DD2434 - Machine Learning, Advanced Course Assignment 1A

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November 2023



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1 Exponential Family

Question 1.1

$$p(x|\theta) = h(x) \exp(\eta(\theta) \cdot T(x) - A(\eta))$$

$$= h(x) \exp(\eta(\lambda) \cdot T(x) - A(\eta(\lambda)))$$

$$= h(x) \exp(\log \lambda \cdot x - A(\log \lambda))$$

$$= h(x) \exp(\log \lambda \cdot x - \lambda)$$

$$= h(x) \exp(\log \lambda \cdot x) \exp(-\lambda)$$

$$= e^{-\lambda} \frac{\lambda^x}{x!}$$
(1)

We can see that the distribution correspond to a Poisson distribution of parameter λ .

Question 1.2

$$p(x|\theta) = h(x) \exp(\eta(\theta) \cdot T(x) - A(\eta))$$

$$= \exp(\eta([\alpha, \beta]) \cdot [\log x, x] - A(\alpha - 1, -\beta))$$

$$= \exp([\alpha - 1, -\beta] \cdot [\log x, x] - \log \Gamma(\alpha) + \alpha \log(\beta))$$

$$= \exp((\alpha - 1) \log x - \beta x - \log \Gamma(\alpha) + \alpha \log(\beta))$$

$$= \frac{\beta^{\alpha}}{\Gamma(\alpha)} x^{\alpha - 1} e^{-\beta x}$$
(2)

We can see that the distribution correspond to a Gamma distribution of parameters α and β .

Question 1.3

$$p(x|\theta) = h(x) \exp(\eta(\theta) \cdot T(x) - A(\eta))$$

$$= \frac{\exp(\eta([\mu, \sigma^{2}]) \cdot [x, x^{2}] - A(\eta([\mu, \sigma^{2}])))}{\sqrt{2\pi}}$$

$$= \frac{\exp([\frac{\mu}{\sigma^{2}}, -\frac{1}{2\sigma^{2}}] \cdot [x, x^{2}] - A([\frac{\mu}{\sigma^{2}}, -\frac{1}{2\sigma^{2}}]))}{\sqrt{2\pi}}$$

$$= \frac{\exp(\frac{\mu x}{\sigma^{2}} - \frac{x^{2}}{2\sigma^{2}} - \frac{\mu^{2}}{2\sigma^{2}} - \log \sigma)}{\sqrt{2\pi}}$$

$$= \frac{\exp(-\frac{(x-\mu)^{2}}{2\sigma^{2}})}{\sigma\sqrt{2\pi}}$$
(3)

We can see that the distribution correspond to a Normal distribution of parameters μ and σ^2 .

Question 1.4

$$p(x|\theta) = h(x) \exp(\eta(\theta) \cdot T(x) - A(\eta))$$

$$= 2 \exp(\eta(\lambda) \cdot x - A(\eta(\lambda)))$$

$$= 2 \exp(-\lambda x - A(-\lambda))$$

$$= 2 \exp\left(-\lambda x + \log\left(\frac{\lambda}{2}\right)\right)$$

$$= \lambda e^{-\lambda x}$$
(4)

We can see that the distribution correspond to a Exponential distribution of parameter λ .

Question 1.5

$$p(x|\theta) = h(x) \exp(\eta(\theta) \cdot T(x) - A(\eta))$$

$$= \exp(\eta([\psi_1, \psi_2]) \cdot [\log x, \log(1 - x)] - A(\eta([\psi_1, \psi_2])))$$

$$= \exp([\psi_1 - 1, \psi_2 - 1] \cdot [\log x, \log(1 - x)] - A([\psi_1 - 1, \psi_2 - 1]))$$

$$= \exp((\psi_1 - 1) \log x + (\psi_2 - 1) \log(1 - x) - \log \Gamma(\psi_1) - \log \Gamma(\psi_2) + \log \Gamma(\psi_1 + \psi_2))$$

$$= \frac{\Gamma(\psi_1 + \psi_2)}{\Gamma(\psi_1)\Gamma(\psi_2)} x^{\psi_1 - 1} (1 - x)^{\psi_2 - 1}$$
(5)

We can see that the distribution correspond to a Beta distribution of parameters ψ_1 and ψ_2 .

2 Dependencies in a Directed Graphical Model

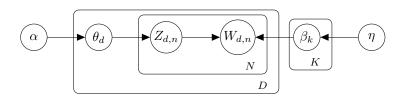
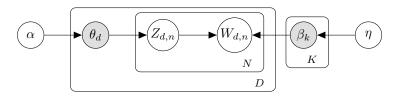


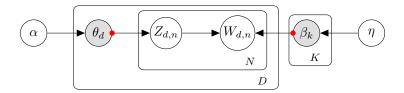
Figure 1: Graphical model of smooth LDA.

Question 2.6

The Bayes net take this form:



Then, if we use the method using the d-separation, we obtain this:



Therefore, we can see that $W_{d,n} \perp W_{d,n+1} | \theta_d, \beta_{1:K}$ is <u>false</u>.

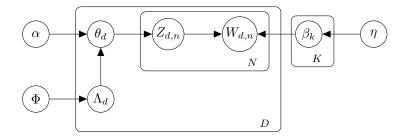


Figure 2: Graphical model of Labeled LDA.

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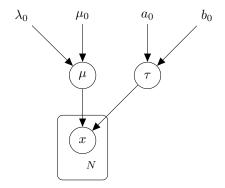


Figure 3: DGM

Question 3.12

In the bishop book, we can see that:

$$p(X|\mu,\tau) = \left(\frac{\tau}{2\pi}\right)^{N/2} \exp\left\{-\frac{\tau}{2} \sum_{n=1}^{N} (x_n - \mu)^2\right\}$$
 (6)

$$p(\mu|\tau) = \mathcal{N}(\mu|\mu_0, (\lambda_0 \tau)^{-1}) \tag{7}$$

$$p(\tau) = \operatorname{Gam}(\tau|a_0, b_0) \tag{8}$$

Then, by using the code in appendix A.1, we obtain:

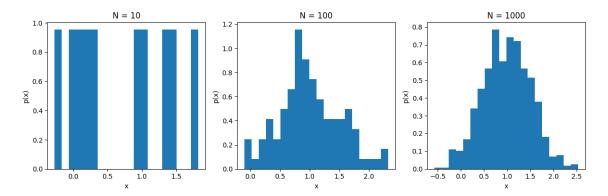


Figure 4: Generated Data.

Question 3.13

Let's find the ML estimates of μ and τ . We know that $\log(q^*(\mu)) = \mathbb{E}_{-\mu}[\log p(X, \mu, \tau)]$. Then, we can write:

$$\log(q^{*}(\mu)) = \mathbb{E}_{-\mu}[\log p(X, \mu, \tau)]$$

$$\stackrel{!}{=} \mathbb{E}_{\tau}[\log p(X|\mu, \tau) + \log p(\mu|\tau)]$$

$$= \mathbb{E}_{\tau} \left[\frac{N}{2} \log \left(\frac{\tau}{2\pi} \right) + \frac{\tau}{2} \sum_{n=1}^{N} (x_{n} - \mu)^{2} + \frac{1}{2} \log \left(\frac{\lambda_{0}\tau}{2\pi} \right) + \frac{\lambda_{0}\tau}{2} (\mu - \mu_{0}) \right]$$

$$\stackrel{!}{=} \frac{\mathbb{E}_{\tau}[\tau]}{2} \left(\lambda_{0}(\mu - \mu_{0}) + \sum_{n=1}^{N} (x_{n} - \mu)^{2} \right)$$

$$= -\frac{\mathbb{E}_{\tau}[\tau]}{2} \left(\lambda_{0}\mu^{2} - 2\lambda_{0}\mu\mu_{0} + \lambda_{0}\mu_{0}^{2} + \sum_{n=1}^{N} x_{n}^{2} - 2\mu \sum_{n=1}^{N} x_{n} + N\mu^{2} \right)$$

$$= -\frac{\mathbb{E}_{\tau}[\tau]}{2} \left((\lambda_{0} + N)\mu^{2} - 2(\lambda_{0}\mu_{0} + \sum_{n=1}^{N} x_{n})\mu + \lambda_{0}\mu_{0}^{2} + \sum_{n=1}^{N} x_{n}^{2} \right)$$

$$= -\frac{\mathbb{E}_{\tau}[\tau](\lambda_{0} + N)}{2} \left(\mu^{2} - 2\mu^{\epsilon} \frac{\lambda_{0}\mu_{0} + \sum_{n=1}^{N} x_{n}}{\lambda_{0} + N} + \frac{\lambda_{0}\mu_{0}^{2} + \sum_{n=1}^{N} x_{n}^{2}}{\lambda_{0} + N} \right)$$

Therefore we can conclude that $q^*(\mu) = \mathcal{N}(\mu|\mu', \lambda'^{-1})$ with :

$$\mu' = \frac{\lambda_0 \mu_0 + \sum_{n=1}^{N} x_n}{\lambda_0 + N} \tag{10}$$

$$\lambda' = (\lambda_0 + N)\mathbb{E}[\tau] \tag{11}$$

And for τ we have :

$$\log(q^{*}(\tau)) = \mathbb{E}_{-\tau}[\log p(X,\mu,\tau)]$$

$$\stackrel{+}{=} \mathbb{E}_{\mu}[\log p(X|\mu,\tau) + \log p(\mu|\tau)] + \log p(\tau)$$

$$\stackrel{+}{=} (a_{0} - 1)\log \tau - b_{0}\tau + \frac{N}{2}\log \tau - \frac{\tau}{2}\mathbb{E}_{\mu}\left[\sum_{n=1}^{N}(x_{n} - \mu)^{2} + \lambda_{0}(\mu - \mu_{0})^{2}\right]$$

$$= (a_{0} + \frac{N}{2} - 1)\log \tau - \left(b_{0} + \frac{1}{2}\mathbb{E}_{\mu}\left[\sum_{n=1}^{N}(x_{n} - \mu)^{2} + \lambda_{0}(\mu - \mu_{0})^{2}\right]\right)\tau$$
(12)

Therefore we can conclude that $q^*(\tau) = \operatorname{Gam}(\tau|a',b')$ with :

$$a' = a_0 + \frac{N}{2} \tag{13}$$

$$b' = b_0 + \frac{1}{2} \mathbb{E}_{\mu} \left[\sum_{n=1}^{N} (x_n - \mu)^2 + \lambda_0 (\mu - \mu_0)^2 \right]$$

$$b' = b_0 + \frac{1}{2} \left(\sum_{n=1}^{N} x_n^2 + N \mathbb{E}_{\mu} [\mu^2] - 2 \mathbb{E}_{\mu} [\mu] \sum_{n=1}^{N} x_n + \lambda_0 \left(\mathbb{E}_{\mu} [\mu^2] + \mu_0^2 - 2\mu_0 \mathbb{E}_{\mu} [\mu] \right) \right)$$
(14)

Question 3.14

The equation (10.24) in the Bishop is the mean-field approximation which is:

$$q(\mu, \tau) = q(\mu)q(\tau) \tag{15}$$

We have also some formulas to detail:

$$\mathbb{E}_{\mu}[\mu] = \mu'$$

$$\mathbb{E}_{\mu}[\mu^{2}] = \frac{1}{\lambda'} + {\mu'}^{2}$$

$$\mathbb{E}_{\tau}[\tau] = \frac{a'}{b'}$$
(16)

The rest of the answer is in the code in appendix A.1.

A Appendix

A.1 Python Code

```
import numpy as np
import scipy.special as sp_spec
import scipy.stats as sp_stats
import numpy.random as np_rand
import matplotlib.pyplot as plt
from sklearn.cluster import KMeans
from sklearn.metrics import adjusted_rand_score
import seaborn as sns
# Question 1.3.12
def generate_data(N: int, mu: float, tau: float) -> np.ndarray:
   return np_rand.normal(mu, tau, N)
def plot_data(X: np.ndarray, ax: plt.Axes) -> None:
   ax.hist(X, bins=20, density=True)
   ax.set_xlabel('x')
   ax.set_ylabel('p(x)')
   ax.set_title(f'N = {len(X)}')
MU = 1
TAU = 0.5
N = [10, 100, 1000]
Xs = [generate_data(n, MU, TAU) for n in N]
fig, axs = plt.subplots(1, 3, figsize=(12, 4))
for i in range(len(Xs)):
   plot_data(Xs[i], axs[i])
plt.tight_layout()
plt.savefig('12_data.png')
# Question 1.3.14
def update_tau(x, mu_prime, lambda_prime, a_0, b_0, mu_0, lambda_0):
    E_mu = mu_prime
   E_mu2 = 1 / lambda_prime + mu_prime ** 2
    a_prime = a_0 + x.shape[0] / 2
   b_{prime} = b_{0} + 0.5 * (np.sum(x ** 2) - 2 * np.sum(x) * E_mu +
                           x.shape[0] * E_mu2 + lambda_0*(E_mu2 + mu_0 ** 2
                               - 2 * E_mu * mu_0))
   return a_prime, b_prime
```

```
def update_mu(x, a_prime, b_prime, mu_0, lambda_0):
    E_tau = a_prime / b_prime
    mu_prime = (lambda_0 * mu_0 + np.sum(x)) / (lambda_0 + x.shape[0])
    lambda_prime = (lambda_0 + x.shape[0])*E_tau
    return mu_prime, lambda_prime
```