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urn containing  $M$  white &  $N-M$  balls of some other color

$n$  balls drawn at a time or one after another without replacement

$X$  : number of white balls in the sample

minimum value of  $X$ :  $\max(0, n - (N - M))$

maximum value of  $x$ :  $\min(n, M)$

$$P(X=x) = \begin{cases} \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}, & x = \max(0, n-(N-M)), \dots, \min(n, M) \\ 0, & \text{o.w.} \end{cases}$$

Note: Same setup with replacement sampling -  $\text{Bin}(n, \frac{M}{N})$

m.g.f.  $E(e^{tx}) = \sum_{x=0}^n e^{tx} \frac{\binom{M}{x} \binom{N-M}{n-x}}{\binom{N}{n}}$   
(for  $n > M, n > N-M$ )

$$E(x) = 2 \frac{1}{2}$$

$$V(x) = n \frac{M}{Z} \left(1 - \frac{M}{Z}\right) \left(\frac{Z-n}{Z-1}\right)$$

use  $\sum_{k=0}^n \binom{a}{k} \binom{b}{n-k} = \binom{a+b}{n}$  to derive  $E(X)$ ,  
 $E(X(X-1))$  to get  $V(X)$

## VII Poisson

$$X \sim P(\lambda) \quad \lambda > 0$$

$$P(X=x) = \begin{cases} \frac{e^{-\lambda} \lambda^x}{x!} & ; \quad x = 0, 1, 2, \dots \\ 0 & ; \quad \text{o/w} \end{cases}$$

$$\begin{aligned} \text{m.g.f. : } E(e^{tx}) &= \sum_{x=0}^{\infty} \frac{e^{tx} \lambda^x}{x!} e^{-\lambda} \\ &= e^{\lambda e^t} e^{-\lambda} = e^{-\lambda(1-e^t)} \end{aligned}$$

$$E(X) = V(X) = \lambda$$

Note: Poisson dist<sup>n</sup> is applicable to model count of events, change of states, failures etc.

Note: Poisson approximation to Binomial

$$X \sim \text{Bin}(n, p)$$

$$\begin{aligned} P(X=x) &= \binom{n}{x} p^x (1-p)^{n-x} \\ &= \frac{n(n-1)\dots(n-x+1)}{x!} p^x (1-p)^{n-x} \\ &= \frac{n}{n} \cdot \frac{n-1}{n} \dots \frac{(n-x+1)}{n} \frac{(np)^x}{x!} (1-p)^n (1-p)^{-x} \end{aligned}$$

As  $n \rightarrow \infty, p \rightarrow 0 \ni np = \lambda$  (fixed)

$$P(X=x) \rightarrow \frac{\lambda^x}{x!} e^{-\lambda}$$

## Some standard continuous distributions

### (I) Uniform dist<sup>n</sup>

$$X \sim U[a, b]$$

$$\text{p.d.f. } f(x) = \begin{cases} \frac{1}{b-a}, & a \leq x \leq b \\ 0, & \text{o/w.} \end{cases}$$

$$\text{d.f. } F(x) = \begin{cases} 0, & x < a \\ \frac{x-a}{b-a}, & a \leq x \leq b \\ 1, & x > b \end{cases}$$

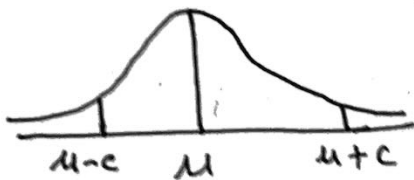
$$\text{m.g.f. } M_X(t) = \frac{1}{b-a} \int_a^b e^{tx} dx = \begin{cases} \frac{e^{tb} - e^{ta}}{t(b-a)} & t \neq 0 \\ 1 & t = 0 \end{cases}$$

### ii Normal dist<sup>n</sup>

$$X \sim N(\mu, \sigma^2)$$

$$f(x) = \frac{1}{\sqrt{2\pi} \sigma} \exp\left(-\frac{1}{2\sigma^2} (x-\mu)^2\right); \quad x \in \mathbb{R}$$

$\mu \in \mathbb{R}, \sigma > 0$



Dist<sup>n</sup> is symmetric around  $\mu \forall (\mu, \sigma)$

$$P(X \leq \mu - c) = P(X \geq \mu + c)$$

$$f(\mu - c) = f(\mu + c) \quad \forall c$$

$$\text{m.g.f. } M_X(t) = E(e^{tx}) = \int_{-\infty}^{\infty} e^{tx} \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2\sigma^2} (x-\mu)^2} dx$$

$$\frac{x-\mu}{\sigma} = z; \quad dx = \sigma dz$$

$$\begin{aligned} M_X(t) &= \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{t(\mu+\sigma z)} e^{-z^2/2} dz \\ &= \frac{e^{t\mu}}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(z^2 - 2t\sigma z + t^2\sigma^2 - t^2\sigma^2)} dz \\ &= \frac{e^{t\mu}}{\sqrt{2\pi}} e^{t^2\sigma^2/2} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(z-t\sigma)^2} dz \\ &= e^{t\mu + t^2\sigma^2/2} \quad t \in (-\infty, \infty) \end{aligned}$$

$$E(X) = \left. \frac{\partial}{\partial t} M_X(t) \right|_{t=0} = e^{t\mu + t^2\sigma^2/2} (\mu + t\sigma^2) \Big|_{t=0} = \mu$$

$$\begin{aligned} E(X^2) &= \left. \frac{\partial^2}{\partial t^2} M_X(t) \right|_{t=0} = e^{t\mu + t^2\sigma^2/2} (\sigma^2) + (\mu + t\sigma^2) e^{t\mu + t^2\sigma^2/2} (\mu + t\sigma^2) \Big|_{t=0} \\ &= \sigma^2 + \mu^2 \end{aligned}$$

$$E(X^2) = \mu^2 + \sigma^2$$

$$V(X) = \sigma^2$$

Standard normal dist<sup>n</sup>  $N(0,1)$

$$Z \sim N(0,1) \quad \phi(z) = f_Z(z) = \frac{1}{\sqrt{2\pi}} e^{-z^2/2} \quad z \in (-\infty, \infty)$$



$$\Phi(z) = P(Z \leq z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-z^2/2} dz \quad (\text{values are tabulated})$$

$$\Phi(-z) = P(Z \leq -z) = P(Z \geq z)$$

$$= 1 - \Phi(z)$$

$$\Phi(-z) + \Phi(z) = 1 \quad ; \quad \Phi(0) = \frac{1}{2}$$

Note: If  $X \sim N(\mu, \sigma^2)$ , then  $\frac{X-\mu}{\sigma} \sim N(0, 1)$

$$\text{Let } Z = \frac{X-\mu}{\sigma}$$

$$\text{d.f. of } Z: P(Z \leq z) = P\left(\frac{X-\mu}{\sigma} \leq z\right)$$

$$= P(X \leq \mu + \sigma z)$$

$$= \int_{-\infty}^{\mu + \sigma z} \frac{1}{\sqrt{2\pi} \sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2} dx$$

$$\frac{x-\mu}{\sigma} = y; \quad = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^z e^{-y^2/2} dy = \Phi(z)$$

$$\text{i.e. } Z \sim N(0, 1)$$

Alt: use m.g.f.