

Note: We can also define a autocorrelation matrix function for a covariance stationary process as

$$R_X(h) = \begin{pmatrix} \rho_{11}(h) & \rho_{12}(h) & \dots & \rho_{1m}(h) \\ & \ddots & & \\ & & \ddots & \\ & & & \rho_{mm}(h) \end{pmatrix}$$

$$\rho_{ij}(h) = \frac{\gamma_{ij}(h)}{(\gamma_{ii}(0) \gamma_{jj}(0))^{1/2}}$$

$$R_X(h) = D_0^{-1/2} \Gamma_X(h) D_0^{-1/2}$$

$$D_0 = \text{diag}(\gamma_{11}(0), \dots, \gamma_{mm}(0))$$

Standard multivariate processes

~~Def~~ (I) Vector White noise (VWN)

$$\tilde{Z}_t \sim \text{VWN}(\underline{0}, \Sigma)$$

(\tilde{Z}_t is a VWN with mean vector $\underline{0}$ and covariance matrix Σ)

$E \tilde{Z}_t = \underline{0}$ and covariance matrix Σ as

$$\Gamma(h) = \begin{cases} \Sigma, & h=0 \\ 0, & \text{o/w} \end{cases} \quad \text{i.e. } E(\tilde{Z}_t \tilde{Z}_s') = \begin{cases} \Sigma, & t=s \\ 0, & t \neq s \end{cases}$$

Note : Note that for VWN process the vector process is uncorrelated; however, It is not necessary that the components are uncorrelated.

II : Vector Moving Average (VMA)

VMA(1): $\underset{K \times 1}{\tilde{X}}_t = \underset{K \times K}{\textcircled{H}} \underset{K \times 1}{\tilde{\epsilon}}_{t-1} + \underset{K \times 1}{\tilde{\epsilon}}_t$

$$\underset{K \times 1}{\tilde{\epsilon}}_t \sim \text{VWN}(0, \Sigma)$$

\textcircled{H} : $K \times K$ matrix of constants

$$\underset{K \times 1}{\tilde{X}}_t = (I_K + \textcircled{H} B) \underset{K \times 1}{\tilde{\epsilon}}_t$$

i.e. $\underset{K \times 1}{\tilde{X}}_t = \textcircled{H}(B) \underset{K \times 1}{\tilde{\epsilon}}_t$

$\textcircled{H}(B) = I_K + \textcircled{H} B$: MA matrix polynomial

$$E \underset{K \times 1}{\tilde{X}}_t = \underset{K \times 1}{0} \quad \forall t$$

$$\begin{aligned} \text{Cov}(\underset{K \times 1}{\tilde{X}}_t, \underset{K \times 1}{\tilde{X}}_{t+h}) &= E(\underset{K \times 1}{\tilde{X}}_t \underset{K \times 1}{\tilde{X}}_{t+h}') \\ &= E(\underset{K \times 1}{\tilde{\epsilon}}_t + \textcircled{H} \underset{K \times 1}{\tilde{\epsilon}}_{t-1} (\underset{K \times 1}{\tilde{\epsilon}}_{t+h} + \textcircled{H} \underset{K \times 1}{\tilde{\epsilon}}_{t+h-1})') \\ &= E(\underset{K \times 1}{\tilde{\epsilon}}_t \underset{K \times 1}{\tilde{\epsilon}}_{t+h}') + E(\underset{K \times 1}{\tilde{\epsilon}}_t \underset{K \times 1}{\tilde{\epsilon}}_{t+h-1}' \textcircled{H}') \\ &\quad + E(\textcircled{H} \underset{K \times 1}{\tilde{\epsilon}}_{t-1} \underset{K \times 1}{\tilde{\epsilon}}_{t+h}') \\ &\quad + E(\textcircled{H} \underset{K \times 1}{\tilde{\epsilon}}_{t-1} \underset{K \times 1}{\tilde{\epsilon}}_{t+h-1}' \textcircled{H}') \end{aligned}$$

i.e. $\Gamma_X(h) = \sum I_0(h) + \sum \textcircled{H}' I_1(h) + \textcircled{H} \sum I_{-1}(h) + \textcircled{H} \sum \textcircled{H}' I_0(h)$

$$\text{i.e. } \Gamma_X(h) = \begin{cases} \Sigma + \Theta \Sigma \Theta', & h = 0 \\ \Sigma \Theta', & h = 1 \\ \Theta \Sigma, & h = -1 \\ 0, & |h| > 1 \end{cases}$$

X_t is always covariance stationary $\forall K \times K \Theta$.

VMA(q)

$$X_t = \epsilon_t + \Theta_1 \epsilon_{t-1} + \dots + \Theta_q \epsilon_{t-q}$$

$$\Theta_q \neq 0; \epsilon_t \sim \text{VWN}(0, \Sigma)$$

$$X_t = (I_K + \Theta_1 B + \dots + \Theta_q B^q) \epsilon_t$$

$$\text{i.e. } X_t = \Theta(B) \epsilon_t$$

$$\Theta(B) = I_K + \Theta_1 B + \dots + \Theta_q B^q \text{ is the MA matrix polynomial}$$

$$E X_t = 0$$

$$\text{Cov}(X_t, X_{t+h}) = E(X_t X_{t+h}') =$$

$$\left(\Theta_0 = I_K \right) \rightarrow E \left(\Theta_0 \epsilon_t + \Theta_1 \epsilon_{t-1} + \dots + \Theta_q \epsilon_{t-q} \right) \left(\Theta_0 \epsilon_{t+h} + \Theta_1 \epsilon_{t+h-1} + \dots + \Theta_q \epsilon_{t+h-q} \right)'$$

$$\Gamma_0 = \text{Cov}(X_t, X_t) = \sum_{j=0}^q \Theta_j \Sigma \Theta_j'$$

$$\Gamma(1) = \text{Cov}(X_t, X_{t+1}) = \sum_{j=0}^{q-1} \Theta_j \Sigma \Theta_{j+1}'$$

$$\Gamma(-1) = \Gamma(1)'$$

$$\Gamma(2) = \text{Cov}(X_{\tilde{t}}, X_{\tilde{t}+2}) = \sum_{j=0}^{q-2} \textcircled{H}_j \sum \textcircled{H}'_{j+2} \quad (99)$$

$$\Gamma(-2) = \Gamma(2)'$$

$$\forall h \leq q$$

$$\Gamma(h) = \sum_{j=0}^{q-h} \textcircled{H}_j \sum \textcircled{H}'_{j+h}$$

$$\Gamma(-h) = \Gamma(h)' = \sum_{j=0}^{q-h} \textcircled{H}_{j+h} \sum \textcircled{H}'_j$$

$$\text{e.g. } \Gamma(-1) = E \left(\textcircled{H}_0 \epsilon_{\tilde{t}} + \textcircled{H}_1 \epsilon_{\tilde{t}-1} + \dots + \textcircled{H}_q \epsilon_{\tilde{t}-q} \right) \\ \left(\epsilon'_{\tilde{t}-1} \textcircled{H}'_0 + \epsilon'_{\tilde{t}-2} \textcircled{H}'_1 + \dots + \epsilon'_{\tilde{t}-q-1} \textcircled{H}'_q \right)$$

$$= \textcircled{H}_1 \sum \textcircled{H}'_0 + \dots + \textcircled{H}_q \sum \textcircled{H}'_{q-1}$$

$$\text{i.e. } \Gamma(-1) = \sum_{j=0}^{q-1} \textcircled{H}_{j+1} \sum \textcircled{H}'_j = \Gamma(1)'$$

$$\text{Also } \Gamma(h) = \text{Cov}(X_{\tilde{t}}, X_{\tilde{t}+h}) \\ = 0 \quad \forall |h| > q$$

VMA(q) is always covariance stationary

$\forall \textcircled{H}_1, \dots, \textcircled{H}_q$, matrices of constants
(a property like the univariate MA processes)

VMA(α)

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$$\tilde{X}_t = \sum_{j=0}^{\alpha} \textcircled{H}_j \tilde{\epsilon}_{t-j} ; \tilde{\epsilon}_t \sim \text{VWN}(0, \Sigma)$$

[Defⁿ: A seq of $n \times m$ matrices $\{\textcircled{H}_s\}_{s=0}^{\alpha}$ is said to be absolutely summable if each of its elements forms an absolutely summable scalar sequence
i.e., $\sum_{s=0}^{\alpha} |\textcircled{H}_{ij(s)}| < \alpha \quad \forall i, j; \begin{matrix} i=1(1)n \\ j=1(1)m \end{matrix}$]

Suppose that for the VMA(α) process the sequence of matrices $\{\textcircled{H}_j\}_{j=0}^{\alpha}$ is absolutely summable,

$$\text{i.e.} \quad \sum_{s=0}^{\alpha} |\textcircled{H}_{ij(s)}| < \alpha \quad \forall i, j = 1(1)K$$

($\textcircled{H}_{ij(s)}$: (i, j)th element of \textcircled{H}_s)

Then VMA(α) is covariance stationary

Note: If $\{\textcircled{H}_j\}$ is square summable, then also VMA(α) would be covariance stationary.

If VMA(α) is covariance stationary, then it's auto covariance matrix function is

$$M(h) = \sum_{j=0}^{\alpha} \textcircled{H}_j \Sigma \textcircled{H}_{j+h}' ; h=0, 1, 2, \dots$$

$$M(-h) = M(h)'$$

VAR models

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Consider a VAR(p) process (say K-variate)

$$\underset{K \times 1}{\tilde{X}}_t = \underset{K \times K}{\tilde{\Phi}}_1 \underset{K \times 1}{\tilde{X}}_{t-1} + \underset{K \times K}{\tilde{\Phi}}_2 \underset{K \times 1}{\tilde{X}}_{t-2} + \dots + \underset{K \times K}{\tilde{\Phi}}_p \underset{K \times 1}{\tilde{X}}_{t-p} + \underset{K \times 1}{\tilde{\epsilon}}_t;$$

$$\underset{K \times K}{\tilde{\Phi}}_p \neq 0, \quad \underset{K \times 1}{\tilde{\epsilon}}_t \sim \text{VWN}(\underset{K \times 1}{0}, \underset{K \times K}{\Sigma})$$

$$\text{Cov}(\underset{K \times 1}{\tilde{\epsilon}}_t, \underset{K \times 1}{\tilde{X}}_{t-j}') = 0 \quad \forall j > 0$$

$$\text{i.e. } E(\underset{K \times 1}{\tilde{\epsilon}}_t \underset{K \times 1}{\tilde{X}}_{t-j}') = 0 \quad \forall j > 0$$

$$\left[\begin{array}{l} \underset{K \times 1}{\tilde{X}}, \underset{q \times 1}{\tilde{Y}} \text{ random vectors} \\ \text{Cov}(\underset{K \times 1}{\tilde{X}}, \underset{q \times 1}{\tilde{Y}}) = E(\underset{K \times 1}{\tilde{X}} - E(\underset{K \times 1}{\tilde{X}})(\underset{q \times 1}{\tilde{Y}} - E(\underset{q \times 1}{\tilde{Y}}))' \end{array} \right]$$

$\underset{K \times K}{\tilde{\Phi}}_1, \underset{K \times K}{\tilde{\Phi}}_2, \dots, \underset{K \times K}{\tilde{\Phi}}_p$ are $K \times K$ matrices of

constants - VAR matrices of parameters

Model can be written in terms of AR operator matrix polynomial

$$\underset{K \times K}{\tilde{\Phi}}(B) \underset{K \times 1}{\tilde{X}}_t = \underset{K \times 1}{\tilde{\epsilon}}_t$$

$$\text{Where } \underset{K \times K}{\tilde{\Phi}}(B) = \underset{K \times K}{I}_K - \underset{K \times K}{\tilde{\Phi}}_1 B - \dots - \underset{K \times K}{\tilde{\Phi}}_p B^p.$$

VAR matrix polynomial

Let us see what the VAR(p) gives us. The l^{th} row of the VAR(p) system is (at time point t)

$$\begin{aligned}
 X_{l,t} = & \left(\Phi_{l1}^{(1)} X_{1,t-1} + \Phi_{l2}^{(1)} X_{2,t-1} + \dots + \Phi_{lK}^{(1)} X_{K,t-1} \right) \\
 & + \left(\Phi_{l1}^{(2)} X_{1,t-2} + \Phi_{l2}^{(2)} X_{2,t-2} + \dots + \Phi_{lK}^{(2)} X_{K,t-2} \right) \\
 & \vdots \\
 & + \left(\Phi_{l1}^{(p)} X_{1,t-p} + \Phi_{l2}^{(p)} X_{2,t-p} + \dots + \Phi_{lK}^{(p)} X_{K,t-p} \right) \\
 & + \epsilon_{lt}
 \end{aligned}$$

$$l = 1(1)K ; t = 1(1)n$$

i.e.

$$\begin{aligned}
 X_{l,t} = & \left(\Phi_{l1}^{(1)} X_{1,t-1} + \Phi_{l1}^{(2)} X_{1,t-2} + \dots + \Phi_{l1}^{(p)} X_{1,t-p} \right) \\
 & + \left(\Phi_{l2}^{(1)} X_{2,t-1} + \Phi_{l2}^{(2)} X_{2,t-2} + \dots + \Phi_{l2}^{(p)} X_{2,t-p} \right) \\
 & \vdots \\
 & + \left(\Phi_{lK}^{(1)} X_{K,t-1} + \Phi_{lK}^{(2)} X_{K,t-2} + \dots + \Phi_{lK}^{(p)} X_{K,t-p} \right) \\
 & + \epsilon_{lt}
 \end{aligned}$$

Thus the model eqⁿ for the l^{th} variable is expressed in terms of p lags of the l^{th} variable (as it would have been for AR(p)) + p lags of all the remaining $K-1$ variables

present inside the VAR(p) system
+ instantaneous noise variable for the l^{th} variable
system

Thus the VAR(p) system takes care of using inputs
from related variables for a predictive model.