

Incorporating structural priors in Gaussian Random Field models

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Acknowledgements: a number of co-authors, notably appearing via citations!

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Outline

1 Introduction: Background and motivations

2 Covariance kernels and invariances

- Kernels invariant under a combination of compositions
- Further operators in the Gaussian case. Applications.

3 On ANOVA decompositions of kernels and GRF paths

- Subjective state of the art
- Some recent contributions

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Black box functions

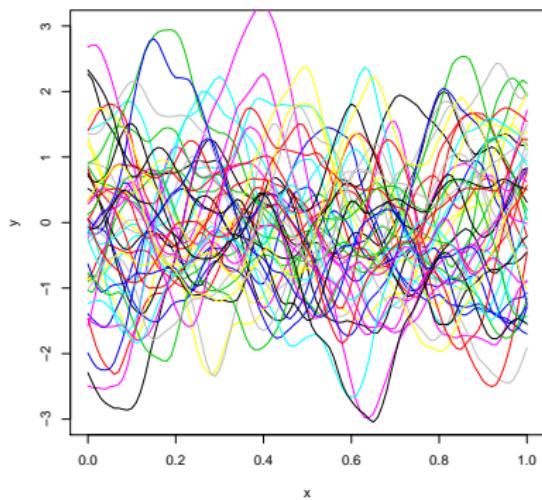
Here we mainly focus on cases where a system of interest can be modelled as (or involves) a costly-to-evaluate deterministic function:

$$f : \mathbf{x} \in D \subset E \longrightarrow f(\mathbf{x}) \in F$$

for some given *input space* E and *output space* F –often $E \subset \mathbb{R}^d$ and $F \subset \mathbb{R}$.

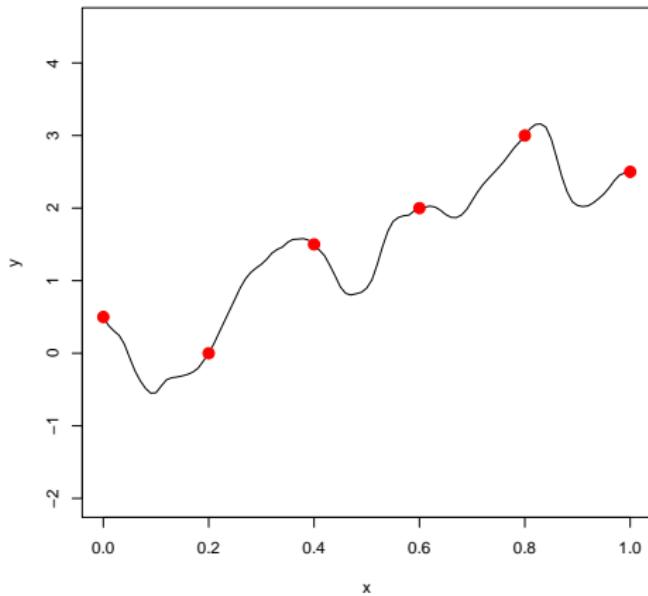
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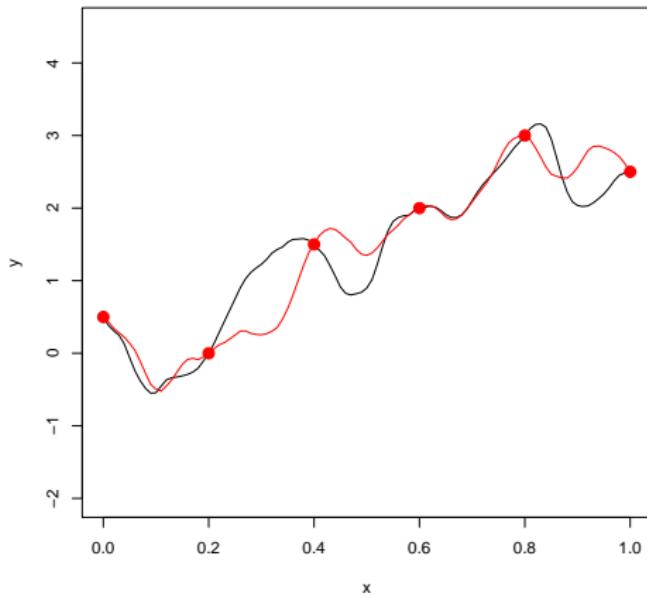


Actually, they can serve as prior distribution on function spaces!

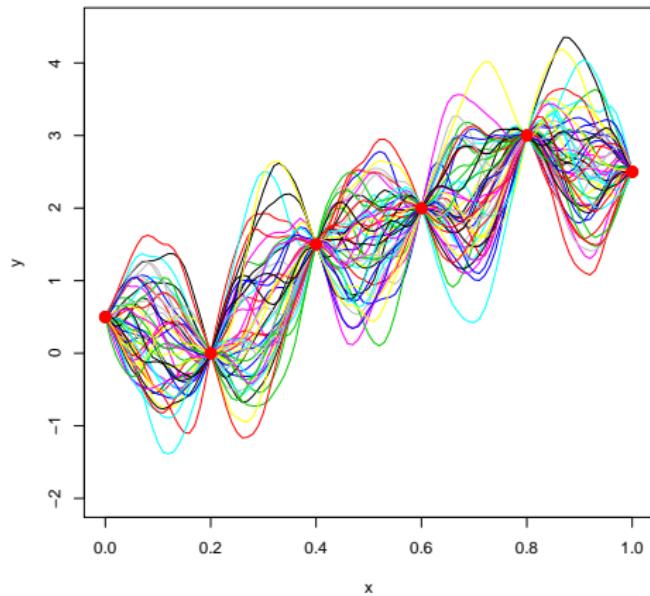
Conditional simulations (1D)



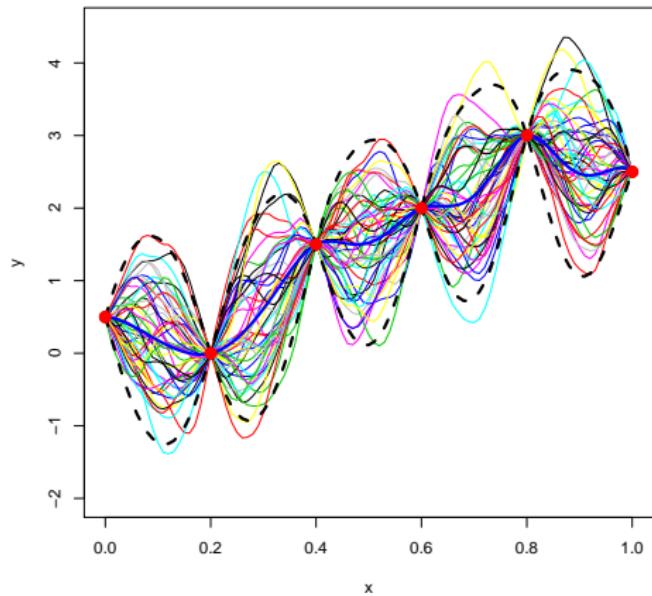
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Conditional simulations and Kriging (1D)



Kriging at a glance: from geostats to machine learning

Originally, Kriging refers to "optimal" linear prediction of a random field $(Z(\mathbf{x}))_{\mathbf{x} \in D}$ ($D \subset \mathbb{R}^2$ or \mathbb{R}^3) based on observations at $\mathbf{X}_n := \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$, i.e.

$$\mathcal{A}_n := \{(Z(\mathbf{x}_1), \dots, Z(\mathbf{x}_n)) = \mathbf{z}_n\}$$

where $\mathbf{z}_n = (z(\mathbf{x}_1), \dots, z(\mathbf{x}_n))$ with $z(\cdot) = Z(\cdot; \omega)$ for some $\omega \in \Omega$.

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Kriging may be cast as an ancestor/a particular or more general case of various contemporary methods from different fields, including

- Gaussian Process Regression
- Interpolation Splines
- Kernel methods and regularization in RKHS

A few references about those facets



C. E. Rasmussen and C. K. I. Williams (2006)

Gaussian Processes for Machine Learning

The MIT Press



G. Wahba (1990)

Spline Models for Observational Data

CBMS-NSF Regional Conference Series in Applied Mathematics



M.L. Stein (1999)

Interpolation of Spatial Data - Some Theory for Kriging

Springer-Verlag New York

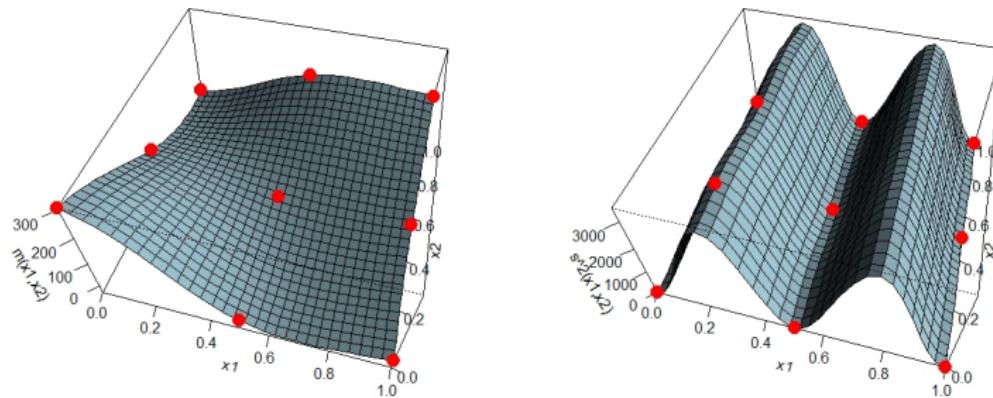


A. Berlinet, C. Thomas-Agnan (2004)

Reproducing Kernel Hilbert Spaces in Probability and Statistics

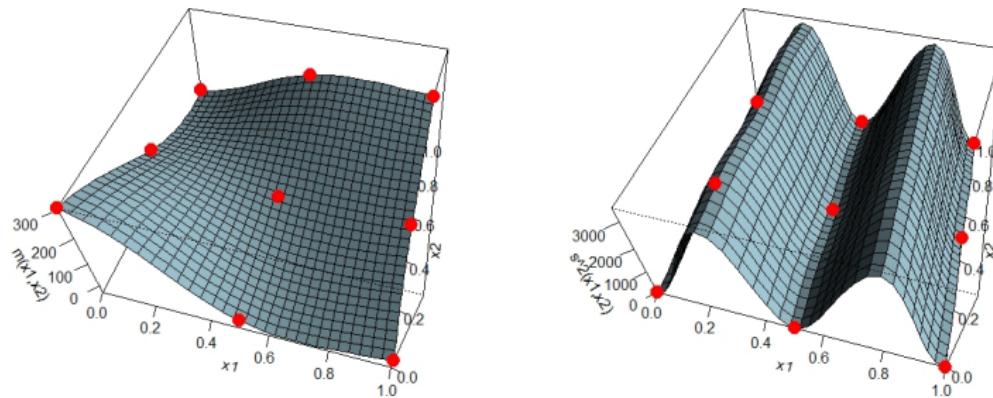
Kluwer Academic Publishers

Interpolating deterministic functions by Kriging



Prediction by Kriging (based on 9 points) of the Branin-Hoo function.

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The covariance is here a **stationary** anisotropic Matérn kernel ($\nu = 5/2$) with scale and range parameters estimated by Maximum Likelihood.

Ordinary Kriging Equations –for completeness!–

Assume Z has a covariance kernel k , and constant mean $\mu \in \mathbb{R}$

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$$\mathbf{K}_n = \begin{pmatrix} k(\mathbf{x}_1, \mathbf{x}_1) & k(\mathbf{x}_1, \mathbf{x}_2) & \dots & k(\mathbf{x}_1, \mathbf{x}_n) \\ k(\mathbf{x}_2, \mathbf{x}_1) & k(\mathbf{x}_2, \mathbf{x}_2) & \dots & k(\mathbf{x}_2, \mathbf{x}_n) \\ \dots & \dots & \dots & \dots \\ k(\mathbf{x}_n, \mathbf{x}_1) & \dots & \dots & k(\mathbf{x}_n, \mathbf{x}_n) \end{pmatrix}, \mathbf{k}_n(\mathbf{x}) = \begin{pmatrix} k(\mathbf{x}, \mathbf{x}_1) \\ k(\mathbf{x}, \mathbf{x}_2) \\ \dots \\ k(\mathbf{x}, \mathbf{x}_n) \end{pmatrix}, \hat{\mu}_n = \frac{\mathbf{1}_n^T \mathbf{K}_n^{-1} \mathbf{z}_n}{(\mathbf{1}_n^T \mathbf{K}_n^{-1} \mathbf{1}_n)}.$$

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If μ is known (or with improper uniform prior) and $Z - \mu$ is assumed Gaussian, then m_n and k_n are Z 's conditional mean and covariance and

$$\mathcal{L}(Z|A_n) = \mathcal{GRF}(m_n(\cdot), k_n(\cdot, \cdot))$$

More on the Bayesian approach: selected references

-  H. Omre and K. Halvorsen (1989).
The bayesian bridge between simple and universal kriging.
Mathematical Geology, 22 (7):767-786.
-  M. S. Handcock and M. L. Stein (1993).
A bayesian analysis of kriging.
Technometrics, 35(4):403-410.
-  A. O'Hagan (2006)
Bayesian analysis of computer code outputs: a tutorial.
Reliability Engineering and System Safety, 91:1290-1300.
-  A.W. Van der Vaart and J. H. Van Zanten (2008)
Rates of contraction of posterior distributions based on Gaussian process priors.
Annals of Statistics, 36:1435-1463.

Conditional simulations of the Branin-Hoo function

In second-order random field models with constant mean, prior assumptions on f are implicitly accounted for through the choice of the covariance

$$k : (\mathbf{x}, \mathbf{x}') \in D \times D \longrightarrow k(\mathbf{x}, \mathbf{x}') = \text{cov}(Z_{\mathbf{x}}, Z_{\mathbf{x}'}) \in \mathbb{R}$$

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Classical invariance notions for k

- 2nd order stationarity (k invariant wrt simult. translations of \mathbf{x} and \mathbf{x}')
- Isotropy (k invariant wrt simultaneous rigid motions of \mathbf{x} and \mathbf{x}').

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The main focus here is on functional properties of random field paths driven by k , both in Gaussian and in more general second-order settings.

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Property

Let G be a finite group acting measurably on D via

$$\Phi : (\mathbf{x}, g) \in D \times G \longrightarrow \Phi(\mathbf{x}, g) = g.\mathbf{x} \in D$$

and Z be a second-order random field indexed by D with constant mean.

$$(\forall \mathbf{x} \in D, \mathbb{P}(\forall g \in G, Z_{\mathbf{x}} = Z_{g.\mathbf{x}}) = 1) \Leftrightarrow (\forall \mathbf{x} \in D, \forall g \in G, k(g.\mathbf{x}, \cdot) = k(\mathbf{x}, \cdot))$$

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D. G., X. Bay, O. Roustant and L. Carraro

Argumentwise invariant kernels for the approximation of invariant functions

Annales de la faculté des sciences de Toulouse. Mathématiques (2012)

Invariant kernels enable invariant simulations

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Remark: See, e.g., help files of the *kergp* R package for more examples.

Another invariance: random fields with additive paths

Let $D = \prod_i^d D_i$ where $D_i \subset \mathbb{R}$. $f \in \mathbb{R}^D$ is called **additive** when there exists $f_i \in \mathbb{R}^{D_i}$ ($1 \leq i \leq d$) such that $f(\mathbf{x}) = \sum_{i=1}^d f_i(x_i)$ ($\mathbf{x} = (x_1, \dots, x_d) \in D$).

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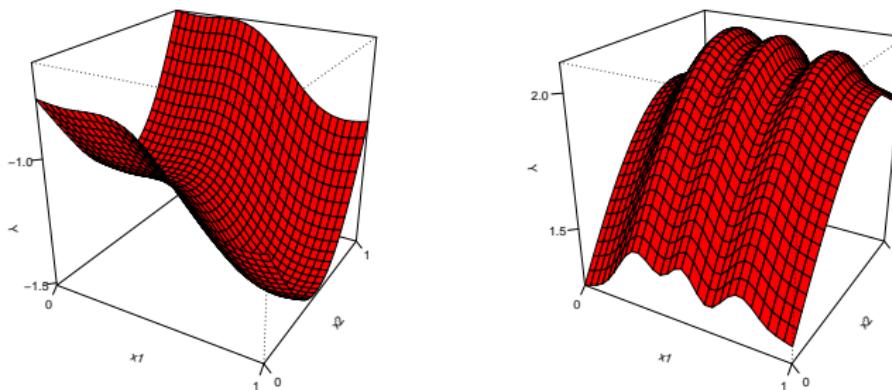
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GRF models possessing additive paths (with $k(\mathbf{x}, \mathbf{x}') = \sum_{i=1}^d k_i(x_i, x'_i)$) have been considered in Nicolas Durrande's Ph.D. thesis (2011):

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A few selected/related references



N. Durrande

Étude de classes de noyaux adaptés à la simplification et à l'interprétation des modèles d'approximation. Une approche fonctionnelle et probabiliste

PhD thesis, Ecole des Mines de Saint-Etienne (2011)



D. Duvenaud, H. Nickisch, C. E. Rasmussen

Additive Gaussian Processes

NIPS (2011)



N. Durrande, D. G., and O. Roustant

Additive Covariance kernels for high-dimensional Gaussian Process modeling

Annales de la faculté des sciences de Toulouse. Mathématiques (2012)



D. G., N. Durrande and O. Roustant

Kernels and designs for modelling invariant functions: From group invariance to additivity.

In mODa 10 - Advances in Model-Oriented Design and Analysis. Contributions to Statistics (2013)

Pathwise properties of fields with invariant kernels

Definition: Composition operator

Let us consider a (non-necessarily bi/in/sur-jective) function

$$\nu : \mathbf{x} \in D \longrightarrow \nu(\mathbf{x}) \in D.$$

$$T_\nu : f \in \mathbb{R}^D \longrightarrow T_\nu(f) = f \circ \nu \in \mathbb{R}^D$$

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Property

Let Z be a centred second-order RF with covariance kernel k and T be a finite linear combination of composition operators. Then **k is T -invariant**, i.e.

$$T(k(., \mathbf{x}')) = k(., \mathbf{x}') \quad (\mathbf{x}' \in D)$$

If and only if $\mathbb{P}(Z_{\mathbf{x}} = T(Z)_{\mathbf{x}}) = 1 \quad (\mathbf{x} \in D)$.

Particular case of additivity

One can show that f is additive $\iff f$ is invariant under

$$T(f)(\mathbf{x}) = \sum_{i=1}^d f(\mathbf{v}_i(\mathbf{x})) - (d-1)f(\mathbf{a})$$

where $\mathbf{a} \in D$ is arbitrary and $\mathbf{v}_i(\mathbf{x}) = (a_1, \dots, a_{i-1}, \underbrace{x_i}_{\text{i-th coordinate}}, a_{i+1}, \dots, a_d)$

This leads to Z additive **if and only if** k (is positive definite and) writes

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Particular case of group invariance

$$T(f)(\mathbf{x}) = \sum_{i=1}^{\#G} \frac{1}{\#G} f(\mathbf{v}_i(\mathbf{x})) \text{ with } \mathbf{v}_i(\mathbf{x}) := g_i \cdot \mathbf{x}$$

leads to Z **Φ -invariant if and only if k is argumentwise invariant.**

Extension to further operators in the Gaussian case

In the Gaussian case, the last results can be extended to a wider class of operators using the Loève isometry Ψ between $\mathcal{L}(Z)$ (The Hilbert space generated by Z) and the RKHS associated with k , $\mathcal{H}(k)$.

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$T(Z)_x \in \mathcal{L}(Z)$ ($x \in D$). T induces an operator \mathcal{T} from $\mathcal{H}(k)$ to \mathbb{R}^D , defined by

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D. G., O. Roustant and N. Durrande

Invariances of random fields paths, with applications in Gaussian Process Regression

(<http://arxiv.org/abs/1308.1359> ; a restructured version is currently in revision)

Examples (Gaussian case)

a) Let ν be a measure on D s.t. $\int_D \sqrt{k(\mathbf{u}, \mathbf{u})} d\nu(\mathbf{u}) < +\infty$.
 Then Z has centred paths iff $\int_D k(\mathbf{x}, \mathbf{u}) d\nu(\mathbf{u}) = 0, \forall \mathbf{x} \in D$.

For instance, given any p.d. kernel k , k_0 defined by

$$k_0(\mathbf{x}, \mathbf{y}) = k(\mathbf{x}, \mathbf{y}) - \int k(\mathbf{x}, \mathbf{u}) d\nu(\mathbf{u}) - \int k(\mathbf{y}, \mathbf{u}) d\nu(\mathbf{u}) + \int k(\mathbf{u}, \mathbf{v}) d\nu(\mathbf{u}) d\nu(\mathbf{v})$$

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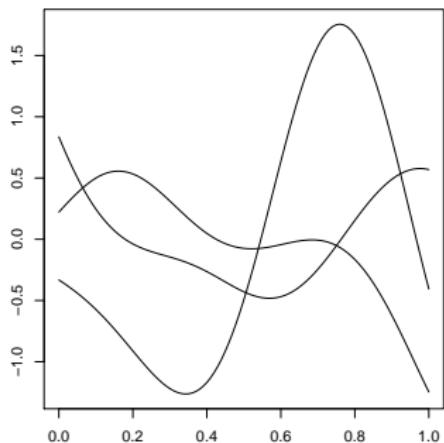
satisfies the above condition.

b) Solutions to the *Laplace equation* are called harmonic functions. Let us call harmonic any p.d. kernel solving the Laplace equation argumentwise:
 $(\Delta k(\cdot, \mathbf{x}')) = 0 (\mathbf{x}' \in D)$.

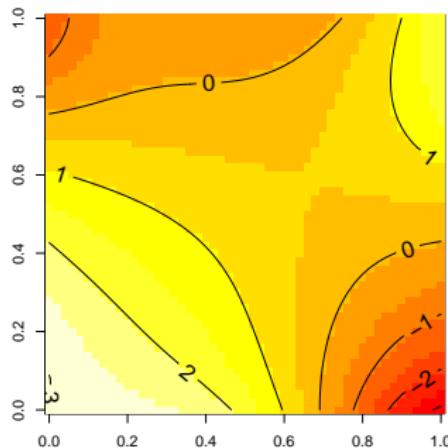
An example of such harmonic kernel over $\mathbb{R}^2 \times \mathbb{R}^2$ can be found in the recent literature (Schaback et al. 2009):

$$k_{harm}(\mathbf{x}, \mathbf{y}) = \exp\left(\frac{x_1 y_1 + x_2 y_2}{\theta^2}\right) \cos\left(\frac{x_2 y_1 - x_1 y_2}{\theta^2}\right).$$

Example sample paths invariant under various T 's



(a) Zero-mean paths of the centred GP with kernel k_0 .



(b) Harmonic path of a GRF with kernel k_{harm} .

Kriging with invariant kernels: example a)

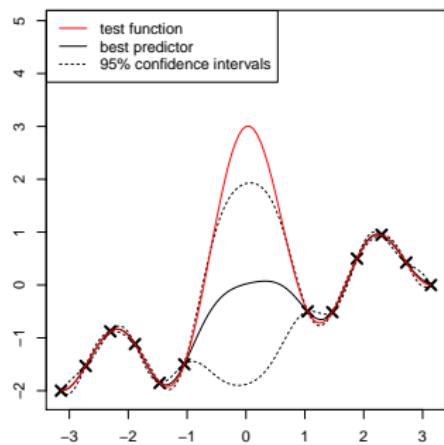
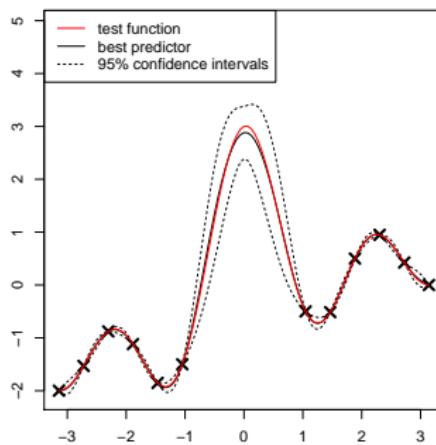
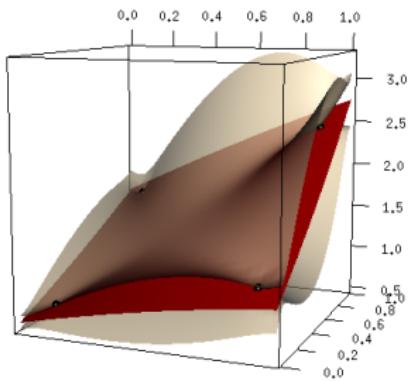
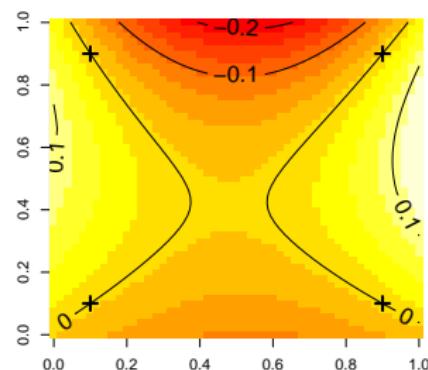
(c) GPR with kernel k (d) GPR with kernel k_0

Figure: Comparison of two kriging models. The left one is based on a Gaussian kernel. The right one incorporates the zero-mean property.

Kriging with invariant kernels: example b)



(a) Mean predictor and 95%
confidence intervals



(b) prediction error

Figure: Example of kriging model based on a harmonic kernel.

Outline

1 Introduction: Background and motivations

2 Covariance kernels and invariances

- Kernels invariant under a combination of compositions
- Further operators in the Gaussian case. Applications.

3 On ANOVA decompositions of kernels and GRF paths

- Subjective state of the art
- Some recent contributions

Set up: Functional ANOVA decomposition

Specific assumptions on D and f : $D = D_1 \times \cdots \times D_d$, where each $D_i \subset \mathbb{R}$ is endowed with a probability measure μ_i , D is endowed with the product measure $\mu := \mu_1 \times \cdots \times \mu_d$, and $f \in L^2(\mu)$.

The **Functional ANOVA** (or *Sobol'-Hoeffding*) decomposition consists in expanding f into a sum of orthogonal terms of increasing complexity:

$$f = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} f_{\mathbf{u}}$$

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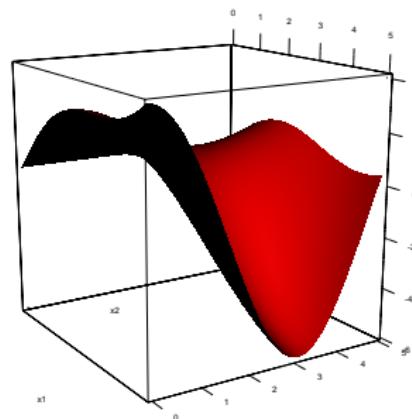
By orthogonality, $\|f\|_{L^2(\mu)}^2 = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} \sigma_{\mathbf{u}}^2$, where $\sigma_{\mathbf{u}}^2 := \|f_{\mathbf{u}}\|_{L^2(\mu)}^2$.

The ratios $S_{\mathbf{u}} := \sigma_{\mathbf{u}}^2 / (\sum_{\mathbf{v} \neq \emptyset} \sigma_{\mathbf{v}}^2)$ are referred to as **Sobol' indices** ($\mathbf{u} \neq \emptyset$).

Example on a simple test function

We consider the following test function over $[0, 5]^2$:

$$f(\mathbf{x}) = \sin(x_1) + x_1 \cos(x_2)$$



Example on a simple test function

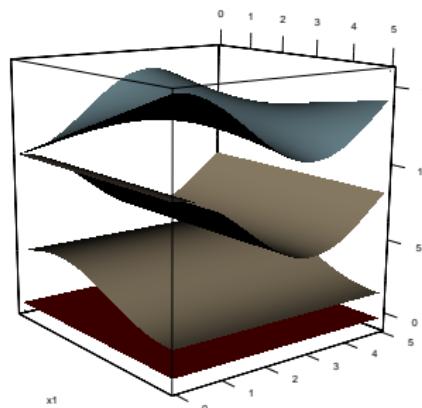
The FANOVA decomposition of f can be obtained analytically:

$$f_{\emptyset}(\mathbf{x}) = 0.2(1 - \cos(5)) + 0.5 \sin(5)$$

$$f_{\{1\}}(\mathbf{x}) = \sin(x_1) + 0.2x_1 \sin(5) - f_{\emptyset}$$

$$f_{\{2\}}(\mathbf{x}) = 0.2(1 - \cos(5)) + 0.1 \cos(x_2) - f_{\emptyset}$$

$$f_{\{1,2\}}(\mathbf{x}) = f(\mathbf{x}) - f_{\{1\}}(\mathbf{x}) - f_{\{2\}}(\mathbf{x}) - f_{\emptyset}$$



Before going further: A few fundamental references



W. Hoeffding (1948)

A class of statistics with asymptotically normal distribution

Annals of Mathematical Statistics, 19, 293-325



B. Efron and C. Stein (1981)

The jackknife estimate of variance

The Annals of Statistics, 9:586-596



A. Antoniadis (1984)

Analysis of variance on function spaces

Math. Oper. Forsch. und Statist., series Statistics, 15(1):59-71



I.M. Sobol' (1993)

Sensitivity estimates for nonlinear mathematical models

Mathematical Modelling and Computational Experiments, 1:407-414.

Revisiting FANOVA: an operator approach

Let \mathcal{F} be a subspace of \mathbb{R}^D , P_j ($1 \leq j \leq d$) be set of commuting projections on \mathcal{F} s.t. $P_j(f) = f$ if f does not depend on x_j , and $P_j(f)$ does not depend on x_j .

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The identity operator $I_{\mathcal{F}} : \mathcal{F} \longrightarrow \mathcal{F}$ can be decomposed as

$$I_{\mathcal{F}} = \prod_{j=1}^d [(I - P_j) + P_j] = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} \overbrace{\left(\prod_{j \in \mathbf{u}} (I - P_j) \right)}^{T_{\mathbf{u}}} \left(\prod_{j \in \{1, \dots, d\} \setminus \mathbf{u}} P_j \right),$$

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whereof $\forall f \in \mathcal{F}, f = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} f_{\mathbf{u}}$ with $f_{\mathbf{u}} := T_{\mathbf{u}}(f)$.



F.Y. Kuo, I.H. Sloan, G.W. Wasilkowski, and H. Wozniakowski (2010)

On decompositions of multivariate functions

Mathematics of Computation, 79, 953 - 966

The standard FANOVA is obtained when \mathcal{F} is the set of square integrable functions on $D = [0, 1]^d$, and the P_j 's are partial integration operators:

$$P_j(f)(\mathbf{x}) = \int_0^1 f(x_1, \dots, x_{j-1}, t, x_{j+1}, \dots, x_d) d\mu_j(t),$$

leading to the following operators:

$$T_{\mathbf{u}}(f) = \sum_{\mathbf{v} \subseteq \mathbf{u}} (-1)^{|\mathbf{u}| - |\mathbf{v}|} \int_{[0,1]^{d-|\mathbf{v}|}} f(\mathbf{x}) d\mu_{-\mathbf{v}}(\mathbf{x}_{-\mathbf{v}})$$

where $\mu_{-\mathbf{v}}$ denotes $\prod_{j \in [1\dots d] \setminus \mathbf{v}} \mu_j$ and $\mathbf{x}_{-\mathbf{v}}$ is defined accordingly.

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Example: Low order terms of the decomposition ($i, j \in \{1, \dots, d\}$)

$$T_\emptyset(f) = \int_{[0,1]^d} f(\mathbf{x}) d\mu(\mathbf{x})$$

$$T_{\{i\}}(f)(x_i) = \int_{[0,1]^{d-1}} f(\mathbf{x}) d\mu_{-\{i\}}(\mathbf{x}_{-\{i\}}) - \int_{[0,1]^d} f(\mathbf{x}) d\mu(\mathbf{x})$$

$$\begin{aligned} T_{\{i,j\}}(f)(x_i, x_j) &= \int_{[0,1]^{d-2}} f(\mathbf{x}) d\mu_{-\{i,j\}}(\mathbf{x}_{-\{i,j\}}) - \int_{[0,1]^{d-1}} f(\mathbf{x}) d\mu_{-\{i\}}(\mathbf{x}_{-\{i\}}) \\ &\quad - \int_{[0,1]^{d-1}} f(\mathbf{x}) d\mu_{-\{j\}}(\mathbf{x}_{-\{j\}}) + \int_{[0,1]^d} f(\mathbf{x}) d\mu(\mathbf{x}) \end{aligned}$$

Back to the set up: how do deal with a costly f ?

Assuming that the value of f at points $\mathbf{X}_n = \{\mathbf{x}_1, \dots, \mathbf{x}_n\} \subset D$ is known, how to estimate ANOVA decompositions terms and Sobol' indices?

Back to the set up: how do deal with a costly f ?

Assuming that the value of f at points $\mathbf{X}_n = \{\mathbf{x}_1, \dots, \mathbf{x}_n\} \subset D$ is known, how to estimate ANOVA decompositions terms and Sobol' indices?

Popular workflow: replace f in ANOVA decompositions and global SA by a (cheaper) approximation \tilde{f} based on $\{(\mathbf{x}_i, f(\mathbf{x}_i)), 1 \leq i \leq n\}$, e.g.,

- Standard linear models
- Polynomial chaos models (Sudret et al.)
- Smoothing spline models (Wahba et al.)
- Kriging and GRF models

About Oakley and O'Hagan's contributions

O&O'H have suggested to estimate ANOVA terms in the Bayesian framework, where a GRF model, say $(Z_x)_{x \in D}$, is assumed for f . [Analytical expressions of the posterior means of ANOVA terms were derived in](#)



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Probabilistic Sensitivity Analysis of Complex Models: A Bayesian Approach

Journal of the Royal Statistical Society (Series B), 66(3):751-769

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- Posterior means computed by multi-dimensional numerical integration.
- k is by default a stationary Gaussian kernel
- S_u 's are estimated through a **ratio of posterior means**.

About Marrel et al.'s contribution

Marrel et al. have proposed to investigate (approximate) **posterior distributions of the S_u 's** by appealing to conditional simulations.



- A. Marrel, B. Iooss, B. Laurent, and O. Roustant (2009)
Calculations of Sobol indices for the Gaussian process metamodel
Reliability Engineering and System Safety 94:742-751

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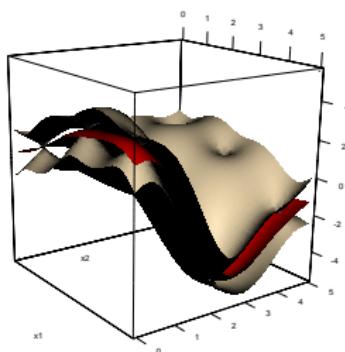
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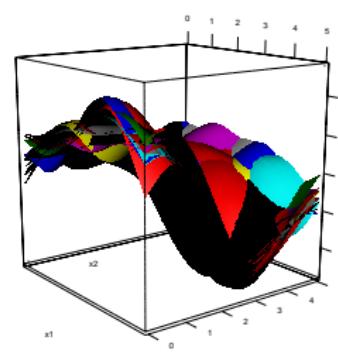
- Again, k is chosen among standard stationary covariance kernels
- The approach proves useful on a 20-dimensional test case
- Links between the chosen k and Sobol' indices are not discussed

Back to our simple test function

Here the test function $f(\mathbf{x}) = \sin(x_1) + x_1 \cos(x_2)$ is evaluated at a 9-point grid design, and a GRF model with Gaussian kernel is fitted to the data.



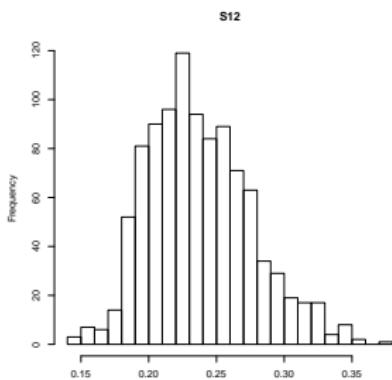
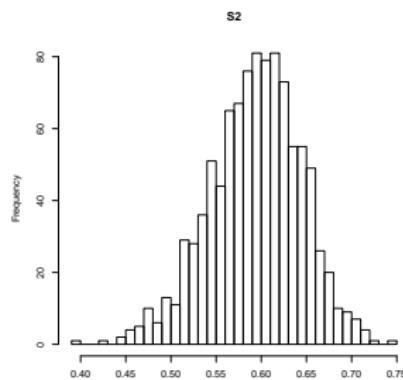
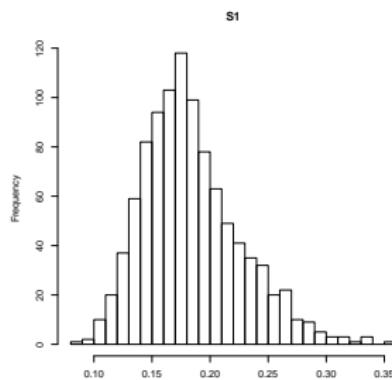
GRF model



GRF conditional simulations

Back to our simple test function

Following Marrel et al., posterior distributions of Sobol' indices are approximated relying on numerical integration and Monte Carlo :



About Durrande et al.'s contribution

Durrande et al. have focused on the choice of k , and showed that for a particular class of so-called ANOVA kernels

$$k(\mathbf{x}, \mathbf{x}') = \prod_{i=1}^d (1 + k_0(x_i, x'_i)),$$

the FANOVA decomposition of the kriging mean predictor m can be calculated without numerical integration.

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the FANOVA decomposition of the kriging mean predictor m can be calculated without numerical integration. Further, m 's Sobol' indices write:

$$S_{\mathbf{u}}(m) = \frac{Z_{\mathbf{x}_n}^t K^{-1} (\bigodot_{i \in \mathbf{u}} \Gamma_i) K^{-1} Z_{\mathbf{x}_n}}{Z_{\mathbf{x}_n}^t K^{-1} \left(\bigodot_{i=1}^d (1_{n \times n} + \Gamma_i) - 1_{n \times n} \right) K^{-1} Z_{\mathbf{x}_n}}$$

where Γ_i is the $n \times n$ matrix $\Gamma_i = \int_{D_i} \mathbf{k}_0^i(x_i) \mathbf{k}_0^i(x_i)^t d\mu_i(x_i)$.



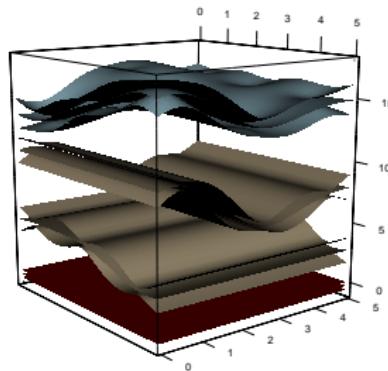
N. Durrande, D. Ginsbourger, O. Roustant, and L. Carraro (2013)

ANOVA kernels and RKHS of zero mean functions for model-based sensitivity analysis

Journal of Multivariate Analysis, 155:57-67

Back to our simple example

Using an ad hoc ANOVA kernel, the FANOVA decomposition of the GRF posterior mean is obtained analytically:



Focus and starting research questions

Claim: Assuming that f is some realization of a centred GRF Z with kernel k goes with implicit assumptions concerning f 's FANOVA decomposition...

Focus and starting research questions

Claim: Assuming that f is some realization of a centred GRF Z with kernel k goes with implicit assumptions concerning f 's FANOVA decomposition...

- How are the different FANOVA terms (jointly) distributed?
- What is the interplay between k and this distribution?
- How does conditioning on data affect those results?
- What consequences for Sobol' indices estimation under a GRF prior?

A fundamental decomposition result for GRFs

Let $(Z_x)_{x \in D}$ be a centred GRF with a.s. squared-integrable paths and denote by $k : D \times D \rightarrow \mathbb{R}$ its covariance kernel (say D compact, k continuous).

Then Z admits the following *pathwise ANOVA decomposition* almost surely:

$$Z = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} T_{\mathbf{u}}(Z).$$

where the $(T_{\mathbf{u}}(Z)_x)_{x \in D}$ are centred GRFs with respective covariance kernels $T_{\mathbf{u}} \otimes T_{\mathbf{u}}(k)$ ($\mathbf{u} \subseteq \{1, \dots, d\}$).

Moreover, $(T_{\mathbf{u}}(Z)_x; \mathbf{u} \subseteq \{1, \dots, d\})_{x \in D}$ defines a 2^d -dimensional vector-valued GRF with cross-covariances $T_{\mathbf{u}} \otimes T_{\mathbf{v}}(k)$ ($(\mathbf{u}, \mathbf{v}) \subseteq \{1, \dots, d\}^2$).

Example: The 1-dimensional Brownian Motion

Let $(B_t)_{t \in [0,1]}$ be a Brownian Motion on $[0, 1]$, i.e. a centred Gaussian field ($d = 1 \rightarrow$ “process”) with kernel $(s, t) \in [0, 1]^2 \rightarrow k(s, t) = \min(s, t)$.

The 1-dimensional pathwise ANOVA decomposition of B writes

$$B_t = \underbrace{\int_0^1 B_v dv}_{(\tau_{\emptyset} B)_t} + \underbrace{\left(B_t - \int_0^1 B_v dv \right)}_{(\tau_{\{1\}} B)_t}$$

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$T_\emptyset B$ is a centred Gaussian process with covariance kernel

$$(T_\emptyset \otimes T_\emptyset k)(s, t) = \left(\int_0^1 \int_0^1 \min(s, t) ds dt \right) \mathbf{1}_{[0,1]^2}(s, t) = \frac{1}{3} \mathbf{1}_{[0,1]^2}(s, t)$$

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$T_{\{1\}} B$ is a centred Gaussian process with covariance kernel

$$(T_{\{1\}} \otimes T_{\{1\}} k)(s, t) = \left(\min(s, t) - \left(t - \frac{t^2}{2} \right) - \left(s - \frac{s^2}{2} \right) + \frac{1}{3} \right) \mathbf{1}_{[0,1]^2}(s, t)$$

Associated ANOVA decomposition of p.d. kernels

Likewise, starting from a squared-integrable p.d. kernel $k : D \times D \rightarrow \mathbb{R}$, we obtain the following tensor product decomposition, named KANOVA :

$$k = \sum_{\mathbf{u} \subseteq \{1, \dots, d\}} \sum_{\mathbf{v} \subseteq \{1, \dots, d\}} T_{\mathbf{u}} \otimes T_{\mathbf{v}}(k).$$

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D. G., O. Roustant, D. Schuhmacher, N. Durrande and N. Lenz (2014).

On ANOVA decompositions of kernels and Gaussian random field paths

<http://arxiv.org/abs/1409.6008>

Comments

- This decomposition consists of 2^{2d} terms! (e.g., 16 terms for $d = 2$)
- $T_{\mathbf{u}} \otimes T_{\mathbf{v}}(k)$ can actually be interpreted as the orthogonal projection of k onto $\text{Ran}(T_{\mathbf{u}} \otimes T_{\mathbf{v}})$, in the sense of the tensor product structure.

Example: Back to the 1-dimensional Brownian Motion

The two projected processes $(T_\emptyset B)_t$ and $(T_{\{1\}} B)_t$ are correlated, with cross-covariance kernels

$$(T_\emptyset \otimes T_{\{1\}} k)(s, t) = \left(t - \frac{t^2}{2} - \frac{1}{3} \right) \mathbf{1}_{[0,1]^2}(s, t)$$

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To sum up, the double FANOVA decomposition of $k(s, t) = \min(s, t)$ writes

$$\begin{aligned} k(s, t) &= \frac{1}{3} \mathbf{1}_{[0,1]^2}(s, t) + \left(s - \frac{s^2}{2} - \frac{1}{3} \right) \mathbf{1}_{[0,1]^2}(s, t) + \left(t - \frac{t^2}{2} - \frac{1}{3} \right) \mathbf{1}_{[0,1]^2}(s, t) \\ &\quad + \left(\min(s, t) - \left(t - \frac{t^2}{2} \right) - \left(s - \frac{s^2}{2} \right) + \frac{1}{3} \right) \mathbf{1}_{[0,1]^2}(s, t) \end{aligned}$$

Sparsity and independence properties

Let $(Z_x)_{x \in D}$ be a centred GRF as before. Then, for any given $\mathbf{u} \subseteq \{1, \dots, d\}$ the two following assertions are equivalent:

- $T_{\mathbf{u}} \otimes T_{\mathbf{u}}(k) = \mathbf{0}$ ($\mu \otimes \mu - \text{a.e.}$)
- $\mathbb{P}(T_{\mathbf{u}} Z = \mathbf{0}) = 1$

Furthermore, for any $\mathbf{v} \subseteq \{1, \dots, d\}$, we have the second equivalence

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Corollary: stability of sparsity by conditioning

Let $\mathbf{X}_n = \{\mathbf{x}_1, \dots, \mathbf{x}_n\} \subset D$, $Z_{\mathbf{X}_n}$ be the values of Z at \mathbf{X}_n , and $\mathbf{u} \subset \{1, \dots, n\}$. The sparsity of Z with respect to $T_{\mathbf{u}}$ is preserved by conditioning, i.e.

$$\text{If } \mathbb{P}(T_{\mathbf{u}} Z = \mathbf{0}) = 1, \text{ then } \mathbb{P}(T_{\mathbf{u}} Z = \mathbf{0} | Z_{\mathbf{X}_n} = \mathbf{z}_n) = 1 \quad (\forall \mathbf{z}_n \in \mathbb{R}^n)$$

Spectral interpretation

Let $(Z_{\mathbf{x}})_{\mathbf{x} \in D}$ be decomposable as $Z_{\mathbf{x}} = \sum_{i=1}^{+\infty} \xi_i \phi_i(\mathbf{x})$, where $\xi_i \sim \mathcal{N}(0, \lambda_i)$ independently, and the ϕ_i 's form an orthonormal basis of $L^2(\mu)$. Then,

$$(T_{\mathbf{u}} Z)_{\mathbf{x}} = \sum_{i=1}^{+\infty} \xi_i (T_{\mathbf{u}} \phi_i)(\mathbf{x}) \quad (\mathbf{u} \subset \{1, \dots, n\})$$

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$$(T_{\mathbf{u}} Z)_x = \sum_{i=1}^{+\infty} \xi_i (T_{\mathbf{u}} \phi_i)(x) \quad (\mathbf{u} \subset \{1, \dots, n\})$$

From $\sigma_{\mathbf{u}}^2(Z) := \|T_{\mathbf{u}} Z\|_{L^2(\mu)}^2 = \sum_{i=1}^{+\infty} \sum_{j=1}^{+\infty} \xi_i \xi_j \langle T_{\mathbf{u}} \phi_i, T_{\mathbf{u}} \phi_j \rangle_{L^2(\mu)}$, we then get

$$\mathbb{E}[\sigma_{\mathbf{u}}^2(Z)] = \sum_{i=1}^{+\infty} \lambda_i \|T_{\mathbf{u}} \phi_i\|_{L^2(\mu)}^2$$

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Let $(Z_x)_{x \in D}$ be decomposable as $Z_x = \sum_{i=1}^{+\infty} \xi_i \phi_i(x)$, where $\xi_i \sim \mathcal{N}(0, \lambda_i)$ independently, and the ϕ_i 's form an orthonormal basis of $L^2(\mu)$. Then,

$$(T_{\mathbf{u}} Z)_x = \sum_{i=1}^{+\infty} \xi_i (T_{\mathbf{u}} \phi_i)(x) \quad (\mathbf{u} \subset \{1, \dots, n\})$$

From $\sigma_{\mathbf{u}}^2(Z) := \|T_{\mathbf{u}} Z\|_{L^2(\mu)}^2 = \sum_{i=1}^{+\infty} \sum_{j=1}^{+\infty} \xi_i \xi_j \langle T_{\mathbf{u}} \phi_i, T_{\mathbf{u}} \phi_j \rangle_{L^2(\mu)}$, we then get

$$\mathbb{E}[\sigma_{\mathbf{u}}^2(Z)] = \sum_{i=1}^{+\infty} \lambda_i \|T_{\mathbf{u}} \phi_i\|_{L^2(\mu)}^2$$

Using that $(\mathbb{P}(T_{\mathbf{u}} Z = \mathbf{0}) = 1) \Leftrightarrow (\mathbb{E}[\sigma_{\mathbf{u}}^2(Z)] = 0)$, both are then equivalent to

$$(\forall i : \lambda_i \neq 0, \quad T_{\mathbf{u}} \phi_i = \mathbf{0} \text{ } \mu - \text{a.e.})$$

which is in turn also equivalent to $T_{\mathbf{u}} \otimes T_{\mathbf{u}}(k) = \mathbf{0}$ ($\mu \otimes \mu - \text{a.e.}$).

Conclusions and perspectives

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Thank you for your attention!