

## 0.1 Perspective of Probability

### 0.1.1 Abstract

In this issue, we will observe multivariate Gaussian distribution from the perspective of probability.

### 0.1.2 Prior Knowledge

$$\begin{aligned}x &\sim N(\mu, \sigma^2) \\ \mu &\in \mathcal{R}^p, \sigma \in \mathcal{R}^p \\ x_i &\sim N(\mu_i, \sigma_i) \\ p(x_i) &= \frac{1}{\sqrt{2\pi}\sigma_i} \exp\left(-\frac{(x_i - \mu_i)^2}{2\sigma_i^2}\right)\end{aligned}$$

### 0.1.3 Derivation

First, let's assume that each  $x_i$  is *iid*(*independent identically distribution*) as below:

$$\begin{aligned}p(x) &= \prod_{i=1}^p p(x_i) \\ &= \frac{1}{(2\pi)^{\frac{p}{2}} \prod_{i=1}^p \sigma_i} \exp\left(-\frac{1}{2} \sum_{i=1}^p \left(\frac{(x_i - \mu_i)^2}{\sigma_i^2}\right)\right) \\ &= \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} \begin{pmatrix} x_1 - \mu_1 & x_2 - \mu_2 & \dots & x_p - \mu_p \end{pmatrix} \begin{pmatrix} \frac{1}{\sigma_1^2} & 0 & \dots & 0 \\ \dots & \dots & \dots & \dots \\ 0 & \dots & 0 & \frac{1}{\sigma_p^2} \end{pmatrix} \begin{pmatrix} x_1 - \mu_1 \\ \dots \\ \dots \\ x_p - \mu_p \end{pmatrix}\right) \\ &= \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu)\right)\end{aligned}$$

The above is the probability density function of multivariate Gaussian distribution.

We know that  $\sigma$  is a positive semidefinite matrix, so we can perform singu-

lar value decomposition. So we have:

$$\begin{aligned}
\Sigma &= UVU^T \\
&= (u_1 \quad \dots \quad u_p) \begin{pmatrix} \lambda_1 & 0 & \dots & 0 \\ \dots & 0 & \dots & \dots \\ 0 & \dots & \dots & \lambda_p \end{pmatrix} \begin{pmatrix} u_1^T \\ \vdots \\ u_p^T \end{pmatrix} \\
&= (u_1 \lambda_1 \quad \dots \quad u_p \lambda_p) \begin{pmatrix} u_1^T \\ \vdots \\ u_p^T \end{pmatrix} \\
&= \sum_{i=1}^p u_i \lambda_i u_i^T
\end{aligned}$$

then

$$\begin{aligned}
\Sigma^{-1} &= (UVU^T)^{-1} \\
&= (U^T)^{-1} V^{-1} U^{-1} \\
&= UV^{-1}U^T \\
&= \sum_{i=1}^p u_i \frac{1}{\lambda_i} u_i^T
\end{aligned}$$

Let's set  $\Delta = (x - \mu)^T \Sigma^{-1} (x - \mu)$

Substitute the results derived above into:

$$\begin{aligned}
\Delta &= (x - \mu)^T \Sigma^{-1} (x - \mu) \\
&= (x - \mu)^T \sum_{i=1}^p u_i \frac{1}{\lambda_i} u_i^T (x - \mu) \\
&= \sum_{i=1}^p (x - \mu)^T u_i \frac{1}{\lambda_i} u_i^T (x - \mu)
\end{aligned}$$

Let's set  $y_i = (x - \mu)^T u_i$

Here,  $y_i$  represents the coordinate value of  $x$  projected onto the new orthogonal basis  $u_i$  after centralization.

so:

$$\Delta = \sum_{i=1}^p \frac{y_i^2}{\lambda_i}$$

Next, let's look at the probability density function of multivariate Gaussian distribution:

$$p(x) = \frac{1}{(2\pi)^{\frac{p}{2}} |\Sigma|^{\frac{1}{2}}} \exp\left(-\frac{1}{2} (x - \mu)^T \Sigma^{-1} (x - \mu)\right)$$

You can see that only the exponential part of the formula is related to the variable  $x$ . The previous factor is to make the probability sum 1. Therefore, the probability of Gaussian distribution is directly related to the value of  $\Delta$ .

We assume  $p = 2$ , then:

$$\frac{y_1^2}{\lambda_1} + \frac{y_2^2}{\lambda_2} = \Delta$$

We were surprised to find that this is very similar to the elliptic equation. The value of  $\Delta$  is not fixed, so for different  $x$ , these sample points form concentric ellipses in the plane. This is one of the properties of Gaussian distribution.