

(1)

$$\textcircled{1} \quad X_t = \epsilon_t + \theta \epsilon_{t-1} : |\theta| < 1 \quad \epsilon_t \sim WN(0, \sigma^2)$$

$$\epsilon_t = X_t - \theta \epsilon_{t-1}$$

$$\begin{aligned} \Rightarrow X_t &= \epsilon_t + \theta (X_{t-1} - \theta \epsilon_{t-2}) \\ &= \epsilon_t + \theta X_{t-1} - \theta^2 \epsilon_{t-2} \\ &= \epsilon_t + \theta X_{t-1} - \theta^2 (X_{t-2} - \theta \epsilon_{t-3}) \\ &= \epsilon_t + \theta X_{t-1} - \theta^2 X_{t-2} + \theta^3 \epsilon_{t-3} \end{aligned}$$

K^{th} step :

$$X_t \stackrel{\downarrow}{=} \epsilon_t - \sum_{i=1}^K (-\theta)^i X_{t-i} - (-\theta)^{K+1} \epsilon_{t-(K+1)}$$

$$\begin{aligned} \Rightarrow E \left(X_t - \epsilon_t + \sum_{i=1}^K (-\theta)^i X_{t-i} \right)^2 &= E \left((-\theta)^{2(K+1)} \epsilon_{t-(K+1)}^2 \right) \\ &= \sigma^2 (-\theta)^{2(K+1)} \end{aligned}$$

$$\Rightarrow \lim_{K \rightarrow \infty} E \left(X_t - \epsilon_t + \sum_{i=1}^K (-\theta)^i X_{t-i} \right)^2 = 0$$

$$\textcircled{2} \quad X_t = \phi X_{t-1} + \epsilon_t \quad |\phi| < 1 \quad \epsilon_t \sim WN(0, \sigma^2)$$

$$\begin{aligned} &= \phi (\phi X_{t-2} + \epsilon_{t-1}) + \epsilon_t \\ &= \phi^2 X_{t-2} + \phi \epsilon_{t-1} + \epsilon_t \end{aligned}$$

N^{th} step :

$$\downarrow = \phi^{N+1} X_{t-(N+1)} + \sum_{j=0}^N \phi^j \epsilon_{t-j}$$

$$\begin{aligned} \Rightarrow E \left(X_t - \sum_{j=0}^N \phi^j \epsilon_{t-j} \right)^2 &= E \left(\phi^{2(N+1)} X_{t-N-1}^2 \right) \\ &= \phi^{2(N+1)} E \left(X_{t-N-1}^2 \right) \end{aligned}$$

Ans $\{X_t\}$ is covariance stationary, $E[X_{t-N-1}^2] < \infty$ (2)

$$\Rightarrow \lim_{N \rightarrow \infty} E \left(X_t - \sum_{j=0}^N \phi^j \epsilon_{t-j} \right)^2 = 0 \quad (\text{note that } |\phi| < 1)$$

(3) $Y_t = X_{2t}$

$$X_t = \phi X_{t-1} + \epsilon_t; \quad \epsilon_t \sim WN(0, \sigma^2)$$

$$\begin{aligned} Y_t = X_{2t} &= \phi X_{2t-1} + \epsilon_{2t} \\ &= \phi (\phi X_{2t-2} + \epsilon_{2t-1}) + \epsilon_{2t} \end{aligned}$$

$$\text{i.e. } Y_t = \phi^2 X_{2(t-1)} + (\phi \epsilon_{2t-1} + \epsilon_{2t})$$

$$\text{i.e. } Y_t = \phi^2 Y_{t-1} + \eta_t$$

$$\eta_t = \epsilon_{2t} + \phi \epsilon_{2t-1} \sim WN(0, \sigma^2(1 + \phi^2))$$

$$\Rightarrow Y_t = \delta Y_{t-1} + \eta_t \text{ is a stationary AR(1)}$$

$$\delta = \phi^2; \quad |\delta| < 1$$

Ans $\{Y_t\}$ is covariance stationary, $\{Y_t\}$ is causal also.

causal representation of $\{Y_t\}$:

$$Y_t = \sum_{j=0}^{\infty} (\phi^2)^j \eta_{t-j}$$

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$$(4) \quad X_t = z_t z_{t-1} + z_{t-1} z_{t-2} + \frac{1}{4} z_{t-2} z_{t-3}$$

$$\text{Let } \eta_t = z_t z_{t-1}; \quad \eta_t \text{ 's } \Rightarrow E \eta_t = 0$$

$$\begin{aligned} \& \text{Cov}(\eta_t, \eta_s) &= E(z_t z_{t-1} z_s z_{s-1}) \\ &= \begin{cases} \sigma^4, & t = s \\ 0, & t \neq s \end{cases} \end{aligned}$$

$$\Rightarrow \eta_t \sim WN(0, \sigma^4)$$

$$\Rightarrow X_t = \eta_t + \eta_{t-1} + \frac{1}{4} \eta_{t-2} \text{ is MA}(2)$$

$$X_t = \theta(B) \eta_t; \quad \theta(B) = 1 + B + \frac{1}{4} B^2$$

$$\text{roots of } \theta(z) = 0; \text{ i.e. of } 1 + z + \frac{1}{4} z^2 = \left(1 + \frac{1}{2} z\right)^2$$

$$\text{are } -2, -2$$

$$\Rightarrow X_t \text{ is invertible MA}(2)$$

$$(5) \quad \begin{aligned} X_t &= \phi X_{t-1} + \epsilon_t; \quad \phi = \frac{1}{4} \quad \epsilon_t \sim WN(0, 1) \\ Y_t &= \delta_t - \delta_{t-1}; \quad \delta_t \sim WN(0, 1) \end{aligned} \xrightarrow{\text{indep}}$$

$$\begin{aligned} \text{Note that } X_t &= \sum_{j=0}^{\infty} \phi^j \epsilon_{t-j} \\ Y_t &= \delta_t - \delta_{t-1} \end{aligned} \Rightarrow \{X_t\} \& \{Y_t\} \text{ are indep}$$

$$E z_t = (E(1 - X_t))(E(1 + Y_t)) = 1$$

$$\text{Cov}(z_t, z_{t+h}) = E(z_t z_{t+h}) - 1$$

$$E(z_t z_{t+h}) = E(1 - X_t)(1 - X_{t+h}) E(1 + Y_t)(1 + Y_{t+h})$$

$$E(z_t z_{t+h}) = (1 + \gamma_x(h))(1 + \gamma_y(h))$$

$$\Rightarrow \text{Cov}(z_t, z_{t+h}) = \gamma_x(h) \gamma_y(h) + \gamma_x(h) + \gamma_y(h) \neq 0 \quad \forall h \neq 0$$

$\{z_t\}$ is not a WN

Remark: Note that $E z_t \neq 0$ and WN is a 0 mean process

So we can simply say that $E z_t = 1 \Rightarrow z_t \not\sim \text{WN}$

$$(6) \quad \phi(B) X_t = \theta(B) \epsilon_t$$

$$\phi(B) = 1 - .5B; \quad \theta(B) = 1 + .4B$$

Root of $\phi(z) = 0$ is $\frac{1}{.5} \Rightarrow \{X_t\}$ is stationary & causal

Root of $\theta(z) = 0$ is $-\frac{1}{.4} \Rightarrow \{X_t\}$ is invertible

AR(∞) representation of $\{X_t\}$

$$\theta(B)^{-1} \phi(B) X_t = \epsilon_t$$

$$\psi(B) X_t = \epsilon_t; \quad \psi(B) = \psi_0 + \psi_1 B + \dots$$

$$\Rightarrow \theta(B)^{-1} \phi(B) = \psi(B)$$

$$\text{i.e. } \phi(B) = \theta(B) \psi(B)$$

$$(1 - .5B) = (1 + .4B)(\psi_0 + \psi_1 B + \psi_2 B^2 + \dots)$$

Comparing coeffs of B^j we get:

$$\psi_0 = 1$$

$$\psi_1 = -(.5 + .4)$$

$$\psi_2 = -.4 \psi_1 = -.4(-.5 + .4)$$

$$\psi_3 = (-.4)^2 (-.5 + .4)$$

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$$\psi_j = (-.4)^{j-1} (-(.5 + .4))$$

$$\sum_{j=0}^{\infty} \psi_j X_{t-j} = \epsilon_t$$

$$X_t = - \sum_{j=1}^{\infty} \psi_j X_{t-j} + \epsilon_t \leftarrow \text{AR}(\infty)$$

Why the causal MA(∞) representation can be obtained

$$(7) \quad X_t \sim \text{MA}(\infty)$$

$$X_t = \theta(B) \epsilon_t \quad - \quad \epsilon_t \sim \text{WN}(0, \sigma^2), \quad \text{~~indep~~}$$

$$Y_t = \sum_{j=0}^{\infty} a_j X_{t-j} + \epsilon_t: \quad \sum_j |a_j| < \infty$$

$$Y_t = P_t + \epsilon_t, \text{ say}$$

$$P_t = \sum_{j=0}^{\infty} a_j X_{t-j} = a(B) X_t \leftarrow \text{filtered process}$$

$$a(B) = a_0 + a_1 B + a_2 B^2 + \dots$$

Note that since ϵ_t & $\{X_t\}$ are indep, P_t & ϵ_t are indep

$$\text{Also, } \gamma_Y(h) = E Y_t Y_{t+h}$$

$$= E \left(\sum_{l=0}^{\infty} a_l X_{t-l} \right) \left(\sum_{k=0}^{\infty} a_k X_{t+h-k} \right) + E \epsilon_t \epsilon_{t+h}$$

$$= \sum_l \sum_k a_l a_k \gamma_X(h-k+l) + \gamma_\epsilon(h)$$

$$\text{i.e. } \gamma_Y(h) = \gamma_P(h) + \gamma_\epsilon(h)$$

$$A \subset GF \text{ of } \{y_t\}$$

$$g_y(z) = \sum_{j=-r}^r z^j \gamma_y(j)$$

$$= \sum_{j=-r}^r z^j (\gamma_p(j) + \gamma_e(j))$$

$$= g_p(z) + g_e(z)$$

$$= a(z) a(\bar{z}^{-1}) g_x(z) + g_e(z)$$

$$\Rightarrow g_y(z) = \left(\sum_l a_l z^l \right) \left(\sum_k a_k \bar{z}^{-k} \right) \theta(z) \theta(\bar{z}^{-1}) \sigma^2 + 1$$

Take $a_l = 2, a_0 = 1, a_1 = 2 \wedge a_j = 0 \forall j > 1$

$$X_t = (1 + \theta_1 B + \theta_2 B^2) \epsilon_t$$

$$g_y(z) = (1 + 2\bar{z}^{-1})(1 + 2z) g_x(z) + 1$$

$$= (1 + 2\bar{z}^{-1} + 2z + 4) g_x(z) + 1$$

$$g_x(z) = (1 + \theta_1 z + \theta_2 z^2)(1 + \theta_1 \bar{z}^{-1} + \theta_2 \bar{z}^{-2}) \sigma^2$$

$$= \left[(1 + \theta_1^2 + \theta_2^2) + (\theta_1 + \theta_1 \theta_2)(z + \bar{z}^{-1}) + \theta_2(z^2 + \bar{z}^{-2}) \right] \sigma^2$$

$$\gamma_y(2) = \text{Coeff of } z^2 \text{ in } g_y(z)$$

⑧

$$X_t = \frac{1}{2} X_{t-1} + \epsilon_t$$

$$Y_t = \delta_t - \delta_{t-1} - \delta_{t-2}$$

$$\epsilon_t \sim WN(0, \sigma^2)$$

$$\delta_t \sim WN(0, \sigma^2)$$

> indep.

Note that $X_t = \sum_{j=0}^{\infty} \left(\frac{1}{2}\right)^j \epsilon_{t-j}$ - f'n of $\{\epsilon_t\}$ seq

A $Y_t = \delta_t - \delta_{t-1} - \delta_{t-2}$ - f'n of $\{\delta_t\}$ seq

$\Rightarrow X_t$ & Y_t are indep

$$Z_t = X_t - X_{t-1} + Y_t \quad (*)$$

$$\begin{aligned} \gamma_Z(h) &= \text{Cov}(X_t - X_{t-1} + Y_t, X_{t+h} - X_{t+h-1} + Y_{t+h}) \\ &= 2\gamma_X(h) - \gamma_X(h-1) - \gamma_X(h+1) + \gamma_Y(h) \end{aligned}$$

A CF of $\{Z_t\}$:

$$g_Z(y) = \sum_{h=-\infty}^{\infty} y^h \gamma_Z(h)$$

$$= \sum_{h=-\infty}^{\infty} y^h (2\gamma_X(h) - \gamma_X(h-1) - \gamma_X(h+1) + \gamma_Y(h)) y^h$$

$$= 2g_X(y) - y g_X(y) - y^{-1} g_X(y) + g_Y(y)$$

i.e. $g_Z(y) = (1-y)(1-y^{-1})g_X(y) + g_Y(y)$

Note that \nearrow can be obtained from (*) directly by using filtering argument

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$$X_t = \epsilon_t + \epsilon_{t-1}; \quad \epsilon_t \sim WN(0, \sigma^2); \quad |\theta| > 1$$

$$Y_t = \sum_{j=0}^{\infty} \left(-\frac{1}{\theta}\right)^j X_{t-j}$$

$$\Rightarrow \left(1 + \frac{1}{\theta} B\right) Y_t = X_t \quad \left|\frac{1}{\theta}\right| < 1 \Rightarrow \left(1 + \frac{1}{\theta}\right)^{-1} \text{ exists}$$

ACGF of L.H.S.

$$\left(1 + \frac{z}{\theta}\right) \left(1 + \frac{\bar{z}^{-1}}{\theta}\right) g_Y(z)$$

$$= \text{ACGF of r.h.s} = (1 + \theta z)(1 + \theta \bar{z}^{-1}) \sigma^2$$

$$\Rightarrow g_Y(z) = \sigma^2 \frac{(1 + \theta z)(1 + \theta \bar{z}^{-1})}{\left(1 + \frac{z}{\theta}\right) \left(1 + \frac{\bar{z}^{-1}}{\theta}\right)}$$

$$= \sigma^2 \frac{(1 + \theta z)(1 + \theta \bar{z}^{-1})}{(z + \theta)(z\theta + 1) / \theta^2 z}$$

$$= \sigma^2 \frac{(1 + \theta z)(z + \theta) / \theta z}{(z + \theta)(z\theta + 1) / \theta^2 z} = \theta^2 \sigma^2$$

$$\text{i.e. } g_Y(z) = \theta^2 \sigma^2 \leftarrow \text{const}$$

$$\Rightarrow Y \sim WN(0, \theta^2 \sigma^2)$$