

Problem class 3: Hypothesis tests ; Inference under model uncertainty ; Hierarchical Bayes

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1 Hypothesis test

Exercise 1. (**) Consider a Bayesian model

$$\begin{cases} x_i | \lambda & \stackrel{\text{iid}}{\sim} \text{Pn}(\lambda), \quad \forall i = 1, \dots, n \\ \lambda & \sim \Pi(\lambda) \end{cases}$$

Hint-1 Poisson distribution has PMF: $\text{Pn}(x|\lambda) = \frac{1}{x!} \lambda^x \exp(-\lambda) 1_{\mathbb{N}}(x)$

Hint-2 Gamma distribution has PDF: $\text{Ga}(x|a, b) = \frac{b^a}{\Gamma(a)} x^{a-1} \exp(-bx) 1_{(0, \infty)}(x)$, with $E(x) = a/b$, $\text{Var}(x) = a/b^2$.

Hint-3 Negative Binomial distribution has PMF: $\text{Nb}(x|r, \theta) = \binom{r+x-1}{r-1} \theta^r (1-\theta)^x 1_{\mathbb{N}}(x)$. with $\theta \in (0, 1)$, $r \in \mathbb{N}$.

Consider that we are interested in testing the hypothesis whether $\lambda = \lambda_0$, (where λ_0 is a fixed known number), or not.

Let $\pi_j = P(H_j)$ be the marginal prior probability of hypothesis H_j .

1. Design the test of hypotheses in Bayesian framework: Namely, set pair of hypotheses, specify priors, and compute the associated Bayes Factor.
 2. Compute the posterior probability that $\lambda = \lambda_0$.
 3. Perform the hypothesis test to test if $\lambda = 2$ or not based on the Jeffrey's scaling rule, by considering that
 - we have collected two observations $x_1 = 2, x_2 = 3$,
 - a priori the probability that $\{\lambda = 2\}$ is 0.5,
 - given $\{\lambda \neq 2\}$, the prior distr. of λ is a conjugate one with $E(\lambda) = 2$, and $\text{Var}(\lambda) = 1$.
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2 Inference under model uncertainty

Exercise 2. (★★) Let $B_{k,j}(y)$ be the Bayes factor of model \mathcal{M}_k against model \mathcal{M}_j , for all $\forall k, i, j \in \mathcal{K}$. . Show that $B_{k,j}(y) = B_{k,i}(y)B_{i,j}(y)$, for all $\forall k, i, j \in \mathcal{K}$.

3 Hierarchical Bayes

Exercise 3. (★★)[Relevance Vector Machine]

Regarding the statistical model: Long story short (supplementary material)

Consider that we are interested in recovering the mapping

$$x \mapsto \eta(x)$$

in the sense that $y \in \mathbb{R}$ is the response (output quantity) that depends on $x = (x_1, \dots, x_d) \in \mathcal{X} \subseteq \mathbb{R}^d$ which is the independent variable (input quantity) in a procedure; E.g.:

- y : precipitation in log scale
- $x = (\text{longitude}, \text{latitude})$: geographical coordinates.

Consider a set of observed data $\{(y_i, x_i)\}_{i=1}^n$, which may be contaminated by additive noise of unknown variance; i.e.

$$y_i = \eta(x_i) + \epsilon_i,$$

where $\epsilon_i \stackrel{\text{iid}}{\sim} \mathcal{N}(0, \sigma^2)$ and $\sigma^2 > 0$ is unknown. We wish to recover $\eta(x)$ by using the Tikhonov regularization on the functional space \mathcal{H} such that

$$\eta = \arg \min_{\tilde{\eta} \in \mathcal{H}} \left\{ \sum_{i=1}^n L(y_i - \tilde{\eta}(x_i)) + \lambda \|\tilde{\eta}\|_{\mathcal{H}}^2 \right\} \quad (1)$$

By assuming that \mathcal{H} is a Reproducing Kernel Hilbert Space (RKHS), the solution to (1) is such that

$$\eta(x) = \beta_0 + \sum_{j=1}^n k(x, x_j) \beta_j = k(x)^\top \beta$$

where $k(x) = (1, k(x, x_1), \dots, k(x, x_n))^\top$, $k(x, x_j)$ is the reproducing kernel (such as $k_\phi(x, x_j) = \exp(-\phi \|x - x_j\|^2)$ for some known parameter $\phi > 0$), and $\beta \in \mathbb{R}^{n+1}$ is an unknown vector.

Consider the following Bayesian model¹

$$\begin{cases} y|\beta, \sigma^2 & \sim \mathcal{N}(K\beta, I\sigma^2) \\ \beta|\lambda & \sim \mathcal{N}(0, D^{-1}), \quad D = (\lambda_0, \lambda_1, \dots, \lambda_n) \\ \lambda_i & \stackrel{\text{iid}}{\sim} d\Pi(\lambda_i) \propto \lambda_i^{a-1} \exp(-b\lambda_i), \quad \forall i = 1, \dots, n \\ \sigma^2 & \sim d\Pi(\sigma^2) \propto (\sigma^2)^{c-1} \exp(-\frac{1}{\sigma^2}d) d\sigma^2 \\ \beta, \sigma^2 & \text{a priori independent} \end{cases}$$

where K is a known matrix with size $n \times (n+1)$ such that

$$K = \begin{bmatrix} 1 & k(x_1, x_1) & \cdots & k(x_1, x_n) \\ \vdots & \vdots & \ddots & \vdots \\ 1 & k(x_n, x_1) & \cdots & k(x_n, x_n) \end{bmatrix}.$$

The quantities $a > 0$, $b > 0$, $c > 0$, $d > 0$, and $\phi > 0$ are considered as fixed.

¹Dixit, A., & Roy, V. (2021). Posterior impropriety of some sparse Bayesian learning models. Statistics & Probability Letters, 171, 109039.

1. When $b = 0$, show that a necessary condition for a valid posterior inference is $a \in (-1/2, 0)$ for any choice of prior for τ (i.e. any choice of (c, d)).
2. Let $P = K (K^\top K)^{-1} K^\top$. Show that (2a) and (2b) are sufficient conditions for the Bayesian model to lead to a valid posterior inference
 - (a) if $a > 0$ and $b > 0$, or
 - (b) if $y^\top (I - P) y + 2d > 0$ and $c > -\frac{n}{2}$
3. Does the the improper Uniform prior on the joint $\log(\lambda_i)$ and $\log(\sigma^2)$, i.e. $\pi(\log(\lambda_i), \log(\sigma^2)) \propto 1$, lead to a valid inference?
4. Does the Jeffreys' prior $\pi(\lambda_i) \propto 1/\lambda_i$ lead to a valid inference?

Hint-1:

$$(y - K\beta)^\top (y - K\beta) + (\beta - \mu)^\top V^{-1}(\beta - \mu) = (\beta - \mu^*)^\top (V^*)^{-1}(\beta - \mu^*) + S^*;$$

$$S^* = \mu^\top V^{-1}\mu - (\mu^*)^\top (V^*)^{-1}(\mu^*) + y^\top y; \quad V^* = (V^{-1} + K^\top K)^{-1}; \quad \mu^* = V^* (V^{-1}\mu + K^\top y)$$

Hint-2: Sherman-Morrison-Woodbury formula:

$$(A + UCV)^{-1} = A^{-1} - A^{-1}U(C^{-1} + VA^{-1}U)^{-1}VA^{-1}$$

Hint-3:

$$-\frac{y^\top y}{2\sigma^2} \leq -\frac{y^\top (I\sigma^2 + KD^{-1}K^\top)^{-1}y}{2} \leq -\frac{1}{2\sigma^2}y^\top (I - P)y$$

where $P = K (K^\top K)^{-1} K$.

Hint-4: It is given that $\int_{(0,\infty)} \frac{t^{-(a+1)}}{(\xi+t)^{1/2}} dt < \infty$ if and only if $a \in (-1/2, 0)$.
