LOGARITHMIC TOTAL VARIATION REGULARIZATION FOR CROSS-VALIDATION IN PHOTON-LIMITED IMAGING

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ABSTRACT

In fields such as astronomy and medicine, many imaging modalities operate in the photon-limited realm because of the low photon counts available over a reasonable exposure time. Photon-limited observations are often modeled as the composite of a linear operator, such as a blur or tomographic projection, applied to a scene of interest, followed by Poisson noise draws for each pixel. One method to reconstruct the underlying scene intensity is to minimize a penalized Poisson negative log-likelihood. This paper presents a new model that solves for and regularizes the *logarithm* of the true scene, and focuses on the special case of total variation regularization. This method yields considerable gains when used in conjunction with cross-validation, where weighting of the regularization term is automatically determined using observed data.

Index Terms— Convex optimization, image denoising, image reconstruction, total variation, cross-validation

1. INTRODUCTION

In photon-limited imaging, which arises in astronomy, medical imaging, microscopy, and night vision, an image can be represented using a Poisson model of a discrete number of photons hitting a detector. Observations from this model can be expressed as

$$y \sim \text{Poisson}(Af^*),$$

where $\operatorname{Poisson}(\cdot)$ draws from a Poisson distribution element-wise, $y \in \mathbb{Z}_+^m$ is an m-dimensional vector of observed photon counts, $f^\star \in \mathbb{R}_+^n$ is an m-dimensional vector of the ground truth image of interest (assumed to be piecewise smooth), and $A \in \mathbb{R}_+^{m \times n}$ is a linear projection of the truth to the observation set, such as an operation with an imaging system's point-spread function. From observations y, we are interested in recovering an estimate of f^\star .

One approach to this problem is to solve the following constrained optimization problem:

$$\hat{f}_{\tau} = \underset{f}{\operatorname{arg\,min}} - \log \ p(y|Af) + \tau ||f||_{\text{TV}}$$
 subject to $f \ge 0$, (1)

where p(y|Af) is the likelihood of observing y if the underlying intensity were Af and $\|\cdot\|_{\mathrm{TV}}$ is the total variation (TV) semi-norm with tuning parameter $\tau \geq 0$ [1].

The total variation semi-norm is an operator that computes the magnitude of its argument's discrete gradient. It is used in image recovery because natural images tend to have sparse gradients, as they are often piecewise smooth with sparse edge pixels. The TV norm has been shown to have good signal reconstruction in practice, but

theoretical guarantees about its robustness have only recently been found [2]. There are different types of TV, a popular one being isotropic TV for images [3], which will be used in this paper. For an image $f \in \mathbb{R}_{+}^{n_1 \times n_2}$ (note that this is the unvectorized version, where $n_1 n_2 = n$), it is defined as

$$||f||_{\text{TV}} = \sum_{i=1}^{n_1-1} \sum_{i=1}^{n_2-1} \sqrt{(f_{i,j} - f_{i+1,j})^2 + (f_{i,j} - f_{i,j+1})^2} + \sum_{i=1}^{n_1-1} |f_{i,n_2} - f_{i+1,n_2}| + \sum_{j=1}^{n_2-1} |f_{n_1,j} - f_{n_1,j+1}|.$$

The best value of τ (in terms of either mean squared error or visual quality) differs depending on the data set, and it can be found empirically through trial-and-error parameter tuning, often involving manual choice of the perceived optimal value. One way to automate this process is through cross-validation (CV) via hold out, where we split the data into a training and a validation set. For each candidate value of τ , we solve our optimization problem on the training set, then see how well the resulting \hat{f}_{τ} fits the validation set using the negative Poisson log-likelihood, ultimately choosing the value of τ which makes this measure smallest.

In the context of Poisson reconstruction problems, however, cross-validation has a significant pitfall. In particular, say there is a pixel that has zero counts in the training set and a nonzero number of counts in the validation set, which is a common scenario in low-light settings. Then for moderate values of τ , the estimate \hat{f}_{τ} will be zero in this location, and the value of the corresponding validation measure (the negative log-likelihood) will be infinite, even if \hat{f}_{τ} is an excellent estimate for every other pixel in the image. This value of τ will be discarded in favor of a larger τ which keeps all estimated pixel values away from zero by over-smoothing the estimate. The practical effect of this is a CV-selected τ which is far larger than what would minimize the RMSE. We want to address this practical problem because it would allow fully automated code to be provided to experimentalists.

In this paper, we present an alternative to the optimization formulation described in (1) which (a) yields accurate and visually appealing results, and (b) facilitates cross-validation for selecting tuning parameters. Essentially, instead of estimating f^* directly, we estimate (and regularize) its log, $x^* \equiv \log(f^*)$. This formulation yields a simple convex optimization problem in the denoising setting (where A is the identity matrix), and yields a more complex, non-convex optimization problem in the more general inverse problem setting. We describe approaches to both optimization problems.

2. EXPONENTIAL FORMULATION

Using the relationship $x = \log(f)$, we may write our observation model as

$$y \sim \text{Poisson}(A \exp(x^*)),$$

where the exponentiation operator is applied element-wise. If *A* were the identity matrix, this model would correspond to a generalized linear model [4].

In this Poisson model, the likelihood of observing y is

$$p(y|A\exp(x^{*})) = \prod_{i=1}^{m} \frac{(e_{i}^{T} A \exp(x^{*}))^{y_{i}}}{y_{i}!} \exp(-e_{i}^{T} A \exp(x^{*}))$$

where e_i is the $i^{\rm th}$ unit vector in the standard basis, so the negative Poisson log-likelihood is proportional to

$$F(x) = \sum_{i=1}^{m} e_i^T A \exp(x) - y_i \log(e_i^T A \exp(x)),$$

where we neglect $\log(y_i!)$ because it is constant with respect to x. We may then estimate x^* using this negative log-likelihood function and TV penalty:

$$x_{\tau} = \underset{\tau}{\operatorname{arg \, min}} \quad \Phi(x) \equiv F(x) + \tau ||x||_{\text{TV}},$$

which enforces TV regularization on the *logarithm* of the image of interest.

This negative log-likelihood function, however, is non-convex for general A. To handle this, we adopt the following approach: using the current iterate, we compute an approximation of F(x) which is used to formulate a convex optimization problem that can be solved to yield the next iterate. Known convex optimization techniques can then be used to find the minimizer of this convex surrogate.

The key to our approximation is to use a first-order Taylor series expansion on $h_i(x) \equiv \log(e_i^T A \exp(x))$ in the negative log-likelihood expression. The approximation of $h_i(x)$ centered about the j^{th} iterate x^j , denoted $h_i^j(x)$, is

$$h_i^j(x) \equiv h_i(x^j) + \nabla h_i(x^j)^T (x - x^j)$$

$$= \left[h_i(x^j) - \nabla h_i(x^j)^T x^j \right] + \nabla h_i(x^j)^T x$$

$$= u_i^j + e_i^T B^j x,$$

$$(2)$$

where we define

$$B^{j} \equiv \operatorname{diag}(A \exp(x^{j}))^{-1} A \operatorname{diag}(\exp(x^{j})) \tag{3}$$

$$u^{j} \equiv \log(A\exp(x^{j})) - B^{j}x^{j},\tag{4}$$

and where $diag(\cdot)$ is a function that diagonalizes its input and $log(\cdot)$ is an element-wise logarithm operator.

After the linear approximation, we compute the next iterate x_{τ}^{j+1} by solving the optimization problem

$$x_{\tau}^{j+1} = \underset{x}{\arg\min} \quad \Phi'(x) \equiv F'_{j}(x) + \tau ||x||_{\text{TV}},$$
 (5)

where

$$F'_{j}(x) = \sum_{i=1}^{m} \exp(u_{i}^{j} + e_{i}^{T} B^{j} x) - y_{i}(u_{i}^{j} + e_{i}^{T} B^{j} x), \quad (6)$$

is a convex negative Poisson log-likelihood with gradient

$$\nabla F'_j(x) = (B^j)^T \left[\exp(u + B^j x) - y \right].$$

Note that (6) is convex because the composition of a non-decreasing convex function (i.e. $\exp(\cdot)$) with any convex function (i.e. its linear argument) is convex [5]. Because the TV semi-norm is convex as well, this new objective function is therefore convex.

To solve the subproblem in (5), we use the FISTA algorithm presented by Beck and Teboulle [6]. This algorithm iteratively solves for x_{τ}^{j+1} using a gradient descent step followed by a TV denoising step. With k as an iteration counter for this inner loop and \tilde{x}^k as the subproblem's iterates:

$$s^{k} = w^{k} - \frac{1}{L} \nabla F_{j}'(w^{k})$$

$$\tilde{x}^{k} = \arg\min_{x} \frac{1}{2} \|x - s^{k}\|^{2} + \frac{\tau}{L} \|x\|_{\text{TV}},$$
(7)

where w^k is an update variable introduced in FISTA and specified in Algorithm 1, and L is an upper bound on the Lipschitz constant of $\nabla F'$ (and is also a step-size parameter in practice). The subproblem in (7) can be solved using the TV-based denoising and deblurring framework by Beck and Teboulle in [3].

We refer to our approach as "log-TV reconstruction", which is summarized in Algorithm 1 below.

Algorithm 1 log-TV reconstruction

```
Choose initialization f^0. Set \tau, A, L, and tolerance. Set j=0, x^0=\log(f^0) repeat

Compute B^j and u^j via (3) and (4)
j \leftarrow j+1
Set k=0, \tilde{x}^0=x^j, t^0=1, w^1=x^j repeat
k \leftarrow k+1
if using FISTA then
\tilde{x}^k \leftarrow \text{solution of } (7)
t^k=\frac{1+\sqrt{1+(4t^{k-1})^2}}{2}
w^{k+1}=\tilde{x}^k+(\frac{t^{k-1}-1}{t^k})(\tilde{x}^k-\tilde{x}^{k-1})
else if using monotonic FISTA then
z^k \leftarrow \text{solution of } (7)
t^k=\frac{1+\sqrt{1+(4t^{k-1})^2}}{2}
\tilde{x}^k=\arg\min\{\Phi'(q): q=z^k, \tilde{x}^{k-1}\}
w^{k+1}=\tilde{x}^k+(\frac{t^{k-1}}{t^k})(z^k-\tilde{x}^k)+(\frac{t^{k-1}-1}{t^k})(\tilde{x}^k-\tilde{x}^{k-1})
end if
until \|\tilde{x}^k-\tilde{x}^{k-1}\|_2/\|\tilde{x}^k\|_2 \leq \text{tolerance}
x^j\leftarrow\tilde{x}^k
until \|x^j-x^{j-1}\|_2/\|x^j\|_2 \leq \text{tolerance}
```

2.1. Stationary points

Although we are trying to solve a non-convex objective, the convex Taylor series approximation ensures stationarity of gradient-step iterations if at a local minimum. This is due to the fact that the gradient of our approximation matches the gradient of the original objective at the current iterate. To see this, note that using the notation of

(2), we have that $h_i^j(x^j) = h_i(x^j)$ and $\nabla h_i^j(x^j) = \nabla h_i(x^j)$. Examining the gradients of the objective and our approximation, we respectively have

$$\nabla F(x) = \sum_{i=1}^{m} \left[\exp(h_i(x)) - y_i \right] \nabla h_i(x)$$
 (8)

$$\nabla F_j'(x) = \sum_{i=1}^m \left[\exp(h_i^j(x)) - y_i \right] \nabla h_i^j(x), \tag{9}$$

and therefore our approximation guarantees $\nabla F(x^j) = \nabla F_j'(x^j)$. If we are at a stationary point for the original problem, we will have $0 \in \nabla F(x^j) + \tau \partial \|x^j\|_{\mathrm{TV}}$, and because (8) and the approximation in (9) are equivalent at $x = x^j$, we have that $0 \in \nabla F_j'(x^j) + \tau \partial \|x^j\|_{\mathrm{TV}}$. Hence, the point x^j also optimizes the subproblem with the approximate objective, thus guaranteeing stationarity of our method.

2.2. Special case: denoising

When the linear projection A is the identity matrix, our problem reduces to a Poisson denoising problem. In this special case, u=0 and $B=I_n$, meaning the outer Taylor series expansion loop is unnecessary because F(x)=F'(x). This denoising problem is convex, so only the FISTA inner loop is needed to solve for x.

3. CROSS-VALIDATION (CV)

Performance depends heavily on the choice of τ , and selecting this value accurately is important. τ can be found by manually tuning parameters and adjusting based on trial-and-error, but it is more advantageous to select this automatically using cross-validation via hold out, as presented in [7]. In this setup, we assume that each photon count is independent of all others, so a new model can be written where a certain fraction p of the observations are used as a training set y^{train} , and the remaining fraction 1-p are used as a validation set y^{val} :

$$y^{\text{train}} \sim \text{Poisson}(pAf)$$

 $y^{\text{val}} \sim \text{Poisson}((1-p)Af).$

To separate photons into these two groups, we take the original set of observations and draw training data from an element-wise Binomial distribution so that $y^{\text{train}} \sim \text{Binomial}(y,p)$, where the validation set is the complement $y^{\text{val}} = y - y^{\text{train}}$. For a collection of T test values of τ , we can use the training set to compute

$$\hat{f}_{\tau} = \underset{f}{\operatorname{arg\,min}} - \log \ p(y^{\operatorname{train}} | pAf) + \tau ||f||_{\text{TV}},$$

For each $\tau \in T$, the negative Poisson log-likelihood of the validation set is calculated, and the τ corresponding to the minimum is determined to be the best tuning parameter:

$$\hat{\tau} = \underset{\tau \in T}{\operatorname{arg\,min}} - \log \ p(y^{\operatorname{val}}|(1-p)A\hat{f}_{\tau}).$$

The final estimate is then $\hat{f}^{\text{TV}} := \hat{f}_{\hat{\tau}}$. Similarly, we can perform cross-validation using log-TV regularization via

$$\hat{x}_{\tau} = \underset{x}{\operatorname{arg\,min}} - \log \ p(y^{\operatorname{train}} | pA \exp(x)) + \tau ||x||_{\text{TV}}$$

$$\hat{\tau} = \underset{\tau \in T}{\operatorname{arg\,min}} - \log \ p(y^{\operatorname{val}} | (1 - p)A \exp(\hat{x}_{\tau}))$$

$$\hat{f}^{\log \text{TV}} := \exp(\hat{x}_{\hat{\tau}}).$$

4. EXPERIMENTAL RESULTS

To demonstrate the performance of our algorithm, we used the Shepp-Logan head phantom available in Matlab via the Image Processing Toolbox; the grayscale image used was size 128×128 , scaled to have a mean 4.432 photons per pixel. Performance was evaluated visually and by the root mean squared error (RMSE) percentage given by $100 \times \|\hat{f} - f^*\|_2 / \|f^*\|_2$; note that the truth was known in these simulations for the sake of evaluation, but is not known in a real setting. We compare our log-TV method with the TV reconstruction described in [1] that solves (1), where both methods were run until the relative difference between consecutive iterates was $\|\Delta f\|_2 / \|f\|_2 \le 10^{-5}$ and $\|\Delta x\|_2 / \|x\|_2 \le 10^{-5}$, respectively.

4.1. Deblurring

Results for the deblurring case can be seen in Figure 1, where all of the reconstructions are from the same Poisson noise realization and are displayed on the same color scale as the truth. The truth was convolved with a Gaussian blur matrix and was subjected to Poisson noise, and these resulting observations had a mean value of 4.426 photons per pixel. The TV reconstruction was run using the monotonic option, while log-TV reconstruction used monotonic FISTA with L=100.

TV and log-TV reconstructions were used to estimate \hat{f} clair-voyantly by choosing values of τ minimizing the RMSE. It can be seen that log-TV yields a reconstruction that more accurately matches the true intensities; also, the TV reconstruction has more noise artifacts in both the background and foreground, even estimating some pixels within the main body as having small value. Cross-validation (with p=0.9 and the candidate collection of τ values considered spaced 0.02 apart for TV and 0.001 apart for log-TV) shows that TV offers a poorer reconstruction that is very oversmoothed; log-TV with CV, on the other hand, gives a better estimate that is comparable to the clairvoyant log-TV result. The RMSE averaged over 50 different Poisson noise realizations for these comparisons can be seen in Table 1, where it is seen that log-TV has less error than TV.

The huge practical benefits of log-TV reconstruction can be seen in the cross-validation behavior. As discussed in Section 3, CV selects the optimal τ value based on the minimum negative Poisson log-likelihood from various validation sets. A more accurate method would be to select the value that minimizes the RMSE, but this cannot be the criteria in a real setting because the truth is unknown. Figure 2 shows that for TV, there is a large difference between the τ chosen via the validation set negative log-likelihood and via the training set RMSE known in simulation, while the difference is much smaller for log-TV reconstruction. This highlights the advantage that log-TV reconstruction with CV has much less over-smoothing because of its smaller τ value.

	TV reconstruction	log-TV reconstruction
Tuning $ au$	31.66%	28.48%
${ m CV}~ au$	43.96%	28.49%

Table 1: RMSE comparison for deblurring, averaged over 50 trials

4.2. Denoising

The results of the denoising experiments can be seen in Figure 3, where the reconstructions are from the same Poisson noise realiza-

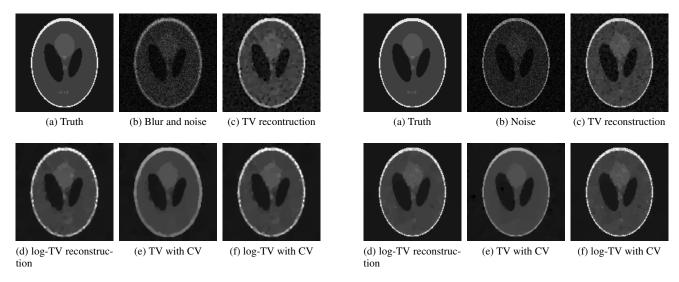


Fig. 1: Deblurring results

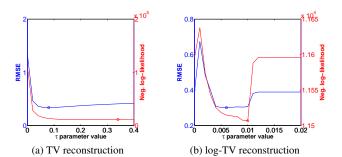


Fig. 2: Cross-validation behavior (deblurring)

0.3
0.3
0.28
0.4
0.20
0.5
1.15
0.10
0.20
0.10
0.005
0.01
0.015
0.029
0.10
0.005
0.01
0.015
0.029
0.10
0.005
0.01
0.015
0.029
0.10
0.005
0.01
0.015
0.029

Fig. 3: Denoising results

Fig. 4: Cross-validation behavior (denoising)

tion and are shown on the same scale as the truth. After the truth was subjected to only Poisson noise, the observation mean was $4.450\,$ photons per pixel. TV was run using monotonic iterations, while log-TV reconstruction used non-monotonic FISTA iterations with L=200.

In the clairvoyant, manual parameter tuning case, log-TV performed better than TV, yielding a smoother reconstruction that matched the true intensities better. For TV, noise artifacts are seen in the background and certain pixels within the body are estimated to be near-zero. For cross-validation (with p=0.9 and the candidate collection of τ values considered spaced 0.05 apart for TV and 0.001 apart for log-TV), TV gives an over-smoothed reconstruction due to a large selected parameter value; its intensities are much lower than that of the truth, especially around the outer edge of the phantom. The log-TV reconstruction is also comparable to its clairvoyantly-chosen counterpart, which cannot be said about TV. The RMSE averaged over 50 trials for both comparisons can be seen in Table 2, where it is seen that log-TV reconstruction empirically performs better than TV.

Cross-validation again shows the practical benefits of this new formulation. As seen in Figure 4, there is a large difference between the τ value chosen via the validation set negative log-likelihood for TV and that of the training set reconstruction RMSE; log-TV reconstruction chose a more reasonable τ that avoids over-smoothing.

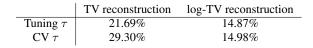


Table 2: RMSE comparison for denoising, averaged over 50 trials

5. CONCLUSION

We have presented a new exponential formulation for photon-limited imaging inverse problems, involving logarithmic TV regularization. This method has been shown to perform better than its untransformed counterpart in visual quality and empirical error of reconstructions. We have also shown significant improvements in selecting optimal regularization parameters via cross-validation, where this new formulation prevents over-smoothing of reconstructions. This has beneficial implications in experimental settings, where the truth is unknown and must be accurately estimated. This work has further extensions beyond TV and can be extended to other convex regularizations, such as ℓ_1 and learned dictionaries (*e.g.*, [8]), and to more general frameworks not limited to images.

6. REFERENCES

- [1] R. Willett, Z. Harmany, and R. Marcia, "Poisson image reconstruction with total variation regularization," in *Proc. IEEE International Conference on Image Processing*, 2010.
- [2] D. Needell and R. Ward, "Stable image reconstruction using total variation minimization," arXiv:1202.6429, 2012.
- [3] A. Beck and M. Teboulle, "Fast gradient-based algorithms for constrained total variation image denoising and deblurring problems," *IEEE Transactions on Image Processing*, vol. 18, no. 11, pp. 2419–2434, November 2009.
- [4] A.J. Dobson, *Introduction to Generalized Linear Models*, Chapman and Hall/CRC, 2nd edition, 1945.
- [5] S. Boyd and L. Vandenberghe, Convex Optimization, Cambridge University Press, Cambridge, 2004.
- [6] A. Beck and M. Teboulle, "A fast iterative shrinkage-thresholding algorithm for linear inverse problems," *SIAM Journal on Imaging Sciences*, vol. 2, no. 1, pp. 183–202, 2009.
- [7] Z.T. Harmany, Computational Optical Imaging Systems: Sensing Strategies, Optimization Methods, and Performance Bounds, Ph.D. thesis, Duke University, Durham, NC, 2012.
- [8] J. Salmon, Z. Harmany, C. Deledalle, and R. Willett, "Poisson noise reduction with non-local PCA," submitted to *Journal of Mathematical Imaging and Vision*. arXiv:1206:0338, 2012.