



Inequality in Health

Lecture VII: Policy Interventions Affecting Early Life Health

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1 Introduction

2 Empirical Study: Infant Care in Sweden

- Background
- Data and Methods
- Results: Mortality
- Results: Education and Labour Market Outcomes
- Summary

3 Summary and Conclusions

How to Reach Us

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- If they did, later-life effects would be of **secondary importance**.
- What evidence do we have to guide **positive policies**?

Examples of Historical Interventions

- Federal tax reforms (1986-93) in the US (Hoynes et al., 2015):
 - Reduction in low birth weight incidence; increase in mean birth weight
- Environmental regulations (1998) in China (Tanaka, 2015):
 - Infant mortality rate decrease of 20%
- Health care reform (2001) in Thailand (Gruber et al., 2014):
 - Reduction in infant mortality inequalities between provinces
- Introduction of compulsory health insurance (1884) by Bismarck in the German Empire (Bauernschuster et al., 2017):
 - Significant reduction in child mortality

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- In utero exposure to shocks have adult health impacts (see last lecture).
- Therefore also interventions during early childhood potentially have long-term health effects.

Long-Term Trends

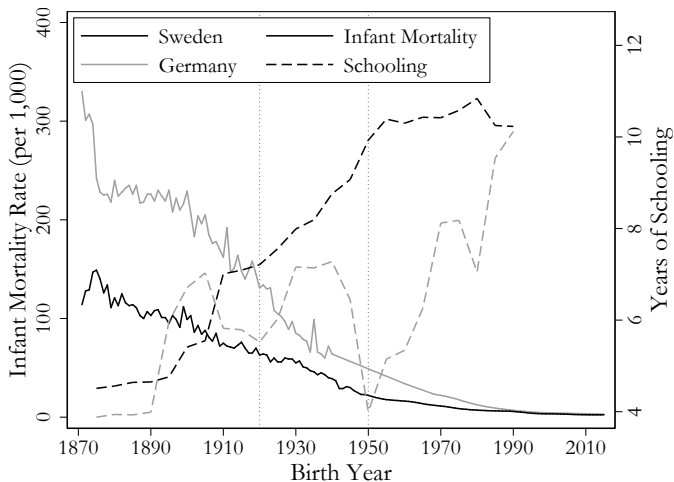


Figure 1. Long-Term Trends in Human Capital and Health, Sweden and Germany.

Empirical Study: Infant Care in Sweden

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 - Districts quasi “randomly” chosen to reflect diversity in local conditions.

Activities

- **Three main activities:**

- ① Examination of babies at centers \Rightarrow follow-up if problems identified.
- ② Home visits to provide support and monitoring.
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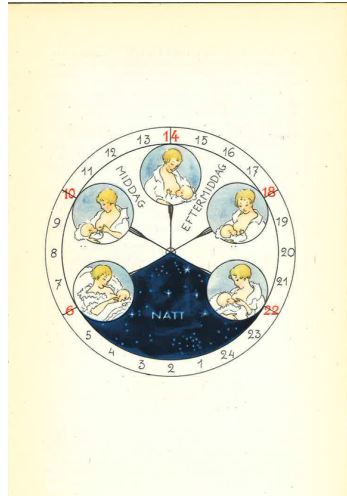
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- Norway (Bütikofer et al., 2018) and Denmark (Hjort et al., 2017; Wüst, 2012) rolled out similar programs from 1936 and 1937.

Activities



Figure 2. Advice on appropriate feeding of infants from leaflet provided within the infant care intervention.



Identification – Aided by Programme Features

- Districts selected “randomly” to be representative of Sweden.

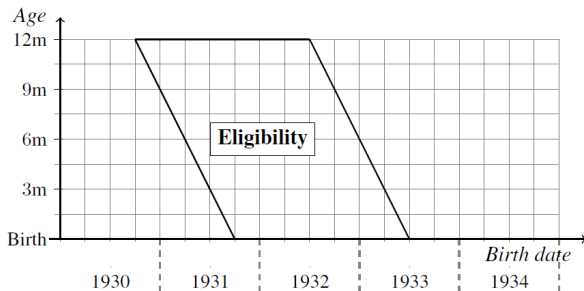


Figure 3. Eligibility by birth date.

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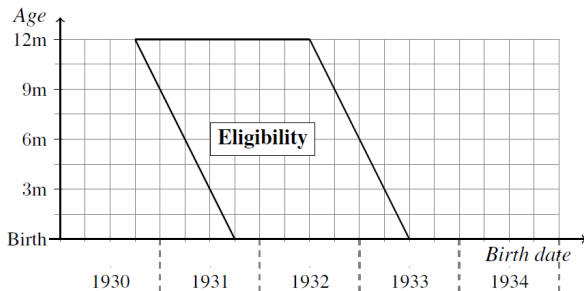


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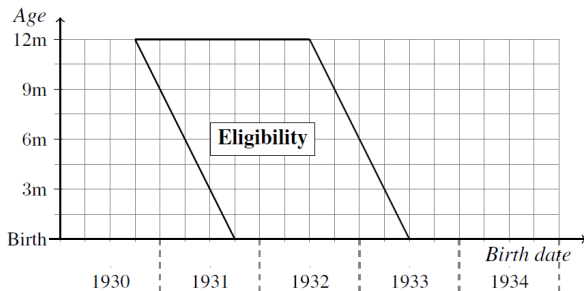


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- Narrow window of eligibility limits confounding unobserved trends.

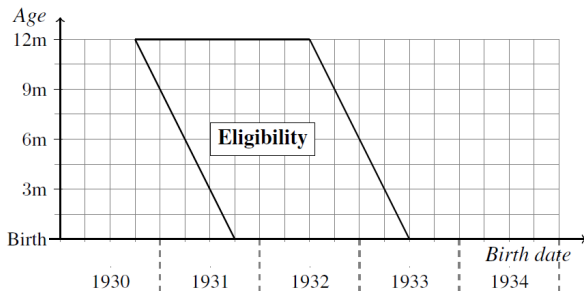


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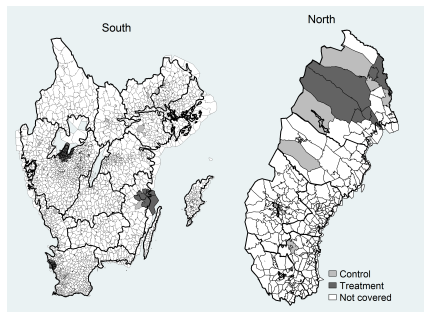


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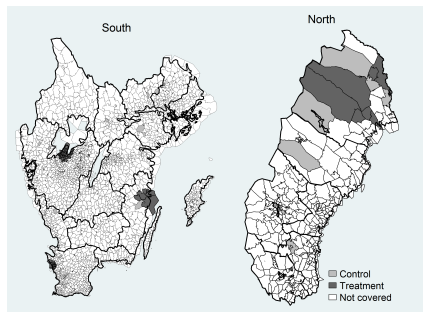


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- Identify 2 control cities and 57 control parishes using Mahalanobis matching estimator.

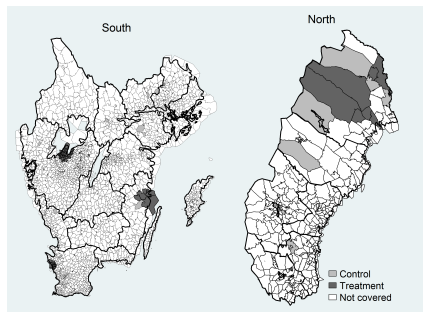


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Research Questions

- Determining the impact of the intervention **over the life course**.
- Main outcomes:
 - **Mortality** until ages 1, 5, 40, and 75.
 - **Performance in school** and **education**.
 - **Labour market** performance and **life-cycle earnings**.
- Understanding the process **connecting** early-life health with adult outcomes.

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- **4 subjects**: math, writing, reading and speaking, religion.

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- Outcome variables:
 - Mortality by different ages.
 - Cause of death

Empirical Strategy: Main Specification

- Difference-in-differences approach:

$$y_{icj} = \alpha + \beta T_c + \gamma D_j + \delta D_j T_c + \kappa_c + \varepsilon_{icj}$$

y_{icj} survival outcome of child i born on date c in parish j

T_c (theoretical) duration in months of eligibility if born on day c

D_j treatment status of parish j

κ quarter of birth times year of birth fixed effects

- δ estimates the intent-to-treat (ITT) effect, i.e., the effect – for each additional month of eligibility – of making the service available.

Empirical Strategy: Additional Specification

A richer specification may account for **diverging trends**, e.g.

- Changes in the composition of births possibly related to the intervention
- Regional variation over time due to other interventions.

$$y_{icj} = \alpha + \beta T_c + \gamma_j + \tau_j c + \delta D_j T_c + \lambda X_i + \kappa_c + \varepsilon_{icj}$$

γ_j parish-fixed effects

$\tau_j c$ parish-specific linear trends

X_i individual covariates

Results: Mortality

Table 1. Infant and future survival chances (basic specification). Source: Bhalotra et al. (2017)

	d_{0-1}	d_{0-5}	d_{0-40}	d_{0-75}
ITT	-0.1414** (0.061)	-0.0872 (0.057)	-0.1295* (0.076)	-0.2972** (0.143)
AITT	-1.0889	-0.6716	-0.9979	-2.2891
Pre-Mean	6.617	8.257	11.221	36.535

Standard errors clustered at the parish level. d_{0-x} denotes mortality before age x . AITT is the intent-to-treat effect for the average eligible individual (i.e., the product of the DID point estimates and the average eligibility period conditional on enrollment). Pre-Mean represents the mortality rate for children born before the start of the eligibility period (starting 2 October 1930). * $p < 0.1$, ** $p < 0.05$, *** $p < 0.01$.

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- Compare **actual estimates** to this distribution.

Randomisation Inference: Results

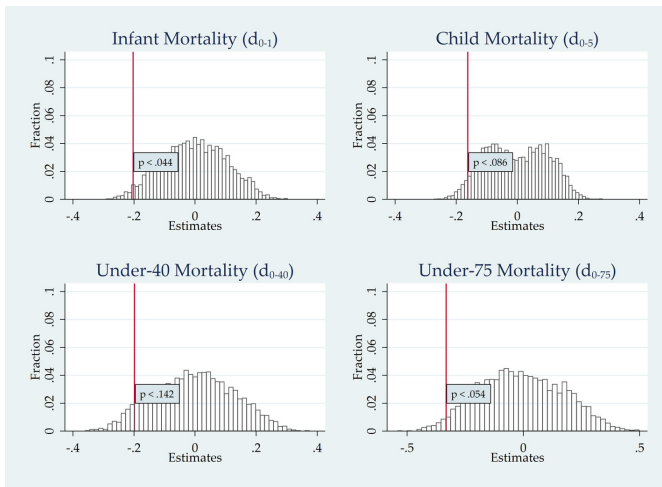


Figure 5. Randomisation inference based on 5,000 permutations of treatment status. Source: Bhalotra et al. (2017).

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- Children of **young mothers** benefited more; however, this effect is not persistent beyond the age of 5.
- There is **no** significant difference in effects by **child sex**.

Adult Death Causes

Table 2. Adult results by death cause. Source: Bhalotra et al. (2017)

	All-cause	Infect	External	Cancer	Cardio
Mortality between ages 50 and 75					
ITT	-0.3527*** (0.134)	-0.0433* (0.022)	-0.0339 (0.044)	-0.2835** (0.134)	-0.0958 (0.091)
AITT	-2.7165	-0.3337	0.2608	-2.1842	-0.7381
Pre-Mean	26.249	0.210	1.579	7.373	7.778
Average age at death	65.008	68.392	60.636	64.911	64.951

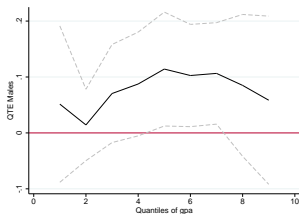
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Education and Labour Market Outcomes

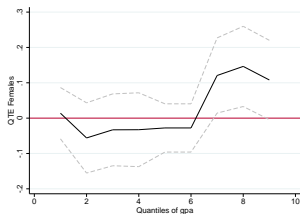
Table 3. DID Estimates: Education and Earnings.

	Women				Men			
	N	Mean	(1)	(2)	N	Mean	(3)	(4)
Top GPA	6,465	0.227	0.1000* (0.058)	0.1243* (0.070)	6,607	0.116 (0.033)	0.0400 (0.028)	0.0275
GPA	6,465	0.098	0.0410 (0.048)	0.0617 (0.053)	6,607	-0.200 (0.056)	0.1213** (0.070)	0.1084
Secondary	8,071	0.198	0.0353** (0.016)	0.0350** (0.014)	8,301	0.172 (0.029)	-0.0468 (0.021)	-0.0289
Top Income 1970	10,301	0.244	0.0655** (0.026)	0.0788** (0.031)	10,619	0.210 (0.034)	-0.0445 (0.029)	-0.0361
Log Income	10,301	8.990	0.1204* (0.068)	0.1947** (0.074)	10,619	10.222 (0.036)	-0.0596 (0.033)	-0.0464
Log Pensions (age 71)	8,284	11.609	0.0293 (0.021)	0.0711*** (0.018)	7,680	11.995 (0.017)	-0.0400** (0.022)	-0.0400*
Parish FE			✓	✓			✓	✓
QOB×YOB FE			✓	✓			✓	✓
SES Effects			✓	✓			✓	✓
School Reforms			✓	✓			✓	✓
Parish Trends				✓				✓

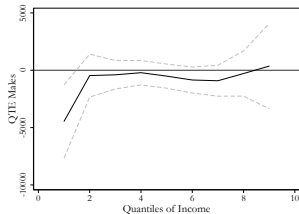
Results: Grade Point Average



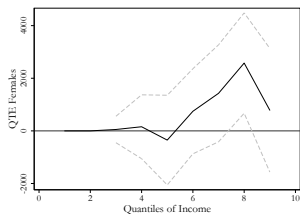
(a) QTE GPA Males



(b) QTE GPA Females



(c) QTE Income Males



(d) QTE Income Females

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$$Y = T\tau_{base} + X\lambda + \epsilon$$
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where Z are potential mediators.

- The difference $\hat{\delta} = \hat{\tau}_{base} - \hat{\tau}_{full}$ captures the impact of the mediators on the estimated effect.
- One variable's contribution: $\hat{\delta}_k = \hat{\Gamma}_k \hat{\beta}_k$.

Results: Female Schooling

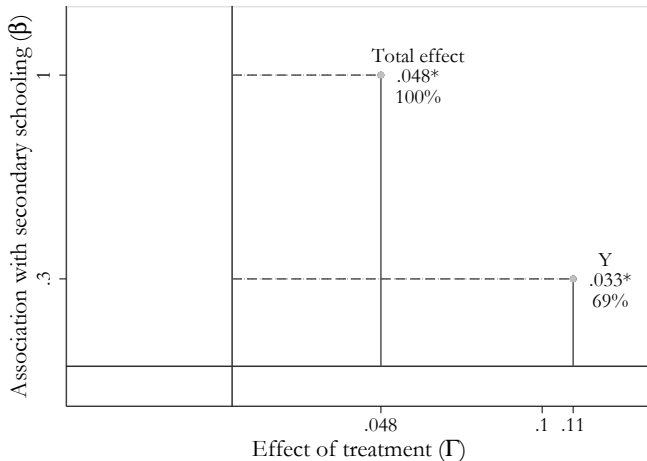


Figure 7. Mediation Analysis: Secondary Schooling.

Results: Female Earnings

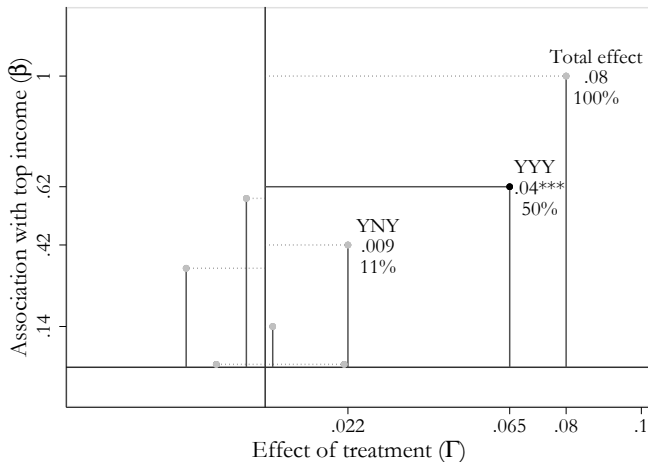


Figure 8. Mediation Analysis: Labour Market Earnings.

Results: Female Occupation

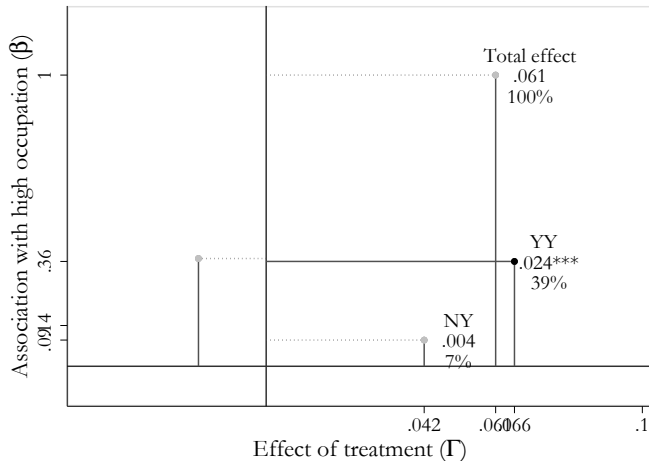


Figure 9. Mediation Analysis: High-Ranking Occupation.

Why Females?

- Utilisation?
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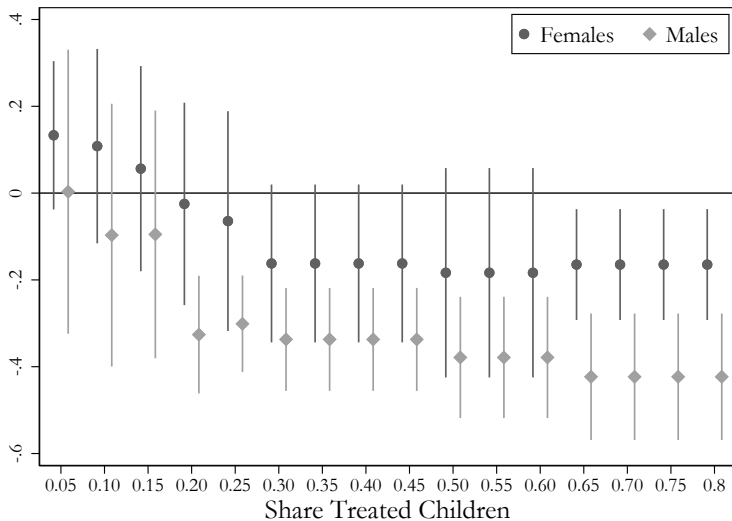
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- Labour demand?
 - Labour demand for qualified females **expanded rapidly** (welfare state) but not for males.

Skill Acquisition: Secondary Schooling



Growth in Labor Market Opportunities

Table 4. Treatment Effect Heterogeneity by Bartik Instrument for Skilled Workers, Adult Index

	Females (N=10,301)		Males (N=10,619)	
	(1)	(2)	(3)	(4)
Treated × Duration Eligibility	0.0732*** (0.022)	0.0758*** (0.021)	-0.0129 (0.017)	-0.0126 (0.017)
Treated × Own Skilled Bartik	0.0072 (0.052)	0.0035 (0.052)	0.0373** (0.018)	0.0373** (0.018)
Own Skilled Bartik	0.0404 (0.040)	0.0447 (0.039)	-0.0006 (0.008)	-0.0002 (0.008)
Duration Eligibility × Own Skilled Bartik	-0.0311** (0.014)	-0.0313** (0.014)	-0.0211* (0.013)	-0.0218* (0.013)
Treated × Duration Eligibility × Own Skilled Bartik	0.0577*** (0.018)	0.0583*** (0.018)	0.0148 (0.018)	0.0173 (0.019)
Treated × Other Skilled Bartik		0.0231 (0.017)		0.0318 (0.041)
Other Skilled Bartik		-0.0319*** (0.010)		-0.0145 (0.022)
Duration Eligibility × Other Skilled Bartik		0.0160 (0.014)		0.0006 (0.009)
Treated × Duration Eligibility × Other Skilled Bartik		-0.0100 (0.019)		-0.0007 (0.012)
Parish FE	✓	✓	✓	✓
QOB×YOB FE	✓	✓	✓	✓
SES Effects	✓	✓	✓	✓
School Reforms	✓	✓	✓	✓
Parish Trends	✓	✓	✓	✓

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- The intervention is associated with significant **reductions in mortality** for all age thresholds.
- The effect does not **fade over time** but is rather persistent
- ⇒ Consistent with infancy being a critical stage of development and with programme-driven learning within mothers that persists over time.
- Moderate gains in school performance translate into large gains in earnings for females.

Further Studies

- Several studies evaluate similar interventions in Denmark and Norway.
- Hjort et al. (2017):
 - The authors estimate the effects of a Danish home visiting programme in 1937.
 - They find higher survival rates during ages 45 to 64, fewer hospital nights, as well as a reduction in cardiovascular disease diagnoses.
- Bütikofer et al. (2015):
 - This study evaluates an intervention in Norway from 1936 to 1955.
 - The results indicate positive effects on education and earnings of affected children.
 - Further, a reduction in health risks at age 40 is found.

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- Several studies confirm shocks during childhood to have old-age impacts on health.
- Relatively low-cost home visiting programmes in Scandinavia during the 1930s were found to not only decrease infant mortality but also reduce deaths at old-age due to cardiovascular diseases or cancer.
- The conclusions of these studies might be important for today's developing countries.