

**The Illusion Game: A Novel Experimental Paradigm Provides Evidence in
Favour of a General Factor of Visual Illusion Sensitivity and Personality
Correlates**

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24

Abstract

25 Visual illusions strikingly highlight how the brain uses contextual and prior information to
26 inform our perception of reality. Unfortunately, illusion research has been hampered by the
27 difficulty of adapting these stimuli to experimental settings, which ideally require a
28 controlled and gradual modulation of the effects of interest. In this set of studies, we used
29 the parametric framework for visual illusions implemented in the *Pyllusion* software to
30 generate 10 different classic illusions (Delboeuf, Ebbinghaus, Rod and Frame,
31 Vertical-Horizontal, Zöllner, White, Müller-Lyer, Ponzo, Poggendorff, Contrast) varying in
32 strength. We tested the objective effect of the illusions on errors and reaction times in a
33 perceptual discrimination task, from which we extracted participant-level performance
34 scores ($n=250$). Our results provide evidence in favour of the existence of a general factor
35 (labelled Factor *i*) underlying the sensitivity to different illusions. Moreover, we report a
36 positive relationship between illusion sensitivity and personality traits such as
37 Agreeableness, Honesty-Humility, and negative relationships with Psychoticism,
38 Antagonism, Disinhibition, and Negative Affect. All the materials are available in
39 open-access (<https://github.com/RealityBending/IllusionGameValidation>). We invite
40 researchers to re-analyze the data using alternative approaches to provide complimentary
41 findings on the effect, structure and correlates, of illusion sensitivity.

42 *Keywords:* visual illusions, illusion game, Pyllusion, personality, general factor

43 Word count: 1156

44 **The Illusion Game: A Novel Experimental Paradigm Provides Evidence in**
45 **Favour of a General Factor of Visual Illusion Sensitivity and Personality**
46 **Correlates**

47 **Introduction**

48 Visual illusions are fascinating stimuli capturing a key feature of our neurocognitive
49 systems. They eloquently show that our brains did not evolve to be perfect perceptual
50 devices providing veridical accounts of physical reality, but integrate prior knowledge and
51 contextual information - blended together in our subjective conscious experience (Carbon,
52 2014). Despite the historical and intensive interest within the fields of visual perception
53 (Day, 1972; Eagleman, 2001; Gomez-Villa et al., 2022), consciousness science (Caporuscio
54 et al., 2022; Lamme, 2020), and psychiatry (Gori et al., 2016; Notredame et al., 2014;
55 Razeghi et al., 2022; Teufel et al., 2015), several important issues remain open.

56 Notably, the presence of a common mechanism underlying the effect of different
57 illusions has been contested (Cretenoud, Francis, et al., 2020; Cretenoud et al., 2019a;
58 Hamburger, 2016); and the nature of the underlying processes - whether related to
59 low-level features of the visual processing system (Cretenoud et al., 2019b; Gori et al.,
60 2016) or to top-down influences of prior beliefs (Caporuscio et al., 2022; Teufel et al., 2018)
61 are strongly debated. The existence of dispositional correlates of illusion sensitivity - for
62 example, higher illusion resistance has been reported in schizophrenia and autism (Giaouri
63 & Alevriadou, 2011; Keane et al., 2014; Notredame et al., 2014; Park et al., 2022; Pessoa et
64 al., 2008), as well as in individuals with stronger aggression and narcissism traits (Konrath
65 et al., 2009a; Zhang et al., 2017) - is another area of controversy.

66 One key challenge hindering the further development of illusion research is the
67 relative difficulty in adapting visual illusions to an experimental setting, which typically
68 requires the controlled modulation of the specific variables of interest. To address this
69 issue, we first developed a parametric framework to manipulate visual illusions, which we

⁷⁰ implemented and made accessible in the open-source software *Pyllusion* (Makowski et al.,
⁷¹ 2021). This software allows us to generate different types of classic visual illusions (e.g.,
⁷² Müller-Lyer, Ponzo, Delboeuf, Ebbinghaus, . . .) with a continuous and independent
⁷³ modulation of two parameters: *illusion strength* and *task difficulty* (see **Figure 1**).

Parametric Framework for Visual Illusions

Example with the Müller-Lyer Illusion



The Müller-Lyer Illusion is traditionally presented as two segments (the **red targets**), which perception is biased by the **context** (the arrows). Here, the lower segment appears longer despite being of the same length.

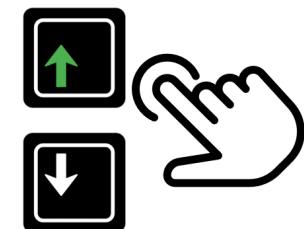


In this illusion, the **task difficulty** corresponds to the difference between the lengths of the red target segments, and the **illusion strength** corresponds to the angle of the arrows.

Example of Stimuli



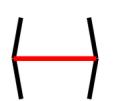
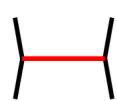
- ✓ Task difficulty: **easy**
(upper line is 2 times longer)
- ✓ Illusion Strength: **strong**
(angle is sharp)
- ✓ Illusion Direction (left): **incongruent**
(the illusion makes the task harder)
- ✓ Illusion Direction (right): **congruent**
(the illusion makes the task easier)



- ✓ Task difficulty: **hard**
(upper line is only 1.1 times longer)
- ✓ Illusion Strength: **weak**
(angle is flat)
- ✓ Illusion Direction (left): **incongruent**
(the illusion makes the task harder)
- ✓ Illusion Direction (right): **congruent**
(the illusion makes the task easier)



Task: For these stimuli, the correct response is always the « up » arrow, indicating the longer red segment. We measured the reaction time and the errors (in this case, the « down » arrow).



Stimuli created with the open-source software *Pyllusion* (Makowski et al., 2021)

Figure 1. The parametric framework for visual illusions (Makowski et al., 2021) applied to the Müller-Lyer illusion (above). Below are examples of stimuli showcasing the manipulation of two parameters, task difficulty and illusion strength.

⁷⁴ Indeed, many visual illusions can be seen as being composed of *targets* (e.g.,

75 same-length lines), of which perception is biased by the *context* (e.g., in the Müller-Lyer
76 illusion, the same-length line segments appear to have different lengths when they end with
77 inwards or outwards pointing arrows). Past illusion studies traditionally employed
78 paradigms focusing on participants' subjective experience, by asking them to what extent
79 they perceive two identical targets as different (Lányi et al., 2022), or having them adjust
80 the targets to a reference stimulus relying only on their perception (Grzeczkowski et al.,
81 2018; Mylniec & Bednarek, 2016). Alternatively, *Pyllusion* allows the creation of illusions
82 in which the targets are objectively different (e.g., one segment is truly more or less longer
83 than the other), and in which the illusion varies in strength (the biasing angle of the arrows
84 is more or less acute).

85 This opens the door for an experimental task in which participants make perceptual
86 judgments about the targets (e.g., which segment is the longest) under different conditions
87 of objective difficulty and illusion strength. Moreover, the illusion effect can be either
88 “incongruent” (making the task more difficult by biasing the perception in the opposite
89 way) or “congruent” (making the task easier). Although visual illusions are inherently tied
90 to subjective perception, this framework allows a reversal of the traditional paradigm to
91 potentially quantify the “objective” effect of illusions by measuring its behavioral effect
92 (error rate and reaction times) on the performance in a perceptual task.

93 In the present set of preregistered studies, we will first test this novel paradigm by
94 investigating if the effect of illusion and task difficulty can be manipulated continuously,
95 and separately modeled statistically. Then, we will further utilize the paradigm to assess
96 whether 10 different classic illusions (Delboeuf, Ebbinghaus, Rod and Frame,
97 Vertical-Horizontal, Zöllner, White, Müller-Lyer, Ponzo, Poggendorff, Contrast) share a
98 common latent factor. Finally, we will investigate how the the inter-individual sensitivity
99 to illusions relates to dispositional variables, such as demographic characteristics and
100 personality.

101 In line with open-science standards, all the material (stimuli generation code,
102 experiment code, raw data, analysis script with complementary figures and analyses,
103 preregistration, etc.) is available at
104 <https://github.com/RealityBending/IllusionGameValidation>.

105 **Study 1**

106 **Aim**

107 Study 1 can be seen as a pilot experiment aiming to gather some preliminary data to
108 assess if the stimuli generated by *Pyllusion* behaves as expected for each of the 10 illusion
109 types (i.e., whether an increase of task difficulty and illusion strength leads to an increase
110 of errors); and develop an intuition about the magnitude of effects, to refine the stimuli
111 parameters to a more sensible range (i.e., not overly easy and not impossibly hard) for the
112 next study.

113 **Procedure**

114 We generated 56 stimuli for each of the 10 illusion types. These stimuli resulted from
115 the combination of 8 linearly-spread levels of task difficulty (e.g., [1, 2, 3, 4, 5, 6, 7], where
116 1 corresponds to the highest difficulty - i.e., the smallest objective difference between
117 targets) and 7 levels of illusion strength (3 values of strength on the congruent side, 3 on
118 the incongruent side, and 0; e.g., [-3, -2, -1, 0, 1, 2, 3], where negative values correspond to
119 congruent illusion strengths).

120 The 10 illusion blocks were randomly presented, and the order of the 56 stimuli
121 within the blocks was also randomized. After the first series of 10 blocks, another series
122 was done (with new randomized order of blocks and trials). In total, each participant saw
123 56 different trials per 10 illusion type, repeated 2 times (total = 1120 trials), to which they
124 had to respond “as fast as possible without making errors” (i.e., an explicit double
125 constraint to mitigate the inter-individual variability in the speed-accuracy trade off). The

¹²⁶ task was implemented using *jsPsych* (De Leeuw, 2015). The instructions for each illusion
¹²⁷ type are available in the experiment code.

¹²⁸ Participants

¹²⁹ Fifty-two participants were recruited via *Prolific* (www.prolificacademic.co.uk), a
¹³⁰ crowd-sourcing platform providing high data quality (Peer et al., 2022). The only inclusion
¹³¹ criterion was a fluent proficiency in English to ensure that the task instructions would be
¹³² well-understood. Participants were incentivised with a reward of about £7.5 for completing
¹³³ the task, which took about 50 minutes to finish.

¹³⁴ We removed 6 participants upon inspection of the average error rage (when close to
¹³⁵ 50%, suggesting random answers), and when the reaction time distribution was implausibly
¹³⁶ fast. For the remaining participants, we discarded blocks where the error rate was higher
¹³⁷ than 50% (possibly indicating that instructions got misunderstood; e.g., participants were
¹³⁸ selecting the shorter line instead of the longer one). Finally, we removed 692 (1.37%) trials
¹³⁹ based on an implausibly short or long response time (< 150 ms or > 3000 ms).

¹⁴⁰ The final sample included 46 participants (Mean age = 26.7, SD = 7.7, range: [19,
¹⁴¹ 60]; Sex: 39.1% females, 56.5% males).

¹⁴² Data Analysis

¹⁴³ The analysis of study 1 focused on the probability of errors as the main outcome
¹⁴⁴ variable. For each illusion, we started by visualizing the average effect of task difficulty and
¹⁴⁵ illusion strength to gain some intuition on the underlying generative model. Next, we
¹⁴⁶ tested the performance of various logistic models differing in their specifications, such as:
¹⁴⁷ with or without a transformation of the task difficulty (log, square root or cubic root), with
¹⁴⁸ or without a 2nd order polynomial term for the illusion strength, and with or without the
¹⁴⁹ illusion side (up *vs.* down or left *vs.* right) as an additional predictor. We then fitted the
¹⁵⁰ best performing model under a Bayesian framework, and compared its visualization with

151 that of a General Additive Model (GAM), which has an increased ability of mapping
152 underlying potential non-linear relationships (at the expense of model simplicity).

153 The analysis was carried out using *R 4.2* (R Core Team, 2022), *brms* (Bürkner, 2017),
154 the *tidyverse* (Wickham et al., 2019), and the *easystats* collection of packages (Lüdecke et
155 al., 2021, 2019; Makowski et al., 2020; Makowski, Ben-Shachar, & Lüdecke, 2019).

156 Results

157 The statistical models suggested that the effect of task difficulty had a cubic
158 relationship with error rate for the Delboeuf and Ebbinghaus illusions (both composed of
159 circular shapes), square relationship for the Rod and Frame and Vertical-Horizontal
160 illusions, cubic relationship for the Zöllner and Poggendorff illusions, exponential
161 relationship for the White illusion, cubic relationship for the Müller-Lyer and Ponzo
162 illusions (both based on line lengths), and linear relationship for the Contrast illusion. All
163 models suggested a significant effect of illusion strength and task difficulty. See details and
164 figures in the analysis script.

165 Discussion

166 This study provided a clearer understanding of the magnitude of the parametric
167 effects at stake and the type of interaction between them. Furthermore, it allowed us to
168 better understand and test the stimuli generated by *Pyillusion*, as well as uncover technical
169 bugs and issues (for instance, the specification direction of the illusion strength was
170 reversed for a few illusions), which were fixed by a new software release. Crucially, this
171 study allowed us to refine the range of task difficulty and illusion strength values in order
172 to maximize information gain.

173 In most illusions, the task difficulty exhibited monotonic power-law scaled effects,
174 which is in line with the psychophysics literature on perceptual decisions (Bogacz et al.,
175 2006; Ditzinger, 2010; Shekhar & Rahnev, 2021). One notable result was the illusion effect

176 pattern for the Zöllner illusion, which suggested a non-linear relationship. By generating a
177 wider range of illusion strength values, the next study will attempt at clarifying this point.

178 **Study 2**

179 **Aim**

180 The aim of study 2 was two-fold. In the first part, we carefully modeled the error rate
181 and the reaction time of each illusion type in order to validate our novel paradigm and
182 show that the effect of illusions can be manipulated continuously. In the second part, we
183 derived the participant-level scores from the models (i.e., the effect of illusion strength for
184 each individual) and analyzed their latent factors structure.

185 **Procedure**

186 The paradigm of study 2 was similar to that of study 1, with the following changes:
187 the illusory stimuli were re-generated within a refined space of parameters based on the
188 results of study 1. Moreover, taking into account the findings of study 1, we used
189 non-linearly spaced difficulty levels, depending on the best underlying model (i.e., with an
190 exponential, square or cubic spacing depending on the relationship). For instance, a linear
191 space of [0.1, 0.4, 0.7, 1.0] can be transformed to an exponential space of [0.1, 0.34, 0.64,
192 1.0].

193 Additionally, instead of repeating each stimulus two times, we generated illusions
194 using more levels of difficulty and illusion strength. As such, for each illusion type, we
195 generated a total of 134 stimuli that were split into two groups (67 stimuli per illusion
196 block). Furthermore, instead of a simple break screen, we added two personality
197 questionnaires between the two series of 10 illusion blocks (see study 3).

198 Participants

199 Using the same recruitment procedure as in study 1, we recruited 256 participants,
200 out of which 6 were identified as outliers and excluded, leaving a final sample of 250
201 participants (Mean age = 26.5, SD = 7.6, range: [18, 69]; Sex: 48% females, 52% males).
202 Please see study 3 for the full demographic breakdown. We discarded blocks with more
203 than 50% of errors (2.16% of trials) and 0.76% trials with extreme response times (< 125
204 ms or > 4 SD above mean).

205 Data Analysis

206 The first part of the analysis focused on modelling the effect of illusion strength and
207 task difficulty on errors and reaction time (RT), within each illusion. In order to achieve
208 that, we started by fitting General Additive Models (GAMs), which can accommodate
209 possible non-linear effects and interactions. Errors were analyzed using Bayesian logistic
210 mixed models, and RTs of correct responses were analyzed using an ex-Gaussian family
211 with the same fixed effects entered for the location μ (mean), scale σ (spread) and
212 tail-dominance τ of the RT distribution (Balota & Yap, 2011; Matzke & Wagenmakers,
213 2009).

214 Using GAMs as the “ground-truth” models, we attempted at approximating them
215 using general linear models, which have the advantage of estimating the participant-level
216 variability of the effects (via random slopes). Following a comparison of models with a
217 combination of transformations (raw, log, square root or cubic root) on the main predictors
218 (task *difficulty* and illusion *strength*), we selected and fitted the best model (best on their
219 indices of fit), and compared their output visually (see **Figure 2**).

220 We then extracted the inter-individual variability in the effect of illusion strength and
221 its interaction with task difficulty, and used it as participant-level scores. Finally, We
222 explored the relationship of these indices across different illusions using exploratory factor

223 analysis (EFA) and structural equation modelling (SEM).

224 **Results**

225 The best models were $\log(\text{diff}) * \text{strength}$ for Delboeuf; $\sqrt{\text{diff}} * \text{strength}$ for
226 Ebbinghaus; $\log(\text{diff}) * \log(\text{strength})$ for Rod and Frame; $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ for
227 Vertical-Horizontal; $\text{cbrt}(\text{diff}) * \text{strength}$ for Zöllner; $\text{diff} * \sqrt{\text{strength}}$ and
228 $\log(\text{diff}) * \text{strength}$ respectively for errors and RT in White; $\sqrt{\text{diff}} * \sqrt{\text{strength}}$
229 and $\sqrt{\text{diff}} * \text{strength}$ respectively for errors and RT in Müller-Lyer;
230 $\text{cbrt}(\text{diff}) * \text{strength}$ for Ponzo; $\text{cbrt}(\text{diff}) * \sqrt{\text{strength}}$ and $\text{cbrt}(\text{diff}) * \text{strength}$
231 respectively for errors and RT in Poggendorff; $\sqrt{\text{diff}} * \sqrt{\text{strength}}$ for Contrast. In
232 all of these models, the effects of illusion strength, task difficulty and their interaction were
233 significant.

234 For errors, most of the models closely matched their GAMs counterpart (see **Figure**
235 **2**), with the exception of Delboeuf (for which the GAM suggested a non-monotonic effect
236 of illusion strength with a local minimum at 0) and Zöllner (for which theoretically
237 congruent illusion effects were related to increased error rate).

238 For RTs, the GAMs suggested a consistent non-linear relationship between RT and
239 illusion strength: as the illusion strength increase beyond a certain threshold, the
240 participants respond faster. While this is not surprising (strong illusions are likely so
241 effective in biasing perception that it is “easier”, i.e., faster, to make the wrong decision),
242 the linear models were not designed to capture this - likely quadratic - pattern and hence
243 are not good representatives of the underlying dynamics. As such, we decided not to use
244 them for the individual scores analysis.

245 Though imperfect, we believe that the random-slope models capture inter-individual
246 differences with more accuracy (and are also more conservative estimates due shrinkage)
247 than basic empirical scores, such as the total number of errors, or the average RT. Thus, for

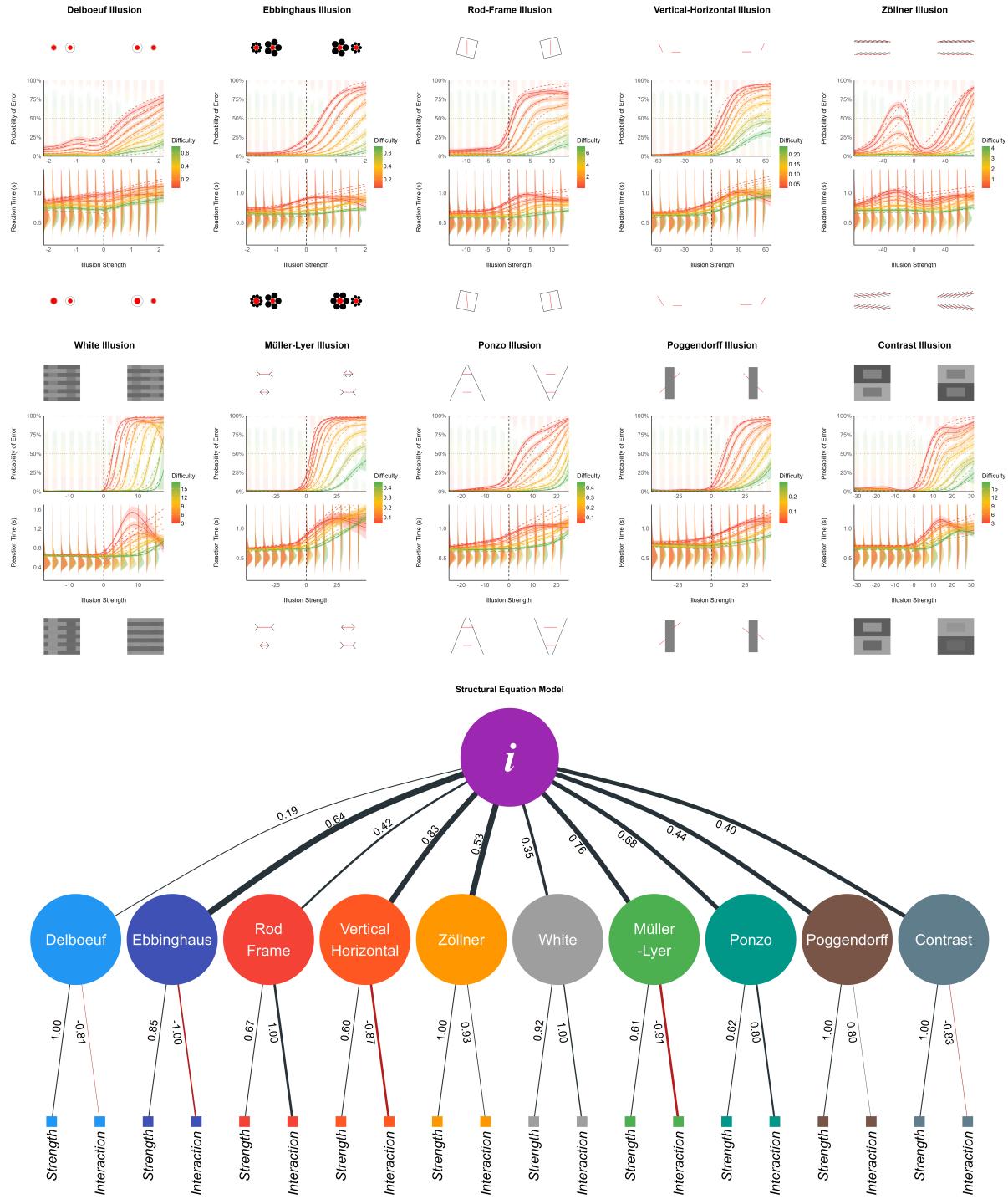


Figure 2. Top: the effect of illusion strength and task difficulty on the error rage and reaction time (RT) for each individual illusions. The solid line represent the General Additive Model (GAM), and the dashed line correspond to its approximation via linear models. Descriptive data is shown with stacked dots (errors are hanging from the top) and distributions for RTs. Negative values for illusion strength correspond to congruent (i.e., facilitating) illusion effects. Task difficulty (the objective difference between the targets of perceptual decision) levels are shown as colors, with lowest values corresponding to harder trials. Each illusion type is surrounded by 4 extreme examples of stimuli, corresponding to the hardest difficulty (on top) and the strongest illusion (on the right for incongruent illusions). Bottom: We extracted the effect slope of the illusion strength and its interaction with task difficulty for each participant. We fitted a Structural Equation Model (SEM) suggesting that these manifest variables group to first-level illusion-specific latent factors, which then load on a general factor of illusion sensitivity (Factor *i*).

248 each illusion and within each participant, we extracted the effect of illusion strength and its
249 interaction with task difficulty when the illusion effect was incongruent. These twenty
250 participant-level scores were subjected to exploratory factor analysis (EFA). The Method
251 Agreement Procedure (Lüdecke et al., 2020) suggested the presence of 7 latent factors. An
252 oblique (*oblimin* rotation) factor solution explaining 66.69% of variance suggested separate
253 dimensions for the effect of Zöllner, White, Poggendorff, Contrast, Ebbinghaus, Delboeuf,
254 and a common factor for the parameters related to Müller-Lyer, Vertical-Horizontal, Ponzo
255 and Rod and Frame. We submitted these factors to a second-level analysis and extracted
256 two orthogonal (*varimax* rotation) factors. The first factor was loaded by all the previous
257 dimensions with the exception of Delboeuf, which formed its own separate factor.

258 Finally, we tested this data-driven model (*m0*) against four other structural models
259 using structural equation modelling (SEM): one in which the two parameters of each of the
260 10 illusions (illusion strength and interaction with task difficulty) loaded on separate
261 factors, which then all loaded on a common factor (*m1*); one which the parameters were
262 grouped by illusion type (lines, circles, contrast and angle) before loading on a common
263 factor (*m2*); one in which all the parameters related to strength, and all the parameters
264 related to the interaction loaded onto two respective factors, which then loaded on a
265 common factor (*m3*); and one in which there was no intermediate level: all 20 parameters
266 loaded directly on a common factor (*m4*).

267 The model *m1*, in which the parameters loaded on a first level of 10 illusion-specific
268 factors, which then all loaded on a common factor significantly outperformed the other
269 models. Its indices of fit were ranging from acceptable to satisfactory (CFI = .92; SRMR =
270 .08; NNFI = .91; PNFI = .74; RMSEA = .08), and all the specified effects were significant.
271 The illusion-specific latent factors were loaded positively by the sensitivity to illusion
272 strength, and positively by the interaction effect with task difficulty (with the exception of
273 Delboeuf, Ebbinghaus, Vertical-Horizontal, Müller-Lyer and Contrast, for which the

loading was negative). The general factor of illusion sensitivity, labelled Factor i (i- for illusion), explained 48.02% of the total variance of the initial dataset, and was strongly related to Vertical-Horizontal ($\beta_{std.} = 0.83$), Müller-Lyer ($\beta_{std.} = 0.76$), Ponzo ($\beta_{std.} = 0.65$), Ebbinghaus ($\beta_{std.} = 0.64$); moderately to Zöllner ($\beta_{std.} = 0.53$), Poggendorff ($\beta_{std.} = 0.44$), Rod and Frame ($\beta_{std.} = 0.42$), Contrast ($\beta_{std.} = 0.40$) and White ($\beta_{std.} = 0.35$); and weakly to Delboeuf ($\beta_{std.} = 0.19$). We then computed, for each participant, its score for the 10 illusion-specific factors and for the general Factor i .

We have to keep in mind that these individual scores are the result of several layers of simplification: 1) the individual coefficient is that of simpler models that sometimes do not perfectly capture the underlying dynamics (especially in the case of Delboeuf and Zöllner); 2) we only used the models on error rate, which could be biased by the speed-accuracy decision criterion used by participants; 3) the structural equation model used to compute the scores also incorporated multiple levels of abstractions. Thus, in order to validate the individual scores, we computed the correlation between them and simple empirical scores, such as the average error rate and the mean RT in the task. This analysis revealed strong and significant correlations between each illusion-specific factor and the average amount of errors in its respective task. Moreover, each individual score was strongly associated with the average RT across multiple illusion types. This suggests that the individual scores obtained from the structural equation model do capture the sensitivity of each participant to visual illusions, manifesting in both the number of errors and high reaction times.

294 Discussion

This study confirmed that it was possible to continuously manipulate the effect of illusion strength for 10 classical illusions. Increasing the illusion strength increased the likelihood of errors, as well as the average and spread of RTs (but only up to a point, after which participants become faster at responding with the wrong answer). Future studies are needed to explore reaction times and try to identify the most appropriate models, and / or

300 use models that integrate errors and reaction time (e.g., drift diffusion models).

301 The effect on errors was monotonic for most illusions, with the exception of Delboeuf
302 and Zöllner. For both of them, mildly congruent illusion strengths (which theoretically
303 were supposed to be associated with less errors than incongruent effects) were related to
304 small and strong increases of errors, respectively. For the Delboeuf illusion, we believe that
305 this was due to an artifact caused by the illusion generation algorithm: the outline of the
306 target circles was always created as slightly bigger, which made the difference between
307 them more obvious at an illusion strength of 0. This was fixed in latest release of *Pyllusion*
308 (v1.2), which now generate outlines of the same size as the target circle. For the Zöllner
309 illusion, the observed non-monotonic pattern is actually consistent with previous reports
310 (Kitaoka, 2007; Kitaoka & Ishihara, 2000), suggesting an acute angle contraction effect at
311 very small as well as at sufficiently large angles (below 10 degrees for the former and
312 between 50 to 90 degrees for the latter) between the target horizontal line and the biasing
313 horizontal bars when the illusion strength is weak.

314 Finally, this study provided evidence for both the existence of illusion-specific factors,
315 as well as for a common latent factor (labelled Factor *i*) that explained about half of the
316 total variance. These participant-level scores were positively related to the error rate and
317 average reaction time, and can thus be interpreted as indices of illusion sensitivity.

318 Study 3

319 Aim

320 Study 3 aimed at investigating the links between the inter-individual scores of illusion
321 sensitivity (obtained in study 2), contextual variables (pertaining to the experiment
322 setting), such as screen size, demographic features (such as sex and age), and stable
323 dispositional variables such as “general” personality traits. Indeed, despite the abundant
324 literature on visual illusions, relatively few studies have investigated its ties with

325 participants' characteristics. Research examining the influence of demographic variables
326 such as gender and age have generally found inconsistent results (Cretenoud, Grzeczkowski,
327 et al., 2020; Grzeczkowski et al., 2017; Lo & Dinov, 2011; Papageorgiou et al., 2020).
328 Regarding links with personality, most works focused on traits associated with
329 psychopathology, such as impulsivity or sensation-seeking (Hlavatá et al., 2018; Lányi et
330 al., 2022; Pessoa et al., 2008; Zhang et al., 2017).

331 **Procedure**

332 This study was based on the data collected in study 2. The variables of interest here
333 were taken from the questionnaires that were inserted in between the two series of illusion
334 blocks. We used the *IPIP6* (24 items, Sibley et al., 2011) to measure 6 “normal”
335 personality traits (Extraversion, Openness, Conscientiousness, Agreeableness, Neuroticism
336 and Honesty-humility), and the *PID-5* (25 items, Hopwood et al., 2012) to measure
337 “pathological” personality traits (Disinhibition, Antagonism, Detachment, Negative Affect
338 and Psychoticism). The participants were the same as in study 2 (see **Figure 3**). However,
339 due to a technical issue, no personality data was recorded for the first eight participants.

340 **Data Analysis**

341 For each of the individual illusion sensitivity scores (10 illusion-specific factors and
342 the general Factor i), we tested the effect of contextual variables (screen size, screen refresh
343 rate), demographic variables (sex, education, age) and personality. As the supplementary
344 material contains the detailed results, we will here only report the significant results (based
345 on the Bayes Factor BF or the Probability of Direction pd , see Makowski, Ben-Shachar,
346 Chen, et al., 2019).

347 **Results**

348 The Bayesian correlation analysis (with narrow priors centered around a null effect)
349 between the illusion scores and contextual variables (screen size and refresh rate) provided

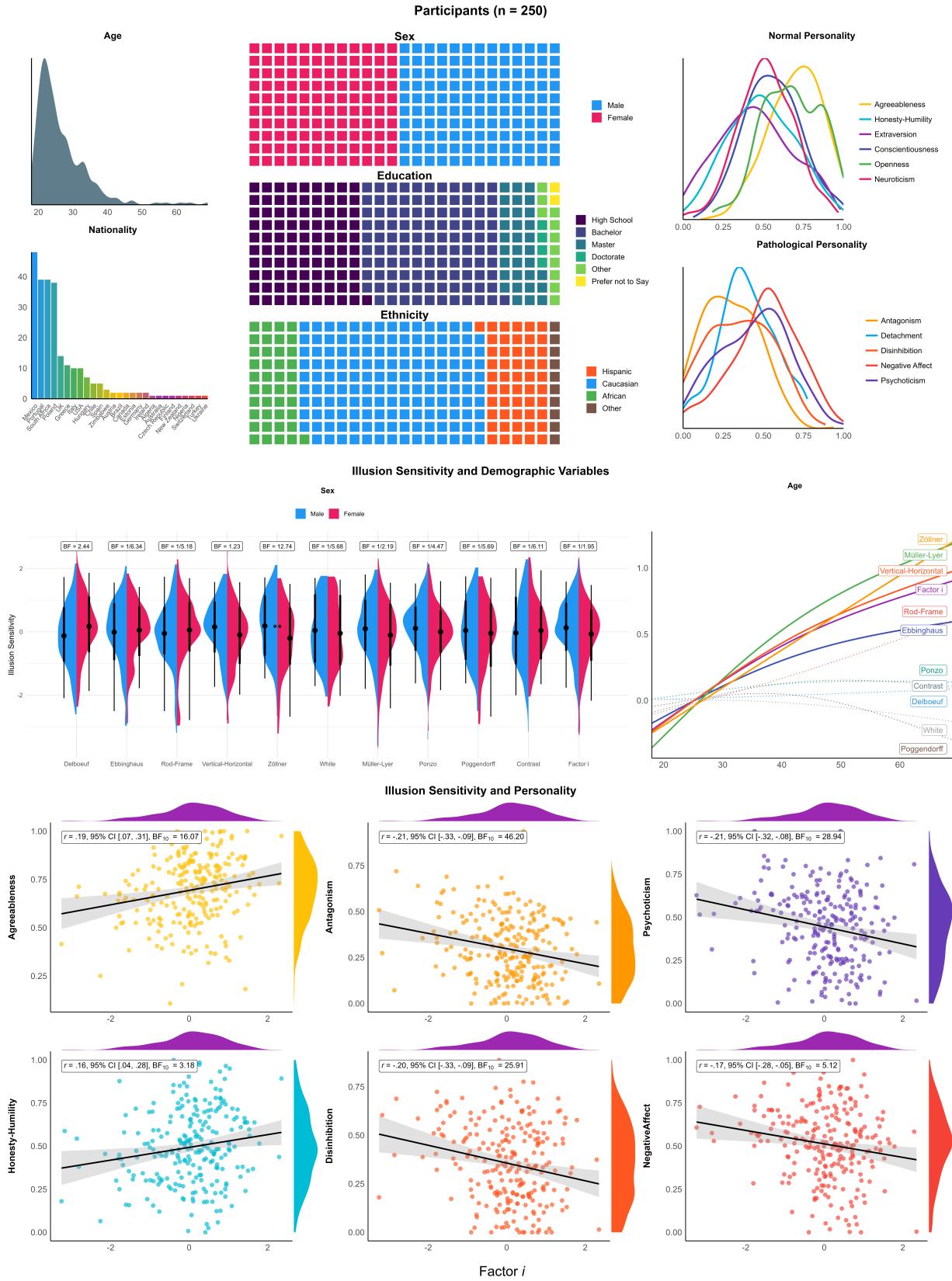


Figure 3. The upper plots show the distribution of demographic and dispositional variables. The middle plots show the relationship between illusion sensitivity scores, sex, and age (solid lines indicate significant relationships). Bottom plots show the correlation between the general factor of illusion sensitivity (Factor i), and personality traits.

350 weak evidence in favor of an absence of effect, with the exception of the two contrast-based
351 illusions. Anecdotal ($BF_{10} = 2.05$) and moderate evidence ($BF_{10} = 4.11$) was found for a
352 negative correlation between screen size and the sensitivity to the White and the Contrast
353 illusion, respectively. To test whether this result could be an artifact related to the highly
354 skewed screen size distribution (caused by very few participants with extreme screen sizes),
355 we re-ran a robust correlation (with rank-transformed values), which provided even
356 stronger evidence in favor of the effect existence ($BF_{10} = 28.19$, $BF_{10} = 4.31$ for White and
357 Contrast, respectively).

358 The Bayesian t-tests on the effect of sex suggested anecdotal to moderate evidence in
359 favour of the null effect for all scores, with the exception of the sensitivity to the Zöllner
360 illusion, which was higher in males as compared to females ($\Delta = -0.37$, 95% CI [-0.62,
361 -0.13], $BF_{10} = 12.74$). We fitted Bayesian linear models with the education level entered as
362 a monotonic predictor (appropriate for ordinal variables, Bürkner & Charpentier, 2020),
363 which yielded no significant effects. For age, we fitted two types of models for each score,
364 one general additive models (GAM) and a 2nd order polynomial model. These consistently
365 suggested a significant positive linear relationship between age and Factor i ($pd = 100\%$),
366 as well as the sensitivity to Müller-Lyer ($pd = 100\%$), Vertical-Horizontal ($pd = 100\%$),
367 Zöllner ($pd = 100\%$) and Ebbinghaus ($pd = 99\%$) illusions.

368 Regarding “normal” personality traits, Bayesian correlations suggested substantial
369 evidence in favor of a positive relationship between *Honesty-Humility* and Zöllner
370 ($BF_{10} > 100$), Vertical-Horizontal ($BF_{10} = 9.78$) and the Factor i ($BF_{10} = 4.00$); as well as
371 between *Agreeableness* and Vertical-Horizontal ($BF_{10} = 25.06$), Ponzo ($BF_{10} = 4.88$) and
372 the Factor i ($BF_{10} = 19.65$).

373 Regarding “pathological” personality traits, the results yielded strong evidence in
374 favor of a negative relationship between multiple illusion scores and multiple traits.
375 *Antagonism* was associated with the sensitivity to Vertical-Horizontal ($BF_{10} > 100$),

376 Müller-Lyer ($BF_{10} = 21.57$), Ponzo ($BF_{10} = 17.97$) illusions, and the Factor *i*
377 ($BF_{10} = 55.45$); *Psychoticism* was associated with the sensitivity to Vertical-Horizontal
378 ($BF_{10} = 66.63$) and Müller-Lyer ($BF_{10} = 35.59$) illusions, and the Factor *i* ($BF_{10} = 35.02$);
379 *Disinhibition* was associated with the sensitivity to Vertical-Horizontal ($BF_{10} = 25.38$),
380 Zöllner ($BF_{10} = 7.59$), Müller-Lyer ($BF_{10} = 5.89$) illusions, and the Factor *i*
381 ($BF_{10} = 31.42$); and *Negative Affect* was associated with Zöllner ($BF_{10} = 62.04$),
382 Vertical-Horizontal ($BF_{10} = 12.65$), Müller-Lyer ($BF_{10} = 3.17$), and the Factor *i*
383 ($BF_{10} = 6.39$). The last remaining trait, *Detachment*, did not share any relationship with
384 illusion sensitivity.

385 **Discussion**

386 We report significant links between inter-individual indices of illusion sensitivity and
387 varialbes related to experimental context, demographic characteristics and personality.
388 Firstly, screen size was found to have a significant negative relationship with the sensitivity
389 to the two contrast-based illusions, namely the White and Contrast illusions. One possible
390 explanation can be found in the mechanism by which visual systems filter through more
391 low spatial frequencies when the size of the target object is small (Dixon et al., 2014). As
392 this filtering process excludes illumination information from visual processing, smaller
393 screen sizes could yield artifactual changes in brightness perception, which in turn could
394 attenuate the illusory effect of luminance-related illusions.

395 Our results suggested an inconsistent pattern of non-significant sex-differences, with
396 the exception of greater sensitivity of males as compared to females reported to the Zöllner
397 illusion. As we do not consider this result as significant given its specificity, we note that
398 the existing literature reports, if any differences, that females exhibited greater illusion
399 sensitivity (Lo & Dinov, 2011; Miller, 2001; Papageorgiou et al., 2020). This inconsistency
400 could be due to past studies using a measure of illusion sensitivity that conflates the effect
401 of illusions *per se* with the perceptual abilities involved in the task, for which

402 gender-related differences can be found (in fact, the authors mention sex-differences in
403 visuospatial strategies as the potential mechanism underlying their findings). On the
404 contrary, the perceptual difficulty of the task and the illusion effect was independently
405 modulated in our paradigm, and statistically dissociated. Our scores of illusion sensitivity
406 might thus be less loaded with perceptual skills, thereby mitigating its effect.

407 Our findings also suggested a positive relationship between illusion sensitivity and
408 two “normal” personality traits, namely *Honesty-Humility* and *Agreeableness*, and a
409 negative link with *Antagonism*. Although the past literature regarding the links between
410 illusion sensitivity and personality traits remain scarce, convergent evidence can be found
411 in studies reporting a negative relationship between illusion sensitivity and hostility,
412 aggression and narcissism [Zhang et al. (2017); Konrath et al. (2009a); pessoa2008]. While
413 this result’s interpretation is challenging, one possible explanation could be drawn from the
414 literature on the cognitive style known as field dependence. Since narcissism and aggression
415 tendencies are correlated with lower field dependence (D’Amour et al., 2021; Ohmann &
416 Burgmer, 2016; i.e., a lesser reliance on external cues in ambiguous contexts, Witkin &
417 Goodenough, 1976), opposite traits, such as *Honesty-Humility* and *Agreeableness*, could
418 conversely be more biased by contextual cues and thus more sensitive to illusions.

419 The positive relationship between illusion sensitivity and “positive” personality traits
420 is mirrored by a negative relationship with several other pathological traits, including
421 *Psychoticism*, *Disinhibition*, and *Negative Affect*. These results are, in general, consistent
422 with past findings and theories, suggesting a negative relationship between egocentric
423 cognitive styles and context processing (including illusion sensitivity, Konrath et al.,
424 2009b). For instance, pathological egocentric beliefs (often observed alongside
425 *Psychoticism*, Fox, 2006) have been related to reduced context integration (Fox, 2006;
426 Konrath et al., 2009b; manifesting for instance in a tendency to separate objects from their
427 surroundings when processing visual stimuli, Ohmann & Burgmer, 2016). As such, it is

428 possible to relate this higher resistance to illusions to a self-centered, decontextualized and
429 disorganized information processing style, which can be found across the aforementioned
430 maladaptive personality traits [REF].

431 Furthermore, these results in favour of a link between illusion sensitivity and
432 maladaptive personality traits in a non-clinical population could be put in relation with
433 clinical findings, which could be seen as extreme cases where the relationship with illusion
434 sensitivity is the most manifest. In line with our results (in particular on *Psychoticism* and
435 *Disinhibition*), prior research has found greater illusion resistance in schizophrenia
436 (Grzeczkowski et al., 2018; Notredame et al., 2014; Pessoa et al., 2008), and in particular,
437 in association with schizotypal traits, such as cognitive disorganization (Cretenoud et al.,
438 2019b; Lányi et al., 2022).

439 **General Discussion**

440 The parametric illusion generation framework developed in Makowski et al. (2021)
441 proposes to conceptualize illusions as made of targets and distractors, both of which can be
442 manipulated independently and continuously. In the present study, we have shown that
443 such gradual modulation of illusion strength is effectively possible across 10 different types
444 of classic visual illusions. This important methodological step opens the door for new
445 illusions-based paradigms and tasks, to study the effect of illusions under different
446 conditions and to quantify illusion sensitivity using objective behavioral outcomes, such as
447 accuracy or speed.

448 The participants' sensitivities to 10 different types of visual illusions shared a
449 common part of variance, suggesting the existence of a general factor of illusion sensitivity
450 (Factor *i*). This result comes in a field of mixed findings. In fact, contrary to early studies
451 on visual illusions, more recent research have generally not found any significant evidence
452 for a common stable factor across illusions within individuals (Cretenoud, Francis, et al.,

453 2020; Cretenoud et al., 2019b; Grzeczkowski et al., 2017, 2018; Yang et al., 2012). Instead,
454 past findings suggests illusory effects are highly specific to the perceptual features of the
455 illusions at stake (Cretenoud et al., 2019b; Grzeczkowski et al., 2017). It is to note,
456 however, that most of these studies were low-powered and/or relied on conventional
457 paradigms, such as the adjustment procedure to measure the participants' subjective
458 perception. We believe that our study presents several methodological improvements,
459 including statistical power (high number of trials per participant), homogeneous stimuli
460 (with minimal and highly controlled features) and tasks (decision-making reaction-time
461 task), and more reliable participant-level score extraction method (based on random-factors
462 models), which in our opinion contributed to the emergence of the common factor.

463 However, although the illusions were relatively different in terms of the perceptual
464 task (contrast-based, size-estimation, angle-perception), The possibility of our general
465 factor being driven by inter-individual perceptual skills variability (or other cognitive skills)
466 cannot be discarded. Future studies should investigate the relationship between perceptual
467 abilities (in a similar task, but without illusions) and illusion sensitivity, and assess the
468 psychometric properties of similar paradigms, including stability (e.g., test-retest
469 reliability) and validity.

470 Finally, we found that the sensitivity to illusions was positively associated with
471 “positive” personality traits, such as *Agreeableness* and *Honesty-Humility*, and negatively
472 associated with maladaptive traits such as *Antagonism*, *Psychoticism*, *Disinhibition* and
473 *Negative Affect*. Beyond highlighting the relevance of illusions beyond the field of visual
474 perception, these results point towards an association with high-level general-domain
475 mechanisms. While the search for the exact mechanism(s) underlying these links is an
476 important goal of future research, our findings unlock the potential of illusion-based tasks
477 as sensitive tools to capture specific inter-individual neuro-cognitive differences.

478 As a conclusion, we strongly invite researchers to explore and re-analyze our dataset

⁴⁷⁹ with other approaches and methods to push the understanding of visual illusions and
⁴⁸⁰ illusion sensitivity further. The task, data and analysis script are available in open-access
⁴⁸¹ at <https://github.com/RealityBending/IllusionGameValidation>.

⁴⁸² **Acknowledgments**

⁴⁸³ We would like to thank Tam Pham and Zen J. Lau for their contribution to
⁴⁸⁴ *Pyillusion*, as well as Prof Dólos for the inspiration.

485

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