

Exercises on characteristic functions

Exercise [3.3.10]

1. Denoting by $\Phi_X(\theta)$ the ch.f. of X , since X and \tilde{X} are i.i.d., the ch.f. of $-\tilde{X}$ is $\Phi_X(-\theta) = \overline{\Phi_X(\theta)}$. Hence, by Lemma 3.3.8,

$$\Phi_Z(\theta) = \Phi_X(\theta)\overline{\Phi_X(\theta)} = |\Phi_X(\theta)|^2 \geq 0.$$

2. If $U = X - \tilde{X}$ for some i.i.d. X and \tilde{X} , then by part (a), its ch.f. $\Phi_U(\theta)$ must be a real-valued non-negative function. Recall that the ch.f. of the uniform random variable on (a, b) is $\Phi_U(\theta) = e^{i\theta(a+b)/2} \sin(c\theta)/(c\theta)$ for $c = (b - a)/2$ (see Example 3.3.7). This function is real-valued only when $a = -b$ and even then $\sin(b\theta) = -1$ for $\theta = 3\pi/(2b) > 0$, leading to the stated conclusion.

Exercise [3.3.20]

1. Recall Example 3.3.7 showing that the Uniform Distribution on $(-1, 1)$, which is of bounded probability density function, has the ch.f. $\sin(\theta)/\theta$. Clearly, $\int_{\mathbb{R}} (|\sin \theta|/|\theta|) d\theta = \infty$ (consider $\theta \in [\pi n + \pi/4, \pi n + 3\pi/4]$, $n = 0, 1, 2, \dots$ for which $|\sin \theta| \geq 1/\sqrt{2}$).
2. Recall Example 3.3.13 showing that the Cauchy distribution has the ch.f. $\exp(-|\theta|)$ which is not differentiable at $\theta = 0$.

Exercise [3.3.21]

Combining Lemma 3.3.8 and Example 3.3.7, we deduce that

$$\Phi_{S_n}(\theta) = (\sin \theta / \theta)^n.$$

For any $n \geq 2$, the integral $\int |(\sin \theta)/\theta|^n d\theta$ is finite. Thus, by the inversion formula (3.3.7), the r.v. S_n has the bounded continuous probability density function

$$\begin{aligned} f_{S_n}(s) &= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-i\theta s} (\sin \theta / \theta)^n d\theta \\ &= \frac{1}{\pi} \int_0^{\infty} \cos(\theta s) (\sin \theta / \theta)^n d\theta \end{aligned}$$

with the latter identity due to the fact that $(\sin \theta)/\theta$ is invariant under the change of variable $\theta \mapsto -\theta$. Since $S_n \leq n$, the continuous p.d.f. $f_{S_n}(\cdot)$ must be identically zero for $s > n$, yielding the stated conclusion that $\int_0^{\infty} \cos(\theta s) (\sin \theta / \theta)^n d\theta = 0$ for all $s > n \geq 2$.

An exercise on weak convergence of measures

1. Indeed $|A_n| = \binom{n}{n/2}$ is finite and

$$\nu_n = \frac{1}{|A_n|} \sum_{\xi \in A_n} \delta_{\xi} \tag{1}$$

(this identity can be checked on the π -system $\mathcal{P} = \{N_\ell(\omega) : \ell \in \mathbb{N}, \omega \in \Omega\}$). Each δ_ξ is a probability measure, hence ν_n is a probability measure.

2. We claim that ν_n converges weakly to the uniform measure ν_∞ , defined by

$$\nu_\infty(N_\ell(\omega)) = \frac{1}{2^\ell}. \quad (2)$$

In order to prove this fact, we will show that, for any bounded continuous function $h : \{0, 1\}^\mathbb{N} \rightarrow \mathbb{R}$, $\lim_{n \rightarrow \infty} \nu_n(h) = \nu_\infty(h)$. Let us start by a function h measurable on $\sigma(\{N_\ell(\omega) : \omega \in \Omega\})$, i.e. depending only on the first ℓ coordinates of ω . For such a function we have $h(\omega) = h_\ell(\omega_1^\ell)$, $h_\ell : \{0, 1\}^\ell \rightarrow \mathbb{R}$. Therefore

$$\nu_n(h) = \sum_{\xi_1^\ell} h_\ell(\xi_1^\ell) \nu_n(\{\omega : \omega_1^\ell = \xi_1^\ell\}). \quad (3)$$

But, letting $k = \xi_1 + \dots + \xi_\ell$, we have

$$\lim_{n \rightarrow \infty} \nu_n(\{\omega : \omega_1^\ell = \xi_1^\ell\}) = \lim_{n \rightarrow \infty} \binom{n}{n/2}^{-1} \binom{n-\ell}{n/2-k} = \frac{1}{2^\ell}, \quad (4)$$

where the last equality is a straightforward application of Stirling formula. This proves the claim for $h \in \text{m}\sigma(\{N_\ell(\omega) : \omega \in \Omega\})$.

Consider now a general bounded continuous function h , and let for $\ell \in \mathbb{N}$, $\hat{h}_\ell \in \text{m}\sigma(\{N_\ell(\omega) : \omega \in \Omega\})$ be defined by $\hat{h}_\ell(\omega) = h(\omega_1^\ell, 0, 0, 0 \dots)$. By Fact 1, we have, for any probability measure μ , $|\mu(h) - \mu(\hat{h}_\ell)| \leq \int |h(\omega) - \hat{h}_\ell(\omega)| d\mu(\omega) \leq \delta(\ell)$. Therefore

$$|\nu_n(h) - \nu_\infty(h)| \leq |\nu_n(h) - \nu_n(\hat{h}_\ell)| + |\nu_n(\hat{h}_\ell) - \nu_\infty(\hat{h}_\ell)| + |\nu_\infty(h) - \nu_\infty(\hat{h}_\ell)| \quad (5)$$

$$\leq |\nu_n(\hat{h}_\ell) - \nu_\infty(\hat{h}_\ell)| + 2\delta(\ell). \quad (6)$$

By letting $n \rightarrow \infty$, and using the above result, we get

$$\lim_{n \rightarrow \infty} |\nu_n(h) - \nu_\infty(h)| \leq 2\delta(\ell). \quad (7)$$

The thesis follows because ℓ can be taken arbitrarily large.