Handout 2: Revision in mixture of probability distributions

Lecturer & author: Georgios P. Karagiannis georgios.karagiannis@durham.ac.uk

Aim

To practice on probability calculations on compound distribution functions. To become familiar with distributions, Inverted Gamma, multivariate Normal, and multivariate Student T distributions.

References:

- DeGroot, M. H. (1970, or 2005). Optimal statistical decisions (Vol. 82). John Wiley & Sons.
  - Part one: Survey of probability theory. Chapters 1-5 ; However the treatment of the Normal and Student T distributions is different than ours.
- Raiffa, H., & Schlaifer, R. (1961). Applied statistical decision theory.
  - Chapters 8.2, 8.3 ; However the treatment of the Normal and Student T distributions is different than ours.

Web-applets

- `https://georgios-stats-3.shinyapps.io/demo_multivariatenormaldistribution/`
- `https://github.com/georgios-stats/Shiny_applets/tree/master/demo_MultivariateNormalDistribution`

1 Mixture probability distribution

Definition 1. Consider a random variable x distributed according to $F(x|z)$ with a pdf/pmf $f(x|z)$ labeled by an unknown parameter z. Consider that z is distributed according to distribution $\Pi(z)$ with pdf/pmf $\pi(z)$. Then:

1. This dependency can be represented as

$$x|z \sim F(x|z)$$
$$z \sim \Pi(z)$$

   where $F(x|z)$ is called conditional (or parametrized) distribution, $\Pi(z)$ is called mixing (or latent) distribution; and z is called mixing (or latent) variable.

2. Distribution $G(x)$ that results by integrating the conditional distribution $F(x|z)$ with respect to $\Pi(z)$ as

$$G(x) = \int F(x|z)\mathrm{d}\Pi(z)$$

   is called the mixture (or compound) distribution of x. The mixture (or compound) distribution $G(x)$ has PDF/PMF

$$g(x) = \int f(x|z)\mathrm{d}\Pi(z) = \begin{cases} \int f(x|z)\pi(z)\mathrm{d}z & , z \text{ cont.} \\ \sum_{\forall z} f(x|z)\pi(z) & , z \text{ discr.} \end{cases}$$

Definition 2. $G(x)$ is also called continuous mixture distribution when the mixing/latent variable z is continuous. $G(x)$ is also called finite mixture when z is discrete.

Remark 3. Recall from Handout 1 that

$$\mathrm{E}_G(x) = \mathrm{E}_\Pi\left(\mathrm{E}_F(x|z)\right)$$
$$\mathrm{Var}_G(x) = \mathrm{E}_\Pi\left(\mathrm{Var}(x|z)\right) + \mathrm{Var}_\Pi\left(\mathrm{E}_F(x|z)\right)$$

2 Inverted Gamma distribution $x|a,b \sim \mathbf{IG}(a,b)$

Definition 4. The random variable $x \in (0,+\infty)$ follows an Inverted Gamma distribution $x \sim \mathrm{IG}(a,b)$, if and only if $x = \frac{1}{y}$ follows a Gamma distribution, $y \sim \mathrm{Ga}(a,b)$, with $a > 0$ and $b > 0$.

Example 5. Let $x \sim \mathrm{IG}(a,b)$, then the PDF of x is

$$f_{\mathrm{IG}(a,b)}(x) = \frac{b^a}{\Gamma(a)} x^{-a-1} \exp\left(-\frac{b}{x}\right) 1_{(0,+\infty)}(x) \tag{1}$$

Solution. It is

$$f_{\mathrm{IG}(a,b)}(x) = f_{\mathrm{G}(a,b)}(\frac{1}{x}) \left|\frac{\mathrm{d}}{\mathrm{d}x}(\frac{1}{x})\right| = \frac{b^a}{\Gamma(a)} \left(\frac{1}{x}\right)^{a-1} \exp\left(-\frac{b}{x}\right) 1_{(0,+\infty)}\left(\frac{1}{x}\right) \left|-\frac{1}{x^2}\right|$$

Example 6. Let a random variable $x \sim \mathrm{IG}(a,b)$, then

$$\mathrm{E}_{\mathrm{IG}(a,b)}(x) = \frac{b}{a-1};\ a > 1 \qquad \text{and} \qquad \mathrm{Var}_{\mathrm{IG}(a,b)}(x) = \frac{b^2}{(a-1)^2(a-2)};\ a > 2$$

Solution. It is

$$\mathrm{E}_{\mathrm{IG}(a,b)}(x) = \int x f_{\mathrm{IG}(a,b)}(x)\mathrm{d}x \quad = \int_{(0,+\infty)} x \frac{b^a}{\Gamma(a)} x^{-a-1} \exp(-\frac{b}{x})\mathrm{d}x$$

Assume that $a > 1$. Then

$$\begin{aligned}
\mathrm{E}_{\mathrm{IG}(a,b)}(x) &= \int_{(0,+\infty)} \frac{b^{a-1}}{\Gamma(a)} b \frac{\Gamma(a-1)}{\Gamma(a-1)} x^{-a+1-1} \exp(-\frac{b}{x})\mathrm{d}x \quad = b\frac{\Gamma(a-1)}{\Gamma(a)} \int_{(0,+\infty)} \frac{b^{a-1}}{\Gamma(a-1)} x^{-a+1-1} \exp(-\frac{b}{x})\mathrm{d}x \\
&= b\frac{\Gamma(a-1)}{\Gamma(a)} \int_{(0,+\infty)} \frac{b^{a-1}}{\Gamma(a-1)} x^{-a+1-1} \exp(-\frac{b}{x})\mathrm{d}x = b\frac{\Gamma(a-1)}{(a-1)\Gamma(a-1)} \int f_{\mathrm{IG}(a-1,b)}(x)\mathrm{d}x \quad = \frac{b}{a-1}
\end{aligned}$$

Similarly

$$\begin{aligned}
\mathrm{E}_{\mathrm{IG}(a,b)}(x^2) &= \int_{(0,+\infty)} x^2 \frac{b^a}{\Gamma(a)} x^{-a-1} \exp\left(-\frac{b}{x}\right) \mathrm{d}x = \ldots = b\frac{\Gamma(a-1)}{(a-1)\Gamma(a-1)} \int x f_{\mathrm{IG}(a-1,b)}(x)\mathrm{d}x \\
&= \frac{b}{a-1}\frac{b}{a-2};\ a > 2
\end{aligned}$$

So

$$\mathrm{Var}_{\mathrm{IG}(a,b)}(x) = \mathrm{E}_{\mathrm{IG}(a,b)}(x^2) - \left(\mathrm{E}_{\mathrm{IG}(a,b)}(x)\right)^2 = \frac{b^2}{(a-1)^2(a-2)}$$

3 Multivariate Normal distribution $x|\mu,\Sigma \sim \mathbf{N}_d(\mu,\Sigma)$

Definition 7. A d-dimensional random variable $x \in \mathbb{R}^d$ is said to have a multivariate Normal (Gaussian) distribution, if for every d-dimensional fixed vector $\alpha \in \mathbb{R}^d$, the random variable $\alpha^\top x$ has a univariate Normal (Gaussian) distribution.

Proposition 8. *A random vector $x \in \mathbb{R}^d$ has a d-dimensional Normal distribution with mean $\mu = E(x)$ and covariance matrix $\Sigma = Var(x)$ if and only if random vector $x \in \mathbb{R}^d$ has a characteristic function*

$$\varphi_x(t) = \exp\left(it^\top \mu - \frac{1}{2} t^\top \Sigma t\right) \tag{2}$$

Hence: the d-dimensional Normal distribution is uniquely defined by the mean and the covariance matrix.

Proof. ($\Longrightarrow$) If x has a d-dimensional distribution then the characteristic function is $\varphi_x(t) = \varphi_{t^\top x}(1)$. Since x has a d-dimensional Normal distribution with mean $\mu = \mathrm{E}(x)$ and covariance matrix $\Sigma = \mathrm{Var}(x)$, $t^\top x$ has a Normal distribution with mean $\mathrm{E}\left(t^\top x\right) = t^\top \mu$ and variance $\mathrm{Var}\left(t^\top x\right) = t^\top \Sigma t$. Then

$$\varphi_x(t) = \varphi_{t^\top x}(1) = \exp\left(i\mathrm{E}\left(t^\top x\right) - \frac{1}{2}\mathrm{Var}\left(t^\top x\right)\right) = \exp\left(it^\top \mathrm{E}\left(x\right) - \frac{1}{2} t^\top \mathrm{Var}\left(x\right) t\right) = \exp\left(it^\top \mu - \frac{1}{2} t^\top \Sigma t\right)$$

($\Longleftarrow$) If random vector $x \in \mathbb{R}^d$ has a characteristic function $\varphi_x(t) = \exp(it^\top \mu - \frac{1}{2} t^\top \Sigma t)$, then for every d-dimensional fixed vector $\alpha \in \mathbb{R}^d$ the characteristic function of $\alpha^\top x$ is

$$\varphi_{\alpha^\top x}(t) = \varphi_x(t\alpha) = \exp\left(it\alpha^\top \mu - \frac{1}{2} t\alpha^\top \Sigma \alpha t\right) = \exp\left(it\left(\alpha^\top \mu\right) - \frac{1}{2}\left(\alpha^\top \Sigma \alpha\right) t^2\right)$$

which defines that $\alpha^\top x$ has a univariate Normal distribution with mean $\alpha^\top \mu$ and variance $\alpha^\top \Sigma \alpha$. □

Notation 9. We denote the d-dimensional Normal distribution with mean μ and covariance matrix $\Sigma \geq 0$ as $\mathrm{N}_d(\mu, \Sigma)$.

Notation 10. The d-dimensional standardized Normal distribution is $\mathrm{N}_d(0, I)$.

Proposition 11. *Let random variable $x \sim N_d(\mu, \Sigma)$* , fixed vector $c \in \mathbb{R}^q$ and fixed matrix $A \in \mathbb{R}^q \times \mathbb{R}^d$. The random vector $y = c + Ax$ *has distribution* $y \sim N_q\left(c + A\mu, A\Sigma A^\top\right)$.

Proof. First I show that y is Normally distributed. Let $\alpha \in \mathbb{R}^q$ any fixed vector. Then $\alpha^\top y = \tilde{\alpha}^\top x + \alpha^\top c$ where $\tilde{\alpha} = A^\top b$. Because x is multivariate Normal, then $\tilde{\alpha}^\top x$ is univariate Normal (by Definition 7), then $\alpha^\top y$ is univariate Normal. So y is q-variate Normal. Also, $\mathrm{E}(y) = \mathrm{E}(c + Ax) = c + A\mathrm{E}(x)$, and $\mathrm{Var}(y) = \mathrm{Var}(c + Ax) = A\mathrm{Var}(x)A^\top$. □

Proposition 12. *Let a d-dimensional random vector $x \sim N_{(any)}(\mu, \Sigma)$.*

1. *Let $y = Ax$ and $z = Bx$, where $A \in \mathbb{R}^{q \times d}$ and $B \in \mathbb{R}^{k \times d}$: The vectors $y = Ax$ and $z = Bx$ are independent if and only if $A\Sigma B^\top = 0$.*
2. *Let $x = (x_1, ..., x_d)^\top$: The $x_1, ..., x_d$ are mutually independent if and only if the corresponding off diagonal parts of the Σ are zero.*

Proof. In both cases, the CF (2) factorizes as $\varphi_x(t) = \prod_j \varphi_{x_j}(t_j)$ only when the corresponding of diagonal parts of Σ are zero. □

Proposition 13. *Any sub-vector of a vector with multivariate Normal distribution has a multivariate Normal distribution.*

Proof. Let $x \sim \mathrm{N}_d(\mu, \Sigma)$. Any sub-vector y of x can be expressed as $y = 0 + Px$, where $P \in \mathbb{R}^{q \times d}$ is a suitable projection matrix. Then $y \sim \mathrm{N}_d(P\mu, P\Sigma P^\top)$. □

Proposition 14. *[Marginalization & conditioning] Let $x \sim$* $\mathrm{N}_d(\mu, \Sigma)$. Consider partition such that

$$x = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}; \qquad \mu = \begin{bmatrix} \mu_1 \\ \mu_2 \end{bmatrix}; \qquad \Sigma = \begin{bmatrix} \Sigma_1 & \Sigma_{21}^\top \\ \Sigma_{21} & \Sigma_2 \end{bmatrix},$$

where $x_1 \in \mathbb{R}^{d_1}$, and $x_2 \in \mathbb{R}^{d_2}$ Then:

1. For the marginal, it is $x_1 \sim \mathrm{N}_{d_1}(\mu_1, \Sigma_1)$.

2. For $x_{2.1} = x_2 - \Sigma_{21}\Sigma_1^{-1}x_1$, with $\Sigma_1 > 0$, it is $x_{2.1} \sim \mathrm{N}_{d_2}(\mu_{2.1}, \Sigma_{2.1})$ where

$$\mu_{2.1} = \mu_2 - \Sigma_{21}\Sigma_1^{-1}\mu_1 \text{ and } \Sigma_{2.1} = \Sigma_2 - \Sigma_{21}\Sigma_1^{-1}\Sigma_{21}^\top \tag{3}$$

3. Random variables x_1 and $x_{2.1}$ are independent.

4. For the conditional, if $\Sigma_1 > 0$, it is

$$x_2|x_1 \sim \mathrm{N}_{d_2}(\mu_{2|1}, \Sigma_{2|1})$$

where

$$\mu_{2|1} = \mu_2 - \Sigma_{21}\Sigma_1^{-1}(x_1 - \mu_1) \text{ and } \Sigma_{2|1} = \Sigma_2 - \Sigma_{21}\Sigma_1^{-1}\Sigma_{21}^\top \tag{4}$$

Hint: *If that was a Homework it will be given as a hint to use , in (1.): $x_1 = Ax$ with $A = [I, 0]$, and in (2.): $x_{2.1} = Bx$ with $[-\Sigma_{21}\Sigma_1^{-1}, I]$.*

Solution.

1. It is $x_1 = Ax$ with $A = [I, 0]$. Then $x_1 \sim \mathrm{N}(A\mu, A\Sigma A^\top)$ where

$$A\mu = [I, 0]\begin{bmatrix}\mu_1\\ \mu_2\end{bmatrix} = \mu_1\ ; \qquad A\Sigma A^\top = [I, 0]\begin{bmatrix}\Sigma_1 & \Sigma_{21}^\top\\ \Sigma_{21} & \Sigma_2\end{bmatrix}\begin{bmatrix}I\\ 0\end{bmatrix} = \Sigma_1$$

2. It is $x_{2.1} = Bx$ with $B = [-\Sigma_{21}\Sigma_1^{-1}, I]$. Then $x_{2.1} \sim \mathrm{N}(B\mu, B\Sigma B^\top)$ where

$$\begin{aligned}
B\mu &= \left[-\Sigma_{21}\Sigma_1^{-1}, I\right]\left[\mu_1, \mu_2\right]^\top = -\Sigma_{21}\Sigma_1^{-1}\mu_1 + \mu_2;\\
B\Sigma B^\top &= \left[-\Sigma_{21}\Sigma_1^{-1}, I\right]\begin{bmatrix}\Sigma_1 & \Sigma_{21}^\top\\ \Sigma_{21} & \Sigma_2\end{bmatrix}\begin{bmatrix}-\Sigma_1^{-1}\Sigma_{21}^\top\\ I\end{bmatrix} = \left[0, -\Sigma_{21}\Sigma_1^{-1}\Sigma_{21}^\top + \Sigma_2\right]\begin{bmatrix}-\Sigma_{21}\Sigma_1^{-1}\\ I\end{bmatrix}\\
&= -\Sigma_{21}\Sigma_1^{-1}\Sigma_{21}^\top + \Sigma_2
\end{aligned}$$

3. x_1 and $x_{2.1}$ are independent, because (i.) x_1 and x_2 are Normally distributed and (ii.) for $x_1 = Ax$ with $A = [I, 0]$ and $x_{2.1} = Bx$ with $[\Sigma_{21}\Sigma_1^{-1}, 0]$ are

$$\mathrm{Cov}(x_1, x_{2.1}) = \mathrm{Cov}(Ax, Bx) = A\Sigma B^\top = \begin{bmatrix}I & 0\end{bmatrix}\begin{bmatrix}\Sigma_1 & \Sigma_{21}^\top\\ \Sigma_{21} & \Sigma_2\end{bmatrix}\begin{bmatrix}-\Sigma_1^{-1}\Sigma_{21}^\top\\ I\end{bmatrix} =$$

$$= \begin{bmatrix}\Sigma_1, & \Sigma_{21}^\top\end{bmatrix}\begin{bmatrix}-\Sigma_1^{-1}\Sigma_{21}^\top\\ I\end{bmatrix} = -\Sigma_{21}^\top + \Sigma_{21}^\top = 0$$

4. From part **??**, I observe that it is

$$x_{2.1} = x_2 - \Sigma_{21}\Sigma_1^{-1}x_1 \iff x_2 = x_{2.1} + \Sigma_{21}\Sigma_1^{-1}x_1. \tag{5}$$

In (5), if I condition x_2 on a given value for x_1, the term $\Sigma_{21}\Sigma_1^{-1}x_1$ is a constant, namely I have $x_2|x_1 = x_{2.1}|x_1 + \text{const}$ where variation comes from $x_{2.1}|x_1$. From part **??**, $x_{2.1}$ is independent on x_1, and hence $F(x_{2.1}|x_1) = F(x_{2.1})$. Furthermore, $x_{2.1}|x_1 \in$ Normal because $x_{2.1} \in$ Normal,. Consequently, $x_2|x_1 \in$ Normal as a linear transformation of a Normal variate. Now, about the moments

$$\mathrm{E}(x_2|x_1) = \mathrm{E}(x_{2.1} + \Sigma_{21}\Sigma_1^{-1}x_1|x_1) = \mathrm{E}(x_{2.1}|x_1) + \mathrm{E}(\Sigma_{21}\Sigma_1^{-1}x_1|x_1) = \left[\mu_2 - \Sigma_{21}\Sigma_1^{-1}\mu_1\right] + \left[\Sigma_{21}\Sigma_1^{-1}x_1\right]$$

$$\mathrm{Var}(x_2|x_1) = \mathrm{Var}(x_{2.1} + \Sigma_{21}\Sigma_1^{-1}x_1|x_1) = \mathrm{Var}(x_{2.1}|x_1) = \Sigma_2 - \Sigma_{21}\Sigma_1^{-1}\Sigma_{21}^\top$$

Proposition 15. *The density function of the d-dimensional Normal distribution with mean μ and covariance matrix Σ, when Σ is symmetric positive definite matrix ($\Sigma > 0$), exists and it is equal to* Supplementary rial A

$$f(x) = (2\pi)^{-\frac{d}{2}} \det(\Sigma)^{-\frac{1}{2}} \exp\left(-\frac{1}{2}(x-\mu)^\top \Sigma^{-1}(x-\mu)\right) \tag{6}$$

Proof. Let $x \sim \mathrm{N}(\mu, \Sigma)$. Because $\Sigma > 0$, we use Cholesky decomposition to define L such that $\Sigma = LL^\top$. Let $z = L^{^{-1}}(x-\mu)$. It is $\mathrm{E}(z) = 0$, $\mathrm{Var}(z) = I$, $z \sim \mathrm{N}_d(0, I)$, and hence $z_1, ..., z_d$ are mutually independent So

$$f_z(z) = \prod_{i=1}^{d} (2\pi)^{-\frac{1}{2}} \exp\left(-\frac{1}{2}z_i^2\right) = (2\pi)^{-\frac{d}{2}} \exp\left(-\frac{1}{2}z^\top z\right)$$

Then

$$\begin{aligned} f_x(x) &= f_z(z)\left|\frac{\mathrm{d}z}{\mathrm{d}x}\right| = f_z\left(L^{^{-1}}(x-\mu)\right)\left|\det\left(\frac{\mathrm{d}}{\mathrm{d}x}L^{^{-1}}(x-\mu)\right)\right| \\ &= (2\pi)^{-\frac{d}{2}} \exp\left(-\frac{1}{2}(x-\mu)^\top \left(L^{-1}\right)^\top L^{-1}(x-\mu)\right)\det(L^{-1}) \\ &= (2\pi)^{-\frac{d}{2}} \det(\Sigma)^{-\frac{1}{2}} \exp\left(-\frac{1}{2}(x-\mu)^\top \Sigma^{-1}(x-\mu)\right)\det(\Sigma)^{-\frac{1}{2}} \end{aligned}$$

□

4 Multivariate Student's T distribution $x \sim \mathbf{T}_d(\mu, \Sigma, v)$

Definition 16. A d-dimensional random variable $x \in \mathbb{R}^d$ is said to have a multivariate Student's T distribution with location parameter μ, scale matrix Σ, and degrees of freedom v, and it is denoted as $x \sim \mathrm{T}_d(\mu, \Sigma, v)$, if and only if

$$x = \mu + y\sqrt{v\xi}$$

where $y \sim \mathrm{N}_d(0, \Sigma)$ and $\xi \sim \mathrm{IG}(\frac{v}{2}, \frac{1}{2})$ are independent random variables.

Example 17. Definition 16 can be equivalently represented in terms of a mixture probability distribution (Section 1) as the compound (or mixture) distribution

$$x \sim \mathrm{T}_d(\mu, \Sigma, v)$$

of

$$\begin{aligned} x|\xi &\sim \mathrm{N}_d(\mu, \Sigma v\xi) \\ \xi &\sim \mathrm{IG}(\frac{v}{2}, \frac{1}{2}) \end{aligned}$$

Solution. Straightforward derivation from Definition 16. We have that $y \sim \mathrm{N}_d(0, \Sigma)$, $x = \mu + y\sqrt{v\xi}$, and consider the rest Σ, μ, v, and ξ as constant/known. So $x|\mu, \Sigma, v, \xi$ follows a Normal distribution with mean $\mathrm{E}(x|\xi) = \mu$ and covariance matrix $\mathrm{Var}(x|\xi) = \Sigma v\xi$. Hence, we can just write down $x|\xi \sim \mathrm{N}_d(\mu, \Sigma v\xi)$ by suppress the uninteresting constants Σ, μ, and v from the conditioning.

Proposition 18. *If $x \sim T_d(\mu, \Sigma, v)$ and $\Sigma > 0$ then*

1. *The expected value is*

$$E_{T_d(\mu,\Sigma,\nu)}(X) = \mu$$

2. *The covariance matrix is*

$$Var_{T_d(\mu,\Sigma,\nu)}(X) = \begin{cases} \frac{\nu}{\nu-2}\Sigma & , \textit{if } \nu > 2 \\ \textit{undefined} & , \textit{else} \end{cases} \tag{7}$$

Proof. Given Definition 16, $x \sim \mathrm{T}_d(\mu, \Sigma, v)$ results as the marginal distribution of (x, ξ) where $x|\xi \sim \mathrm{N}_d(\mu, \Sigma\xi v)$ and $\xi \sim \mathrm{IG}(\frac{v}{2}, \frac{1}{2})$.

1. It is

$$\mathrm{E}_{\mathrm{t}_d(\mu,\Sigma,\nu)}(x) = \mathrm{E}_{\mathrm{IG}(\frac{v}{2},\frac{1}{2})}\left(\mathrm{E}_{\mathrm{N}_d(\mu,\Sigma\xi v)}(x|\xi)\right) = \mathrm{E}_{\mathrm{IG}(\frac{v}{2},\frac{1}{2})}(\mu) = \mu \tag{8}$$

2. It is

$$\begin{aligned} \mathrm{Var}_{\mathrm{t}_d(\mu,\Sigma,\nu)}(x) =& \mathrm{E}_{\mathrm{IG}(\frac{v}{2},\frac{1}{2})}\left(\mathrm{Var}_{\mathrm{N}_d(\mu,\Sigma\xi v)}(x|\xi)\right) + \mathrm{Var}_{\mathrm{IG}(\frac{v}{2},\frac{1}{2})}\left(\mathrm{E}_{\mathrm{N}_d(\mu,\Sigma\xi v)}(x|\xi)\right) \\ \\ =& \mathrm{E}_{\mathrm{IG}(\frac{v}{2},\frac{1}{2})}(\Sigma\xi v) + \cancel{\mathrm{Var}_{\mathrm{IG}(\frac{v}{2},\frac{1}{2})}(\mu)}^{=0} = \Sigma v \mathrm{E}_{\mathrm{IG}(\frac{v}{2},\frac{1}{2})}(\xi) + 0 \\ \\ =& \begin{cases} \Sigma v \frac{\frac{1}{2}}{\frac{v}{2}-1} & , \text{if } \frac{v}{2} > 1 \\ \text{undefined} & , \text{else} \end{cases} \end{aligned} \tag{9}$$

$\square$

Example 19. If $x \sim \mathrm{T}_d(\mu, \Sigma, v)$ and $\Sigma > 0$ the PDF of x is

$$f_X(x|\mu,\Sigma,v) = \frac{\Gamma(\frac{\nu+d}{2})}{\Gamma(\frac{\nu}{2})\nu^{\frac{d}{2}}\pi^{\frac{d}{2}}\det(\Sigma)^{\frac{1}{2}}}\left(1 + \frac{1}{v}(t-\mu)^{\mathrm{T}}\Sigma^{-1}(t-\mu)\right)^{-\frac{\nu+d}{2}} \tag{10}$$

Solution. Given Definition 16, $x \sim \mathrm{T}_d(\mu, \Sigma, v)$ results as the marginal distribution of (x, ξ) where $x|\xi \sim \mathrm{N}_d(\mu, \Sigma\xi v)$ and $\xi \sim \mathrm{IG}(\frac{v}{2}, \frac{1}{2})$. So it is

$$\begin{aligned} f_x(x) &= \int f_{x|\xi}(x|\xi)f_\xi(\xi)\mathrm{d}\xi = \int f_{\mathrm{N}_d(\mu,\Sigma v\xi)}(x|\xi)f_{\mathrm{IG}(\frac{v}{2},\frac{1}{2})}(\xi)\mathrm{d}\xi \\ &= \int \underbrace{\left(\frac{1}{2\pi}\right)^{\frac{d}{2}}\frac{1}{\sqrt{\det(\Sigma v\xi)}}\exp\left(-\frac{1}{2}(x-\mu)^\top\frac{\Sigma^{-1}}{v\xi}(x-\mu)\right)}_{=\mathrm{N}_d(x|\mu,\Sigma v\xi)}\underbrace{\frac{\frac{1}{2}^{\frac{v}{2}}}{\Gamma(\frac{v}{2})}\xi^{-\frac{v}{2}-1}\exp\left(-\frac{1}{\xi}\frac{1}{2}\right)1_{(0,\infty)}(\xi)}_{=\mathrm{IG}(\xi|\frac{v}{2},\frac{1}{2})}\mathrm{d}\xi \\ &= \left(\frac{1}{2\pi}\right)^{\frac{d}{2}}\frac{1}{\sqrt{\det(\Sigma v)}}\frac{\frac{1}{2}^{\frac{v}{2}}}{\Gamma(\frac{v}{2})}\underbrace{\int \xi^{-\frac{v}{2}-\frac{d}{2}-1}\exp\left(-\frac{1}{\xi}\left[\frac{1}{2v}(x-\mu)^\top\Sigma^{-1}(x-\mu)+\frac{1}{2}\right]\right)\mathrm{d}\xi}_{=\Gamma\left(\frac{v}{2}+\frac{d}{2}\right)\left[\frac{1}{2v}(x-\mu)^\top\Sigma^{-1}(x-\mu)+\frac{1}{2}\right]^{-\left(\frac{v}{2}+\frac{d}{2}\right)}} \\ &= \left(\frac{1}{2\pi}\right)^{\frac{d}{2}}\frac{1}{\sqrt{\det(\Sigma v)}}\frac{\frac{1}{2}^{\frac{v}{2}}}{\Gamma(\frac{v}{2})}\Gamma\left(\frac{v}{2}+\frac{d}{2}\right)\left[\frac{1}{2v}(x-\mu)^\top\Sigma^{-1}(x-\mu)+\frac{1}{2}\right]^{-\left(\frac{v}{2}+\frac{d}{2}\right)} \\ &= \left(\frac{1}{2\pi}\right)^{\frac{d}{2}}\frac{1}{\sqrt{\det(\Sigma v)}}\frac{\frac{1}{2}^{\frac{v}{2}}}{\Gamma(\frac{v}{2})}\Gamma\left(\frac{v}{2}+\frac{d}{2}\right)\left(\frac{1}{2}\right)^{-\frac{(v+d)}{2}}\left[\frac{1}{v}(x-\mu)^\top\Sigma^{-1}(x-\mu)+1\right]^{-\frac{v+d}{2}} \\ &= \left(\frac{1}{\pi}\right)^{\frac{d}{2}}\frac{1}{\sqrt{\det(\Sigma)}}\left(\frac{1}{v}\right)^{\frac{d}{2}}\frac{1}{\Gamma(\frac{v}{2})}\Gamma\left(\frac{v+d}{2}\right)\left[\frac{1}{v}(x-\mu)^\top\Sigma^{-1}(x-\mu)+1\right]^{-\frac{v+d}{2}} \end{aligned} \tag{11}$$

where the integral in (11) was calculated by recognizing the IG density from (1).

5 Practice

Question 20. *For practice try the Exercises 18, 19, and, 21, from the Exercise Sheet.*

Supplementary material

A The useful formulas regarding Normal PDF

The following formulas will be given as hints in the exercises and there is no need to be memorized. I present them for your information –no need to memorize.

Fact 21. *If* $\Sigma_1 > 0$ *and* $\Sigma_2 > 0$ *symmetric*

$$-\frac{1}{2}\left(x-\mu_1\right)\Sigma_1^{-1}\left(x-\mu_1\right)^\top-\frac{1}{2}\left(x-\mu_2\right)\Sigma_2^{-1}\left(x-\mu_2\right)^\top=-\frac{1}{2}\left(x-m\right)V^{-1}\left(x-m\right)^\top+C$$

where

$$V^{-1}=\Sigma_1^{-1}+\Sigma_2^{-1};\quad m=V\left(\Sigma_1^{-1}\mu_1+\Sigma_2^{-1}\mu_2\right);\quad C=\frac{1}{2}m^\top V^{-1}m-\frac{1}{2}\left(\mu_1^\top\Sigma_1^{-1}\mu_1+\mu_2^\top\Sigma_2^{-1}\mu_2\right)$$

Proof. It is derived by $\pm$ing terms and doing matrix calculations. □

Fact 22. *If* $f_{N_d(\mu,\Sigma)}(x)$ *denotes the PDF of* $N_d(\mu,\Sigma)$*, then*

$$f_{N_d(\mu_1,\Sigma_1)}(x)\,f_{N_d(\mu_2,\Sigma_2)}(x)\;=\;f_{N_d(m,V)}(x)\,f_{N_d(\mu_2,\Sigma_1+\Sigma_2)}(\mu_1)$$

where

$$V^{-1}=\Sigma_1^{-1}+\Sigma_2^{-1};\quad m=V\left(\Sigma_1^{-1}\mu_1+\Sigma_2^{-1}\mu_2\right)$$

Proof. It is derived by exponentiation □

Fact 23. *If* $\Sigma_i > 0$ *symmetric for* $i = 1, ..., n$

$$-\frac{1}{2}\sum_{i=1}^{n}\left(x-\mu_i\right)\Sigma_i^{-1}\left(x-\mu_i\right)^\top=-\frac{1}{2}\left(x-m\right)V^{-1}\left(x-m\right)^\top+C \tag{12}$$

where

$$V^{-1}=\sum_{i=1}^{n}\Sigma_i^{-1};\quad m=V\left(\sum_{i=1}^{n}\Sigma_i^{-1}\mu_i\right);\quad C=\frac{1}{2}mV^{-1}m^\top-\frac{1}{2}\left(\sum_{i=1}^{n}\mu_i\Sigma_i^{-1}\mu_i^\top\right) \tag{13}$$

Proof. It is shown by induction from the (1.). □

B Examples of mixture distributions

The following is distributions below are given as examples of mixture distributions –no need to memorize.

B.1 Fisher-Snedecor distribution $x \sim \mathbf{F}(d_1, d_2)$

Definition 24. A random variable $x \in \mathbb{R}$ is said to have a Fisher-Snedecor distribution distribution with degrees of freedom n, m, and it is denoted as $x \sim \mathrm{F}(n, m)$, if and only if

$$x = \frac{y_1/d_1}{y_2/d_2}$$

where $y_1 \sim \mathrm{G}(d_1/2, 1/2)$ and $y_2 \sim \mathrm{G}(d_2/2, 1/2)$ are independent random variables.

Example 25. The distribution function $x \sim \mathrm{L}(\mu, b)$ is such that

$$\mathrm{d}F(x) = \left[x\mathrm{B}\left(\frac{d_1}{2}, \frac{d_2}{2}\right)\right]^{-1} \sqrt{\frac{(d_1 x)^{d_1} d_2^{d_2}}{(d_1 x + d_2)^{d_1 + d_2}}}\mathrm{d}x$$

with mean $\mathrm{E}(x) = d_2/(d_2 - 1)$.

Solution. See Statistical Concepts II (Exercises) which is proved by transformation of random variables and marginalization.

Example 26. The distribution function $x \sim \mathrm{F}(d_1, d_2)$ can be re-presented as a mixture distribution as

$$\begin{cases} x|v \sim \mathrm{Ga}(d_1/2, zd_1/2) \\ z \sim \mathrm{Ga}(d_2/2, d_2/2) \end{cases}$$

Solution. Show that $f(x) = \int f(x|z)f(z)\mathrm{d}z$; see Exercise 19 in the Exercise sheet.

B.2 Laplace distribution $x \sim \mathbf{L}(\mu, b)$

Definition 27. A random variable $x \in \mathbb{R}$ is said to have a Laplace (or Double exponential) distribution with location parameter μ, and scale parameter b, and it is denoted as $x \sim \mathrm{L}_d(\mu, b)$, if and only if

$$x = \mu + y\sqrt{2v}$$

where $y \sim \mathrm{N}_d(0, b^2)$ and $v \sim \mathrm{G}(1, 1)$ are independent random variables.

Example 28. The distribution function $x \sim \mathrm{L}(\mu, b)$ can be re-presented as a mixture distribution as

$$\begin{cases} x|v \sim \mathrm{N}(\mu, 2vb^2) \\ v \sim \mathrm{Ga}(1, 1) \equiv \mathrm{Ex}(1) \end{cases} \tag{14}$$

Show that the distribution function $x \sim \mathrm{L}(\mu, b)$ is such that

$$\mathrm{d}F(x) = \frac{1}{2b}\exp\left(-\frac{|x - \mu|}{b}\right)\mathrm{d}x$$

with mean $\mathrm{E}(x) = \mu$ and variance $\mathrm{Var}(x) = 2b^2$.

Solution. By transformation of random variables you show (14). To show that $f(x) = \int f(x|v)f(v)\mathrm{d}v$. Perform change of variable to compute the integral. See C. Robert's solution: `https://stats.stackexchange.com/questions/175458/show-that-a-scale-mixtures-of-normals-is-a-power-exponential`