Bayesian Statistics III/IV (MATH3341/4031)

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# **Handout 7: Mixture priors**

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Aim: Explain the mixture distribution. Explain, theorize, and construct conjugate mixture prior distribution.

#### **References:**

- Berger, J. O. (2013; Sections 4.2.2). Statistical decision theory and Bayesian analysis. Springer Science & Business Media.
- Robert, C. (2007; Sections 3, pp. 105-123, & pp. 127-141). The Bayesian choice: from decision-theoretic foundations to computational implementation. Springer Science & Business Media.

Web applets: https://georgios-stats-1.shinyapps.io/demo\_mixturepriors/

# 1 Finite mixture distributions

**Definition 1.** Let  $\mathcal{P}_m = \{\Pi_l(\theta|\chi_l); l = 1,...,m\}$  be a collection of probability distributions where  $\{\chi_l\}_{l=1}^m$  are parameters of the l-th component  $\Pi_l(\theta|\chi_l)$ . Let  $\{\varpi_l\}_{l=1}^m$  be a set of weights where  $\varpi_l > 0$  and  $\sum_{l=1}^m \varpi_l = 1$ . The mixture distribution derived from the aforementioned collections is

$$\Pi(\theta|\varpi,\chi) = \sum_{l=1}^{m} \varpi_{l} \Pi_{l}(\theta|\chi_{l}), \qquad \theta \in \Theta$$
(1)

where  $\chi := (\chi_l, l = 1 : m)$  and  $\varpi := (\varpi_l, l = 1, ..., m)$ .  $\Pi_l(\theta|\chi_l)$  is called l-th mixture component with mixture weight  $\varpi_l$ .

**Example 2.** Let  $h(\cdot)$  be a function defined on  $\Theta$ . The expectation of  $h(\cdot)$  with respect to (1) is

$$\mathbf{E}_{\Pi}(h(\theta)|\varpi,\chi) = \int \sum_{l=1}^{m} \varpi_{l}h(\theta) \mathrm{d}\Pi_{l}(\theta|\chi_{l}) = \sum_{l=1}^{m} \varpi_{l} \int h(\theta) \mathrm{d}\Pi_{l}(\theta|\chi_{l}) = \sum_{l=1}^{m} \varpi_{l}\mathbf{E}_{\Pi_{l}}(h(\theta)|\chi_{l})$$

where  $E_{\Pi_l}(h(\theta)|\chi_l) = \int h(\theta) d\Pi_l(\theta|\chi_l)$ .

**Example 3.** A mixture of probability distributions (1) is a probability distribution; i.e.

$$\int_{\Theta} \pi(\theta | \varpi, \chi) d\theta = \sum_{l=1}^{m} \varpi_{l} \underbrace{\int_{\Theta} d\pi_{l}(\theta | \chi_{l}) d\theta}_{=1} = \sum_{l=1}^{m} \varpi_{l} = 1$$

**Definition 4.** The mixture is called finite mixture if  $m < \infty$ , and countably infinite mixture if  $m \to \infty$ . Here, we focus on finite mixtures.

**Definition 5.** A mixture model is called is called parametric mixture model if its components are members of the same parametric family of distributions eg.  $\mathcal{P}_m = \{\Pi(\theta|\chi_l); l = 1, ..., m\}$ , and hence

$$\Pi(\theta|\varpi,\chi) = \sum_{l=1}^{m} \varpi_l \Pi(\theta|\chi_l)$$

**Example 6.** An example of a parametric mixture model is the Normal mixture model

$$\pi(\theta|\varpi,\mu,\sigma^2) = \sum_{l=1}^{m} \varpi_l N(\theta|\mu_l,\sigma_l^2),$$

where all components belong to the Normal distribution family.

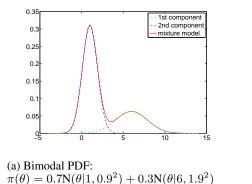
Note 7. Mixture distributions are useful because (among others):

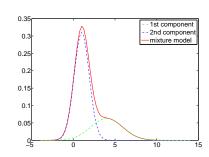
- they can approximate complicate other distributions by using a combination of simpler distributions  $\{\Pi_l(\theta|\chi_l)\}$  which lead to more convenient computations.
- they can naturally model heterogeneity. E.g. consider a population which is heterogeneous in the sense that there are multiple sub-groups labeled by  $\ell \in \{1,...,m\}$ , each group represented in the population with proportion  $\varpi_{\ell}$ , and distributed as  $\Pi_{l}(\theta|\chi_{l})$ . Then  $y \sim \Pi(\theta|\chi)$  can be realized by drawing  $\ell$  with probability  $P(\ell = l) = \varpi_{l}$  and drawing  $\theta$  from  $\Pi_{\ell}(\theta|\chi_{\ell})$  given  $\ell$ , aka

$$\begin{cases} \theta|\ell & \sim \Pi_{\ell}(\theta|\chi_{\ell}) \\ \ell & \sim \mathsf{P}(\ell) \end{cases} \quad \text{which implies } \theta \sim \Pi(\theta|\chi) \quad \text{since} \quad \Pi(\theta|\chi) = \sum_{\ell=1}^{m} \Pi(\theta|\chi_{\ell}) \mathsf{P}(\ell) = \sum_{l=1}^{m} \varpi_{l} \Pi(\theta|\chi_{l})$$

See Section 1 in Handout 2

**Example 8.** A bimodal, or a right skewed (non-symmetric) distribution can be approximated by a Mixture of (unimodal and symmetric) Normal distributions with different parameter values. Also it describes a population with two groups Normally distributed with different parameters.





(b) Right skewed PDF:  $\pi(\theta) = 0.7 \text{N}(\theta|1,0.9^2) + 0.3 \text{N}(\theta|4,1.9^2)$ 

Figure 1: Normal mixture models.  $N(x|\mu, \sigma^2)$  denotes the Normal distribution density at value x with mean  $\mu$  and variance  $\sigma^2$ .

# 2 Mixture prior distributions

*Note* 9. Mixture models can be used to specify priors distributions, either as a mean to approximate Your actual prior distribution with simpler & tractable distributions, or as a mean to represent heterogeneous prior believes.

**Theorem 10.** Let  $y := (y_1, ..., y_n)$  be observables generated from the sampling distribution  $F(y|\theta)$ . Prior mixture distribution  $\Pi(\theta|\varpi)$  is called the prior with pdf/pmf

$$\pi(\theta|\varpi) = \sum_{l=1}^{m} \varpi_l \pi_l(\theta), \tag{2}$$

where,  $\mathcal{P}_m = \{\pi_l(\theta), \theta \in \Theta\}_{l=1}^m$  is a collection of distributions, and  $\{\varpi_l\}$  are weights such that  $\sum_{l=1}^m \varpi_l = 1$  and  $\varpi_l > 0$ . Then:

1. the posterior distribution  $\Pi(\theta|y,\varpi)$  has pdf/pmf

$$\pi(\theta|y,\varpi) = \sum_{l=1}^{m} \varpi_l^* \pi_l(\theta|y), \qquad (3)$$

2. the predictive distribution of a future outcome z has pdf/pmf

$$g(z|y,\varpi) = \sum_{l=1}^{m} \varpi_l^* g_l(z|y)$$

where

$$\begin{split} \varpi_l^* &= \frac{\varpi_l f_l(y)}{\sum_{l=1}^m \varpi_l f_l(y)} \propto \varpi_l f_l(y) \\ f_l(y) &= \int_{\Theta} f(y|\theta) \mathrm{d}\Pi_l(\theta) \\ \pi_l(\theta|y) &= \frac{f(y|\theta)\pi_l(\theta)}{\int_{\Theta} f(y|\theta) \mathrm{d}\Pi_l(\theta)} \propto f(y|\theta)\pi_l(\theta) \\ g_l(z|y) &= \int_{\Theta} f(z|\theta)\pi_l(\theta|y) \mathrm{d}\theta \end{split}$$

*Proof.* From Bayes theorem, we have:

$$\begin{split} \mathrm{d}\Pi(\theta|y) &= \frac{f(y|\theta)\mathrm{d}\Pi(\theta)}{\int_{\Theta} f(y|\vartheta)\mathrm{d}\Pi(\theta)} = \frac{f(y|\theta)\sum_{l=1}^{m} \varpi_{l}\mathrm{d}\Pi_{l}(\theta)}{\int_{\Theta} f(y|\vartheta)\sum_{l'=1}^{m} \varpi_{l'}\mathrm{d}\Pi_{l'}(\vartheta)} = \frac{\sum_{l=1}^{m} \varpi_{l}f(y|\theta)\mathrm{d}\Pi_{l}(\theta)}{\sum_{l'=1}^{m} \int_{\Theta} \varpi_{l'}f(y|\vartheta)\mathrm{d}\Pi_{l'}(\vartheta)} \\ &= \frac{\sum_{l=1}^{m} \varpi_{l}f_{l}(y)}{\underbrace{\int_{l'=1}^{m} \int_{\Theta} \varpi_{l'}f_{l}'(y)}\underbrace{\underbrace{\int_{l'=1}^{m} \int_{\theta} \varpi_{l'}f_{l}'(y)}_{=\mathrm{d}\Pi_{l'}(\vartheta|y)}} = \frac{\sum_{l=1}^{m} \varpi_{l}f_{l}(y)\mathrm{d}\Pi_{l}(\theta|y)}{\sum_{l'=1}^{m} \varpi_{l'}f_{l'}(y)\underbrace{\underbrace{\int_{\theta} \mathrm{d}\Pi_{l'}(\vartheta|y)}_{=\mathrm{d}\Pi_{l'}(\vartheta|y)}}_{=\mathrm{d}\Pi_{l'}(\vartheta|y)} \\ &= \sum_{l=1}^{m} \underbrace{\underbrace{\sum_{l'=1}^{m} \varpi_{l'}f_{l}'(y)}_{=\mathrm{d}\Pi_{l'}(\vartheta|y)}}_{=\mathrm{d}\Pi_{l}(\theta|y)} = \sum_{l=1}^{m} \varpi_{l}^{*}\mathrm{d}\Pi_{l}(\theta|y). \end{split}$$

Also

$$g(z|y,\varpi) = \int_{\Theta} f(z|\theta) \mathrm{d}\Pi_l(\theta|y) = \int_{\Theta} f(y|\theta) \sum_{l=1}^m \varpi_l^* \mathrm{d}\Pi_l(\theta|y) = \sum_{l=1}^m \varpi_l^* g_l(z|y)$$

Remark 11. Theorem 10 shows that the posterior distribution  $\Pi(\theta|y)$  (derived by a mixture prior) is a mixture of 'individual posterior distributions'  $\Pi_l(\theta|y)$  weighted by  $\varpi_l^*$ . It is determined not only by the observables but also by the weights of the individual distributions. Note that, for l=1,...,m, the prior weights  $\varpi_l$  and posterior weights  $\varpi_l^*$  my differ a lot, however they can be close each other.

Remark 12. Mixture priors  $\Pi(\theta)$  whose components  $\{\Pi_l(\theta)\}$  are conjugate to the likelihood  $f(y|\theta)$  can facilitate tractable Bayesian inference. In Theorem 10, if each  $\Pi_l(\theta)$  is conjugate to the likelihood  $f(y|\theta)$ , then obviously each  $\Pi_l(\theta|y)$  will belong to the same distribution family as the corresponding  $\Pi_l(\theta)$  because  $\pi_l(\theta|y) \propto f(y|\theta)\pi_l(\theta)$ . Then, provided that components  $\pi_l(\theta)$  are tractable, the components  $\pi_l(\theta|y)$  will be tractable too (as they belong to the same distr. family). Likewise, the posterior weights can be calculated in closed form i.e.,  $\varpi_l^* \propto \varpi_l f_l(y)$  since the integral  $f_l(y) = \int_{\Theta} f(y|\theta) d\Pi_l(\theta)$  will be tractable. Hence, the produced posterior mixture pdf/pmf will be tractable.

*Remark* 13. Mixtures of conjugate priors are as easy to manipulate as regular conjugate distributions, while leading to a greater freedom in the modeling of the prior information.

**Example 14.** Let  $y = (y_1, ..., y_n)$  observable quantities, generated iid from a Bernoulli sampling distribution with unknown parameter  $\theta$ ; aka  $y_i | \theta \stackrel{\text{iid}}{\sim} \text{Br}(\theta), \ i = 1, ..., n.$ 

- 1. Find the likelihood function
- 2. Find the PDF of the conjugate prior mixture prior  $\pi(\theta) = \sum_{l=1}^{m} \varpi_l \pi_l(\theta)$ , with m components  $\{\pi_l(\theta)\}_{l=1}^m$ .
- 3. Compute the pdf of the posterior distribution, and recognize it.
- 4. Compute the predictive pdf for the next outcome  $z = (y_{n+1}, ..., y_{n+m})$  given we have observed y? What do you observe?

**Hint:** Beta distribution:  $x \sim \text{Be}(a, b)$  has pdf

$$f(x) = \frac{1}{B(a,b)} x^{a-1} (1-x)^{b-1} 1(x \in [0,1]); \quad \text{where} \qquad B(a,b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)}, a > 0, b > 0$$

# Solution.

1. The likelihood function is

$$f(y|\theta) = \prod_{i=1}^{n} f(y_i|\theta) = \prod_{i=1}^{n} Br(x_i|\theta) = \theta^{\sum_{i=1}^{n} x_i} (1-\theta)^{n-\sum_{i=1}^{n} y_i}$$

- 2. In previous examples, we found that
  - the sampling distribution  $F(\cdot|\theta)$  is a Bernoulli distribution which is a member of the Exponential family.
  - the conjugate prior is Beta distribution, aka  $\theta \sim \text{Be}(a, b)$  a > 0 and b > 0.

Therefore, the components  $\{\pi_l(\theta)\}_{l=1}^m$  in the prior mixture distribution will be from the family of Beta distributions  $\mathcal{P}_m = \{\text{Be}(\theta|a_l,b_l); l=1,...,m\}$ .

Therefore, the conjugate mixture prior has pdf

$$\pi(\theta) = \sum_{l=1}^{m} \varpi_{l} \operatorname{Be}(\theta|a_{l}, b_{l})$$

where  $Be(\theta|a_l, b_l)$  is the pdf of  $Be(a_l, b_l)$ , and  $\{(a_l, b_l)\}$  are fixed prior hyper-parameters.

## 3. The posterior mixture posterior has PDF

$$\pi(\theta|y) = \sum_{l=1}^{m} \varpi_l^* \pi_l(\theta|y)$$

According to Theorem 10, the l-th components of the mixture posterior is

$$\pi_l(\theta|y) \propto f(y|\theta)\pi_l(\theta) = \prod_{i=1}^n \operatorname{Br}(y_i|\theta) \times \operatorname{Be}(\theta|a_l,b_l) \propto \prod_{i=1}^n [\theta^{y_i}(1-\theta)^{1-y_i}] \times \theta^{a_l-1}(1-\theta)^{b_l-1}$$

$$r_n = \sum_{i=1}^n y_i \theta^{r_n+a_l-1}(1-\theta)^{n-r_n+b_l-1} \propto \operatorname{Be}(\theta|a_l^*,b_l^*)$$

with 
$$a_l^* = r_n + a_l$$
,  $b_l^* = n - r_n + b_l$ , and  $r_n = \sum_{i=1}^n y_i$ .

According to Theorem 10, the posterior weights can be calculated as

$$\varpi_{l}^{*} \propto \varpi_{l} f_{l}(y) \propto \varpi_{l} \int_{(0,\infty)} \prod_{i=1}^{n} f(x_{i}|\theta) \pi_{l}(\theta) d\theta = \varpi_{l} \int \prod_{i=1}^{n} \operatorname{Br}(y_{i}|\theta) \operatorname{Be}(\theta|a_{l}, b_{l}) d\theta 
= \varpi_{l} \int_{(0,\infty)} \theta^{r_{n}} (1-\theta)^{n-r_{n}} \frac{\Gamma(a_{l}+b_{l})}{\Gamma(a_{l})\Gamma(b_{l})} \theta^{a_{l}-1} (1-\theta)^{b_{l}-1} d\theta 
= \varpi_{l} \frac{\Gamma(a_{l}+b_{l})}{\Gamma(a_{l})\Gamma(b_{l})} \int_{(0,1)} \theta^{r_{n}+a_{l}-1} (1-\theta)^{n-r_{n}+b_{l}-1} d\theta = \varpi_{l} \frac{\Gamma(a_{l}+b_{l})}{\Gamma(a_{l})\Gamma(b_{l})} \frac{\Gamma(a_{l}^{*})\Gamma(b_{l}^{*})}{\Gamma(a_{l}^{*}+b_{l}^{*})} \tag{4}$$

namely,

$$\varpi_l^* = \frac{\varpi_l \frac{\Gamma(a_l + b_l)}{\Gamma(a_l)\Gamma(b_l)} \frac{\Gamma(a_l^*)\Gamma(b_l^*)}{\Gamma(a_l^* + b_l^*)}}{\sum_{l=1}^m \varpi_l \frac{\Gamma(a_l + b_l)}{\Gamma(a_l)\Gamma(b_l)} \frac{\Gamma(a_l^*)\Gamma(b_l^*)}{\Gamma(a_l^* + b_l^*)}}$$

4. It is ...

$$g(z|y) = \int_{\Theta} f(z|\theta) \mathrm{d}\Pi(\theta|y) = \sum_{i=1}^{m} \varpi_{l}^{*}(y) \underbrace{\int_{\Theta} f(z|\theta) \mathrm{d}\Pi_{l}(\theta|y)}_{=g_{l}(z|y)}$$

where the following is just copy-paste from Handout 3...

$$\begin{split} g_l(z|y) &= \int_{\Theta} f(z|\theta) \pi_l(\theta|y) \mathrm{d}\theta \ = \int_{\Theta} \prod_{i=1}^m f(z_i|\theta) \pi_l(\theta|y) \mathrm{d}\theta \ = \int_{(0,\infty)} \prod_{i=1}^m \mathrm{Br}(z_i|\theta) \mathrm{Be}(\theta|a_l^*,b_l^*) \mathrm{d}\theta \\ &= \int_0^1 \left[ \theta^{\sum_{i=1}^m z_i} (1-\theta)^{m-\sum_{i=1}^m z_i} \right] \left[ \frac{\theta^{a_l^*-1} (1-\theta)^{b_l^*-1}}{B(a_l^*,b_l^*)} \right] \mathrm{d}\theta \ 1(z \in \{0,1\}^m) \\ &= \frac{1}{B(a_l^*,b_l^*)} \int_0^1 \theta^{\sum_{i=1}^m z_i + a_l^*-1} (1-\theta)^{m-\sum_{i=1}^m z_i + b_l^*-1} \mathrm{d}\theta \ 1(z \in \{0,1\}^m) \\ &= \frac{B(\sum_{i=1}^m z_i + a_l^*, m - \sum_{i=1}^m z_i + b_l^*)}{B(a_l^*,b_l^*)} 1(z \in \{0,1\}^m) \end{split}$$

## 3 Practice

**Question 15.** Try the Exercise 51 from the Exercise sheet.