

¹ **A Novel Visual Illusion Paradigm Provides Evidence for a General Factor of
2 Illusion Sensitivity and Personality Correlates**

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22

Abstract

23 Visual illusions are a gateway to understand how we construct our experience of reality.
24 Unfortunately, important questions remain open, such as the hypothesis of a common
25 factor underlying the sensitivity to different types of illusions, as well as of personality
26 correlates of illusion sensitivity. In this study, we used a novel parametric framework for
27 visual illusions to generate 10 different classic illusions (Delboeuf, Ebbinghaus, Rod and
28 Frame, Vertical-Horizontal, Zöllner, White, Müller-Lyer, Ponzo, Poggendorff, Contrast)
29 varying in strength, embedded in a perceptual discrimination task. We tested the objective
30 effect of the illusions on errors and response times, and extracted participant-level
31 performance scores ($n=250$) for each illusion. Our results provide evidence in favour of a
32 general factor underlying the sensitivity to different illusions (labelled Factor i). Moreover,
33 we report a positive link between illusion sensitivity and personality traits such as
34 Agreeableness, Honesty-Humility, and negative relationships with Psychoticism,
35 Antagonism, Disinhibition, and Negative Affect.

36 *Keywords:* visual illusions, illusion game, Pyllusion, personality, general factor

37 Word count: 4085

38 A Novel Visual Illusion Paradigm Provides Evidence for a General Factor of
39 Illusion Sensitivity and Personality Correlates

Significance Statement. A novel paradigm to study the objective effect of visual illusions yielded evidence in favor of a common factor to visual illusions (Factor *i*) and a relationship between illusion resistance and maladaptive personality traits, such as antagonism, psychoticism and disinhibition.

Introduction

Visual illusions are fascinating stimuli capturing a key feature of our neurocognitive systems. They eloquently show that our brains did not evolve to be perfect perceptual devices providing veridical accounts of physical reality, but integrate prior knowledge and contextual information - blended together in our subjective conscious experience¹. Despite the long-standing interest within the fields of visual perception²⁻⁴, consciousness science^{5,6}, and psychiatry⁷⁻¹⁰, several important issues remain open.

51 One area of contention concerns the presence of a common mechanism underlying the
52 effect of different illusions^{11,12}. While early research has suggested a common factor of
53 illusion sensitivity indexed by overall vision proficiency^{13,14}, recent empirical studies
54 observed at most weak correlations between inter-individual resistance to distinct
55 illusions^{15,16}. The existence of dispositional correlates of illusion sensitivity has also been
56 controversial, with evidence suggesting a lower illusion sensitivity in patients with
57 schizophrenia and autism^{7-9,16,17}, as well as individuals with stronger aggression and
58 narcissism traits^{18,19}.

59 Although the nature of the processes underlying illusion perception - whether related
60 to low-level features of the visual processing system^{8,20} or to top-down influences^{5,21} -
61 remains debated, a growing body of literature proposes to conceptualize illusions under the
62 Bayesian brain hypothesis²². In this context, illusions are conceptualized as **non-veridical**

63 **perceptual experiences that result from giving ample weight to prior knowledge**
64 **to minimize prediction error in the face of biasing sensory evidence.** The
65 predictive coding account further provides an explanation regarding the observations from
66 clinical populations. Certain dispositional traits or characteristics (e.g., psychoticism) are
67 seen as driven by alterations in the system's metacognitive components²³, resulting in an
68 underweighting of priors during perceptual inferences, and manifesting as a decreased
69 sensitivity to illusions²⁴.

70 Despite strong theoretical foundations and hypotheses, the empirical evidence remains
71 scarce, clouded by methodological hurdles. For instance, one key challenge can be found in
72 the difficulty of adapting visual illusions to an experimental setting, which typically
73 requires the controlled modulation of the specific variables of interest. Instead, existing
74 studies typically use only one or a small subset of illusion types, with few contrasting
75 conditions, restricting the findings' generalizability^{12,20,25}. Moreover, conventional
76 paradigms often focus on the participants' subjective experience, by asking them the extent
77 to which they perceive two identical targets as different²⁶, having them estimate the
78 targets' physical properties²⁷, **or through the method of adjustment, which involves**
79 **having them adjust the targets to perceptually match a reference**
80 **stimulus**^{16,28–30}. This reliance on meta-cognitive judgements about one's subjective
81 experience likely distorts the measurand, limiting the ability to reliably obtain more direct
82 and objective measures of illusion sensitivity³¹. **While some recent efforts have some**
83 **made to implement more empirically rigorous paradigms**³², **most of the applied**
84 **manipulations only focus on varying the physical dimensions of the illusion's**
85 **target features without modulating its contextual elements, hence limiting the**
86 **variability in the illusory effects of the stimuli presented.** Furthermore, such
87 **prior studies have typically generated stimuli whose targets' physical attributes**
88 **vary over a relatively narrow range, thus further constraining the reliability of**
89 **their findings.** As such, it is possible that the recent evidence reported against

90 a common factor of illusions could be due to the low stimulus variance instead
91 of a true reflection of a lack of common mechanism.

92 To address these issues, we first developed a parametric framework to manipulate
93 visual illusions that we implemented and made accessible in the open-source software
94 *Pyllusion*³³. This software allows us to generate different types of classic visual illusions
95 with a continuous and independent modulation of two parameters: *illusion strength* and
96 *task difficulty* (**Figure 1**). Indeed, many visual illusions can be seen as being composed of
97 *targets* (e.g., same-length lines), of which perception is biased by the *context* (e.g., in the
98 Müller-Lyer illusion, the same-length line segments appear to have different lengths if they
99 end with inwards vs. outwards pointing arrows). Past illusion studies traditionally
100 employed paradigms focusing on participants' subjective experience, by asking them the
101 extent to which they perceive two identical targets as different²⁶, or having them adjust the
102 targets to match a reference stimulus relying only on their perception^{16,28}. Alternatively,
103 *Pyllusion* allows the creation of illusions in which the targets are objectively different (e.g.,
104 one segment is truly more or less longer than the other), and in which the illusion varies in
105 strength (the biasing angle of the arrows is more or less acute).

106 This systematic calibration of the stimuli enables the creation of experimental tasks
107 in which participants make perceptual judgments about the targets (e.g., which segment is
108 the longest) under different conditions of objective difficulty and illusion strength.
109 Moreover, the illusion effect can be specified as either “incongruent” (making the task more
110 difficult by biasing the perception in the opposite way) or “congruent” (making the task
111 easier). Although visual illusions are inherently tied to subjective perception, this
112 framework allows a reversal of the traditional paradigm to potentially quantify the
113 “objective” effect of illusions by measuring its behavioral effect (error rate and reaction
114 times) on the performance in a perceptual task.

115 The aim of the present preregistered (<https://osf.io/5d6xp>) study is three-fold. First,

Parametric Framework for Visual Illusions

Example with the Müller-Lyer Illusion



The Müller-Lyer Illusion is traditionally presented as two segments (the **red targets**), which perception is biased by the **context** (the arrows). Here, the lower segment appears longer despite being of the same length.



In this illusion, the **task difficulty** corresponds to the difference between the lengths of the red target segments, and the **illusion strength** corresponds to the angle of the arrows. The **illusion direction** corresponds to the facilitating or impeding effect with regards to the task at hand.

Example of Stimuli



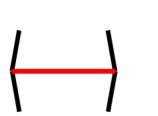
❖ Task difficulty: **easy**
(top line is 2 times longer)



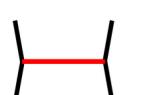
❖ Illusion Strength: **strong**
(angle is sharp)



❖ Illusion Direction: **incongruent**
(the illusion makes the task harder: red lines look more similar)



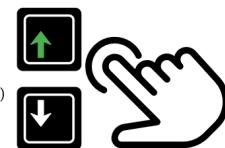
❖ Task difficulty: **hard**
(top line is only 1.1 times longer)



❖ Illusion Strength: **weak**
(angle is flat)



❖ Illusion Direction: **congruent**
(the illusion makes the task easier: red lines look more different)



Task: For these stimuli, the correct response is always the « up » arrow, indicating the longer red segment. We measured the reaction time and the errors (in this case, the « down » arrow).



Stimuli created with the open-source software PyMusion (Makowski et al., 2021)

Figure 1. The parametric framework for visual illusions (Makowski et al., 2021) applied to the Müller-Lyer illusion (above). Below are examples of stimuli showcasing the manipulation of two parameters, task difficulty and illusion strength.

we will test this novel paradigm by investigating if the effect of illusion strength and task difficulty can be manipulated continuously for 10 different classic illusions (Delboeuf, Ebbinghaus, Rod and Frame, Vertical-Horizontal, Zöllner, White, Müller-Lyer, Ponzo, Poggendorff, Simultaneous Brightness Contrast). Next, we will investigate the factor structure of illusion-specific performance scores and test the existence of a common latent factor of illusion sensitivity. Finally, we will explore how illusion sensitivity relates to demographic characteristics, contextual variables, and personality traits.

123

Methods

124 **Ethics Statement**

This study was approved by the NTU Institutional Review Board (NTU IRB-2022-187) and all procedures performed were in accordance with the ethical standards of the institutional board and with the 1964 Helsinki Declaration. All participants provided their informed consent prior to participation and were incentivized after completing the study.

130 **Stimuli**

We investigated the effect of 10 different classic illusions (**Figure 2**). A pilot study ($n = 46$), of which a full description is available at <https://github.com/RealityBending/IllusionGameValidation>, was first conducted to determine a sensitive range of stimuli parameters. Then, for each of the 10 illusion types, we generated a total of 134 stimuli. These stimuli resulted from the combination of 15 equally-spaced levels of illusion *strength* (7 negative, i.e., congruent effects; 7 positive, i.e., incongruent effects) overlapped with 16 non-linearly spaced task *difficulty* levels (i.e., with an exponential, square or cubic spacing depending on the pilot results). For instance, a linear space of [0.1, 0.4, 0.7, 1.0] can be transformed to an exponential space of [0.1, 0.34, 0.64, 1.0], where 0.1 corresponds to the highest difficulty - i.e., the smallest objective

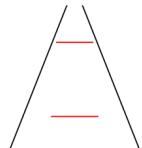
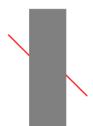
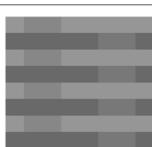
Illusion	Example	Task	Description
Delboeuf		Which red circle is bigger? 	Two circles surrounded by outer rings. The circle with the closer outline appears larger than the one with the distant outline.
Ebbinghaus		Which red circle is bigger? 	Two circles surrounded by other circles. The circle surrounded by smaller circles appears larger than the one with the larger surrounding circles.
Müller-Lyer		Which red line is longer? 	Two parallel segments that end with inwards/outwards pointing arrow-like fins. The segment with inward-pointing fins is typically perceived to be longer.
Ponzo Illusion		Which red line is longer? 	Two parallel segments embedded between a pair of converging lines. The segment on the convergence side is typically perceived to be longer.
Vertical-Horizontal		Which red line is longer? 	Two line segments, one horizontal and one angled. The angled line is usually perceived as longer.
Zöllner		Which direction are the red lines converging towards? 	Represented by two long lines crossed with short, repeated lines, that appear as converging to one side.
Rod and Frame		Which direction is the red line leaning towards? 	A segment ("rod") enclosed in a tilted square ("frame"), that appears to be leaning to one side.
Poggendorff		Is the right line above or below the left line? 	Two oblique segments separated by a rectangle. The right segment appears to be vertically misplaced.
Simultaneous Contrast		Which small rectangle is lighter? 	Consists of two smaller grey rectangles embedded in backgrounds of different contrasts. The rectangle with the darker background appears lighter.
White		Which vertical rectangle is lighter? 	Composed of a series of contrasting horizontal bars superimposed with vertical grey rectangles. Rectangles superimposed onto darker bars appear lighter.

Figure 2. Ten different illusions were used as stimuli in a perceptual task, where participants had to answer as fast as possibly, without making errors, according to specific instructions. For each illusion type, two parameters were experimentally manipulated, 1) the task difficulty (e.g., how large was the difference between the bigger and the smaller red circles in the Delboeuf illusion), and 2) the illusion strength (e.g., the size of the black circles in the Ebbinghaus illusion).

¹⁴¹ difference between targets). For each illusion type, the stimuli were split into two series (56
¹⁴² and 72 stimuli per series) with alternating parameter values to maintain their homogeneity.
¹⁴³ Additionally, 6 stimuli per illusion type were generated for a practice series using
¹⁴⁴ parameters with more extreme variations (i.e., containing very easy trials to help cement
¹⁴⁵ the task instructions).

¹⁴⁶ **Procedure**

¹⁴⁷ Participants were first given a brief demographic survey, which collected information
¹⁴⁸ regarding their **age, gender, country of birth, ethnicity and highest attained**
¹⁴⁹ **education level**. This was followed by a practice series of illusions, after which the first
¹⁵⁰ series of 10 illusion blocks was presented in a randomized order, with a further
¹⁵¹ randomization of the stimuli order within each block. Following this first series of blocks,
¹⁵² two personality questionnaires were administered, the *IPIP6*³⁴ - measuring 6 “normal”
¹⁵³ personality traits (Extraversion, Openness, Conscientiousness, Agreeableness, Neuroticism
¹⁵⁴ and Honesty-Humility), and the *PID-5*³⁵ - measuring 5 “pathological” personality traits
¹⁵⁵ (Disinhibition, Antagonism, Detachment, Negative Affect and Psychoticism). Next, the
¹⁵⁶ second series of 10 illusion blocks was presented (with new randomized orders of blocks and
¹⁵⁷ trials). In total, each participant underwent 1340 trials of which they had to respond “as
¹⁵⁸ fast as possible without making errors” (i.e., an explicit double constraint to mitigate the
¹⁵⁹ inter-individual variability in the speed-accuracy trade off) by pressing the correct arrow
¹⁶⁰ key (left/right, or up/down depending on the illusion type). For instance, in the
¹⁶¹ Müller-Lyer block, participants had to answer which one of the upper or bottom target line
¹⁶² was the longest. **All trials were required to be completed within a single-session**
¹⁶³ **(total experiment duration: ~55 minutes)**. The task was implemented using
¹⁶⁴ *jsPsych*³⁶ and was hosted on *Pavlovia* (<https://pavlovia.org/>). The set of
¹⁶⁵ instructions for each illusion type is available in the experiment code.

166 **Participants**

167 Participants were recruited via *Prolific*, a crowd-sourcing platform recognized for
168 providing high quality data³⁷. The only inclusion criterion was a fluent proficiency in
169 English to ensure that the task instructions would be well-understood. Participants were
170 incentivised with a reward of about £7.50 for completing the task, which took
171 approximately 50 minutes to finish. Demographic variables (age, gender, and ethnicity)
172 were self-reported on a voluntary basis.

173 We excluded 6 participants upon inspection of the average error rate (when close to
174 50%, suggesting random answers), and reaction time distribution (when implausibly fast
175 relative to the average RT distribution). For the remaining participants, we discarded
176 blocks with more than 50% of errors (2.16% of trials), possibly indicating that instructions
177 were misunderstood (e.g., participants focused on the shorter line instead of the longer
178 one), and 0.76% trials with extreme response times (< 125 ms or > 4 SD above mean RT).
179 Additionally, due to a technical issue, no personality data was recorded for the first eight
180 participants.

181 The final sample included 250 participants (Mean age = 26.5, SD = 7.6, range: [18 -
182 69]; Sex: 48% females, 52% males).

183 **Data Analysis**

184 The first part of the analysis focused on modelling the effect of illusion strength and
185 task difficulty on errors and response time (RT) **separately for each illusion** under a
186 Bayesian framework. We started by fitting General Additive Models (GAMs), which can
187 parsimoniously accommodate possible non-linear effects and interactions. Errors were
188 analyzed using logistic mixed models (**suites to estimate the error rate**), and RTs of
189 correct responses were analyzed using an ex-Gaussian family with the same fixed effects
190 entered for the location μ (mean), scale σ (spread) and tail-dominance τ of the RT

191 distribution^{38,39}.

192 Using GAMs as the “ground-truth” models, we attempted at approximating them
193 using general linear mixed models, which can be used to estimate the effects’
194 participant-level variability (via random slopes). Following a comparison of models with a
195 combination of transformations (raw, log, square root or cubic root; **which are types of**
196 **relationship commonly found in perceptual tasks**) on the main predictors (task
197 *difficulty* and illusion *strength*), we fitted the best model (**based on their BIC and R2**),
198 and compared their output visually (**Figure 3**). Note that the model comparison and
199 **the parameters used in the resulting models were not pre-registered**.

200 The inter-individual variability in the effect of illusion strength and its interaction
201 with task difficulty (**diff**) was extracted from the models and used as participant-level
202 scores. We then explored the relationship of these indices across different illusions using
203 exploratory factor analysis (**EFA, to gain insights into the structure**), and structural
204 equation modelling (**SEM, to model and test different hierarchical models**), and
205 tested the existence of a general factor of illusion sensitivity (Factor *i*).

206 Finally, for each of the individual illusion sensitivity scores (10 illusion-specific factors
207 and the general Factor *i*), we tested the effect of contextual variables (screen size, screen
208 refresh rate), demographic variables (sex, education, age), and personality traits. It should
209 be noted that the measure of screen size used (measured using the number of pixels) is
210 **only a proxy of the true physical screen size**.

211 The analysis was carried out using *R* 4.2⁴⁰, *brms*⁴¹, the *tidyverse*⁴², and the *easystats*
212 collection of packages^{43–46}. As all the full results have been made available (see **Data**
213 **Availability**), we will focus here on the significant results (based on the Bayes Factor *BF*
214 or the Probability of Direction *pd*⁴⁷).

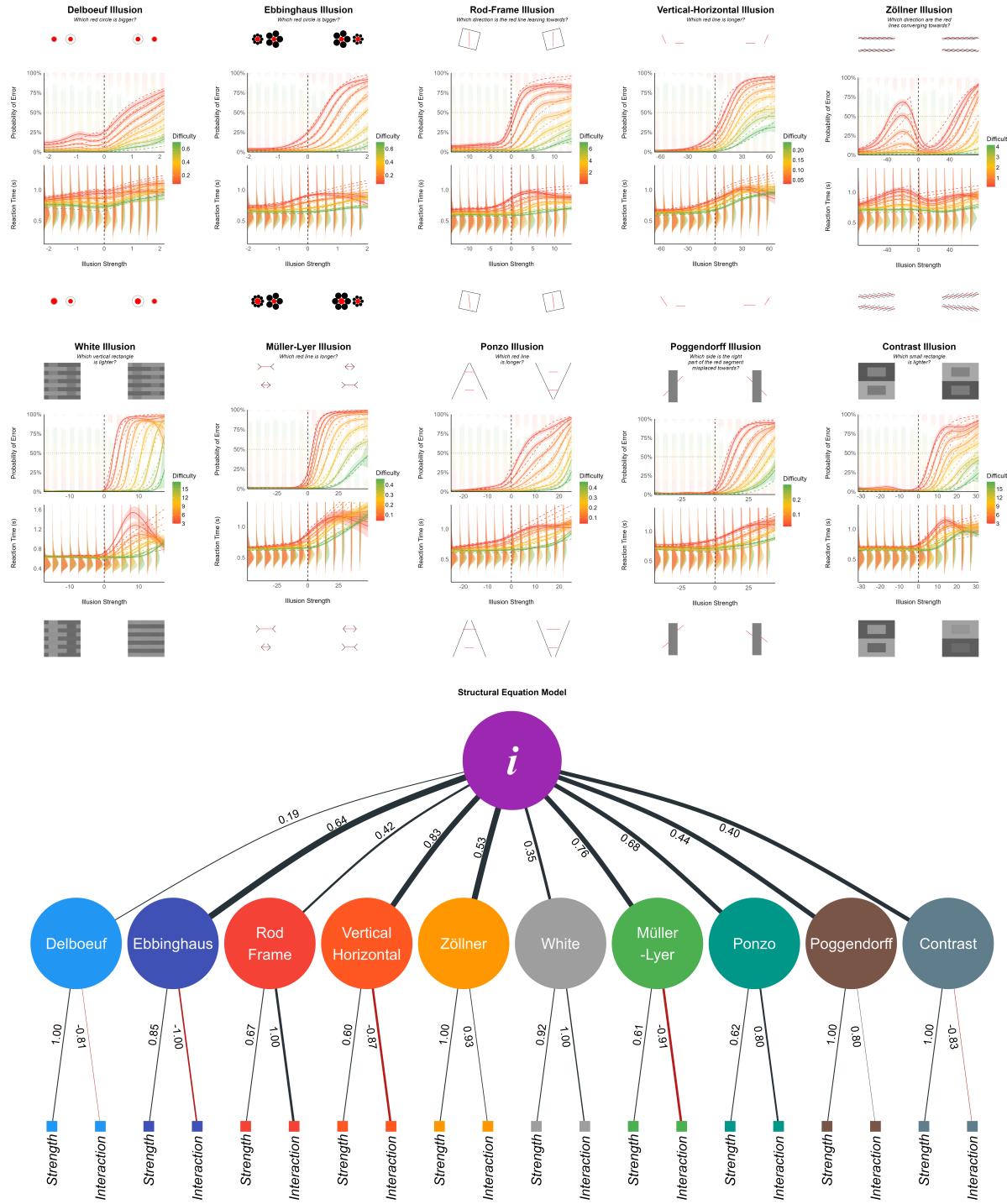


Figure 3. Top: the effect of illusion strength and task difficulty on the error rate and reaction time (RT) for each individual illusion. The solid line represents the General Additive Model (GAM), and the dashed line corresponds to its approximation via linear models. Descriptive data is shown with stacked dots (for which errors start from the top) and distributions for RTs. Negative values for illusion strength correspond to congruent (i.e., facilitating) illusion effects. Task difficulty (the objective difference between the targets of perceptual decision) levels are shown as colors, with lower values corresponding to harder trials. The results for each illusion are surrounded by 4 extreme examples of stimuli, corresponding to lowest to the hardest difficulty (bottom to top) and the weakest to strongest illusion (from left to right for incongruent illusions). Bottom: We extracted the effect slope of the illusion strength and its interaction with task difficulty for each participant. We fitted a Structural Equation Model (SEM) suggesting that these manifest variables group to first-level illusion-specific latent factors, which then load on a general factor of illusion sensitivity (Factor i).

215

Results

216 **Effects of Illusion Strength and Task Difficulty**

217 The best model specifications were $\log(\text{diff}) * \text{strength}$ for Delboeuf;
218 $\sqrt{\text{diff}} * \text{strength}$ for Ebbinghaus; $\log(\text{diff}) * \log(\text{strength})$ for Rod and Frame;
219 $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ for Vertical-Horizontal; $\text{cbrt}(\text{diff}) * \text{strength}$ for Zöllner;
220 $\text{diff} * \sqrt{\text{strength}}$ and $\log(\text{diff}) * \text{strength}$ respectively for errors and RT in White;
221 $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ and $\sqrt{\text{diff}} * \text{strength}$ respectively for errors and RT in
222 Müller-Lyer; $\text{cbrt}(\text{diff}) * \text{strength}$ for Ponzo; $\text{cbrt}(\text{diff}) * \sqrt{\text{strength}}$ and
223 $\text{cbrt}(\text{diff}) * \text{strength}$ respectively for errors and RT in Poggendorff; and
224 $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ for Contrast. For all of these models, the effects of illusion
225 strength, task difficulty and their interaction were significant.

226 For error rates, most of the models closely matched their GAMs counterpart, with
227 the exception of Delboeuf (for which the GAM suggested a non-monotonic effect of illusion
228 strength with a local minimum at 0) and Zöllner (for which theoretically congruent illusion
229 effects were related to increased error rate). A specific discussion regarding these 2 illusions
230 is available in the study documentation (part 1) at
231 <https://github.com/RealityBending/IllusionGameValidation>.

232 For RTs, the GAMs suggested a consistent non-linear relationship between RT and
233 illusion strength: as the illusion strength increases beyond a certain threshold, the
234 participants responded faster. While this is not surprising (strong illusions are likely so
235 effective in biasing perception that it is “easier”, i.e., faster, to make the wrong decision),
236 the linear models were not designed to capture this - likely quadratic - pattern and hence
237 are not good representatives of the underlying dynamics. As such, we decided not to use
238 them for the individual scores analysis.

239 Factor Structure

240 Though imperfect, we believe that the random-slope models capture inter-individual
241 differences with more accuracy (and also provide more conservative estimates due to
242 shrinkage) than basic empirical scores, such as the total number of errors, or the average
243 RT. Thus, for each illusion and within each participant, we extracted the effect of illusion
244 strength and its interaction with task difficulty when the illusion effect was incongruent.
245 These twenty participant-level scores were subjected to exploratory factor analysis (EFA).
246 The Method Agreement Procedure⁴⁸ suggested the presence of 7 latent factors. An oblique
247 (*oblimin* rotation) factor solution explaining 66.69% of variance suggested separate
248 dimensions for the effect of Zöllner, White, Poggendorff, Contrast, Ebbinghaus, Delboeuf,
249 and a common factor for the parameters related to Müller-Lyer, Vertical-Horizontal, Ponzo
250 and Rod and Frame. We submitted these factors to a second-level analysis and extracted
251 two orthogonal (*varimax* rotation) factors. The first factor was loaded by all the previous
252 dimensions with the exception of Delboeuf, which formed its own separate factor.

253 Finally, we tested this data-driven model (*m0*) against four other structural models
254 using structural equation modelling (SEM): one in which the two parameters of each of the
255 10 illusions (illusion strength and interaction with task difficulty) loaded on separate
256 factors, which then all loaded on a common factor (*m1*); one in which the parameters were
257 grouped by illusion type (lines, circles, contrast and angle) before loading on a common
258 factor (*m2*); one in which all the parameters related to strength, and all the parameters
259 related to the interaction loaded onto two respective factors, which then loaded on a
260 common factor (*m3*); and one in which there was no intermediate level: all 20 parameters
261 loaded directly on a common factor (*m4*).

262 The model *m1*, in which the parameters loaded on a first level of 10 illusion-specific
263 factors, which then all loaded on a common factor, significantly outperformed the other
264 models. Its indices of fit ranged from acceptable to satisfactory (CFI = .92; SRMR = .08;

265 NNFI = .91; PNFI = .74; RMSEA = .08), and all the specified effects were significant.
266 The illusion-specific latent factors were loaded positively by the sensitivity to illusion
267 strength, as well as by the interaction effect with task difficulty (with the exception of
268 Delboeuf, Ebbinghaus, Vertical-Horizontal, Müller-Lyer and Contrast, for which the
269 loading was negative). The general factor of illusion sensitivity, labelled Factor i (i- for
270 illusion), explained 48.02% of the total variance of the initial dataset, and was strongly
271 related to Vertical-Horizontal ($\beta_{std.} = 0.83$), Müller-Lyer ($\beta_{std.} = 0.76$), Ponzo
272 ($\beta_{std.} = 0.65$), Ebbinghaus ($\beta_{std.} = 0.64$); moderately to Zöllner ($\beta_{std.} = 0.53$), Poggendorff
273 ($\beta_{std.} = 0.44$), Rod and Frame ($\beta_{std.} = 0.42$), Contrast ($\beta_{std.} = 0.40$) and White
274 ($\beta_{std.} = 0.35$); and weakly to Delboeuf ($\beta_{std.} = 0.19$). We then computed, for each
275 participant, the score for the 10 illusion-specific factors and for the general Factor i .

276 It is important to note that these individual scores are the result of several layers of
277 simplification: 1) the individual coefficient is that of simpler models that sometimes do not
278 perfectly capture the underlying dynamics (especially in the case of Delboeuf and Zöllner);
279 2) we only used the models on error rate, which could be biased by the speed-accuracy
280 decision criterion used by participants; 3) the structural equation model used to compute
281 the scores also incorporated multiple levels of abstractions. Thus, in order to validate the
282 individual scores, we computed the correlation between them and simple empirical scores,
283 such as the average error rate and the mean RT in the task. This analysis revealed strong
284 and significant correlations between each illusion-specific factor and the average amount of
285 errors in its corresponding task. Moreover, each individual score was strongly associated
286 with the average RT across multiple illusion types. This suggests that the individual scores
287 obtained from the structural equation model do capture the sensitivity of each participant
288 to visual illusions, manifesting in both the number of errors and long reaction times.

289 Correlations with Inter-individual Characteristics

290 The Bayesian correlation analysis (with narrow priors centered around a null effect)
291 between the illusion scores and contextual variables (screen size and refresh rate) provided
292 weak evidence in favor of an absence of effect, with the exception of the two contrast-based
293 illusions. Anecdotal ($BF_{10} = 2.05$) and moderate evidence ($BF_{10} = 4.11$) was found for a
294 negative correlation between screen size and the sensitivity to the White and the Contrast
295 illusion, respectively. To test whether this result could be an artifact related to the highly
296 skewed screen size distribution (caused by very few participants with extreme screen sizes),
297 we re-ran a robust correlation (with rank-transformed values), which provided even
298 stronger evidence in favor of the effect existence ($BF_{10} = 28.19$, $BF_{10} = 4.31$ for White and
299 Contrast, respectively).

300 The Bayesian t-tests on the effect of sex suggested anecdotal to moderate evidence in
301 favour of the null effect for all scores, with the exception of the sensitivity to the Zöllner
302 illusion, which was higher in males as compared to females ($\Delta = -0.37$, 95% CI [-0.62,
303 -0.13], $BF_{10} = 12.74$). We fitted Bayesian linear models with the education level entered as
304 a monotonic predictor (appropriate for ordinal variables^{49]}) which yielded no significant
305 effects. For age, we fitted two types of models for each score, one general additive models
306 (GAM) and a 2nd order polynomial model. These consistently suggested a significant
307 positive linear relationship between age and Factor i ($pd = 100\%$), as well as the sensitivity
308 to Müller-Lyer ($pd = 100\%$), Vertical-Horizontal ($pd = 100\%$), Zöllner ($pd = 100\%$) and
309 Ebbinghaus ($pd = 99\%$) illusions (**Figure 4**).

310 Regarding “normal” personality traits, Bayesian correlations suggested substantial
311 evidence in favor of a positive relationship between *Honesty-Humility* and Zöllner
312 ($BF_{10} > 100$), Vertical-Horizontal ($BF_{10} = 9.78$) and the Factor i ($BF_{10} = 4.00$); as well as
313 between *Agreeableness* and Vertical-Horizontal ($BF_{10} = 25.06$), Ponzo ($BF_{10} = 4.88$) and
314 the Factor i ($BF_{10} = 19.65$).

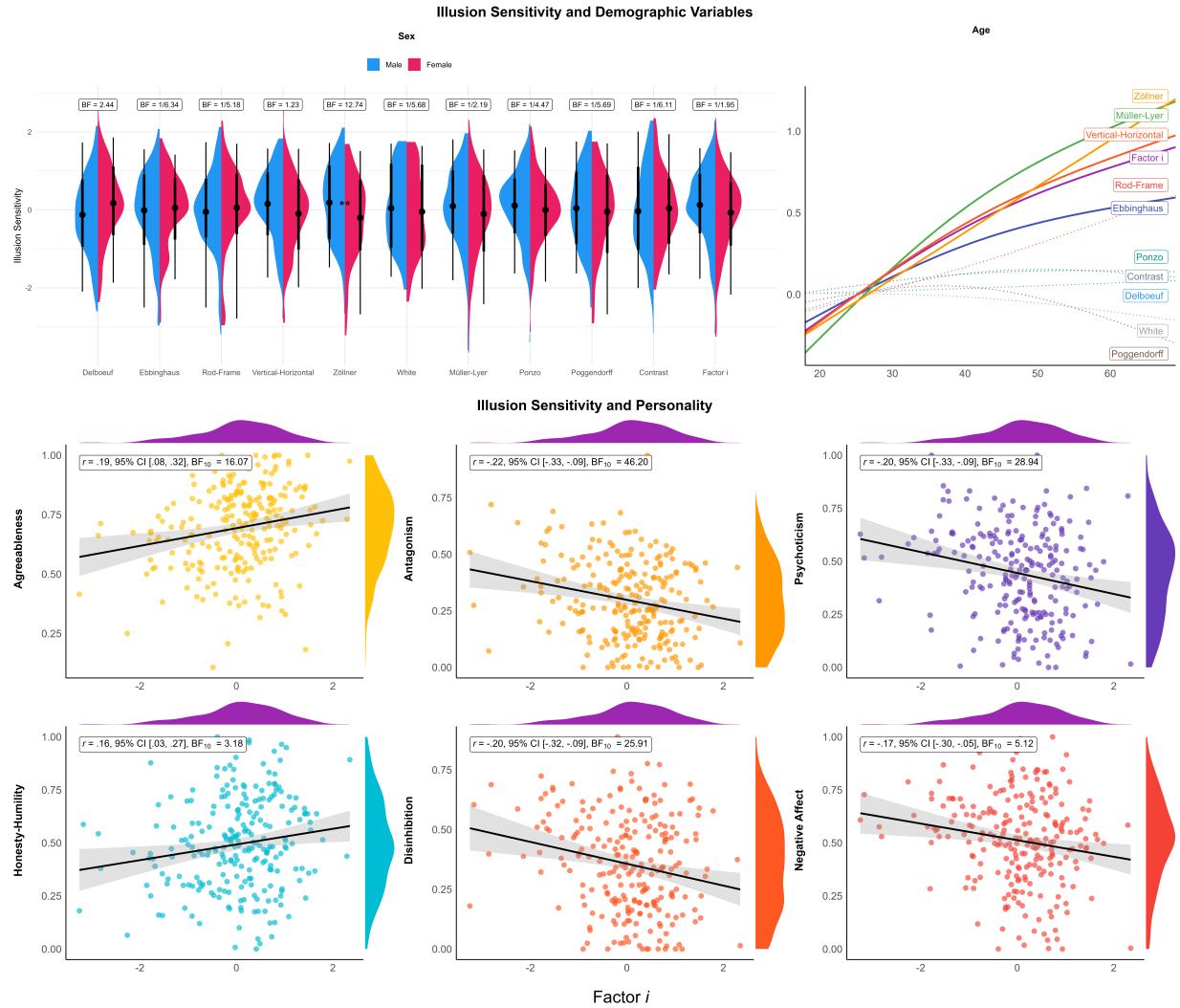


Figure 4. The upper plots show the illusion sensitivity scores as a function of sex and age (solid lines indicate significant relationships). Bottom plots show the correlation between the general factor (Factor *i*) of illusion sensitivity (on the x-axes) and personality traits.

315 Regarding “pathological” personality traits, the results yielded strong evidence in
316 favor of a negative relationship between illusion scores and multiple traits. *Antagonism* was
317 associated with the sensitivity to Vertical-Horizontal ($BF_{10} > 100$), Müller-Lyer
318 ($BF_{10} = 21.57$), Ponzo ($BF_{10} = 17.97$) illusions, and the Factor *i* ($BF_{10} = 55.45$);
319 *Psychoticism* was associated with the sensitivity to Vertical-Horizontal ($BF_{10} = 66.63$) and
320 Müller-Lyer ($BF_{10} = 35.59$) illusions, and the Factor *i* ($BF_{10} = 35.02$); *Disinhibition* was
321 associated with the sensitivity to Vertical-Horizontal ($BF_{10} = 25.38$), Zöllner
322 ($BF_{10} = 7.59$), Müller-Lyer ($BF_{10} = 5.89$) illusions, and the Factor *i* ($BF_{10} = 31.42$); and
323 *Negative Affect* was associated with Zöllner ($BF_{10} = 62.04$), Vertical-Horizontal
324 ($BF_{10} = 12.65$), Müller-Lyer ($BF_{10} = 3.17$), and the Factor *i* ($BF_{10} = 6.39$). The last
325 remaining trait, *Detachment*, did not share any significant relationship with illusion
326 sensitivity.

327

Discussion

328 This study tested a novel illusion sensitivity task paradigm based on the parametric
329 illusion generation framework³³. Using the carefully generated stimuli in a perceptual
330 decision task, we have shown that a gradual modulation of illusion strength is effectively
331 possible across 10 different types of classic visual illusions. Increasing the illusion strength
332 led to an increase in error likelihood, as well as the average and spread of RTs (but only up
333 to a point, after which participants become faster at responding with the wrong answer).
334 Using mixed models, we were able to statistically quantify the effect of illusions for each
335 illusion and each participant separately. This important methodological step opens the
336 door for new illusions-based paradigms and tasks to study the effect of illusions under
337 different conditions and to measure illusion sensitivity using objective behavioral outcomes
338 - such as accuracy or speed - instead of subjective meta-cognitive reports. This new and
339 complementary approach will hopefully help address some of the longstanding literature
340 gaps, as well as cement illusions as valuable stimuli for the study of cognition.

Our findings suggest that the sensitivity to 10 different types of visual illusions share a common part of variance, supporting the existence of a general factor of illusion sensitivity (Factor *i*). This result comes in a field of mixed findings. In fact, contrary to early studies on visual illusions, more recent research have generally not found any significant evidence for a common stable factor across illusions within individuals^{12,15,16,20,50}. Instead, past findings suggest illusory effects are highly specific to the perceptual features of the illusions at stake^{15,20}. It should be noted, however, that most of these studies were low-powered and/or relied on conventional paradigms, such as the adjustment procedure to measure the participants' subjective perception. We believe that our study presents several methodological improvements, including statistical power (high number of trials per participant), homogeneous stimuli (with minimal and highly controlled features) and tasks (decision-making reaction-time task), and a more reliable participant-level score extraction method (based on random-factors models), which in our opinion contributed to the emergence of the common factor.

Finally, we found illusion sensitivity to be positively associated with "positive" personality traits, such as agreeableness and honesty-humility, and negatively associated with maladaptive traits such as antagonism, psychoticism, disinhibition, and negative affect. Although the existing evidence investigating links between illusion sensitivity and personality traits is scarce, these results are consistent with past findings relating pathological egocentric beliefs (often associated with psychoticism⁵¹) to reduced context integration, manifesting in a tendency to separate objects from their surroundings when processing visual stimuli^{19,51,52}. As such, the association between maladaptive traits and lower illusion sensitivity could be linked to a self-centered, decontextualized and disorganized information processing style. Conversely, the relationship between illusion sensitivity and adaptive personality traits is in line with the decreased field dependence (the tendency to rely on external cues in ambiguous contexts) associated with traits negatively correlated with agreeableness and honesty-humility, such as hostility, aggression

368 and narcissism^{18,19,53}.

369 Importantly, these findings highlight the relevance of illusions beyond the field of
370 visual perception, pointing towards an association with high-level domain-general
371 mechanisms. In particular, the evidence in favor of a relationship between maladaptive
372 personality traits and illusion sensitivity is in line with clinical observations, in which a
373 greater resistance to illusions have been reported among patients with schizophrenia^{7,16,53},
374 especially in association with schizotypal traits such as cognitive disorganization^{20,26}.
375 While the search for the exact mechanism(s) underlying these links is an important goal of
376 future research, our findings unlock the potential of illusion-based tasks as sensitive tools to
377 capture specific inter-individual neuro-cognitive differences.

378 Future research is needed to address several limitations. One key question concerns
379 the relationship of illusion sensitivity with perceptual abilities (e.g., using similar tasks, but
380 without illusions). Although the illusions used in the present study did differ in terms of
381 the perceptual task (contrast-based, size-estimation, angle-perception), the possibility of
382 our general factor being driven by inter-individual perceptual skills variability (or other
383 cognitive skills) cannot be discarded. Moreover, using only the error rate models to extract
384 individual-level scores might fail in capturing the whole range of behavioral dynamics.
385 Future work should attempt at integrating the reaction times data (e.g., by jointly
386 analyzing them using drift diffusion models), and assess the psychometric properties - such
387 as stability (e.g., test-retest reliability) and validity - of similar illusion-based paradigms.
388 Finally, while the personality measures used in this study highlight illusion sensitivity as an
389 interesting measure rather than a mere perceptual artifact, further studies should test its
390 relationship with more specific dispositional characteristics (e.g., autistic or schizotypal
391 traits), cognitive styles and abilities, to help understand the potential underlying
392 mechanisms of these associations.

393

Data Availability

394

The datasets generated and/or analysed during the current study are available in the

395

GitHub repository <https://github.com/RealityBending/IllusionGameValidation>

396

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