

Remark: Ex 1, 2, 3 are direct approaches to prove convergence in law/distribution.

Convergence in law can also be proved using m.g.f. convergence.

(4) $X_n \sim \text{Bin}(n, \theta)$

Suppose $n \rightarrow \infty$, $\exists n\theta = \lambda$ is fixed i.e. $\theta = \frac{\lambda}{n}$

$$\begin{aligned} M_{X_n}(t) &= ((1-\theta) + \theta e^t)^n \\ &= \left(1 + \frac{\lambda}{n}(e^t - 1)\right)^n \end{aligned}$$

$$\rightarrow e^{\lambda(e^t - 1)} \text{ as } n \rightarrow \infty$$

$$\Rightarrow X_n \xrightarrow{L} X; \text{ where } X \sim P(\lambda)$$

$$(5) X_1, \dots \text{ i.i.d. } N(0, 1)$$

$$\bar{X}_n \sim N(0, \frac{1}{n})$$

$$M_{\bar{X}_n}(t) = e^{\frac{t^2}{2n}} \rightarrow 1 \text{ as } n \rightarrow \infty$$

$$\Rightarrow \bar{X}_n \xrightarrow{L} X; \text{ where } X \text{ is degenerate at } 0$$

$$(6) X_n \sim \chi_n^2; E X_n = n; V X_n = 2n$$

$$M_{X_n}(t) = (1 - 2t)^{-n/2} \quad t < \frac{1}{2}$$

$$Y_n = \frac{X_n - n}{\sqrt{2n}}$$

$$M_{Y_n}(t) = E \left(e^{t \left(\frac{X_n - n}{\sqrt{2n}} \right)} \right)$$

$$= e^{-\frac{tn}{\sqrt{2n}}} E \left(e^{\frac{t X_n}{\sqrt{2n}}} \right)$$

$$= e^{-\frac{tn}{\sqrt{2n}}} \left(1 - \frac{2t}{\sqrt{2n}} \right)^{-n/2} \quad t < \frac{\sqrt{n}}{2}$$

$$= \left(e^{t\sqrt{\frac{2}{n}}} \right)^{-n/2} \left(1 - \sqrt{\frac{2}{n}} t \right)^{-n/2}$$

$$= \left[\left(1 + t\sqrt{\frac{2}{n}} + \frac{t^2 2/n}{2!} + \frac{t^3 2^{3/2}/n^{3/2}}{3!} + \dots \right) - t\sqrt{\frac{2}{n}} \left(1 + t\sqrt{\frac{2}{n}} + \frac{t^2 2/n}{2!} + \dots \right) \right]^{-n/2}$$

$$= \left(1 - \frac{t^2}{n} + \frac{K_1}{n^{3/2}} - \dots \right)^{-n/2}$$

$$= \left(1 - \frac{t^2}{n} + o\left(\frac{1}{n}\right) \right)^{-n/2}$$

$$\rightarrow e^{t^2/2} \leftarrow \text{m.g.f. of } N(0, 1)$$

$$\Rightarrow \frac{X_n - n}{\sqrt{2n}} \xrightarrow{L} N(0, 1) \text{ r.v.}$$

Some important results

(i) If $X_n \xrightarrow{p} X$ then $X_n \xrightarrow{L} X$

Converse is not true in general.

However, if $X_n \xrightarrow{L} c$ (a const), then

$$X_n \xrightarrow{p} c$$

(ii) Slutsky's Lemma

If $X_n \xrightarrow{L} X$ and $Y_n \xrightarrow{p} c$, then

(i) $X_n \pm Y_n \xrightarrow{L} X \pm c$

(ii) $X_n Y_n \xrightarrow{L} cX$

(iii) $\frac{X_n}{Y_n} \xrightarrow{L} X/c \quad (c \neq 0)$

(iii) Δ -method or Δ -rule

Let $\{X_n\}$ be a seq of random variables \Rightarrow

$$\sqrt{n}(X_n - \theta) \xrightarrow{L} N(0, \sigma^2)$$

Suppose g be real valued f^n differentiable at $\theta \Rightarrow$

$g'(\theta) \neq 0$, then

$$\sqrt{n}(g(X_n) - g(\theta)) \xrightarrow{L} N(0, (g'(\theta))^2 \sigma^2)$$

Central Limit Theorem (CLT)

In WLLN, we investigated

$$\frac{S_n - a_n}{b_n} \xrightarrow{P} 0$$

i.e. convergence of $\frac{S_n - a_n}{b_n}$ to a degenerate distⁿ
(degenerate at 0)

In CLT, we investigate convergence of $\frac{S_n - a_n}{b_n}$ to a non-degenerate distⁿ.

Defⁿ: Let X_1, X_2, \dots be a seq of i.i.d. r.v.s with common d.f F . We say that F belongs to the domain of attraction of a distⁿ V if there exists centering constants $\{A_n\}$ and norming constants $\{B_n\}$ ($B_n > 0$) \Rightarrow

as $n \rightarrow \infty$

$$P\left(\frac{S_n - A_n}{B_n} \leq x\right) \rightarrow V(x) \quad \text{at all continuity points of } V$$

$$S_n = \sum_{i=1}^n X_i$$

$$\text{i.e. } \frac{S_n - A_n}{B_n} \xrightarrow{L} X; \text{ where d.f. of } X \text{ is } V(\cdot)$$

Lindeberg-Levy CLT

Let $\{X_n\}$ be a seq of i.i.d. r.v.s with $E(X_i) = \mu$ and $V(X_i) = \sigma^2 < \infty$, then for $S_n = \sum_{i=1}^n X_i$

$$\frac{S_n - ES_n}{\sqrt{V(S_n)}} \xrightarrow{L} N(0, 1) \text{ r.v.}$$

$$\text{i.e. } \frac{\sum_{i=1}^n X_i - n\mu}{\sqrt{n} \sigma} \xrightarrow{L} X ; X \sim N(0,1)$$

$$\text{i.e. } \frac{n\bar{X}_n - n\mu}{\sqrt{n} \sigma} \xrightarrow{L} X ; X \sim N(0,1)$$

$$\text{i.e. } \frac{\sqrt{n}(\bar{X}_n - \mu)}{\sigma} \xrightarrow{L} X ; X \sim N(0,1)$$

$$\text{i.e. } \sqrt{n}(\bar{X}_n - \mu) \xrightarrow{L} Y ; Y \sim N(0, \sigma^2)$$

Applications of CLT

(1) X_1, \dots, X_n are i.i.d. $\text{Exp}(\theta)$ (exponential with mean θ)

$$\text{i.e. } f_X(x) = \begin{cases} \frac{1}{\theta} e^{-x/\theta}, & x > 0 \\ 0, & \text{o.w.} \end{cases}$$

$$E(X_i) = \theta ; V(X_i) = \theta^2$$

By CLT

$$\frac{\sqrt{n}(\bar{X}_n - \theta)}{\theta} \xrightarrow{L} N(0,1)$$

$$\text{i.e. } \sqrt{n}(\bar{X}_n - \theta) \xrightarrow{L} N(0, \theta^2)$$

Further, suppose we are interested in asymptotic dist ($n \rightarrow \infty$)

of $\frac{1}{\bar{X}_n}$; we can apply Δ -rule on the CLT result with

$$g(x) = \frac{1}{x} \quad g'(x) = -\frac{1}{x^2} \quad g'(\theta) \neq 0$$

$$\Delta\text{-rule} \Rightarrow \sqrt{n} \left(\frac{1}{\bar{X}_n} - \frac{1}{\theta} \right) \xrightarrow{L} N \left(0, \left(\frac{1}{\theta^4} \right) \cdot \theta^2 \right)$$

\uparrow
 $(g'(\theta))^2$

$$\text{i.e. } \sqrt{n} \left(\frac{1}{\bar{X}_n} - \frac{1}{\theta} \right) \xrightarrow{L} N \left(0, \frac{1}{\theta^2} \right)$$