

Frequency Domain Analysis

Aim: To study the frequency (corresponding to periodic component) properties of time series and identify dominant frequencies that drive the time series

Tool: Spectral density function

Defⁿ: Spectral density

Suppose that $\{x_t\}$ is a stationary zero mean time series with autocovariance $f^{\wedge} \gamma(\cdot)$ satisfying $\sum_h |\gamma(h)| < \infty$. The spectral density of $\{x_t\}$ is the function $f(\cdot)$ defined by

$$f(\lambda) = \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} e^{-i\lambda h} \gamma(h); \quad -\infty < \lambda < \infty$$

Important properties

(1) Alternate form

$$\begin{aligned} f(\lambda) &= \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} \gamma(h) (\cos \lambda h - i \sin \lambda h) \\ &= \frac{1}{2\pi} \gamma(0) (\cos 0 - i \sin 0) + \frac{1}{2\pi} \left(\sum_{h=1}^{\infty} \gamma(h) (\cos \lambda h + \cos(-\lambda h) - i \sin(\lambda h) - i \sin(-\lambda h)) \right) \end{aligned}$$

$$\text{i.e. } f(\lambda) = \frac{1}{2\pi} \left(\gamma(0) + 2 \sum_{h=1}^{\infty} \gamma(h) \cos \lambda h \right)$$

(2) $f(\cdot)$ is even

$$\begin{aligned} f(\lambda) &= \frac{1}{2\pi} \left(\gamma(0) + 2 \sum_{h=1}^{\infty} \gamma(h) \cos \lambda h \right) \\ &= \frac{1}{2\pi} \left(\gamma(0) + 2 \sum_{h=1}^{\infty} \gamma(h) \cos(-\lambda h) \right) \\ &= f(-\lambda) \quad \forall \lambda \end{aligned}$$

(3) $f(\lambda)$ is periodic with period 2π

$$f(\lambda) = \frac{1}{2\pi} \left(\gamma(0) + 2 \sum_{h=1}^{\infty} \gamma(h) \cos \lambda h \right)$$

Note that $\cos((\lambda + 2\pi k)h) = \cos \lambda h$ for any integer k & h

$$f(\lambda + 2\pi k) = \frac{1}{2\pi} \left(\gamma(0) + 2 \sum_{h=1}^{\infty} \gamma(h) \cos((\lambda + 2\pi k)h) \right)$$

$$= \frac{1}{2\pi} \left(\gamma(0) + 2 \sum_{h=1}^{\infty} \gamma(h) \cos \lambda h \right)$$

$$= f(\lambda) \quad \text{for any integer } k$$

Hence, it suffices to confine attention to the values of f on the interval $-\pi$ to π as if we know the values of $f(\lambda) \forall \lambda \in [-\pi, \pi]$, we can infer the values of $f(\lambda)$ for any λ .

(4) $f(\lambda) \geq 0 \quad \forall \lambda$

For any positive integer N , define

$$f_N(\lambda) = \frac{1}{2\pi N} E \left(\left| \sum_{r=1}^N X_r e^{-ir\lambda} \right|^2 \right)$$

Note that $f_N(\lambda) \geq 0 \quad \forall \lambda$

$$\begin{aligned}
f_N(\lambda) &= \frac{1}{2\pi N} E \left(\left| \sum_{r=1}^N X_r e^{-ir\lambda} \right|^2 \right) \\
&= \frac{1}{2\pi N} E \left(\sum_{r=1}^N X_r e^{-ir\lambda} \right) \left(\sum_{s=1}^N X_s e^{is\lambda} \right) \\
&= \frac{1}{2\pi N} E \left(X_1 e^{-i\lambda} + X_2 e^{-2i\lambda} + \dots + X_N e^{-Ni\lambda} \right) \\
&\quad \left(X_1 e^{i\lambda} + X_2 e^{2i\lambda} + \dots + X_N e^{Ni\lambda} \right)
\end{aligned}$$

$$f_N(\lambda) = \frac{1}{2\pi N} \sum_{|h| < N} (N - |h|) e^{-ih\lambda} \gamma(h)$$

[you may recall a result from Lecture #

$$NV(\bar{X}_N) = \frac{1}{N} \sum_{|h| < N} (N - |h|) \gamma(h) \quad \text{with } \sum |\gamma(h)| < \infty$$

$$\longrightarrow \sum_{h=-\infty}^{\infty} \gamma(h) \quad \text{as } N \rightarrow \infty$$

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Proceeding exactly along the same lines of the proof of this result, it follows that

$$f_N(\lambda) \longrightarrow \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} e^{-ih\lambda} \gamma(h) = f(\lambda)$$

Thus, $f_N(\lambda) \geq 0 \quad \forall N \quad \forall \lambda$ as $N \rightarrow \infty$

and $f_N(\lambda) \rightarrow f(\lambda)$ as $N \rightarrow \infty$

$$\Rightarrow f(\lambda) \geq 0 \quad \forall \lambda$$

(5) Inversion formula

Let $\{r(h)\}$ be an absolutely summable sequence of auto covariances associated with a covariance stationary process and let the spectral density of the process $\{X_t\}$ be

$$f(\lambda) = \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} e^{-i h \lambda} r(h)$$

Then

$$r(k) = \int_{-\pi}^{\pi} f(\lambda) e^{i \lambda k} d\lambda$$

Consider,

$$\begin{aligned} \int_{-\pi}^{\pi} e^{i \lambda k} f(\lambda) d\lambda &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \left(\sum_{h=-\infty}^{\infty} r(h) e^{-i \lambda h} \right) e^{i \lambda k} d\lambda \\ &= \frac{1}{2\pi} \int_{-\pi}^{\pi} \sum_{h=-\infty}^{\infty} r(h) e^{i \lambda (k-h)} d\lambda \\ &= \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} r(h) \int_{-\pi}^{\pi} e^{i \lambda (k-h)} d\lambda \quad (*) \end{aligned}$$

Note that

$$\begin{aligned} &\int_{-\pi}^{\pi} e^{i \lambda (k-h)} d\lambda \\ &= \int_{-\pi}^{\pi} (\cos(\lambda(k-h)) + i \sin(\lambda(k-h))) d\lambda \\ &= \begin{cases} 2\pi, & \text{if } h=k \\ 0, & \text{o/w (i.e. } h \neq k) \end{cases} \end{aligned}$$

$$\Rightarrow (*) = r(k)$$

Characterization of spectral density function

A real valued function $f(\cdot)$ defined on $[-\pi, \pi]$ is the spectral density of a stationary process with absolutely summable autocovariance function iff

$$(i) f(\lambda) = f(-\lambda) \quad \forall \lambda$$

$$(ii) f(\lambda) \geq 0 \quad \forall \lambda$$

$$\& (iii) \int_{-\pi}^{\pi} f(\lambda) d\lambda < \infty$$

Alternate characterization of ACVF

An absolutely summable $\gamma(\cdot)$ defined on the set of integers is the ACVF of a stationary process iff it is

(i) even and

$$(ii) f(\lambda) = \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} e^{-ikh\lambda} \gamma(h) \geq 0 \quad \forall \lambda \in [-\pi, \pi]$$

in which case $f(\lambda)$ is the corresponding spectral density function.

Spectral density of standard models

(I) White noise process

$$\text{Let } X_t \sim WN(0, \sigma^2)$$

$$\gamma(h) = \begin{cases} \sigma^2, & h=0 \\ 0, & \forall h \end{cases}$$

$$f(\lambda) = \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} \gamma(h) e^{-i\lambda h} = \frac{1}{2\pi} \gamma(0) = \frac{\sigma^2}{2\pi} \quad \forall \lambda$$

Alternatively, given $f(\lambda) = \frac{\sigma^2}{2\pi}$

$$\gamma(h) = \int_{-\pi}^{\pi} e^{ih\lambda} f(\lambda) d\lambda = \frac{\sigma^2}{2\pi} \int_{-\pi}^{\pi} e^{ih\lambda} d\lambda$$

$$= \frac{\sigma^2}{2\pi} \begin{cases} 2\pi, & \text{if } h=0 \\ 0, & \forall h \end{cases}$$

$$\text{i.e. } \gamma(h) = \begin{cases} \sigma^2, & h=0 \\ 0, & \forall h \end{cases}$$

II AR(1) process

$$X_t = \phi X_{t-1} + \epsilon_t; \quad \epsilon_t \sim WN(0, \sigma^2)$$

$$\gamma(h) = \frac{\sigma^2}{1-\phi^2} \phi^{|h|}$$

$$f(\lambda) = \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} e^{-i\lambda h} \left(\frac{\sigma^2}{1-\phi^2} \phi^{|h|} \right)$$

$$= \frac{1}{2\pi} \frac{\sigma^2}{1-\phi^2} \left(1 + \sum_{h=1}^{\infty} \phi^h e^{-i\lambda h} + \sum_{h=1}^{\infty} \phi^h e^{i\lambda h} \right)$$

$$= \frac{\sigma^2}{2\pi(1-\phi^2)} \left(1 + \frac{\phi e^{-i\lambda}}{1-\phi e^{-i\lambda}} + \frac{\phi e^{i\lambda}}{1-\phi e^{i\lambda}} \right)$$

$$\text{i.e. } f(\lambda) = \frac{\sigma^2}{2\pi(1-\phi^2)} \left(\frac{1 + \cancel{\phi} - \cancel{\phi} e^{-i\lambda} - \cancel{\phi} e^{i\lambda} + \cancel{\phi} e^{-i\lambda} - \cancel{\phi} + \cancel{\phi} e^{i\lambda} - \phi^2}{(1-\phi e^{-i\lambda})(1-\phi e^{i\lambda})} \right)$$

$$= \frac{\sigma^2}{2\pi(1-\phi^2)} \cdot \frac{(1-\phi^2)}{(1-\phi e^{-i\lambda})(1-\phi e^{i\lambda})}$$

$$\text{i.e. } f(\lambda) = \frac{\sigma^2}{2\pi} \cdot \frac{1}{(1-\phi e^{-i\lambda})(1-\phi e^{i\lambda})} = \frac{\sigma^2}{2\pi} \frac{1}{(1-2\phi \cos \lambda + \phi^2)}$$

$$f(\lambda) = \frac{\sigma^2}{2\pi} \cdot \frac{1}{\phi(e^{-i\lambda})\phi(e^{i\lambda})}$$

iii) MA(1) process

$$X_t = \epsilon_t + \theta \epsilon_{t-1}; \quad \epsilon_t \sim WN(0, \sigma^2)$$

$$r(h) = \begin{cases} \sigma^2(1+\theta^2), & h=0 \\ \theta\sigma^2, & h=\pm 1 \\ 0, & |h| \geq 2 \end{cases}$$

$$f(\lambda) = \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} e^{ih\lambda} r(h)$$

$$= \frac{1}{2\pi} \left(\sigma^2(1+\theta^2) + \sigma^2 \theta e^{-i\lambda} + \sigma^2 \theta e^{i\lambda} \right)$$

$$= \frac{\sigma^2}{2\pi} (1 + \theta^2 + 2\theta \cos \lambda)$$

$$= \frac{\sigma^2}{2\pi} (1 + \theta e^{-i\lambda})(1 + \theta e^{i\lambda}) = \frac{\sigma^2}{2\pi} \theta(e^{-i\lambda}) \theta(e^{i\lambda})$$

Remark : Note that

$$f(\lambda) = \frac{1}{2\pi} \sum_{h=-\infty}^{\infty} (e^{-i\lambda})^h r(h) = \frac{1}{2\pi} g_X(e^{-i\lambda})$$

where $g_X(\cdot)$ is ACVF of $\{X_t\}$

$f(\cdot)$ of WN, AR(1) and MA(1) are accordingly in term of $g_X(\cdot)$

We can use the above connection between spectral density function and AC GF to obtain spectral densities of standard model for which we have earlier derived AC GF.

Spectral density f_λ of AR(p)

$$X_t = \phi_1 X_{t-1} + \dots + \phi_p X_{t-p} + \epsilon_t; \quad \epsilon_t \sim WN(0, \sigma^2)$$

$$\phi(B) X_t = \epsilon_t$$

$$\phi(B) = 1 - \phi_1 B - \dots - \phi_p B^p$$

$$\text{AC GF} \quad g_X(z) = \frac{\sigma^2}{\phi(z) \phi(z^{-1})}$$

Spectral density

$$f_X(\lambda) = \frac{1}{2\pi} g_X(e^{-i\lambda})$$

$$\text{i.e. } f_X(\lambda) = \frac{\sigma^2}{2\pi} \frac{1}{\phi(e^{-i\lambda}) \phi(e^{i\lambda})}$$

$$\text{e.g. AR(1)} \quad X_t = \phi X_{t-1} + \epsilon_t$$

$$f_X(\lambda) = \frac{\sigma^2}{2\pi} \frac{1}{(1 - \phi e^{-i\lambda})(1 - \phi e^{i\lambda})}$$

$$f_X(\lambda) = \frac{\sigma^2}{2\pi} \frac{1}{1 - \phi e^{-i\lambda} - \phi e^{i\lambda} + \phi^2}$$

Note that the spectral density f^n of $AR(p)$ can be expressed in terms of roots of AR polynomial.

Suppose n_1, \dots, n_p are the roots of

$$n^p - \phi_1 n^{p-1} - \dots - \phi_p = 0$$

Then
$$f_x(\lambda) = \frac{\sigma^2}{2\pi} \left[\prod_{j=1}^p (e^{i\lambda} - n_j)(e^{-i\lambda} - n_j) \right]^{-1}$$

Spectral density of $MA(q)$

$$Y_t = \sum_{j=0}^q \theta_j \epsilon_{t-j} \quad ; \quad \epsilon_t \sim WN(0, \sigma^2)$$

ACGF
$$g_Y(z) = \sigma^2 \theta(z) \theta(z^{-1}) \quad ; \quad \theta(B) = \theta_0 + \theta_1 B + \dots + \theta_q B^q$$

$$\theta_0 = 1$$

Spectral density f^n

$$\begin{aligned} f_Y(\lambda) &= \frac{\sigma^2}{2\pi} \theta(e^{-i\lambda}) \theta(e^{i\lambda}) \\ &= \frac{\sigma^2}{2\pi} \left(\sum_{j=0}^q \theta_j e^{-ij\lambda} \right) \left(\sum_{j=0}^q \theta_j e^{ij\lambda} \right) \end{aligned}$$

if m_1, m_2, \dots, m_q are the roots of

$$m^q + \theta_1 m^{q-1} + \dots + \theta_q = 0$$

Then
$$f_Y(\lambda) = \frac{\sigma^2}{2\pi} \left(\prod_{j=1}^q (e^{-i\lambda} - m_j)(e^{i\lambda} - m_j) \right)$$

Spectral density of ARMA(p, q)

$$X_t = \phi_1 X_{t-1} + \dots + \phi_p X_{t-p} + \epsilon_t + \theta_1 \epsilon_{t-1} + \dots + \theta_q \epsilon_{t-q}$$

$$\epsilon_t \sim WN(0, \sigma^2)$$

$$\phi(B)X_t = \theta(B)\epsilon_t$$

$$\phi(B) = 1 - \phi_1 B - \phi_2 B^2 - \dots - \phi_p B^p$$

$$\theta(B) = 1 + \theta_1 B + \dots + \theta_q B^q$$

$$\text{ACGF: } \gamma_X(z) = \sigma^2 \frac{\theta(z)\theta(\bar{z}')}{\phi(z)\phi(\bar{z}')}.$$

Spectral density f^n :

$$f_X(\lambda) = \frac{\sigma^2}{2\pi} \cdot \frac{\theta(e^{-i\lambda})\theta(e^{i\lambda})}{\phi(e^{-i\lambda})\phi(e^{i\lambda})}.$$

If n_1, \dots, n_p are the roots of

$$n^p - \phi_1 n^{p-1} - \dots - \phi_p = 0$$

And m_1, \dots, m_q are the roots of

$$m^q + \theta_1 m^{q-1} + \dots + \theta_q = 0$$

Then

$$f_X(\lambda) = \frac{\sigma^2}{2\pi} \frac{\left(\prod_{j=1}^q (e^{i\lambda} - m_j)(e^{-i\lambda} - \bar{m}_j) \right)}{\left(\prod_{j=1}^p (e^{i\lambda} - n_j)(e^{-i\lambda} - \bar{n}_j) \right)}.$$