

Note: If  $\Sigma > 0$ , then p.d.f. of  $\underline{X} \sim N_p(\underline{\mu}, \Sigma)$  is

$$f_{\underline{X}}(\underline{x}) = \frac{1}{(2\pi)^{p/2} |\Sigma|^{1/2}} e^{-\frac{1}{2}(\underline{x}-\underline{\mu})' \Sigma^{-1}(\underline{x}-\underline{\mu})} \quad \underline{x} \in \mathbb{R}^p.$$

$p = 2$ ;  $N_2(\underline{\mu}, \Sigma)$   $\underline{\mu} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix}$ ;  $\Sigma = \begin{pmatrix} \sigma_{11} & \sigma_{12} \\ \sigma_{21} & \sigma_{22} \end{pmatrix}$

or  $BVN(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$

$\sigma_{ii} = V(X_i) = \sigma_i^2$  (var)

$\sigma_{12} = \sigma_{21} = \text{Cov}(X_1, X_2)$

$\rho = \frac{\sigma_{12}}{[\sigma_{11} \cdot \sigma_{22}]^{1/2}}$

$\Sigma = \begin{pmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{pmatrix}$

$$f_{\underline{X}}(\underline{x}) = \frac{1}{2\pi \sigma_1 \sigma_2 \sqrt{1-\rho^2}} \exp \left\{ -\frac{1}{2(1-\rho^2)} \left\{ \left( \frac{x_1 - \mu_1}{\sigma_1} \right)^2 + \left( \frac{x_2 - \mu_2}{\sigma_2} \right)^2 - 2\rho \left( \frac{x_1 - \mu_1}{\sigma_1} \right) \left( \frac{x_2 - \mu_2}{\sigma_2} \right) \right\} \right\}$$

$\underline{x} \in \mathbb{R}^2$

$$X_1 \sim N(\mu_1, \sigma_1^2)$$

$$X_2 \sim N(\mu_2, \sigma_2^2)$$

Remark:

Note that if  $X_1$  &  $X_2$  are uncorrelated, i.e.  $\rho = 0$ , then

$$f_{X_1, X_2}(x_1, x_2) = f_{X_1}(x_1) f_{X_2}(x_2)$$

$\Rightarrow X_1$  &  $X_2$  are indep.

(such a conclusion is not true in general;  
i.e. uncorrelated  $\nRightarrow$  independence in general)

For  $\underline{X} \sim N_p(\underline{\mu}, \Sigma)$

$$\underline{X} = \begin{pmatrix} X_1 \\ \vdots \\ X_q \\ \vdots \\ X_{q+1} \\ \vdots \\ X_p \end{pmatrix} = \begin{pmatrix} \underline{X}^{(1)} \\ \vdots \\ \underline{X}^{(2)} \end{pmatrix} \quad \underline{\mu} = \begin{pmatrix} \underline{\mu}^{(1)} \\ \vdots \\ \underline{\mu}^{(2)} \end{pmatrix}$$

$$\Sigma = \begin{pmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{pmatrix} \quad \Sigma_{ii} = \text{cov}(\tilde{X}^{(i)})$$

$\tilde{X}^{(1)}$  &  $\tilde{X}^{(2)}$  are indep iff  $\Sigma_{12} = 0$ .

Conditional dist<sup>n</sup>

$$\tilde{X}^{(1)} | \tilde{X}^{(2)} \sim N_q \left( \tilde{\mu}^{(1)} + \Sigma_{12} \Sigma_{22}^{-1} (\tilde{X}^{(2)} - \tilde{\mu}^{(2)}), \Sigma_{11.2} \right)$$

$$\Sigma_{11.2} = \Sigma_{11} - \Sigma_{12} \Sigma_{22}^{-1} \Sigma_{21}$$

shy

$$\tilde{X}^{(2)} | \tilde{X}^{(1)} \sim N_{p-q} \left( \tilde{\mu}^{(2)} + \Sigma_{21} \Sigma_{11}^{-1} (\tilde{X}^{(1)} - \tilde{\mu}^{(1)}), \Sigma_{22.1} \right)$$

$$\Sigma_{22.1} = \Sigma_{22} - \Sigma_{21} \Sigma_{11}^{-1} \Sigma_{12}$$

For  $p=2$ , i.e.  $N_2(\mu_1, \mu_2, \sigma_1^2, \sigma_2^2, \rho)$

$$X_1 | X_2 \sim N_1 \left( \mu_1 + (\rho \sigma_1 \sigma_2) (\sigma_2^2)^{-1} (x_2 - \mu_2), \sigma_1^2 - \frac{(\rho \sigma_1 \sigma_2)^2}{\sigma_2^2} \right)$$

$$\Sigma = \begin{pmatrix} \sigma_1^2 & \rho \sigma_1 \sigma_2 \\ \rho \sigma_1 \sigma_2 & \sigma_2^2 \end{pmatrix}$$

$$\text{i.e. } X_1 | X_2 \sim N_1 \left( \mu_1 + \rho \frac{\sigma_1}{\sigma_2} (x_2 - \mu_2), \sigma_1^2 (1 - \rho^2) \right)$$

shy

$$X_2 | X_1 \sim N_1 \left( \mu_2 + \rho \frac{\sigma_2}{\sigma_1} (x_1 - \mu_1), \sigma_2^2 (1 - \rho^2) \right)$$

Note : If  $\tilde{X} \sim N_p(\underline{\mu}, \Sigma)$  ;  $\Sigma > 0$

$$\text{Then } (\underline{X} - \underline{\mu})' \Sigma^{-1} (\underline{X} - \underline{\mu}) \sim \chi_p^2$$

Proof of  $\tilde{X}^{(1)}$  &  $\tilde{X}^{(2)}$  indep iff  $\Sigma_{12} = 0$

If  $\tilde{X}^{(1)}$  &  $\tilde{X}^{(2)}$  are indep then for any  $X_i$  in  $\tilde{X}^{(1)}$  and  $X_j$  in  $\tilde{X}^{(2)}$ ,

$$X_i \text{ \& } X_j \text{ are indep} \Rightarrow \text{cov}(X_i, X_j) = 0 \Rightarrow \Sigma_{12} = 0$$

Alternately suppose  $\Sigma_{12} = 0$ , then

$$\Sigma = \begin{pmatrix} \Sigma_{11} & 0 \\ 0 & \Sigma_{22} \end{pmatrix}; \quad \Sigma^{-1} = \begin{pmatrix} \Sigma_{11}^{-1} & 0 \\ 0 & \Sigma_{22}^{-1} \end{pmatrix}$$

$$\& |\Sigma| = |\Sigma_{11}| |\Sigma_{22}|$$

$$f_{\tilde{X}}(\tilde{x}^{(1)}, \tilde{x}^{(2)}) = \frac{1}{(2\pi)^{p/2} |\Sigma_{11}|^{1/2} |\Sigma_{22}|^{1/2}} \exp\left(-\frac{1}{2}(\tilde{x} - \underline{\mu})' \begin{pmatrix} \Sigma_{11}^{-1} & 0 \\ 0 & \Sigma_{22}^{-1} \end{pmatrix} (\tilde{x} - \underline{\mu})\right)$$

$$= \left( \frac{1}{(2\pi)^{q/2} |\Sigma_{11}|^{1/2}} \exp\left(-\frac{1}{2}(\tilde{x}^{(1)} - \underline{\mu}^{(1)})' \Sigma_{11}^{-1} (\tilde{x}^{(1)} - \underline{\mu}^{(1)})\right) \right)$$

$$\times \left( \frac{1}{(2\pi)^{\frac{p-q}{2}} |\Sigma_{22}|^{1/2}} \exp\left(-\frac{1}{2}(\tilde{x}^{(2)} - \underline{\mu}^{(2)})' \Sigma_{22}^{-1} (\tilde{x}^{(2)} - \underline{\mu}^{(2)})\right) \right)$$

$$= f_{\tilde{X}^{(1)}}(\tilde{x}^{(1)}) f_{\tilde{X}^{(2)}}(\tilde{x}^{(2)})$$

$$\tilde{X}^{(1)} \sim N_q(\underline{\mu}^{(1)}, \Sigma_{11}); \quad \tilde{X}^{(2)} \sim N_{p-q}(\underline{\mu}^{(2)}, \Sigma_{22})$$

$\Rightarrow \tilde{X}^{(1)}$  &  $\tilde{X}^{(2)}$  are indep.

$$\underline{X} \sim N_p(\underline{\mu}, \Sigma) ; \Sigma > 0$$

Derivation of conditional dist<sup>n</sup>

Let

$$\underline{Z} = \underbrace{\begin{bmatrix} I_q & -\Sigma_{12}\Sigma_{22}^{-1} \\ 0 & I_{p-q} \end{bmatrix}}_A \begin{pmatrix} \underline{X}^{(1)} - \underline{\mu}^{(1)} \\ \underline{X}^{(2)} - \underline{\mu}^{(2)} \end{pmatrix} = \begin{pmatrix} \underline{X}^{(1)} - \underline{\mu}^{(1)} - \Sigma_{12}\Sigma_{22}^{-1}(\underline{X}^{(2)} - \underline{\mu}^{(2)}) \\ \underline{X}^{(2)} - \underline{\mu}^{(2)} \end{pmatrix} \sim N_p(\underline{0}, A\Sigma A')$$

$$\begin{aligned} A\Sigma A' &= \begin{bmatrix} I_q & -\Sigma_{12}\Sigma_{22}^{-1} \\ 0 & I_{p-q} \end{bmatrix} \begin{bmatrix} \Sigma_{11} & \Sigma_{12} \\ \Sigma_{21} & \Sigma_{22} \end{bmatrix} \begin{bmatrix} I_q & 0 \\ -\Sigma_{22}^{-1}\Sigma_{21} & I_{p-q} \end{bmatrix} \\ &= \begin{bmatrix} \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21} & 0 \\ 0 & \Sigma_{22} \end{bmatrix} \begin{bmatrix} I_q & 0 \\ -\Sigma_{22}^{-1}\Sigma_{21} & I_{p-q} \end{bmatrix} \\ \Sigma_{11.2} &\rightarrow \begin{bmatrix} \Sigma_{11} - \Sigma_{12}\Sigma_{22}^{-1}\Sigma_{21} & 0 \\ 0 & \Sigma_{22} \end{bmatrix} \end{aligned}$$

$\Rightarrow (\underline{X}^{(1)} - \underline{\mu}^{(1)} - \Sigma_{12}\Sigma_{22}^{-1}(\underline{X}^{(2)} - \underline{\mu}^{(2)})) \text{ \& } (\underline{X}^{(2)} - \underline{\mu}^{(2)}) \text{ are indep}$

$\Rightarrow (\underline{X}^{(1)} - \underline{\mu}^{(1)} - \Sigma_{12}\Sigma_{22}^{-1}(\underline{X}^{(2)} - \underline{\mu}^{(2)})) \mid \underline{X}^{(2)} \text{ and unconditional dist<sup>n</sup> are identical}$

$\Rightarrow \underline{X}^{(1)} - \underline{\mu}^{(1)} - \Sigma_{12}\Sigma_{22}^{-1}(\underline{X}^{(2)} - \underline{\mu}^{(2)}) \mid \underline{X}^{(2)} \sim N_q(\underline{0}, \Sigma_{11.2})$

$\Rightarrow \underline{X}^{(1)} \mid \underline{X}^{(2)} \sim N_q(\underline{\mu}^{(1)} + \Sigma_{12}\Sigma_{22}^{-1}(\underline{X}^{(2)} - \underline{\mu}^{(2)}), \Sigma_{11.2})$

# joint moment generating $f^n$

$$\tilde{X} = (X_1, \dots, X_p)'$$

joint m.g.f.

$$M_{\tilde{X}}(\tilde{t}) = E(e^{\tilde{t}' \tilde{X}}) = E(e^{t_1 X_1 + \dots + t_p X_p})$$

$\downarrow$   
 $(t_1, \dots, t_p)$

provided the expectation exists  
in some nbd of  $\tilde{0}_{p \times 1}$

$$= \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} e^{\sum_{i=1}^p t_i x_i} f_{\tilde{X}}(\underline{x}) dx_1 \dots dx_p.$$

for continuous case

$$= \sum_{x_1} \dots \sum_{x_p} e^{\sum t_i x_i} P(\tilde{X} = \underline{x})$$

for discrete case