

**The Illusion Game: A Novel Experimental Paradigm Provides Evidence in
Favour of a General Factor of Visual Illusion Sensitivity and Personality
Correlates**

Dominique Makowski¹, An Shu Te¹, Stephanie Kirk¹, Ngoi Zi Liang¹, & S.H. Annabel
Chen^{1, 2, 3, 4}

¹ School of Social Sciences, Nanyang Technological University, Singapore

² LKC Medicine, Nanyang Technological University, Singapore

³ National Institute of Education, Singapore

⁴ Centre for Research and Development in Learning, Nanyang Technological University,
Singapore

¹² Correspondence concerning this article should be addressed to Dominique Makowski,

¹³ HSS 04-18, 48 Nanyang Avenue, Singapore (dom.makowski@gmail.com).

¹⁴ The authors made the following contributions. Dominique Makowski:

¹⁵ Conceptualization, Data curation, Formal Analysis, Funding acquisition, Investigation,

¹⁶ Methodology, Project administration, Resources, Software, Supervision, Validation,

¹⁷ Visualization, Writing – original draft; An Shu Te: Project administration, Resources,

¹⁸ Investigation, Writing – original draft; Stephanie Kirk: Project administration, Resources,

¹⁹ Writing – original draft; Ngoi Zi Liang: Project administration, Resources, Writing –

²⁰ review & editing; S.H. Annabel Chen: Project administration, Supervision, Writing –

²¹ review & editing.

²² Correspondence concerning this article should be addressed to Dominique Makowski,

²³ HSS 04-18, 48 Nanyang Avenue, Singapore. E-mail: dom.makowski@gmail.com

24

Abstract

25 Visual illusions highlight how the brain uses contextual and prior information to inform our
26 perception of reality. Unfortunately, illusion research has been hampered by the difficulty
27 of adapting these stimuli to experimental settings. In this study, we used the novel
28 parametric framework for visual illusions to generate 10 different classic illusions (Delboeuf,
29 Ebbinghaus, Rod and Frame, Vertical-Horizontal, Zöllner, White, Müller-Lyer, Ponzo,
30 Poggendorff, Contrast) varying in strength, embedded in a perceptual discrimination task.
31 We tested the objective effect of the illusions on errors and reaction times, and extracted
32 participant-level performance scores ($n=250$). Our results provide evidence in favour of a
33 general factor (labelled Factor i) underlying the sensitivity to different illusions. Moreover,
34 we report a positive relationship between illusion sensitivity and personality traits such as
35 Agreeableness, Honesty-Humility, and negative relationships with Psychoticism,
36 Antagonism, Disinhibition, and Negative Affect. The experiment, data and code are fully
37 available for re-use at <https://github.com/RealityBending/IllusionGameValidation>.

38 *Keywords:* visual illusions, illusion game, Pyllusion, personality, general factor

39 Word count: 3621

40 **The Illusion Game: A Novel Experimental Paradigm Provides Evidence in**
41 **Favour of a General Factor of Visual Illusion Sensitivity and Personality**
42 **Correlates**

43 **Introduction**

44 Visual illusions are fascinating stimuli capturing a key feature of our neurocognitive
45 systems. They eloquently show that our brains did not evolve to be perfect perceptual
46 devices providing veridical accounts of physical reality, but integrate prior knowledge and
47 contextual information - blended together in our subjective conscious experience (Carbon,
48 2014). Despite the longstanding interest within the fields of visual perception (Day, 1972;
49 Eagleman, 2001; Gomez-Villa et al., 2022), consciousness science (Caporuscio et al., 2022;
50 Lamme, 2020), and psychiatry (Gori et al., 2016; Notredame et al., 2014; Razeghi et al.,
51 2022; Teufel et al., 2015), several important issues remain open.

52 Notably, the presence of a common mechanism underlying the effects of different
53 illusions has been contested (Cretenoud et al., 2019; Cretenoud et al., 2020; Hamburger,
54 2016); and the nature of the underlying processes - whether related to low-level features of
55 the visual processing system (Cretenoud et al., 2019; Gori et al., 2016) or to top-down
56 influences of prior beliefs (Caporuscio et al., 2022; Teufel et al., 2018) are strongly debated.
57 The existence of dispositional correlates of illusion sensitivity is another area of
58 controversy, with some studies reporting higher illusion resistance in patients with
59 schizophrenia and autism (Giaouri & Alevriadou, 2011; Keane et al., 2014; Notredame et
60 al., 2014; Park et al., 2022; Pessoa et al., 2008), and in individuals with stronger aggression
61 and narcissism traits (Konrath et al., 2009; Zhang et al., 2017).

62 One key challenge hindering the further development of illusion research is the
63 relative difficulty of adapting visual illusions to an experimental setting, which typically
64 requires the controlled modulation of the specific variables of interest. To address this
65 issue, we first developed a parametric framework to manipulate visual illusions that we

66 implemented and made accessible in the open-source software *Pyllusion* (Makowski et al.,
 67 2021). This software allows us to generate different types of classic visual illusions with a
 68 continuous and independent modulation of two parameters: *illusion strength* and *task*
 69 *difficulty* (Figure 1).

Parametric Framework for Visual Illusions

Example with the Müller-Lyer Illusion



The Müller-Lyer Illusion is traditionally presented as two segments (the **red targets**), which perception is biased by the **context** (the arrows). Here, the lower segment appears longer despite being of the same length.

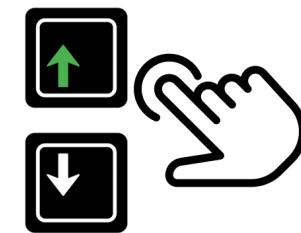


In this illusion, the **task difficulty** corresponds to the difference between the lengths of the red target segments, and the **illusion strength** corresponds to the angle of the arrows.

Example of Stimuli



- ✓ Task difficulty: **easy**
(top line is 2 times longer)
- ✓ Illusion Strength: **strong**
(angle is sharp)
- ← Illusion Direction (left): **incongruent**
(the illusion makes the task harder)
- Illusion Direction (right): **congruent**
(the illusion makes the task easier)



- ✓ Task difficulty: **hard**
(top line is only 1.1 times longer)
- ✓ Illusion Strength: **weak**
(angle is flat)
- ← Illusion Direction (left): **incongruent**
(the illusion makes the task harder)
- Illusion Direction (right): **congruent**
(the illusion makes the task easier)



Task: For these stimuli, the correct response is always the « up » arrow, indicating the longer red segment. We measured the reaction time and the errors (in this case, the « down » arrow).



Stimuli created with the open-source software *Pyllusion* (Makowski et al., 2021)

Figure 1. The parametric framework for visual illusions (Makowski et al., 2021) applied to the Müller-Lyer illusion (above). Below are examples of stimuli showcasing the manipulation of two parameters, task difficulty and illusion strength.

70 Indeed, many visual illusions can be seen as being composed of *targets* (e.g.,

71 same-length lines), of which perception is biased by the *context* (e.g., in the Müller-Lyer
72 illusion, the same-length line segments appear to have different lengths if they end with
73 inwards vs. outwards pointing arrows). Past illusion studies traditionally employed
74 paradigms focusing on participants' subjective experience, by asking them the extent to
75 which they perceive two identical targets as different (Lányi et al., 2022), or having them
76 adjust the targets to match a reference stimulus relying only on their perception
77 (Grzeczkowski et al., 2018; Mylniec & Bednarek, 2016). Alternatively, *Pyllusion* allows the
78 creation of illusions in which the targets are objectively different (e.g., one segment is truly
79 more or less longer than the other), and in which the illusion varies in strength (the biasing
80 angle of the arrows is more or less acute).

81 This opens the door for an experimental task in which participants make perceptual
82 judgments about the targets (e.g., which segment is the longest) under different conditions
83 of objective difficulty and illusion strength. Moreover, the illusion effect can be either
84 “incongruent” (making the task more difficult by biasing the perception in the opposite
85 way) or “congruent” (making the task easier). Although visual illusions are inherently tied
86 to subjective perception, this framework allows a reversal of the traditional paradigm to
87 potentially quantify the “objective” effect of illusions by measuring its behavioral effect
88 (error rate and reaction times) on the performance in a perceptual task.

89 The aim of the present preregistered study is three-fold. First, we will test this novel
90 paradigm by investigating if the effect of illusion strength and task difficulty can be
91 manipulated continuously for 10 different classic illusions (Delboeuf, Ebbinghaus, Rod and
92 Frame, Vertical-Horizontal, Zöllner, White, Müller-Lyer, Ponzo, Poggendorff, Contrast).
93 Next, we will investigate the latent factors structure of participant-level performance scores
94 to these illusions and test the existence of a common factor of illusion sensitivity. Finally,
95 we will explore how illusion sensitivity relates to demographic characteristics, contextual
96 variables (pertaining to the experiment setting), and personality traits.

97 In line with open-science standards, all the material (stimuli generation code,
98 experiment code, raw data, analysis script with complementary figures and analyses,
99 preregistration, etc.) is available as **Supplementary Materials** at
100 <https://github.com/RealityBending/IllusionGameValidation>.

101 **Methods**

102 **Stimuli**

103 A pilot study ($n = 46$), of which full description is available in the Supplementary
104 Materials, was first conducted to determine a sensitive range of stimuli parameters. Then,
105 for each of the 10 illusion types, we generated a total of 134 stimuli. These stimuli resulted
106 from the combination of 15 equally-spaced levels of illusion *strength* (7 negative, i.e.,
107 congruent effects; 7 positive, i.e., incongruent effects; and 0) overlapped with 16
108 non-linearly spaced task *difficulty* levels (i.e., with an exponential, square or cubic spacing
109 depending on the pilot results). For instance, a linear space of [0.1, 0.4, 0.7, 1.0] can be
110 transformed to an exponential space of [0.1, 0.34, 0.64, 1.0], where 0.1 corresponds to the
111 highest difficulty - i.e., the smallest objective difference between targets). For each illusion
112 type, the stimuli were split into two series (56 and 72 stimuli per series) with alternating
113 parameter values to maintain their homogeneity. Additionally, 6 stimuli per illusion type
114 was generated for a practice series, with more extreme variations (i.e., containing very easy
115 trials to help cement the task instructions).

116 **Procedure**

117 After a brief demographic survey and a practice series of illusions, the first series of
118 10 illusion blocks was presented in a randomized order, with a further randomization of the
119 stimuli order within each block. Following this first series of blocks, two personality
120 questionnaires were administered, the *IPIP6* (24 items, Sibley et al., 2011) - measuring 6
121 “normal” personality traits (Extraversion, Openness, Conscientiousness, Agreeableness,

122 Neuroticism and Honesty-Humility), and the *PID-5* (25 items, Hopwood et al., 2012) -
123 measuring 5 “pathological” personality traits (Disinhibition, Antagonism, Detachment,
124 Negative Affect and Psychoticism). Next, the second series of 10 illusion blocks was
125 presented (with new randomized orders of blocks and trials). In total, each participant
126 underwent 1340 trials of which they had to respond “as fast as possible without making
127 errors” (i.e., an explicit double constraint to mitigate the inter-individual variability in the
128 speed-accuracy trade off) by pressing the correct arrow key (left/right, or up/down
129 depending on the illusion type). For instance, in the Müller-Lyer block, participants had to
130 answer which one of the upper or bottom target line was the longest. The task was
131 implemented using *jsPsych* (De Leeuw, 2015), and the set of instructions for each illusion
132 type is available in the experiment code.

133 Participants

134 Participants were recruited via *Prolific*, a crowd-sourcing platform recognized for
135 providing high quality data (Peer et al., 2022). The only inclusion criterion was a fluent
136 proficiency in English to ensure that the task instructions would be well-understood.
137 Participants were incentivised with a reward of about £7.5 for completing the task, which
138 took about 50 minutes to finish. Demographic variables (age, gender, and ethnicity) were
139 self-reported on a voluntary basis.

140 We excluded 6 participants upon inspection of the average error rate (when close to
141 50%, suggesting random answers), and reaction time distribution (when implausibly fast).
142 For the remaining participants, we discarded blocks with more than 50% of errors (2.16%
143 of trials), possibly indicating that instructions were misunderstood (e.g., participants
144 focused on the shorter line instead of the longer one), and 0.76% trials with extreme
145 response times (< 125 ms or > 4 SD above mean). Additionally, due to a technical issue,
146 no personality data was recorded for the first eight participants.

¹⁴⁷ The final sample included 250 participants (Mean age = 26.5, SD = 7.6, range: [18,
¹⁴⁸ 69]; Sex: 48% females, 52% males).

¹⁴⁹ **Data Analysis**

¹⁵⁰ The first part of the analysis focused on modelling the effect of illusion strength and
¹⁵¹ task difficulty on errors and reaction time (RT) within each illusion. We started by fitting
¹⁵² General Additive Models (GAMs), which can parsimoniously accommodate possible
¹⁵³ non-linear effects and interactions. Errors were analyzed using Bayesian logistic mixed
¹⁵⁴ models, and RTs of correct responses were analyzed using an ex-Gaussian family with the
¹⁵⁵ same fixed effects entered for the location μ (mean), scale σ (spread) and tail-dominance τ
¹⁵⁶ of the RT distribution (Balota & Yap, 2011; Matzke & Wagenmakers, 2009).

¹⁵⁷ Using GAMs as the “ground-truth” models, we attempted at approximating them
¹⁵⁸ using general linear mixed models, which can be used to estimate the effects'
¹⁵⁹ participant-level variability (via random slopes). Following a comparison of models with a
¹⁶⁰ combination of transformations (raw, log, square root or cubic root) on the main predictors
¹⁶¹ (task *difficulty* and illusion *strength*), we fitted the best model (based on their indices of
¹⁶² fit), and compared their output visually (**Figure 2**).

¹⁶³ The inter-individual variability in the effect of illusion strength and its interaction
¹⁶⁴ with task difficulty was extracted from the models and used as participant-level scores. We
¹⁶⁵ then explored the relationship of these indices across different illusions using exploratory
¹⁶⁶ factor analysis (EFA) and structural equation modelling (SEM), and tested the existence of
¹⁶⁷ a general of illusion sensitivity (Factor i).

¹⁶⁸ Finally, for each of the individual illusion sensitivity scores (10 illusion-specific factors
¹⁶⁹ and the general Factor i), we tested the effect of contextual variables (screen size, screen
¹⁷⁰ refresh rate), demographic variables (sex, education, age), and personality traits.

The analysis was carried out using *R* 4.2 (R Core Team, 2022), *brms* (Bürkner, 2017), the *tidyverse* (Wickham et al., 2019), and the *easystats* collection of packages (Lüdecke et al., 2021, 2019; Makowski et al., 2020; Makowski, Ben-Shachar, & Lüdecke, 2019). As the full results are available as supplementary materials, we will focus here on the significant results (based on the Bayes Factor *BF* or the Probability of Direction *pd*, see Makowski, Ben-Shachar, Chen, et al., 2019).

Results

Effects of Illusion Strength and Task Difficulty

The best model specifications were $\log(\text{diff}) * \text{strength}$ for Delboeuf; $\sqrt{\text{diff}} * \text{strength}$ for Ebbinghaus; $\log(\text{diff}) * \log(\text{strength})$ for Rod and Frame; $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ for Vertical-Horizontal; $\text{cbrt}(\text{diff}) * \text{strength}$ for Zöllner; $\text{diff} * \sqrt{\text{strength}}$ and $\log(\text{diff}) * \text{strength}$ respectively for errors and RT in White; $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ and $\sqrt{\text{diff}} * \text{strength}$ respectively for errors and RT in Müller-Lyer; $\text{cbrt}(\text{diff}) * \text{strength}$ for Ponzo; $\text{cbrt}(\text{diff}) * \sqrt{\text{strength}}$ and $\text{cbrt}(\text{diff}) * \text{strength}$ respectively for errors and RT in Poggendorff; and $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ for Contrast. For all of these models, the effects of illusion strength, task difficulty and their interaction were significant.

For error rates, most of the models closely matched their GAMs counterpart, with the exception of Delboeuf (for which the GAM suggested a non-monotonic effect of illusion strength with a local minimum at 0) and Zöllner (for which theoretically congruent illusion effects were related to increased error rate). A specific discussion regarding these 2 illusions is available in the Supplementary Materials (Part 1 - Discussion).

For RTs, the GAMs suggested a consistent non-linear relationship between RT and illusion strength: as the illusion strength increases beyond a certain threshold, the participants responded faster. While this is not surprising (strong illusions are likely so

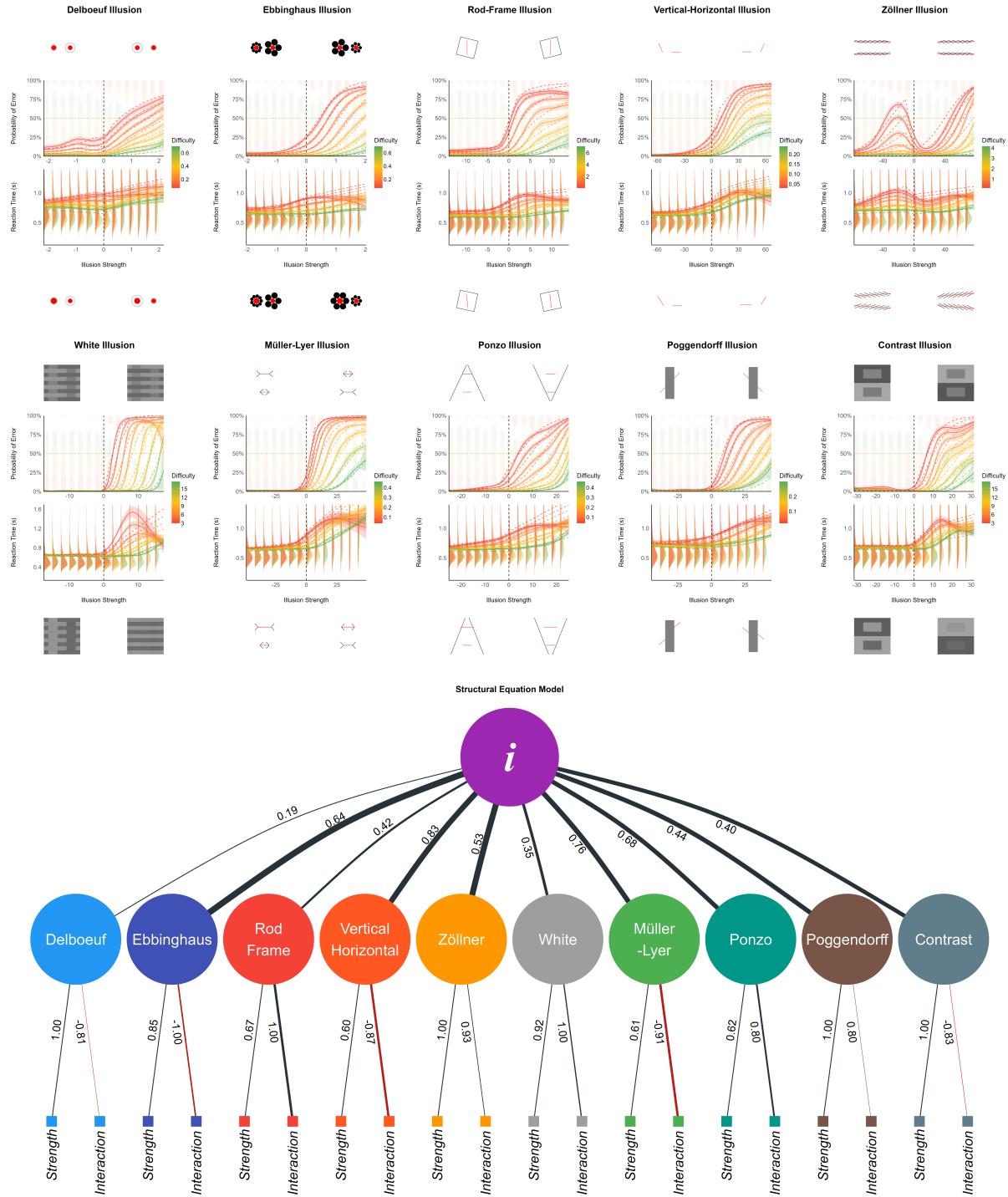


Figure 2. Top: the effect of illusion strength and task difficulty on the error rate and reaction time (RT) for each individual illusion. The solid line represents the General Additive Model (GAM), and the dashed line corresponds to its approximation via linear models. Descriptive data is shown with stacked dots (for which errors start from the top) and distributions for RTs. Negative values for illusion strength correspond to congruent (i.e., facilitating) illusion effects. Task difficulty (the objective difference between the targets of perceptual decision) levels are shown as colors, with lower values corresponding to harder trials. The results for each illusion are surrounded by 4 extreme examples of stimuli, corresponding to the hardest difficulty (on top) and the strongest illusion (on the right for incongruent illusions). Bottom: We extracted the effect slope of the illusion strength and its interaction with task difficulty for each participant. We fitted a Structural Equation Model (SEM) suggesting that these manifest variables group to first-level illusion-specific latent factors, which then load on a general factor of illusion sensitivity (Factor i).

196 effective in biasing perception that it is “easier”, i.e., faster, to make the wrong decision),
197 the linear models were not designed to capture this - likely quadratic - pattern and hence
198 are not good representatives of the underlying dynamics. As such, we decided not to use
199 them for the individual scores analysis.

200 **Factor Structure**

201 Though imperfect, we believe that the random-slope models capture inter-individual
202 differences with more accuracy (and also provide more conservative estimates due to
203 shrinkage) than basic empirical scores, such as the total number of errors, or the average
204 RT. Thus, for each illusion and within each participant, we extracted the effect of illusion
205 strength and its interaction with task difficulty when the illusion effect was incongruent.
206 These twenty participant-level scores were subjected to exploratory factor analysis (EFA).
207 The Method Agreement Procedure (Lüdecke et al., 2020) suggested the presence of 7 latent
208 factors. An oblique (*oblimin* rotation) factor solution explaining 66.69% of variance
209 suggested separate dimensions for the effect of Zöllner, White, Poggendorff, Contrast,
210 Ebbinghaus, Delboeuf, and a common factor for the parameters related to Müller-Lyer,
211 Vertical-Horizontal, Ponzo and Rod and Frame. We submitted these factors to a
212 second-level analysis and extracted two orthogonal (*varimax* rotation) factors. The first
213 factor was loaded by all the previous dimensions with the exception of Delboeuf, which
214 formed its own separate factor.

215 Finally, we tested this data-driven model (*m0*) against four other structural models
216 using structural equation modelling (SEM): one in which the two parameters of each of the
217 10 illusions (illusion strength and interaction with task difficulty) loaded on separate
218 factors, which then all loaded on a common factor (*m1*); one in which the parameters were
219 grouped by illusion type (lines, circles, contrast and angle) before loading on a common
220 factor (*m2*); one in which all the parameters related to strength, and all the parameters
221 related to the interaction loaded onto two respective factors, which then loaded on a

common factor ($m3$); and one in which there was no intermediate level: all 20 parameters loaded directly on a common factor ($m4$).

The model $m1$, in which the parameters loaded on a first level of 10 illusion-specific factors, which then all loaded on a common factor, significantly outperformed the other models. Its indices of fit ranged from acceptable to satisfactory (CFI = .92; SRMR = .08; NNFI = .91; PNFI = .74; RMSEA = .08), and all the specified effects were significant. The illusion-specific latent factors were loaded positively by the sensitivity to illusion strength, as well as by the interaction effect with task difficulty (with the exception of Delboeuf, Ebbinghaus, Vertical-Horizontal, Müller-Lyer and Contrast, for which the loading was negative). The general factor of illusion sensitivity, labelled Factor i (i- for illusion), explained 48.02% of the total variance of the initial dataset, and was strongly related to Vertical-Horizontal ($\beta_{std.} = 0.83$), Müller-Lyer ($\beta_{std.} = 0.76$), Ponzo ($\beta_{std.} = 0.65$), Ebbinghaus ($\beta_{std.} = 0.64$); moderately to Zöllner ($\beta_{std.} = 0.53$), Poggendorff ($\beta_{std.} = 0.44$), Rod and Frame ($\beta_{std.} = 0.42$), Contrast ($\beta_{std.} = 0.40$) and White ($\beta_{std.} = 0.35$); and weakly to Delboeuf ($\beta_{std.} = 0.19$). We then computed, for each participant, the score for the 10 illusion-specific factors and for the general Factor i .

It is important to note that these individual scores are the result of several layers of simplification: 1) the individual coefficient is that of simpler models that sometimes do not perfectly capture the underlying dynamics (especially in the case of Delboeuf and Zöllner); 2) we only used the models on error rate, which could be biased by the speed-accuracy decision criterion used by participants; 3) the structural equation model used to compute the scores also incorporated multiple levels of abstractions. Thus, in order to validate the individual scores, we computed the correlation between them and simple empirical scores, such as the average error rate and the mean RT in the task. This analysis revealed strong and significant correlations between each illusion-specific factor and the average amount of errors in its corresponding task. Moreover, each individual score was strongly associated

248 with the average RT across multiple illusion types. This suggests that the individual scores
249 obtained from the structural equation model do capture the sensitivity of each participant
250 to visual illusions, manifesting in both the number of errors and long reaction times.

251 **Correlations with Inter-individual Characteristics**

252 The Bayesian correlation analysis (with narrow priors centered around a null effect)
253 between the illusion scores and contextual variables (screen size and refresh rate) provided
254 weak evidence in favor of an absence of effect, with the exception of the two contrast-based
255 illusions (**Figure 3**). Anecdotal ($BF_{10} = 2.05$) and moderate evidence ($BF_{10} = 4.11$) was
256 found for a negative correlation between screen size and the sensitivity to the White and
257 the Contrast illusion, respectively. To test whether this result could be an artifact related
258 to the highly skewed screen size distribution (caused by very few participants with extreme
259 screen sizes), we re-ran a robust correlation (with rank-transformed values), which provided
260 even stronger evidence in favor of the effect existence ($BF_{10} = 28.19$, $BF_{10} = 4.31$ for
261 White and Contrast, respectively).

262 The Bayesian t-tests on the effect of sex suggested anecdotal to moderate evidence in
263 favour of the null effect for all scores, with the exception of the sensitivity to the Zöllner
264 illusion, which was higher in males as compared to females ($\Delta = -0.37$, 95% CI [-0.62,
265 -0.13], $BF_{10} = 12.74$). We fitted Bayesian linear models with the education level entered as
266 a monotonic predictor (appropriate for ordinal variables, Bürkner & Charpentier, 2020),
267 which yielded no significant effects. For age, we fitted two types of models for each score,
268 one general additive models (GAM) and a 2nd order polynomial model. These consistently
269 suggested a significant positive linear relationship between age and Factor i ($pd = 100\%$),
270 as well as the sensitivity to Müller-Lyer ($pd = 100\%$), Vertical-Horizontal ($pd = 100\%$),
271 Zöllner ($pd = 100\%$) and Ebbinghaus ($pd = 99\%$) illusions.

272 Regarding “normal” personality traits, Bayesian correlations suggested substantial

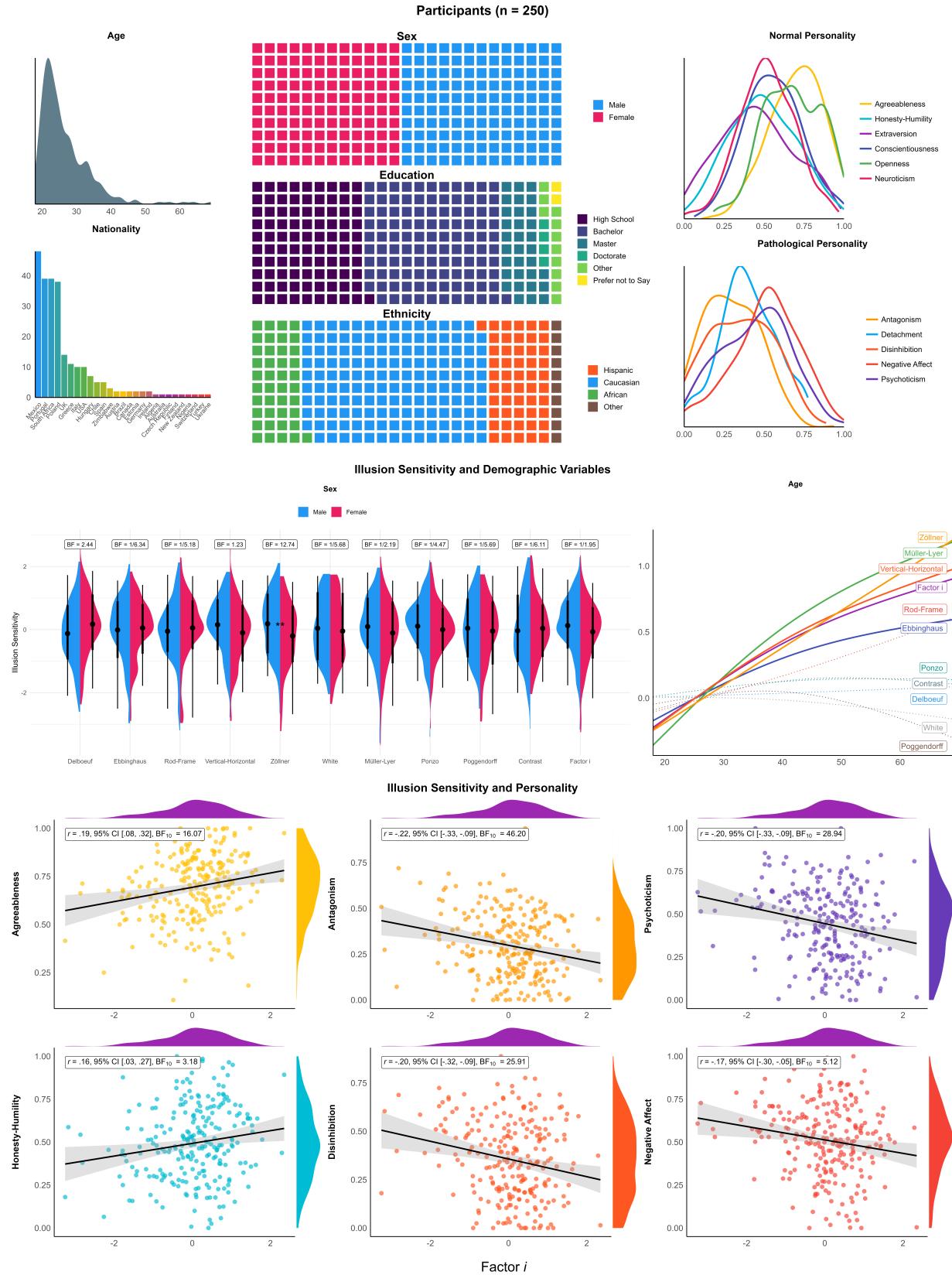


Figure 3. The upper plots show the distribution of demographic and dispositional variables. The middle plots shows the illusion sensitivity scores as a function of sex and age (solid lines indicate significant relationships). Bottom plots show the correlation between the general factor of illusion sensitivity (Factor i) and personality traits.

273 evidence in favor of a positive relationship between *Honesty-Humility* and Zöllner
274 ($BF_{10} > 100$), Vertical-Horizontal ($BF_{10} = 9.78$) and the Factor *i* ($BF_{10} = 4.00$); as well as
275 between *Agreeableness* and Vertical-Horizontal ($BF_{10} = 25.06$), Ponzo ($BF_{10} = 4.88$) and
276 the Factor *i* ($BF_{10} = 19.65$).

277 Regarding “pathological” personality traits, the results yielded strong evidence in
278 favor of a negative relationship between illusion scores and multiple traits. *Antagonism* was
279 associated with the sensitivity to Vertical-Horizontal ($BF_{10} > 100$), Müller-Lyer
280 ($BF_{10} = 21.57$), Ponzo ($BF_{10} = 17.97$) illusions, and the Factor *i* ($BF_{10} = 55.45$);
281 *Psychoticism* was associated with the sensitivity to Vertical-Horizontal ($BF_{10} = 66.63$) and
282 Müller-Lyer ($BF_{10} = 35.59$) illusions, and the Factor *i* ($BF_{10} = 35.02$); *Disinhibition* was
283 associated with the sensitivity to Vertical-Horizontal ($BF_{10} = 25.38$), Zöllner
284 ($BF_{10} = 7.59$), Müller-Lyer ($BF_{10} = 5.89$) illusions, and the Factor *i* ($BF_{10} = 31.42$); and
285 *Negative Affect* was associated with Zöllner ($BF_{10} = 62.04$), Vertical-Horizontal
286 ($BF_{10} = 12.65$), Müller-Lyer ($BF_{10} = 3.17$), and the Factor *i* ($BF_{10} = 6.39$). The last
287 remaining trait, *Detachment*, did not share any significant relationship with illusion
288 sensitivity. See Supplementary Materials (Part 2 - Discussion) for a detailed discussion
289 regarding these associations.

290

Discussion

291 The parametric illusion generation framework developed in Makowski et al. (2021)
292 proposes to conceptualize illusions as composed of targets and distractors that can be
293 manipulated independently and continuously. In the present study, we have shown that
294 such gradual modulation of illusion strength is effectively possible across 10 different types
295 of classic visual illusions. Specifically, increasing the illusion strength led to an increase in
296 the likelihood of errors, as well as the average and spread of RTs (but only up to a point,
297 after which participants become faster at responding with the wrong answer). This
298 important methodological step opens the door for new illusions-based paradigms and tasks

299 to study the effect of illusions under different conditions and to quantify illusion sensitivity
300 using objective behavioral outcomes - such as accuracy or speed - instead of subjective
301 meta-cognitive reports. This new and complementary approach will hopefully help address
302 some of the longstanding literature gaps, as well as cement illusions as valuable stimuli for
303 the study of cognition.

304 Our findings suggest that the sensitivity to 10 different types of visual illusions share a
305 common part of variance, supporting the existence of a general factor of illusion sensitivity
306 (Factor *i*). This result comes in a field of mixed findings. In fact, contrary to early studies
307 on visual illusions, more recent research have generally not found any significant evidence
308 for a common stable factor across illusions within individuals (Cretenoud et al., 2019;
309 Cretenoud et al., 2020; Grzeczkowski et al., 2017, 2018; Yang et al., 2012). Instead, past
310 findings suggest illusory effects are highly specific to the perceptual features of the illusions
311 at stake (Cretenoud et al., 2019; Grzeczkowski et al., 2017). It should be noted, however,
312 that most of these studies were low-powered and/or relied on conventional paradigms, such
313 as the adjustment procedure to measure the participants' subjective perception. We believe
314 that our study presents several methodological improvements, including statistical power
315 (high number of trials per participant), homogeneous stimuli (with minimal and highly
316 controlled features) and tasks (decision-making reaction-time task), and a more reliable
317 participant-level score extraction method (based on random-factors models), which in our
318 opinion contributed to the emergence of the common factor.

319 However, although the illusions did differ in terms of the perceptual task
320 (contrast-based, size-estimation, angle-perception), the possibility of our general factor
321 being driven by inter-individual perceptual skills variability (or other cognitive skills)
322 cannot be discarded. Future studies should investigate the relationship of illusion
323 sensitivity with perceptual abilities (e.g., using similar tasks, but without illusions), and
324 assess the psychometric properties - such as stability (e.g., test-retest reliability) and

325 validity - of similar illusion-based paradigms.

326 Finally, we found the sensitivity to illusions to be positively associated with
327 “positive” personality traits, such as agreeableness and honesty-humility, and negatively
328 associated with maladaptive traits such as antagonism, psychoticism, disinhibition, and
329 negative affect. Although the existing evidence investigating links between illusion
330 sensitivity and personality traits is scarce, these results are consistent with past findings
331 relating pathological egocentric beliefs (often associated with psychoticism, Fox, 2006) to
332 reduced context integration, manifesting in a tendency to separate objects from their
333 surroundings when processing visual stimuli (Fox, 2006; Konrath et al., 2009; Ohmann &
334 Burgmer, 2016). As such, the association between maladaptive traits and lower illusion
335 sensitivity could be linked to a self-centered, decontextualized and disorganized information
336 processing style. Conversely, the relationship between illusion sensitivity and adaptive
337 personality traits is in line with the decreased field dependence (the tendency to rely on
338 external cues in ambiguous contexts) associated with traits negatively correlated with
339 agreeableness and honesty-humility, such as hostility, aggression and narcissism (Konrath
340 et al., 2009; Pessoa et al., 2008; Zhang et al., 2017).

341 Importantly, these findings highlight the relevance of illusions beyond the field of
342 visual perception, pointing towards an association with high-level domain-general
343 mechanisms. In particular, the evidence in favor of a relationship between maladaptive
344 personality traits and illusion sensitivity is in line with clinical observations, in which a
345 greater resistance to illusions have been reported among patients with schizophrenia
346 (Grzeczkowski et al., 2018; Notredame et al., 2014; Pessoa et al., 2008), especially in
347 association with schizotypal traits such as cognitive disorganization (Cretenoud et al., 2019;
348 Lányi et al., 2022). While the search for the exact mechanism(s) underlying these links is
349 an important goal of future research, our findings unlock the potential of illusion-based
350 tasks as sensitive tools to capture specific inter-individual neuro-cognitive differences.

351 In conclusion, we strongly invite researchers to explore and re-analyze our dataset
352 with other approaches and methods to push the understanding of visual illusions and
353 illusion sensitivity further. The task, data and analysis script are available in open-access
354 at <https://github.com/RealityBending/IllusionGameValidation>.

355 **Acknowledgments**

356 We would like to thank Zen J. Lau, Tam Pham, and W. Paul Boyce for their
357 contribution to *Pyllusion*, as well as Prof Dólos for the inspiration.

358

References

- 359 Balota, D. A., & Yap, M. J. (2011). Moving beyond the mean in studies of mental
360 chronometry: The power of response time distributional analyses. *Current Directions in
361 Psychological Science*, 20(3), 160–166.
- 362 Bürkner, P.-C. (2017). brms: An R package for Bayesian multilevel models using Stan.
363 *Journal of Statistical Software*, 80(1), 1–28. <https://doi.org/10.18637/jss.v080.i01>
- 364 Bürkner, P.-C., & Charpentier, E. (2020). Modelling monotonic effects of ordinal
365 predictors in bayesian regression models. *British Journal of Mathematical and
366 Statistical Psychology*, 73(3), 420–451.
- 367 Caporuscio, C., Fink, S. B., Sterzer, P., & Martin, J. M. (2022). When seeing is not
368 believing: A mechanistic basis for predictive divergence. *Consciousness and Cognition*,
369 102, 103334. <https://doi.org/10.1016/j.concog.2022.103334>
- 370 Carbon, C.-C. (2014). Understanding human perception by human-made illusions.
371 *Frontiers in Human Neuroscience*, 8.
372 <https://www.frontiersin.org/articles/10.3389/fnhum.2014.00566>
- 373 Cretenoud, A. F., Francis, G., & Herzog, M. H. (2020). When illusions merge. *Journal of
374 Vision*, 20(8), 12–12.
- 375 Cretenoud, A. F., Karimpur, H., Grzeczkowski, L., Francis, G., Hamburger, K., & Herzog,
376 M. H. (2019). Factors underlying visual illusions are illusion-specific but not
377 feature-specific. *Journal of Vision*, 19(14), 12. <https://doi.org/10.1167/19.14.12>
- 378 Day, R. H. (1972). Visual Spatial Illusions: A General Explanation: A wide range of visual
379 illusions, including geometrical distortions, can be explained by a single principle.
380 *Science*, 175(4028), 1335–1340. <https://doi.org/10.1126/science.175.4028.1335>
- 381 De Leeuw, J. R. (2015). jsPsych: A JavaScript library for creating behavioral experiments
382 in a web browser. *Behavior Research Methods*, 47(1), 1–12.
- 383 Eagleman, D. M. (2001). Visual illusions and neurobiology. *Nature Reviews Neuroscience*,
384 2(12), 920–926. <https://doi.org/10.1038/35104092>

- 385 Fox, A. (2006). *Adolescent self-development and psychopathology: Anorexia nervosa and*
386 *psychosis* [PhD thesis].
- 387 Giaouri, S., & Alevriadou, A. (2011). Are children with down syndrome susceptible to
388 visual illusions? *Procedia - Social and Behavioral Sciences*, 15, 1988–1992.
389 <https://doi.org/10.1016/j.sbspro.2011.04.040>
- 390 Gomez-Villa, A., Martín, A., Vazquez-Corral, J., Bertalmío, M., & Malo, J. (2022). On the
391 synthesis of visual illusions using deep generative models. *Journal of Vision*, 22(8), 2.
392 <https://doi.org/10.1167/jov.22.8.2>
- 393 Gori, S., Molteni, M., & Facoetti, A. (2016). Visual illusions: An interesting tool to
394 investigate developmental dyslexia and autism spectrum disorder. *Frontiers in Human*
395 *Neuroscience*, 10, 175. <https://doi.org/10.3389/fnhum.2016.00175>
- 396 Grzeczkowski, L., Clarke, A. M., Francis, G., Mast, F. W., & Herzog, M. H. (2017). About
397 individual differences in vision. *Vision Research*, 141, 282–292.
398 <https://doi.org/10.1016/j.visres.2016.10.006>
- 399 Grzeczkowski, L., Roinishvili, M., Chkonia, E., Brand, A., Mast, F. W., Herzog, M. H., &
400 Shaqiri, A. (2018). Is the perception of illusions abnormal in schizophrenia? *Psychiatry*
401 *Research*, 270, 929–939. <https://doi.org/10.1016/j.psychres.2018.10.063>
- 402 Hamburger, K. (2016). Visual Illusions Based on Processes: New Classification System
403 Needed: *Perception*. <https://doi.org/10.1177/0301006616629038>
- 404 Hopwood, C. J., Thomas, K. M., Markon, K. E., Wright, A. G. C., & Krueger, R. F.
405 (2012). DSM-5 personality traits and DSM-IV personality disorders. *Journal of*
406 *Abnormal Psychology*, 121(2), 424–432. <https://doi.org/10.1037/a0026656>
- 407 Keane, B. P., Joseph, J., & Silverstein, S. M. (2014). Late, not early, stages of Kanizsa
408 shape perception are compromised in schizophrenia. *Neuropsychologia*, 56, 302–311.
409 <https://doi.org/10.1016/j.neuropsychologia.2014.02.001>
- 410 Konrath, S., Bushman, B. J., & Grove, T. (2009). Seeing my world in a million little
411 pieces: Narcissism, self-construal, and cognitive-perceptual style. *Journal of*

- 412 *Personality*, 77(4), 1197–1228.
- 413 Lamme, V. A. F. (2020). Visual functions generating conscious seeing. *Frontiers in*
414 *Psychology*, 11. <https://www.frontiersin.org/articles/10.3389/fpsyg.2020.00083>
- 415 Lányi, O., Keri, S., Pálffy, Z., & Polner, B. (2022). *Can you believe your eyes? Positive*
416 *schizotypy is associated with increased susceptibility to the müller-lyer illusion.*
- 417 Lüdecke, D., Ben-Shachar, M., Patil, I., & Makowski, D. (2020). Extracting, computing
418 and exploring the parameters of statistical models using R. *Journal of Open Source*
419 *Software*, 5(53), 2445. <https://doi.org/10.21105/joss.02445>
- 420 Lüdecke, D., Ben-Shachar, M., Patil, I., Waggoner, P., & Makowski, D. (2021).
421 performance: An R package for assessment, comparison and testing of statistical
422 models. *Journal of Open Source Software*, 6(60), 3139.
423 <https://doi.org/10.21105/joss.03139>
- 424 Lüdecke, D., Waggoner, P., & Makowski, D. (2019). Insight: A unified interface to access
425 information from model objects in R. *Journal of Open Source Software*, 4(38), 1412.
426 <https://doi.org/10.21105/joss.01412>
- 427 Makowski, D., Ben-Shachar, M. S., Chen, S. A., & Lüdecke, D. (2019). Indices of effect
428 existence and significance in the bayesian framework. *Frontiers in Psychology*, 10, 2767.
- 429 Makowski, D., Ben-Shachar, M., & Lüdecke, D. (2019). bayestestR: Describing effects and
430 their uncertainty, existence and significance within the Bayesian framework. *Journal of*
431 *Open Source Software*, 4(40), 1541. <https://doi.org/10.21105/joss.01541>
- 432 Makowski, D., Ben-Shachar, M., Patil, I., & Lüdecke, D. (2020). Methods and algorithms
433 for correlation analysis in R. *Journal of Open Source Software*, 5(51), 2306.
434 <https://doi.org/10.21105/joss.02306>
- 435 Makowski, D., Lau, Z. J., Pham, T., Paul Boyce, W., & Annabel Chen, S. H. (2021). A
436 Parametric Framework to Generate Visual Illusions Using Python. *Perception*, 50(11),
437 950–965. <https://doi.org/10.1177/03010066211057347>
- 438 Matzke, D., & Wagenmakers, E.-J. (2009). Psychological interpretation of the ex-gaussian

- 439 and shifted wald parameters: A diffusion model analysis. *Psychonomic Bulletin &
440 Review*, 16(5), 798–817.
- 441 Mylniec, A., & Bednarek, H. (2016). Field dependence, efficiency of information processing
442 in working memory and susceptibility to orientation illusions among architects. *Polish
443 Psychological Bulletin*, 47(1), 112–122. <https://doi.org/10.1515/ppb-2016-0012>
- 444 Notredame, C.-E., Pins, D., Deneve, S., & Jardri, R. (2014). What visual illusions teach us
445 about schizophrenia. *Frontiers in Integrative Neuroscience*, 8, 63.
446 <https://doi.org/10.3389/fnint.2014.00063>
- 447 Ohmann, K., & Burgmer, P. (2016). Nothing compares to me: How narcissism shapes
448 comparative thinking. *Personality and Individual Differences*, 98, 162–170.
449 <https://doi.org/10.1016/j.paid.2016.03.069>
- 450 Park, S., Zikopoulos, B., & Yazdanbakhsh, A. (2022). Visual illusion susceptibility in
451 autism: A neural model. *European Journal of Neuroscience*, 56.
452 <https://doi.org/10.1111/ejn.15739>
- 453 Peer, E., Rothschild, D., Gordon, A., Evernden, Z., & Damer, E. (2022). Data quality of
454 platforms and panels for online behavioral research. *Behavior Research Methods*, 54(4),
455 1643–1662. <https://doi.org/10.3758/s13428-021-01694-3>
- 456 Pessoa, V. F., Monge-Fuentes, V., Simon, C. Y., Suganuma, E., & Tavares, M. C. H.
457 (2008). The müller-lyer illusion as a tool for schizophrenia screening. *Reviews in the
458 Neurosciences*, 19(2-3). <https://doi.org/10.1515/REVNEURO.2008.19.2-3.91>
- 459 R Core Team. (2022). *R: A language and environment for statistical computing*. R
460 Foundation for Statistical Computing. <https://www.R-project.org/>
- 461 Razeghi, R., Arsham, S., Movahedi, A., & Sammaknejad, N. (2022). The effect of visual
462 illusion on performance and quiet eye in autistic children. *Early Child Development and
463 Care*, 192(5), 807–815. <https://doi.org/10.1080/03004430.2020.1802260>
- 464 Sibley, C., Luyten, N., Wolfman, M., Mobberley, A., Wootton, L. W., Hammond, M.,
465 Sengupta, N., Perry, R., West-Newman, T., Wilson, M., McLellan, L., Hoverd, W. J., &

- 466 Robertson, A. (2011). The mini-IPIP6: Validation and extension of a short measure of
467 the big-six factors of personality in new zealand. *New Zealand Journal of Psychology*,
468 *40*, 142–159.
- 469 Teufel, C., Dakin, S. C., & Fletcher, P. C. (2018). Prior object-knowledge sharpens
470 properties of early visual feature-detectors. *Scientific Reports*, *8*(1), 10853.
471 <https://doi.org/10.1038/s41598-018-28845-5>
- 472 Teufel, C., Subramaniam, N., Dobler, V., Perez, J., Finnemann, J., Mehta, P. R., Goodyer,
473 I. M., & Fletcher, P. C. (2015). Shift toward prior knowledge confers a perceptual
474 advantage in early psychosis and psychosis-prone healthy individuals. *Proceedings of the*
475 *National Academy of Sciences*, *112*(43), 13401–13406.
476 <https://doi.org/10.1073/pnas.1503916112>
- 477 Wickham, H., Averick, M., Bryan, J., Chang, W., McGowan, L., François, R., Grolemund,
478 Hayes, A., Henry, L., Hester, J., Kuhn, M., Pedersen, T., Miller, E., Bache, S.,
479 Müller, K., Ooms, J., Robinson, D., Seidel, D., Spinu, V., ... Yutani, H. (2019).
480 Welcome to the tidyverse. *Journal of Open Source Software*, *4*(43), 1686.
481 <https://doi.org/10.21105/joss.01686>
- 482 Yang, E., Tadin, D., Glasser, D. M., Hong, S. W., Blake, R., & Park, S. (2012). Visual
483 Context Processing in Schizophrenia: *Clinical Psychological Science*.
484 <https://doi.org/10.1177/2167702612464618>
- 485 Zhang, Y., Liu, J., Wang, Y., Huang, J., Wei, L., Zhang, B., Wang, W., & Chen, W.
486 (2017). Personality traits and perception of Müller-Lyer illusion in male Chinese
487 military soldiers and university students. *Translational Neuroscience*, *8*(1), 15–20.
488 <https://doi.org/10.1515/tnsci-2017-0004>