

Homework2

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1

It is true that a nonlinear transformation of a normal distribution is no longer normal. as $n \rightarrow \infty$, $Var(\theta|y) \rightarrow 0$ can still hold local linearity property. So for local one-to-one transformation, as $n \rightarrow \infty$, $\phi|y \sim Normal$

2

assume for simplicity that the posterior distribution is continuous and that needed moments exist.

a

$$\begin{aligned}
 E(L(a|y)) &= \int (\theta - a)^2 P(\theta|y) d\theta \\
 \Rightarrow \frac{dE}{da} &= 2 \int (\theta - a) p(\theta|y) d\theta = 0 \\
 \Rightarrow \int \theta p(\theta|y) d\theta - a \cdot \int p(\theta|y) d\theta &= 0 \\
 \Rightarrow a &= E(\theta|y)
 \end{aligned}$$

b

$$\begin{aligned}
 E(L(a|y)) &= \int (\theta - a)^2 p(\theta|y) d\theta \\
 &= \int_{-\infty}^a (a - \theta)^2 p(\theta|y) d\theta + \int_a^{\infty} (\theta - a)^2 p(\theta|y) d\theta
 \end{aligned}$$

We can have:

$$\begin{aligned}
 \frac{dE}{da} &= \int_{-\infty}^a p(\theta|y) d\theta - \int_a^{\infty} p(\theta|y) d\theta = 0 \\
 &\text{so that } a \text{ is median of } \theta|y
 \end{aligned}$$

c

similarly we have:

$$\frac{dE}{da} = k_1 \int_{-\infty}^a p(\theta|y) d\theta - k_0 \int_a^{\infty} p(\theta|y) d\theta = 0$$

We have that $p(\cdot|y)$ is $\frac{k_0}{k_0+k_1}$ quantile, from c we can also have the conclusion of b

3

Unbiasedness: $E(E(\theta|y)|\theta) = \theta$

$$E(\theta E(\theta|y)) = E[E(\theta E(\theta|y)|\theta)] = E[\theta^2]$$

At the same time

$$E(\theta E(\theta|y)) = E[E(\theta E(\theta|y)|y)] = E[E(\theta|y)^2]$$

It must follow that $E[(E(\theta|y) - \theta)^2] = 0$ which assumes θ is constant.

4

$$\begin{aligned}
 p(\mu, \sigma^2|y) &\propto \sigma^{-n-2} \exp\left\{-\frac{1}{2\sigma^2}[(n-1)s^2 + n(\bar{y} - \mu)^2]\right\} \\
 \log p(\mu, \sigma^2|y) &= -\frac{(n+2)}{2} \log \sigma^2 - \frac{1}{2\sigma^2}[(n-1)s^2 + n(\bar{y} - \mu)^2] \\
 \frac{d}{d\mu} \log p &= \frac{n(\bar{y} - \mu)}{\sigma^2} \\
 \frac{d}{d\sigma^2} \log p &= -\left(\frac{n}{2} + 1\right) \cdot \frac{1}{\sigma^2} + \frac{1}{2(\sigma^2)^2}[(n-1)s^2 + n(\bar{y} - \mu)^2]
 \end{aligned}$$

So we have $\hat{\mu} = \bar{y}$, $\hat{\sigma}^2 = \frac{(n-1)s^2}{n+2}$ To derive $I(\theta)$, calculate derivatives of $\log p$ with respectives to $d\mu^2, d(\sigma^2)^2, d\mu d\sigma^2$

$$\begin{aligned}
 I(\hat{\theta}) &= \begin{bmatrix} \frac{n}{\sigma^2} & 0 \\ 0, & \frac{(n+2)^3}{2(n-1)^2 s^4} \end{bmatrix} \\
 p(\mu, \sigma^2|y) &\sim N\left(\begin{pmatrix} \hat{\mu} \\ \hat{\sigma}^2 \end{pmatrix} \cdot I^{-1}(\theta)\right) \\
 &= N\left(\begin{pmatrix} \bar{y} \\ \frac{(n-1)s^2}{n+2} \end{pmatrix} \cdot \begin{bmatrix} \frac{\sigma^2}{n} & 0 \\ 0, & \frac{2(n-1)^2 s^4}{(n+2)^3} \end{bmatrix}\right)
 \end{aligned}$$