

**A Novel Experimental Paradigm Provides Evidence for a General Factor of
Visual Illusion Sensitivity and Personality Correlates**

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23

Abstract

24 Visual illusions highlight how the brain uses contextual and prior information to inform our
25 perception of reality. Unfortunately, illusion research has been hampered by the difficulty
26 of adapting these stimuli to experimental settings, leaving open the questions pertaining to
27 a potential unique factor underlying the sensitivity to different types of illusions. In this
28 study, we used a novel parametric framework for visual illusions to generate 10 different
29 classic illusions (Delboeuf, Ebbinghaus, Rod and Frame, Vertical-Horizontal, Zöllner,
30 White, Müller-Lyer, Ponzo, Poggendorff, Contrast) varying in strength, embedded in a
31 perceptual discrimination task. We tested the objective effect of the illusions on errors and
32 reaction times, and extracted participant-level performance scores ($n=250$). Our results
33 provide evidence in favour of a general factor (labelled Factor i) underlying the sensitivity
34 to different illusions. Moreover, we report a positive relationship between illusion
35 sensitivity and personality traits such as Agreeableness, Honesty-Humility, and negative
36 relationships with Psychoticism, Antagonism, Disinhibition, and Negative Affect.

37 *Keywords:* visual illusions, illusion game, Pyllusion, personality, general factor

38 Word count: 3978

39 **A Novel Experimental Paradigm Provides Evidence for a General Factor of**
40 **Visual Illusion Sensitivity and Personality Correlates**

41 **Significance Statement.** A novel paradigm to study the objective effect of visual
42 illusions yielded evidence in favor of a common factor to visual illusions (Factor *i*) and a
43 relationship between illusion resistance and maladaptive personality traits, such as
44 antagonism, psychoticism and disinhibition.

45 **Introduction**

46 Visual illusions are fascinating stimuli capturing a key feature of our neurocognitive
47 systems. They eloquently show that our brains did not evolve to be perfect perceptual
48 devices providing veridical accounts of physical reality, but integrate prior knowledge and
49 contextual information - blended together in our subjective conscious experience (Carbon,
50 2014). Despite the long-standing interest within the fields of visual perception (Day, 1972;
51 Eagleman, 2001; Gomez-Villa et al., 2022), consciousness science (Caporuscio et al., 2022;
52 Lamme, 2020), and psychiatry (Gori et al., 2016; Notredame et al., 2014; Razeghi et al.,
53 2022; Teufel et al., 2015), several important issues remain open.

54 One area of contention concerns the presence of a common mechanism underlying the
55 effect of different illusions (Cretenoud et al., 2020; Hamburger, 2016). While early research
56 has suggested a common factor of illusion sensitivity indexed by overall vision proficiency
57 (Halpern et al., 1999; Thurstone, 1944), recent empirical studies observed at most weak
58 correlations between resistance to distinct illusions (Grzeczkowski et al., 2017, 2018). The
59 existence of dispositional correlates of illusion sensitivity has also been controversial, with
60 evidence suggesting a lower illusion sensitivity in patients with schizophrenia and autism
61 (Gori et al., 2016; Grzeczkowski et al., 2018; Notredame et al., 2014; Park et al., 2022;
62 Razeghi et al., 2022), as well as individuals with stronger aggression and narcissism traits
63 (Konrath et al., 2009; Zhang et al., 2017).

64 Although the nature of the processes underlying illusion perception - whether related
65 to low-level features of the visual processing system (Cretenoud et al., 2019; Gori et al.,
66 2016) or to top-down influences (Caporuscio et al., 2022; Teufel et al., 2018) - remains
67 debated, a growing body of literature proposes to conceptualize illusions under the
68 Bayesian brain hypothesis (Friston, 2010), as ambiguous percepts (noisy sensory evidence)
69 giving ample weight to prior knowledge to minimize prediction error and provide a
70 coherent perceptual experience. In this framework, certain dispositional traits or
71 characteristics (e.g., psychotism) are seen as driven by alterations in the system's
72 metacognitive components (Adams et al., 2013), resulting in an underweighting of priors
73 during perceptual inferences, and manifesting as a decreased sensitivity to illusions (Koethe
74 et al., 2009).

75 Despite strong theoretical foundations and hypotheses, the empirical evidence
76 remains scarce, clouded by methodological hurdles. For instance, one key challenge can be
77 found in the difficulty of adapting visual illusions to an experimental setting, which
78 typically requires the controlled modulation of the specific variables of interest. Instead,
79 existing studies typically use only one or a small subset of illusion types, with few
80 contrasting conditions, restricting the findings' generalizability (Bressan & Kramer, 2021;
81 Cretenoud et al., 2019; Cretenoud et al., 2020). Moreover, conventional paradigms often
82 focus on the participants' subjective experience, by asking them the extent to which they
83 perceive two identical targets as different (Lányi et al., 2022), or having them adjust the
84 targets to perceptually match a reference stimulus (Grzeczkowski et al., 2018; Mylniec &
85 Bednarek, 2016). This reliance on meta-cognitive judgements about one's subjective
86 experience likely distorts the measurand, limiting the ability to reliably obtain more direct
87 and objective measures of illusion sensitivity (Skottun & Skoyles, 2014).

88 To address these issues, we first developed a parametric framework to manipulate
89 visual illusions that we implemented and made accessible in the open-source software

90 *Pyllusion* (Makowski et al., 2021). This software allows us to generate different types of
91 classic visual illusions with a continuous and independent modulation of two parameters:
92 *illusion strength* and *task difficulty* (**Figure 1**). Indeed, many visual illusions can be seen
93 as being composed of *targets* (e.g., same-length lines), of which perception is biased by the
94 *context* (e.g., in the Müller-Lyer illusion, the same-length line segments appear to have
95 different lengths if they end with inwards vs. outwards pointing arrows). Past illusion
96 studies traditionally employed paradigms focusing on participants' subjective experience,
97 by asking them the extent to which they perceive two identical targets as different (Lányi
98 et al., 2022), or having them adjust the targets to match a reference stimulus relying only
99 on their perception (Grzeczkowski et al., 2018; Mylniec & Bednarek, 2016). Alternatively,
100 *Pyllusion* allows the creation of illusions in which the targets are objectively different (e.g.,
101 one segment is truly more or less longer than the other), and in which the illusion varies in
102 strength (the biasing angle of the arrows is more or less acute).

103 This systematic calibration of the stimuli enables the creation of experimental tasks
104 in which participants make perceptual judgments about the targets (e.g., which segment is
105 the longest) under different conditions of objective difficulty and illusion strength.
106 Moreover, the illusion effect can be specified as either “incongruent” (making the task more
107 difficult by biasing the perception in the opposite way) or “congruent” (making the task
108 easier). Although visual illusions are inherently tied to subjective perception, this
109 framework allows a reversal of the traditional paradigm to potentially quantify the
110 “objective” effect of illusions by measuring its behavioral effect (error rate and reaction
111 times) on the performance in a perceptual task.

112 The aim of the present preregistered study is three-fold. First, we will test this novel
113 paradigm by investigating if the effect of illusion strength and task difficulty can be
114 manipulated continuously for 10 different classic illusions (Delboeuf, Ebbinghaus, Rod and
115 Frame, Vertical-Horizontal, Zöllner, White, Müller-Lyer, Ponzo, Poggendorff, Contrast).

Parametric Framework for Visual Illusions

Example with the Müller-Lyer Illusion



The Müller-Lyer Illusion is traditionally presented as two segments (the **red targets**), which perception is biased by the **context** (the arrows). Here, the lower segment appears longer despite being of the same length.

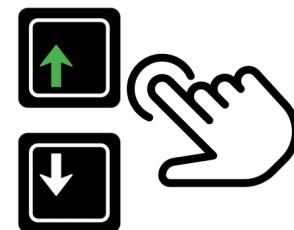


In this illusion, the **task difficulty** corresponds to the difference between the lengths of the red target segments, and the **illusion strength** corresponds to the angle of the arrows.

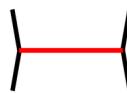
Example of Stimuli



- ✓ Task difficulty: **easy**
(top line is 2 times longer)
- ✓ Illusion Strength: **strong**
(angle is sharp)
- ← Illusion Direction (left): **incongruent**
(the illusion makes the task harder)
- Illusion Direction (right): **congruent**
(the illusion makes the task easier)



- ✓ Task difficulty: **hard**
(top line is only 1.1 times longer)
- ✓ Illusion Strength: **weak**
(angle is flat)
- ← Illusion Direction (left): **incongruent**
(the illusion makes the task harder)
- Illusion Direction (right): **congruent**
(the illusion makes the task easier)



Task: For these stimuli, the correct response is always the « up » arrow, indicating the longer red segment. We measured the reaction time and the errors (in this case, the « down » arrow).

Stimuli created with the open-source software PyMusion (Makowski et al., 2021)

Figure 1. The parametric framework for visual illusions (Makowski et al., 2021) applied to the Müller-Lyer illusion (above). Below are examples of stimuli showcasing the manipulation of two parameters, task difficulty and illusion strength.

116 Next, we will investigate the factor structure of illusion-specific performance scores and test
117 the existence of a common latent factor of illusion sensitivity. Finally, we will explore how
118 illusion sensitivity relates to demographic characteristics, contextual variables, and
119 personality traits.

120 Following open-science standards, all the material (stimuli generation code,
121 experiment code, raw data, analysis script with complementary figures and analyses,
122 preregistration, etc.) is available as **Supplementary Materials** at
123 <https://github.com/RealityBending/IllusionGameValidation>.

124 Methods

125 Stimuli

126 A pilot study ($n = 46$), of which full description is available in the Supplementary
127 Materials, was first conducted to determine a sensitive range of stimuli parameters. Then,
128 for each of the 10 illusion types, we generated a total of 134 stimuli. These stimuli resulted
129 from the combination of 15 equally-spaced levels of illusion *strength* (7 negative, i.e.,
130 congruent effects; 7 positive, i.e., incongruent effects; and 0) overlapped with 16
131 non-linearly spaced task *difficulty* levels (i.e., with an exponential, square or cubic spacing
132 depending on the pilot results). For instance, a linear space of [0.1, 0.4, 0.7, 1.0] can be
133 transformed to an exponential space of [0.1, 0.34, 0.64, 1.0], where 0.1 corresponds to the
134 highest difficulty - i.e., the smallest objective difference between targets). For each illusion
135 type, the stimuli were split into two series (56 and 72 stimuli per series) with alternating
136 parameter values to maintain their homogeneity. Additionally, 6 stimuli per illusion type
137 were generated for a practice series using parameters with more extreme variations (i.e.,
138 containing very easy trials to help cement the task instructions).

139 Procedure

140 After a brief demographic survey and a practice series of illusions, the first series of
141 10 illusion blocks was presented in a randomized order, with a further randomization of the
142 stimuli order within each block. Following this first series of blocks, two personality
143 questionnaires were administered, the *IPIP6* (24 items, Sibley et al., 2011) - measuring 6
144 “normal” personality traits (Extraversion, Openness, Conscientiousness, Agreeableness,
145 Neuroticism and Honesty-Humility), and the *PID-5* (25 items, Hopwood et al., 2012) -
146 measuring 5 “pathological” personality traits (Disinhibition, Antagonism, Detachment,
147 Negative Affect and Psychoticism). Next, the second series of 10 illusion blocks was
148 presented (with new randomized orders of blocks and trials). In total, each participant
149 underwent 1340 trials of which they had to respond “as fast as possible without making
150 errors” (i.e., an explicit double constraint to mitigate the inter-individual variability in the
151 speed-accuracy trade off) by pressing the correct arrow key (left/right, or up/down
152 depending on the illusion type). For instance, in the Müller-Lyer block, participants had to
153 answer which one of the upper or bottom target line was the longest. The task was
154 implemented using *jsPsych* (De Leeuw, 2015), and the set of instructions for each illusion
155 type is available in the experiment code.

156 Participants

157 Participants were recruited via *Prolific*, a crowd-sourcing platform recognized for
158 providing high quality data (Peer et al., 2022). The only inclusion criterion was a fluent
159 proficiency in English to ensure that the task instructions would be well-understood.
160 Participants were incentivised with a reward of about £7.5 for completing the task, which
161 took about 50 minutes to finish. Demographic variables (age, gender, and ethnicity) were
162 self-reported on a voluntary basis.

163 We excluded 6 participants upon inspection of the average error rate (when close to
164 50%, suggesting random answers), and reaction time distribution (when implausibly fast).

165 For the remaining participants, we discarded blocks with more than 50% of errors (2.16%
166 of trials), possibly indicating that instructions were misunderstood (e.g., participants
167 focused on the shorter line instead of the longer one), and 0.76% trials with extreme
168 response times (< 125 ms or > 4 SD above mean). Additionally, due to a technical issue,
169 no personality data was recorded for the first eight participants.

170 The final sample included 250 participants (Mean age = 26.5, SD = 7.6, range: [18,
171 69]; Sex: 48% females, 52% males).

172 Data Analysis

173 The first part of the analysis focused on modelling the effect of illusion strength and
174 task difficulty on errors and reaction time (RT) within each illusion. We started by fitting
175 General Additive Models (GAMs), which can parsimoniously accommodate possible
176 non-linear effects and interactions. Errors were analyzed using Bayesian logistic mixed
177 models, and RTs of correct responses were analyzed using an ex-Gaussian family with the
178 same fixed effects entered for the location μ (mean), scale σ (spread) and tail-dominance τ
179 of the RT distribution (Balota & Yap, 2011; Matzke & Wagenmakers, 2009).

180 Using GAMs as the “ground-truth” models, we attempted at approximating them
181 using general linear mixed models, which can be used to estimate the effects'
182 participant-level variability (via random slopes). Following a comparison of models with a
183 combination of transformations (raw, log, square root or cubic root) on the main predictors
184 (task *difficulty* and illusion *strength*), we fitted the best model (based on their indices of
185 fit), and compared their output visually (**Figure 2**).

186 The inter-individual variability in the effect of illusion strength and its interaction
187 with task difficulty was extracted from the models and used as participant-level scores. We
188 then explored the relationship of these indices across different illusions using exploratory
189 factor analysis (EFA) and structural equation modelling (SEM), and tested the existence of

190 a general factor of illusion sensitivity (Factor i).

191 Finally, for each of the individual illusion sensitivity scores (10 illusion-specific factors
192 and the general Factor i), we tested the effect of contextual variables (screen size, screen
193 refresh rate), demographic variables (sex, education, age), and personality traits.

194 The analysis was carried out using *R 4.2* (R Core Team, 2022), *brms* (Bürkner,
195 2017), the *tidyverse* (Wickham et al., 2019), and the *easystats* collection of packages
196 (Lüdecke et al., 2021, 2019; Makowski et al., 2020; Makowski, Ben-Shachar, & Lüdecke,
197 2019). As the full results are available as supplementary materials, we will focus here on
198 the significant results (based on the Bayes Factor BF or the Probability of Direction pd ,
199 see Makowski, Ben-Shachar, Chen, et al., 2019).

200 **Results**

201 **Effects of Illusion Strength and Task Difficulty**

202 The best model specifications were $\log(\text{diff}) * \text{strength}$ for Delboeuf;
203 $\sqrt{\text{diff}} * \text{strength}$ for Ebbinghaus; $\log(\text{diff}) * \log(\text{strength})$ for Rod and Frame;
204 $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ for Vertical-Horizontal; $\text{cbrt}(\text{diff}) * \text{strength}$ for Zöllner;
205 $\text{diff} * \sqrt{\text{strength}}$ and $\log(\text{diff}) * \text{strength}$ respectively for errors and RT in White;
206 $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ and $\sqrt{\text{diff}} * \text{strength}$ respectively for errors and RT in
207 Müller-Lyer; $\text{cbrt}(\text{diff}) * \text{strength}$ for Ponzo; $\text{cbrt}(\text{diff}) * \sqrt{\text{strength}}$ and
208 $\text{cbrt}(\text{diff}) * \text{strength}$ respectively for errors and RT in Poggendorff; and
209 $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ for Contrast. For all of these models, the effects of illusion
210 strength, task difficulty and their interaction were significant.

211 For error rates, most of the models closely matched their GAMs counterpart, with
212 the exception of Delboeuf (for which the GAM suggested a non-monotonic effect of illusion
213 strength with a local minimum at 0) and Zöllner (for which theoretically congruent illusion

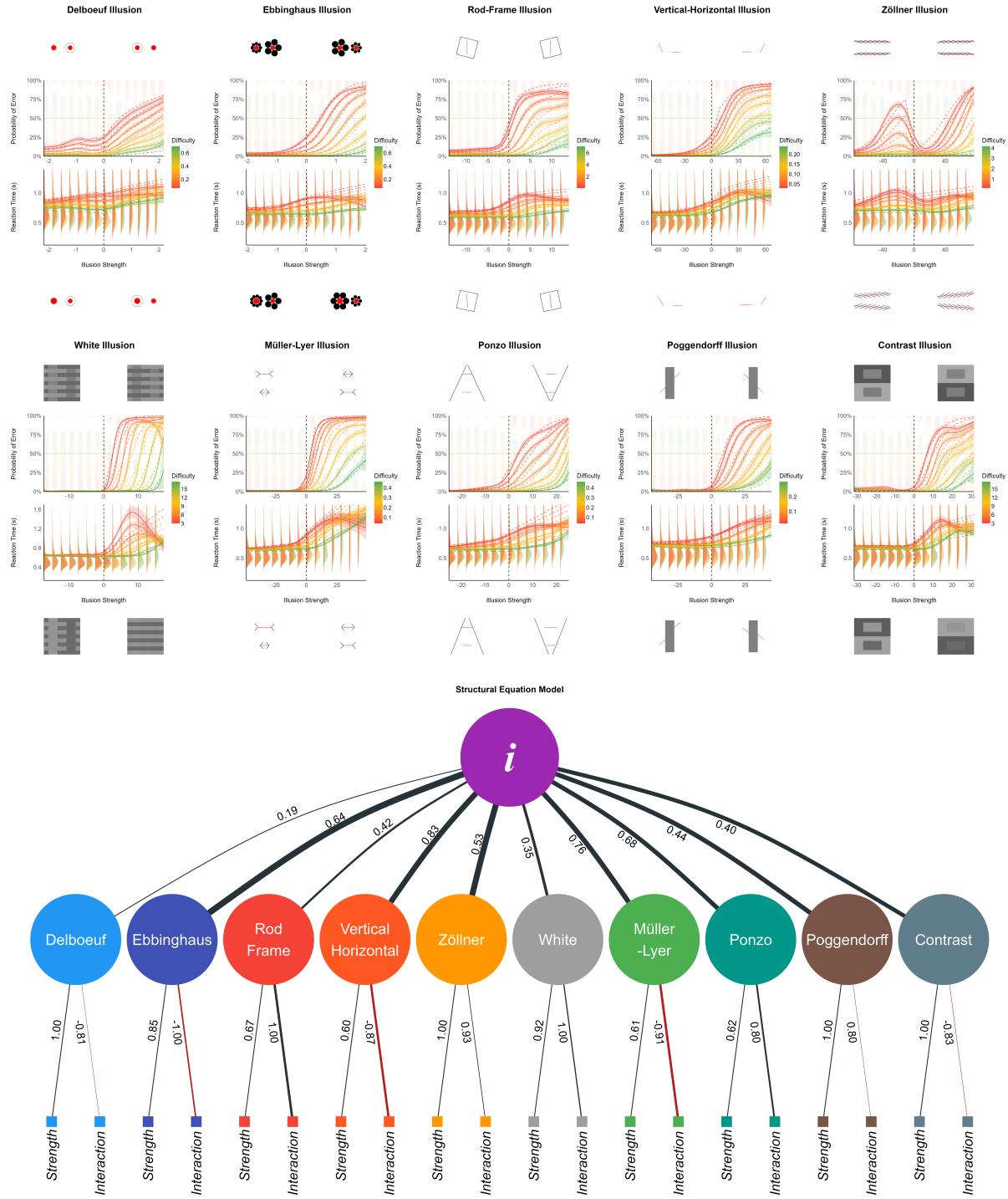


Figure 2. Top: the effect of illusion strength and task difficulty on the error rate and reaction time (RT) for each individual illusion. The solid line represents the General Additive Model (GAM), and the dashed line corresponds to its approximation via linear models. Descriptive data is shown with stacked dots (for which errors start from the top) and distributions for RTs. Negative values for illusion strength correspond to congruent (i.e., facilitating) illusion effects. Task difficulty (the objective difference between the targets of perceptual decision) levels are shown as colors, with lower values corresponding to harder trials. The results for each illusion are surrounded by 4 extreme examples of stimuli, corresponding to the hardest difficulty (on top) and the strongest illusion (on the right for incongruent illusions). Bottom: We extracted the effect slope of the illusion strength and its interaction with task difficulty for each participant. We fitted a Structural Equation Model (SEM) suggesting that these manifest variables group to first-level illusion-specific latent factors, which then load on a general factor of illusion sensitivity (Factor *i*).

214 effects were related to increased error rate). A specific discussion regarding these 2 illusions
215 is available in the Supplementary Materials (Part 1 - Discussion).

216 For RTs, the GAMs suggested a consistent non-linear relationship between RT and
217 illusion strength: as the illusion strength increases beyond a certain threshold, the
218 participants responded faster. While this is not surprising (strong illusions are likely so
219 effective in biasing perception that it is “easier”, i.e., faster, to make the wrong decision),
220 the linear models were not designed to capture this - likely quadratic - pattern and hence
221 are not good representatives of the underlying dynamics. As such, we decided not to use
222 them for the individual scores analysis.

223 Factor Structure

224 Though imperfect, we believe that the random-slope models capture inter-individual
225 differences with more accuracy (and also provide more conservative estimates due to
226 shrinkage) than basic empirical scores, such as the total number of errors, or the average
227 RT. Thus, for each illusion and within each participant, we extracted the effect of illusion
228 strength and its interaction with task difficulty when the illusion effect was incongruent.
229 These twenty participant-level scores were subjected to exploratory factor analysis (EFA).
230 The Method Agreement Procedure (Lüdecke et al., 2020) suggested the presence of 7 latent
231 factors. An oblique (*oblimin* rotation) factor solution explaining 66.69% of variance
232 suggested separate dimensions for the effect of Zöllner, White, Poggendorff, Contrast,
233 Ebbinghaus, Delboeuf, and a common factor for the parameters related to Müller-Lyer,
234 Vertical-Horizontal, Ponzo and Rod and Frame. We submitted these factors to a
235 second-level analysis and extracted two orthogonal (*varimax* rotation) factors. The first
236 factor was loaded by all the previous dimensions with the exception of Delboeuf, which
237 formed its own separate factor.

238 Finally, we tested this data-driven model ($m0$) against four other structural models

239 using structural equation modelling (SEM): one in which the two parameters of each of the
240 10 illusions (illusion strength and interaction with task difficulty) loaded on separate
241 factors, which then all loaded on a common factor ($m1$); one in which the parameters were
242 grouped by illusion type (lines, circles, contrast and angle) before loading on a common
243 factor ($m2$); one in which all the parameters related to strength, and all the parameters
244 related to the interaction loaded onto two respective factors, which then loaded on a
245 common factor ($m3$); and one in which there was no intermediate level: all 20 parameters
246 loaded directly on a common factor ($m4$).

247 The model $m1$, in which the parameters loaded on a first level of 10 illusion-specific
248 factors, which then all loaded on a common factor, significantly outperformed the other
249 models. Its indices of fit ranged from acceptable to satisfactory (CFI = .92; SRMR = .08;
250 NNFI = .91; PNFI = .74; RMSEA = .08), and all the specified effects were significant.
251 The illusion-specific latent factors were loaded positively by the sensitivity to illusion
252 strength, as well as by the interaction effect with task difficulty (with the exception of
253 Delboeuf, Ebbinghaus, Vertical-Horizontal, Müller-Lyer and Contrast, for which the
254 loading was negative). The general factor of illusion sensitivity, labelled Factor i (i - for
255 illusion), explained 48.02% of the total variance of the initial dataset, and was strongly
256 related to Vertical-Horizontal ($\beta_{std.} = 0.83$), Müller-Lyer ($\beta_{std.} = 0.76$), Ponzo
257 ($\beta_{std.} = 0.65$), Ebbinghaus ($\beta_{std.} = 0.64$); moderately to Zöllner ($\beta_{std.} = 0.53$), Poggendorff
258 ($\beta_{std.} = 0.44$), Rod and Frame ($\beta_{std.} = 0.42$), Contrast ($\beta_{std.} = 0.40$) and White
259 ($\beta_{std.} = 0.35$); and weakly to Delboeuf ($\beta_{std.} = 0.19$). We then computed, for each
260 participant, the score for the 10 illusion-specific factors and for the general Factor i .

261 It is important to note that these individual scores are the result of several layers of
262 simplification: 1) the individual coefficient is that of simpler models that sometimes do not
263 perfectly capture the underlying dynamics (especially in the case of Delboeuf and Zöllner);
264 2) we only used the models on error rate, which could be biased by the speed-accuracy

265 decision criterion used by participants; 3) the structural equation model used to compute
266 the scores also incorporated multiple levels of abstractions. Thus, in order to validate the
267 individual scores, we computed the correlation between them and simple empirical scores,
268 such as the average error rate and the mean RT in the task. This analysis revealed strong
269 and significant correlations between each illusion-specific factor and the average amount of
270 errors in its corresponding task. Moreover, each individual score was strongly associated
271 with the average RT across multiple illusion types. This suggests that the individual scores
272 obtained from the structural equation model do capture the sensitivity of each participant
273 to visual illusions, manifesting in both the number of errors and long reaction times.

274 Correlations with Inter-individual Characteristics

275 The Bayesian correlation analysis (with narrow priors centered around a null effect)
276 between the illusion scores and contextual variables (screen size and refresh rate) provided
277 weak evidence in favor of an absence of effect, with the exception of the two contrast-based
278 illusions. Anecdotal ($BF_{10} = 2.05$) and moderate evidence ($BF_{10} = 4.11$) was found for a
279 negative correlation between screen size and the sensitivity to the White and the Contrast
280 illusion, respectively. To test whether this result could be an artifact related to the highly
281 skewed screen size distribution (caused by very few participants with extreme screen sizes),
282 we re-ran a robust correlation (with rank-transformed values), which provided even
283 stronger evidence in favor of the effect existence ($BF_{10} = 28.19$, $BF_{10} = 4.31$ for White and
284 Contrast, respectively).

285 The Bayesian t-tests on the effect of sex suggested anecdotal to moderate evidence in
286 favour of the null effect for all scores, with the exception of the sensitivity to the Zöllner
287 illusion, which was higher in males as compared to females ($\Delta = -0.37$, 95% CI [-0.62,
288 -0.13], $BF_{10} = 12.74$). We fitted Bayesian linear models with the education level entered as
289 a monotonic predictor (appropriate for ordinal variables, Bürkner & Charpentier, 2020),
290 which yielded no significant effects. For age, we fitted two types of models for each score,

one general additive models (GAM) and a 2nd order polynomial model. These consistently suggested a significant positive linear relationship between age and Factor i ($pd = 100\%$), as well as the sensitivity to Müller-Lyer ($pd = 100\%$), Vertical-Horizontal ($pd = 100\%$), Zöllner ($pd = 100\%$) and Ebbinghaus ($pd = 99\%$) illusions (**Figure 3**).

Regarding “normal” personality traits, Bayesian correlations suggested substantial evidence in favor of a positive relationship between *Honesty-Humility* and Zöllner ($BF_{10} > 100$), Vertical-Horizontal ($BF_{10} = 9.78$) and the Factor i ($BF_{10} = 4.00$); as well as between *Agreeableness* and Vertical-Horizontal ($BF_{10} = 25.06$), Ponzo ($BF_{10} = 4.88$) and the Factor i ($BF_{10} = 19.65$).

Regarding “pathological” personality traits, the results yielded strong evidence in favor of a negative relationship between illusion scores and multiple traits. *Antagonism* was associated with the sensitivity to Vertical-Horizontal ($BF_{10} > 100$), Müller-Lyer ($BF_{10} = 21.57$), Ponzo ($BF_{10} = 17.97$) illusions, and the Factor i ($BF_{10} = 55.45$); *Psychoticism* was associated with the sensitivity to Vertical-Horizontal ($BF_{10} = 66.63$) and Müller-Lyer ($BF_{10} = 35.59$) illusions, and the Factor i ($BF_{10} = 35.02$); *Disinhibition* was associated with the sensitivity to Vertical-Horizontal ($BF_{10} = 25.38$), Zöllner ($BF_{10} = 7.59$), Müller-Lyer ($BF_{10} = 5.89$) illusions, and the Factor i ($BF_{10} = 31.42$); and *Negative Affect* was associated with Zöllner ($BF_{10} = 62.04$), Vertical-Horizontal ($BF_{10} = 12.65$), Müller-Lyer ($BF_{10} = 3.17$), and the Factor i ($BF_{10} = 6.39$). The last remaining trait, *Detachment*, did not share any significant relationship with illusion sensitivity. See Supplementary Materials (Part 2 - Discussion) for a detailed discussion regarding these associations.

Discussion

This study tested a novel illusion sensitivity task paradigm based on the parametric illusion generation framework (Makowski et al., 2021). Using the carefully generated

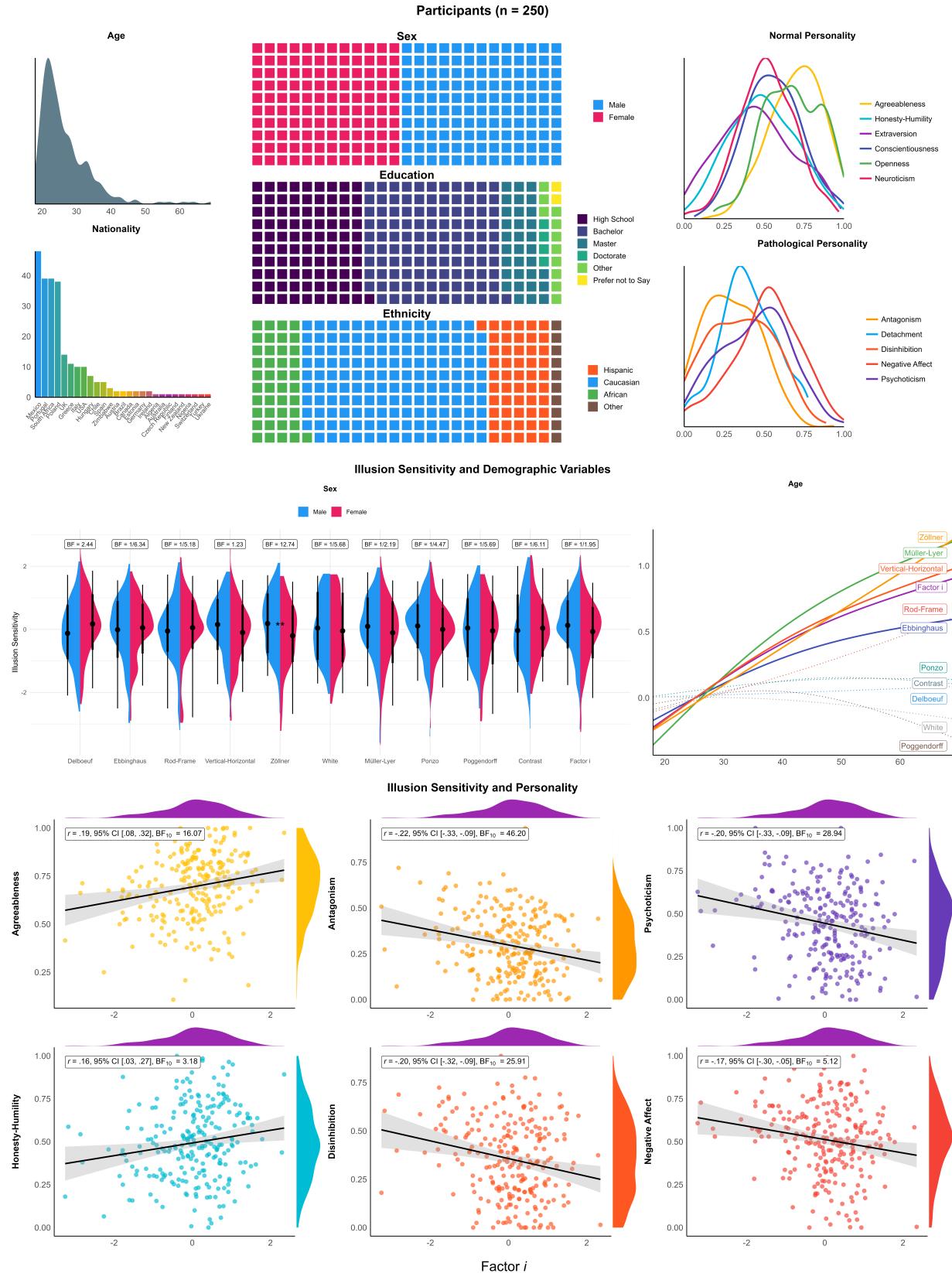


Figure 3. The upper plots show the distribution of demographic and dispositional variables. The middle plots show the illusion sensitivity scores as a function of sex and age (solid lines indicate significant relationships). Bottom plots show the correlation between the general factor of illusion sensitivity (Factor i) and personality traits.

316 stimuli in a perceptual decision task, we have shown that a gradual modulation of illusion
317 strength is effectively possible across 10 different types of classic visual illusions. Increasing
318 the illusion strength led to an increase in error likelihood, as well as the average and spread
319 of RTs (but only up to a point, after which participants become faster at responding with
320 the wrong answer). Using mixed models, we were able to statistically quantify the effect of
321 illusions for each illusion and each participant separately. This important methodological
322 step opens the door for new illusions-based paradigms and tasks to study the effect of
323 illusions under different conditions and to measure illusion sensitivity using objective
324 behavioral outcomes - such as accuracy or speed - instead of subjective meta-cognitive
325 reports. This new and complementary approach will hopefully help address some of the
326 longstanding literature gaps, as well as cement illusions as valuable stimuli for the study of
327 cognition.

328 Our findings suggest that the sensitivity to 10 different types of visual illusions share a
329 common part of variance, supporting the existence of a general factor of illusion sensitivity
330 (Factor *i*). This result comes in a field of mixed findings. In fact, contrary to early studies
331 on visual illusions, more recent research have generally not found any significant evidence
332 for a common stable factor across illusions within individuals (Cretenoud et al., 2019;
333 Cretenoud et al., 2020; Grzeczkowski et al., 2017, 2018; Yang et al., 2012). Instead, past
334 findings suggest illusory effects are highly specific to the perceptual features of the illusions
335 at stake (Cretenoud et al., 2019; Grzeczkowski et al., 2017). It should be noted, however,
336 that most of these studies were low-powered and/or relied on conventional paradigms, such
337 as the adjustment procedure to measure the participants' subjective perception. We believe
338 that our study presents several methodological improvements, including statistical power
339 (high number of trials per participant), homogeneous stimuli (with minimal and highly
340 controlled features) and tasks (decision-making reaction-time task), and a more reliable
341 participant-level score extraction method (based on random-factors models), which in our
342 opinion contributed to the emergence of the common factor.

Finally, we found illusion sensitivity to be positively associated with “positive” personality traits, such as agreeableness and honesty-humility, and negatively associated with maladaptive traits such as antagonism, psychoticism, disinhibition, and negative affect. Although the existing evidence investigating links between illusion sensitivity and personality traits is scarce, these results are consistent with past findings relating pathological egocentric beliefs (often associated with psychoticism, Fox, 2006) to reduced context integration, manifesting in a tendency to separate objects from their surroundings when processing visual stimuli (Fox, 2006; Konrath et al., 2009; Ohmann & Burgmer, 2016). As such, the association between maladaptive traits and lower illusion sensitivity could be linked to a self-centered, decontextualized and disorganized information processing style. Conversely, the relationship between illusion sensitivity and adaptive personality traits is in line with the decreased field dependence (the tendency to rely on external cues in ambiguous contexts) associated with traits negatively correlated with agreeableness and honesty-humility, such as hostility, aggression and narcissism (Konrath et al., 2009; Pessoa et al., 2008; Zhang et al., 2017).

Importantly, these findings highlight the relevance of illusions beyond the field of visual perception, pointing towards an association with high-level domain-general mechanisms. In particular, the evidence in favor of a relationship between maladaptive personality traits and illusion sensitivity is in line with clinical observations, in which a greater resistance to illusions have been reported among patients with schizophrenia (Grzeczkowski et al., 2018; Notredame et al., 2014; Pessoa et al., 2008), especially in association with schizotypal traits such as cognitive disorganization (Cretenoud et al., 2019; Lányi et al., 2022). While the search for the exact mechanism(s) underlying these links is an important goal of future research, our findings unlock the potential of illusion-based tasks as sensitive tools to capture specific inter-individual neuro-cognitive differences.

Future research is needed to address several limitations. One key question concerns

369 the relationship of illusion sensitivity with perceptual abilities (e.g., using similar tasks, but
370 without illusions). Although the illusions used in the present study did differ in terms of
371 the perceptual task (contrast-based, size-estimation, angle-perception), the possibility of
372 our general factor being driven by inter-individual perceptual skills variability (or other
373 cognitive skills) cannot be discarded. Moreover, using only the error rate models to extract
374 individual-level scores might fail in capturing the whole range of behavioral dynamics.
375 Future work should attempt at integrating the reaction times data (e.g., by jointly
376 analyzing them using drift diffusion models), and assess the psychometric properties - such
377 as stability (e.g., test-retest reliability) and validity - of similar illusion-based paradigms.
378 Finally, while the personality measures used in this study highlight illusion sensitivity as an
379 interesting measure rather than a mere perceptual artifact, further studies should test its
380 relationship with more specific dispositional characteristics (e.g., autistic or schizotypal
381 traits), cognitive styles and abilities, to help understanding the potential underlying
382 mechanisms of these associations.

383 In conclusion, we strongly invite researchers to explore and re-analyze our dataset
384 with other approaches and methods to push the understanding of visual illusions and
385 illusion sensitivity further. The task, data and analysis script are available in open-access
386 at <https://github.com/RealityBending/IllusionGameValidation>.

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390

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