in nervous activity](#)) suggested a simple model, with **threshold** b

$$y_i = f \left(\sum_{j=1}^p x_{j,i} \right) \text{ where } f(x) = \mathbf{1}(x \geq b)$$

where $\mathbf{1}(x \geq b) = +1$ if $x \geq b$ and 0 otherwise, or (equivalently)

$$y_i = f \left(\omega + \sum_{j=1}^p x_{j,i} \right) \text{ where } f(x) = \mathbf{1}(x \geq 0)$$

with **weight** $\omega = -b$. The trick of adding 1 as an input was very important !

$\omega = -1$ is the **or** logical operator : $y_i = 1$ if $\exists j$ such that $x_{j,i} = 1$
 $\omega = -p$, is the **and** logical operator : $y_i = 1$ if $\forall j$, $x_{j,i} = 1$

Binary Threshold Neuron & Perceptron

but not possible for the **xor** logical operator : $y_i = 1$ if $x_{1,i} \neq x_{2,i}$

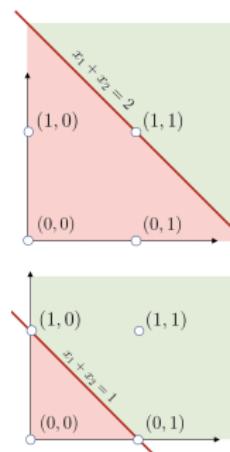
Rosenblatt (1961, *Principles of Neurodynamics: Perceptron & Theory of Brain Mechanism*) considered the extension where x 's are real-valued, with **weight** $\omega \in \mathbb{R}^p$

$$y_i = f \left(\sum_{j=1}^p \omega_j x_{j,i} \right) \text{ where } f(x) = \mathbf{1}(x \geq b)$$

where $\mathbf{1}(x \geq b) = +1$ if $x \geq b$ and 0 otherwise, or (equivalently)

$$y_i = f \left(\omega_0 + \sum_{j=1}^p \omega_j x_{j,i} \right) \text{ where } f(x) = \mathbf{1}(x \geq 0)$$

with **weights** $\omega \in \mathbb{R}^{p+1}$



Binary Threshold Neuron & Perceptron

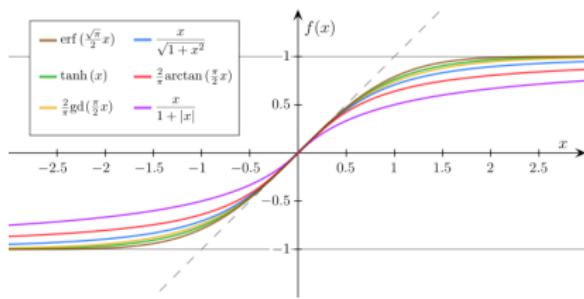
Minsky & Papert (1969, [Perceptrons: an Introduction to Computational Geometry](#)) proved that perceptron were a linear separator, not very powerful

Define the **sigmoid** function $f(x) = \frac{1}{1 + e^{-x}} = \frac{e^x}{e^x + 1}$ (= **logistic** function).

This function f is called the **activation function**.

If $y \in \{-1, +1\}$, one can consider the hyperbolic tangent

$f(x) = \frac{(e^x - e^{-x})}{(e^x + e^{-x})}$ or the inverse tangent function (see [wikipedia](#)).



Statistics & ML Toolbox : Neural Nets

So here, for a classification problem,

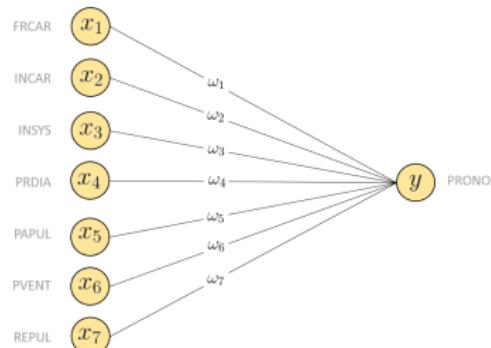
$$y_i = f \left(\omega_0 + \sum_{j=1}^p \omega_j x_{j,i} \right) = p(\mathbf{x}_i)$$

The error of that model can be computed using the quadratic loss function,

$$\sum_{i=1}^n (y_i - p(\mathbf{x}_i))^2$$

or cross-entropy

$$\sum_{i=1}^n (y_i \log p(\mathbf{x}_i) + [1 - y_i] \log [1 - p(\mathbf{x}_i)])$$



Statistics & ML Toolbox : Neural Nets

Consider a single hidden layer, with 3 different neurons, so that

$$p(\mathbf{x}) = f \left(\omega_0 + \sum_{h=1}^3 \omega_h f_h \left(\omega_{h,0} + \sum_{j=1}^p \omega_{h,j} x_j \right) \right)$$

or

$$p(\mathbf{x}) = f \left(\omega_0 + \sum_{h=1}^3 \omega_h f_h \left(\omega_{h,0} + \mathbf{x}^\top \boldsymbol{\omega}_h \right) \right)$$

